Introduction of SVCJ

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P measure

According to Hou et al. (2020) and Asgharian et al. (2006), the stochastic volatility with correlated jumps (SVCJ) model can be written as:

•
$$d\log S_t = \mu dt + \sqrt{V_t} dW_t^S + Z_t^y dN_t$$

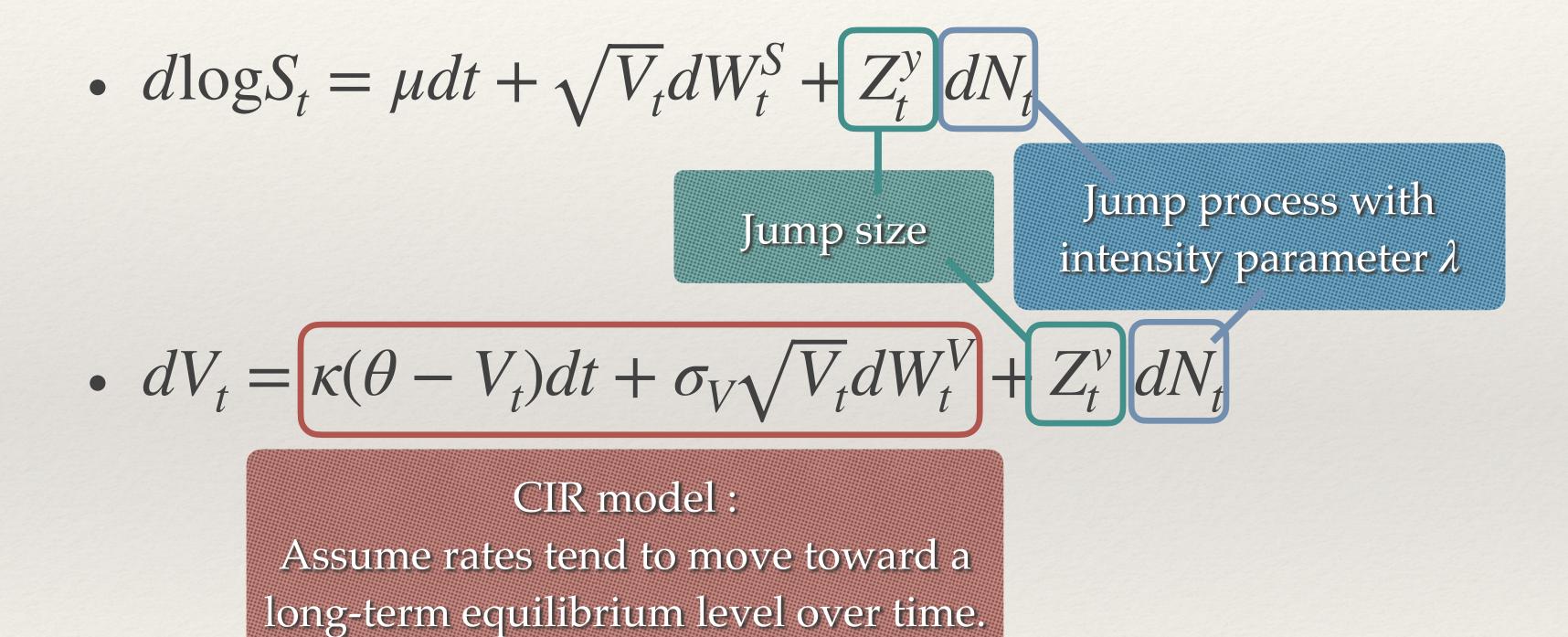
•
$$dV_t = \kappa(\theta - V_t)dt + \sigma_V \sqrt{V_t}dW_t^V + Z_t^v dN_t$$

•
$$Cov(dW_t^S, dW_t^V) = \rho dt$$

•
$$P(dN_t = 1) = \lambda dt$$

•
$$Z_t^y | Z_t^v \sim N(\mu_y + \rho_j Z_t^v, \sigma_y^2); Z_t^v \sim \text{Exp}(\mu_v)$$

