

# DIGITAL COMMUNICATIONS

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IITH Future Wireless Communication (FWC)

ASSIGNMENT

January 5, 2023

## 1 Random numbers

1.4 The mean of  $U$  is defined as

$$E[U] = \frac{1}{N} \sum_{i=1}^N U_i \quad (2)$$

### 1.1 Uniform Random Numbers

Let  $U$  be a uniform random variable between 0 and 1.

- 1.1 Generate  $10^6$  samples of  $U$  using a C program and save into a file called uni.dat .

<https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-5/5.1.1/uni.dat>

- 1.2 Load the uni.dat file into python and plot the empirical CDF of  $U$  using the samples in uni.dat. The CDF is defined as

$$F_U(x) = \Pr(U \leq x) \quad (1)$$

**Solution:** The following code plots Fig. ??

[https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-5/5.1.2/cdf\\_plot.py](https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-5/5.1.2/cdf_plot.py)

images/uni\_cdf.png

Figure 1: CDF

- 1.3 Find a theoretical expression for  $F_U(x)$ .

Uniform Random variable:

$$f_U(x) = \begin{cases} \frac{1}{b-a} & a \leq x \leq b \\ 0 & \text{elsewhere} \end{cases} \quad \left. \begin{array}{l} \text{where } a \text{ and } b \text{ are real} \\ \text{constants} \\ -\alpha < a < \alpha \\ \& b > a \end{array} \right\}$$

$$\begin{aligned} F_U(x) &= \int_{-\infty}^x f_x(x) dx \\ &= \int_{-\infty}^x \frac{1}{b-a} \cdot dx \\ &= \frac{1}{b-a} \cdot x \Big|_a^x \\ &= \frac{x-a}{b-a} \end{aligned}$$

$$F_U(x) = \begin{cases} 0 & x < a \\ (x-a)/(b-a) & a \leq x < b \\ 1 & b \leq x \end{cases}$$

and its variance as

$$\text{var}[U] = E[U - E[U]]^2 \quad (3)$$

Write a C program to find the mean and variance of  $U$ .

```
#include <stdio.h>
#include <stdlib.h>
#include <math.h>
#include "coeffs.h"

int main(void) //main function begins
{

//Uniform random numbers
uniform("uni.dat", 1000000);

//Mean of uniform
printf("%lf\n", mean("uni.dat"));
//Variance of uniform
printf("%lf", variance("uni.dat"));
return 0;
}
```

```
double mean(char *str)
{
int i=0,c;
FILE *fp;
double x, temp=0.0;

fp = fopen(str,"r");
//get numbers from file
while(fscanf(fp,"%lf",&x)!=EOF)
{
//Count numbers in file
i=i+1;
//Add all numbers in file
temp = temp+x;
}
fclose(fp);
temp = temp/(i-1);
return temp;
}
```

```
double variance(char *str)
{
int i=0,j=0,c;
FILE *fp;
double x, temp=0.0,value,sumsq=0,variance=0.0;

fp = fopen(str,"r");
//get numbers from file
while(fscanf(fp,"%lf",&x)!=EOF)
{
//Count numbers in file
i=i+1;
//Add all numbers in file
temp = temp+x;
}
fclose(fp);
temp = temp/(i-1);
fp = fopen(str,"r");
while(fscanf(fp,"%lf",&x)!=EOF)
{
j=j+1;
value=x-temp;
sumsq=sumsq+value*value;
}
fclose(fp);
variance = sumsq/(j-1);
return variance;
}
```

