

Advanced Computer Graphics

Rendering Equation

Matthias Teschner

Computer Science Department
University of Freiburg

Albert-Ludwigs-Universität Freiburg



Outline

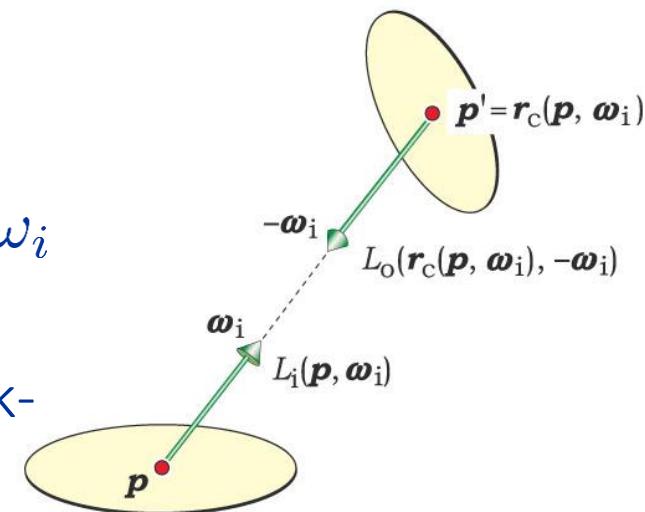
- rendering equation
- Monte Carlo integration
- sampling of random variables

Reflection and Rendering Equation

- reflection equation at point p for reflective surfaces
 - $L_o(p, \omega_o) = \int_{2\pi^+} f_r(p, \omega_i, \omega_o) L_i(p, \omega_i) \cos \theta_i d\omega_i$
 - incident radiance - weighted with the BRDF - is integrated over the hemisphere to compute the outgoing radiance
 - expresses energy balance between surfaces
 - outgoing radiance from a surface can be incident to another surface
- rendering equation at point p for reflective surfaces
 - $L_o(p, \omega_o) = L_e(p, \omega_o) + \int_{2\pi^+} f_r(p, \omega_i, \omega_o) L_i(p, \omega_i) \cos \theta_i d\omega_i$
 - adds emissive surfaces to the reflection equation
 - exitant radiance is the sum of emitted and reflected radiance
 - expresses the steady state of radiance in a scene including light sources

Ray-Casting Operator

- in general, the incoming radiance is not only determined by light sources, but also by outgoing radiance of reflective surfaces
- incident radiance $L_i(p, \omega_i)$ can be computed by tracing a ray from p into direction ω_i
- ray-casting operator $r_c(p, \omega_i)$
 - nearest hit point from p into direction ω_i
 - $L_i(p, \omega_i) = L_o(r_c(p, \omega_i), -\omega_i)$
 - if no surface is hit, radiance from a background or light source can be returned



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Rendering Equation with Ray-Casting Operator

- using the ray-casting operator,

$$L_o(p, \omega_o) = L_e(p, \omega_o) + \int_{2\pi^+} f_r(p, \omega_i, \omega_o) L_i(p, \omega_i) \cos \theta_i d\omega_i$$

can be rewritten as

$$L_o(p, \omega_o) = L_e(p, \omega_o) + \int_{2\pi^+} f_r(p, \omega_i, \omega_o) L_o(r_c(p, \omega_i), -\omega_i) \cos \theta_i d\omega_i$$

- goal: computation of the outgoing radiance $L_o(p, \omega_o)$ at all points p into all directions ω_o
 - towards the camera to compute the image
 - towards other surface points to account for indirect illumination

Forms of the Rendering Equation

- hemisphere form

$$L_o(p, \omega_o) = L_e(p, \omega_o) + \int_{2\pi^+} f_r(p, \omega_i, \omega_o) L_o(r_c(p, \omega_i), -\omega_i) \cos \theta_i d\omega_i$$

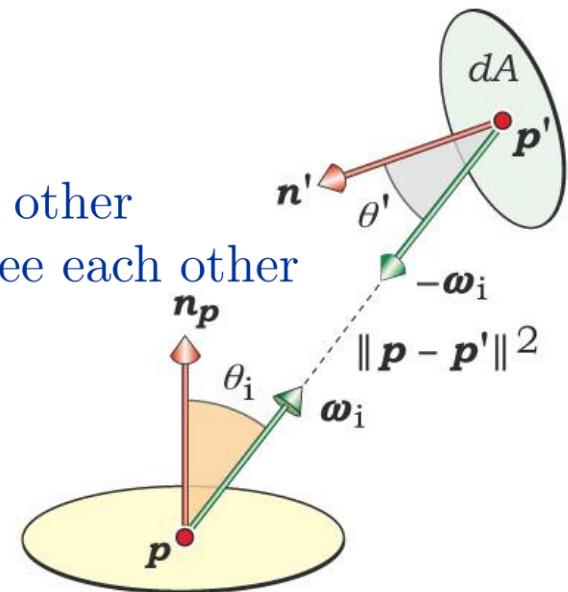
- area form

- p is a sample point on a surface dA
- visibility function

$$\forall (p, p') : V(p, p') = \begin{cases} 1 & \text{if } p \text{ and } p' \text{ see each other} \\ 0 & \text{if } p \text{ and } p' \text{ do not see each other} \end{cases}$$

- solid angle vs. area $d\omega_i = \frac{\cos \theta' dA}{\|p' - p\|^2}$
- $\cos \theta' = n' \cdot -\omega_i$

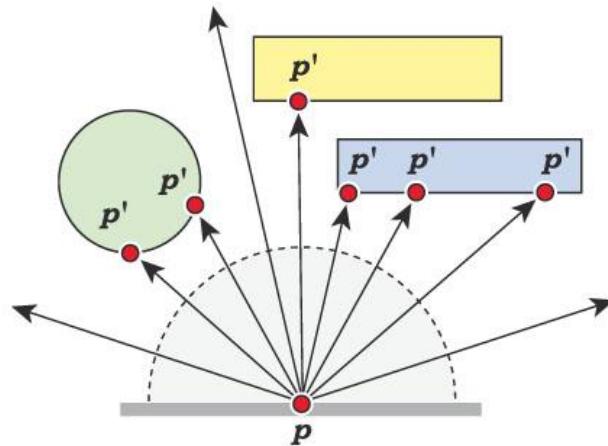
$$L_o(p, \omega_o) = L_e(p, \omega_o) + \int_A f_r(p, \omega_i, \omega_o) L_o(p', -\omega_i) \frac{\cos \theta_i \cos \theta'}{\|p' - p\|^2} V(p, p') dA$$



