$(x^{(t+1)} - x^*) = (x^{(t)} - x^*) - \eta (\nabla f(x^{(t)}) - \nabla f(x^*))$ $= \left(x^{(t)} - x^*\right) - \eta \nabla^2 f(\zeta^{(t)}) \left(x^{(t)} - x^*\right)$ = $\left(I - \eta \nabla^2 f(\zeta^{(t)})\right) \left(x^{(t)} - x^*\right)$.

orm inequality $||Az||_2 \le ||A|| ||z||_2$

 $\|x^{(t+1)} - x^*\|_2 \leq \|I - \eta \nabla^2 f(\zeta^{(t)}) \big\| \|x^{(t)} - x^*\|_2$ $\leq \max{(|1-\eta\mu|,|1-\eta L|)\,\|x^{(t)}-x^*\|_2}$

the definition of PSD matrices and the properties of the

$$\frac{2}{1+\mu}$$
, we have

• A differentiable function
$$f$$
 is strongly convex if there exists a constant $\mu>0$ such that:
$$f(y)\geq f(x)+\nabla f(x)^T(y-x)+\frac{\mu}{2}\|y-x\|^2,\quad \forall x,y\in\mathbb{R}^n.$$

Strongly convex functions are easier to optimize and allow for faster convergence guarantees.

$$\nabla^2 f(x) \succeq \mu I$$

where $\mu > 0$ is a constant. This means that the smallest eigenvalue of the Hessian is at least μ .

• f(x) is strongly convex, if there is a $\mu > 0$, such that the function $h(x) = f(x) - \frac{\mu}{2} ||x||^2$ is convex

$$\|x^{(t+1)} - x^*\|_2 \leq \left(\frac{L - \mu}{L + \mu}\right) \cdot \|x^{(t)} - x^*\|_2,$$

$$\|x^{(N)} - x^*\|_2 \leq \left(\frac{L - \mu}{L + \mu}\right)^N \cdot \|x^{(0)} - x^*\|_2, \\ = \mathbb{E} x^{(t)} - \alpha_t \frac{1}{N} \sum_{i=1}^N \nabla f(x^{(t)}; y_i).$$

Definition 2.1. A probability space is the triple $(\Omega, \mathcal{F}, \mathbb{P})$. Here,

- Ω is called the *sample space*, which is the set of all possible outcomes of a random experiment.
- \mathcal{F} is called the *event space*, which is a set whose elements $A \in \mathcal{F}$ are subsets of Ω , i.e. $\mathcal{F} \subseteq 2^{\Omega}$. Further,
 - 1. $\emptyset \in \mathcal{F}$, i.e., the empty set is an element of \mathcal{F} .

 - 3. If $A_1,A_2,...\in\mathcal{F}\Rightarrow \cup_i A_i\in\mathcal{F},$ i.e., \mathcal{F} is closed under countable unions
- $\mathbb{P}: \mathcal{F} \to \mathbb{R}$ is called the probability measure, which should satisfy the axioms of probability

Proposition 2 (Properties of CDFs)

- $0 \le F_X(x) \le 1$.
- $\lim_{x\to\infty} F_X(x) = 1$.
- lim_{x→-∞} F_X(x) = 0.
- $x \le y \Rightarrow F_X(x) \le F_X(y)$.
- $F_X(x)$ is right continuous

Definition 2.4 (Random variables). A random variable X is a measurable function from Ω , to (for example)

$$X : \Omega \rightarrow \mathbb{R}$$
.

Definition 2.6 (Probability density function). Suppose we have a real valued random variable X defined on the probability space $(\Omega, \mathcal{F}, \mathbb{P})$. If the derivative of the CDF $F_X(x)$ exists, then we define the probability density function $(pdf) f_X(x)$ of X to be the derivative of $F_X(x)$. Namely,

$$f_X(x) = \frac{dF_X(x)}{dx}$$
.

