

Estimating Number of Distinct Elements

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Streaming Model

- The input consists of m objects/items/tokens e_1, e_2, \dots, e_m that are seen one by one by the algorithm.
- The algorithm has “limited” memory say for B tokens where $B < m$ (often $B \ll m$) and hence cannot store all the input
- Want to compute interesting functions over input

Distinct Elements

How many distinct items in the stream of integers? Here we know that each token is a positive integer from $[n] = \{1, 2, \dots, n\}$.

- Input stream: e_1, \dots, e_m .
- We associate a frequency vector $f = (f_1, \dots, f_n)$.
- f_i is the frequency of the element i in the input stream.
- We want to estimate $|\{f_i > 0 : i \in [n]\}|$

$n = 7$

2 4 2 3 1 6

$f_1 \quad f_2 \quad f_3 \quad f_4 \quad f_5 \quad f_6 \quad f_7$
(0 2 1 1 0 1 0) = f .

$$e_1 \ e_2 \ \dots \ e_m$$

$$f = (f_1 \ \dots \ f_m)$$

$$0 \rightarrow 0$$

$$x^0 \geq 1$$

$$\|f\|_1 = \sum_{i=1}^m |f_i| \leftarrow \text{length of the stream}$$

$$O(\log \log m)$$

$$e, \delta$$

$$\|f\|_0 = \sum_{i=1}^m f_i^0 \leftarrow \# \text{ distinct elements}$$

$$\|f\|_2 = \left(\sum_{i=1}^m f_i^2 \right)^{1/2}$$

$$\|f\|_\infty = \max_i |f_i|$$

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Obvious: counter for each $i \in \{1, \dots, n\}$.

Space complexity $O(n \log m)$. — bits.

$$O(\log n + \log m)$$

Distinct Elements: Our objective

$$d = \|f\|_0$$

- We will discuss a $(O(1), \delta)$ -estimate.

$$\Pr \left[\frac{3}{4}d \leq \hat{d} \leq \frac{5}{4}d \right] \geq 1 - \delta$$

Distinct Elements: Our objective

- We will discuss a $(O(1), \delta)$ -estimate.
- Let d be the no. of distinct elements.
- Then the algorithm will output \hat{d} with the following guarantee.

$$\Pr\left[\frac{d}{3} \leq \hat{d} \leq 3d\right] \geq 1 - \delta.$$

[Flajolet and Martin' 85], [Alon, Matias and Szegedy' 99]

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- In this algorithm, we use
 - Pairwise Independent Hash family ✓
 - Median Trick and Chernoff bound ✓

Probabilistic Inequalities and Pairwise Independent Hash Family

Markov's Inequality

Let X be a **non-negative** random variable over a probability space (Ω, \Pr) . For any $a > 0$,

$$\Pr[X \geq a] \leq \frac{\mathbf{E}[X]}{a}.$$

In other words, for any $t > 0$,

$$\Pr[X \geq t\mathbf{E}[X]] \leq \frac{1}{t}.$$

Chebyshev's Inequality: Variance

Variance

Given a random variable X over probability space (Ω, \Pr) , variance of X is the measure of how much does it deviate from its mean value. Formally,

$$\text{Var}(X) = \mathbf{E}[(X - \mathbf{E}[X])^2] = \mathbf{E}[X^2] - \mathbf{E}[X]^2. \checkmark$$

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Independence

Random variables X and Y are called mutually independent if

$$\forall x, y \in \mathbb{R}, \Pr[X = x \wedge Y = y] = \Pr[X = x] \Pr[Y = y]$$

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Lemma

If X and Y are independent random variables then

$$\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y).$$

Chebyshev's Inequality

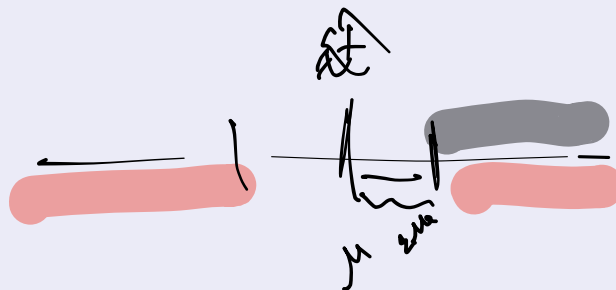
If $\text{Var}(X) < \infty$, then for any $a \geq 0$,

$$\Pr[|X - \mathbf{E}[X]| \geq a] \leq \frac{\text{Var}(X)}{a^2}.$$

Recap: Chernoff bound

Let X_1, \dots, X_k be k independent random variables such that, for each $i \in \{1, \dots, k\}$, X_i equals 1 with probability p_i , and 0 with probability $(1 - p_i)$. Let $X = \sum_{i=1}^k X_i$ and $\mu = \mathbf{E}[X] = \sum_i p_i$. For any $0 < \varepsilon < 1$, it holds that:

- $\Pr[X \geq (1 + \varepsilon)\mu] \leq e^{\frac{-\varepsilon^2 \mu}{3}}$
- $\Pr[X \leq (1 - \varepsilon)\mu] \leq e^{\frac{-\varepsilon^2 \mu}{2}}$



For $0 < \varepsilon < 1$ and $\mu_{\min} < \mu < \mu_{\max}$,

- $\Pr[X \geq (1 + \varepsilon)\mu_{\max}] \leq e^{\frac{-\varepsilon^2 \mu_{\max}}{3}}$
- $\Pr[X \leq (1 - \varepsilon)\mu_{\min}] \leq e^{\frac{-\varepsilon^2 \mu_{\min}}{2}}$



Pairwise Independent Hash Family

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- Want the hash function to behave like a “random function” and has a compact representation.

Let $Z_1, Z_2 \dots Z_n$
 are n random variables
 over (Ω, \mathcal{F}) . We say
 that $Z_1, Z_2 \dots Z_n$ are
 pairwise independent if
 for any two distinct i and
 j and any two values
 a and b .

$$P[Z_i = a \wedge Z_j = b] = P[Z_i = a] \cdot P[Z_j = b]$$

Pairwise Independent Hash Family

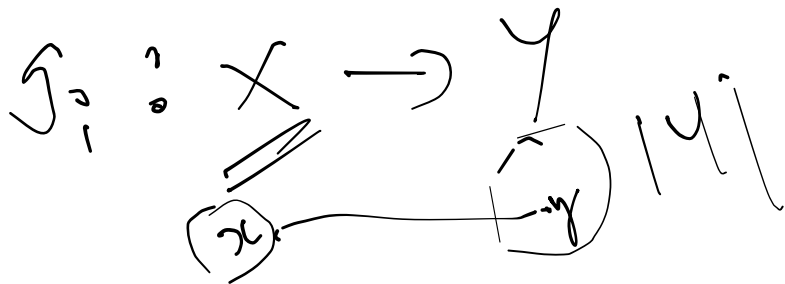
- Hash functions are used in various fields in the CS.
- Want the hash function to behave like a “random function” and has a compact representation.
- A family of hash functions $\mathcal{H} \subseteq \{f: X \rightarrow Y\}$, is a Pairwise Independent Hash Family if the following two conditions hold.

$$\begin{aligned} & \Omega = \mathbb{Z}_2 \\ & \forall \text{ any } h \in \mathcal{H}, p_\delta[h] = \frac{1}{|\mathbb{Z}_2|} \\ & (\Omega, p_\delta) \end{aligned}$$

Pairwise Independent Hash Family

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- Want the hash function to behave like a “random function” and has a compact representation. $\nabla \quad \neg$
- A family of hash functions $\mathcal{H} \subseteq \{g: X \rightarrow Y\}$, is a Pairwise Independent Hash Family if the following two conditions hold.
 - Uniformly distributed: for ~~any~~ any $x \in X$ and $y \in Y$,

$$\mathcal{H} = \{g_1, g_2, \dots, g_k\} \quad \Pr_{h \sim \mathcal{H}} [h(x) = y] = \frac{1}{|Y|}.$$

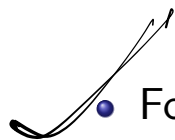
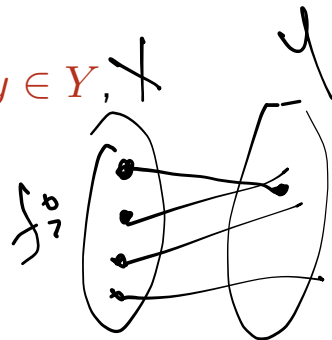