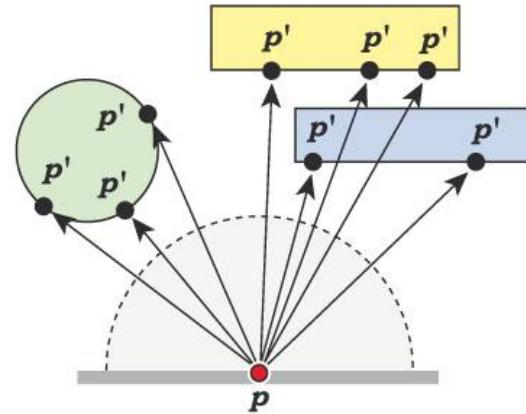
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Forms of the Rendering Equation

- the area form works with a visibility term
 - useful for direct illumination from area lights
- the hemisphere form works with the ray-casting operator
 - useful for indirect illumination



hemisphere form



area form

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Solving the Rendering Equation

- recursively cast rays into the scene
- maximum recursion depth due to absorption of light
- for point lights, directional lights, perfect reflection and transmission, the integrals reduce to simple sums
 - radiance from only a few directions contributes to the outgoing radiance
- for area lights and indirect illumination, i. e. diffuse-diffuse light transport, Monte Carlo techniques are used to numerically evaluate the multi-dimensional integrals

Outline

- rendering equation
- Monte Carlo integration
- sampling of random variables

Introduction

- approximately evaluate the integral
$$\int_{2\pi^+} f_r(p, \omega_i, \omega_o) L_o(r_c(p, \omega_i), -\omega_i) \cos \theta_i d\omega_i$$
 by
 - randomly sampling the hemisphere
 - tracing rays into the sample directions
 - computing the incoming radiance from the sample directions
- challenge
 - approximate the integral as exact as possible
 - trace as few rays as possible
 - trace relevant rays
 - for diffuse surfaces, rays in normal direction are more relevant than rays perpendicular to the normal
 - for specular surfaces, rays in reflection direction are relevant
 - rays to light sources are relevant

Properties

- benefits
 - processes only evaluations of the integrand at arbitrary points in the domain
 - works for a large variety of integrands, e.g., it handles discontinuities
 - appropriate for integrals of arbitrary dimensions
- drawbacks
 - using n samples, the scheme converges to the correct result with $O(n^{1/2})$, i.e. to half the error, $4n$ samples are required
 - errors are perceived as noise, i.e. pixels are arbitrarily too bright or dark
 - evaluation of the integrand at a point is expensive

Continuous Random Variables

- continuous random variables X infinite number of possible values
(in contrast to discrete random variable)
- canonical uniform random variable $0 \leq \xi < 1$
 - samples from arbitrary distributions can be computed from ξ
- probability density function (PDF) $p(x)$
 - the probability of a random variable taking certain value ranges
 - $Pr(x_0 \leq X \leq x_1) = \int_{x_0}^{x_1} p(x)dx$ The probability, that the random variable has a certain exact value, is 0.
 - $p(x) \geq 0 \quad \forall x \in [a, b]$
 - $\int_a^b p(x)dx = 1$ The probability, that the random variable is in the specified domain, is 1.
- cumulative distribution function (CDF) $P(x)$
 - describes the probability of a random variable to be less or equal to x
 - $Pr(X \leq x) = P(x)$ $Pr(x_0 \leq X \leq x_1) = P(x_1) - P(x_0)$
 - $0 \leq P(x) \leq 1$

Expected Value

- motivation: expected value of an estimator function is equal to the integral in the rendering equation
- expected value $E_p[f(x)]$ of a function $f(x)$ is defined as the weighted average value of the function over a domain D
- $E_p[f(x)] = \int_D f(x) p(x) dx$ with $\int_D p(x) dx = 1$
processes an infinite number of samples x according to a PDF $p(x)$
- properties
 - $E[af(x)] = aE[f(x)]$
 - $E[\sum_i f(X_i)] = \sum_i E[f(X_i)]$ for independent random variables X_i
- example for uniform p

$$E_p[\cos(x)] = \int_0^\pi \cos(x) \frac{1}{\pi} dx = \frac{1}{\pi}(-\sin \pi + \sin 0) = 0$$

Variance

- motivation: quantifies the error of a Monte Carlo algorithm
- variance V of a function is the expected deviation of the function from its expected value $V[f(x)] = E[(f(x) - E[f(x)])^2]$
- properties
 - $V[af(x)] = a^2V[f(x)]$
 - $V[f(x)] = E[(f(x))^2] - E[f(x)]^2$
 - $\sum_i V[f(X_i)] = V[\sum_i f(X_i)]$ for independent random variables X_i

Monte Carlo Estimator

Uniform Random Variables

- motivation: approximation of the integral in the rendering equation
- goal: computation of $\int_a^b f(x)dx$
- uniformly distributed random variables $X_i \in [a, b]$
- probability density function $p(x) = \frac{1}{b-a}$ constant and integration to one
- Monte Carlo estimator $F_N = \frac{b-a}{N} \sum_{i=1}^N f(X_i)$
- expected value of F_N is equal to the integral $\int_a^b f(x)dx$
 - $E[F_N] = \int_a^b f(x)dx$
- variance $V = \frac{1}{N-1} \sum_{i=1}^N [f(X_i) - E[F_N]]^2$
 - convergence rate of $O(\sqrt{N})$
 - independent from the dimensionality
⇒ appropriate for high-dimensional integrals

Monte Carlo Estimator

Uniform Random Variables

- $$\begin{aligned} E[F_N] &= E \left[\frac{b-a}{N} \sum_{i=1}^N f(X_i) \right] \\ &= \frac{b-a}{N} \sum_{i=1}^N E[f(X_i)] \\ &= \frac{b-a}{N} \sum_{i=1}^N \int_a^b f(x) p(x) dx \\ &= \frac{b-a}{N} \sum_{i=1}^N \int_a^b f(x) \frac{1}{b-a} dx \\ &= \frac{1}{N} \sum_{i=1}^N \int_a^b f(x) dx \\ &= \int_a^b f(x) dx \end{aligned}$$