Result:

```
mean=0.500137
variance=0.83251
```

<https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-5/5.1.4/uniform.c>

1.5 Verify your result theoretically given that

$$E[U^k] = \int_{-\infty}^{\infty} x^k dF_U(x) \quad (4)$$

images/gau\_cdf.pdf

Figure 2: CDF of X

$$\text{Mean : } E[U] = \int_a^b x \cdot \frac{1}{b-a} dx = \frac{b^2 - a^2}{2} \cdot \frac{1}{b-a} \quad (5)$$

$$= \frac{(b-a)(b+a)}{2} \cdot \frac{1}{b-a} \quad (6)$$

$$= \frac{b+a}{2} \quad (7)$$

$$\text{Here } a = 0, b = 1 \quad (8)$$

$$\mu = E[X] = \frac{1}{2} = 0.5 \quad (9)$$

$$E[U^2] = \int_a^b x^2 \cdot \frac{1}{b-a} dx = \frac{b^3 - a^3}{3} \cdot \frac{1}{b-a} \quad (10)$$

$$= \frac{a^2 + ab + b^2}{3} \quad (11)$$

$$\text{Variance : } \sigma^2 = E(U^2) - [E(U)]^2 = \frac{(a-b)^2}{12} \quad (12)$$

$$\sigma^2 = 0.834 \quad (13)$$

## 1.2 Central Limit Theorem

1. Generate  $10^6$  samples of the random variable

$$X = \sum_{i=1}^{12} U_i - 6 \quad (14)$$

using a C program, where  $U_i, i = 1, 2, \dots, 12$  are a set of independent uniform random variables between 0 and 1 and save in a file called gau.dat

<https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-5/5.2.1/gau.dat>

2. Load gau.dat in python and plot the empirical CDF of  $X$  using the samples in gau.dat. What properties does a CDF have?

**Solution:** The CDF of  $X$  is plotted in Fig.

[https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-5/5.2.2/cdf\\_plot.py](https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-5/5.2.2/cdf_plot.py)

3. Load gau.dat in python and plot the empirical PDF of  $X$  using the samples in gau.dat. The PDF of  $X$  is

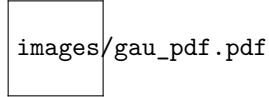


Figure 3: The PDF of  $X$

defined as

$$p_X(x) = \frac{d}{dx} F_X(x) \quad (15)$$

What properties does the PDF have?

**Solution:** The PDF of  $X$  is plotted in Fig. ?? using the code below

```
https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-5/5.2.3/pdf\_plot.py
```

4. Find the mean and variance of  $X$  by writing a C program.

```
#include <stdio.h>
#include <stdlib.h>
#include <math.h>
#include "coeffs.h"

int main(void) //main function begins
{

//gaussian random numbers
gaussian("gau.dat", 1000000);

//Mean ,variance of gaussian
printf("%lf\n",mean("gau.dat"));
printf("%lf",variance("gau.dat"));
return 0;
}
```

```
mean=-0.000283
variance=0.999702
```

```
https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-5/5.2.4/gauss.c
```

5. Given that

$$p_X(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right), -\infty < x < \infty, \quad (16)$$

repeat the above exercise theoretically.

$$\begin{aligned} E(X) &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} x e^{-\frac{x^2}{2}} dx \\ &= 0 \quad (\text{odd function}) \\ E(X^2) &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} x^2 e^{-\frac{x^2}{2}} dx \quad (\text{evenfunction}) \\ &= \frac{2}{\sqrt{2\pi}} \int_0^{\infty} x^2 e^{-\frac{x^2}{2}} dx \\ &= \frac{2}{\sqrt{2\pi}} \int_0^{\infty} \sqrt{2u} e^{-u} du \quad \left( \text{Let } \frac{x^2}{2} = u \right) \\ &= \frac{2}{\sqrt{\pi}} \int_0^{\infty} e^{-u} u^{\frac{3}{2}-1} du \\ &= \frac{2}{\sqrt{\pi}} \Gamma\left(\frac{3}{2}\right) \\ &= \frac{1}{\sqrt{\pi}} \Gamma\left(\frac{1}{2}\right) \\ &= 1 \end{aligned}$$

where we have used the fact that

$$\Gamma(n) = (n-1)\Gamma(n-1); \Gamma\left(\frac{1}{2}\right) = \sqrt{\pi}$$

Thus, the variance is

$$\sigma^2 = E(X)^2 - E^2(X) = 1$$

### 1.3 From Uniform to Other

1. Generate samples of

$$V = -2 \ln(1 - U) \quad (17)$$

and plot its CDF.