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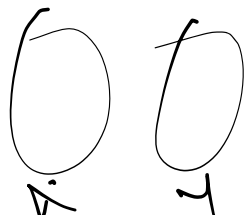
$$\Pr_{h \sim \mathcal{H}}[h(x) = y] = \frac{1}{|Y|}.$$



- For any $x, x' \in X$ and $y, y' \in Y$ s.t $x \neq x'$,

$$\Pr_{h \sim \mathcal{H}}[h(x) = y \wedge h(x') = y'] = \frac{1}{|Y|^2}.$$

$$= \Pr_{h \sim \mathcal{H}}[h(x) = y] \cdot \Pr_{h \sim \mathcal{H}}[h(x') = y']$$



Example: Pairwise Independent Hash Family

Let $X = \{0, 1\}^N$ and $Y = \{0, 1\}^K$ where $K \leq N$.

- For a matrix $A \in \{0, 1\}^{K \times N}$ and vector $\underline{b} \in \{0, 1\}^K$, define $h_{A,b}: X \rightarrow Y$ as follows:

$$h_{A,b}(x) = (Ax + b) \bmod 2.$$

$N=3$

x

000
001
010
011
100
101
110
111

00
01
10
11

$K=2$

y

$$A = \begin{bmatrix} 0 & 0 & 0 \\ 0 & 0 & 0 \end{bmatrix}$$

$$b = \begin{bmatrix} 1 \\ 1 \end{bmatrix}$$

$$(A \cdot x + b) \bmod 1$$

$$h_{A,b} = \begin{bmatrix} (0 \ 0 \ 0)^T \\ (0 \ 0 \ 1)^T \\ (0 \ 1 \ 0)^T \\ (0 \ 1 \ 1)^T \end{bmatrix}$$

$$(0 \ 0)^T$$

$$(0 \ 1)^T$$

$$\begin{bmatrix} 0 & 0 & 0 \\ 0 & 0 & 0 \end{bmatrix} \begin{bmatrix} 0 \\ 0 \\ 1 \end{bmatrix} \rightarrow \begin{bmatrix} 1 \\ 1 \end{bmatrix}$$

$$= \begin{bmatrix} 1 \\ 1 \end{bmatrix} \bmod 1$$

$$1 \ 0 \ 0$$

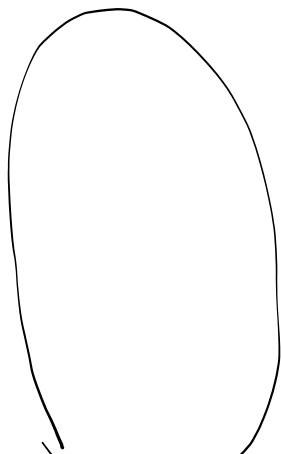
$$1 \ 0 \ 1$$

$$1 \ 1 \ 0$$

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$$1 \ 0$$

$$1 \ 1$$



Example: Pairwise Independent Hash Family

$$l = \log n \quad l = \log n.$$

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$$h_{A,b}(x) = (Ax + b) \bmod 2.$$

- $\mathcal{H} = \{h_{A,b}: A \in \{0, 1\}^{K \times N}, b \in \{0, 1\}^K\}$ is a pairwise independent hash family.

$$| \mathcal{H} | = 2^{K \times N} \cdot 2^K$$

$$H \quad (\Omega, p_x)$$

$$\Omega = 27, \quad p_x[h_{A,b}] = \frac{1}{|27|}$$

Tidemark Algorithm

Notation

Q. 10101110010000
zeros 4.

For an integer $p > 0$, $\text{zeros}(p)$ is the number of zeros that the binary representation of p ends with. That is,

$$\text{zeros}(p) = \max\{i : 2^i \text{ divides } p\}.$$

$$p = 5$$

$$101$$

$$\text{zeros}(5) = 0$$

$$p = 6$$

$$110$$

$$\text{zeros}(6) = 1$$

$$p = 8$$

$$1000$$

$$\text{zeros}(8) = 3$$

$$\{0,1\}^l \rightarrow \{0,1\}^d, \quad n = 2^d.$$

Algorithm 1: Tidemark Algorithm

\mathcal{H} is a pairwise independent hash family from $[n]$ to $[n]$;

choose h at random from \mathcal{H} ;

