

### Theorem

## Fisher's Inequality

Barring trivial ~~expected~~ exceptions, the inequality  $b \geq 0$  holds for all connected, equireplicate, variance-balanced design.

Proof  $\gg$  For a connected, variance-balanced design

$$C = \theta \left( I_v - \frac{J_v}{v} \right)$$

The eigen values of  $C$  are  $\theta$  (with multiplicity  $v-1$ ) and  
0 ( " " " 1 )

Again, since the design is equireplicate,  $D_r = rI_v$  and hence the eigen values of  $D_r$  are  $r$  with multiplicity  $v$ .

Define  $p = D_r - C = N D_K^+ N'$

Thus, the eigenvalues of  $P$  are  $\lambda = 0$ , with multiplicity 2 and  $\lambda = 1$ , with multiplicity 1.

$$\therefore \text{rank}(p) = v$$

*James*

$P$  is singular when  $r=0$ , in this case  $\text{rank}(P)=1$  and hence  $\text{rank}(N)=1$

The columns of  $P$  and hence ~~rows~~<sup>those</sup> of  $N$  are spanned by the vector  $\mathbf{1}$  which is the eigen vector corresponding to the zero eigen value of  $C$ . Thus it follows that in case of  $r=0$  the rows of  $N$  are identical. If we exclude designs with incidence matrices having identical rows,

then we find that for any other equireplicate, variance-balanced design,  $P$  is non singular

$$\text{rank}(P) = v$$

$$v = \text{rank}(P) \leq \text{rank}(N) \leq b$$

$$\text{So, } v \leq b \text{ (Proved)}$$

## Recovery of Inter-block Information

While discussing the intra-block analysis of block design, it was stated that the block as well as the trt. effects are fixed. If the block effect is regarded as a random variable, the analysis is called inter block analysis or recovery of inter block information.

In the context of incomplete block design Yates noticed that since the allocated trt. to the incomplete blocks is made at random, it is reasonable to assume that the block effect themselves are random variables instead of fixed. If the experimental material is fairly heterogeneous, treating the block effects as fixed quantities results in the loss of information contained in the block designs.

Assume proper and binary design

$$y_{ij} = \mu + \tau_i + \beta_j + e_{ij} \quad \begin{pmatrix} i = 1(1)v \\ j = 1(1)b \end{pmatrix}$$

Assumptions ①  $e_{ij}$  s are i.i.d with  $E(e_{ij}) = 0$   
and  $V(e_{ij}) = \sigma^2 \neq 0$

②  $\beta_j$  are r.v. with

$$E(\beta_j) = 0, \quad V(\beta_j) = \sigma_b^2 \quad \forall j$$

$$\text{Cov}(\beta_j, \beta_{j'}) = 0 \quad \forall j \neq j'$$

③  $\beta_j$  s are uncorrelated with the error terms  $e_{ij}$  s.

we may regard the block totals as obs.

$$B_j = \sum_{i=1}^v y_{ij} = k\mu + \sum_{i=1}^v n_{ij} \tau_i + \left( k\beta_j + \sum_{i=1}^v e_{ij} \right), \quad j=1(1)b$$

$$\text{New error terms } d_j = k\beta_j + \sum_{i=1}^v e_{ij}, \quad j=1(1)b$$

$$E(d_j) = 0 \quad \forall j=1(1)b$$

$$\text{Var}(d_j) = k^2 \sigma_b^2 + k \sigma^2 = \sigma_d^2$$

Inter block estimates are obtained by minimizing the SS due to new errors

$$\begin{aligned} S &= \sum_{j=1}^b d_j^2 = \sum_{j=1}^b \left( B_j - k\mu - \sum_{i=1}^v n_{ij} \tau_i \right)^2 \\ &= \left( \underset{\sim}{B} - k\mu \underset{\sim}{1}_b - N' \underset{\sim}{\tau} \right)' \left( \underset{\sim}{B} - k\mu \underset{\sim}{1}_b - N' \underset{\sim}{\tau} \right) \end{aligned}$$

Normal equations

$$\frac{\partial S}{\partial \mu} = 0 \Rightarrow k \sum_{j=1}^b \left( B_j - k\mu - \sum_{i=1}^v n_{ij} \tau_i \right) = 0$$

$$\text{or, } bk\mu + \sum_{j=1}^b \sum_{i=1}^v n_{ij} \tau_i = \sum_{j=1}^b B_j$$

$$\text{or, } bk\mu + \sum_{j=1}^b \sum_{i=1}^v n_{ij} \tau_i = G$$

$$\text{or, } bk\mu + \underset{\sim}{1}'_b N' \underset{\sim}{\tau} = G \quad \text{--- (1)}$$

$$\frac{\partial \mathcal{L}}{\partial \tau_i} = 0, \quad i=1(1)v$$

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$$b_k \mu + \tilde{\tau}' \tilde{\tau} = G \quad [\tilde{L}'_v N' = \tilde{\tau}']$$

$$\text{or, } b_k \mu + \tilde{L}'_v D_r \tilde{\tau} = G \quad \text{--- (1)}$$

$$\frac{\partial \mathcal{L}}{\partial \tau_i} = 0 \Rightarrow \sum_{j=1}^b n_{ij} (B_j - k\mu - \sum_{i=1}^v n_{ij} \tau_i) = 0, \quad i=1(1)v$$

$$\text{or, } \sum_{j=1}^b n_{ij} + \sum_{j=1}^b n_{ij} \left( \sum_{i=1}^v n_{ij} \tau_i \right) = \sum_{j=1}^b n_{ij} B_j, \quad i=1(1)v$$

$$k\mu \begin{pmatrix} \sum_{j=1}^b n_{1j} \\ \sum_{j=1}^b n_{2j} \\ \vdots \\ \sum_{j=1}^b n_{vj} \end{pmatrix} + \begin{pmatrix} \sum_{j=1}^b n_{1j} \left( \sum_{i=1}^v n_{ij} \tau_i \right) \\ \sum_{j=1}^b n_{2j} \left( \sum_{i=1}^v n_{ij} \tau_i \right) \\ \vdots \\ \sum_{j=1}^b n_{vj} \left( \sum_{i=1}^v n_{ij} \tau_i \right) \end{pmatrix} = \begin{pmatrix} \sum_{j=1}^b n_{1j} B_j \\ \sum_{j=1}^b n_{2j} B_j \\ \vdots \\ \sum_{j=1}^b n_{vj} B_j \end{pmatrix}$$

$$\text{or, } k\mu \begin{pmatrix} \gamma_1 \\ \gamma_2 \\ \vdots \\ \gamma_v \end{pmatrix} + \begin{pmatrix} \tilde{n}'_1 (N' \tilde{\tau}) \\ \tilde{n}'_2 (N' \tilde{\tau}) \\ \vdots \\ \tilde{n}'_v (N' \tilde{\tau}) \end{pmatrix} = \begin{pmatrix} \tilde{n}'_1 B \\ \tilde{n}'_2 B \\ \vdots \\ \tilde{n}'_v B \end{pmatrix}$$

where  
 $\tilde{n}'_i = \begin{pmatrix} n_{i1} \\ n_{i2} \\ \vdots \\ n_{ib} \end{pmatrix}$   
 $i=1(1)v$   
 is the  $i$ th row of  $N$

$$\text{or, } k\mu \tilde{\gamma} + N N' \tilde{\tau} = N B$$

$$\text{or, } k\mu D_r \tilde{L}'_v + N N' \tilde{\tau} = N B \quad \text{--- (2)}$$

In matrix notation

$$\begin{pmatrix} b_k & \tilde{L}'_v D_r \\ k D_r \tilde{L}'_v & N N' \end{pmatrix} \begin{pmatrix} \mu \\ \tilde{\tau} \end{pmatrix} = \begin{pmatrix} G \\ N B \end{pmatrix}$$



