DIGITAL COMMUNICATIONS

V.GOKULKUMAR

velicharla@outlook.com
IITH Future Wireless Communication (FWC)

ASSIGNMENT

January 5, 2023

1 Random numbers

FWC22034

1.1 Uniform Random Numbers

Let U be a uniform random variable between 0 and 1.

 $1.1\,$ Generate 10^6 samples of U using a C program and save into a file called uni.dat .

https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-5/5.1.1/uni.dat

1.2 Load the uni.dat file into python and plot the empirical CDF of U using the samples in uni.dat. The CDF is defined as

$$F_U(x) = \Pr\left(U \le x\right) \tag{1}$$

Solution: The following code plots Fig. ??

https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-5/5.1.2/cdf_plot.py

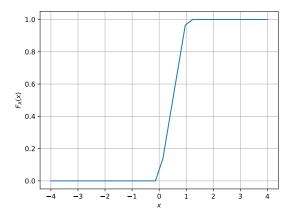


Figure 1: CDF

1.3 Find a theoretical expression for $F_U(x)$.

Uniform Random variable:

$$f_U(x) = \frac{1}{b-a} \quad a \le x \le b \\ = 0 \quad \text{elsewhere} \quad \begin{cases} b \text{ are real} \\ \text{constants} \\ -\alpha < a < \alpha \\ 8tb > a \end{cases}$$

$$F_U(x) = \int_{-\infty}^x f_x(x)dx$$
$$= \int_{-\infty}^x \frac{1}{b-a} \cdot dx$$
$$= \frac{1}{b-a} \cdot x \Big|_a^x$$
$$= \frac{x-a}{b-a}$$

$$F_U(x) = \begin{cases} 0 & x < a \\ (x - a)/(b - a) & a \le x < b \\ 1 & b \le x \end{cases}$$

1.4 The mean of \boldsymbol{U} is defined as

$$E[U] = \frac{1}{N} \sum_{i=1}^{N} U_i$$
 (2)

and its variance as

$$var[U] = E[U - E[U]]^{2}$$
 (3)

Write a C program to find the mean and variance of $U. \label{eq:U.def}$

```
#include <stdio.h>
#include <stdlib.h>
#include <math.h>
#include "coeffs.h"

int main(void) //main function begins
{
    //Uniform random numbers
    uniform("uni.dat", 1000000);

    //Mean of uniform
    printf("%lf\n",mean("uni.dat"));
    //Variance of uniform
    printf("%lf",variance("uni.dat"));
    return 0;
}
```

```
double mean(char *str)
{
  int i=0,c;
  FILE *fp;
  double x, temp=0.0;

  fp = fopen(str,"r");
  //get numbers from file
  while(fscanf(fp,"%lf",&x)!=EOF)
  {
  //Count numbers in file
  i=i+1;
  //Add all numbers in file
  temp = temp+x;
  }
  fclose(fp);
  temp = temp/(i-1);
  return temp;
}
```

```
double variance(char *str)
int i=0,j=0,c;
FILE *fp;
double x, temp=0.0, value, sumsqr=0, variance=0.0;
fp = fopen(str,"r");
//get numbers from file
while(fscanf(fp,"%lf",&x)!=EOF)
//Count numbers in file
i=i+1;
//Add all numbers in file
temp = temp + x;
fclose(fp):
temp = temp/(i-1);
fp = fopen(str,"r");
while(fscanf(fp,"%|f",&x)!=EOF)
 j=j+1;
 value=x—temp;
     sumsqr=sumsqr+value*value;
fclose(fp);
variance = sumsqr/(j-1);
return variance;
```

Result:

```
mean=0.500137
variance=0.83251
```

https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-5/5.1.4/uniform.c

1.5 Verify your result theoretically given that

$$E\left[U^{k}\right] = \int_{-\infty}^{\infty} x^{k} dF_{U}(x) \tag{4}$$

$$Mean : \mathbf{E}[U] = \int_{a}^{b} x \cdot \frac{1}{b-a} dx = \frac{b^{2} - a^{2}}{2} \cdot \frac{1}{b-a}$$