$z \leftarrow 0$;

while a new token e_j arrives **do**

if $\text{zeros}(h(e_j)) > 0$ **then**

$z \leftarrow \text{zeros}(h(e_j))$

end

end

return $2^{z+\frac{1}{2}}$

$$(A, b)$$

$$\{1, 2, \dots, n\} = \log^2 n + \log n.$$

$$\{1, 2, \dots, n\}.$$

$$Ax + b =$$

$$h(e_j)$$

$$e_1, \dots, e_m$$

$$\{1, 2, \dots, n\}$$

Intuition

$$\begin{array}{cccccc}
 & e_1 & e_2 & - & e_3 & e_4 & e_5 - e_6 \\
 h: & 1 & 2 & & 1 & 2 & 3 & 4
 \end{array}$$

$$\begin{array}{cccccc}
 & 3 & 3 & & 3 & 3 & 2 & 4 \\
 \text{zeros} & 0 & 0 & & 0 & 0 & 1 & 2 \\
 \hline
 & & & & 0 & 0 & 0 & 1, 2
 \end{array}$$

$2 = 0, \quad 2 = 0$

$$\begin{array}{cc}
 e_7: & 1 \\
 h(i) & 3 \\
 & 0
 \end{array}$$

$2 + \frac{1}{2}$

$2L - 2$

$2 = 4 \cancel{+ 5} 2 \quad 4 \cancel{+ 5} 2$

Space complexity

$h_{A,b}$

$A \in \{0,1\}^{d \times d}$

Algorithm 2: Tidemark Algorithm

$d = \log n$

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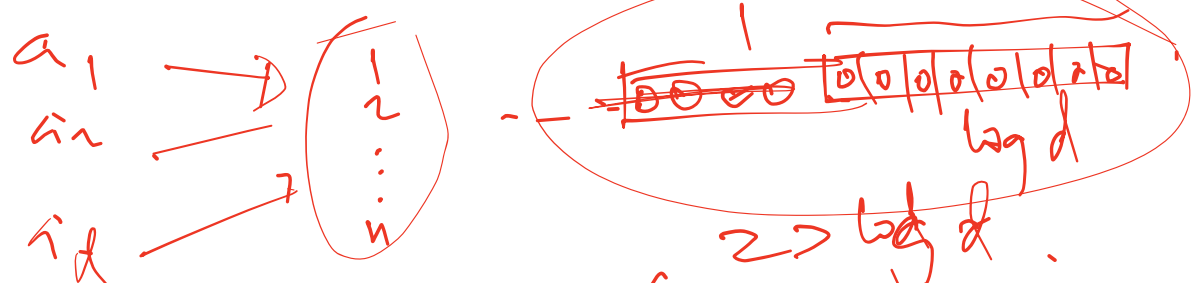
end

return $2^{z+\frac{1}{2}}$

$h: [n] \rightarrow [n]$
 $O(\log^2 n)$
 $\log n \times \log n$

$2^{\log d}$
 $d \cdot 52$

distinct elements in d



Space complexity = $O(\log^2 n)$

Analysis: $(\underline{O(1)}, \frac{1}{\sqrt{2}})$ -Estimate

- For each integer $t \in [n]$ and each integer $r \geq 0$, $X_{r,t}$ be the indicator random variable s.t.

$$X_{r,t} = \begin{cases} 1 & \text{if } \underline{\text{zeros}(h(t))} \geq r \\ 0 & \text{Otherwise} \end{cases}$$

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- $Y_r = \sum_{t: f_t > 0} X_{r,t}$.

1, 5, 8, 9

$$Y_r = X_{r,1} + X_{r,5} + X_{r,8} + X_{r,9}$$

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$2^{T+1/2}$

- Let T be the value of z at the end of the algorithm.

Analysis: $(O(1), \frac{1}{\sqrt{2}})$ -Estimate

- For each integer $t \in [n]$ and each integer $r \geq 0$, $X_{r,t}$ be the indicator random variable s.t.

$$X_{r,t} = \begin{cases} 1 & \text{if } \text{zeros}(h(t)) \geq r \\ 0 & \text{Otherwise} \end{cases}$$

$$\underline{\underline{x_2 = 3}}$$

- $Y_r = \sum_{t: f_t > 0} X_{r,t}$

$$X_{3,1} = 0$$

$$X_{1,1} = 1$$

$$X_{1,8} = 1$$

$$X_{3,4} = 1$$

- Let T be the value of z at the end of the algorithm.

