

Debt Issuance Efficiency and the Cross Section of Stock Returns

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Abstract

This paper studies the asset pricing implications of Debt Issuance Efficiency (DIE), and its robustness in predicting returns in the cross-section of equities using the protocol proposed by [Novy-Marx and Velikov \(2023\)](#). A value-weighted long/short trading strategy based on DIE achieves an annualized gross (net) Sharpe ratio of 0.50 (0.39), and monthly average abnormal gross (net) return relative to the [Fama and French \(2015\)](#) five-factor model plus a momentum factor of 25 (22) bps/month with a t-statistic of 3.31 (2.92), respectively. Its gross monthly alpha relative to these six factors plus the six most closely related strategies from the factor zoo (Net debt financing, Change in financial liabilities, Net external financing, Inventory Growth, change in ppe and inv/assets, Asset growth) is 24 bps/month with a t-statistic of 3.19.

1 Introduction

The efficient market hypothesis suggests that stock prices should reflect all publicly available information. However, a growing body of literature documents persistent patterns in stock returns that appear to contradict market efficiency (Harvey et al., 2016). While research has extensively examined how firms' financing decisions affect their stock returns, the efficiency with which companies manage their debt issuance remains understudied. This gap is particularly notable given that debt financing represents the primary source of external capital for most public firms (Denis and Mihov, 2012).

Prior research has focused primarily on the level or changes in leverage (?), but has largely overlooked how efficiently firms time and execute their debt issuance decisions. This distinction is crucial because similar levels of debt can have dramatically different implications for firm value depending on the efficiency of the issuance process, including factors such as timing, pricing, and structure of debt offerings.

We propose that Debt Issuance Efficiency (DIE) captures valuable information about management quality and future firm performance. Our first theoretical channel builds on Ross (1977)'s signaling theory, suggesting that managers with private information about strong future prospects have both the incentive and ability to structure debt issuance more efficiently. Efficient debt issuance requires sophisticated market timing, strong banking relationships, and deep understanding of the firm's future cash flows - capabilities that correlate with superior management.

The second mechanism operates through the disciplining effect of debt (Jensen and Meckling, 1976). Firms that issue debt efficiently face lower borrowing costs and more favorable covenant terms, which creates stronger incentives for operational efficiency. This enhanced discipline should lead to improved future performance and higher stock returns.

Finally, building on Myers and Majluf (1984)'s pecking order theory, we argue

that firms demonstrating high debt issuance efficiency are likely to have exhausted internal financing options in a value-maximizing way. This suggests they are pursuing positive NPV projects that the market may not fully appreciate, leading to predictable future returns.

Our empirical analysis reveals that Debt Issuance Efficiency strongly predicts future stock returns. A value-weighted long-short portfolio strategy based on DIE quintiles generates monthly abnormal returns of 25 basis points (t-statistic = 3.31) relative to the Fama-French six-factor model. The strategy achieves an annualized gross Sharpe ratio of 0.50, placing it in the top 8% of documented cross-sectional predictors.

Importantly, the predictive power of DIE persists among large, liquid stocks. In the largest market capitalization quintile, the DIE strategy earns monthly abnormal returns of 35 basis points (t-statistic = 3.81). This suggests that the effect is not driven by small, illiquid stocks that are costly to trade.

The DIE signal maintains its predictive power even after controlling for related anomalies. When we control for the six most closely related predictors, including net debt financing and asset growth, the strategy still generates an alpha of 24 basis points per month (t-statistic = 3.19), indicating that DIE captures unique information about future returns.

Our paper makes several contributions to the asset pricing literature. First, we introduce a novel measure of financing efficiency that provides incremental predictive power beyond existing metrics. While prior work by [Bradshaw et al. \(2006\)](#) and [Lewis et al. \(2003\)](#) examined aspects of debt financing, our DIE measure uniquely captures the efficiency dimension of debt issuance decisions.

Second, we contribute to the growing literature on investment-based asset pricing ([Cochrane and Saa-Requejo, 2016](#)) by showing how the efficiency of financing decisions affects expected returns. Our findings suggest that markets do not fully

incorporate the information content of debt issuance efficiency, creating a profitable trading opportunity that survives standard risk adjustments.

Finally, our results have important implications for both corporate finance and investment practice. For corporate managers, our findings highlight the importance of debt issuance efficiency as a signal of firm quality. For investors, we document a robust return predictor that remains profitable after accounting for transaction costs and works well among large, liquid stocks.

2 Data

Our study investigates the predictive power of Debt Issuance Efficiency, a financial signal derived from accounting data for cross-sectional returns. We obtain accounting and financial data from COMPUSTAT, covering firm-level observations for publicly traded companies. To construct our signal, we use COMPUSTAT's item DLTIS, which represents the long-term debt issuance of a firm. This variable captures the amount of new long-term debt issued by companies during their fiscal year, providing insight into their financing activities and capital structure decisions. The total assets variable (AT) represents the firm's total assets, serving as a scaling factor to ensure comparability across firms of different sizes. The construction of the signal follows a dynamic approach, where we calculate the year-over-year change in debt issuance (DLTIS minus its lagged value) and scale this difference by the previous year's total assets (lagged AT). This scaled difference captures the relative magnitude of changes in debt issuance activity, normalized by firm size. By focusing on the change in debt issuance rather than absolute levels, our signal aims to identify shifts in firms' financing patterns that might signal important changes in their investment opportunities or financial constraints. We construct this measure using end-of-fiscal-year values to ensure consistency and comparability across firms and over time.

3 Signal diagnostics

Figure 1 plots descriptive statistics for the DIE signal. Panel A plots the time-series of the mean, median, and interquartile range for DIE. On average, the cross-sectional mean (median) DIE is -0.03 (-0.00) over the 1974 to 2023 sample, where the starting date is determined by the availability of the input DIE data. The signal's interquartile range spans -0.06 to 0.05. Panel B of Figure 1 plots the time-series of the coverage of the DIE signal for the CRSP universe. On average, the DIE signal is available for 6.30% of CRSP names, which on average make up 7.46% of total market capitalization.

4 Does DIE predict returns?

Table 1 reports the performance of portfolios constructed using a value-weighted, quintile sort on DIE using NYSE breaks. The first two lines of Panel A report monthly average excess returns for each of the five portfolios and for the long/short portfolio that buys the high DIE portfolio and sells the low DIE portfolio. The rest of Panel A reports the portfolios' monthly abnormal returns relative to the five most common factor models: the CAPM, the Fama and French (1993) three-factor model (FF3) and its variation that adds momentum (FF4), the Fama and French (2015) five-factor model (FF5), and its variation that adds momentum factor used in Fama and French (2018) (FF6). The table shows that the long/short DIE strategy earns an average return of 0.26% per month with a t-statistic of 3.53. The annualized Sharpe ratio of the strategy is 0.50. The alphas range from 0.25% to 0.32% per month and have t-statistics exceeding 3.31 everywhere. The lowest alpha is with respect to the FF6 factor model.

Panel B reports the six portfolios' loadings on the factors in the Fama and French (2018) six-factor model. The long/short strategy's most significant loading is 0.24,

with a t-statistic of 4.79 on the CMA factor. Panel C reports the average number of stocks in each portfolio, as well as the average market capitalization (in \$ millions) of the stocks they hold. In an average month, the five portfolios have at least 540 stocks and an average market capitalization of at least \$1,323 million.

Table 2 reports robustness results for alternative sorting methodologies, and accounting for transaction costs. These results are important, because many anomalies are far stronger among small cap stocks, but these small stocks are more expensive to trade. Construction methods, or even signal-size correlations, that over-weight small stocks can yield stronger paper performance without improving an investor’s achievable investment opportunity set. Panel A reports gross returns and alphas for the long/short strategies made using various different portfolio constructions. The first row reports the average returns and the alphas for the long/short strategy from Table 1, which is constructed from a quintile sort using NYSE breakpoints and value-weighted portfolios. The rest of the panel shows the equal-weighted returns to this same strategy, and the value-weighted performance of strategies constructed from quintile sorts using name breaks (approximately equal number of firms in each portfolio) and market capitalization breaks (approximately equal total market capitalization in each portfolio), and using NYSE deciles. The average return is lowest for the quintile sort using NYSE breakpoints and equal-weighted portfolios, and equals 22 bps/month with a t-statistics of 4.80. Out of the twenty-five alphas reported in Panel A, the t-statistics for twenty-three exceed two, and for twenty exceed three.