Definition



Definition

Expected log return

•
$$d\log S_t = \mu dt + \sqrt{V_t} dW_t^S + Z_t^y dN_t$$

Mean reversion rate

Volatility of process

•
$$dV_t = \kappa \theta - V_t dt + \sigma_V \sqrt{V_t} dW_t^V + Z_t^V dN_t$$

Mean reversion level

Definition

•
$$d\log S_t = \mu dt + \sqrt{V} dW_t^S + Z_t^y dN_t$$

Brownian motion with correlation ρ $dW \sim N(0,1)$

•
$$dV_t = \kappa(\theta - V_t)dt + \sigma_V \sqrt{V_t} dW_t^V + Z_t^v dN_t$$

- Utilizing Bayesian inference to obtain the posterior distribution of its parameters by Markov Chain Monte Carlo (MCMC) method.
- Discretization interval in one day, then the models become:

$$\begin{pmatrix} Y_{(t+1)\Delta} \\ V_{(t+1)\Delta} \end{pmatrix} = \begin{pmatrix} \mu \\ \alpha + \left(\frac{1}{\Delta} + \beta\right) V_{t\Delta} \end{pmatrix} \Delta + \sqrt{V_{t\Delta}\Delta} \begin{pmatrix} \varepsilon_{(t+1)\Delta}^{Y} \\ \sigma_{V} \varepsilon_{(t+1)\Delta}^{V} \end{pmatrix} + \begin{pmatrix} Z_{(t+1)\Delta}^{Y} \\ Z_{(t+1)\Delta}^{V} \end{pmatrix} J_{(t+1)\Delta}$$

Parameters: $\Theta = \{\mu, \mu_y, \sigma_y, \lambda, \alpha, \beta, \sigma_v, \rho, \rho_j, \mu_v\}$

Latent variables : $X_t = \{V_t, Z_t^y, Z_t^y, J_t\}$

Parameters

$$\begin{split} p(\mu) &= \frac{1}{\sqrt{50\pi}} e^{-\frac{\mu^2}{50}} \to \mu \sim N(0,25) \\ p(\alpha,\beta) &= \frac{1}{2\pi} e^{-\frac{1}{2}(\alpha^2 + \beta^2)} \to (\alpha,\beta) \sim N(0_{2\times 1},I_{2\times 2}) \\ p(\sigma_v^2) &= \frac{0.1^{2.5}}{\Gamma(2.5)} (\frac{1}{\sigma_v^2})^{\alpha+1} e^{-\frac{\beta}{\sigma_v^2}} \to \sigma_v^2 \sim IG(2.5,0.1) \\ p(\mu_y) &= \frac{1}{\sqrt{200\pi}} e^{-\frac{\mu_y^2}{200}} \to \mu_y \sim N(0,100) \\ p(\sigma_y^2) &= \frac{40^{10}}{\Gamma(10)} (\frac{1}{\sigma_y^2})^{\alpha+1} e^{-\frac{\beta}{\sigma_y^2}} \to \sigma_y^2 \sim IG(10,40) \\ p(\rho) &= \frac{1}{2}, \quad -1 \le \rho \le 1 \to \rho \sim U(-1,1) \\ p(\rho_j) &= \frac{1}{\sqrt{8\pi}} e^{-\frac{\rho_j^2}{8}} \to \rho_j \sim N(0,4) \\ p(\mu_v) &= \frac{20^{10}}{\Gamma(10)} (\frac{1}{\mu_v})^{\alpha+1} e^{-\frac{\beta}{\mu_v}} \to \mu_v \sim IG(10,20) \\ p(\lambda) &= \frac{\Gamma(42)}{\Gamma(2)\Gamma(40)} \lambda (1 - \lambda)^{39}, \ 0 \le \lambda \le 1 \to \lambda \sim Beta(2,40) \end{split}$$

Latent variables

$$\begin{split} p(V_{t}|\,V_{t-1},Z_{t}^{v},J_{t},\alpha,\beta,\sigma_{v}) &= \frac{1}{\sqrt{2\pi(\sigma_{v}^{2}V_{t-1})}}e^{-\frac{(V_{t}-(\alpha+\beta V_{t-1}+Z_{t}^{y}J_{t}))^{2}}{2(\sigma_{v}^{2}V_{t-1})}} \to V_{t} \sim N(\alpha+\beta V_{t-1}+Z_{t}^{v}J_{t},\sigma_{v}^{2}V_{t-1}) \\ p(Z_{t}^{v}|\,\mu_{v}) &= \frac{1}{\mu_{v}}e^{-\frac{Z_{t}^{v}}{\mu_{v}}} \to Z_{t}^{v} \sim Exp(\mu_{v}) \\ p(Z_{t}^{v}|\,Z_{t}^{v},\mu_{y},\rho_{j},\sigma_{y}^{2}) &= \frac{1}{\sqrt{2\pi\sigma_{y}^{2}}}e^{-\frac{(Z_{t}^{v}-(\mu_{y}+\rho_{j}Z_{t}^{y}))^{2}}{2\sigma_{y}^{2}}} \to Z_{t}^{v}|\,Z_{t}^{v} \sim N(\mu_{y}+\rho_{j}Z_{t}^{v},\sigma_{y}^{2}) \\ p(J_{t}=j\,|\,\lambda) &= \lambda^{j}(1-\lambda)^{1-j} \to J_{t} \sim Ber(\lambda) \end{split}$$