Pre-multiply both sides by the non singular matrix —

$$\begin{pmatrix} 1 & \tilde{O}'_v \\ -\frac{1}{b} D_r \tilde{L}_v & I_v \end{pmatrix}_{(v+1) \times (v+1)}$$

$$\begin{pmatrix} 1 & \tilde{O}'_v \\ -\frac{1}{b} D_r \tilde{L}_v & I_v \end{pmatrix} \begin{pmatrix} bK & \tilde{L}'_v D_r \\ K D_r \tilde{L}_v & NN' \end{pmatrix} = \begin{pmatrix} 1 & \tilde{O}'_v \\ -\frac{1}{b} D_r \tilde{L}_v & I_v \end{pmatrix} \begin{pmatrix} G \\ NB_{\tilde{v}} \end{pmatrix}$$

$$\begin{pmatrix} bK & \tilde{L}'_v D_r \\ -K D_r \tilde{L}_v + K D_r \tilde{L}_v & -\frac{1}{b} D_r \tilde{L}_v \tilde{L}'_v D_r + NN' \end{pmatrix} \begin{pmatrix} \mu \\ \tilde{\tau} \end{pmatrix} = \begin{pmatrix} G \\ -\frac{G}{b} D_r \tilde{L}_v + NB_{\tilde{v}} \end{pmatrix}$$

$$\begin{pmatrix} bK & \tilde{L}'_v D_r \\ \tilde{O}_v & NN' - \frac{1}{b} D_r \tilde{L}_v \tilde{L}'_v D_r \end{pmatrix} \begin{pmatrix} \mu \\ \tilde{\tau} \end{pmatrix} = \begin{pmatrix} G \\ NB_{\tilde{v}} - \frac{G}{b} D_r \tilde{L}_v \end{pmatrix}$$

A soln<sup>2</sup> of the system of eqn<sup>s</sup> is obtained by taking the

condition  $\tilde{L}'_v D_r \tilde{\tau} = 0$

i.e.  $\sum_{i=1}^v \tau_i \tau_i = 0$

and assuming that  $NN'$  is non-singular:

we have,  $bK\mu + \tilde{L}'_v D_r \tilde{\tau} = G$  — (3)

$NN' \tilde{\tau} - \frac{1}{b} D_r \tilde{L}_v \tilde{L}'_v D_r \tilde{\tau} = NB_{\tilde{v}} - \frac{G}{b} D_r \tilde{L}_v$  — (4)

subject to the above condition

$bK\hat{\mu} = G$

$NN'\hat{\tilde{\tau}} = NB_{\tilde{v}} - \frac{G}{b} D_r \tilde{L}_v$

$$\text{or } \hat{\mu} = \frac{G}{bk}$$

$$\hat{\tau} = (NN')^{-1} \left( NB_{\sim} - \frac{G}{b} D_{\sim} 1_v \right)$$

$$= (NN')^{-1} \left( NB_{\sim} - \frac{G}{b} \tau_{\sim} \right)$$

$$= (NN')^{-1} \left( NB_{\sim} - \frac{G}{b} N' 1_b \right)$$

$$= (NN')^{-1} \left( NB_{\sim} - \frac{G}{bk} NK_{\sim} 1_b \right) \left[ \because \text{Design is proper} \cdot K_{\sim} 1_b = k \right]$$

$$= (NN')^{-1} \left( NB_{\sim} - \frac{G}{bk} NN' 1_v \right)$$

$$= (NN')^{-1} \left[ NB_{\sim} - \frac{G}{bk} 1_v \right]$$

We have two estimators of  $\tau_{\sim}$

$$\begin{aligned} \textcircled{1} \text{ Intra-Block estimator } & \hat{\tau}_{\sim} = C^{-1} Q \\ \textcircled{2} \text{ Inter-Block estimator } & \hat{\tau}_{\sim} = (NN')^{-1} NB_{\sim} - \frac{G}{bk} 1_v \end{aligned}$$

Let  $\psi = p' \tau_{\sim}$  be a contrast of trt. effects.

Intra-block estimate of  $\psi$  is  $\psi_1 = p' C^{-1} Q$

$$\begin{aligned} \text{With variance } V(\psi_1) &= p' C^{-1} \text{Disp}(Q) C^{-1} p \\ &= p' C^{-1} (C \sigma^2) C^{-1} p \\ &= \sigma^2 (p' C^{-1} p) \end{aligned}$$

Inter-block estimate of  $\psi$  is  $\psi_2 = p' \left[ (NN')^{-1} NB_{\sim} - \frac{G}{bk} 1_v \right]$  with variance

$$\begin{aligned} V(\psi_2) &= p' \text{Disp} \left( (NN')^{-1} NB_{\sim} \right) p = p' (NN')^{-1} NB_{\sim} \left[ \because p' 1_v = 0 \right] \\ &= p' (NN')^{-1} N \text{Disp}(B_{\sim}) N' (NN')^{-1} p \\ &= p' (NN')^{-1} N (\sigma^2 I_b) N' (NN')^{-1} p \left[ \because \text{Disp}(B_{\sim}) = \sigma^2 I_b \right] \\ &= \sigma^2 p' (NN')^{-1} p \end{aligned}$$

$\Psi_1$  and  $\Psi_2$  are uncorrelated

$$\Psi_1 = \underset{\sim}{p}' \underset{\sim}{c} \underset{\sim}{g}$$

$$\Psi_2 = \underset{\sim}{p}' (NN')^T \underset{\sim}{N} \underset{\sim}{B}$$

We'll show that  $\underset{\sim}{g}$  &  $\underset{\sim}{B}$  are uncorrelated

$$\underset{\sim}{g} = \underset{\sim}{T} - \underset{\sim}{N} \underset{\sim}{D} \underset{\sim}{K}^T \underset{\sim}{B}$$

$$\text{cov}(\underset{\sim}{g}, \underset{\sim}{B}) = \text{cov}(\underset{\sim}{T}, \underset{\sim}{B}) - \underset{\sim}{N} \underset{\sim}{D} \underset{\sim}{K}^T \text{Disp}(\underset{\sim}{B})$$

$$\begin{aligned} \text{cov}(\underset{\sim}{T}, \underset{\sim}{B}) &= \text{cov}(X' \underset{\sim}{Y}, X' \underset{\sim}{B}) \\ &= X' \underset{\sim}{C} \text{Disp}(\underset{\sim}{Y}) X \underset{\sim}{B} \end{aligned}$$

$$y_{ij} = \mu + \tau_i + \beta_j + e_{ij}$$

$$v(y_{ij}) = \sigma_b^2 + \sigma_e^2 \quad \forall i, j$$

$$\text{cov}(y_{ij}, y_{i'j'}) = 0$$

$$\begin{aligned} \text{cov}(y_{ij}, y_{i'j}) &= \text{cov}(\beta_j, \beta_j) \quad ; \quad i \neq i' \\ &= \sigma_b^2 \end{aligned}$$

$$\underset{\sim}{T}_p = \sum_{j=1}^b y_{ij}$$

$$= \sum_{j=1}^b (\mu + \tau_i + \beta_j + e_{ij})$$

$$\underset{\sim}{B}_j = k\mu + \sum_{i=1}^v n_{ij} \tau_i + k\beta_j + \sum_{i=1}^v e_{ij}$$

$$\text{cov}(T_i, \beta_j) = \text{cov}\left(\sum_{j=1}^b \beta_j + \sum_{j=1}^b e_{ij}, k\beta_j + \sum_{i=1}^v e_{ij}\right)$$