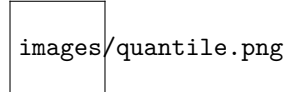


Figure 4: CDF

2. Find a theoretical expression for  $F_V(x)$ . The CDF of  $V$  is defined as

$$F_V(v) = \text{pr}(V \leq v) \quad (18)$$

$$= \text{pr}(-2 \ln(1 - U) \leq v) \quad (19)$$

$$= \text{pr}(\ln(1 - U) \geq -\frac{v}{2}) \quad (20)$$

$$= \text{pr}(1 - U \geq \exp\left(-\frac{v}{2}\right)) \quad (21)$$

$$= \text{pr}(U \leq 1 - \exp\left(-\frac{v}{2}\right)) \quad (22)$$

$$= F_U\left(1 - \exp\left(-\frac{v}{2}\right)\right) \quad (23)$$

$$F_U(x) = \begin{cases} 0 & x < 0 \\ x & 0 \leq x \leq 1 \\ 1 & x > 1 \end{cases} \quad (24)$$

Substituting the above in (??),

$$F_U \left( 1 - \exp \left( -\frac{v}{2} \right) \right) = \begin{cases} 0 & 1 - \exp \left( -\frac{v}{2} \right) < 0 \\ 1 - \exp \left( -\frac{v}{2} \right) & 0 \leq 1 - \exp \left( -\frac{v}{2} \right) \leq 1 \\ 1 & 1 - \exp \left( -\frac{v}{2} \right) > 1 \end{cases} \quad (25)$$

After some algebra, the above conditions yield

$$F_V(v) = \begin{cases} 0 & v < 0 \\ 1 - \exp \left( -\frac{v}{2} \right) & v \geq 0 \end{cases} \quad (26)$$

which is the CDF of the exponential distribution with parameter  $\frac{1}{2}$ .

## 1.4 Triangular Distribution

1.1 Generate

$$T = U_1 + U_2 \quad (27)$$

1.2 Find the CDF of  $T$ .

1.3 Find the PDF of  $T$ .

1.4 Find the theoretical expressions for the PDF and CDF of  $T$ .

1.5 Verify your results through a plot.

## 2 transformations of R.V

### 2.1 Gaussian to Other

1. Let  $X_1 \sim (0, 1)$  and  $X_2 \sim (0, 1)$ . Plot the CDF and PDF of

$$V = X_1^2 + X_2^2 \quad (28)$$

[https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-7/7.1.1/7.1.1\\_CDF.py](https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-7/7.1.1/7.1.1_CDF.py)

[https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-7/7.1.1/7.1.1\\_PDF.py](https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-7/7.1.1/7.1.1_PDF.py)

2. If

$$F_V(x) = \begin{cases} 1 - e^{-\alpha x} & x \geq 0 \\ 0 & x < 0, \end{cases} \quad (29)$$

find  $\alpha$ .

For the value  $\alpha = 0.5$ , the theory matches the simulation. The following code generates the CDF of  $V$  in Fig. Fig.

3. Plot the CDF and PDF of

$$A = \sqrt{V} \quad (30)$$

The CDF and PDF of  $A$  are plotted in Figs. ?? and ?? using the codes below.

[https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-7/7.1.3/7.1.3\\_CDF.py](https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-7/7.1.3/7.1.3_CDF.py)