Examples - Uniform Random Variables

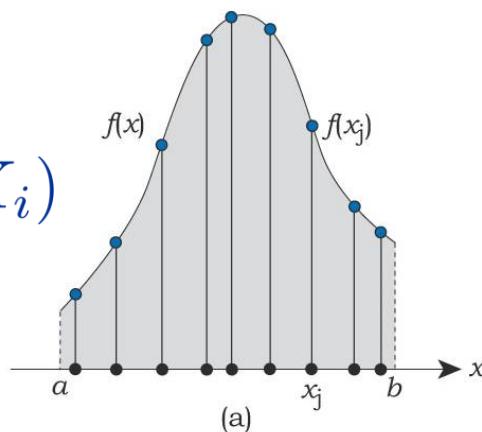
- integral $\int_0^1 5x^4 dx = 1$
- estimator $F_N = \frac{1}{N} \sum_{i=1}^N 5X_i^4$
- for an increasing number of uniformly distributed random variables X_i , the estimator converges to one

$$F_N = \frac{b-a}{N} \sum_{i=1}^N f(X_i)$$

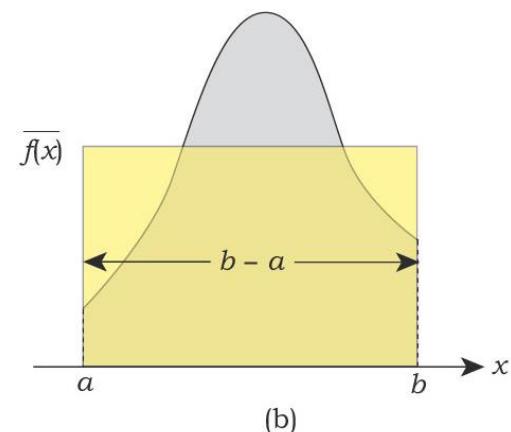
$$F_N = (b-a) \frac{1}{N} \sum_{i=1}^N f(X_i)$$

$$F_N = (b-a) \overline{f(x)}$$

$$E[F_N] = \int_a^b f(x) dx$$



uniformly distributed
random samples



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Monte Carlo Estimator

Non-uniform Random Variables

- Monte Carlo estimator $F_N = \frac{1}{N} \sum_{i=1}^N \frac{f(X_i)}{p(X_i)}$ $p(X_i) \neq 0$
- $$\begin{aligned} E[F_N] &= E \left[\frac{1}{N} \sum_{i=1}^N \frac{f(X_i)}{p(X_i)} \right] \\ &= \frac{1}{N} \sum_{i=1}^N \int_a^b \frac{f(x)}{p(x)} p(x) dx \\ &= \frac{1}{N} \sum_{i=1}^N \int_a^b f(x) dx \\ &= \int_a^b f(x) dx \end{aligned}$$

Monte Carlo Estimator

Multiple Dimensions

- samples X_i are multidimensional
- e.g. $\int_{x_0}^{x_1} \int_{y_0}^{y_1} \int_{z_0}^{z_1} f(x, y, z) dx dy dz$
- uniformly distributed random samples
 $(x_0, y_0, z_0) \leq X_i = (x_i, y_i, z_i) \leq (x_1, y_1, z_1)$
- probability density function $p(X_i) = \frac{1}{x_1 - x_0} \frac{1}{y_1 - y_0} \frac{1}{z_1 - z_0}$
- Monte Carlo estimator
$$F_N = \frac{(x_1 - x_0)(y_1 - y_0)(z_1 - z_0)}{N} \sum_{i=1}^N f(X_i)$$
- N can be arbitrary, N is independent from the dimensionality

Monte Carlo Estimator

Integration over a Hemisphere

- approximate computation of the irradiance at a point

$$\begin{aligned} E_i(p) &= \int_{2\pi+} L_i(p, \omega) \cos \theta d\omega \\ &= \int_0^{2\pi} \int_0^{\frac{\pi}{2}} L_i(p, \theta, \phi) \cos \theta \sin \theta d\theta d\phi \end{aligned}$$

- estimator

$$F_N = \frac{1}{N} \sum_{i=1}^N \frac{f(X_i)}{p(X_i)} = \frac{1}{N} \sum_{i=1}^N \frac{L_i(p, \theta, \phi) \cos \theta \sin \theta}{p(\theta, \phi)}$$

- probability distribution

- should be similar to the shape of the integrand
- as incident radiance is weighted with $\cos \theta$,
it is appropriate to generate more samples
close to the top of the hemisphere
- $p(\omega) \propto \cos \theta$

Monte Carlo Estimator

Integration over a Hemisphere

- probability distribution (cont.)

$$\int_{2\pi^+} c p(\omega) d\omega = 1$$

$$\int_0^{2\pi} \int_0^{\frac{\pi}{2}} c \cos \theta \sin \theta d\theta d\phi = 1$$

$$c \frac{2\pi}{1+1} = 1$$

$$c = \frac{1}{\pi}$$

$$p(\theta, \phi) = \frac{\cos \theta \sin \theta}{\pi}$$

- estimator

$$F_N = \frac{1}{N} \sum_{i=1}^N \frac{L_i(p, \theta, \phi) \cos \theta \sin \theta}{p(\theta, \phi)}$$

$$= \frac{\pi}{N} \sum_{i=1}^N L_i(p, \theta, \phi)$$

Monte Carlo Integration

Steps

- choose an appropriate probability density function
- generate random samples according to the PDF
- evaluate the function for all samples
- average the weighted function values

Monte Carlo Estimator

Error

- variance

$$V = \frac{1}{N} \int_a^b \left(\frac{f(x)}{p(x)} - F_N \right)^2 p(x) dx$$

- estimator

$$V_N = \frac{1}{N-1} \sum_{i=1}^N [f(X_i) - F_N]^2$$

- for increasing N

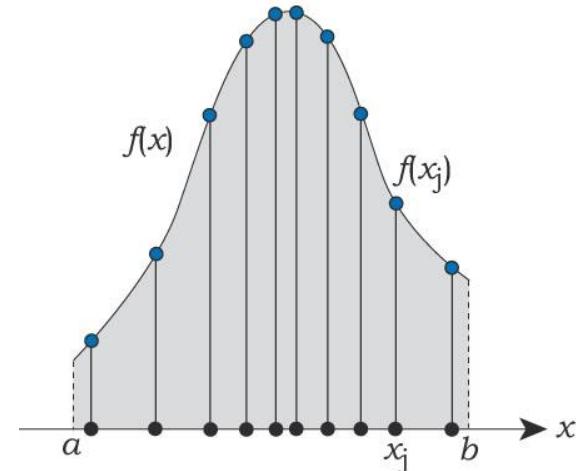
- the variance decreases with $O(N)$
 - the standard deviation decreases with $O(N^{1/2})$

- variance is perceived as noise

Monte Carlo Estimator

Variance Reduction / Error Reduction

- importance sampling
 - motivation: contributions of larger function values are more important
 - PDF should be similar to the shape of the function
 - optimal PDF $p(x) = \frac{f(x)}{\int f(x)dx}$
 - e.g., if incident radiance is weighted with $\cos \theta$, the PDF should choose more samples close to the normal direction
- stratified sampling
 - domain subdivision into strata does not increase the variance
- multi-jittered sampling
 - alternative to random samples for, e.g., uniform sampling of area lights