- Then, $Y_r \geq 0$ iff $T \geq r$.

$$\text{seq: } 1, 8, 2, 1, 2$$

- Equivalently, $Y_r = 0$ iff $T \leq r - 1$.

$$h(): 2 \quad 1 \quad 3 \quad 2 \quad 3$$

$$\text{zero: } 1 \quad 0 \quad 0 \quad 1 \quad 0$$

$$4$$

$$h[4] = 9$$

Fix an γ :

Suppose $\gamma_\gamma = 0$

Then all the numbers γ_α seen
in the stream is mapped to
numbers that has less than or
equal to $\gamma-1$ # of zeros

\Rightarrow The value in z
will be at most
 $\gamma-1$

Expectation and Variance of Y_r

- $\mathbf{E}[X_{r,t}] = \Pr[\underbrace{\text{zeros}(h(t)) \geq r}_{\substack{\text{0} \\ \text{0 0 0 0 0} \\ \text{2}}}] = \Pr[2^r \text{ divides } h(t)] = \frac{1}{2^r}$

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- $\mathbf{E}[Y_r] = \sum_{t: f_t > 0} \mathbf{E}[X_{r,t}] = \frac{d}{2^r}$ (Here d is the no. of distinct elements)

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$\text{Var}[Y_r] = \sum_{t: f_t > 0} \text{Var}[X_{r,t}] \leq \sum_{t: f_t > 0} \mathbf{E}[X_{r,t}^2] = \frac{d}{2^r}$

~~ok~~

~~1, 5, 10~~

are pairwise independent.

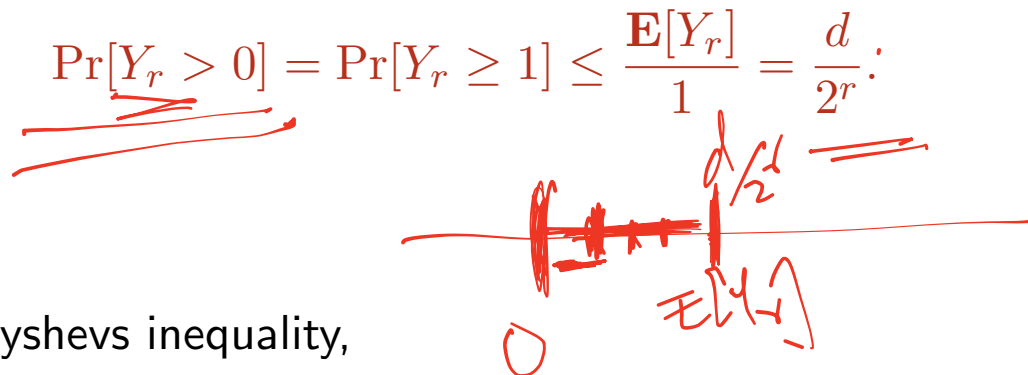
Applying Markov's and Chebyshevs inequalities

- By Markov's inequality,

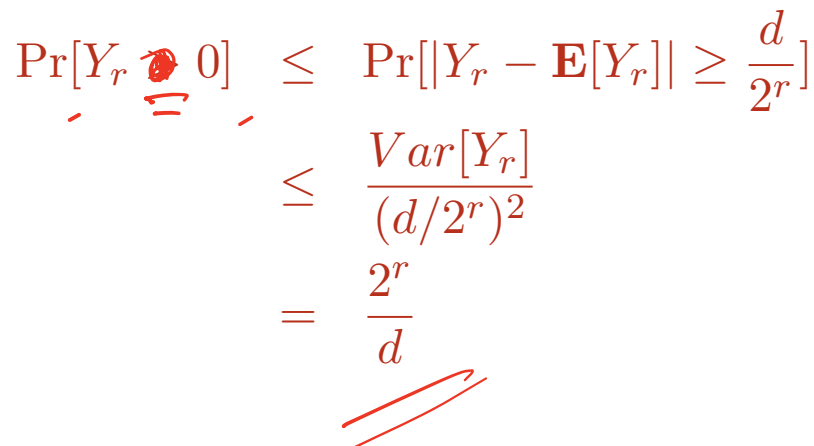
$$\Pr[Y_r > 0] = \Pr[Y_r \geq 1] \leq \frac{\mathbf{E}[Y_r]}{1} = \frac{d}{2^r}.$$