Panel B reports for these same strategies the average monthly net returns and the generalized net alphas of [Novy-Marx and Velikov \(2016\)](#). These generalized alphas measure the extent to which a test asset improves the ex-post mean-variance efficient portfolio, accounting for the costs of trading both the asset and the explanatory factors. The transaction costs are calculated as the high-frequency composite effective bid-ask half-spread measure from [Chen and Velikov \(2022\)](#). The net aver-

age returns reported in the first column range between -4-21bps/month. The lowest return, (-4 bps/month), is achieved from the quintile sort using NYSE breakpoints and equal-weighted portfolios, and has an associated t-statistic of -0.68. Out of the twenty-five construction-methodology-factor-model pairs reported in Panel B, the DIE trading strategy improves the achievable mean-variance efficient frontier spanned by the factor models in twenty cases, and significantly expands the achievable frontier in eighteen cases.

Table 3 provides direct tests for the role size plays in the DIE strategy performance. Panel A reports the average returns for the twenty-five portfolios constructed from a conditional double sort on size and DIE, as well as average returns and alphas for long/short trading DIE strategies within each size quintile. Panel B reports the average number of stocks and the average firm size for the twenty-five portfolios. Among the largest stocks (those with market capitalization greater than the 80th NYSE percentile), the DIE strategy achieves an average return of 35 bps/month with a t-statistic of 3.81. Among these large cap stocks, the alphas for the DIE strategy relative to the five most common factor models range from 26 to 38 bps/month with t-statistics between 2.78 and 4.16.

5 How does DIE perform relative to the zoo?

Figure 2 puts the performance of DIE in context, showing the long/short strategy performance relative to other strategies in the “factor zoo.” It shows Sharpe ratio histograms, both for gross and net returns (Panel A and B, respectively), for 212 documented anomalies in the zoo.¹ The vertical red line shows where the Sharpe ratio for the DIE strategy falls in the distribution. The DIE strategy’s gross (net) Sharpe ratio of 0.50 (0.39) is greater than 92% (95%) of anomaly Sharpe ratios,

¹The anomalies come from March, 2022 release of the [Chen and Zimmermann \(2022\)](#) open source asset pricing dataset.

respectively.

Figure 3 plots the growth of a \$1 invested in these same 212 anomaly trading strategies (gray lines), and compares those with the growth of a \$1 invested in the DIE strategy (red line).² Ignoring trading costs, a \$1 invested in the DIE strategy would have yielded \$3.62 which ranks the DIE strategy in the top 5% across the 212 anomalies. Accounting for trading costs, a \$1 invested in the DIE strategy would have yielded \$2.26 which ranks the DIE strategy in the top 5% across the 212 anomalies.

Figure 4 plots percentile ranks for the 212 anomaly trading strategies in terms of gross and Novy-Marx and Velikov (2016) net generalized alphas with respect to the CAPM, and the Fama-French three-, four-, five-, and six-factor models from Table 1, and indicates the ranking of the DIE relative to those. Panel A shows that the DIE strategy gross alphas fall between the 60 and 75 percentiles across the five factor models. Panel B shows that, accounting for trading costs, a large fraction of anomalies have not improved the investment opportunity set of an investor with access to the factor models over the 197406 to 202306 sample. For example, 45% (53%) of the 212 anomalies would not have improved the investment opportunity set for an investor having access to the Fama-French three-factor (six-factor) model. The DIE strategy has a positive net generalized alpha for five out of the five factor models. In these cases DIE ranks between the 79 and 90 percentiles in terms of how much it could have expanded the achievable investment frontier.

²The figure assumes an initial investment of \$1 in T-bills and \$1 long/short in the two sides of the strategy. Returns are compounded each month, assuming, as in Detzel et al. (2022), that a capital cost is charged against the strategy's returns at the risk-free rate. This excess return corresponds more closely to the strategy's economic profitability.

6 Does DIE add relative to related anomalies?

With so many anomalies, it is possible that any proposed, new cross-sectional predictor is just capturing some combination of known predictors. It is consequently natural to investigate to what extent the proposed predictor adds additional predictive power beyond the most closely related anomalies. Closely related anomalies are more likely to be formed on the basis of signals with higher absolute correlations. Figure 5 plots a name histogram of the correlations of DIE with 210 filtered anomaly signals.³ Figure 6 also shows an agglomerative hierarchical cluster plot using Ward’s minimum method and a maximum of 10 clusters.

A closely related anomaly is also more likely to price DIE or at least to weaken the power DIE has predicting the cross-section of returns. Figure 7 plots histograms of t-statistics for predictability tests of DIE conditioning on each of the 210 filtered anomaly signals one at a time. Panel A reports t-statistics on β_{DIE} from Fama-MacBeth regressions of the form $r_{i,t} = \alpha + \beta_{DIE}DIE_{i,t} + \beta_X X_{i,t} + \epsilon_{i,t}$, where X stands for one of the 210 filtered anomaly signals at a time. Panel B plots t-statistics on α from spanning tests of the form: $r_{DIE,t} = \alpha + \beta r_{X,t} + \epsilon_t$, where $r_{X,t}$ stands for the returns to one of the 210 filtered anomaly trading strategies at a time. The strategies employed in the spanning tests are constructed using quintile sorts, value-weighting, and NYSE breakpoints. Panel C plots t-statistics on the average returns to strategies constructed by conditional double sorts. In each month, we sort stocks into quintiles based one of the 210 filtered anomaly signals. Then, within each quintile, we sort stocks into quintiles based on DIE. Stocks are finally grouped into five DIE portfolios by combining stocks within each anomaly sorting portfolio. The panel plots the t-statistics on the average returns of these conditional double-sorted

³When performing tests at the underlying signal level (e.g., the correlations plotted in Figure 5), we filter the 212 anomalies to avoid small sample issues. For each anomaly, we calculate the common stock observations in an average month for which both the anomaly and the test signal are available. In the filtered anomaly set, we drop anomalies with fewer than 100 common stock observations in an average month.

DIE trading strategies conditioned on each of the 210 filtered anomalies.

Table 4 reports Fama-MacBeth cross-sectional regressions of returns on DIE and the six anomalies most closely-related to it. The six most-closely related anomalies are picked as those with the highest combined rank where the ranks are based on the absolute value of the Spearman correlations in Panel B of Figure 5 and the R^2 from the spanning tests in Figure 7, Panel B. Controlling for each of these signals at a time, the t-statistics on the DIE signal in these Fama-MacBeth regressions exceed 0.43, with the minimum t-statistic occurring when controlling for Asset growth. Controlling for all six closely related anomalies, the t-statistic on DIE is -0.69.

Similarly, Table 5 reports results from spanning tests that regress returns to the DIE strategy onto the returns of the six most closely-related anomalies and the six Fama-French factors. Controlling for the six most-closely related anomalies individually, the DIE strategy earns alphas that range from 25-27bps/month. The minimum t-statistic on these alphas controlling for one anomaly at a time is 3.29, which is achieved when controlling for Asset growth. Controlling for all six closely-related anomalies and the six Fama-French factors simultaneously, the DIE trading strategy achieves an alpha of 24bps/month with a t-statistic of 3.19.

7 Does DIE add relative to the whole zoo?

Finally, we can ask how much adding DIE to the entire factor zoo could improve investment performance. Figure 8 plots the growth of \$1 invested in trading strategies that combine multiple anomalies following Chen and Velikov (2022). The combinations use either the 156 anomalies from the zoo that satisfy our inclusion criteria (blue lines) or these 156 anomalies augmented with the DIE signal.⁴ We consider one different methods for combining signals.