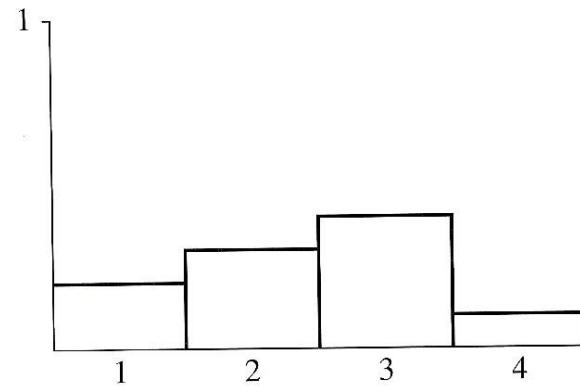


Outline

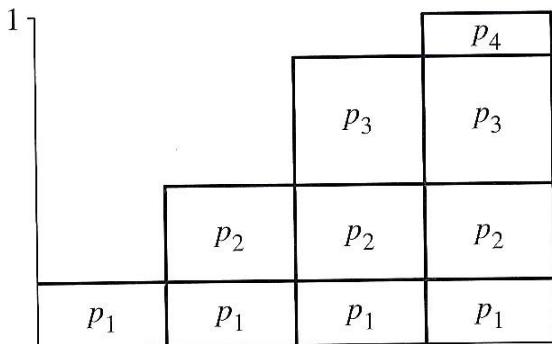
- rendering equation
- Monte Carlo integration
- sampling of random variables
 - inversion method
 - rejection method
 - transforming between distributions
 - 2D sampling
 - examples

Inversion Method

- mapping of a uniform random variable to a goal distribution
- discrete example
 - four outcomes with probabilities p_1, p_2, p_3, p_4 and $\sum_i p_i = 1$

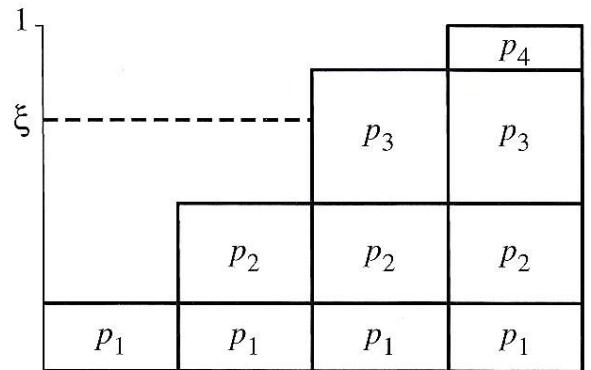


- computation of the cumulative distribution function $P(i) = \sum_{j=1}^i p_j$



Inversion Method

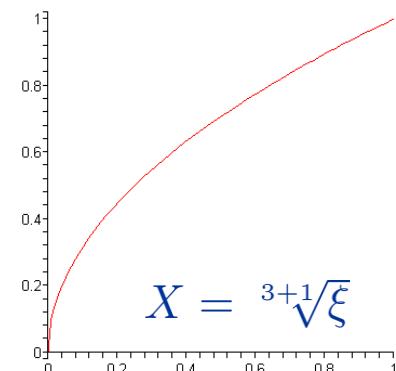
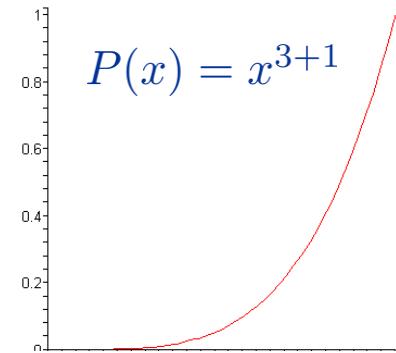
- discrete example cont.
 - take a uniform random variable ξ
 - $P^{-1}(\xi)$ has the desired distribution
- continuous case
 - P and P^{-1} are continuous functions
 - start with the desired PDF $p(x)$
 - compute $P(x) = \int_0^x p(x')dx'$
 - compute the inverse $P^{-1}(x)$
 - obtain a uniformly distributed variable
 - compute $X_i = P^{-1}(\xi)$ which adheres to $p(x)$



Inversion Method

Example 1

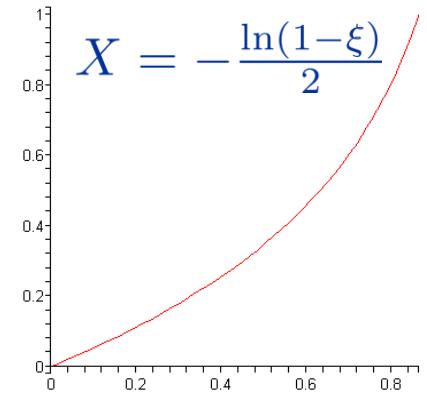
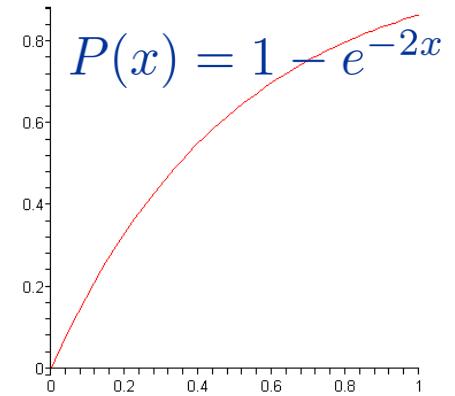
- power distribution $p(x) \propto x^n$
 - e.g., for sampling the Blinn microfacet model
- computation of the PDF
 - $\int_0^1 c x^n dx = 1 \Rightarrow c \frac{x^{n+1}}{n+1} \Big|_0^1 = 1 \Rightarrow c = n + 1$
- PDF $p(x) = (n + 1)x^n$
- CDF $P(x) = \int_0^x p(x')dx' = x^{n+1}$
- inverse of the CDF $P^{-1}(x) = \sqrt[n+1]{x}$
- sample generation
 - generate uniform random samples $0 \leq \xi \leq 1$
 - $X = \sqrt[n+1]{\xi}$ are samples from the power distribution $p(x) = (n + 1)x^n$



Inversion Method

Example 2

- exponential distribution $p(x) \propto e^{-ax}$
 - e.g., for considering participating media
- computation of the PDF
 - $\int_0^\infty c e^{-ax} dx = -\frac{c}{a} e^{-ax} \Big|_0^\infty = \frac{c}{a} = 1$
- PDF $p(x) = a e^{-ax}$
- CDF $P(x) = \int_0^x p(x') dx' = 1 - e^{-ax}$
- inverse of the CDF $P^{-1}(x) = -\frac{\ln(1-x)}{a}$
- sample generation
 - generate uniform random samples $0 \leq \xi \leq 1$
 - $X = -\frac{\ln(1-\xi)}{a}$ are samples from the power distribution $p(x) = a e^{-ax}$

