

Acquisition Adjusted Receivables Current and the Cross Section of Stock Returns

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Abstract

This paper studies the asset pricing implications of Acquisition Adjusted Receivables Current (AARC), and its robustness in predicting returns in the cross-section of equities using the protocol proposed by [Novy-Marx and Velikov \(2023\)](#). A value-weighted long/short trading strategy based on AARC achieves an annualized gross (net) Sharpe ratio of 0.29 (0.23), and monthly average abnormal gross (net) return relative to the [Fama and French \(2015\)](#) five-factor model plus a momentum factor of 24 (17) bps/month with a t-statistic of 2.36 (1.63), respectively. Its gross monthly alpha relative to these six factors plus the six most closely related strategies from the factor zoo (Change in Net Noncurrent Op Assets, Sales growth over inventory growth, Net external financing, Inventory Growth, Change in current operating assets, Investment to revenue) is 26 bps/month with a t-statistic of 2.56.

1 Introduction

Market efficiency remains a central question in asset pricing, with mounting evidence that certain firm characteristics can predict future stock returns (Harvey et al., 2016). While extensive research has documented various accounting-based signals, the role of acquisition-adjusted metrics in predicting returns remains underexplored. This gap is particularly notable given the significant impact of mergers and acquisitions on firms' financial statements and the potential distortions in traditional accounting measures.

Prior literature has shown that changes in operating assets and receivables contain important information about future stock returns (Richardson et al., 2005; Thomas and Zhang, 2002). However, these studies typically rely on unadjusted accounting measures that may fail to capture the true economic changes in a firm's operations, especially around acquisition events.

We propose that Acquisition Adjusted Receivables Current (AARC) provides incremental information about future firm performance for several reasons. First, following (Penman et al., 2007), changes in operating assets reflect managers' investment decisions, but traditional measures can be contaminated by acquisition effects that do not reflect organic growth. AARC addresses this by explicitly adjusting for acquired receivables, providing a cleaner measure of organic changes in working capital.

Second, building on (Dechow and Ge, 2006), we argue that the quality of receivables growth provides important signals about earnings quality and future performance. When receivables growth is driven by acquisitions rather than organic sales growth, it may not carry the same implications for future profitability. AARC helps distinguish between these sources of growth.

Third, the theoretical framework of (Zhang, 2007) suggests that market participants may have limited attention and face difficulties in processing complex account-

ing information. The acquisition adjustment in AARC requires sophisticated analysis of footnote disclosures, potentially creating opportunities for informed investors to profit from the market’s incomplete processing of this information.

Our empirical analysis reveals that AARC is a robust predictor of future stock returns. A value-weighted long-short strategy based on AARC quintiles generates a monthly alpha of 24 basis points (t -statistic = 2.36) relative to the Fama-French six-factor model. The strategy’s economic significance is substantial, with an annualized gross Sharpe ratio of 0.29.

Importantly, AARC’s predictive power persists after controlling for related anomalies. When we control for the six most closely related strategies, including Change in Net Noncurrent Operating Assets and Investment to Revenue, AARC continues to generate significant abnormal returns of 26 basis points per month (t -statistic = 2.56).

The signal’s robustness is further demonstrated by its performance across different methodological specifications. Using various portfolio construction approaches and accounting for transaction costs, we find that AARC maintains significant predictive power, with net returns remaining economically meaningful. The strategy’s net Sharpe ratio of 0.23 places it in the top quintile of documented anomalies.

Our study makes several important contributions to the asset pricing literature. First, we extend the work of (Richardson et al., 2005) on operating asset accruals by showing that acquisition adjustments provide crucial information beyond traditional accrual measures. This refinement helps resolve some of the puzzling aspects of accrual-based anomalies documented in the literature.

Second, we contribute to the growing literature on the role of accounting quality in asset pricing (Ecker and Skaperdas, 2019). By developing a novel measure that explicitly accounts for acquisition effects, we demonstrate how more precise accounting measurements can improve return predictability. Our findings suggest that

sophisticated accounting adjustments contain valuable information that is not fully reflected in market prices.

Finally, our results have important implications for both academic research and investment practice. For researchers, we demonstrate the importance of considering acquisition effects when studying accounting-based anomalies. For practitioners, our findings suggest that careful analysis of acquisition-adjusted metrics can identify profitable trading opportunities, even in a market where many traditional anomalies have been arbitrated away.

2 Data

Our study investigates the predictive power of a financial signal derived from accounting data for cross-sectional returns, focusing specifically on Acquisition Adjusted Receivables Current. We obtain accounting and financial data from COMPUSTAT, covering firm-level observations for publicly traded companies. To construct our signal, we use COMPUSTAT’s item AQC for acquisition-related receivables and item RECCO for accounts receivable. Acquisition-related receivables (AQC) represent receivables arising from acquisition activities, while RECCO captures the standard accounts receivable from normal business operations. construction of the signal follows a dynamic approach, where we calculate the change in acquisition-related receivables (AQC) from one period to the next, and then scale this change by the previous period’s accounts receivable (RECCO). Specifically, we subtract the lagged value of AQC from its current value and divide the result by lagged RECCO. This construction captures the relative change in acquisition-related receivables compared to the firm’s normal receivables base, providing insight into the impact of acquisition activities on a firm’s receivables structure. We use end-of-fiscal-year values for both variables to ensure consistency and comparability across firms and over time.

3 Signal diagnostics

Figure 1 plots descriptive statistics for the AARC signal. Panel A plots the time-series of the mean, median, and interquartile range for AARC. On average, the cross-sectional mean (median) AARC is 3.96 (0.00) over the 1974 to 2023 sample, where the starting date is determined by the availability of the input AARC data. The signal’s interquartile range spans -0.28 to 0.25. Panel B of Figure 1 plots the time-series of the coverage of the AARC signal for the CRSP universe. On average, the AARC signal is available for 2.28% of CRSP names, which on average make up 2.90% of total market capitalization.

4 Does AARC predict returns?

Table 1 reports the performance of portfolios constructed using a value-weighted, quintile sort on AARC using NYSE breaks. The first two lines of Panel A report monthly average excess returns for each of the five portfolios and for the long/short portfolio that buys the high AARC portfolio and sells the low AARC portfolio. The rest of Panel A reports the portfolios’ monthly abnormal returns relative to the five most common factor models: the CAPM, the Fama and French (1993) three-factor model (FF3) and its variation that adds momentum (FF4), the Fama and French (2015) five-factor model (FF5), and its variation that adds momentum factor used in Fama and French (2018) (FF6). The table shows that the long/short AARC strategy earns an average return of 0.20% per month with a t-statistic of 2.07. The annualized Sharpe ratio of the strategy is 0.29. The alphas range from 0.20% to 0.24% per month and have t-statistics exceeding 1.95 everywhere. The lowest alpha is with respect to the FF3 factor model.

Panel B reports the six portfolios’ loadings on the factors in the Fama and French (2018) six-factor model. The long/short strategy’s most significant loading is -0.06,

with a t-statistic of -2.42 on the UMD factor. Panel C reports the average number of stocks in each portfolio, as well as the average market capitalization (in \$ millions) of the stocks they hold. In an average month, the five portfolios have at least 218 stocks and an average market capitalization of at least \$529 million.

Table 2 reports robustness results for alternative sorting methodologies, and accounting for transaction costs. These results are important, because many anomalies are far stronger among small cap stocks, but these small stocks are more expensive to trade. Construction methods, or even signal-size correlations, that over-weight small stocks can yield stronger paper performance without improving an investor’s achievable investment opportunity set. Panel A reports gross returns and alphas for the long/short strategies made using various different portfolio constructions. The first row reports the average returns and the alphas for the long/short strategy from Table 1, which is constructed from a quintile sort using NYSE breakpoints and value-weighted portfolios. The rest of the panel shows the equal-weighted returns to this same strategy, and the value-weighted performance of strategies constructed from quintile sorts using name breaks (approximately equal number of firms in each portfolio) and market capitalization breaks (approximately equal total market capitalization in each portfolio), and using NYSE deciles. The average return is lowest for the quintile sort using name breakpoints and value-weighted portfolios, and equals 13 bps/month with a t-statistics of 1.32. Out of the twenty-five alphas reported in Panel A, the t-statistics for fifteen exceed two, and for six exceed three.