$$\begin{pmatrix} Y_{(t+1)\Delta} \\ V_{(t+1)\Delta} \end{pmatrix} = \begin{pmatrix} \mu \\ \alpha + \left(\frac{1}{\Delta} + \beta\right) V_{t\Delta} \end{pmatrix} \Delta + \sqrt{V_{t\Delta}\Delta} \begin{pmatrix} \varepsilon_{(t+1)\Delta}^{Y} \\ \sigma_{V}\varepsilon_{(t+1)\Delta}^{V} \end{pmatrix} + \begin{pmatrix} Z_{(t+1)\Delta}^{Y} \\ Z_{(t+1)\Delta}^{V} \end{pmatrix} J_{(t+1)\Delta}$$

$$\kappa \theta \qquad 1 - \kappa$$

 $J_{(t+1)\Delta} = 1$ indicates a jump arrival which occurs with probability $\Delta \lambda$

- The Bayesian formula gives the posterior distribution.
- For example, finding the posterior distribution of μ .

$$p(\mu \mid Y, V) \propto p(Y, V \mid \mu) p(\mu)$$

$$\mu^{(i+1)} \mid Y, \alpha^{(i)}, \beta^{(i)}, \mu_y^{(i)}, \sigma_y^{2(i)}, \rho_j^{(i)}, \mu_v^{(i)}, \lambda^{(i)}, \rho^{(i)}, \sigma_v^{2(i)}, X^{(i)} \sim N(a, A)$$

$$, where \ a = A(\frac{\Delta}{1 - \rho^2} \sum_{t=1}^{T} \frac{e_{Y,t}^{\mu} - \frac{\rho}{\sigma_v} e_{V,t}^{\mu}}{V_{t-1}}), \ A = (\frac{\Delta^2}{1 - \rho^2} \sum_{t=1}^{T} \frac{1}{V_{t-1}} + \frac{1}{25})^{-1}$$

- Because our posterior distribution generally has no close-form, using MCMC to approximate probability distribution.
- Since we have many variables, Gibb Sampling can help us to deal with high-dimensional distribution.

MCMC algorithm

$$\begin{split} \mu^{(i+1)} \,|\, Y, \alpha^{(i)}, \beta^{(i)}, \mu_y^{(i)}, \sigma_y^{2(i)}, \rho_j^{(i)}, \mu_v^{(i)}, \lambda^{(i)}, \rho^{(i)}, \sigma_v^{2(i)}, X^{(i)} \sim N(a, A) \\ \alpha^{(i+1)} \,|\, Y, \mu^{(i+1)}, \beta^{(i)}, \mu_y^{(i)}, \sigma_y^{2(i)}, \rho_j^{(i)}, \mu_v^{(i)}, \lambda^{(i)}, \rho^{(i)}, \sigma_v^{2(i)}, X^{(i)} \sim N(b, B) \\ \vdots \end{split}$$

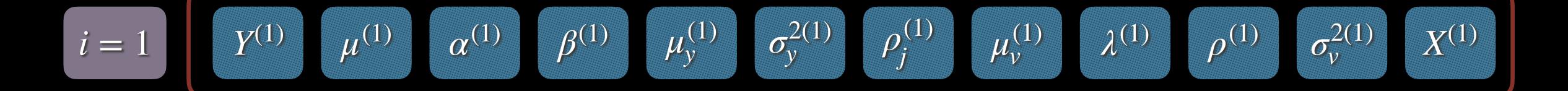
Volatility (for t = 1, ..., T), using a random walk Hastings-Metropolis algorithm for V_t

$$f(V_t) \propto V_t^{-1} e^{-\frac{1}{2} \sum_{i=0}^{1} (\frac{Y_{t+i} - \mu \Delta - Z_{t+i}^y J_{t+i} - \frac{\rho}{\sigma_v} V_{t+i} - V_{t-1+i} \Delta - \beta V_{t-1+i} \Delta - Z_{t+i}^v J_{t+i}}{(1 - \rho^2) V_{t-1+i} \Delta} + \frac{V_{t+i} - V_{t-1+i} - \alpha \Delta - \beta V_{t-1+i} \Delta - Z_{t+i}^v J_{t+i}}{\sigma_v^2 V_{t-1+i} \Delta})}$$