⁴We filter the 207 Chen and Zimmermann (2022) anomalies and require for each anomaly the average month to have at least 40% of the cross-sectional observations available for market capitalization on CRSP in the period for which DIE is available.

Panel A shows results using “Average rank” as the combination method. This method sorts stocks on the basis of forecast excess returns, where these are calculated on the basis of their average cross-sectional percentile rank across return predictors, and the predictors are all signed so that higher ranks are associated with higher average returns. For this method, \$1 investment in the 156-anomaly combination strategy grows to \$935.00, while \$1 investment in the combination strategy that includes DIE grows to \$908.07.

8 Conclusion

Our comprehensive analysis of Debt Issuance Efficiency (DIE) as a predictor of stock returns yields several important conclusions. The empirical evidence strongly supports DIE’s effectiveness as a robust signal for cross-sectional equity returns prediction. The strategy’s impressive performance, achieving an annualized gross Sharpe ratio of 0.50 (0.39 net), demonstrates its practical significance for investment applications. The persistence of significant abnormal returns, even after controlling for established factors and related anomalies, suggests that DIE captures unique information about firm value that is not fully reflected in market prices.

Particularly noteworthy is the signal’s ability to generate substantial alpha (24 bps/month) even when controlling for six closely related strategies from the factor zoo, with a statistically significant t-statistic of 3.19. This finding indicates that DIE provides incremental information beyond existing measures of financing activities and asset growth.

However, several limitations warrant consideration. First, transaction costs and market impact could affect the strategy’s real-world implementation, though our analysis of net returns partially addresses this concern. Second, the signal’s effectiveness might vary across different market conditions and time periods.

Future research could explore several promising directions. Investigating the interaction between DIE and other market anomalies could yield insights into the underlying economic mechanisms. Additionally, examining the signal's performance in international markets and different asset classes would test its broader applicability. Finally, studying the impact of changing market structures and regulations on the signal's effectiveness could provide valuable insights for practitioners and academics alike.

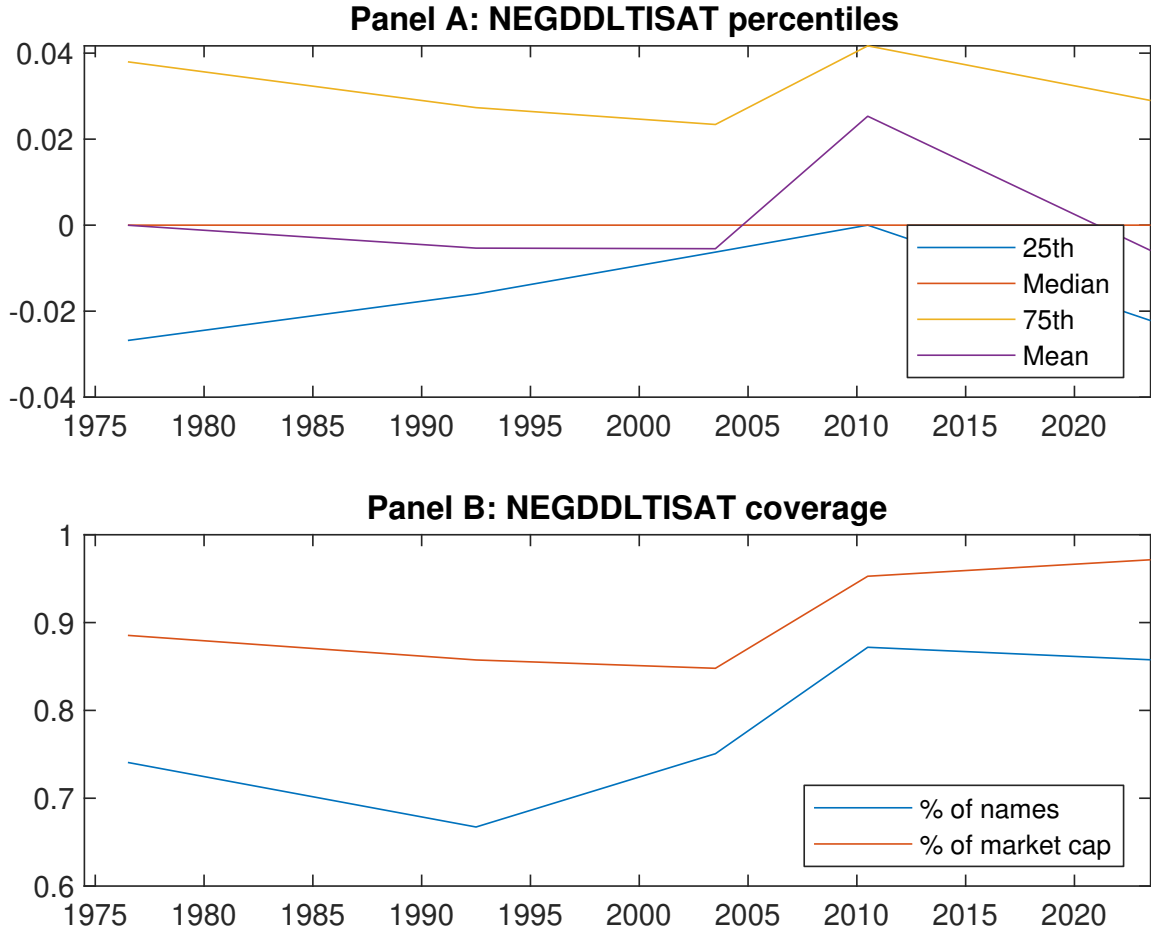


Figure 1: Times series of DIE percentiles and coverage.
This figure plots descriptive statistics for DIE. Panel A shows cross-sectional percentiles of DIE over the sample. Panel B plots the monthly coverage of DIE relative to the universe of CRSP stocks with available market capitalizations.

Table 1: Basic sort: VW, quintile, NYSE-breaks

This table reports average excess returns and alphas for portfolios sorted on DIE. At the end of each month, we sort stocks into five portfolios based on their signal using NYSE breakpoints. Panel A reports average value-weighted quintile portfolio (L,2,3,4,H) returns in excess of the risk-free rate, the long-short extreme quintile portfolio (H-L) return, and alphas with respect to the CAPM, Fama and French (1993) three-factor model, Fama and French (1993) three-factor model augmented with the Carhart (1997) momentum factor, Fama and French (2015) five-factor model, and the Fama and French (2015) five-factor model augmented with the Carhart (1997) momentum factor following Fama and French (2018). Panel B reports the factor loadings for the quintile portfolios and long-short extreme quintile portfolio in the Fama and French (2015) five-factor model. Panel C reports the average number of stocks and market capitalization of each portfolio. T-statistics are in brackets. The sample period is 197406 to 202306.

