- Solar wind and geomagnetic activity influence on the
- <sup>2</sup> equatorial plasma mass density at geosynchronous orbit

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5 Abstract

We consider two types of events, identified by decreases in  $D_{st}$  below a threshold value and increases in the the equatorial mass density at geosynchronous altitudes,  $\rho_{eq}$ , above a threshold value on 1-hour and 1-day time scales using the  $Takahashi\ et\ al.$  [2010] dataset. From the  $D_{st}$  events, we find that there is a weak and small amplitude (< 10%) difference between  $\rho_{eq}$  on the day of the event and the days before and after on both daily and hourly timescales. Although very large geomagnetic storms have been observed to correspond to large increases in  $\rho_{eq}$ , on average,  $\rho_{eq}$  following the onset of a geomagnetic storm does not depend on the storm intensity or north-south component of the interplanetary magnetic field,  $B_z$ . From the  $\rho_{eq}$  events, we find a weak dependence on  $B_z$  after the onset of an event, with higher average  $B_z$  six hours after the event onset corresponding to larger  $\rho_{eq}$  events. In contrast, the average of  $B_z$  six hours prior to  $\rho_{eq}$  events does not have a statistically significant impact on  $\rho_{eq}$  after onset.

#### 1 Introduction

The dependence of the plasmapause location on geomagnetic activity was observed as early as Carpenter [1966] and there are well-validated models of the plasmapause boundary, as a function of L, MLT, and geomagnetic activity (Lemaire [1998], Moldwin et al. [2002], O'Brien and Moldwin [2003]). Models also exist for the plasmasphere density (Gallagher et al. [1988]; Lemaire [1998]) with similar parameterization.

The dependence of the plasma density outside of the nominal plasmapause location on geomagnetic and solar wind conditions has not been extensively studied until the past decade (Takahashi et al. [2006]; Takahashi et al. [2010]; Denton et al. [2016]). Earlier works that considered profiles of hydrogen ion density found a significant dependence of the plasmasphere and plasmapause on geomagnetic activity (Chappell et al. [1970]; Maynard and Grebowsky [1977]; Carpenter and Anderson [1992]), but the relationship between geomagnetic activity

and density outside of the plasmapause was not explicitly considered.

Takahashi et al. [2006] estimated magnetospheric mass density using measurements from the CRRES satellite during a 73-day period in 1991 when CRRES was in an elliptical, near equatorial orbit and primarily beyond the plasmapause. They found that the local average ion mass, M, had some correlation with geomagnetic activity, with more negative hourly  $D_{st}$  values corresponding to higher M and 1.5- and 3-day averages of  $K_p$  corresponding to higher M (3-hour averages of  $K_p$  had little visual correlation with M). The mass density estimates were obtained from CRRES observations in the MLT range of [12:00, 18:00] with most observations at L between 5 and 7  $R_E$ .

Denton et al. [2006] found a weak relationship between  $D_{st}$  and and  $K_p$  and the field line distribution of mass density,  $\rho$ , using observations of the first three toroidal Alfvén harmonic frequencies for L in the range of 6-8  $R_E$ . A more pronounced peak in the distribution along a filed line near the equator was found for lower  $D_{st}$  and larger  $K_p$ . The equatorial value of  $\rho$  appeared identical for  $D_{st}$  bins of [-142, -31] and [-31, 37] nT and  $K_p$  bins of [1.5,3.4] and [3.4, 5.9].

Takahashi et al. [2010] developed a mass density dataset using measurements from the Space Environment Monitor instruments on the Geostationary Operational Environmental Satellites (GOES) satellites from 1980 through 1992, with most measurements in the range of  $L = 6.8 \pm 0.2 R_E$ . The mass density was estimated using the Alfvén wave velocity relationship,  $V_A = B/\sqrt{\mu_0 \rho}$ , a magnetic field model, and a numerical solution to a wave equation with an ionospheric boundary and the assumption of a zero resistance ionosphere. The equatorial mass density,  $\rho_{eq}$ , was derived from the estimated mass density using a power law dependence on the geocentric distance to the field line of the observation, R,  $\rho = \rho_{eq}(LR_E/R)^{1/2}$ . They found a high correlation ( $\sim 0.94$ ) between 27-day averages of F10.7 and 27-day medians of  $\rho_{eq}$ .