$$= n_{ij} k \text{cov}(\beta_j, \beta_j)$$

$$+ n_{ij} \text{cov}(e_{ij}, e_{ij})$$

$$= n_{ij} [k \text{var}(\beta_j) + \text{cov}(e_{ij}, e_{ij})]$$

$$= [k\sigma_b^2 + \sigma_e^2] n_{ij}$$

$$= \frac{n_{ij} \sigma_e^2}{k}$$

$$\begin{bmatrix} \sigma_b^2 & \sigma_e^2 \\ 0 & \sigma_e^2 \end{bmatrix} = k \begin{bmatrix} \sigma_b^2 & 0 \\ 0 & \sigma_e^2 \end{bmatrix}$$

$$\text{cov}(\tilde{I}, \tilde{B}) = \frac{N\sigma_d^2}{k}$$

Since the design is proper,

$$D_k = K I_b$$

$$\therefore D_k^{-1} = \frac{1}{K} I_b$$

$$\text{Again, Disp}(\tilde{B}) = \sigma_d^2 I$$

$$\text{So, } N D_k^{-1} \text{Disp}(\tilde{B})$$

$$= \frac{N}{K} \sigma_d^2$$

$$\therefore \text{cov}(\tilde{\theta}, \tilde{B}) = \mathbf{0}_{v \times b}$$

Hence,  $\psi_1$  and  $\psi_2$  are uncorrelated. If we want to combine these two estimators to obtain an estimator with the smallest variance then the combined estimator is obtained by taking a <sup>weighted</sup> ~~weighted~~ avg. of  $\psi_1$  and  $\psi_2$ , weights being the inverse of the variances of these two estimators. Thus the combined estimator is given by

$$\psi^* = \frac{\theta_1 \psi_1 + \theta_2 \psi_2}{\theta_1 + \theta_2}$$

$$\text{where, } \theta_1 = (\sigma^2 \tilde{p}' \tilde{C} \tilde{p})^{-1}$$

$$\theta_2 = (\sigma_d^2 \tilde{p}' (N N')^{-1} \tilde{p})^{-1}$$



# Mixed effects Model

Let  $y_1, y_2, \dots, y_n$  be  $n$  obs. such that

$$y_i = a_{i1}\beta_1 + a_{i2}\beta_2 + \dots + a_{ip}\beta_p + b_{i1}x_1 + b_{i2}x_2 + \dots + b_{iq}x_q \quad i=1(1)n$$

where,  $\beta_1, \beta_2, \dots, \beta_p$  are unknown parameters (fixed).

$x_1, x_2, \dots, x_q$  are random variables.

$a_{i1}, a_{i2}, \dots, a_{ip}, b_{i1}, b_{i2}, \dots, b_{iq}$  are known constants.

Further let,  $E(\epsilon_i) = 0$ ,  $\text{Var}(\epsilon_i) = \sigma^2 \quad \forall i=1(1)n$

$$\text{Cov}(\epsilon_i, \epsilon_{i'}) = 0 \quad \forall i \neq i'$$

$$\text{and } \text{Cov}(\epsilon_i, x_j) = 0$$

$$E(x_j) = 0, \quad \text{Var}(x_j) = \sigma_b^2 \quad \forall j=1(1)q$$

$$\forall i=1(1)n$$

$$\forall j=1(1)q$$

$$\text{Cov}(x_j, x_{j'}) = 0 \quad \forall j \neq j'$$

A model of this type is called linear mixed effects model.

In matrix notation,

$$\underset{n \times 1}{Y} = \underset{n \times p}{A} \underset{p \times 1}{\beta} + \underset{n \times q}{B} \underset{q \times 1}{x} + \underset{n \times 1}{\epsilon}$$

$$\text{where, } \underset{n}{Y} = \begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{pmatrix}; \quad \underset{p}{\beta} = \begin{pmatrix} \beta_1 \\ \beta_2 \\ \vdots \\ \beta_p \end{pmatrix}; \quad \underset{q}{x} = \begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_q \end{pmatrix}; \quad \underset{n}{\epsilon} = \begin{pmatrix} \epsilon_1 \\ \epsilon_2 \\ \vdots \\ \epsilon_n \end{pmatrix}$$

$$A = \begin{pmatrix} a_{11} & a_{12} & \dots & a_{1p} \\ a_{21} & a_{22} & \dots & a_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ a_{n1} & a_{n2} & \dots & a_{np} \end{pmatrix}; \quad B = \begin{pmatrix} b_{11} & b_{12} & \dots & b_{1q} \\ b_{21} & b_{22} & \dots & b_{2q} \\ \vdots & \vdots & \ddots & \vdots \\ b_{n1} & b_{n2} & \dots & b_{nq} \end{pmatrix}$$

$$E(\underset{n}{\epsilon}) = \underset{n}{0}$$

$$E(\underset{n}{Y}) = A\underset{p}{\beta}$$

$$\text{Disp}(\underset{n}{\epsilon}) = \sigma^2 I_n$$

$$\text{Disp}(\underset{n}{Y}) = B \text{Disp}(\underset{q}{x}) B' + \text{Disp}(\underset{n}{\epsilon})$$

$$E(\underset{q}{x}) = \underset{q}{0}$$

$$= \sigma_b^2 B B' + \sigma^2 I_n$$

$$\text{Disp}(\underset{q}{x}) = \sigma_b^2 I_q$$

$$= \Sigma (\text{say})$$

$$\text{Cov}(\underset{q}{x}, \underset{n}{\epsilon}) = 0$$

Normal eq  $A' \Sigma^{-1} A \beta = A' \Sigma^{-1} Y$

## Combined inter and intra block analysis

Assumption :- Design is proper.

$$K = K I_b$$

$$\& D_k = K I_b$$

Model :- 
$$\underset{n \times 1}{Y} = \underset{n \times 1}{\mu} I + \underset{n \times v}{X} \underset{v \times 1}{\tau} + \underset{n \times b}{X} \underset{b \times 1}{\beta} + \underset{n \times 1}{\epsilon}$$

Treatment effects  $\tau$  are fixed.

Block effects  $\beta$  are random.

Assumption  $E(\epsilon) = 0$ ,  $\text{Disp}(\epsilon) = \sigma^2 I_n$   
 $E(\beta) = 0$ ,  $\text{Disp}(\beta) = \sigma_b^2 I_b$   
 $\text{Cov}(\beta, \epsilon) = 0$

In standard notation,

$$\underset{n \times 1}{Y} = \left( \underset{n \times 1}{I} : \underset{n \times v}{X} \right) \begin{pmatrix} \mu \\ \tau \end{pmatrix} + \underset{n \times b}{X} \beta + \epsilon$$

$$= A \theta + B \beta + \epsilon$$

where  $A_{n \times (v+1)} = \left( I : X \right)$ ;  $B = X \beta$ ;  $\theta = \begin{pmatrix} \mu \\ \tau \end{pmatrix}$

The design is proper,

total no of obs.  $n = b \cdot k$

Assume that the  $n = bk$  obs. are such that the 1st set of  $k$  obs. are from block 1, the ~~the~~ 2nd set of  $k$  obs are from block 2 and finally the last set of  $k$  obs are ~~to~~ from block  $b$ .