| Panel A: Excess returns and alphas on DIE-sorted portfolios | | | | | | |
|---|------------------|------------------|------------------|------------------|------------------|------------------|
| | (L) | (2) | (3) | (4) | (H) | (H-L) |
| r^e | 0.62 [2.90] | 0.66 [3.56] | 0.68 [3.35] | 0.76 [4.16] | 0.88 [4.37] | 0.26 [3.53] |
| α_{CAPM} | -0.13 [-2.26] | 0.02 [0.35] | -0.02 [-0.44] | 0.12 [2.55] | 0.18 [3.29] | 0.31 [4.16] |
| α_{FF3} | -0.12 [-2.32] | -0.03 [-0.61] | 0.02 [0.32] | 0.11 [2.23] | 0.19 [3.64] | 0.32 [4.22] |
| α_{FF4} | -0.11 [-2.00] | -0.00 [-0.01] | 0.07 [1.28] | 0.07 [1.49] | 0.17 [3.25] | 0.28 [3.72] |
| α_{FF5} | -0.13 [-2.41] | -0.06 [-1.51] | 0.07 [1.25] | 0.02 [0.41] | 0.14 [2.63] | 0.27 [3.58] |
| α_{FF6} | -0.12 [-2.19] | -0.04 [-1.02] | 0.10 [1.87] | 0.00 [0.04] | 0.13 [2.47] | 0.25 [3.31] |
| Panel B: Fama and French (2018) 6-factor model loadings for DIE-sorted portfolios | | | | | | |
| β_{MKT} | 1.05 [84.47] | 0.98 [100.13] | 0.99 [78.68] | 0.97 [89.98] | 1.01 [81.89] | -0.04 [-2.33] |
| β_{SMB} | 0.17 [8.86] | -0.14 [-9.40] | -0.01 [-0.35] | -0.04 [-2.61] | 0.16 [8.24] | -0.01 [-0.49] |
| β_{HML} | -0.03 [-1.19] | 0.16 [8.40] | -0.10 [-4.03] | 0.01 [0.65] | -0.14 [-5.74] | -0.11 [-3.22] |
| β_{RMW} | 0.09 [3.68] | 0.10 [5.26] | -0.04 [-1.69] | 0.11 [4.95] | 0.05 [1.83] | -0.05 [-1.33] |
| β_{CMA} | -0.10 [-2.75] | 0.01 [0.41] | -0.10 [-2.61] | 0.16 [5.01] | 0.14 [3.99] | 0.24 [4.79] |
| β_{UMD} | -0.02 [-1.51] | -0.04 [-3.74] | -0.06 [-4.59] | 0.03 [2.77] | 0.01 [1.05] | 0.03 [1.82] |
| Panel C: Average number of firms (n) and market capitalization (me) | | | | | | |
| n | 679 | 540 | 1076 | 594 | 648 | |
| me (\$10 ⁶) | 1349 | 3000 | 2132 | 2972 | 1323 | |

Table 2: Robustness to sorting methodology & trading costs

This table evaluates the robustness of the choices made in the DIE strategy construction methodology. In each panel, the first row shows results from a quintile, value-weighted sort using NYSE break points as employed in Table 1. Each of the subsequent rows deviates in one of the three choices at a time, and the choices are specified in the first three columns. For each strategy construction methodology, the table reports average excess returns and alphas with respect to the CAPM, Fama and French (1993) three-factor model, Fama and French (1993) three-factor model augmented with the Carhart (1997) momentum factor, Fama and French (2015) five-factor model, and the Fama and French (2015) five-factor model augmented with the Carhart (1997) momentum factor following Fama and French (2018). Panel A reports average returns and alphas with no adjustment for trading costs. Panel B reports net average returns and Novy-Marx and Velikov (2016) generalized alphas as prescribed by Detzel et al. (2022). T-statistics are in brackets. The sample period is 197406 to 202306.

| Panel A: Gross Returns and Alphas | | | | | | | | |
|--|--------|---------|--------------------|--------------------------|-------------------------|-------------------------|-------------------------|-------------------------|
| Portfolios | Breaks | Weights | r^e | α_{CAPM} | α_{FF3} | α_{FF4} | α_{FF5} | α_{FF6} |
| Quintile | NYSE | VW | 0.26 [3.53] | 0.31 [4.16] | 0.32 [4.22] | 0.28 [3.72] | 0.27 [3.58] | 0.25 [3.31] |
| Quintile | NYSE | EW | 0.22 [4.80] | 0.24 [5.37] | 0.23 [5.01] | 0.21 [4.57] | 0.22 [4.77] | 0.21 [4.51] |
| Quintile | Name | VW | 0.27 [3.74] | 0.32 [4.46] | 0.33 [4.60] | 0.29 [3.97] | 0.27 [3.73] | 0.25 [3.40] |
| Quintile | Cap | VW | 0.23 [3.60] | 0.27 [4.16] | 0.26 [4.09] | 0.22 [3.38] | 0.19 [2.91] | 0.16 [2.53] |
| Decile | NYSE | VW | 0.26 [2.61] | 0.35 [3.60] | 0.31 [3.25] | 0.27 [2.70] | 0.16 [1.71] | 0.14 [1.52] |
| Panel B: Net Returns and Novy-Marx and Velikov (2016) generalized alphas | | | | | | | | |
| Portfolios | Breaks | Weights | r_{net}^e | α_{CAPM}^* | α_{FF3}^* | α_{FF4}^* | α_{FF5}^* | α_{FF6}^* |
| Quintile | NYSE | VW | 0.20 [2.71] | 0.27 [3.53] | 0.27 [3.58] | 0.25 [3.34] | 0.24 [3.10] | 0.22 [2.92] |
| Quintile | NYSE | EW | -0.04 [-0.68] | | | | | |
| Quintile | Name | VW | 0.21 [2.89] | 0.28 [3.82] | 0.29 [3.92] | 0.26 [3.62] | 0.24 [3.29] | 0.22 [3.07] |
| Quintile | Cap | VW | 0.18 [2.79] | 0.23 [3.53] | 0.22 [3.45] | 0.20 [3.10] | 0.16 [2.48] | 0.15 [2.24] |
| Decile | NYSE | VW | 0.19 [1.86] | 0.29 [2.91] | 0.26 [2.62] | 0.23 [2.34] | 0.13 [1.39] | 0.12 [1.22] |

Table 3: Conditional sort on size and DIE

This table presents results for conditional double sorts on size and DIE. In each month, stocks are first sorted into quintiles based on size using NYSE breakpoints. Then, within each size quintile, stocks are further sorted based on DIE. Finally, they are grouped into twenty-five portfolios based on the intersection of the two sorts. Panel A presents the average returns to the 25 portfolios, as well as strategies that go long stocks with high DIE and short stocks with low DIE. Panel B documents the average number of firms and the average firm size for each portfolio. The sample period is 197406 to 202306.

| Panel A: portfolio average returns and time-series regression results | | | | | | | | | | | | |
|---|---------------|----------------|----------------|----------------|----------------|--|-----------------|----------------|----------------|----------------|----------------|----------------|
| Size quintiles | DIE Quintiles | | | | | DIE Strategies | | | | | | |
| | (L) | (2) | (3) | (4) | (H) | r^e | α_{CAPM} | α_{FF3} | α_{FF4} | α_{FF5} | α_{FF6} | |
| | (1) | 0.71 [2.54] | 0.90 [3.32] | 0.99 [3.59] | 0.88 [3.15] | 0.86 [2.98] | 0.15 [1.56] | 0.18 [1.83] | 0.16 [1.64] | 0.11 [1.15] | 0.12 [1.29] | 0.10 [1.00] |
| | (2) | 0.76 [2.79] | 0.98 [3.87] | 0.82 [3.25] | 0.93 [3.74] | 0.91 [3.55] | 0.15 [1.82] | 0.19 [2.30] | 0.17 [2.10] | 0.17 [2.09] | 0.16 [1.92] | 0.16 [1.94] |
| | (3) | 0.82 [3.19] | 0.85 [3.90] | 0.86 [3.50] | 0.88 [3.93] | 0.93 [3.97] | 0.11 [1.40] | 0.19 [2.36] | 0.18 [2.26] | 0.14 [1.73] | 0.17 [2.02] | 0.14 [1.68] |
| | (4) | 0.76 [3.28] | 0.80 [3.80] | 0.91 [4.08] | 0.74 [3.58] | 0.94 [4.26] | 0.18 [2.32] | 0.21 [2.73] | 0.21 [2.64] | 0.18 [2.24] | 0.19 [2.32] | 0.17 [2.08] |
| | (5) | 0.53 [2.63] | 0.62 [3.41] | 0.62 [3.10] | 0.63 [3.35] | 0.87 [4.53] | 0.35 [3.81] | 0.38 [4.14] | 0.38 [4.16] | 0.30 [3.34] | 0.30 [3.27] | 0.26 [2.78] |
| Panel B: Portfolio average number of firms and market capitalization | | | | | | | | | | | | |
| Size quintiles | DIE Quintiles | | | | | DIE Quintiles | | | | | | |
| | Average n | | | | | Average market capitalization (\$10 ⁶) | | | | | | |
| | (L) | (2) | (3) | (4) | (H) | (L) | (2) | (3) | (4) | (H) | | |
| | (1) | 398 | 400 | 399 | 400 | 395 | 37 | 34 | 33 | 33 | 36 | |
| | (2) | 108 | 108 | 108 | 108 | 108 | 60 | 60 | 59 | 61 | 60 | |
| | (3) | 77 | 77 | 77 | 77 | 77 | 105 | 107 | 102 | 104 | 106 | |
| | (4) | 64 | 65 | 65 | 65 | 64 | 223 | 232 | 223 | 228 | 224 | |
| (5) | 59 | 59 | 59 | 59 | 59 | 1362 | 2105 | 1709 | 2077 | 1396 | | |