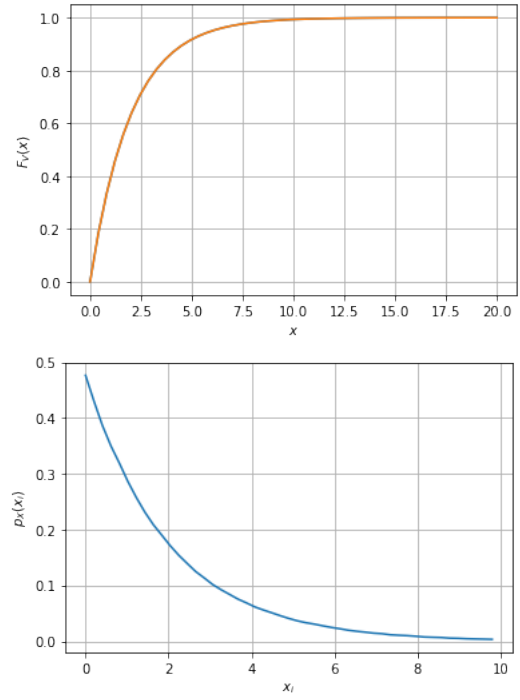


Figure 5: CDF,PDF of  $V$

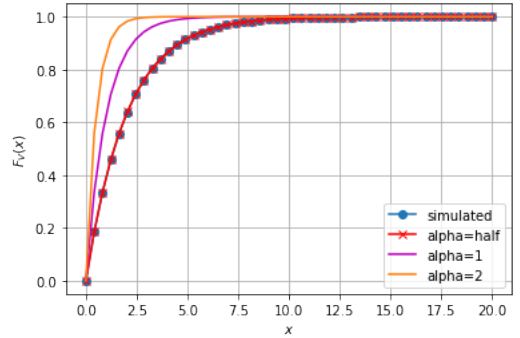


Figure 6: CDF of  $V$

[https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-7/7.1.3/7.1.3\\_PDF.py](https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-7/7.1.3/7.1.3_PDF.py)

### 2.2 Conditional Probability

1. Plot

$$P_e = \Pr(\hat{X} = -1 | X = 1) \quad (31)$$

for

$$Y = AX + N, \quad (32)$$

where  $A$  is Raleigh with  $E[A^2] = \gamma$ ,  $N \sim \mathcal{N}(0, 1) \in (-1, 1)$  for  $0 \leq \gamma \leq 10$  dB.

2. Assuming that  $N$  is a constant, find an expression for  $P_e$ . Call this  $P_e(N)$ .

The estimated value  $\hat{X}$  is given by

$$\hat{X} = \begin{cases} +1 & Y > 0 \\ -1 & Y < 0 \end{cases} \quad (33)$$

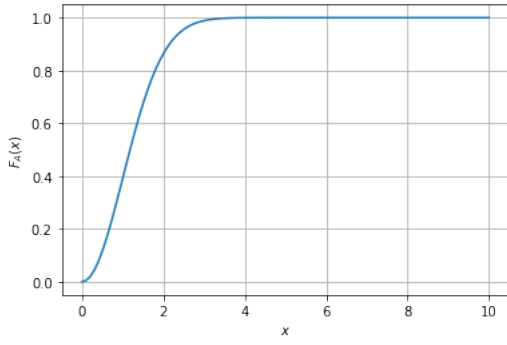


Figure 7: CDF

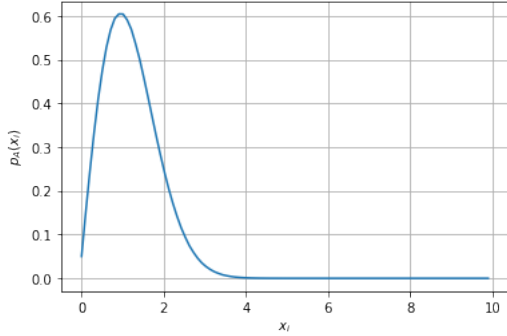


Figure 8: PDF

For  $X = 1$ ,

$$Y = A + N \quad (34)$$

$$P_e = \Pr(\hat{X} = -1 | X = 10) \quad (35)$$

$$= \Pr(Y < 0 | X = 1) \quad (36)$$

$$= \Pr(A < -N) \quad (37)$$

$$= F_A(-N) \quad (38)$$

$$= \int_{-\infty}^{-N} f_A(x) dx \quad (39)$$

By definition

$$f_A(x) = \begin{cases} \frac{x}{\sigma^2} \exp\left(-\frac{x^2}{2\sigma^2}\right) & x \geq 0 \\ 0 & \text{otherwise} \end{cases} \quad (40)$$