Here  $X_B = ((x_{ij}^B))$

where  $x_{ij}^B = \begin{cases} 1 & \text{if the } i\text{th obs. comes from block } j \\ 0 & \text{o.w.} \end{cases}$

$$X_B = \begin{pmatrix} 1 & 0 & \dots & 0 \\ \vdots & \vdots & & \vdots \\ 0 & 1 & \dots & 0 \\ \vdots & \vdots & & \vdots \\ 0 & 0 & \dots & 1 \end{pmatrix}$$

↑  
k  
↓  
↑  
k  
↓  
↑  
k  
↓

$$\text{or, } X_B = \begin{pmatrix} \mathbf{1}_k & \mathbf{0}_k & \dots & \mathbf{0}_k \\ \mathbf{0}_k & \mathbf{1}_k & \dots & \mathbf{0}_k \\ \vdots & \vdots & & \vdots \\ \mathbf{0}_k & \mathbf{0}_k & \dots & \mathbf{1}_k \end{pmatrix}$$

Now,  $\Sigma = X_B X_B' \cdot \sigma_b^2 + I_n \cdot \sigma^2 \quad (\text{Disp}(\gamma))$

$$X_B X_B' = \begin{pmatrix} \mathbf{1}_k & \mathbf{0}_k & \dots & \mathbf{0}_k \\ \mathbf{0}_k & \mathbf{1}_k & \dots & \mathbf{0}_k \\ \vdots & \vdots & & \vdots \\ \mathbf{0}_k & \mathbf{0}_k & \dots & \mathbf{1}_k \end{pmatrix} \begin{pmatrix} \mathbf{1}_k' & \mathbf{0}_k' & \dots & \mathbf{0}_k' \\ \mathbf{0}_k' & \mathbf{1}_k' & \dots & \mathbf{0}_k' \\ \vdots & \vdots & & \vdots \\ \mathbf{0}_k' & \mathbf{0}_k' & \dots & \mathbf{1}_k' \end{pmatrix}$$

$$= \begin{pmatrix} J_k & 0_{k \times k} & 0_{k \times k} & \dots & 0 \\ 0 & J_k & 0 & & 0 \\ \vdots & & & & \\ 0 & 0 & 0 & & J_k \end{pmatrix} \quad \left| \begin{array}{l} J_k = \underline{1}_k \underline{1}_k' \\ = (k \times k) \text{ matrix with all} \\ \text{entries } 1 \end{array} \right.$$

$$= \text{Diag}(J_k, J_k, \dots, J_k)$$

$$\therefore \Sigma = \sigma^2 I_n + \sigma_b^2 \cdot \text{Diag}(J_k, J_k, \dots, J_k)$$

$$= \sigma^2 \underbrace{\text{Diag}(J_k, J_k, \dots, J_k)}_{b \text{ times}} + \sigma_b^2 \underbrace{\text{Diag}(J_k, J_k, \dots, J_k)}_{b \text{ times}}$$

$$= \underbrace{\text{Diag}(\sigma^2 I_k + \sigma_b^2 J_k, \sigma^2 I_k + \sigma_b^2 J_k, \dots, \sigma^2 I_k + \sigma_b^2 J_k)}_{b \text{ times}}$$

$$= \text{Diag}(L, L, \dots, L), \text{ where } L = \sigma^2 I_k + \sigma_b^2 J_k$$

$$\therefore \Sigma^{-1} = \text{Diag}(L^{-1}, L^{-1}, \dots, L^{-1}) ; L^{-1} = (\sigma^2 I_k + \sigma_b^2 J_k)^{-1}$$

$$= \frac{1}{\sigma^2} \left( I_k + \frac{\sigma_b^2}{\sigma^2} \underline{1}_k \underline{1}_k' \right) \left( A + \frac{\underline{u} \underline{u}'}{\sigma^2} \right)^{-1}$$

The matrix  $I_k + \frac{\sigma_b^2}{\sigma^2} \underline{1}_k \underline{1}_k'$  is of the form  $A + \frac{\underline{u} \underline{u}'}{\sigma^2}$  with

$$= A^{-1} - \frac{(A^{-1} \underline{u})(\underline{u}' A^{-1})}{1 + \underline{u}' A^{-1} \underline{u}}$$

$$A = I_k$$

$$\underline{u} = \frac{\sigma_b^2}{\sigma^2} \underline{1}_k$$

$$\underline{v} = \underline{1}_k$$

$$\therefore \left( I_k + \frac{\sigma_b^2}{\sigma^2} \underline{1}_k \underline{1}_k' \right)^{-1} = I_k - \frac{\left( \frac{\sigma_b^2}{\sigma^2} \underline{1}_k \right) \left( \underline{1}_k' \right)}{1 + \underline{1}_k' \underline{1}_k \left( \frac{\sigma_b^2}{\sigma^2} \right)}$$

$$= I_k - \left( \frac{\sigma_b^2}{\sigma^2 + k\sigma_b^2} \right) J_k$$

$$\therefore L^{-1} = \frac{1}{\sigma^2} I_k - \frac{\sigma_b^2}{\sigma^2(\sigma^2 + k\sigma_b^2)} J_k$$

Define  $w_1 = \frac{1}{\sigma^2}$  (reciprocal of intra-block variance)

$w_2 = \frac{1}{\sigma^2 + k\sigma_b^2}$  (reciprocal of inter-block variance)

$$\therefore w_1 - w_2 = \frac{k\sigma_b^2}{\sigma^2(\sigma^2 + k\sigma_b^2)}$$

$$\therefore L^{-1} = w_1 I_k - \left( \frac{w_1 - w_2}{k} \right) J_k$$

$$\therefore \Sigma^{-1} = \text{Diag}(L^{-1}, L^{-1}, \dots, L^{-1})$$

$$= \text{Diag} \left( w_1 I_k - \left( \frac{w_1 - w_2}{k} \right) J_k, \dots, w_1 I_k - \left( \frac{w_1 - w_2}{k} \right) J_k \right)$$

b times.

$$= w_1 I_n - \left( \frac{w_1 - w_2}{k} \right) \text{Diag}(J_k, J_k, \dots, J_k)$$

$$= w_1 I_n - \left( \frac{w_1 - w_2}{k} \right) X_p X_p'$$

Normal equations.

$$A' \Sigma^{-1} A \underline{\theta} = A' \Sigma^{-1} \underline{y}$$

$$A = [1: X_p]; \quad \underline{\theta} = \begin{pmatrix} \mu \\ \tau \end{pmatrix}$$

$$A' \Sigma^{-1} A = \begin{bmatrix} 1' \\ X_p' \end{bmatrix} \Sigma^{-1} \begin{bmatrix} 1_n \\ X_p \end{bmatrix} = \begin{pmatrix} 1_n' \Sigma^{-1} 1_n & 1_n' \Sigma^{-1} X_p \\ X_p' \Sigma^{-1} 1_n & X_p' \Sigma^{-1} X_p \end{pmatrix}$$

$$\therefore \begin{pmatrix} 1_n' \Sigma^{-1} 1_n & 1_n' \Sigma^{-1} X_p \\ X_p' \Sigma^{-1} 1_n & X_p' \Sigma^{-1} X_p \end{pmatrix} \underline{\theta} = \begin{pmatrix} 1_n' \Sigma^{-1} \underline{y} \\ X_p' \Sigma^{-1} \underline{y} \end{pmatrix}$$



We have

$$\frac{1}{n} \mathbf{1}' \Sigma^{-1} \mathbf{1}_n = \frac{1}{n} \left[ w_1 \mathbf{I}_n - \frac{w_1 - w_2}{k} \mathbf{x}_\beta \mathbf{x}_\beta' \right] \frac{1}{n} \mathbf{1}_n$$

$$= \frac{w_1}{n} \mathbf{1}' \mathbf{1}_n - \frac{(w_1 - w_2)}{k} \left( \frac{1}{n} \mathbf{1}' \mathbf{1}_n \right)$$

$$= w_1 \frac{1}{n} \mathbf{1}' \mathbf{1}_n - \frac{(w_1 - w_2)}{k} \left( \frac{1}{n} \mathbf{1}' \mathbf{x}_\beta \right) \left( \mathbf{x}_\beta' \frac{1}{n} \mathbf{1}_n \right)$$

$$= w_1 n - \frac{(w_1 - w_2)}{k} \left( \frac{k'}{n} \right) \left( \frac{k}{n} \right)$$

$$= w_1 n - \frac{(w_1 - w_2)}{k} \left( k \frac{1}{n} \right) \left( \frac{k}{n} \right) [\because \text{proper design}]$$