Takahashi et al. [2010] noted that downward spikes in the  $D_{st}$  index coincided with with significant changes in  $\rho_{eq}$  for L near 6.8  $R_E$ . For five storms, two had  $\rho_{eq}$  spikes after the  $D_{st}$  drop, two had  $\rho_{eq}$  spikes before the drop, and one showed little change in  $\rho_{eq}$ . A key result was that when daily-averaged measurements were considered, the epoch average of  $D_s t$  exhibited for storms with a minimum  $D_{st} < -50$  nT, an enhancement in  $\rho_{eq}$  the same day as minimum  $D_{st}$  epoch averages.

Yao et al. [2008] studied the relationship between  $D_{st}$  and the number density of lowenergy  $O^+$  in different regions (ring current and plasma sheet) using the TEAMS and ESA instruments on the low-altitude and high-inclination orbiting FAST satellite and found that the average  $N_{O^+}$  across the sampled L-shells (2-14) had a strong correlation (0.88) with the minimum  $D_{st}$  of the storm. Although the correlation at geosynchronous distances was not calculated, the data presented for four events show that the variations in  $N_{O^+}$  tended to appear near  $D_{st}$  minimum ( $\pm$  2 hours).

The above results in the papers discussed indicate that (1) the mass density in the at geosynchronous distances is best correlated with F10.7 and (2) there exists a much weaker statistical relationship between mass density and the geomagnetic activity indices  $K_p$  and  $D_{st}$ . In this work we consider the dependence of mass density estimates in the Denton

et al. [2006] data set and attempt to identify and statistically characterize the relationship between geomagnetic activity with mass density at geosynchronous distances. In addition, we attempt to identify solar wind processes that may drive changes in mass density.

## <sub>6</sub> 2 Data Preparation and Overview

The parameters  $\rho_{eq}$  and  $F_{10.7}$  are from the dataset of Takahashi et al. [2010]; data are 77 available from 1980 through 1991 from GOES 2, 3, 5, 6 and 7. All other parameters used are from Kondrashov et al. [2014] which has coverage from 1972 through 2013, which are on a 1-hour time grid.  $\rho_{eq}$  is the inferred equatorial mass density based on the 3rd harmonic torodial frequency of magnetic field measurements as described in Takahashi et al. [2010]. 81 The smallest cadence for  $\rho_{eq}$  values is 10 minutes. To compute an hourly median over the same interval for which the solar wind parameters were averaged, the median of all  $\rho_{eq}$  values in a given hour window was used. GOES 6 had the longest span of available data and the primary results presented are for GOES 6 while measurements from other satellites were used for consistency checks. Figure 2 shows the long-term trends of  $\log_{10}(\rho_{eq})$  from GOES 6 and  $F_{10.7}$  computed using 87 the median values in 27-day non-overlapping windows. A scatter plot of these two lines is shown in the lower panel of Figure 2 along with that for 1-day and 1-hour medians. The linear correlation for GOES 6 using all measurements is found to be 0.94, which is the same value documented by Takahashi et al. [2010] who used measurements from all satellites in the 91 time interval of 1980 through 1991 using the constraint  $0600 \le MLT \le 1200, K_{p3d} \ge 1.0$  and

 $D_{st} \geq -50$  nT. Our 27-day correlations for GOES 2, 5, and 7 are 0.81, 0.78, and 0.87, and

the time range of coverage of measurements from these satellites is 1980-01-01/1983-05-16,

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1983-01-01/1987-02-28, and 1983-05-27/1991-08-29 respectively.

## <sup>96</sup> 3 Epoch Average Results

Two types of events are considered. The first is a drop in the  $D_{st}$  index below the threshold of -50 nT. These types of events are considered in order to determine if there is a well-defined  $\rho_{eq}$  dependence on geomagnetic activity/storms as indicated by the  $D_{st}$  index. It is expected that the plasmasphere will exhibit two responses near these threshold crossings. First is the movement of the plasmapause earthward (*Lemaire* [1998]) and the second is a possible change in density due to magnetospheric processes.