Proposal:
$$V_t^{(i+1)} = V_t^{(i)} + \varepsilon_t$$
 where $\varepsilon_t \sim t(df = 4.5)$

Acceptance rate :
$$A(V_t^{(i)}, V_t^{(i+1)}) = \min(\frac{f(V_t^{(i+1)})q(-\varepsilon_t)}{f(V_t^{(i)})q(\varepsilon_t)}, 1)$$

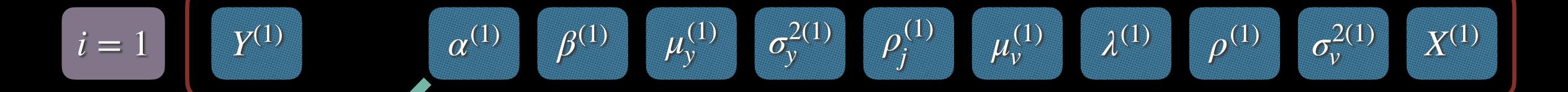
Thus, we can draw $V_1^{i+1}, V_2^i, \dots V_T^i$



 $\mu^{(2)}$

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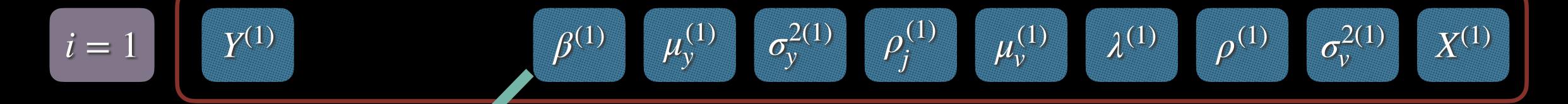
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 $\mu^{(2)}$

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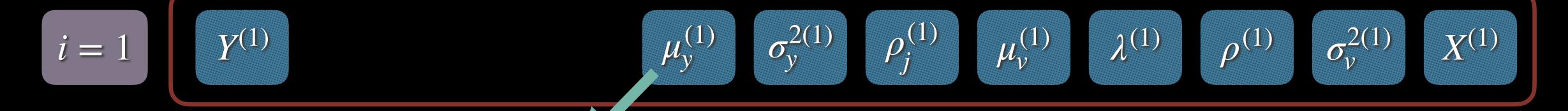
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 $2 \qquad \mu^{(2)} \qquad \alpha^{(2)}$

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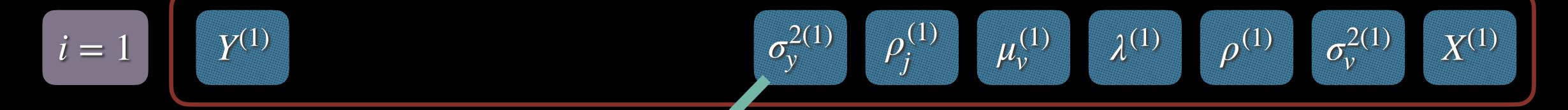
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 $2 \qquad \qquad \beta^{(2)}$

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•



 $2 \qquad \qquad \mu^{(2)} \qquad \alpha^{(2)} \qquad \beta^{(2)} \qquad \qquad \mu_y^{(2)}$

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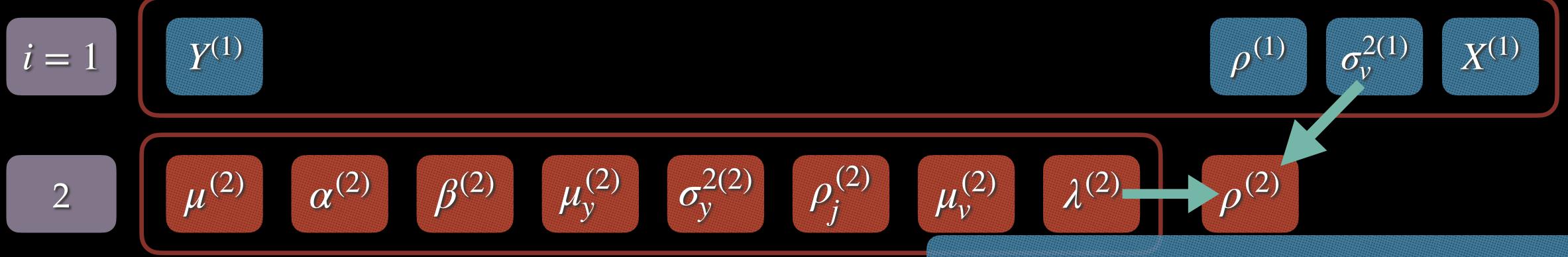