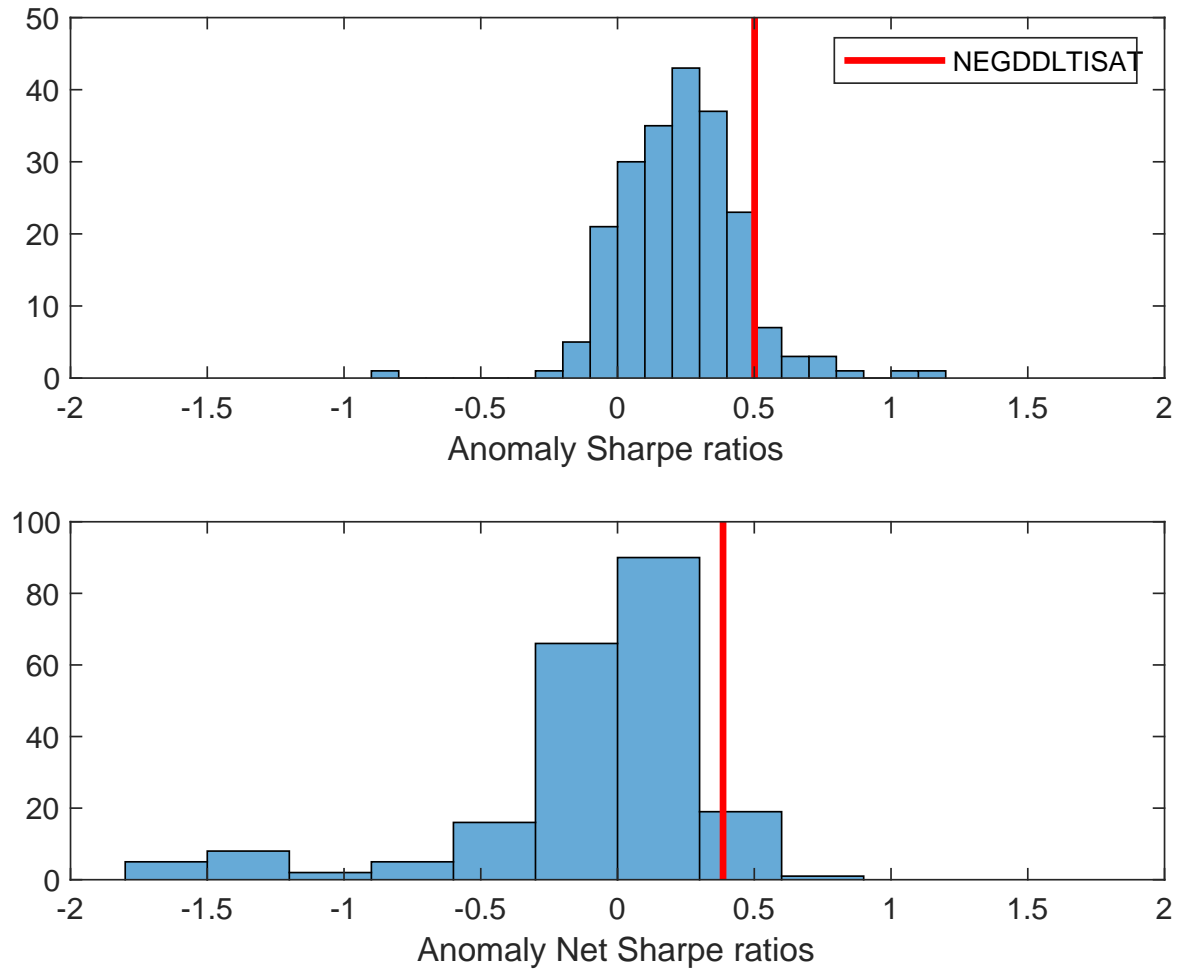


Figure 2: Distribution of Sharpe ratios.
This figure plots a histogram of Sharpe ratios for 212 anomalies, and compares the Sharpe ratio of the DIE with them (red vertical line). Panel A plots results for gross Sharpe ratios. Panel B plots results for net Sharpe ratios.

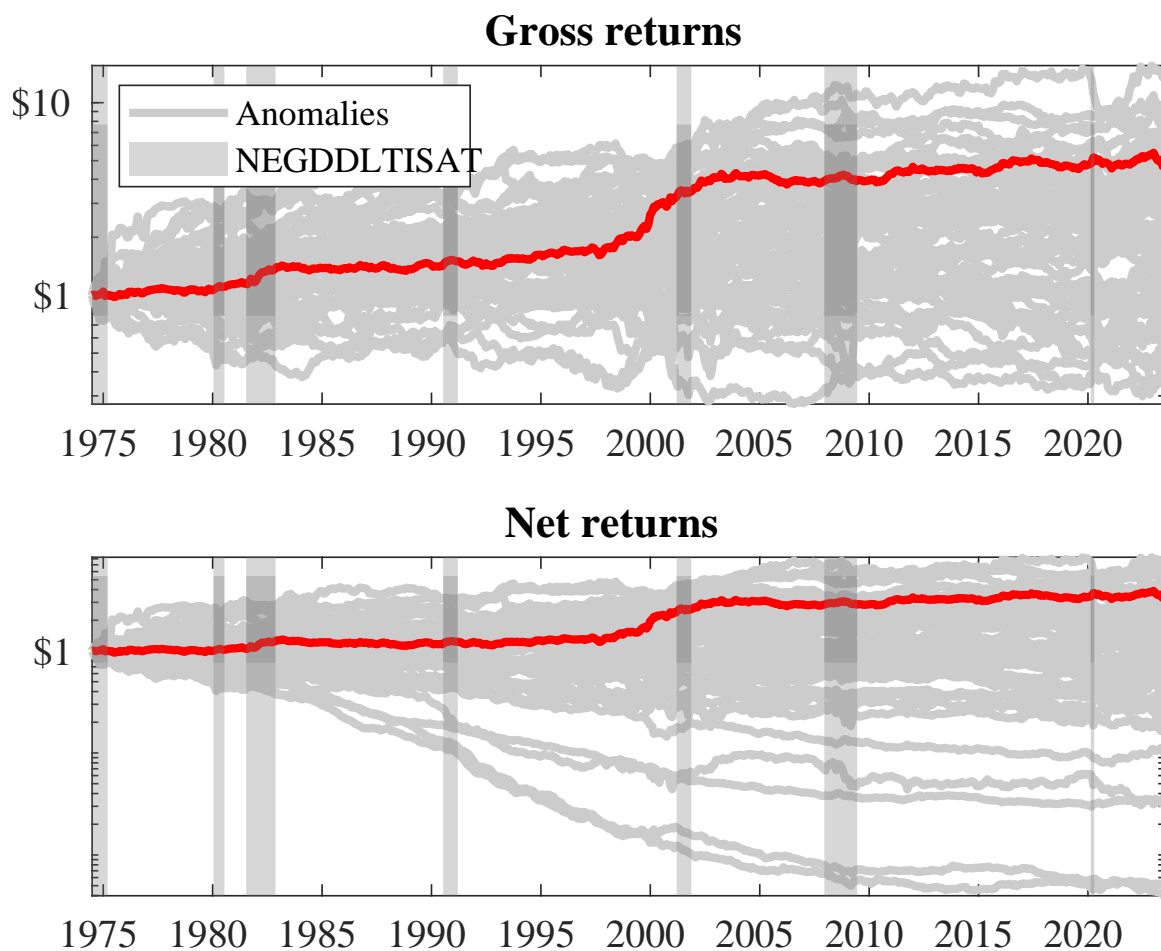


Figure 3: Dollar invested.