$$= w_1 n - k(w_1 - w_2) \left( \frac{1}{n} \right) \left( \frac{1}{n} \right)$$

$$= w_1 n - bk(w_1 - w_2)$$

$$= w_1 n - n(w_1 - w_2)$$

$$= nw_2$$

$$\frac{1}{n} \mathbf{1}' \Sigma^{-1} \mathbf{x}_\tau = w_2 \mathbf{x}'_\tau$$

$$\mathbf{x}'_\tau \Sigma^{-1} \mathbf{x}_\tau = w_1 n - \frac{(w_1 - w_2)}{k} \mathbf{N} \mathbf{N}'$$

$$\frac{1}{n} \mathbf{1}' \Sigma^{-1} \mathbf{x}_\tau = \frac{1}{n} \left[ w_1 \mathbf{I}_n - \frac{w_1 - w_2}{k} \mathbf{x}_\beta \mathbf{x}_\beta' \right] \mathbf{x}_\tau$$

$$= w_1 \frac{1}{n} \mathbf{1}' \mathbf{x}_\tau - \frac{w_1 - w_2}{k} \left( \frac{1}{n} \mathbf{1}' \mathbf{x}_\beta \right) \left( \mathbf{x}_\beta' \mathbf{x}_\tau \right)$$

$$= w_1 \mathbf{x}'_\tau - \frac{w_1 - w_2}{k} \left( \frac{k'}{n} \right) \left( \frac{k}{n} \right)$$

$$= w_1 \mathbf{x}'_\tau - \frac{w_1 - w_2}{k} \left( k \frac{1}{n} \right) \left( \frac{k}{n} \right)$$

$$= w_1 \mathbf{x}'_\tau - \frac{(w_1 - w_2)}{k} \mathbf{N} \mathbf{N}'$$

$$= w_1 \mathbf{x}'_\tau - \frac{(w_1 - w_2)}{k} \left( \frac{k}{n} \right) \left( \frac{k}{n} \right)$$

$$= w_1 \mathbf{x}'_\tau - w_1 \mathbf{x}'_\tau + w_2 \mathbf{x}'_\tau$$

$$= w_2 \mathbf{x}'_\tau$$

$$\mathbf{x}'_\tau \Sigma^{-1} \mathbf{x}_\tau = \mathbf{x}'_\tau \left[ w_1 \mathbf{I}_n - \frac{w_1 - w_2}{k} \mathbf{x}_\beta \mathbf{x}_\beta' \right] \mathbf{x}_\tau$$

$$= w_1 \mathbf{x}'_\tau \mathbf{x}_\tau - \frac{w_1 - w_2}{k} \left( \mathbf{x}'_\tau \mathbf{x}_\beta \right) \left( \mathbf{x}_\beta' \mathbf{x}_\tau \right)$$

$$= w_1 n - \frac{w_1 - w_2}{k} \mathbf{N} \mathbf{N}'$$

$$\begin{aligned}
 \frac{1}{n} \Sigma^{-1} Y &= \frac{1}{n} \left[ w_1 I_n - \frac{(w_1 - w_2)}{k} x_\beta x_\beta' \right] Y \\
 &= w_1 \frac{1}{n} Y - \frac{(w_1 - w_2)}{k} \left( \frac{1}{n} x_\beta \right) \left( x_\beta' Y \right) \\
 &= w_1 \frac{1}{n} Y - \frac{(w_1 - w_2)}{k} (k \frac{1}{n} \beta) \left( \frac{1}{n} \beta \right) \\
 &= w_1 G_T - (w_1 - w_2) G_T \quad [\because \frac{1}{n} Y = \frac{1}{n} \beta = G_T]
 \end{aligned}$$

$$\begin{aligned}
 x_\gamma' \Sigma^{-1} Y &= x_\gamma' \left[ w_1 I_n - \frac{w_1 - w_2}{k} x_\beta x_\beta' \right] Y \\
 &= w_1 x_\gamma' Y - \frac{w_1 - w_2}{k} (x_\gamma' x_\beta) (x_\beta' Y) \\
 &= w_1 T - \frac{w_1 - w_2}{k} N \beta \\
 &= w_1 T - w_1 \frac{N \beta}{k} + w_2 \frac{N \beta}{k} \\
 &= w_1 \left( T - \frac{N \beta}{k} \right) + w_2 \left( \frac{N \beta}{k} \right) \\
 &= w_1 G + w_2 (T - G)
 \end{aligned}$$

$$\begin{aligned}
 G &= T - \frac{N \beta}{k} \\
 &= T - \frac{N \beta}{k} \\
 &\text{for proper design}
 \end{aligned}$$

$$\therefore \begin{pmatrix} w_2 n & w_2 x_\gamma' \\ w_2 x_\gamma & w_1 I_n - \frac{(w_1 - w_2)}{k} N N' \end{pmatrix} \begin{pmatrix} \mu \\ \gamma \end{pmatrix} = \begin{pmatrix} w_2 G_T \\ w_1 G + w_2 (T - G) \end{pmatrix}$$

$$n w_2 \hat{\mu} + w_2 x_\gamma' \hat{\gamma} = w_2 G_T \quad \text{--- (1)}$$

$$w_2 x_\gamma' \hat{\mu} + w_1 I_n \hat{\gamma} - \frac{(w_1 - w_2)}{k} N N' \hat{\gamma} = w_1 G + w_2 (T - G) \quad \text{--- (2)}$$

premultiply eq<sup>n</sup> (1) by  $x$

$$\underline{w_2 x_\gamma} w_2 x_\gamma' \hat{\mu} + w_2 \frac{x_\gamma x_\gamma'}{n} \hat{\gamma} = \frac{w_2}{n} x_\gamma G_T \quad \text{--- (3)}$$

$w_2 x_\gamma$  (2)-(3) given

$$\left( w_1 I_n - \frac{(w_1 - w_2)}{k} N N' - w_2 \frac{x_\gamma x_\gamma'}{n} \right) \hat{\gamma} = w_1 G + w_2 (T - G) - \frac{w_2}{n} x_\gamma G_T$$

$$\text{or, } \left[ w_1 \left( D_r - \frac{NN'}{k} \right) + w_2 \left( \frac{NN'}{k} - \frac{\tilde{y}\tilde{y}'}{n} \right) \right] \hat{\tilde{\tau}} = w_1 \tilde{Q} + w_2 \left( T - Q - \frac{\tilde{y}\tilde{G}}{n} \right)$$

$$\text{or, } (w_1 C + w_2 C^*) \hat{\tilde{\tau}} = w_1 \tilde{Q} + w_2 \tilde{Q}^*$$

where

$$C^* = \frac{NN'}{k} - \frac{\tilde{y}\tilde{y}'}{n}$$

$$Q^* = T - Q - \frac{\tilde{y}\tilde{G}}{n}$$

$\text{Eq}^n(4)$  is called adjusted inter and intra block normal eq<sup>n</sup>,

23.04.2024

Result 1  $E(\tilde{Q}) = C\tilde{\tau}$

Proof  $\gg \tilde{Q} = T - \frac{NB}{k}$

$$\text{or, } \tilde{Q} = X'\tilde{\tau} - \frac{NX\beta}{k} = \left( X'\tilde{\tau} - \frac{NX\beta}{k} \right) \tilde{y}$$

$$E(\tilde{Q}) = \left( X'\tilde{\tau} - \frac{NX\beta}{k} \right) E(\tilde{y})$$

$$= \left( X'\tilde{\tau} - \frac{NX\beta}{k} \right) (\mu_{\tilde{y}} + X'\tilde{\tau})$$

$$= \mu X'\tilde{\tau} - \frac{\mu}{k} NX\beta + X'\tilde{\tau} X'\tilde{\tau} - \frac{NX\beta X'\tilde{\tau}}{k}$$

$$= \mu_0 \tilde{y} - \frac{\mu}{k} Nk + D_r \tilde{y} - \frac{NN'\tilde{y}}{k}$$

$$= \mu \tilde{y} + \frac{\mu}{k} N(k \perp b) + \left( D_r - \frac{NN'}{k} \right) \tilde{y}$$