If  $N > 0$ ,  $f_A(x) = 0$ . Then,

$$P_e = 0 \quad (41)$$

If  $N < 0$ . Then,

$$P_e(N) = \int_{-\infty}^{-N} f_A(x) dx \quad (42)$$

$$= \int_{-\infty}^0 0 dx + \int_0^{-N} f_A(x) dx \quad (43)$$

$$= \int_0^{-N} \frac{x}{\sigma^2} \exp\left(-\frac{x^2}{2\sigma^2}\right) dx \quad (44)$$

$$= 1 - \exp\left(-\frac{N^2}{2\sigma^2}\right) \quad (45)$$

Therefore,

$$P_e(N) = \begin{cases} 1 - \exp\left(-\frac{N^2}{2\sigma^2}\right) & N < 0 \\ 0 & \text{otherwise} \end{cases} \quad (46)$$

3. For a function  $g$ ,

$$E[g(X)] = \int_{-\infty}^{\infty} g(x) p_X(x) dx \quad (47)$$

Find  $P_e = E[P_e(N)]$ .

Since  $N \sim \mathcal{N}(0, 1)$ ,

$$p_N(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) \quad (48)$$

$$(49)$$

And from (??)

$$P_e(x) = \begin{cases} 1 - \exp\left(-\frac{x^2}{2\sigma^2}\right) & x < 0 \\ 0 & \text{otherwise} \end{cases} \quad (50)$$

$$P_e = E[P_e(N)] = \int_{-\infty}^{\infty} P_e(x) p_N(x) dx \quad (51)$$

If  $x < 0$ ,  $P_e(x) = 0$  and using the fact that for an even function

$$\int_{-\infty}^{\infty} f(x) = 2 \int_{-\infty}^0 f(x) \quad (52)$$

we get

$$P_e = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^0 \exp\left(-\frac{x^2}{2}\right) \left(1 - \exp\left(-\frac{x^2}{2\sigma^2}\right)\right) dx \quad (53)$$

$$= \frac{1}{2\sqrt{2\pi}} \int_{-\infty}^{\infty} \exp\left(-\frac{x^2}{2}\right) dx - \frac{1}{2\sqrt{2\pi}} \int_{-\infty}^{\infty} \exp\left(-\frac{(1+\sigma^2)x^2}{2}\right) dx \quad (54)$$

$$= \frac{\sqrt{2\pi} - \sqrt{\frac{\pi(2\sigma^2)}{1+\sigma^2}}}{2\sqrt{2\pi}} \quad (55)$$

$$= \frac{1}{2} - \frac{1}{2} \sqrt{\frac{\sigma^2}{1+\sigma^2}} \quad (56)$$

For a Rayleigh Distribution with scale  $= \sigma$ ,

$$E[A^2] = 2\sigma^2 \quad (57)$$

$$\gamma = 2\sigma^2 \quad (58)$$

$$\therefore P_e = \frac{1}{2} - \frac{1}{2} \sqrt{\frac{\gamma}{2+\gamma}} \quad (59)$$

4. Plot  $P_e$  in problems ?? and ?? on the same graph w.r.t  $\gamma$ . Comment.

<https://github.com/velicharlagokulkumar/digital-communications/blob/main/codes/chapter-7/7.2/7.2.py>

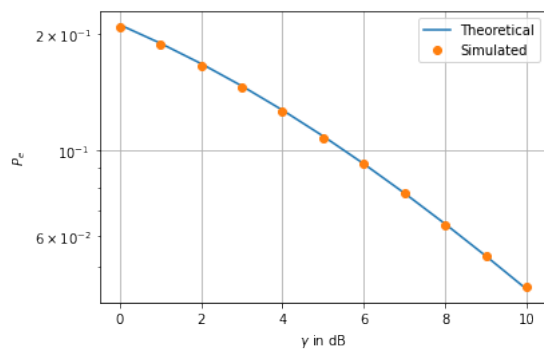


Figure 9: PDF