$$= \frac{(b-a)(b+a)}{2} \cdot \frac{1}{b-a}$$

$$= \frac{b+a}{2}$$

$$(7)$$

$$Here \ a = 0, b = 1$$

$$(8)$$

$$\mu = \mathbf{E}[X] = \frac{1}{2} = 0.5$$

$$(9)$$

$$\mathbf{E}[U^{2}] = \int_{a}^{b} x^{2} \cdot \frac{1}{b-a} dx = \frac{b^{3} - a^{3}}{3} \cdot \frac{1}{b-a}$$

$$(10)$$

$$= \frac{a^{2} + ab + b^{2}}{3}$$

$$(11)$$

$$Variance : \sigma^{2} = E(U^{2}) - [E(U)]^{2} = \frac{(a-b)^{2}}{12}$$

$$(12)$$

$$\sigma^{2} = 0.834$$

$$(13)$$

1.2 Central Limit Theorem

1. Generate 10^6 samples of the random variable

$$X = \sum_{i=1}^{12} U_i - 6 \tag{14}$$

using a C program, where $U_i, i=1,2,\ldots,12$ are a set of independent uniform random variables between 0 and 1 and save in a file called gau.dat

https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-5/5.2.1/gau.dat

2. Load gau.dat in python and plot the empirical CDF of X using the samples in gau.dat. What properties does a CDF have?

Solution: The CDF of X is plotted in Fig.

https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-5/5.2.2/cdf_plot.py

3. Load gau.dat in python and plot the empirical PDF of X using the samples in gau.dat. The PDF of X is defined as

$$p_X(x) = \frac{d}{dx} F_X(x) \tag{15}$$

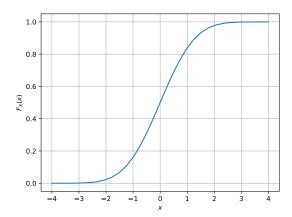


Figure 2: CDF of X

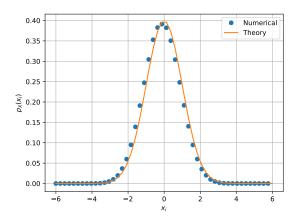


Figure 3: The PDF of \boldsymbol{X}

What properties does the PDF have? **Solution:** The PDF of X is plotted in Fig. 3 using the code below

```
https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-5/5.2.3/pdf_plot.py
```

4. Find the mean and variance of \boldsymbol{X} by writing a C program.

```
#include <stdio.h>
#include <stdib.h>
#include <math.h>
#include "coeffs.h"

int main(void) //main function begins
{
   //gaussian random numbers
   gaussian("gau.dat", 1000000);

   //Mean ,variance of gaussian
   printf("%lf\n",mean("gau.dat"));
   printf("%lf",variance("gau.dat"));
   return 0;
}
```

mean=-0.000283 variance=0.999702

https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-5/5.2.4/gauss.c

5. Given that

$$p_X(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right), -\infty < x < \infty, \quad (16)$$

repeat the above exercise theoretically.

$$\begin{split} E(X) &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} x e^{-\frac{x^2}{2}} dx \\ &= 0 \quad \text{(odd function)} \\ E\left(X^2\right) &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} x^2 e^{-\frac{x^2}{2}} dx \quad \text{(evenfunction)} \\ &= \frac{2}{\sqrt{2\pi}} \int_{0}^{\infty} x^2 e^{-\frac{x^2}{2}} dx \\ &= \frac{2}{\sqrt{2\pi}} \int_{0}^{\infty} \sqrt{2u} e^{-u} du \quad \left(\text{ Let } \frac{x^2}{2} = u \right) \\ &= \frac{2}{\sqrt{\pi}} \int_{0}^{\infty} e^{-u} u^{\frac{3}{2} - 1} du \\ &= \frac{2}{\sqrt{\pi}} \Gamma\left(\frac{3}{2}\right) \\ &= \frac{1}{\sqrt{\pi}} \Gamma\left(\frac{1}{2}\right) \end{split}$$

where we have used the fact that

$$: \Gamma(n) = (n-1)\Gamma(n-1); \Gamma\left(\frac{1}{2}\right) = \sqrt{\pi}$$

Thus, the variance is

$$\sigma^2 = E(X)^2 - E^2(X) = 1$$

1.3 From Uniform to Other

1. Generate samples of

$$V = -2\ln(1 - U) \tag{17}$$

and plot its CDF.