$$= \mu_0 \tilde{y} - \mu(N \perp b) + C\tilde{\tau} \left[ \because k \equiv k \perp b, C = D_r - \frac{NN'}{k} \right]$$

$$= \mu_0 \tilde{y} - \mu \tilde{y} + C\tilde{\tau}$$

$$= C\tilde{\tau}$$

Result 2  $E(\tilde{Q}^*) = C^* \tilde{\tau}$

Proof  $\gg \tilde{Q}^* = T - Q - \frac{\tilde{y}\tilde{G}}{n} = \frac{NB}{k} - \frac{\tilde{y}\tilde{G}}{n}$

$$= \frac{NX\beta}{k} - \frac{\tilde{y}\tilde{y}'\tilde{y}}{n}$$

$$= \left( \frac{N X' P}{k} - \frac{\tilde{x} \tilde{x}' n}{n} \right) \tilde{y}$$

$$E(\tilde{g}^*) = \left( \frac{N X' P}{k} - \frac{\tilde{x} \tilde{x}' n}{n} \right) E(\tilde{y})$$

$$= \left( \frac{N X' P}{k} - \frac{\tilde{x} \tilde{x}' n}{n} \right) (\mu_{\tilde{y}n} + X' \tilde{\tau})$$

$$= \mu_{\tilde{y}n} \frac{N X' P}{k} - \frac{\mu_{\tilde{y}n} \tilde{x} \tilde{x}' n}{n} + \frac{N X' P}{k} X' \tilde{\tau} - \frac{\tilde{x} \tilde{x}' n}{n} X' \tilde{\tau}$$

$$= \frac{\mu_{\tilde{y}n} N k \cdot 1}{k} - \frac{\mu_{\tilde{y}n} \tilde{x} \tilde{x}' n}{n} + \frac{N N' \tilde{\tau}}{k} - \frac{\tilde{x} \tilde{x}' \tilde{\tau}}{n}$$

$$= \mu_{\tilde{y}} - \mu_{\tilde{y}} + \left( \frac{N N'}{k} - \frac{\tilde{x} \tilde{x}'}{n} \right) \tilde{\tau}$$

$$= C^* \tilde{\tau}$$

Result 3

$$\text{Disp} \begin{pmatrix} \tilde{g} \\ \tilde{g}^* \\ g \end{pmatrix} = \begin{pmatrix} \sigma^2 & 0 & 0 \\ 0 & C^*/w_2 & 0 \\ 0 & 0 & \frac{n}{w_2} \end{pmatrix}$$

Proof >> To be shown  $\text{Disp}(\tilde{g}) = \sigma^2$

$$\text{Disp}(\tilde{g}^*) = C^*/w_2$$

$$\text{Var}(g) = \frac{n}{w_2}$$

$$\text{Cov}(\tilde{g}, \tilde{g}^*) = 0$$

$$\text{Cov}(\tilde{g}, g) = 0$$

$$\text{Cov}(\tilde{g}^*, g) = 0$$

$$\tilde{y} = \left( X'_{\tau} - \frac{N X'_{\beta}}{k} \right) \tilde{y}$$

$$\text{Disp}(\tilde{y}) = \left( X'_{\tau} - \frac{N X'_{\beta}}{k} \right) \text{Disp}(y) \left( X_{\tau} - \frac{X_{\beta} N'}{k} \right)$$

$$= \left( X'_{\tau} - \frac{N X'_{\beta}}{k} \right) \Sigma \left( X_{\tau} - \frac{X_{\beta} N'}{k} \right)$$

$$= \left( X'_{\tau} \Sigma X_{\tau} - \frac{N X'_{\beta} \Sigma X_{\tau}}{k} \right)$$

$$= X'_{\tau} \Sigma X_{\tau} - \frac{N X'_{\beta} \Sigma X_{\tau}}{k} - \frac{X'_{\tau} \Sigma X_{\beta} N'}{k} + \frac{N X'_{\beta} \Sigma X_{\beta} N'}{k^2}$$

$$\begin{aligned} X'_{\tau} \Sigma X_{\tau} &= X'_{\tau} [\sigma^2 I_n + \sigma_b^2 X_{\beta} X'_{\beta}] X_{\tau} \\ &= X'_{\tau} X_{\tau} + \sigma_b^2 (X'_{\tau} X_{\beta}) (X'_{\beta} X_{\tau}) \\ &= \sigma^2 D_r + \sigma_b^2 N N' \end{aligned}$$

$$\begin{aligned} X'_{\tau} \Sigma X_{\beta} &= X'_{\tau} [\sigma^2 I_n + \sigma_b^2 X_{\beta} X'_{\beta}] X_{\beta} \\ &= \sigma^2 X'_{\tau} X_{\beta} + \sigma_b^2 (X'_{\tau} X_{\beta}) (X'_{\beta} X_{\beta}) \\ &= \sigma^2 N + k \sigma_b^2 N \quad [X'_{\beta} X_{\beta} = D_k = k I_b] \\ &= N(\sigma^2 + k \sigma_b^2) = \frac{N}{w_2} \quad [\because w_2 = \frac{1}{\sigma^2 + k \sigma_b^2}] \end{aligned}$$

$$\begin{aligned} X'_{\beta} \Sigma X_{\beta} &= X'_{\beta} (\sigma^2 I_n + \sigma_b^2 X_{\beta} X'_{\beta}) X_{\beta} \\ &= \sigma^2 X'_{\beta} X_{\beta} + \sigma_b^2 (X'_{\beta} X_{\beta}) (X'_{\beta} X_{\beta}) \\ &= (k \sigma^2 + k^2 \sigma_b^2) I_b \\ &= k(\sigma^2 + k \sigma_b^2) I_b \\ &= \frac{k I_b}{w_2} \end{aligned}$$

$$\begin{aligned} &\rightarrow \\ &= \sigma^2 D_r + \sigma_b^2 N N' - \frac{(\sigma^2 + k \sigma_b^2) N N'}{k} - \frac{(\sigma^2 + k \sigma_b^2) N N'}{k} \\ &\quad + k \frac{(\sigma^2 + k \sigma_b^2) N N'}{k^2} \end{aligned}$$



$$= \sigma^2 D_r + \sigma_b^2 N N' - \frac{(\sigma^2 + k \sigma_b^2)}{k} N N'$$

$$= \sigma^2 \left( D_r - \frac{N N'}{k} \right)$$

$$= \sigma^2$$

$$g^* = \left( \frac{N x \beta'}{k} - \frac{\sum \tilde{y} \tilde{y}'}{n} \right) \tilde{y}$$

$$\text{Disp}(g^*) = \left( \frac{N x \beta'}{k} - \frac{\sum \tilde{y} \tilde{y}'}{n} \right) \text{Disp}(\tilde{y}) \left( \frac{x \beta' N'}{k} - \frac{\sum \tilde{y} \tilde{y}'}{n} \right)$$

$$= \frac{N x \beta' \sum x \beta N'}{k^2} - \frac{\sum \tilde{y} \tilde{y}' \sum x \beta N'}{n k} - \frac{N x \beta' \sum \tilde{y} \tilde{y}'}{n k} + \frac{\sum \tilde{y} \tilde{y}' \sum \tilde{y} \tilde{y}'}{n^2}$$

$$\sum \tilde{y} \tilde{y}' = \sigma^2 \sum \tilde{y} \tilde{y}' + \sigma_b^2 (\sum \tilde{y} x \beta) (x \beta' \sum \tilde{y})$$

$$= \sigma^2 n + \sigma_b^2 (k') (k)$$

$$= \sigma^2 n + k^2 \sigma_b^2 (\tilde{y} \tilde{y}')$$

$$= \sigma^2 n + k^2 \sigma_b^2$$

$$= n (\sigma^2 + k \sigma_b^2) \quad [bk = n]$$

$$= \frac{n}{w_2}$$

$$x \beta' \sum \tilde{y} = \sigma^2 x \beta' \sum \tilde{y} + \sigma_b^2 (x \beta' x \beta) (x \beta' \sum \tilde{y})$$