Panel B reports for these same strategies the average monthly net returns and the generalized net alphas of [Novy-Marx and Velikov \(2016\)](#). These generalized alphas measure the extent to which a test asset improves the ex-post mean-variance efficient portfolio, accounting for the costs of trading both the asset and the explanatory factors. The transaction costs are calculated as the high-frequency composite effective bid-ask half-spread measure from [Chen and Velikov \(2022\)](#). The net average returns

reported in the first column range between -4-35bps/month. The lowest return, (-4 bps/month), is achieved from the quintile sort using NYSE breakpoints and equal-weighted portfolios, and has an associated t-statistic of -0.53. Out of the twenty-five construction-methodology-factor-model pairs reported in Panel B, the AARC trading strategy improves the achievable mean-variance efficient frontier spanned by the factor models in twenty cases, and significantly expands the achievable frontier in five cases.

Table 3 provides direct tests for the role size plays in the AARC strategy performance. Panel A reports the average returns for the twenty-five portfolios constructed from a conditional double sort on size and AARC, as well as average returns and alphas for long/short trading AARC strategies within each size quintile. Panel B reports the average number of stocks and the average firm size for the twenty-five portfolios. Among the largest stocks (those with market capitalization greater than the 80th NYSE percentile), the AARC strategy achieves an average return of 17 bps/month with a t-statistic of 1.39. Among these large cap stocks, the alphas for the AARC strategy relative to the five most common factor models range from 14 to 20 bps/month with t-statistics between 1.08 and 1.58.

5 How does AARC perform relative to the zoo?

Figure 2 puts the performance of AARC in context, showing the long/short strategy performance relative to other strategies in the “factor zoo.” It shows Sharpe ratio histograms, both for gross and net returns (Panel A and B, respectively), for 212 documented anomalies in the zoo.¹ The vertical red line shows where the Sharpe ratio for the AARC strategy falls in the distribution. The AARC strategy’s gross (net) Sharpe ratio of 0.29 (0.23) is greater than 63% (82%) of anomaly Sharpe ratios,

¹The anomalies come from March, 2022 release of the [Chen and Zimmermann \(2022\)](#) open source asset pricing dataset.

respectively.

Figure 3 plots the growth of a \$1 invested in these same 212 anomaly trading strategies (gray lines), and compares those with the growth of a \$1 invested in the AARC strategy (red line).² Ignoring trading costs, a \$1 invested in the AARC strategy would have yielded \$1.73 which ranks the AARC strategy in the top 12% across the 212 anomalies. Accounting for trading costs, a \$1 invested in the AARC strategy would have yielded \$1.09 which ranks the AARC strategy in the top 9% across the 212 anomalies.

Figure 4 plots percentile ranks for the 212 anomaly trading strategies in terms of gross and Novy-Marx and Velikov (2016) net generalized alphas with respect to the CAPM, and the Fama-French three-, four-, five-, and six-factor models from Table 1, and indicates the ranking of the AARC relative to those. Panel A shows that the AARC strategy gross alphas fall between the 41 and 75 percentiles across the five factor models. Panel B shows that, accounting for trading costs, a large fraction of anomalies have not improved the investment opportunity set of an investor with access to the factor models over the 197406 to 202306 sample. For example, 45% (53%) of the 212 anomalies would not have improved the investment opportunity set for an investor having access to the Fama-French three-factor (six-factor) model. The AARC strategy has a positive net generalized alpha for five out of the five factor models. In these cases AARC ranks between the 61 and 82 percentiles in terms of how much it could have expanded the achievable investment frontier.

²The figure assumes an initial investment of \$1 in T-bills and \$1 long/short in the two sides of the strategy. Returns are compounded each month, assuming, as in Detzel et al. (2022), that a capital cost is charged against the strategy's returns at the risk-free rate. This excess return corresponds more closely to the strategy's economic profitability.

6 Does AARC add relative to related anomalies?

With so many anomalies, it is possible that any proposed, new cross-sectional predictor is just capturing some combination of known predictors. It is consequently natural to investigate to what extent the proposed predictor adds additional predictive power beyond the most closely related anomalies. Closely related anomalies are more likely to be formed on the basis of signals with higher absolute correlations. Figure 5 plots a name histogram of the correlations of AARC with 204 filtered anomaly signals.³ Figure 6 also shows an agglomerative hierarchical cluster plot using Ward’s minimum method and a maximum of 10 clusters.

A closely related anomaly is also more likely to price AARC or at least to weaken the power AARC has predicting the cross-section of returns. Figure 7 plots histograms of t-statistics for predictability tests of AARC conditioning on each of the 204 filtered anomaly signals one at a time. Panel A reports t-statistics on β_{AARC} from Fama-MacBeth regressions of the form $r_{i,t} = \alpha + \beta_{AARC}AARC_{i,t} + \beta_X X_{i,t} + \epsilon_{i,t}$, where X stands for one of the 204 filtered anomaly signals at a time. Panel B plots t-statistics on α from spanning tests of the form: $r_{AARC,t} = \alpha + \beta r_{X,t} + \epsilon_t$, where $r_{X,t}$ stands for the returns to one of the 204 filtered anomaly trading strategies at a time. The strategies employed in the spanning tests are constructed using quintile sorts, value-weighting, and NYSE breakpoints. Panel C plots t-statistics on the average returns to strategies constructed by conditional double sorts. In each month, we sort stocks into quintiles based one of the 204 filtered anomaly signals. Then, within each quintile, we sort stocks into quintiles based on AARC. Stocks are finally grouped into five AARC portfolios by combining stocks within each anomaly sorting portfolio. The panel plots the t-statistics on the average returns of these

³When performing tests at the underlying signal level (e.g., the correlations plotted in Figure 5), we filter the 212 anomalies to avoid small sample issues. For each anomaly, we calculate the common stock observations in an average month for which both the anomaly and the test signal are available. In the filtered anomaly set, we drop anomalies with fewer than 100 common stock observations in an average month.

conditional double-sorted AARC trading strategies conditioned on each of the 204 filtered anomalies.

Table 4 reports Fama-MacBeth cross-sectional regressions of returns on AARC and the six anomalies most closely-related to it. The six most-closely related anomalies are picked as those with the highest combined rank where the ranks are based on the absolute value of the Spearman correlations in Panel B of Figure 5 and the R^2 from the spanning tests in Figure 7, Panel B. Controlling for each of these signals at a time, the t-statistics on the AARC signal in these Fama-MacBeth regressions exceed -0.14, with the minimum t-statistic occurring when controlling for Investment to revenue. Controlling for all six closely related anomalies, the t-statistic on AARC is -0.12.

Similarly, Table 5 reports results from spanning tests that regress returns to the AARC strategy onto the returns of the six most closely-related anomalies and the six Fama-French factors. Controlling for the six most-closely related anomalies individually, the AARC strategy earns alphas that range from 23-27bps/month. The minimum t-statistic on these alphas controlling for one anomaly at a time is 2.20, which is achieved when controlling for Investment to revenue. Controlling for all six closely-related anomalies and the six Fama-French factors simultaneously, the AARC trading strategy achieves an alpha of 26bps/month with a t-statistic of 2.56.

7 Does AARC add relative to the whole zoo?

Finally, we can ask how much adding AARC to the entire factor zoo could improve investment performance. Figure 8 plots the growth of \$1 invested in trading strategies that combine multiple anomalies following [Chen and Velikov \(2022\)](#). The combinations use either the 156 anomalies from the zoo that satisfy our inclusion

criteria (blue lines) or these 156 anomalies augmented with the AARC signal.⁴ We consider one different methods for combining signals.