2. Find a theoretical expression for ${\cal F}_V(x)$. The CDF of ${\cal V}$ is defined as

$$F_V(v) = pr(V \le v) \tag{18}$$

$$= pr(-2\ln(1-U) \le v)$$
 (19)

$$= pr(\ln(1-U) \ge -\frac{v}{2})$$
 (20)

$$= pr(1 - U \ge \exp\left(-\frac{v}{2}\right)) \tag{21}$$

$$= pr(U \le 1 - \exp\left(-\frac{v}{2}\right)) \tag{22}$$

$$=F_U\left(1-\exp\left(-\frac{v}{2}\right)\right) \tag{23}$$

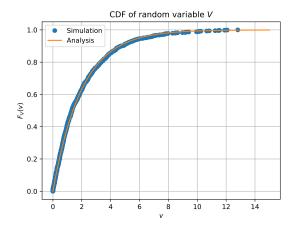


Figure 4: CDF of V

$$F_U(x) = \begin{cases} 0 & x < 0 \\ x & 0 \le x \le 1 \\ 1 & x > 1 \end{cases}$$
 (24)

Substituting the above in (23),

$$F_{U}\left(1 - \exp\left(-\frac{v}{2}\right)\right) = \begin{cases} 0 & 1 - \exp\left(-\frac{v}{2}\right) < 0\\ 1 - \exp\left(-\frac{v}{2}\right) & 0 \le 1 - \exp\left(-\frac{v}{2}\right) \le 1 \end{cases}$$
(25)
$$1 & 1 - \exp\left(-\frac{v}{2}\right) > 1$$

After some algebra, the above conditions yield

$$F_V(v) = \begin{cases} 0 & v < 0\\ 1 - exp\left(-\frac{v}{2}\right) & v \ge 0 \end{cases}$$
 (26)

which is the CDF of the exponential distribution with parameter $\frac{1}{2}$.

1.4 Triangular Distribution

1.1 Generate

$$T = U_1 + U_2 (27)$$

- 1.2 Find the CDF of T.
- 1.3 Find the PDF of T.
- 1.4 Find the theoretical expressions for the PDF and CDF of ${\cal T}.$
- 1.5 Verify your results through a plot.

2 transfromations of R.V

2.1 Gaussian to Other

1. Let $X_1 \sim (0,1)$ and $X_2 \sim (0,1)$. Plot the CDF and PDF of

$$V = X_1^2 + X_2^2 \tag{28}$$

https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-7/7.1.1/7.1.1_CDF.py

https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-7/7.1.1/7.1.1_PDF.py

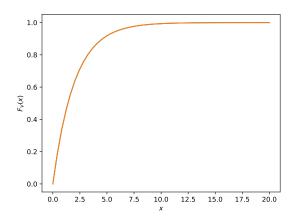


Figure 5: CDF of V

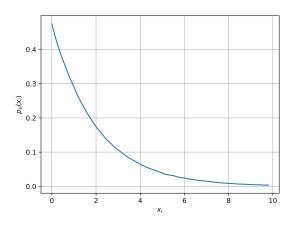


Figure 6: PDF of V

2. If
$$F_V(x) = \begin{cases} 1 - e^{-\alpha x} & x \ge 0 \\ 0 & x < 0, \end{cases} \tag{29}$$

find $\alpha.$ For the value $\alpha=0.5,$ the theory matches the simulation. The following code generates the CDF of V in Fig. Fig.

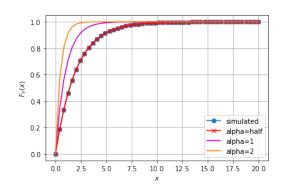


Figure 7: CDF of V

3. Plot the CDF and PDf of

$$A = \sqrt{V} \tag{30}$$

The CDF and PDF of A are plotted in Figs. ?? and ?? using the codes below.

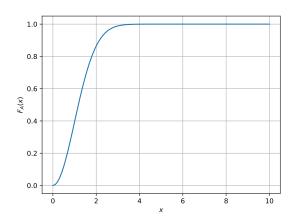


Figure 8: CDF

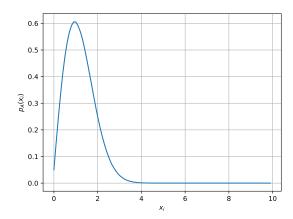


Figure 9: PDF

https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-7/7.1.3/7.1.3_CDF.py

https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-7/7.1.3/7.1.3_PDF.py

2.2 Conditional Probability

1. Plot

$$P_e = \Pr\left(\hat{X} = -1|X = 1\right) \tag{31}$$

for

$$Y = AX + N, (32)$$

where A is Raleigh with $E\left[A^2\right]=\gamma, N\sim\mathcal{N}(0,1)\in(-1,1)$ for $0\leq\gamma\leq10$ dB.