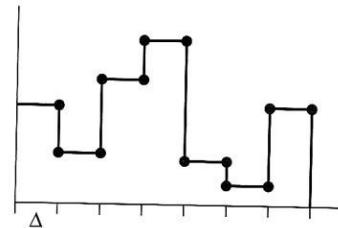


Inversion Method

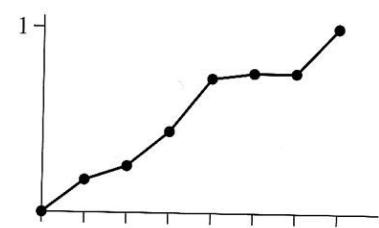
Example 3

- piecewise-constant distribution
 - e.g., for environment lighting

$$f(x) = \begin{cases} v_0 & x_0 \leq x < x_1 \\ v_1 & x_1 \leq x < x_2 \\ \vdots & \vdots \end{cases} \quad x_i = \Delta \cdot i \quad \Delta = \frac{1}{N}$$



$$p(x)$$



$$P(x)$$

- PDF $p(x) = \frac{1}{c} f(x)$

with $c = \int_0^1 f(x)dx = \sum_{i=0}^{N-1} \Delta \cdot v_i = \sum_{i=0}^{N-1} \frac{v_i}{N}$

Inversion Method

Example 3

- CDF

$$P(x_0) = 0$$

$$P(x_1) = \int_{x_0}^{x_1} p(x)dx = \Delta \cdot \frac{v_0}{c} = \frac{v_0}{Nc} = P(x_0) + \frac{v_0}{Nc}$$

$$P(x_2) = \int_{x_0}^{x_2} p(x)dx = \int_{x_0}^{x_1} p(x)dx + \int_{x_1}^{x_2} p(x)dx = P(x_1) + \frac{v_1}{Nc}$$

$$P(x_i) = P(x_{i-1}) + \frac{v_{i-1}}{Nc}$$

- CDF is linear between x_i and x_{i+1} with slope $\frac{v_i}{c}$

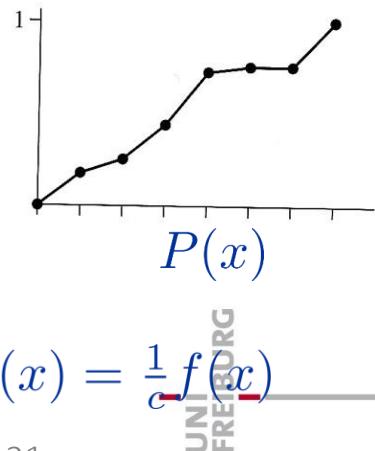
- sample generation

- generate uniform random samples $0 \leq \xi \leq 1$

- compute x_i with $P(x_i) \leq \xi$ and $\xi < P(x_{i+1})$

- compute d with $P(x_i) + d(P(x_{i+1}) - P(x_i)) = \xi$

- $X = x_i + d(x_{i+1} - x_i) = x_i + \frac{d}{N}$ are samples from $p(x) = \frac{1}{c}f(x)$

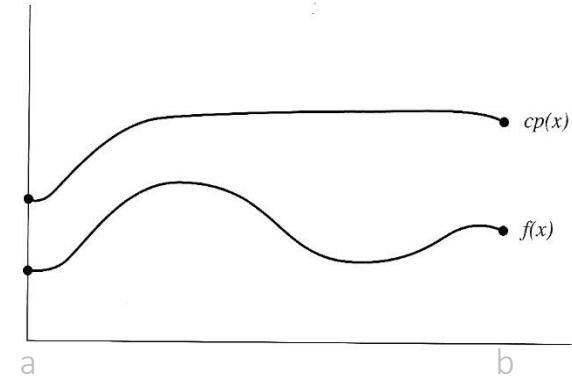


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 - rejection method
 - transforming between distributions
 - 2D sampling
 - examples

Rejection Method

- draws samples according to a function $f(x)$
 - dart-throwing approach
 - works with a PDF $p(x)$ and a scalar c with $f(x) < c \cdot p(x)$
- properties
 - $f(x)$ is not necessarily a PDF
 - PDF, CDF and inverse CDF do not have to be computed
 - simple to implement
 - useful for debugging purposes
- sample generation
 - generate a uniform random sample $0 \leq \xi < 1$
 - generate a sample X according to $p(x)$
 - accept X if $\xi \cdot c \cdot p(X) \leq f(X)$



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Transforming Between Distributions

- computation of a resulting PDF, when a function is applied to samples from an arbitrary distribution
 - random variables X_i are drawn from $p_x(x)$
 - bijective transformation (one-to-one mapping) $Y_i = y(X_i)$
 - How does the distribution $p_y(y)$ look like?
- $\Pr\{Y \leq y(x)\} = \Pr\{X \leq x\}$
 $P_y(y) = P_y(y(x)) = P_x(x)$
 $p_y(y) = \frac{p_x(x)}{|y'(x)|}$
- example $p_x(x) = 2x \quad 0 \leq x \leq 1$
 - $y(x) = \sin x \quad x(y) = \arcsin y$
 - $y'(x) = \cos x$
 - $p_y(y) = \frac{p_x(x)}{|\cos x|} = \frac{2 \arcsin y}{\sqrt{1-y^2}}$

Transforming Between Distributions

- multiple dimensions
 - X_i is an n-dimensional random variable
 - $Y_i = T(X_i)$ is a bijective transformation
- transformation of the PDF

- $p_y(y) = \frac{p_x(x)}{|J_T(x)|} \quad J_T(x) = \begin{pmatrix} \frac{\partial T_1}{\partial x_1} & \dots & \frac{\partial T_1}{\partial x_n} \\ \vdots & \ddots & \vdots \\ \frac{\partial T_n}{\partial x_1} & \dots & \frac{\partial T_n}{\partial x_n} \end{pmatrix}$

- example (polar coordinates)
 - samples (r, θ) with density $p(r, \theta)$
 - corresponding density $p(x, y)$ with $x = r \cos \theta$ and $y = r \sin \theta$
 - $J_T(x) = \begin{pmatrix} \cos \theta & -r \sin \theta \\ \sin \theta & r \cos \theta \end{pmatrix} \quad |J_T(x)| = r(\cos^2 \theta + \sin^2 \theta) = r$
 - $p(x, y) = \frac{1}{r} p(r, \theta) \quad p(r, \theta) = r \cdot p(x, y)$

Transforming Between Distributions

- example (spherical coordinates)

- $x = r \sin \theta \cos \phi$
- $y = r \sin \theta \sin \phi$
- $z = r \cos \theta$
- $p(r, \theta, \phi) = r^2 \sin \theta \cdot p(x, y, z)$