The second type of event is an increase in  $\rho_{eq}$  above a threshold of 20 amu/cm<sup>3</sup>. It is expected that these increases will typically occur after period of quiet geomagnetic activity but may also be the result of increases in geomagnetic activity. In this case, the plasmapause boundary may move outward past geosynchronous orbit and the increase will be due to measurements being made inside the plasmapause. A second possible reason for the increase could be due to magnetospheric processes.

#### $D_{st}$ Events

Our first analysis uses data only from the time interval 1989-1991, which corresponds with 110 the interval used in Takahashi et al. [2010]. In this interval, there were 329  $D_{st}$  events. For 111 both types of events, the hour of the threshold crossing defines the zero epoch time, and for 112 each event,  $4 \cdot 24$  hours were considered before and  $8 \cdot 24$  hours after each event. To compute 113 the epoch averages on a daily time scale, the median value of all available measurements for 114 all events centered on a window of  $\pm 12$  hours of the epoch zero hour was computed, and 115 these averaging windows were shifted in increments of 24 hours. Similar results are obtained 116 if we first reduce each event time series to have a 1-day cadence by computing medians for 117 each event in 1-day bins and then compute the medians across events. 118

The stack plot in the upper panel of Figure 3 shows the epoch averages for the 329 events from 1989-1991. The minimum  $D_{st}$  median is -48 nT. Consistent with Takahashi

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et al. [2010],  $D_{st}$  events correspond to elevated  $\rho_{eq}$  before and after the event, although the 121 magnitude of increase observed here is 1.5 amu/cm<sup>3</sup> (from 18 to 19.5 amu/cm<sup>3</sup>) instead of 122 the increase of  $\sim 10~\mathrm{amu/cm^3}$  found in Figure 11 of Takahashi et al. [2010]. The vertical 123 green bars in the  $\rho_{eq}$  plot show the number of  $\rho_{eq}$  values that were used to compute the 124 medians and red error lines. Error lines for each parameter are the standard deviation of the 125 values used in computing the median divided by the square root of the number of values. 126 For the non- $\rho_{eq}$  medians, the number of measurements used in computing the medians is 127 typically equal to the number of events, which is much larger than the number of  $\rho_{eq}$  values 128 as all available measurements were used instead of restricting to only values where a  $\rho_{eq}$ 129 value also existed. 130

In the lower panel of Figure 3 it is shown that the trends seen in the analysis of 1989-1991 also hold for the entire span of measurements from GOES 6 - a small elevation on the day of the event relative to the day before and day after, with median averages after the day of the event being slightly lower than that before.

To determine if the magnitude of the density medians in the top panel of Figure 3 on epoch days -1, 0, and 1 are statistically different from each other, we used a two-population bootstrap test. A bootstrap sample of the medians for each epoch day was created by sampling the values used to compute the median on each epoch day with replacement. The observed difference between the medians is shown in Table 1 along with the fraction of observations in 1000 bootstrap differences with a different sign than that observed.

If a statistically significant difference in the medians between two days existed, the fraction of bootstrap samples with a different sign is expected to be much smaller (< 5%). In terms of a hypothesis test, we cannot reject the null hypothesis that the medians are the same with a confidence higher than 100 minus the % value shown in the table; the highest confidence level of 87.6% is in the difference between day 0 and day 1.

These tests indicate that there is not a statistically significant change in medians of  $\rho_{eq}$  on the day before and day after a threshold crossing in  $D_{st}$ . In addition, the observed

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Table 1: Results of test on of means of  $\rho_{eq}$  shown in the top panel of Figure 3 between days of threshold crossing near (day = 1 or -1) or on the day of a  $D_{st}$  event (day = 0).

differences are on the order of only 10% of the  $\rho_{eq}$ , which is small compared to the variance in the observations.

We have also considered hourly averages as shown in the top panel of Figure 4. A total of  $668 D_{st}$  events were found in the interval of May 1983 through August 1991 with an average duration of 9 hours and a median duration of 3 hours. In this plot, the overall trend is similar to that found in the bottom panel of Figure 4 -  $\rho_{eq}$  is slightly elevated at the time of the threshold crossing of  $D_{st}$  with respect to the values one day before and after, but the elevation is small relative to the error bars around the time of the event onset.