2. Assuming that N is a constant, find an expression for P_e . Call this $P_e(N)$.

The estimated value \hat{X} is given by

$$\hat{X} = \begin{cases} +1 & Y > 0 \\ -1 & Y < 0 \end{cases} \tag{33}$$

For X=1,

$$Y = A + N \tag{34}$$

$$P_e = \Pr\left(\hat{X} = -1|X = 10\right)$$
 (35)

$$= \Pr(Y < 0 | X = 1) \tag{36}$$

$$=\Pr\left(A<-N\right)\tag{37}$$

$$=F_A(-N) \tag{38}$$

$$= \int_{-\infty}^{-N} f_A(x) dx \tag{39}$$

By definition

$$f_A(x) = \begin{cases} \frac{x}{\sigma^2} \exp\left(-\frac{x^2}{2\sigma^2}\right) & x \ge 0\\ 0 & otherwise \end{cases}$$
 (40)

If $N > 0, f_A(x) = 0$. Then,

$$P_e = 0 (41)$$

If N < 0. Then,

$$P_e(N) = \int_{-\infty}^{-N} f_A(x) dx \tag{42}$$

$$= \int_{-\infty}^{0} 0 dx + \int_{0}^{-N} f_A(x) dx$$
 (43)

$$= \int_0^{-N} \frac{x}{\sigma^2} \exp\left(-\frac{x^2}{2\sigma^2}\right) dx \qquad (44)$$

$$=1-\exp\left(-\frac{N^2}{2\sigma^2}\right) \tag{45}$$

Therefore.

$$P_e(N) = \begin{cases} 1 - \exp\left(-\frac{N^2}{2\sigma^2}\right) & N < 0\\ 0 & otherwise \end{cases} \tag{46}$$

3. For a function g,

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)p_X(x) dx \tag{47}$$

Find $P_e = E[P_e(N)]$. Since $N \sim \mathcal{N}(0, 1)$,

$$p_N(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) \tag{48}$$

(49)

And from (46)

$$P_e(x) = \begin{cases} 1 - \exp\left(-\frac{x^2}{2\sigma^2}\right) & x < 0\\ 0 & otherwise \end{cases}$$
 (50)

$$P_e = E[P_e(N)] = \int_{-\infty}^{\infty} P_e(x) p_N(x) dx$$
 (51)

If $x<0, P_e(x)=0$ and using the fact that for an even function

$$\int_{-\infty}^{\infty} f(x) = 2 \int_{-\infty}^{0} f(x) \tag{52}$$

we get

$$P_e = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^0 \exp\left(-\frac{x^2}{2}\right) \left(1 - \exp\left(-\frac{x^2}{2\sigma^2}\right)\right) dx$$
(53)

$$=\frac{1}{2\sqrt{2\pi}}\int_{-\infty}^{\infty}\exp\left(-\frac{x^2}{2}\right)dx$$

$$-\frac{1}{2\sqrt{2\pi}} \int_{-\infty}^{\infty} \exp\left(-\frac{(1+\sigma^2)x^2}{2\sigma^2}\right) dx \tag{54}$$

$$= \frac{\sqrt{2\pi} - \sqrt{\frac{\pi(2\sigma^2)}{1+\sigma^2}}}{2\sqrt{2\pi}}$$
 (55)

$$=\frac{1}{2} - \frac{1}{2}\sqrt{\frac{\sigma^2}{1+\sigma^2}}\tag{56}$$

For a Rayleigh Distribution with scale $= \sigma$,

$$E\left[A^2\right] = 2\sigma^2 \tag{57}$$

$$\gamma = 2\sigma^2 \tag{58}$$

$$\therefore P_e = \frac{1}{2} - \frac{1}{2} \sqrt{\frac{\gamma}{2+\gamma}} \tag{59}$$

4. Plot P_e in problems 1 and 3 on the same graph w.r.t $\gamma.$ Comment.

https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-7/7.2/7.2.py