- example (solid angle)

- $\Pr\{\omega \in \Omega\} = \int_{\Omega} p(\omega) d\omega$
- $d\omega = \sin \theta \, d\theta \, d\phi$
- $p(\theta, \phi) = \sin \theta \cdot p(\omega)$

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Concept

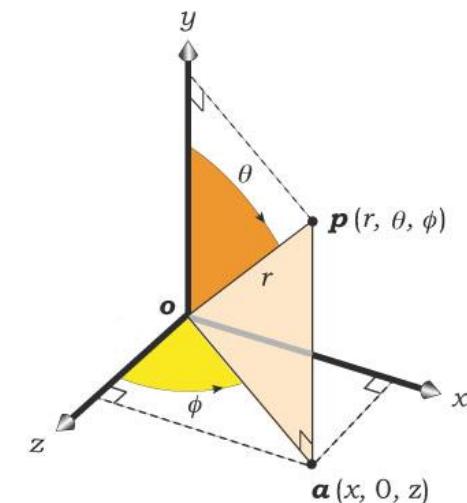
- generation of samples from a 2D joint density function $p(x, y)$
- general case
 - compute the marginal density function $p_x(x) = \int p(x, y) dy$
 - compute the conditional density function $p_y(y|x) = \frac{p(x,y)}{p_x(x)}$
 - generate a sample X according to $p_x(x)$
 - generate a sample Y according to $p_y(y|X) = \frac{p(x,y)}{p_x(X)}$
- marginal density function
 - integral of $p(x, y)$ for a particular x over all y -values
- conditional density function
 - density function for y given a particular x

Outline

- rendering equation
- Monte Carlo integration
- sampling of random variables
 - inversion method
 - rejection method
 - transforming between distributions
 - 2D sampling
 - examples

Uniform Sampling of a Hemisphere

- PDF is constant with respect to a solid angle $p(\omega) = c$
- $\int_{2\pi+} p(\omega)d\omega = 1 \Rightarrow c \int_{2\pi+} d\omega = 1 \Rightarrow c = \frac{1}{2\pi}$
- $p(\omega) = \frac{1}{2\pi} \Rightarrow p(\theta, \phi) = \frac{\sin \theta}{2\pi}$
- marginal density function
 - $p_\theta(\theta) = \int_0^{2\pi} p(\theta, \phi)d\phi = \int_0^{2\pi} \frac{\sin \theta}{2\pi} d\phi = \sin \theta$
- conditional density for ϕ
 - $p_\phi(\phi|\theta) = \frac{p(\theta, \phi)}{p_\theta(\theta)} = \frac{1}{2\pi}$
- inversion method
 - $P_\theta(\theta) = \int_0^\theta \sin \theta' d\theta' = -\cos \theta + 1$
 - $P_\phi(\phi|\theta) = \int_0^\phi \frac{1}{2\pi} d\phi' = \frac{\phi}{2\pi}$

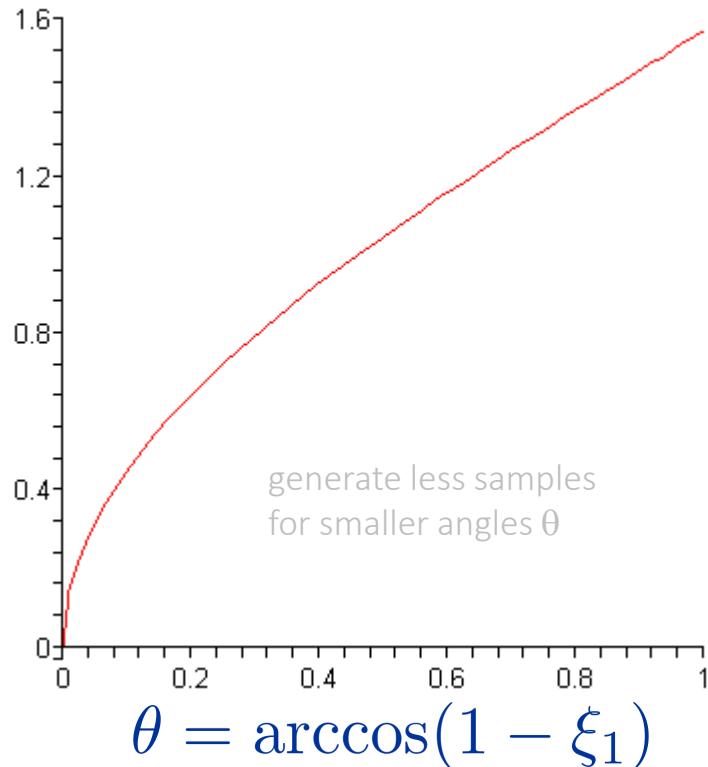


Uniform Sampling of a Hemisphere

- inversion method cont.
 - inverse functions of the cumulative distribution functions
 - $\theta = \arccos(1 - \xi_1)$
 - $\phi = 2\pi\xi_2$
 - generating uniformly sampled random values ξ_1 and ξ_2
 - applying the inverse CDFs to obtain θ and ϕ
- conversion to Cartesian space
 - $x = \sin \theta \cos \phi = \cos(2\pi\xi_2) \sqrt{1 - (1 - \xi_1)^2}$
 - $y = \sin \theta \sin \phi = \sin(2\pi\xi_2) \sqrt{1 - (1 - \xi_1)^2}$
 - $z = \cos \theta = 1 - \xi_1$
- $(x, y, z)^T$ is a normalized direction