The bottom panel of Figure 4 contains the events shown in the top panel under the constraint that  $D_{st}$  remained below -50 nT for at least twelve hours after the event onset time. In the time period of elevated mass density within 12 hours of the start of the event, there are only  $\sim$ 5 samples per hourly interval. This result is consistent with previous observations of large mass density spikes in the plasma trough region during large geomagnetic storms (Yao et al. [2008]; Takahashi et al. [2010]).

To determine of this relationship can be also observed in the events of the top panel of Figure 4, we sorted the events by the average  $D_{st}$  values in the time interval of 0-12 hours, split the events in two and recomputed the epoch averages of  $\rho_{eq}$ . The result is that there is no statistically significant difference in the mass density epoch averages for large versus small events. This indicates that to confirm a statistical  $\rho_{eq}$  dependence on large geomagnetic storms would require a dataset with a much larger number of strong geomagnetic storms.

## $_{68}$ 3.2 $ho_{eq}$ Events

An alternative type of event can be defined that corresponds to large increases in  $\rho_{eq}$ . Figure 5 shows the epoch time series for events in which an increase of  $\rho_{eq}$  above 20 amu/cm<sup>2</sup> defines

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the start of the event. In finding  $\rho_{eq}$  events, missing values were replaced with linearly interpolated values. These events are associated with positive  $B_z$  and low values of  $D_{st}$ . Approximately 8 hours after onset,  $\rho_{eq}$  peaks and then slowly decays after 40 hours to preonset levels.

Figure 5 can be compared with Figure 9 of  $Denton\ et\ al.\ [2016]$ , which shows daily averages of  $\rho_{eq}$  from this dataset after the onset of quiet  $K_p$  intervals. The growth shown here is approximately linear and lasts  $\sim 12$ –18 hours whereas the 10 quiet  $K_p$  events considered by  $Denton\ et\ al.\ [2016]$  showed growth that lasted for at least 48 hours.

To determine if there is an IMF  $B_z$  dependence, the epoch curve for  $\rho_{eq}$  and  $B_z$  was 179 separated into two parts depending whether  $B_z$  was above or below the average four hours before and including the onset hour and four hours after and including the hour of onset; 181 the result is shown in Figure 6. Any statistically significant difference at a 95% confidence 182 level between the two parts is indicated by a green dot on the horizontal axis using the same bootstrap method described previously. (At this confidence level, false positives are expected 184 to randomly occur 5\% of the time or 3.6 times for 72 values.) The largest separation occurs 185 when the separation is performed by Figure 6 after onset, with more positive  $B_z$  leading 186 to larger peak  $\rho_{eq}$  than more negative  $B_z$ . The rate of change of  $\rho_{eq}$  also appears to be 187 insenstive to  $B_z$  for approximately 12 hours prior to onset. We note that the equivalent 188 epoch separation curves for  $P_{dyn}$  had a statistically significant differences similar to that for 189  $B_z$ , with higher dynamic pressure playing a role similar to positive  $B_z$ . 190

Figure 7 shows how the epoch average curve of  $\rho_{eq}$  in Figure 6 depends on F10.7. Statistically significant differences occur independent of if the separating factor the average of F10.7 before or after onset, as expected because the cadence of F10.7 is one day and its
time scale of variation is longer than that of  $\rho_{eq}$ . Events with high F10.7 are associated with
higher  $\rho_{eq}$  before and after event onset but similar peak values. As was the case for  $B_Z$ ,
there is an approximately 12-hour time span of linear growth of  $\rho_{eq}$  near the time of onset
but because the initial starting  $\rho_{eq}$  is higher for high F10.7, the associated growth rate is

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lower.

## 3.3 Conclusions

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# <sup>235</sup> 4 Figures

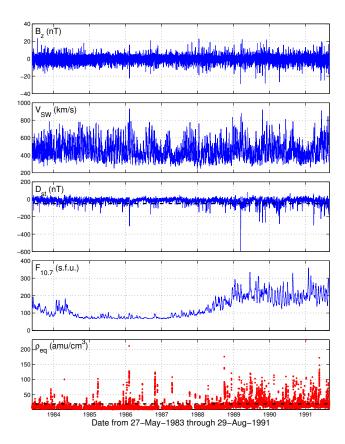


Figure 1: Overview of data used in this article. The top four panels show parameters from Kondrashov et al. [2014] and the bottom panel contains  $\rho_{eq}$  based on GOES 6 measurements from Denton [2007] after interpolation and averaging described in the text. Dashed horizontal lines in the  $D_{st}$  and  $\rho_{eq}$  panels indicate sample event cutoff thresholds of  $D_{st} = -50$  nT and  $\rho_{eq} = 20$  amu/cm<sup>3</sup> considered in Section 3