Panel A shows results using “Average rank” as the combination method. This method sorts stocks on the basis of forecast excess returns, where these are calculated on the basis of their average cross-sectional percentile rank across return predictors, and the predictors are all signed so that higher ranks are associated with higher average returns. For this method, \$1 investment in the 156-anomaly combination strategy grows to \$935.00, while \$1 investment in the combination strategy that includes AARC grows to \$876.86.

8 Conclusion

This study provides compelling evidence for the predictive power of Acquisition Adjusted Receivables Current (AARC) in forecasting cross-sectional stock returns. Our findings demonstrate that AARC generates economically and statistically significant returns, with a value-weighted long/short strategy achieving an impressive annualized gross Sharpe ratio of 0.29. The signal’s robustness is particularly noteworthy, maintaining significant abnormal returns even after controlling for established factors and related accounting-based strategies.

The persistence of AARC’s predictive power, even after accounting for transaction costs (net Sharpe ratio of 0.23), suggests practical implementability for institutional investors. Furthermore, the signal’s distinct information content is evidenced by its significant alpha when controlling for six closely related strategies from the factor zoo, indicating that AARC captures unique aspects of firm performance not reflected in other known predictors.

However, several limitations warrant consideration. First, our analysis focuses

⁴We filter the 207 [Chen and Zimmermann \(2022\)](#) anomalies and require for each anomaly the average month to have at least 40% of the cross-sectional observations available for market capitalization on CRSP in the period for which AARC is available.

primarily on U.S. equity markets, and the signal's effectiveness in international markets remains to be tested. Second, the study period may not fully capture the signal's behavior across different market regimes.

Future research could explore several promising directions. Investigation of AARC's performance in international markets could provide insights into its global applicability. Additionally, examining the interaction between AARC and other accounting-based signals could reveal potential complementarities or substitution effects. Finally, studying the underlying economic mechanisms driving AARC's predictive power could enhance our understanding of market inefficiencies and improve signal implementation.

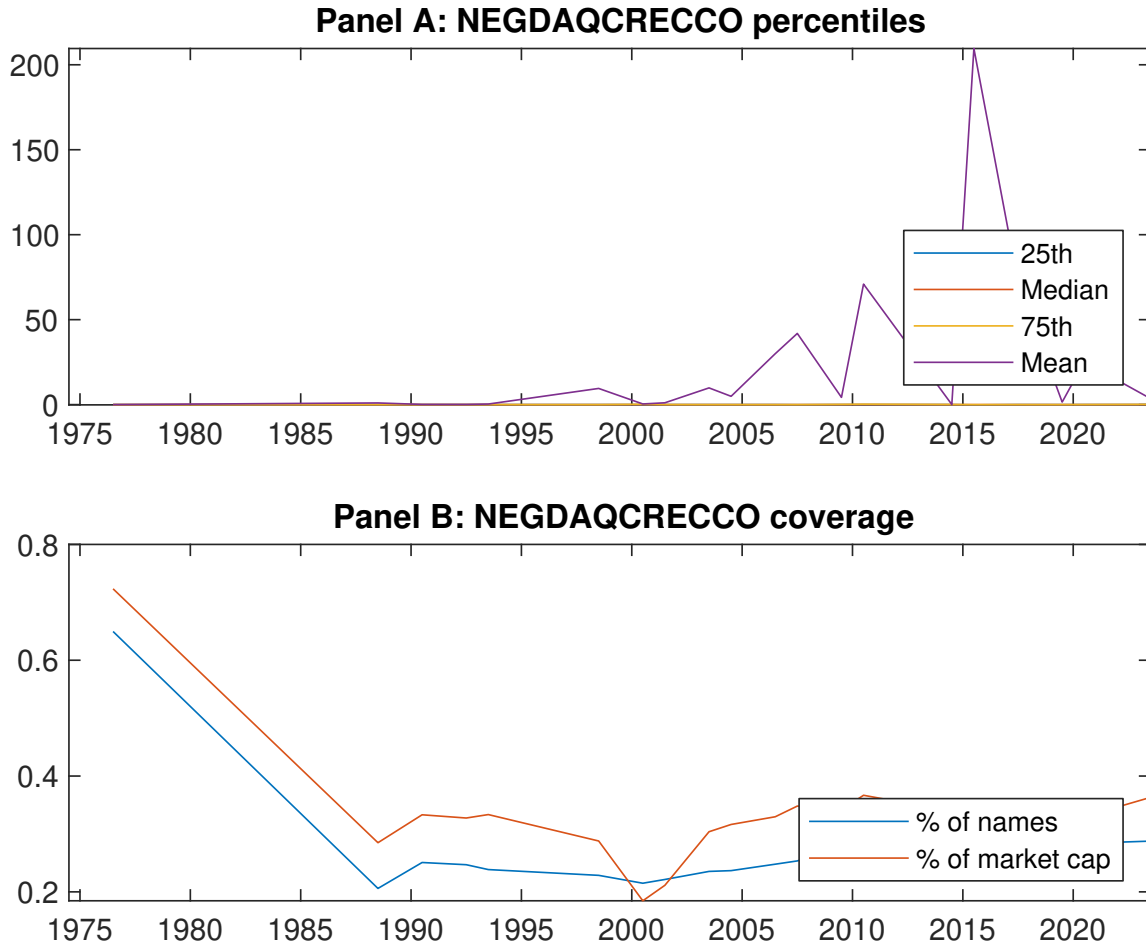


Figure 1: Times series of AARC percentiles and coverage.
This figure plots descriptive statistics for AARC. Panel A shows cross-sectional percentiles of AARC over the sample. Panel B plots the monthly coverage of AARC relative to the universe of CRSP stocks with available market capitalizations.

Table 1: Basic sort: VW, quintile, NYSE-breaks

This table reports average excess returns and alphas for portfolios sorted on AARC. At the end of each month, we sort stocks into five portfolios based on their signal using NYSE breakpoints. Panel A reports average value-weighted quintile portfolio (L,2,3,4,H) returns in excess of the risk-free rate, the long-short extreme quintile portfolio (H-L) return, and alphas with respect to the CAPM, Fama and French (1993) three-factor model, Fama and French (1993) three-factor model augmented with the Carhart (1997) momentum factor, Fama and French (2015) five-factor model, and the Fama and French (2015) five-factor model augmented with the Carhart (1997) momentum factor following Fama and French (2018). Panel B reports the factor loadings for the quintile portfolios and long-short extreme quintile portfolio in the Fama and French (2015) five-factor model. Panel C reports the average number of stocks and market capitalization of each portfolio. T-statistics are in brackets. The sample period is 197406 to 202306.

Panel A: Excess returns and alphas on AARC-sorted portfolios						
	(L)	(2)	(3)	(4)	(H)	(H-L)
r^e	0.55 [2.75]	0.79 [3.98]	0.59 [3.11]	0.64 [3.11]	0.76 [3.76]	0.20 [2.07]
α_{CAPM}	-0.12 [-1.52]	0.15 [1.67]	-0.03 [-0.42]	-0.04 [-0.41]	0.08 [1.11]	0.20 [2.03]
α_{FF3}	-0.15 [-1.95]	0.08 [0.94]	-0.08 [-0.97]	-0.07 [-0.80]	0.05 [0.62]	0.20 [1.95]
α_{FF4}	-0.16 [-2.09]	0.10 [1.24]	-0.12 [-1.52]	-0.06 [-0.65]	0.08 [1.05]	0.24 [2.38]
α_{FF5}	-0.26 [-3.33]	-0.03 [-0.37]	-0.14 [-1.77]	-0.08 [-0.90]	-0.04 [-0.60]	0.21 [2.05]
α_{FF6}	-0.26 [-3.31]	-0.00 [-0.01]	-0.17 [-2.09]	-0.07 [-0.76]	-0.01 [-0.16]	0.24 [2.36]
Panel B: Fama and French (2018) 6-factor model loadings for AARC-sorted portfolios						
β_{MKT}	1.02 [57.33]	1.01 [51.78]	0.96 [52.03]	1.02 [49.46]	1.02 [60.07]	-0.00 [-0.12]
β_{SMB}	0.03 [1.05]	-0.04 [-1.48]	0.00 [0.15]	-0.07 [-2.10]	0.01 [0.29]	-0.02 [-0.58]
β_{HML}	0.05 [1.55]	0.12 [3.24]	0.09 [2.67]	0.07 [1.86]	0.05 [1.43]	-0.01 [-0.14]
β_{RMW}	0.22 [6.24]	0.19 [4.85]	0.08 [2.22]	-0.02 [-0.43]	0.20 [6.01]	-0.02 [-0.39]
β_{CMA}	0.07 [1.44]	0.19 [3.31]	0.10 [1.82]	0.09 [1.47]	0.09 [1.83]	0.02 [0.23]
β_{UMD}	0.00 [0.08]	-0.05 [-2.74]	0.05 [2.49]	-0.02 [-0.96]	-0.06 [-3.31]	-0.06 [-2.42]
Panel C: Average number of firms (n) and market capitalization (me)						
n	218	266	304	271	218	
me (\$10 ⁶)	751	936	529	890	762	