Example 4. Suppose we toss a fair coin. We have the probability space $(\Omega, \mathcal{F}, \mathbb{P})$, where $\Omega = \{H, T\}$, $\mathcal{F} = \{\emptyset, \{H\}, \{T\}, \{H, T\}\}$ and

 $\mathbb{E}x^{(t+1)} = \mathbb{E}x^{(t)} - \alpha_t \mathbb{E}\nabla f(x^{(t)}; y_{i_t})$

$$\mathbb{P}(\emptyset) = 0$$
, $\mathbb{P}(\{H\}) = \mathbb{P}(\{T\}) = 1/2$, $\mathbb{P}(\{H, T\}) = 1$.

$$\mathbb{P}(\{HH, HT\}) = 1/2$$

Theorem 1.1. If F is such that

- $\mathbb{E}_{i_t} \|\nabla f_{i_t}(x^{(t)})\|^2 \le M + M_G \|\nabla F(x^{(t)})\|^2$
- $\bullet \ \ F \ \ is \ strongly \ \ convex, \ with \ \ constant \ c.$
- $\alpha_t = \frac{\beta}{\gamma + t}$, where $\beta > 1/c$, $\gamma > 0$ and $\alpha_1 < \frac{1}{LM_G}$, then

$$\mathbb{E}\left[F(x^{(t)})\right] - F(x^*) \leq \frac{v}{\gamma + t} \to 0 \ \text{as} \ t \to \infty,$$

$$v = \max \left\{ \frac{\beta^2 LM}{2(\beta c - 1)}, (\gamma + 1)(F(x^{(1)}) - F(x^*)) \right\}$$

$$F(x) = \frac{1}{N} \sum_{i=1}^{N} f(x; y_i)$$

$$=: \frac{1}{N} \sum_{i=1}^{N} f_i(x).$$

$$\mathbb{E}[(X - \mathbb{E}[X])^2]$$

$$= \mathbb{E}[X^2 - 2X\mathbb{E}[X] + \mathbb{E}[X]^2]$$

$$= \mathbb{E}[X^2] - 2\mathbb{E}[X]\mathbb{E}[X] + \mathbb{E}[X]^2$$

$$= \mathbb{E}[X^2] - \mathbb{E}[X]^2$$

$$\begin{split} & \mathbb{E}_{i_t} \left[F(x^{(t+1)}) \right] - F(x^{(t)}) \\ & \leq -\alpha_t \nabla F(x^{(t)})^\top \mathbb{E}_{i_t} \left[\nabla f_{i_t}(x^{(t)}) \right] + \frac{1}{2} \alpha_t^2 L \mathbb{E}_{i_t} \left[\| \nabla f_{i_t}(x^{(t)}) \|^2 \right]. \end{split}$$

Proof. Since $x^{(t+1)} = x^{(t)} - \alpha_t \nabla f_{i_t}(x^{(t)})$. Using Taylor expansion, there exists some z such that

$$\begin{split} F(x^{(t+1)}) &= F(x^{(t)} - \alpha_t \nabla f_{i_t}(x^{(t)})) \\ &= F(x^{(t)}) - \alpha_t \nabla F(x^{(t)})^\top \nabla f_{i_t}(x^{(t)}) + \frac{1}{2} \alpha_t^2 \nabla f_{i_t}(x^{(t)})^\top \nabla^2 F(z) \nabla f_{i_t}(x^{(t)}). \end{split}$$

$$\begin{split} F(x^{(t+1)}) &= F(x^{(t)}) - \alpha_t \nabla F(x^{(t)})^\top \nabla f_{i_t}(x^{(t)}) + \frac{1}{2} \alpha_t^2 \nabla f_{i_t}(x^{(t)})^\top \nabla^2 F(z) \nabla f_{i_t}(x^{(t)}) \\ &\leq F(x^{(t)}) - \alpha_t \nabla F(x^{(t)})^\top \nabla f_{i_t}(x^{(t)}) + \frac{1}{2} \alpha_t^2 L \| \nabla f_{i_t}(x^{(t)}) \|^2. \end{split}$$

$$\mathbb{E}_{i_t} \left[F(x^{(t+1)}) \right] - F(x^{(t)}) \leq -\alpha_t \nabla F(x^{(t)})^\top \mathbb{E}_{i_t} \left[\nabla f_{i_t}(x^{(t)}) \right] + \frac{1}{2} \alpha_t^2 L \mathbb{E}_{i_t} \left[\| \nabla f_{i_t}(x^{(t)}) \|^2 \right]$$

$$\mathbb{E}_{i_t} \left[\|\nabla f_{i_t}(x^{(t)}) \right] \|^2 \le M + M_G \|\nabla F(x^{(t)})\|^2,$$

where $M_G > 0$.