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- By Markov's inequality,

$$\Pr[Y_r > 0] = \Pr[Y_r \geq 1] \leq \frac{\mathbf{E}[Y_r]}{1} = \frac{d}{2^r}.$$


- By Chebyshevs inequality,

$$\begin{aligned} \Pr[Y_r \neq 0] &\leq \Pr[|Y_r - \mathbf{E}[Y_r]| \geq \frac{d}{2^r}] \\ &\leq \frac{\text{Var}[Y_r]}{(d/2^r)^2} \\ &= \frac{2^r}{d} \end{aligned}$$


Quality of the estimate \hat{d}

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$$\frac{1}{2} + \frac{1}{2} \geq 3$$

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- $\hat{d} = 2^{T+\frac{1}{2}}$.
- Let a be the smallest integer such that $2^{a+\frac{1}{2}} \geq 4d$.
- Then, $\Pr[\hat{d} \geq 3d] \leq \Pr[T \geq a] = \Pr[Y_a > 0] \leq \frac{d}{2^a} \leq \frac{\sqrt{2}}{4}$.
- Let b be the largest integer such that $2^{b+\frac{1}{2}} \leq d/4$.
- Then, $\Pr[\hat{d} \leq d/4] \leq \Pr[T \leq b] = \Pr[Y_{b+1} = 0] \leq \frac{2^{b+1}}{d} \leq \frac{\sqrt{2}}{4}$.
- Then, by union bound,

$$\Pr[d/4 \leq \hat{d} \leq 4d] \geq 1 - \frac{1}{\sqrt{2}}.$$

Error reduction via median trick

We have:

$$\Pr[\hat{d} \geq 4d \text{ or } \hat{d} \leq d/4] \leq \frac{1}{\sqrt{2}}$$

Want:

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Idea: Repeat independently $\ell = 12 \log(2/\delta)$ times.

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Idea: Repeat independently $\ell = 12 \log(2/\delta)$ times.

Algorithm: Output median of the estimates $Q^{(1)}, Q^{(2)}, \dots, Q^{(\ell)}$.

Error reduction via median trick

Let Z be median of the $\ell = 12 \log(2/\delta)$ independent estimators.

Lemma

$$\Pr[Z > 4d] \leq \delta/2.$$

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- Let X_i be that random variable that takes value 1 when A_i happens and 0 otherwise.

Error reduction via median trick

Let Z be median of the $\ell = 12 \log(2/\delta)$ independent estimators.

Lemma

$$\Pr[Z > 4d] \leq \delta/2.$$

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- For median estimate to be bad, more than half of A_i 's have to be bad.
- Let X_i be that random variable that takes value 1 when A_i happens and 0 otherwise.
- Let $X = \sum_{i=1}^{\ell} X_i$.
- Our output is “bad” if and only if X is at least $\ell/2$.

Applying Chernoff bound

- Let X_1, \dots, X_k be k independent 0/1-random variables,
- $X = \sum_{i=1}^k X_i$, and
- $\mathbf{E}[X] \leq \mu_{\max}$.

Then, for any $0 < \varepsilon < 1$, it holds that:

- $\Pr[X \geq (1 + \varepsilon)\mu] \leq e^{\frac{-\varepsilon^2 \mu_{\max}}{3}}$

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$$\begin{aligned}\Pr[X \geq \ell/2] &\leq \Pr[X \geq 2\mu_{\max}] \\ &\leq \Pr[X \geq (1 + 0.99)\mu_{\max}] \\ &\leq e^{\frac{-(0.99)^2 \mu_{\max}}{3}} \\ &\leq e^{\frac{-(0.99)^2 \sqrt{2}\ell}{12}}\end{aligned}$$

Choose $\ell = 12 \cdot (\log \frac{1}{\delta})$. Then, $\Pr[X \geq \ell/2] \leq \delta$

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Space complexity: $O(\log(1/\delta) \log^2 n)$. ✓

Summary

- We have seen estimating number of distinct elements
- We used pairwise independent hash family
- Median Trick

Thank You.