Table 2: Robustness to sorting methodology & trading costs

This table evaluates the robustness of the choices made in the AARC strategy construction methodology. In each panel, the first row shows results from a quintile, value-weighted sort using NYSE break points as employed in Table 1. Each of the subsequent rows deviates in one of the three choices at a time, and the choices are specified in the first three columns. For each strategy construction methodology, the table reports average excess returns and alphas with respect to the CAPM, Fama and French (1993) three-factor model, Fama and French (1993) three-factor model augmented with the Carhart (1997) momentum factor, Fama and French (2015) five-factor model, and the Fama and French (2015) five-factor model augmented with the Carhart (1997) momentum factor following Fama and French (2018). Panel A reports average returns and alphas with no adjustment for trading costs. Panel B reports net average returns and Novy-Marx and Velikov (2016) generalized alphas as prescribed by Detzel et al. (2022). T-statistics are in brackets. The sample period is 197406 to 202306.

Panel A: Gross Returns and Alphas								
Portfolios	Breaks	Weights	r^e	α_{CAPM}	α_{FF3}	α_{FF4}	α_{FF5}	α_{FF6}
Quintile	NYSE	VW	0.20 [2.07]	0.20 [2.03]	0.20 [1.95]	0.24 [2.38]	0.21 [2.05]	0.24 [2.36]
Quintile	NYSE	EW	0.20 [3.01]	0.22 [3.27]	0.21 [3.03]	0.19 [2.75]	0.18 [2.59]	0.17 [2.47]
Quintile	Name	VW	0.13 [1.32]	0.13 [1.26]	0.13 [1.27]	0.17 [1.64]	0.17 [1.61]	0.19 [1.84]
Quintile	Cap	VW	0.14 [1.40]	0.13 [1.24]	0.14 [1.33]	0.20 [1.86]	0.19 [1.79]	0.23 [2.15]
Decile	NYSE	VW	0.41 [3.36]	0.44 [3.54]	0.42 [3.36]	0.46 [3.63]	0.35 [2.77]	0.39 [3.08]
Panel B: Net Returns and Novy-Marx and Velikov (2016) generalized alphas								
Portfolios	Breaks	Weights	r_{net}^e	α_{CAPM}^*	α_{FF3}^*	α_{FF4}^*	α_{FF5}^*	α_{FF6}^*
Quintile	NYSE	VW	0.16 [1.61]	0.16 [1.56]	0.15 [1.51]	0.18 [1.77]	0.15 [1.44]	0.17 [1.63]
Quintile	NYSE	EW	-0.04 [-0.53]					
Quintile	Name	VW	0.09 [0.88]	0.08 [0.82]	0.08 [0.84]	0.11 [1.07]	0.10 [1.01]	0.12 [1.17]
Quintile	Cap	VW	0.10 [0.97]	0.08 [0.80]	0.09 [0.89]	0.13 [1.22]	0.13 [1.20]	0.15 [1.41]
Decile	NYSE	VW	0.35 [2.89]	0.38 [3.03]	0.36 [2.90]	0.39 [3.09]	0.30 [2.41]	0.32 [2.55]

Table 3: Conditional sort on size and AARC

This table presents results for conditional double sorts on size and AARC. In each month, stocks are first sorted into quintiles based on size using NYSE breakpoints. Then, within each size quintile, stocks are further sorted based on AARC. Finally, they are grouped into twenty-five portfolios based on the intersection of the two sorts. Panel A presents the average returns to the 25 portfolios, as well as strategies that go long stocks with high AARC and short stocks with low AARC. Panel B documents the average number of firms and the average firm size for each portfolio. The sample period is 197406 to 202306.

Panel A: portfolio average returns and time-series regression results												
Size quintiles	AARC Quintiles					AARC Strategies						
	(L)	(2)	(3)	(4)	(H)	r^e	α_{CAPM}	α_{FF3}	α_{FF4}	α_{FF5}	α_{FF6}	
	(1)	0.66 [2.35]	0.91 [3.01]	0.92 [3.16]	0.66 [2.29]	0.98 [2.76]	0.32 [1.44]	0.33 [1.47]	0.30 [1.31]	0.22 [0.96]	0.21 [0.89]	0.16 [0.70]
	(2)	0.81 [3.00]	0.68 [2.59]	0.85 [3.23]	0.97 [3.82]	0.95 [3.53]	0.14 [1.13]	0.17 [1.36]	0.14 [1.10]	0.12 [0.94]	0.09 [0.73]	0.09 [0.67]
	(3)	0.91 [3.74]	0.91 [3.77]	0.88 [3.63]	0.83 [3.49]	0.88 [3.42]	-0.04 [-0.28]	-0.06 [-0.48]	-0.11 [-0.82]	-0.10 [-0.76]	-0.06 [-0.47]	-0.06 [-0.43]
	(4)	0.62 [2.77]	0.85 [3.75]	0.87 [3.83]	0.88 [3.83]	0.79 [3.34]	0.16 [1.37]	0.11 [0.94]	0.13 [1.07]	0.12 [0.94]	0.13 [1.07]	0.12 [0.98]
	(5)	0.61 [2.96]	0.69 [3.39]	0.48 [2.49]	0.56 [2.74]	0.78 [3.86]	0.17 [1.39]	0.18 [1.48]	0.15 [1.25]	0.20 [1.58]	0.14 [1.08]	0.17 [1.35]
Panel B: Portfolio average number of firms and market capitalization												
Size quintiles	AARC Quintiles					AARC Quintiles						
	Average n					Average market capitalization (\$10 ⁶)						
	(L)	(2)	(3)	(4)	(H)	(L)	(2)	(3)	(4)	(H)		
	(1)	148	147	148	147	147	12	10	10	10	12	
	(2)	37	37	37	37	37	19	19	19	19	19	
	(3)	27	26	26	27	27	35	34	35	35	35	
	(4)	22	22	22	22	22	77	81	79	80	78	
(5)	22	22	22	22	22	575	747	477	755	595		