This figure plots the growth of a \$1 invested in 212 anomaly trading strategies (gray lines), and compares those with the DIE trading strategy (red line). The strategies are constructed using value-weighted quintile sorts using NYSE breakpoints. Panel A plots results for gross strategy returns. Panel B plots results for net strategy returns.

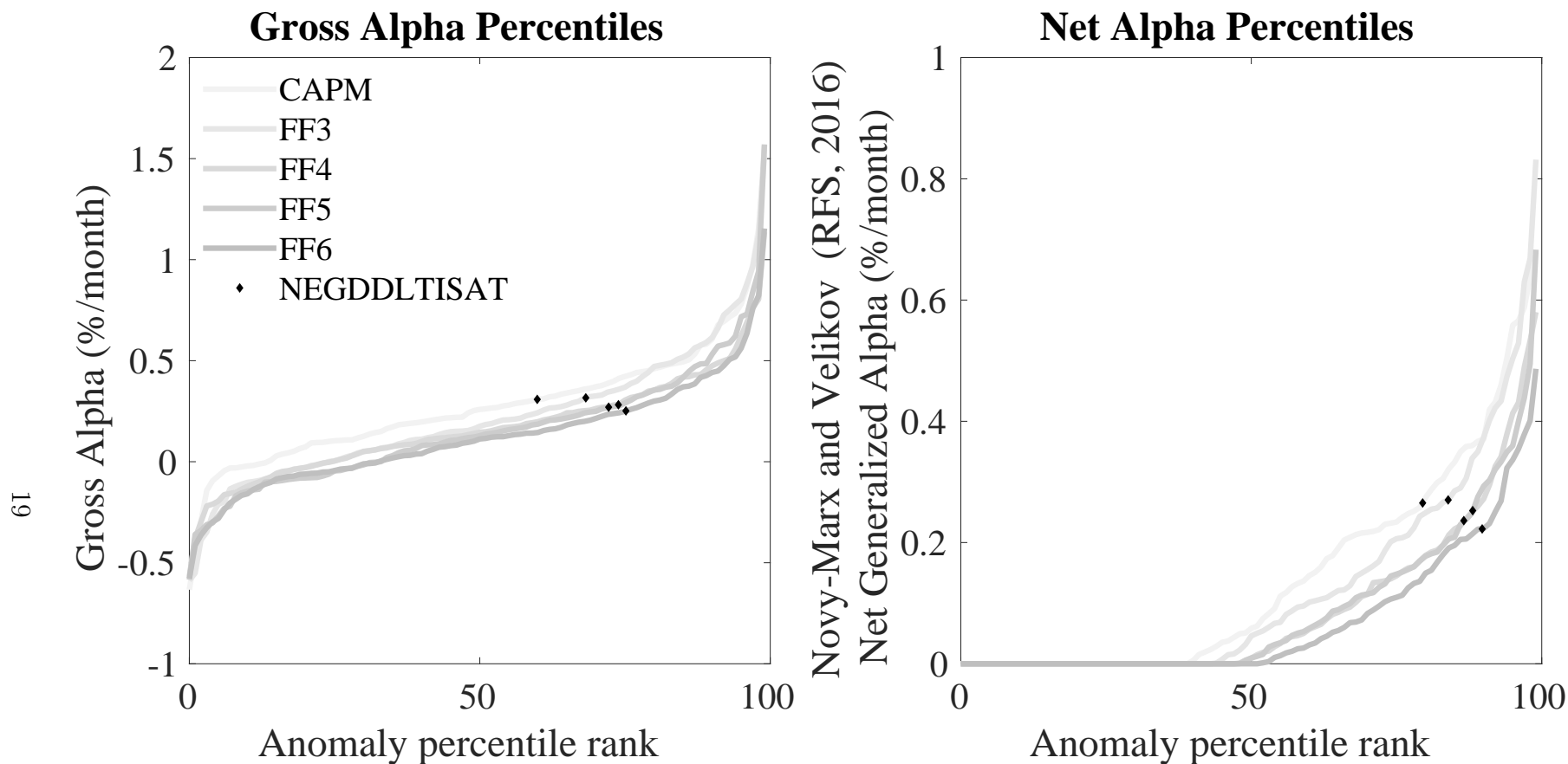


Figure 4: Gross and generalized net alpha percentiles of anomalies relative to factor models

This figure plots the percentile ranks for 212 anomaly trading strategies in terms of alphas (solid lines), and compares those with the DIE trading strategy alphas (diamonds). The strategies are constructed using value-weighted quintile sorts using NYSE breakpoints. The alphas include those with respect to the CAPM, [Fama and French \(1993\)](#) three-factor model, [Fama and French \(1993\)](#) three-factor model augmented with the [Carhart \(1997\)](#) momentum factor, [Fama and French \(2015\)](#) five-factor model, and the [Fama and French \(2015\)](#) five-factor model augmented with the [Carhart \(1997\)](#) momentum factor following [Fama and French \(2018\)](#). The left panel plots alphas with no adjustment for trading costs. The right panel plots [Novy-Marx and Velikov \(2016\)](#) net generalized alphas.

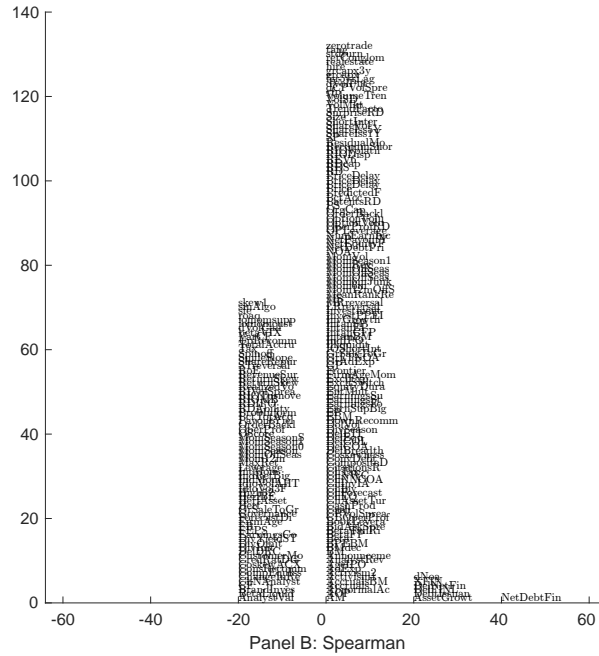
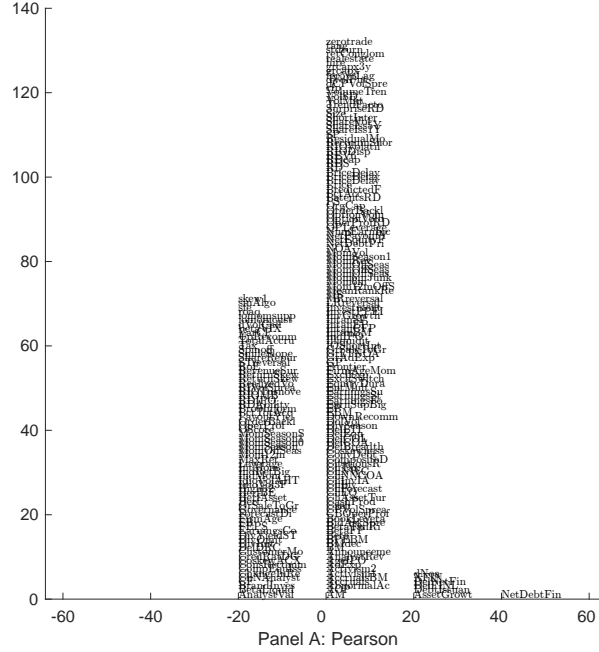


Figure 5: Distribution of correlations.