 $\label{eq:continuous} D\ R\ A\ F\ T \\ \hspace{2cm} D\ R\ A\ F\ T$ 

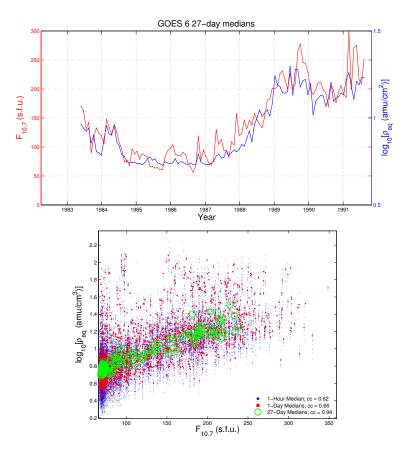


Figure 2: Top: 27-day non-overlapping medians of  $F_{10.7}$  and  $log(\rho_{eq})$  from GOES 6. Bottom: Correlation between  $log(\rho_{eq})$  and  $F_{10.7}$  using medians in non-overlapping hour, day, and 27-day windows.

 $\label{eq:definition} D\ R\ A\ F\ T \\$   $\label{eq:definition} D\ R\ A\ F\ T$ 

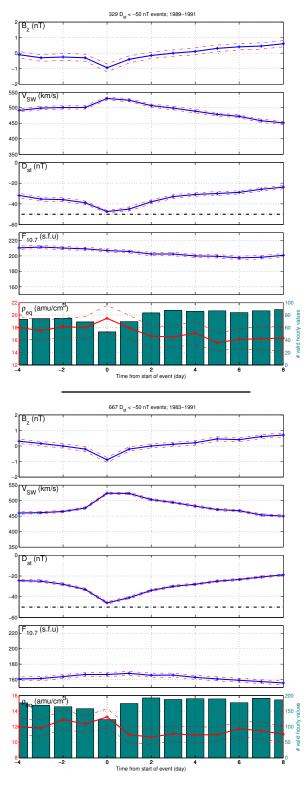


Figure 3:  $D_{st}$  events from GOES 6 using daily medians. Top: Events in the interval 1989-1991; compare to  $Takahashi\ et\ al.\ [2010]$  Figure 11. Bottom: Events in the interval 1983-1991.

 $\label{eq:control_def} D\ R\ A\ F\ T \\$   $\label{eq:control_def} D\ R\ A\ F\ T \\$ 

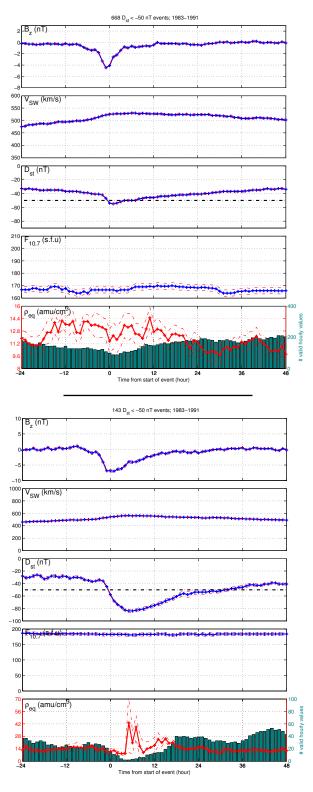


Figure 4:  $D_{st}$  events from GOES 6 using hourly medians. Top: Events in the interval 1983-1991; compare to bottom panel Figure 3. Bottom: Same as Top except for constraint that  $D_{st}$  stayed below -50 nT for at least 12 hours after crossing below -50 nT.