$$= \sigma^2 (k \tilde{y}) + \sigma_b^2 (k I_b) (k \tilde{y})$$

$$= k (\sigma^2 + k \sigma_b^2) \tilde{y}$$

$$= \frac{k}{w_2} \tilde{y}$$

$$= \frac{k D_r N N'}{k^2 w_2} + \frac{k}{w_2} \frac{\sum \tilde{y} \tilde{y}' N'}{n k} - \frac{k}{w_2} \frac{N \sum \tilde{y} \tilde{y}'}{n k} + \frac{n}{w_2} \frac{\sum \tilde{y} \tilde{y}'}{n^2}$$

$$2 \frac{NN'}{KW_2} - \frac{\tilde{y}\tilde{y}'}{nw_2} + \frac{\tilde{y}\tilde{y}'}{nw_2} + \frac{\tilde{y}\tilde{y}'}{w_2 n} \quad \left| N\tilde{1}_b = \tilde{y} \right.$$

$$= \frac{1}{w_2} \left( \frac{NN'}{K} - \frac{\tilde{y}\tilde{y}'}{n} \right)$$

$$= \frac{C^*}{w_2}$$

$$\text{Var}(\hat{\beta}) = \text{Var}(\tilde{1}'\tilde{y})$$

$$= \tilde{1}'\Sigma\tilde{1}_n$$

$$= \frac{n}{w_2} \text{ [already found]}$$

$$\text{Cov}(\hat{\beta}, \hat{\beta}^*) = \text{Cov}\left(\left(\tilde{y}' - \frac{N\tilde{y}\tilde{1}_b'}{K}\right)\tilde{y}, \left(\frac{N\tilde{y}\tilde{1}_b'}{K} - \frac{\tilde{y}\tilde{1}_b'}{n}\right)\tilde{y}\right)$$

$$= \left(\tilde{y}' - \frac{N\tilde{y}\tilde{1}_b'}{K}\right)\Sigma\left(\frac{N\tilde{y}\tilde{1}_b'}{K} - \frac{\tilde{y}\tilde{1}_b'}{n}\right)$$

$$= \frac{\tilde{y}'\Sigma\tilde{y}\tilde{1}_b'}{K} - \frac{N\tilde{y}\tilde{1}_b'\Sigma\tilde{y}\tilde{1}_b'}{K^2} - \frac{\tilde{y}'\Sigma\tilde{y}\tilde{1}_b'}{n} + \frac{N\tilde{y}\tilde{1}_b'\Sigma\tilde{y}\tilde{1}_b'}{nk}$$

$$\tilde{y}'\Sigma\tilde{1}_n = \sigma^2\tilde{y}'\tilde{1}_n + \sigma_b^2(\tilde{y}'\tilde{1}_b)(\tilde{1}_b'\tilde{1}_n)$$

$$= \sigma^2\tilde{y}'\tilde{1}_n + \sigma_b^2 N(K\tilde{1}_b)$$

$$= \sigma^2\tilde{y}'\tilde{1}_n + K\sigma_b^2\tilde{y}'\tilde{1}_n$$

$$= \frac{\tilde{y}'\tilde{1}_n}{w_2}$$

$$= \frac{NN'}{w_2 K} - \frac{NK\tilde{1}_b\tilde{1}_b'N'}{w_2 K^2} - \frac{\tilde{y}'\tilde{y}\tilde{1}_b'}{w_2 n} + \frac{NK\tilde{1}_b\tilde{1}_b'\tilde{y}'\tilde{1}_n}{w_2 nk}$$

$$= \frac{NN'}{w_2 K} - \frac{NN'}{w_2 K}$$

$$- \frac{\tilde{y}\tilde{y}'}{w_2 n} + \frac{\tilde{y}\tilde{y}'}{w_2 nk}$$

$$= 0$$

$$\begin{aligned}
\text{cov}(\mathcal{Q}_2, G_2) &= \text{cov}\left(\left(x'_{\tilde{2}} - \frac{N x' \beta}{k}\right) \tilde{y}_2, \frac{1}{n} \tilde{y}_2\right) \\
&= \left(x'_{\tilde{2}} - \frac{N x' \beta}{k}\right) \Sigma \frac{1}{n} \tilde{y}_2 \\
&= x'_{\tilde{2}} \Sigma \frac{1}{n} \tilde{y}_2 - \frac{N x' \beta \Sigma \frac{1}{n} \tilde{y}_2}{k} \\
&= \frac{\tilde{\gamma}}{w_2} - \frac{k}{w_2 k} N \frac{1}{n} \tilde{y}_2 \\
&= \frac{\tilde{\gamma}}{w_2} - \frac{\tilde{\gamma}}{w_2} \\
&= 0_{2 \times 1}
\end{aligned}$$

$$\begin{aligned}
\text{cov}(\mathcal{Q}_2^*, G_2) &= \text{cov}\left(\left(\cancel{N x' \beta} - \frac{x' \tilde{y}_2}{n}\right) \tilde{y}_2, \frac{1}{n} \tilde{y}_2\right) \\
&= \left(\frac{N x' \beta}{k} - \frac{x' \tilde{y}_2}{n}\right) \Sigma \frac{1}{n} \tilde{y}_2 \\
&= \frac{N x' \beta \Sigma \frac{1}{n} \tilde{y}_2}{k} - \frac{x' \tilde{y}_2 \Sigma \frac{1}{n} \tilde{y}_2}{n} \\
&= \frac{N k \frac{1}{n}}{k w_2} - \frac{x' \tilde{y}_2}{n w_2} \\
&= \frac{\tilde{\gamma}}{w_2} - \frac{\tilde{\gamma}}{w_2} \\
&= 0_{\tilde{2}}
\end{aligned}$$

## Estimation of Variance Component

25.4.24

$$\text{Normal eq}^2 (w_1 c + w_2 c^*) \tilde{y} = w_1 \tilde{y} + w_2 \tilde{y}^*$$

$$w_1 = \frac{1}{\sigma^2}, \quad w_2 = \frac{1}{\sigma^2 + k\sigma_b^2}$$

In reality, we don't know the value of  $\sigma^2$  and  $\sigma_b^2$ , so we need to estimate them from the data. We'll follow the method by R.C. Bose (Assuming proper and binary design)

In intra-Block analysis

$$(1) \text{ Total } SS = S^2 = \sum_i \sum_j y_{ij}^2 = \frac{G^2}{n}$$

$$(2) \text{ Unadjusted block } SS = S_b'^2 = \sum_j \frac{B_j^2}{k} - \frac{G^2}{n}$$

$$(3) \text{ Unadjusted treatment } SS = S_t'^2 = \sum_i \frac{T_i^2}{n_i} - \frac{G^2}{n}$$

$$(4) \text{ Adjusted trt. } SS = S_t^2 = \tilde{y}' \tilde{y} =$$

$$(5) \text{ Residual } SS = R_o^2; \quad S^2 = S_b'^2 + S_t'^2 + R_o^2 = S_b^2 + S_t^2 + R_o^2$$

↓  
Adjusted block SS  
 $\hat{\beta} \hat{\beta}'$

We find the expectations of these SS under mixed effect model.