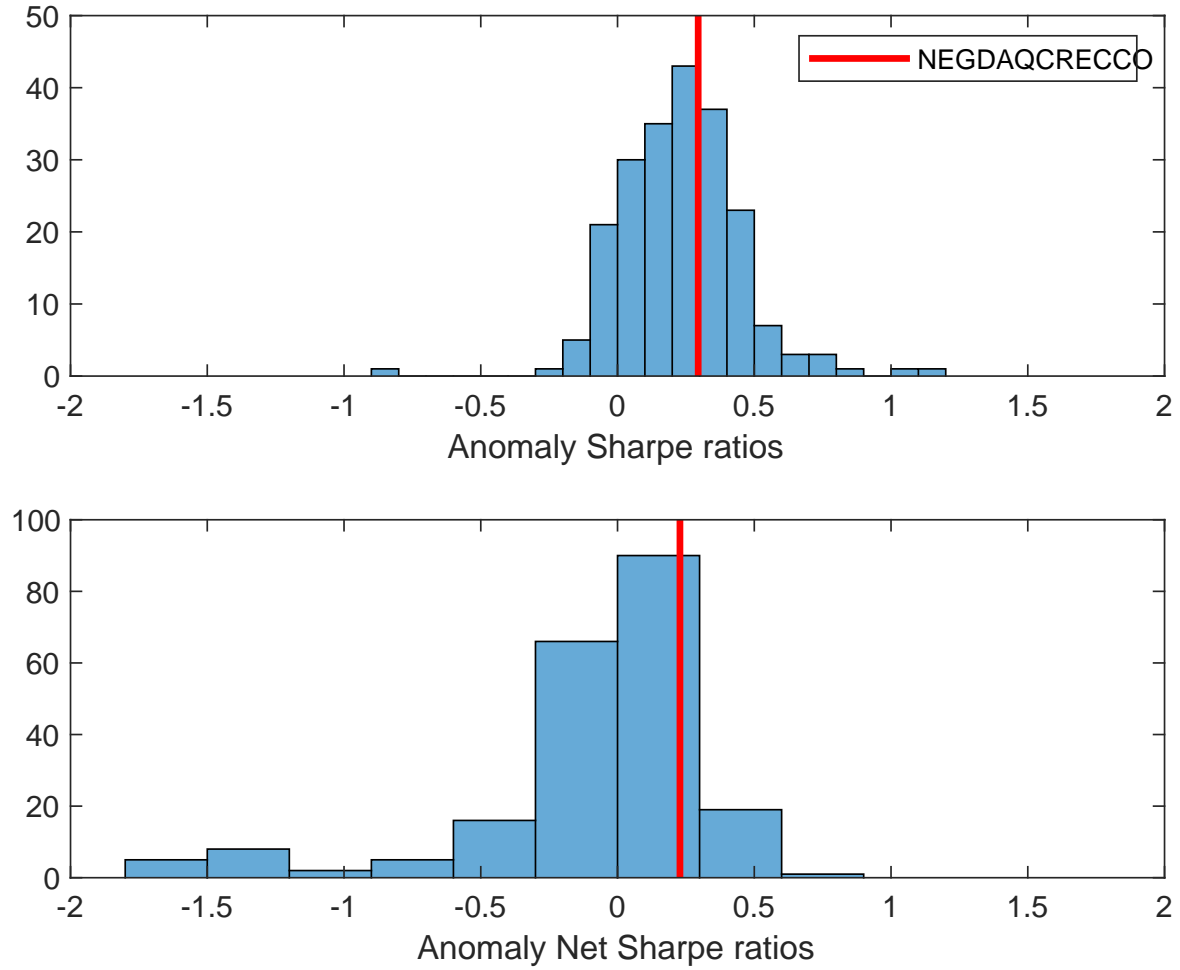


Figure 2: Distribution of Sharpe ratios.

This figure plots a histogram of Sharpe ratios for 212 anomalies, and compares the Sharpe ratio of the AARC with them (red vertical line). Panel A plots results for gross Sharpe ratios. Panel B plots results for net Sharpe ratios.

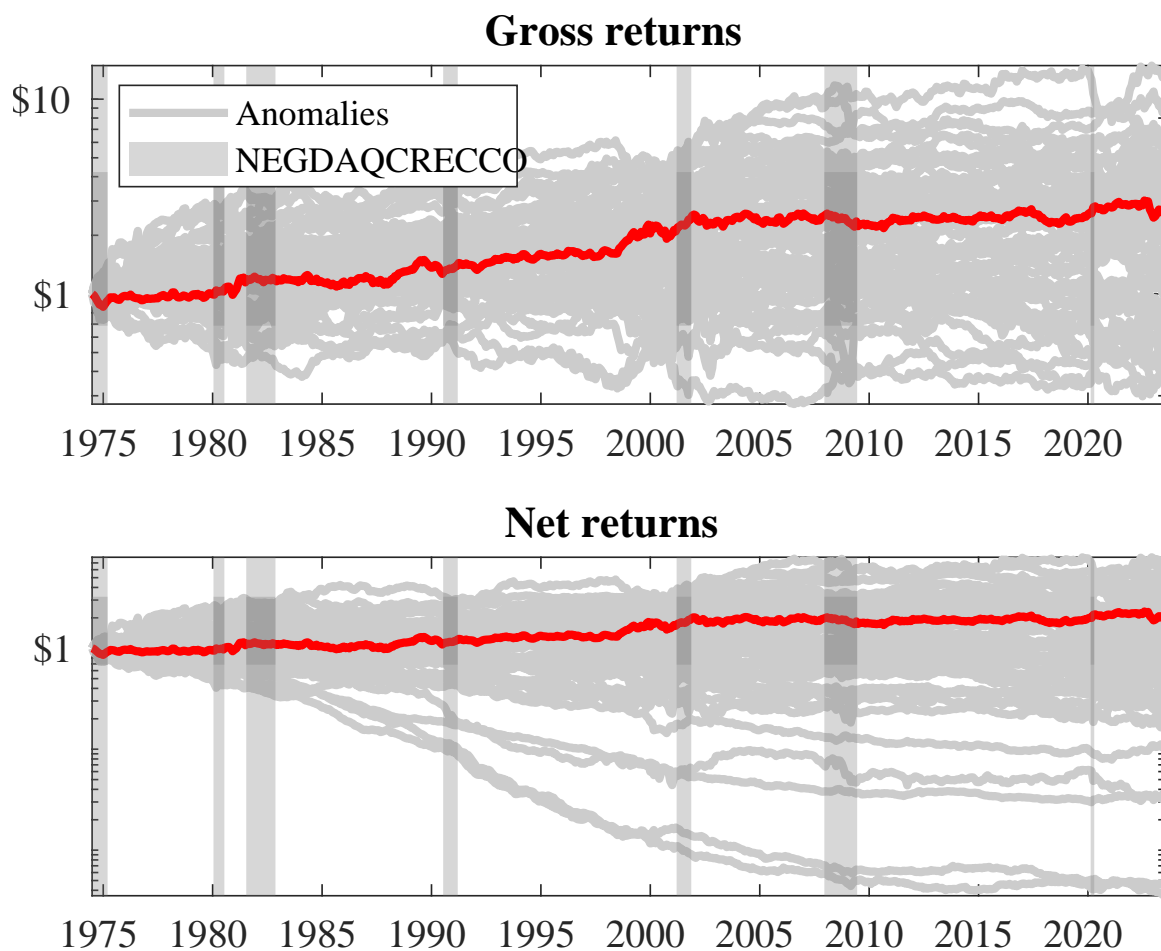


Figure 3: Dollar invested.

This figure plots the growth of a \$1 invested in 212 anomaly trading strategies (gray lines), and compares those with the AARC trading strategy (red line). The strategies are constructed using value-weighted quintile sorts using NYSE breakpoints. Panel A plots results for gross strategy returns. Panel B plots results for net strategy returns.

