

Solution & marking scheme

Name:
Roll No.:

MSO201: PROBABILITY & STATISTICS Quiz 2; Full Marks-20

[1] The joint p.d.f. of X_1 and X_2 is given by

$$f(x_1, x_2) = \frac{1}{4} \left(\frac{1}{2\pi |\Sigma_1|^{\frac{1}{2}}} e^{-\frac{1}{2} x' \Sigma_1^{-1} x} \right) + \frac{3}{4} \left(\frac{1}{2\pi |\Sigma_2|^{\frac{1}{2}}} e^{-\frac{1}{2} x' \Sigma_2^{-1} x} \right);$$

$$x = (x_1, x_2)' \in \mathbb{R}^2; \Sigma_1 = \begin{pmatrix} 1 & 0.5 \\ 0.5 & 1 \end{pmatrix}; \Sigma_2 = \begin{pmatrix} 1 & -0.5 \\ -0.5 & 1 \end{pmatrix}.$$

Prove or disprove the following statements:

- (a) $X_i \sim N(0,1)$ for $i = 1, 2$.
- (b) $\begin{pmatrix} X_1 \\ X_2 \end{pmatrix} \sim N_2$.
- (c) X_1 and X_2 are independent.
- (d) $\text{Cov}(X_1, X_2) = \frac{1}{4}$.

Useful Information: If $X \sim N_p(\mu, \Sigma)$ with $\Sigma > 0$, then p.d.f. of X is

$$f(x) = \frac{1}{(2\pi)^{\frac{p}{2}} |\Sigma|^{\frac{1}{2}}} e^{-\frac{1}{2}(x-\mu)' \Sigma^{-1}(x-\mu)}; x \in \mathbb{R}^p$$

10 marks (2+2+2+4)

[2] X_1, \dots, X_n is i.i.d. random sample from $U(0,1)$ distribution.

- (a) Find C such that $\frac{n}{\sum_{i=1}^n X_i^4} \xrightarrow{p} C$, as $n \rightarrow \infty$.
- (b) Find $\lim_{n \rightarrow \infty} P\left(\frac{1}{n} \sum_{i=1}^n X_i^2 \leq \frac{1}{3}\right)$.

10 marks (4+6)

$$(1) f(x_1, x_2) = \frac{1}{4} f_1(x_1, x_2) + \frac{3}{4} f_2(x_1, x_2)$$

$$(a) f_1(x_1, x_2) : \text{p.d.f. of } N_2(0, \Sigma_1)$$

$$f_2(x_1, x_2) : \text{p.d.f. of } N_2(0, \Sigma_2)$$

$$\begin{aligned} \text{Marginal of } X_1 : \int_{-\infty}^{\infty} f(x_1, x_2) dx_2 &= \frac{1}{4} \int_{-\infty}^{\infty} f_1(x_1, x_2) dx_2 + \frac{3}{4} \int_{-\infty}^{\infty} f_2(x_1, x_2) dx_2 \\ &= \frac{1}{4} (\text{p.d.f. of } N(0,1)) + \frac{3}{4} (\text{p.d.f. of } N(0,1)) \end{aligned}$$

(\because marginals of N_p are N_1)

$$\Rightarrow X_1 \sim N(0,1)$$

$$\text{Marginal of } X_2 : \int_{-\infty}^{\infty} f(x_1, x_2) dx_1 = \text{p.d.f. of } N(0,1)$$

$$\Rightarrow X_2 \sim N(0,1)$$

(a) is proved.

(2)

(b) $f(x_1, x_2)$ is not p.d.f. of N_2
 $\Rightarrow \begin{pmatrix} x_1 \\ x_2 \end{pmatrix} \neq N_2$ (b) is disproved (2)

(c) $f(x_1, x_2) \neq \underbrace{\left(\frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x_1^2} \right)}_{f(x_1)} \underbrace{\left(\frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x_2^2} \right)}_{f(x_2)}$ (2)
 $\Rightarrow X_1$ & X_2 are not independent
 (c) is disproved

(d) $EX_1 = 0$ & $EX_2 = 0$

$\Rightarrow \text{Cov}(X_1, X_2) = EX_1 X_2$

$= \int_{-1}^1 \int_{-1}^1 x_1 x_2 f(x_1, x_2) dx_1 dx_2$

$= \frac{1}{4} \int_{-1}^1 \int_{-1}^1 x_1 x_2 f_1(x_1, x_2) dx_1 dx_2 + \frac{3}{4} \int_{-1}^1 \int_{-1}^1 x_1 x_2 f_2(x_1, x_2) dx_1 dx_2$

$= \frac{1}{4} (0.5) + \frac{3}{4} (-0.5) = -\frac{1}{4}$ (d) is disproved
 (4) $\xrightarrow{\text{Cov}(X_1, X_2) \text{ for } f_1} \quad \xrightarrow{\text{Cov}(X_1, X_2) \text{ for } f_2}$

(2)

X_1, \dots, X_n i.i.d $U(0,1)$

(a) X_1^4, \dots, X_n^4 i.i.d. with $E X_1^4 = \int_0^1 x^4 dx = \frac{1}{5}$

$$WLLN \Rightarrow \frac{1}{n} \sum_{i=1}^n X_i^4 \xrightarrow{p} \frac{1}{5} \text{ as } n \rightarrow \infty \quad (2)$$

$$\Rightarrow \frac{n}{\sum_{i=1}^n X_i^4} \xrightarrow{p} 5 \text{ as } n \rightarrow \infty \quad (2)$$

(b) X_1^2, \dots, X_n^2 are i.i.d with $E X_1^2 = \frac{1}{3}$
& $V X_1^2 = E X_1^4 - (E X_1^2)^2 = \frac{4}{45}$

By CLT,

$$\sqrt{n} \left(\frac{1}{n} \sum_{i=1}^n X_i^2 - \frac{1}{3} \right) \xrightarrow{d} N\left(0, \frac{4}{45}\right) \quad - (2)$$

So

$$\begin{aligned} & P\left(\frac{1}{n} \sum_{i=1}^n X_i^2 \leq \frac{1}{3}\right) \\ &= P\left(\frac{1}{n} \sum_{i=1}^n X_i^2 - \frac{1}{3} \leq 0\right) \\ &= P\left(\frac{\sqrt{n} \left(\frac{1}{n} \sum_{i=1}^n X_i^2 - \frac{1}{3}\right)}{\sqrt{\frac{4}{45}}} \leq 0\right) \end{aligned}$$

$$\rightarrow \Phi(0) = \frac{1}{2} \text{ as } n \rightarrow \infty$$

(4)