$$\begin{aligned} E(G) &= E(\tilde{y}' \tilde{y}) \\ &= \tilde{y}' E(\tilde{y}) \\ &= \tilde{y}' (\tilde{y}' \mu + \tilde{y}' \tau) \begin{pmatrix} \mu \\ \tau \end{pmatrix} \\ &= \tilde{y}' \tilde{y} \mu + \tilde{y}' \tau \tilde{y} \\ &= n\mu + \sum_i \tau_i \tau_i \\ &= n\mu + \sum_{i=1}^v \tau_i \tau_i \\ &= n(\mu + \bar{\tau}) \end{aligned}$$

$$[\bar{\tau} = \sum_{i=1}^v \tau_i \tau_i / \sum_{i=1}^v \tau_i]$$

$$E\left(\frac{G^2}{n}\right) = \frac{1}{n} \left[ \text{Var}(G) + (E(G))^2 \right]$$

$$= \frac{1}{n} \left[ n(\sigma^2 + k\sigma_b^2) + n^2(\mu + \bar{\tau})^2 \right] \quad \text{--- (1)}$$

$$E\left(\frac{T_i^2}{r_i}\right) = \frac{1}{r_i} \left[ \text{Var}(T_i) + (E(T_i))^2 \right]$$

$$\underline{T} = \underline{X}'_T \underline{Y}$$

$$E(\underline{T}) = \underline{X}'_T E(\underline{Y})$$

$$= \underline{X}'_T (\underline{I}_n; \underline{X}_T) \begin{pmatrix} \mu \\ \underline{\tau} \end{pmatrix}$$

$$= \underline{X}'_T \underline{I}_n \mu + \underline{X}'_T \underline{X}_T \cdot \underline{\tau}$$

$$= \underline{r} \mu + D_T \underline{\tau}$$

$$E(T_i) = r_i \mu + r_i \tau_i$$

$$= r_i (\mu + \tau_i)$$

$$\text{Disp}(\underline{T}) = \underline{X}'_T \text{Disp}(\underline{Y}) \underline{X}_T$$

$$= \underline{X}'_T \Sigma \underline{X}_T$$

$$= \underline{X}'_T (\sigma^2 \underline{I}_n + \sigma_b^2 \underline{X}_\beta \underline{X}'_\beta) \underline{X}_T$$

$$= \sigma^2 D_T + \sigma_b^2 N N'$$

$$V(\tau_i) = \sigma^2 r_i + \sigma_b^2 \sum_{j=1}^b n_{ij} \quad ; \quad i=1(1)v$$

$$= \sigma^2 r_i + \sigma_b^2 \sum_{j=1}^b n_{ij} \quad [\because \text{The design is binary}]$$

$$= \sigma^2 r_i + \sigma_b^2 r_i$$

$$= r_i (\sigma^2 + \sigma_b^2)$$

$$E\left(\frac{T_i^2}{r_i}\right) = \frac{1}{r_i} \left[ r_i (\sigma^2 + \sigma_b^2) + r_i^2 (\mu + \tau_i)^2 \right] \quad \text{--- (2)}$$



$$E(S_t'^2) = E\left(\sum_i \frac{T_i^2}{r_i}\right) - E\left(\frac{G^2}{n}\right)$$

$$= \sum_i E\left(\frac{T_i^2}{r_i}\right) - E\left(\frac{G^2}{n}\right)$$

$$= \sum_{i=1}^v \left[ (\sigma^2 + \sigma_b^2) + r_i (\mu + \tau_i)^2 \right] - (\sigma^2 + k\sigma_b^2) - n(\mu + \bar{\tau})^2$$

$$= (v-1)\sigma^2 + (v-k)\sigma_b^2 + \sum_{i=1}^v r_i (\mu + \tau_i)^2 - n(\mu + \bar{\tau})^2$$

$$= (v-1)\sigma^2 + (v-k)\sigma_b^2 + \sum_{i=1}^v r_i (\tau_i - \bar{\tau})^2 \quad \text{--- (3)}$$

$$E(\text{Total SS}) = \sum_{u=1}^n E(Y_u^2) - E\left(\frac{G^2}{n}\right)$$

$$\text{Total SS} = \sum_{i=1}^b \sum_{j=1}^v y_{ij}^2 - \frac{G^2}{n}$$

$$= \sum_{u=1}^n y_u^2 - \frac{G^2}{n}$$

$$= \sum_{u=1}^n \left[ \text{var}(y_u) - \{E(y_u)\}^2 \right] - E\left(\frac{G^2}{n}\right)$$

$$= \sum_{u=1}^n \left[ (\sigma^2 + \sigma_b^2) + \sum_{i=1}^v \sum_{j=1}^v n_{ij} (\mu + \tau_i)^2 \right] - (\sigma^2 + k\sigma_b^2) - n(\mu + \bar{\tau})^2$$

$$= (n-1)\sigma^2 + (n-k)\sigma_b^2 + \sum_{i=1}^v r_i (\tau_i - \bar{\tau})^2 \quad \text{--- (4)}$$

$$B(R_0^2) = (n-b-v+1)\sigma^2 \quad (\text{connected design})$$

$$E(S_b^2) = E(\text{Adjusted block SS})$$

$$= E(\text{Total SS}) - E(S_t'^2) - E(R_0^2)$$

$$= (n-1)\sigma^2 + (n-k)\sigma_b^2 + \sum_{i=1}^v r_i (\tau_i - \bar{\tau})^2 - (v-1)\sigma^2$$

$$- (v-k)\sigma_b^2 - \sum_{i=1}^v r_i (\tau_i - \bar{\tau})^2 - (n-b-v+1)\sigma^2$$

$$= (b-1)\sigma^2 + (n-v)\sigma_b^2 \quad \text{--- (5)}$$

$$\hat{\sigma}^2 = \frac{R_0}{n-b-v+1}$$

$$E\left[\frac{S_b^2}{n-v} - \left(\frac{b-1}{n-v}\right) \frac{R_0^2}{(n-b-v+1)}\right] = \sigma_b^2$$

$$\left[ \sum_{i=1}^b \sum_{j=1}^v n_{ij} (\mu + \tau_i)^2 - n(\mu + \bar{\tau})^2 \right]$$

$$= \sum \sum n_{ij} (\mu^2 + 2\mu\tau_i + \tau_i^2) - n(\mu^2 + 2\mu\bar{\tau} + \bar{\tau}^2)$$

$$= n\mu^2 + 2\mu \sum \sum n_{ij} \tau_i + \sum \sum n_{ij} \tau_i^2 - n\mu^2 - 2n\mu\bar{\tau} - n\bar{\tau}^2$$

$$= 2\mu \left( \sum_i r_i \right) \bar{\tau} + \sum_i r_i \tau_i^2 - n\bar{\tau}^2 - 2n\mu\bar{\tau}$$

$$= 2n\mu\bar{\tau} + \sum_i r_i \tau_i^2 - n\bar{\tau}^2 - 2n\mu\bar{\tau}$$

$$\hat{\sigma}_b^2 = \left( \frac{b-1}{n-v} \right) \left( \frac{S_b^2}{b-1} - \frac{R_b^2}{n-b-v+1} \right)$$

## Balanced Incomplete Block Design (BIBD)

30.4.24

(ii) Each trt. appears in  $r$  blocks.

(iii) Every pair of trt. appear together in  $\lambda$  blocks.

The parameters  $v, b, r, k, \lambda$  are called the parameters of the BIBD. They are related by the following identities:

i)  $b \cdot k = v \cdot r$

ii)  $r(k-1) = \lambda(v-1)$

BIBD is the binary design as in this case the incidence matrix has only two elements 0, 1

Example

1	2	4
2	3	5
3	4	6
4	5	7
5	6	1
6	7	2
7	1	3

$$\begin{aligned} v &= 7 \\ b &= 7 \\ r &= 3 \\ k &= 3 \\ \lambda &= 1 \end{aligned}$$

Result >>

For a BIBD,  $r(k-1) = \lambda(v-1)$

Proof >> As a BIBD is proper and equireplicate,

$$N \mathbf{1}_b = r = r \mathbf{1}_v$$

$$\mathbf{1}_v' N = k' = k \mathbf{1}_b'$$

The  $(i, j)^{th}$  element of  $NN'$  is

$$\sum_{l=1}^b n_{il} n_{jl}$$