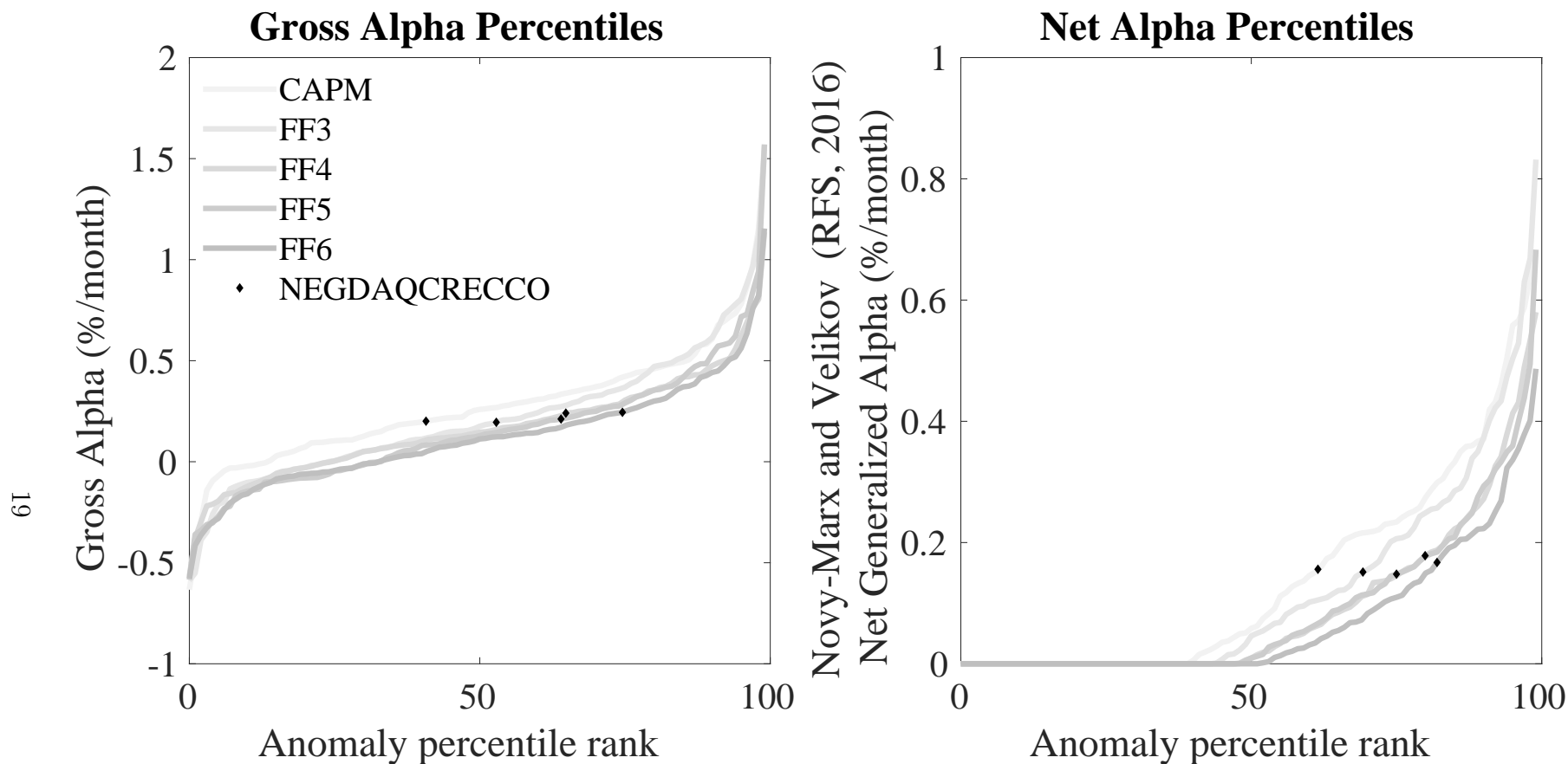


Figure 4: Gross and generalized net alpha percentiles of anomalies relative to factor models

This figure plots the percentile ranks for 212 anomaly trading strategies in terms of alphas (solid lines), and compares those with the AARC trading strategy alphas (diamonds). The strategies are constructed using value-weighted quintile sorts using NYSE breakpoints. The alphas include those with respect to the CAPM, Fama and French (1993) three-factor model, Fama and French (1993) three-factor model augmented with the Carhart (1997) momentum factor, Fama and French (2015) five-factor model, and the Fama and French (2015) five-factor model augmented with the Carhart (1997) momentum factor following Fama and French (2018). The left panel plots alphas with no adjustment for trading costs. The right panel plots Novy-Marx and Velikov (2016) net generalized alphas.

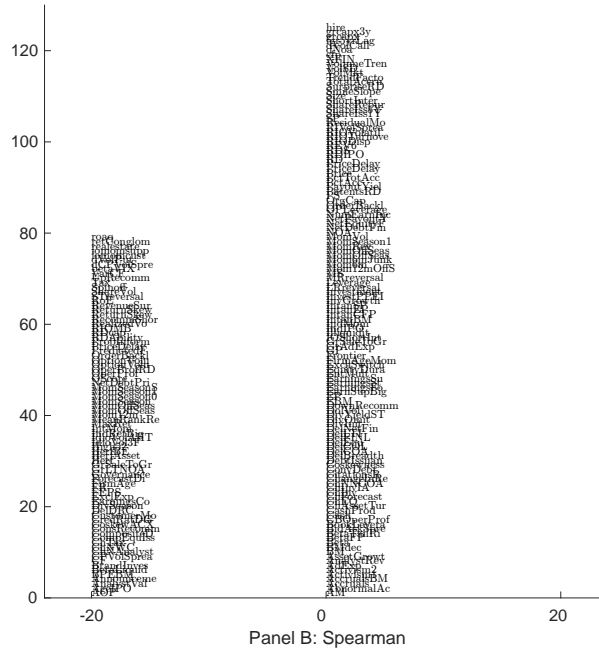
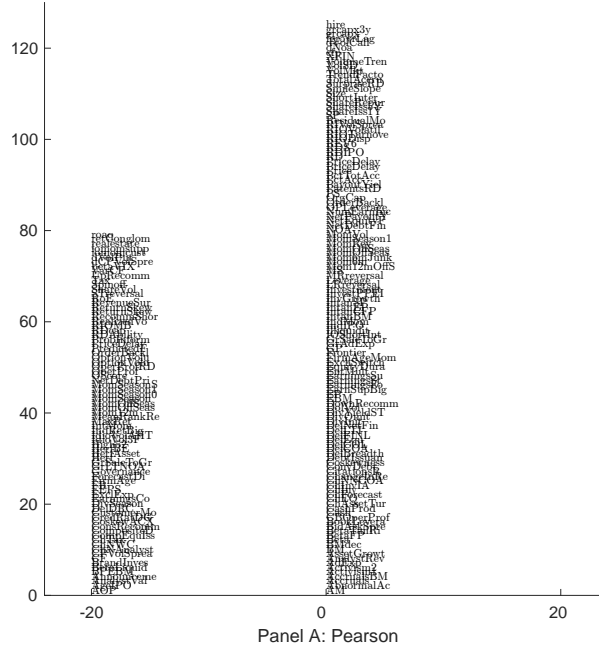


Figure 5: Distribution of correlations.

This figure plots a name histogram of correlations of 204 filtered anomaly signals with AARC. The correlations are pooled. Panel A plots Pearson correlations, while Panel B plots Spearman rank correlations.