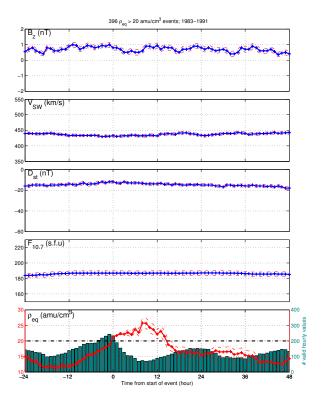


Figure 5:  $\rho_{eq} > 20$  amu/cm<sup>3</sup> events from GOES 6 using hourly medians. As described in the text, there is no statistically significant dependence of the  $\rho_{eq}$  plot on  $B_z$  in the time interval of pm2 hours around the event.

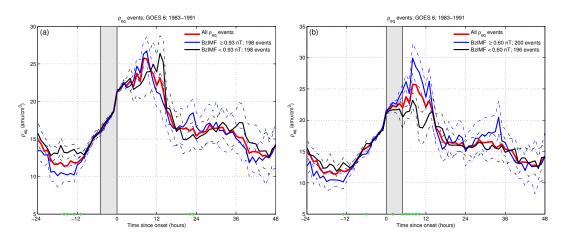


Figure 6:  $\rho_{eq}$  events of Figure 5 separated by the average  $B_z$  value (a) at onset and four hours before, and (b) at onset and four hours after.

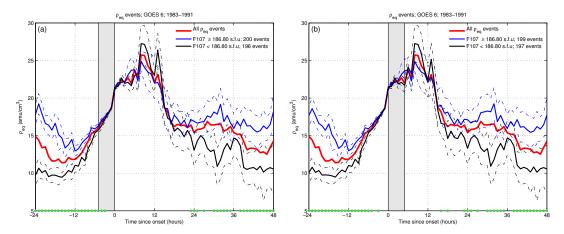


Figure 7:  $\rho_{eq}$  events of Figure 5 separated by the average F10.7 value (a) at onset and four hours before, and (b) at onset and four hours after.

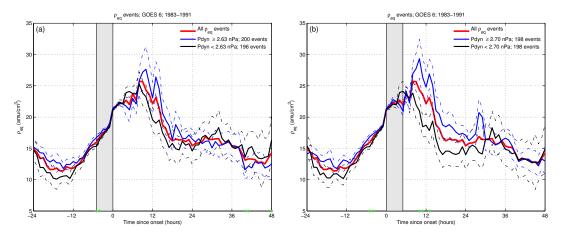


Figure 8: Binning  $\rho_{eq}$  events by  $B_z$  value (a) at onset and four hours before, and (b) at onset and four hours after.

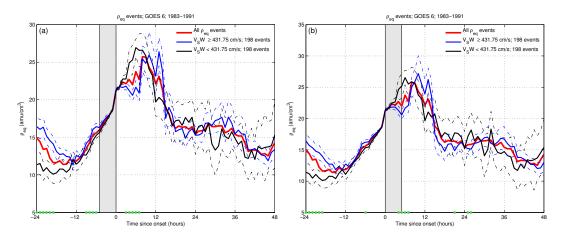


Figure 9: Binning  $\rho_{eq}$  events by  $B_z$  value (a) at onset and four hours before, and (b) at onset and four hours after.

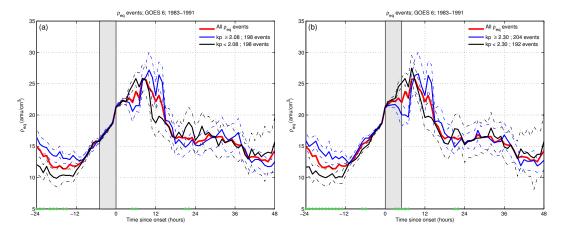


Figure 10: Binning  $\rho_{eq}$  events by  $B_z$  value (a) at onset and four hours before, and (b) at onset and four hours after.

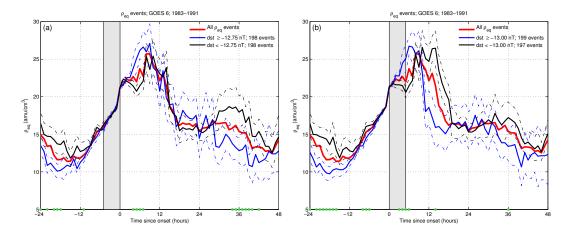


Figure 11: Binning  $\rho_{eq}$  events by  $B_z$  value (a) at onset and four hours before, and (b) at onset and four hours after.