 $Var(aX) = a^2 Var(X)$. Var(X + Y) = Var(X) + Var(Y).

$$\mathbb{E}_{i_t}\left[F(x^{(t+1)})\right] - F(x^{(t)}) \le -\left\lfloor\left(1 - \frac{1}{2}\alpha_t L M_G\right)\alpha_t\right\rfloor \|\nabla F(x^{(t)})\|^2 + \frac{1}{2}\alpha_t^2 L M.$$

$$\begin{split} \mathbb{E}_{i_t} \left[F(x^{(t+1)}) \right] - F(x^{(t)}) &\leq -\alpha_t \|\nabla F(x^{(t)})\|^2 + \frac{1}{2} L \alpha_t^2 (M + M_G \|\nabla F(x^{(t)})\|^2) \\ &= - \left[\left(1 - \frac{1}{2} \alpha_t L M_G \right) \alpha_t \right] \|\nabla F(x^{(t)})\|^2 + \frac{1}{2} \alpha_t^2 L M. \end{split}$$

$$\leq -\alpha_t \|\nabla F(x^{(t)})\|^2 + \frac{1}{2} L \alpha_t^2 (M + M_G \|\nabla F(x^{(t)})\|^2) \\ = -\left[\left(1 - \frac{1}{2} \alpha_t L M_G\right) \alpha_t\right] \|\nabla F(x^{(t)})\|^2 + \frac{1}{2} \alpha_t^2 L M. \qquad F(x) - F(x^*) \leq \frac{\|\nabla F(x)\|^2}{2c}.$$

$$0 < \alpha \le \frac{1}{LM_G}$$
 and $\alpha \le \frac{1}{c}$,

h step-size α and uniform random choice of i_t from the set $\{1, 2, ..., N\}$ guar

$$\mathbb{E}\left[F(x^{(t)}) - F(x^*)\right] \leq \frac{\alpha LM}{2c} + (1-\alpha c)^{t-1}\left[F(x^{(1)}) - F(x^*) - \frac{\alpha LM}{2c}\right].$$

$$\mathbb{E}\left[F(x^{(t)}) - F(x^*)\right] \le \frac{\alpha LM}{2c} + (1 - \alpha c)^{t-1} \left[F(x^{(1)}) - F(x^*) - \frac{\alpha LM}{2c}\right].$$

To optimize $F(x) = \frac{1}{N} \sum_{i=1}^{N} f_i(x)$, the idea is to update

$$x^{(t+1)} = x^{(t)} - \alpha_t g(x^{(t)}),$$

where $g(x) = \frac{1}{m} \sum_{k=1}^{m} \nabla f_{ik}(x)$. Here i_k is randomly drawn independently, with equal probability from $\{1, \dots, N\}$. Here again, as in "vanilla" SGD:

$$\mathbb{E}g(x) = \mathbb{E}\left[\frac{1}{m}\sum_{k=1}^{m}\nabla f_{i_k}(x)\right] = \frac{1}{m}\sum_{k=1}^{m}\mathbb{E}\nabla f_{i_k} = \frac{1}{m}\cdot m\cdot\mathbb{E}\nabla f_{i_k} = \frac{1}{N}\sum_{i=1}^{N}f_i(x) = \nabla F(x).$$

$$\mathbb{E}\left[F(x^{(t+1)}) - F(x^*)\right] - \frac{\alpha LM}{2c} \le (1 - \alpha c)^t \left[E\left(F(x^{(1)}) - F(x^*)\right) - \frac{\alpha LM}{2c}\right].$$

So far, we showed that under assumptions

- $\|\nabla F(x)\| \le L$ for all x.
- $\mathbb{E}_{i_t} \left[\| \nabla f_{i_t}(x^{(t)}) \|^2 \right] \le M + M_G \| \nabla F(x^{(t)}) \|^2$.
- \bullet F is strongly convex with constant c.
- $0 < \alpha \le \frac{1}{LM_G}$ and $\alpha \le \frac{1}{c}$.