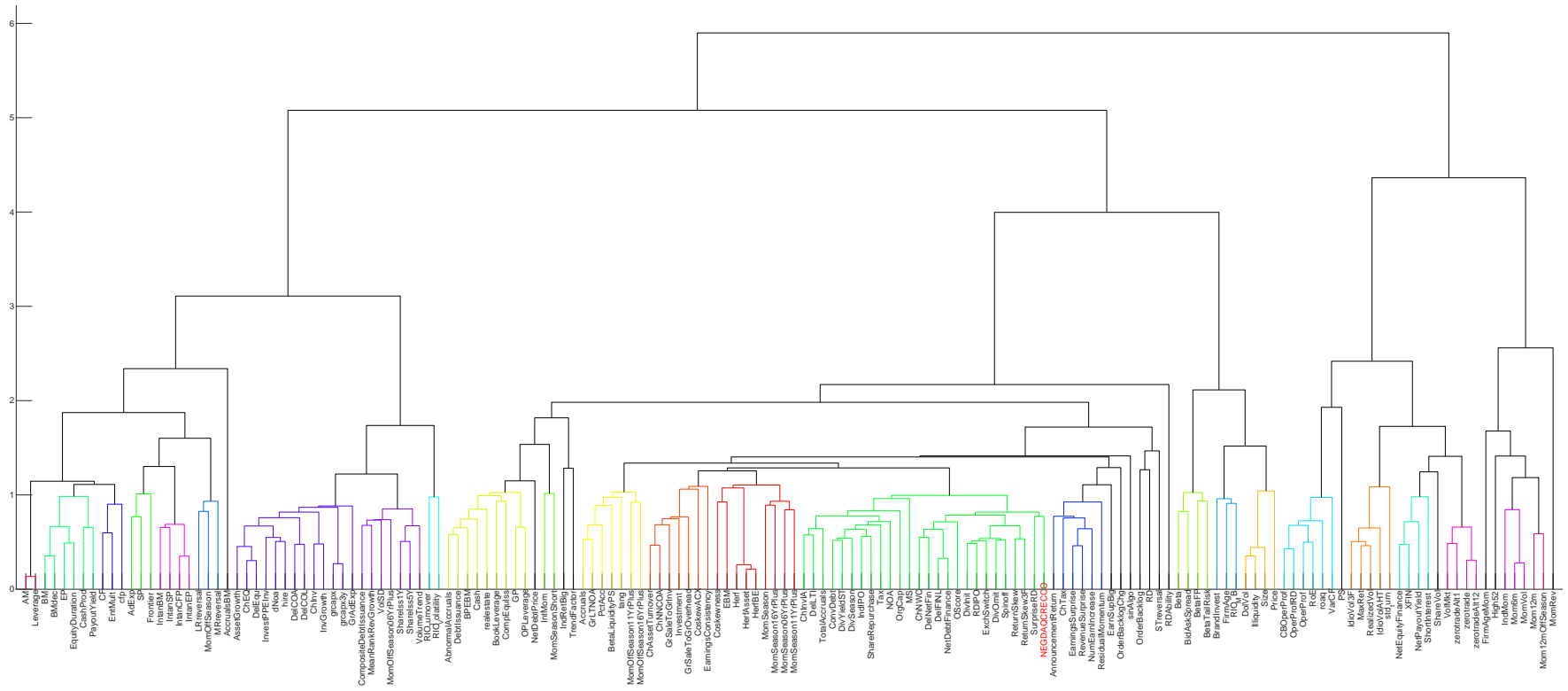


Figure 6: Agglomerative hierarchical cluster plot

This figure plots an agglomerative hierarchical cluster plot using Ward's minimum method and a maximum of 10 clusters.

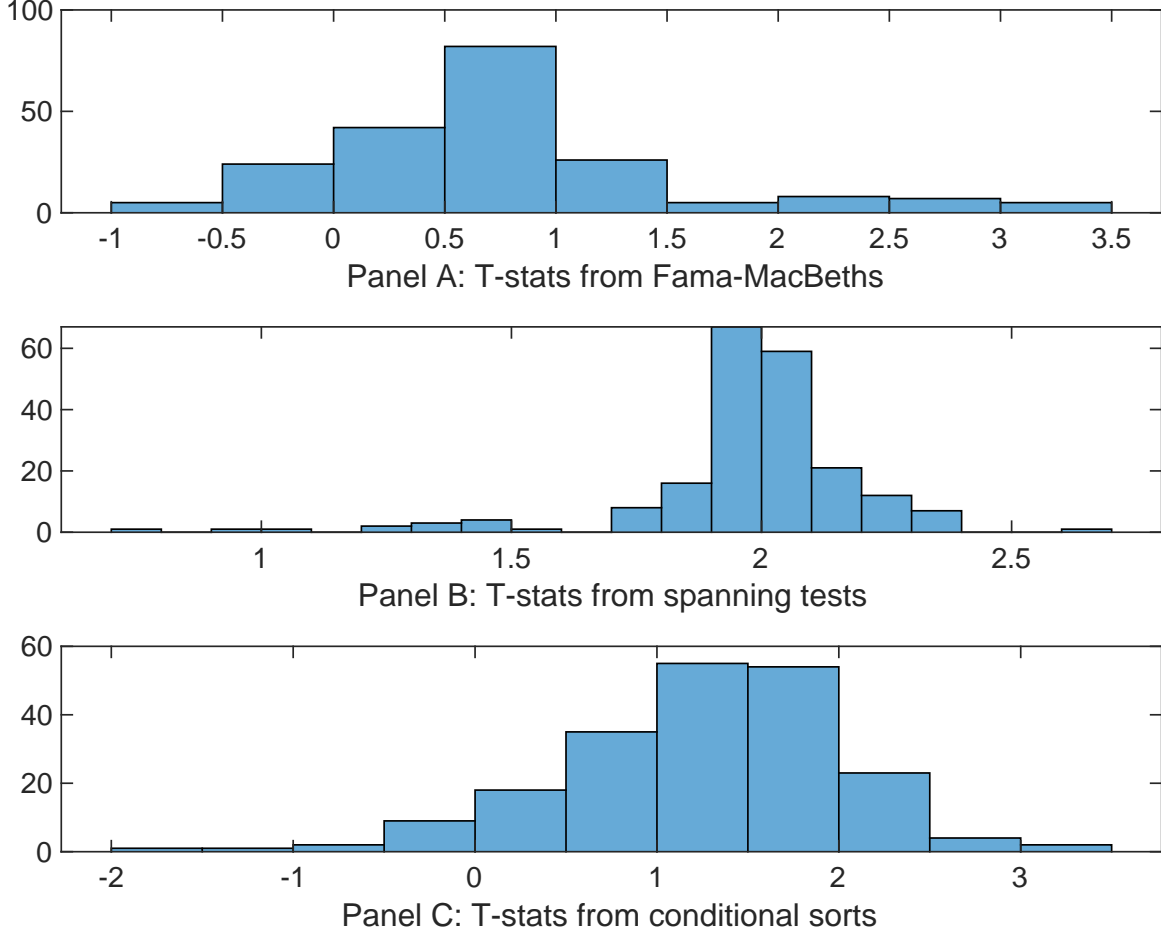


Figure 7: Distribution of t-stats on conditioning strategies

This figure plots histograms of t-statistics for predictability tests of AARC conditioning on each of the 204 filtered anomaly signals one at a time. Panel A reports t-statistics on β_{AARC} from Fama-MacBeth regressions of the form $r_{i,t} = \alpha + \beta_{AARC}AARC_{i,t} + \beta_X X_{i,t} + \epsilon_{i,t}$, where X stands for one of the 204 filtered anomaly signals at a time. Panel B plots t-statistics on α from spanning tests of the form: $r_{AARC,t} = \alpha + \beta r_{X,t} + \epsilon_t$, where $r_{X,t}$ stands for the returns to one of the 204 filtered anomaly trading strategies at a time. The strategies employed in the spanning tests are constructed using quintile sorts, value-weighting, and NYSE breakpoints. Panel C plots t-statistics on the average returns to strategies constructed by conditional double sorts. In each month, we sort stocks into quintiles based on one of the 204 filtered anomaly signals at a time. Then, within each quintile, we sort stocks into quintiles based on AARC. Stocks are finally grouped into five AARC portfolios by combining stocks within each anomaly sorting portfolio. The panel plots the t-statistics on the average returns of these conditional double-sorted AARC trading strategies conditioned on each of the 204 filtered anomalies.

Table 4: Fama-MacBeths controlling for most closely related anomalies

This table presents Fama-MacBeth results of returns on AARC. and the six most closely related anomalies. The regressions take the following form: $r_{i,t} = \alpha + \beta_{AARC} AARC_{i,t} + \sum_{k=1}^s \beta_{X_k} X_{i,t}^k + \epsilon_{i,t}$. The six most closely related anomalies, X , are Change in Net Non-current Op Assets, Sales growth over inventory growth, Net external financing, Inventory Growth, Change in current operating assets, Investment to revenue. These anomalies were picked as those with the highest combined rank where the ranks are based on the absolute value of the Spearman correlations in Panel B of Figure 5 and the R^2 from the spanning tests in Figure 7, Panel B. The sample period is 197406 to 202306.

