

Assignment 1

Given the following probability functions:

$$\xi \sim \text{Poisson}(\lambda) \iff P(\xi = k) = \frac{\lambda^k}{k!} e^{-\lambda} \quad (1)$$

$$\eta \sim \text{Bernoulli}(p) \iff P(\eta|k) = \binom{k}{\eta} p^\eta (1-p)^{k-\eta} \quad (2)$$

where λ and $0 \leq p \leq 1$ are fixed parameters, proof that:

$$\eta \sim \text{Poisson}(p \cdot \lambda) \iff P(\eta) = \frac{(p\lambda)^\eta}{\eta!} e^{-p\lambda} \quad (3)$$

First, note that:

$$P(\eta) = \sum_k P(\eta, k) = \sum_k P(\eta|k) \cdot P(k) = \sum_k P(\eta|k) \cdot P(\xi = k) \quad (4)$$

Hence we fill in equations 1 and 2, execute the sum (over all values of $k = \{\eta, \eta + 1, \eta + 2, \dots\}$) and simplify to proof equation 3.

$$P(\eta) = \sum_{k=\eta}^{\infty} \binom{k}{\eta} p^\eta (1-p)^{k-\eta} \frac{\lambda^k}{k!} e^{-\lambda} \quad (5)$$

$$= e^{-\lambda} \sum_{k=\eta}^{\infty} \frac{1}{\eta!(k-\eta)!} p^\eta \lambda^k (1-p)^{k-\eta} \quad (6)$$

Substitute $h = k - \eta$ to find:

$$P(\eta) = e^{-\lambda} \sum_{h=0}^{\infty} \frac{1}{\eta!h!} p^\eta \lambda^{h+\eta} (1-p)^h \quad (7)$$

$$= e^{-\lambda} \frac{(p\lambda)^\eta}{\eta!} \sum_{h=0}^{\infty} \frac{1}{h!} (\lambda(1-p))^h \quad (8)$$

Finally, use the power series expansion of $e^x = \sum_{x=0}^{\infty} \frac{x^k}{k!}$:

$$P(\eta) = e^{-\lambda} \frac{(p\lambda)^\eta}{\eta!} e^{\lambda(1-p)} = \frac{(p\lambda)^\eta}{\eta!} e^{\lambda(1-p)-\lambda} = \frac{(p\lambda)^\eta}{\eta!} e^{-p\lambda} \quad (9)$$

which corresponds to equation 3 and finalizes the proof.

Assignment 2

There are two Gaussian distributions of review time t for the strict reviewer $R = s$ and the kind reviewer $R = k$:

$$P(t|R) = \frac{1}{\sqrt{2\pi\sigma_R^2}} \exp\left(-\frac{(t - \mu_R)^2}{2\sigma_R^2}\right) \quad (10)$$

where $\mu_s = 30, \mu_k = 20, \sigma_s = 10, \sigma_k = 5$. We are asked to assess the probability that $R = k$ given $t = 10$. Following Bayes we denote:

$$P(R = k|t = 10) = \frac{P(t = 10|R = k) \cdot P(R = k)}{P(t = 10)} \quad (11)$$

$$= \frac{P(t = 10|R = k) \cdot P(R = k)}{P(t = 10|R = k) \cdot P(R = k) + P(t = 10|R = s) \cdot P(R = s)} \quad (12)$$

because $R \in \{s, k\}$ and probabilities are normalized $\sum_R P(t|R) = 1$. Rewriting equation 11 yields:

$$P(R = k|t = 10) = \frac{\frac{1}{\sqrt{2\pi\sigma_k^2}} \exp\left(\frac{-(t-\mu_k)^2}{2\sigma_k^2}\right)}{\frac{1}{\sqrt{2\pi\sigma_k^2}} \exp\left(\frac{-(t-\mu_k)^2}{2\sigma_k^2}\right) + \frac{1}{\sqrt{2\pi\sigma_s^2}} \exp\left(\frac{-(t-\mu_s)^2}{2\sigma_s^2}\right)} \quad (13)$$

$$= \frac{\frac{1}{\sigma_k} \exp\left(\frac{-(t-\mu_k)^2}{\sigma_k^2}\right)}{\frac{1}{\sigma_k} \exp\left(\frac{-(t-\mu_k)^2}{\sigma_k^2}\right) + \frac{1}{\sigma_s} \exp\left(\frac{-(t-\mu_s)^2}{\sigma_s^2}\right)} \quad (14)$$

$$= \frac{1}{1 + \frac{\sigma_k}{\sigma_s} \exp\left(\frac{-(t-\mu_s)^2}{\sigma_s^2} + \frac{(t-\mu_k)^2}{\sigma_k^2}\right)} \quad (15)$$

$$= \frac{1}{1 + \frac{\sigma_k}{\sigma_s} \exp\left(\frac{\sigma_s^2(t-\mu_k)^2 - \sigma_k^2(t-\mu_s)^2}{\sigma_k^2\sigma_s^2}\right)} \quad (16)$$

Solve numerically:

$$P(R = k|t = 10) = \frac{1}{1 + \frac{1}{2} \exp\left(\frac{10^2(10)^2 - 5^2(20)^2}{5^2 10^2}\right)} \quad (17)$$

$$= \frac{1}{1 + \frac{1}{2} \exp\left(\frac{100^2 - 100^2}{5^2 10^2}\right)} = \frac{1}{1 + \frac{1}{2} \exp(0)} = \frac{1}{1 + \frac{1}{2}} = \frac{2}{3} \quad (18)$$