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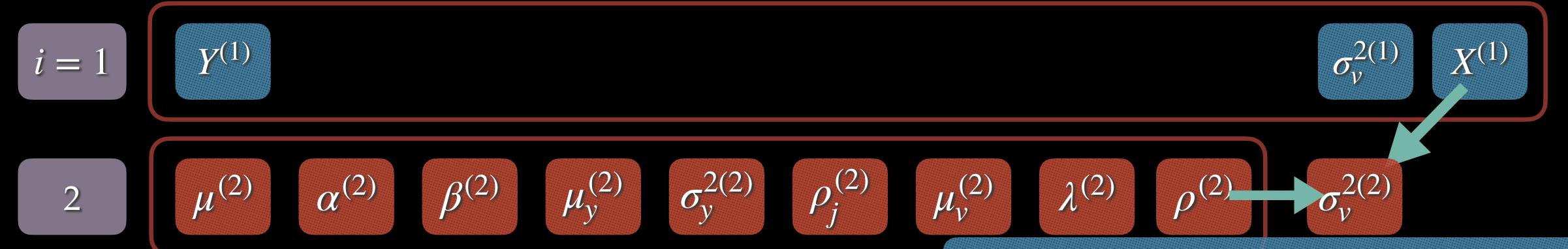


Random walk Hastings-Metropolis algorithm

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Independent Hastings-Metropolis algorithm

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 $\Theta^{(2)}$

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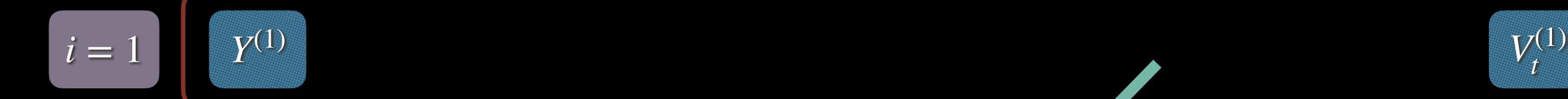
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 $\Theta^{(2)} \qquad J_t^{(2)} \qquad Z_t^{v(2)} \qquad V_1^{(2)}, \cdots, V_T^{(2)} \qquad Y_1^{(2)}, \cdots, Y_T^{(2)}$

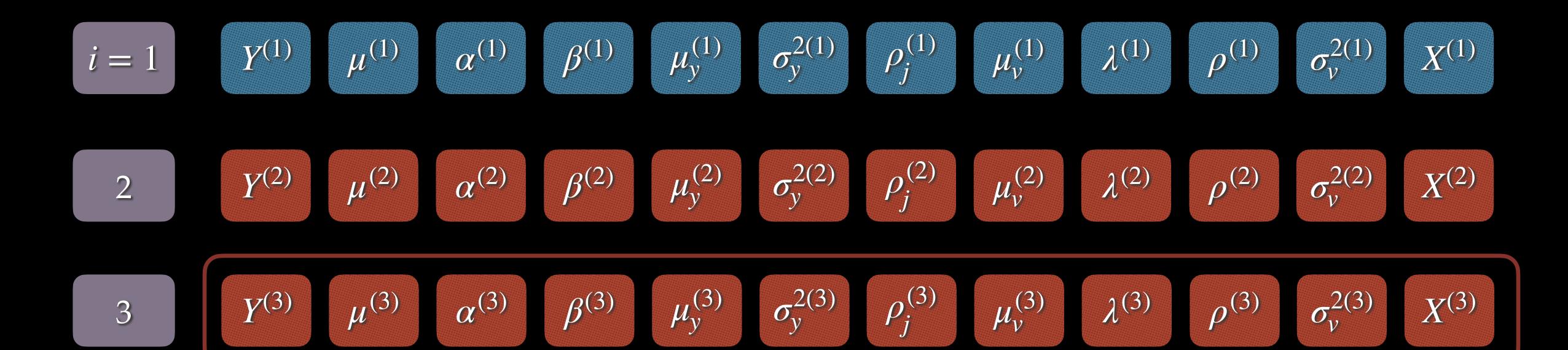
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