Uniform Sampling of a Hemisphere

- illustration for θ

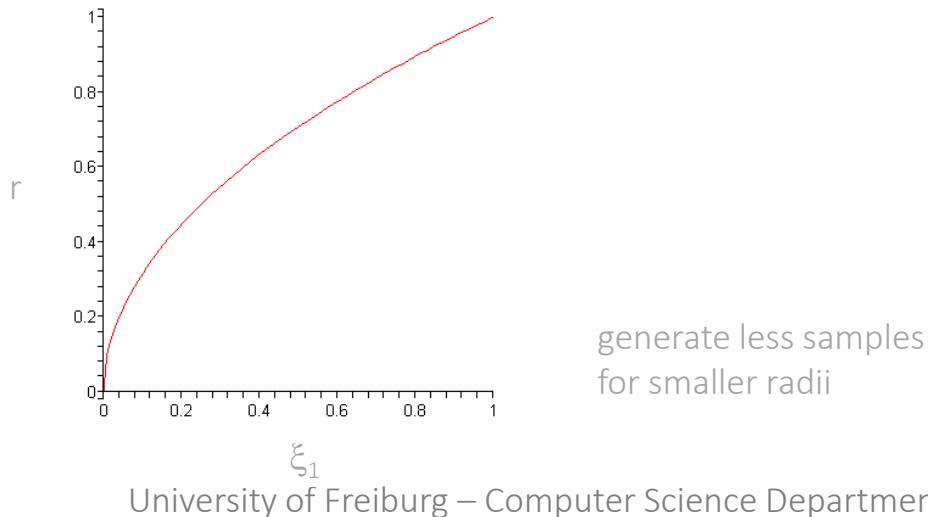


Uniform Sampling of a Unit Disk

- PDF is constant with respect to area $p(x, y) = \frac{1}{\pi}$
- $p(r, \theta) = r \cdot p(x, y) \Rightarrow \frac{r}{\pi}$
- marginal density function
 - $p_r(r) = \int_0^{2\pi} p(r, \theta) d\theta = 2r$
- conditional density
 - $p_\theta(\theta|r) = \frac{p(r, \theta)}{p_r(r)} = \frac{1}{2\pi}$
- inversion method
 - $P_r(r) = \int_0^r 2r' dr' = r^2$
 - $P_\theta(\theta|r) = \int_0^{2\pi} \frac{1}{2\pi} d\theta' = \frac{\theta}{2\pi}$

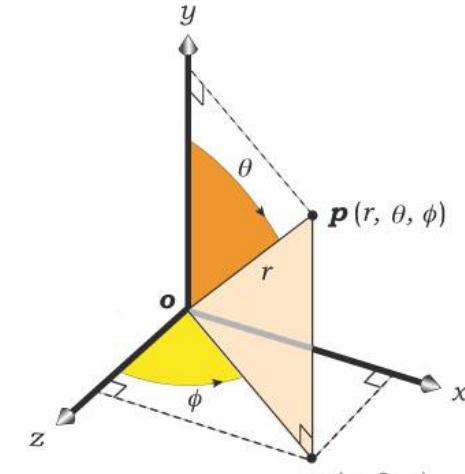
Uniform Sampling of a Unit Disk

- inversion method cont.
 - inverse functions of the cumulative distribution functions
 - $r = \sqrt{\xi_1}$
 - $\theta = 2\pi\xi_2$
 - generating uniformly sampled random values ξ_1 and ξ_2
 - applying the inverse CDFs to obtain r and θ



Uniform Sampling of a Cosine-Weighted Hemisphere

- PDF is proportional to $\cos \theta \quad p(\omega) \propto \cos \theta$
 - $\int_{2\pi} c p(\omega) d\omega = 1 = \int_0^{2\pi} \int_0^{\frac{\pi}{2}} c \cos \theta \sin \theta d\theta d\phi = c 2\pi \int_0^{\frac{\pi}{2}} \cos \theta \sin \theta d\theta = c 2\pi \frac{1}{2} = 1$
- marginal density function
 - $p_\theta(\theta) = \int_0^{2\pi} p(\theta, \phi) d\phi = \int_0^{2\pi} \frac{1}{\pi} \cos \theta \sin \theta d\phi = 2 \cos \theta \sin \theta$
- conditional density for ϕ
 - $p_\phi(\phi|\theta) = \frac{p(\theta, \phi)}{p_\theta(\theta)} = \frac{1}{2\pi}$
- inversion method
 - $P_\theta(\theta) = \int_0^\theta 2 \cos \theta' \sin \theta' d\theta' = 2 \left[-\frac{\cos^2 \theta'}{2} \right]_0^\theta = 2 \left(-\frac{\cos^2 \theta}{2} + \frac{1}{2} \right) = \sin^2 \theta$
 - $P_\phi(\phi|\theta) = \int_0^\phi \frac{1}{2\pi} d\phi' = \frac{\phi}{2\pi}$



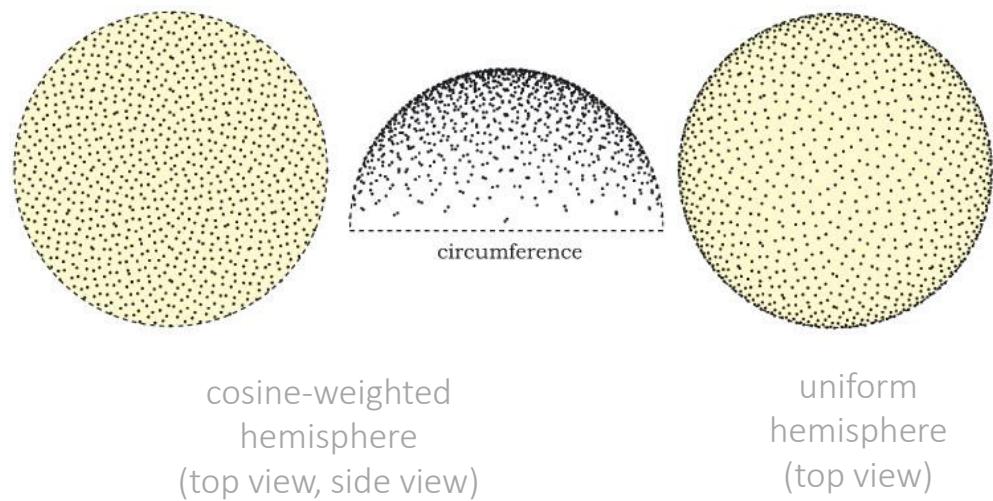
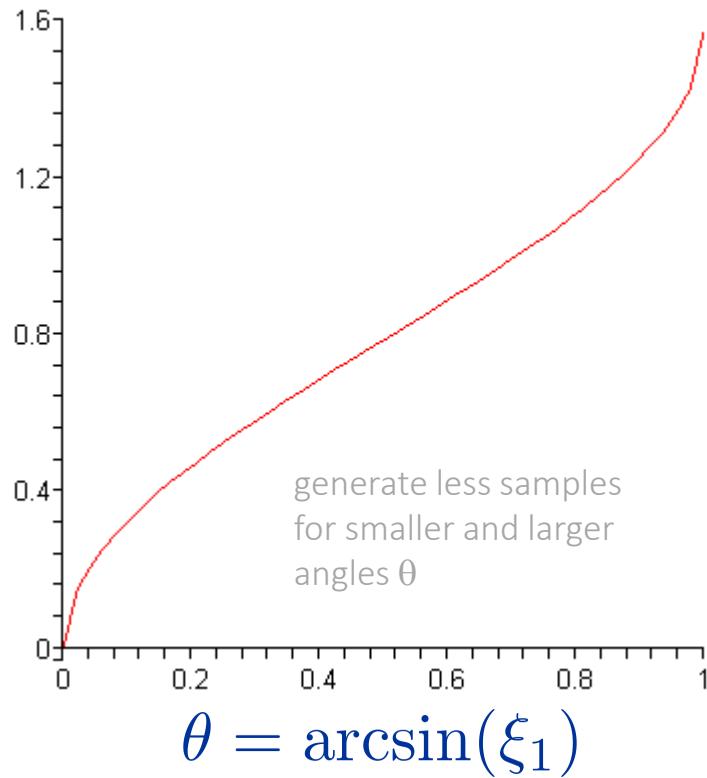
Uniform Sampling of a Cosine-Weighted Hemisphere