This figure plots a name histogram of correlations of 210 filtered anomaly signals with DIE. The correlations are pooled. Panel A plots Pearson correlations, while Panel B plots Spearman rank correlations.

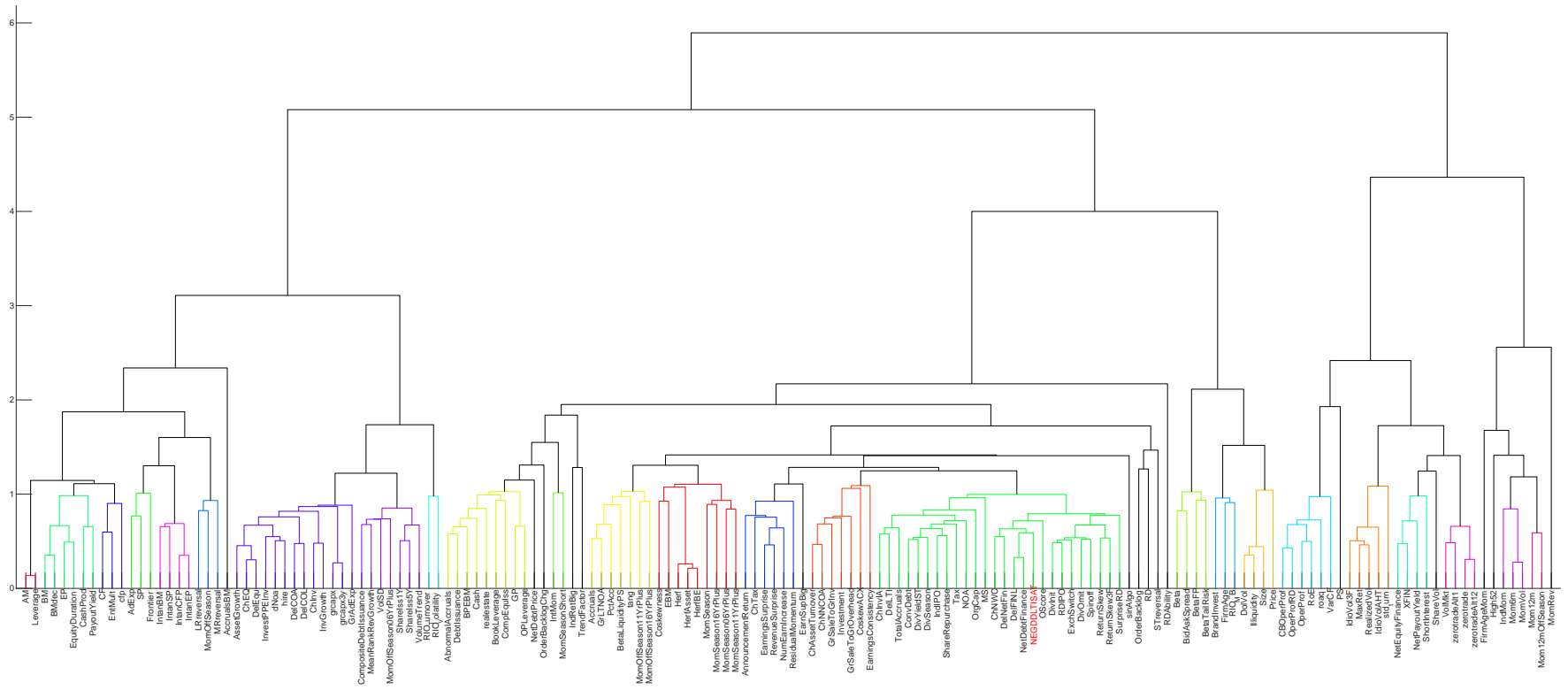


Figure 6: Agglomerative hierarchical cluster plot

This figure plots an agglomerative hierarchical cluster plot using Ward's minimum method and a maximum of 10 clusters.

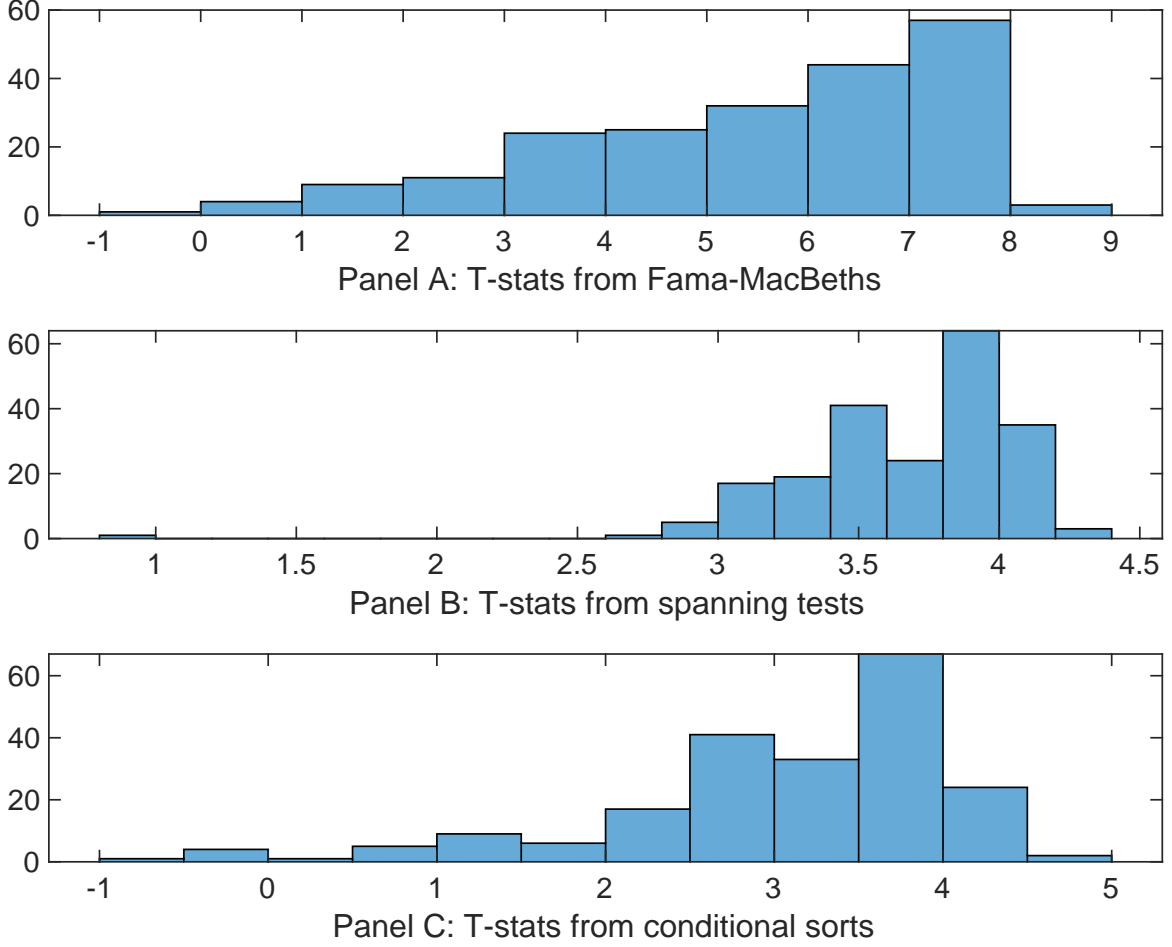


Figure 7: Distribution of t-stats on conditioning strategies

This figure plots histograms of t-statistics for predictability tests of DIE conditioning on each of the 210 filtered anomaly signals one at a time. Panel A reports t-statistics on β_{DIE} from Fama-MacBeth regressions of the form $r_{i,t} = \alpha + \beta_{DIE}DIE_{i,t} + \beta_X X_{i,t} + \epsilon_{i,t}$, where X stands for one of the 210 filtered anomaly signals at a time. Panel B plots t-statistics on α from spanning tests of the form: $r_{DIE,t} = \alpha + \beta r_{X,t} + \epsilon_t$, where $r_{X,t}$ stands for the returns to one of the 210 filtered anomaly trading strategies at a time. The strategies employed in the spanning tests are constructed using quintile sorts, value-weighting, and NYSE breakpoints. Panel C plots t-statistics on the average returns to strategies constructed by conditional double sorts. In each month, we sort stocks into quintiles based one of the 210 filtered anomaly signals at a time. Then, within each quintile, we sort stocks into quintiles based on DIE. Stocks are finally grouped into five DIE portfolios by combining stocks within each anomaly sorting portfolio. The panel plots the t-statistics on the average returns of these conditional double-sorted DIE trading strategies conditioned on each of the 210 filtered anomalies.