Intercept	0.13 [5.12]	0.13 [5.18]	0.14 [5.62]	0.14 [5.24]	0.13 [5.33]	0.15 [5.85]	0.15 [5.43]
AARC	0.11 [0.66]	0.12 [0.78]	0.45 [0.26]	0.56 [0.31]	0.59 [0.04]	-0.27 [-0.14]	-0.26 [-0.12]
Anomaly 1	0.76 [2.45]						0.70 [0.16]
Anomaly 2		0.11 [2.26]					-0.68 [-0.58]
Anomaly 3			0.20 [5.68]				0.17 [3.68]
Anomaly 4				0.46 [6.27]			0.12 [0.82]
Anomaly 5					0.24 [6.70]		0.18 [3.01]
Anomaly 6						0.18 [3.09]	0.80 [1.01]
# months	588	588	588	588	588	588	588
$\bar{R}^2(\%)$	0	0	1	0	0	0	0

Table 5: Spanning tests controlling for most closely related anomalies

This table presents spanning tests results of regressing returns to the AARC trading strategy on trading strategies exploiting the six most closely related anomalies. The regressions take the following form: $r_t^{AARC} = \alpha + \sum_{k=1}^6 \beta_{X_k} r_t^{X_k} + \sum_{j=1}^6 \beta_{f_j} r_t^{f_j} + \epsilon_t$, where X_k indicates each of the six most-closely related anomalies and f_j indicates the six factors from the [Fama and French \(2015\)](#) five-factor model augmented with the [Carhart \(1997\)](#) momentum factor. The six most closely related anomalies, X , are Change in Net Noncurrent Op Assets, Sales growth over inventory growth, Net external financing, Inventory Growth, Change in current operating assets, Investment to revenue. These anomalies were picked as those with the highest combined rank where the ranks are based on the absolute value of the Spearman correlations in Panel B of Figure 5 and the R^2 from the spanning tests in Figure 7, Panel B. The sample period is 197406 to 202306.

Intercept	0.25 [2.43]	0.25 [2.42]	0.23 [2.20]	0.24 [2.35]	0.25 [2.39]	0.27 [2.63]	0.26 [2.56]
Anomaly 1	-10.12 [-2.12]						-3.37 [-0.68]
Anomaly 2		-10.47 [-2.31]					1.01 [0.19]
Anomaly 3			15.41 [2.92]				13.77 [2.58]
Anomaly 4				-8.25 [-2.01]			-11.96 [-2.53]
Anomaly 5					9.18 [1.74]		13.01 [2.38]
Anomaly 6						-19.31 [-4.84]	-16.93 [-4.08]
mkt	-0.79 [-0.33]	-0.53 [-0.22]	1.26 [0.51]	-0.59 [-0.25]	-0.95 [-0.40]	-0.25 [-0.11]	1.65 [0.68]
smb	-1.95 [-0.53]	-2.11 [-0.57]	2.54 [0.63]	-3.26 [-0.88]	-0.66 [-0.17]	1.00 [0.27]	6.40 [1.56]
hml	-0.87 [-0.19]	-1.49 [-0.33]	0.43 [0.09]	-0.90 [-0.20]	-4.94 [-0.98]	-2.69 [-0.60]	-5.65 [-1.13]
rmw	0.53 [0.11]	0.27 [0.06]	-10.03 [-1.76]	-1.98 [-0.41]	0.31 [0.06]	-2.50 [-0.53]	-10.24 [-1.75]
cma	1.07 [0.15]	1.89 [0.27]	-8.86 [-1.14]	8.80 [1.12]	-3.50 [-0.47]	3.34 [0.49]	-2.66 [-0.31]
umd	-4.61 [-1.85]	-4.66 [-1.89]	-5.97 [-2.49]	-5.18 [-2.12]	-5.47 [-2.25]	-3.33 [-1.37]	-1.65 [-0.66]
# months	588	588	588	588	588	588	588
$\bar{R}^2(\%)$	1	1	1	1	1	4	6

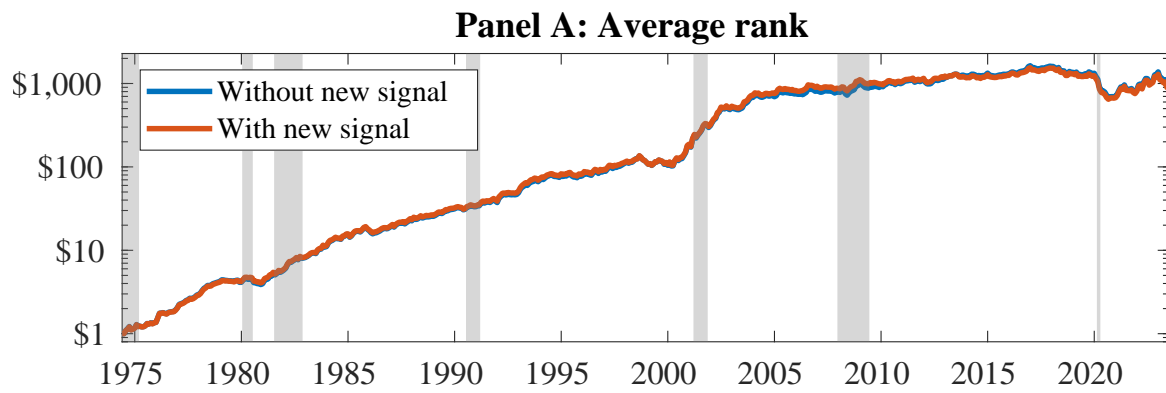


Figure 8: Combination strategy performance

This figure plots the growth of a \$1 invested in trading strategies that combine multiple anomalies following [Chen and Velikov \(2022\)](#). In all panels, the blue solid lines indicate combination trading strategies that utilize 156 anomalies. The red solid lines indicate combination trading strategies that utilize the 156 anomalies as well as AARC. Panel A shows results using "Average rank" as the combination method. See [Section 7](#) for details on the combination methods.

References

- Carhart, M. M. (1997). On persistence in mutual fund performance. *Journal of Finance*, 52:57–82.
- Chen, A. and Velikov, M. (2022). Zeroing in on the expected returns of anomalies. *Journal of Financial and Quantitative Analysis*, Forthcoming.
- Chen, A. Y. and Zimmermann, T. (2022). Open source cross-sectional asset pricing. *Critical Finance Review*, 27(2):207–264.
- Dechow, P. M. and Ge, W. (2006). The persistence of earnings and cash flows and the role of special items: Implications for the accrual anomaly. *Review of Accounting Studies*, 11(2-3):253–296.
- Detzel, A., Novy-Marx, R., and Velikov, M. (2022). Model comparison with transaction costs. *Journal of Finance*, Forthcoming.
- Ecker, F. and Skaperdas, E. (2019). Accounting quality and trade credit. *Journal of Accounting Research*, 57(3):743–790.
- Fama, E. F. and French, K. R. (1993). Common risk factors in the returns on stocks and bonds. *Journal of Financial Economics*, 33(1):3–56.
- Fama, E. F. and French, K. R. (2015). A five-factor asset pricing model. *Journal of Financial Economics*, 116(1):1–22.
- Fama, E. F. and French, K. R. (2018). Choosing factors. *Journal of Financial Economics*, 128(2):234–252.
- Harvey, C. R., Liu, Y., and Zhu, H. (2016). ... and the cross-section of expected returns. *Review of Financial Studies*, 29(1):5–68.

- Novy-Marx, R. and Velikov, M. (2016). A taxonomy of anomalies and their trading costs. *Review of Financial Studies*, 29(1):104–147.
- Novy-Marx, R. and Velikov, M. (2023). Assaying anomalies. *Working paper*.
- Penman, S. H., Richardson, S. A., and Tuna, I. (2007). The book-to-price effect in stock returns: Accounting for leverage. *Journal of Accounting Research*, 45(2):427–467.
- Richardson, S. A., Sloan, R. G., Soliman, M. T., and Tuna, I. (2005). Accrual reliability, earnings persistence and stock prices. *Journal of Accounting and Economics*, 39(3):437–485.
- Thomas, J. K. and Zhang, H. (2002). Inventory changes and future returns. *Review of Accounting Studies*, 7(2-3):163–187.
- Zhang, X. F. (2007). Limited attention, information processing costs, and stock return predictability. *Journal of Finance*, 62(4):1653–1684.