- inversion method cont.
 - inverse functions of the cumulative distribution functions
 - $\theta = \arcsin(\sqrt{\xi_1})$
 - $\phi = 2\pi\xi_2$
 - generating uniformly sampled random values ξ_1 and ξ_2
 - applying the inverse CDFs to obtain θ and ϕ
- conversion to Cartesian space
 - $x = \sin \theta \cos \phi = \cos(2\pi\xi_2)\sqrt{\xi_1}$
 - $y = \sin \theta \sin \phi = \sin(2\pi\xi_2)\sqrt{\xi_1}$
 - $z = \cos \theta = \sqrt{1 - \xi_1}$
- $(x, y, z)^T$ is a normalized direction

x- y- values uniformly sample a unit disk, i. e., cosine-weighted samples of the hemisphere can also be obtained by uniformly sampling a unit sphere and projecting the samples onto the hemisphere

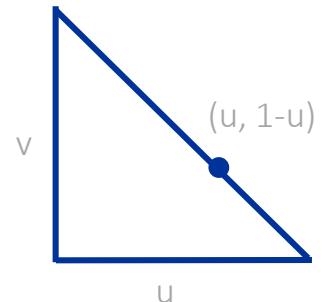
Uniform Sampling of a Cosine-Weighted Hemisphere

- illustration for θ



Uniform Sampling of a Triangle

- sampling an isosceles right triangle of area 0.5
 - u, v can be interpreted as Barycentric coordinates
 - can be used to generate samples for arbitrary triangles
- $p(u, v) = 2$
- marginal density function
 - $p_u(u) = \int_0^{1-u} p(u, v) dv = 2 \int_0^{1-u} dv = 2(1 - u)$
- conditional density
 - $p_v(v|u) = \frac{p(u,v)}{p_u(u)} = \frac{1}{1-u}$
- inversion method
 - $P_u(u) = \int_0^u 2 - 2u' du' = 2u - u^2$
 - $P_v(v|u) = \int_0^v \frac{1}{1-u} dv' = \frac{v}{1-u}$



Uniform Sampling of a Triangle

- inversion method cont.
 - inverse functions of the cumulative distribution functions
 - $u = 1 - \sqrt{\xi_1}$
 - $v = \xi_2 \sqrt{\xi_1}$
 - generating uniformly sampled random values ξ_1 and ξ_2
 - applying the inverse CDFs to obtain u and v

Piecewise-Constant 2D Distribution

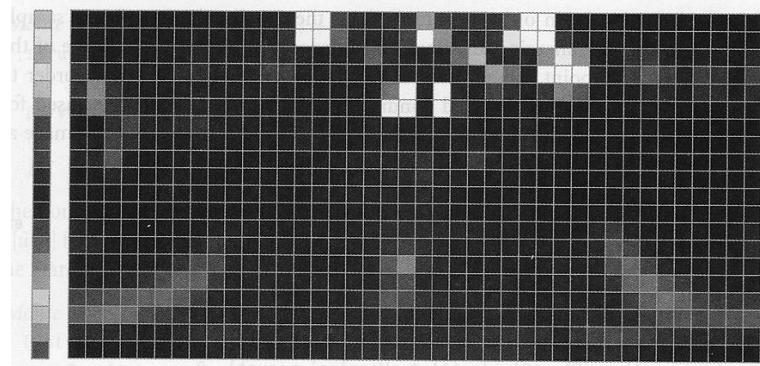
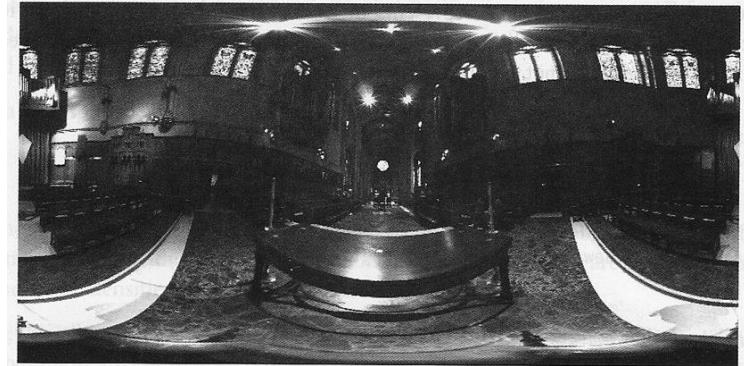
- $n_u \times n_v$ samples defined over $(u, v) \in [0, 1]^2$
 - e.g., an environment map
- $f(u, v)$ is defined by a set of $n_u \times n_v$ values $f[u_i, v_i]$
 - $u_i \in [0, \dots, n_u - 1]$ $v_i \in [0, \dots, n_v - 1]$
 - $f[u_i, v_i]$ is the value of $f(u, v)$ in the range $\left[\frac{i}{n_u}, \frac{i+1}{n_u} \right) \times \left[\frac{j}{n_v}, \frac{j+1}{n_v} \right)$
 - $f(u, v) = f[u_i, v_i]$ with $\tilde{u} = \lfloor n_u u \rfloor$ and $\tilde{v} = \lfloor n_v v \rfloor$
- integral over the domain
 - $I_f = \int \int f(u, v) \, du \, dv = \frac{1}{n_u n_v} \sum_i \sum_j f[u_i, v_j]$
- PDF
 - $p(u, v) = \frac{1}{I_f} f(u, v) = \frac{1}{I_f} f[\tilde{u}, \tilde{v}]$

Piecewise-Constant 2D Distribution

- marginal density function
 - $p_v(v) = \int p(u, v) du = \frac{1}{I_f} \frac{1}{n_u} \sum_i f[u_i, \tilde{v}]$
 - piecewise-constant 1D function
 - defined by n_v values $p_v[\tilde{v}]$
- conditional density
 - $p_u(u|v) = \frac{p(u,v)}{p_v(v)} = \frac{1}{I_f} \frac{f[\tilde{u}, \tilde{v}]}{p[\tilde{v}]}$
 - piecewise-constant 1D function
- sample generation
 - see example 3 of the inversion method

Piecewise-Constant 2D Distribution

- environment map
- low-resolution of the marginal density function and the conditional distributions for rows
- first, a row is selected according to the marginal density function
- then, a column is selected from the row's 1D conditional distribution



Paul Debevec, Grace Cathedral