Table 4: Fama-MacBeths controlling for most closely related anomalies

This table presents Fama-MacBeth results of returns on DIE. and the six most closely related anomalies. The regressions take the following form: $r_{i,t} = \alpha + \beta_{DIE}DIE_{i,t} + \sum_{k=1}^s \beta_{X_k}X_{i,t}^k + \epsilon_{i,t}$. The six most closely related anomalies, X , are Net debt financing, Change in financial liabilities, Net external financing, Inventory Growth, change in ppe and inv/assets, Asset growth. These anomalies were picked as those with the highest combined rank where the ranks are based on the absolute value of the Spearman correlations in Panel B of Figure 5 and the R^2 from the spanning tests in Figure 7, Panel B. The sample period is 197406 to 202306.

| | | | | | | | |
|-----------------|----------------|----------------|----------------|----------------|----------------|----------------|------------------|
| Intercept | 0.14 [5.49] | 0.14 [5.52] | 0.14 [5.86] | 0.14 [5.47] | 0.15 [5.88] | 0.15 [5.95] | 0.15 [5.89] |
| DIE | 0.15 [1.58] | 0.13 [1.40] | 0.24 [2.24] | 0.54 [5.35] | 0.24 [2.68] | 0.43 [0.43] | -0.77 [-0.69] |
| Anomaly 1 | 0.20 [8.25] | | | | | | 0.12 [1.87] |
| Anomaly 2 | | 0.17 [8.92] | | | | | -0.84 [-1.76] |
| Anomaly 3 | | | 0.18 [5.81] | | | | 0.10 [1.82] |
| Anomaly 4 | | | | 0.38 [6.67] | | | 0.11 [0.18] |
| Anomaly 5 | | | | | 0.17 [7.60] | | 0.76 [2.76] |
| Anomaly 6 | | | | | | 0.11 [8.97] | 0.57 [2.72] |
| # months | 588 | 588 | 588 | 588 | 588 | 588 | 588 |
| $\bar{R}^2(\%)$ | 0 | 0 | 1 | 0 | 0 | 0 | 0 |

Table 5: Spanning tests controlling for most closely related anomalies

This table presents spanning tests results of regressing returns to the DIE trading strategy on trading strategies exploiting the six most closely related anomalies. The regressions take the following form: $r_t^{DIE} = \alpha + \sum_{k=1}^6 \beta_{X_k} r_t^{X_k} + \sum_{j=1}^6 \beta_{f_j} r_t^{f_j} + \epsilon_t$, where X_k indicates each of the six most-closely related anomalies and f_j indicates the six factors from the [Fama and French \(2015\)](#) five-factor model augmented with the [Carhart \(1997\)](#) momentum factor. The six most closely related anomalies, X , are Net debt financing, Change in financial liabilities, Net external financing, Inventory Growth, change in ppe and inv/assets, Asset growth. These anomalies were picked as those with the highest combined rank where the ranks are based on the absolute value of the Spearman correlations in Panel B of Figure 5 and the R^2 from the spanning tests in Figure 7, Panel B. The sample period is 197406 to 202306.

| | | | | | | | |
|-----------------|-------------------|------------------|-------------------|-------------------|-------------------|-------------------|------------------|
| Intercept | 0.25 [3.31] | 0.25 [3.30] | 0.25 [3.29] | 0.26 [3.48] | 0.26 [3.45] | 0.27 [3.47] | 0.24 [3.19] |
| Anomaly 1 | 20.21 [4.76] | | | | | | 9.68 [1.65] |
| Anomaly 2 | | 18.42 [4.14] | | | | | 10.43 [1.75] |
| Anomaly 3 | | | 13.99 [3.62] | | | | 8.72 [2.06] |
| Anomaly 4 | | | | 9.55 [3.18] | | | 7.83 [2.49] |
| Anomaly 5 | | | | | 10.64 [3.04] | | 5.14 [1.36] |
| Anomaly 6 | | | | | | 5.63 [1.13] | -4.71 [-0.90] |
| mkt | -4.28 [-2.48] | -4.06 [-2.34] | -2.39 [-1.31] | -4.54 [-2.60] | -4.53 [-2.59] | -4.25 [-2.42] | -3.33 [-1.84] |
| smb | -2.78 [-1.03] | -3.07 [-1.13] | 3.07 [1.04] | -0.45 [-0.17] | -1.24 [-0.46] | -1.99 [-0.72] | 1.17 [0.37] |
| hml | -10.22 [-3.08] | -9.57 [-2.87] | -8.94 [-2.65] | -10.84 [-3.23] | -11.77 [-3.48] | -10.69 [-3.16] | -9.65 [-2.86] |
| rmw | -5.98 [-1.73] | -5.74 [-1.65] | -12.70 [-3.05] | -3.01 [-0.86] | -3.95 [-1.14] | -4.31 [-1.23] | -9.87 [-2.33] |
| cma | 18.32 [3.57] | 17.37 [3.31] | 14.13 [2.49] | 15.00 [2.62] | 15.20 [2.64] | 16.48 [2.05] | 6.62 [0.83] |
| umd | 1.57 [0.88] | 1.45 [0.80] | 3.17 [1.81] | 2.35 [1.32] | 3.12 [1.77] | 3.45 [1.93] | 0.45 [0.25] |
| # months | 588 | 588 | 588 | 588 | 588 | 588 | 588 |
| $\bar{R}^2(\%)$ | 11 | 10 | 9 | 9 | 8 | 7 | 13 |

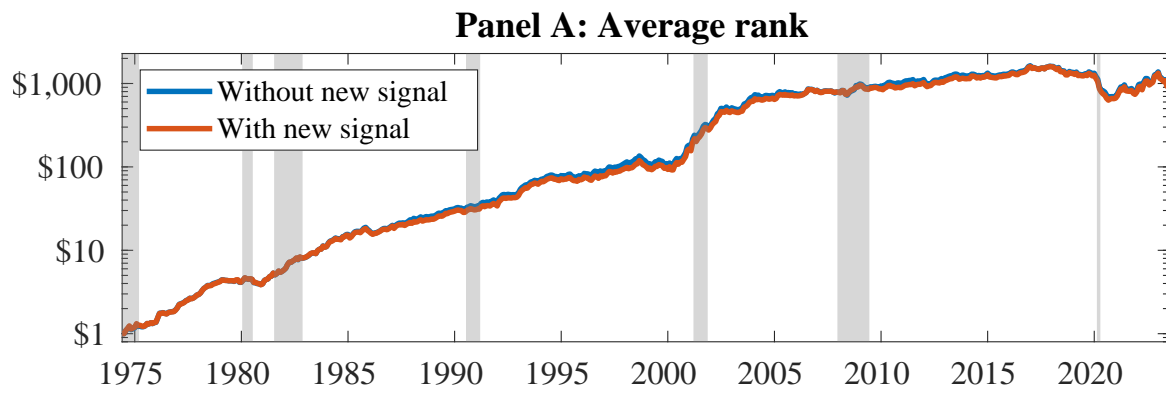


Figure 8: Combination strategy performance

This figure plots the growth of a \$1 invested in trading strategies that combine multiple anomalies following [Chen and Velikov \(2022\)](#). In all panels, the blue solid lines indicate combination trading strategies that utilize 156 anomalies. The red solid lines indicate combination trading strategies that utilize the 156 anomalies as well as DIE. Panel A shows results using "Average rank" as the combination method. See [Section 7](#) for details on the combination methods.

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