Notes for EECS 550: Information Theory

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Chapter 1

Source Codes

1.1 Lossless Coding

Lossless coding is a type of data compression.

GOAL to encode data into bits so that

- 1. bits can be decoded perfectly or with very high accuracy back into original data;
- 2. we use as few bits as possible.

We need to model for data, a measure of decoding accuracy, a measure of compactness.

MODEL FOR DATA

Definition 1.1.1. A *source* is a sequence of i.i.d (discrete) random variables U_1, U_2, \dots

We would like to assume a known alphabet $A = \{a_1, a_2, \dots, a_Q\}$ and known probability distribution either through probability mass functions $p_U(u) = \Pr[U = u]$.

Definition 1.1.2. Source coding

PERFORMANCE MEASURES A measure of compactness (efficiency)

Definition 1.1.3. *Encoding rate,* also called *rate,* is the average number of encoded bits per data symbol.

There are two versions of average rate:

1. Emprical average rate

$$\langle r \rangle := \lim_{N \to \infty} \frac{1}{N} \sum_{k=1}^{N} L_k(U_1, \dots, U_k),$$

2. Statistical average rate

$$\overline{r} := \lim_{N \to \infty} \frac{1}{N} \sum_{k=1}^{N} \mathbb{E}[L_k(U_1, \dots, U_k)]$$

where L_K is the number of bits out of the encoder after U_k and before U_{k+1} .

Definition 1.1.4. The per-letter frequency of error is defined as

$$\langle F_{LE} \rangle := \lim_{N \to \infty} \frac{1}{N} \sum_{k=1}^{N} I(\hat{U}_k = U_k)$$

and per-letter error probability is defined as

$$p_{LE} := \lim_{N \to \infty} \frac{1}{N} \sum_{k=1}^{N} \mathbb{E}[I(\hat{U}_k = U_k)] = \lim_{N \to \infty} \frac{1}{N} \sum_{k=1}^{N} \Pr(\hat{U}_k = U_k)$$

1.1.1 Fixed-length to Fixed-length Block Codes (FFB)

characteristics

Definition 1.1.5. A code is *perfectly lossless* (PL) if the $\beta(\alpha(\underline{u})) = \underline{u}$ for all $\underline{u} \in A_U^k$ (the set of all sequences u_1, \ldots, u_k).

In order to be perfectly loss, α must be one-to-one. Encode must assign a distinct codeword (L bits) to each data sequences. rate = L/K. We seek $R_{PL}^*(k)$ the smallest rate of any PL code.

Number of sequences of size $k=Q^k$, and number binary sequence of size $L=2^L$. We need $2^L\gg Q^K$.

$$\overline{r} = \frac{L}{k} \ge \frac{k \log_2 Q}{k} = \log_2 Q$$

Choose $\lceil k \log_2 Q \rceil$, then we have

$$\begin{split} R_{PL}^*(k) &= \frac{\lceil k \log_2 Q \rceil}{k} \leq \frac{k \log_2 Q + 1}{k} = \log_2 Q + \frac{1}{k}. \\ &\log_2 Q \leq R_{PL}^*(k) \leq \log_2 Q + \frac{1}{k}. \end{split}$$

Let R_{PL}^* be the least rate of any PL FFB code with any k. $R_{PL}^*(k) \to \log_2 Q$ as $k \to \infty$. $R_{PL}^* = \inf_k R_{PL}^*(k)$

Now we want rate less and $\log_2 Q$ almost lossless codes.

 $R_{AL}^* = \inf\{r, \text{there is an FFB code with } \bar{r} \leq r \text{ and arbitrarily small } P_{LE}\}$

= $\inf\{r$, there is an FFB code with $\bar{r} \le r$ and $P_{LE} < \delta$ for all $\delta > 0\}$

Instead of per-letter probability P_{LE} , we focus on block error probability $P_{BE} = \Pr(\underline{\hat{U}} \neq \underline{U})$

Lemma 1.1.1. $P_{BE} \ge P_{LE} \ge \frac{P_{BE}}{k}$

Proof. See homework.

To analyze, we focus on the set of correctly encoded sequences. $G = \{\underline{u} : \beta(\alpha(\underline{u})) = \underline{u}\}$

Then we have

$$P_{BE} = 1 - \Pr[U \in G], |G| \le 2^k, L \ge \lceil \log_2 |G| \rceil.$$

<u>Question</u> How large is the smallest set of sequences with length k form A_U with probability ≈ 1 ?

We need to use weak law of large numbers (WLLN).

Theorem 1.1.1. Suppose $A_x = \{1, 2, ..., Q\}$ with probability $p_1, ..., p_Q$. Given $\underline{u} = (u_1, ..., u_k) \in A_U^k$.

$$n_q(\underline{u}) := \# times \ a_q \ occurs \ in \ \underline{u}, \quad f_q(\underline{u}) = \frac{n_q(\underline{u})}{k} = f requency$$

Fix any $\varepsilon > 0$,

$$\Pr[f_q(\underline{u}) \doteq p_q \pm \varepsilon] \to 1 \text{ as } k \to \infty.$$

Moreover,

$$\Pr[f_q(\underline{u}) \doteq p_q \pm \varepsilon, q = 1, \dots, Q] \to q \text{ as } k \to \infty.$$

NOTATION $a \doteq b \pm \varepsilon \iff |a - b| \le \varepsilon$

Consider subset of A_U^k that corresponds to this event x.

$$T_k = \{\underline{u} : f_q(\underline{u}) \doteq p_q \pm \varepsilon, q = 1, \dots, Q\}.$$

$$\Pr[\underline{U} = \underline{u}] = p(u_1)p(u_2)\dots p(u_k).$$

By WLLN, $\Pr(T_k) \to 1$ as $k \to \infty$.

KEY FACT all sequences in T_k have approximately the same probability.

For $\underline{u} \in T_k$,

$$\begin{split} p(\underline{u}) &= p(u_1)p(u_2)\dots p(u_k) \\ &= p_1^{n_1(u)}p_2^{n_2(u)}\dots p_k^{n_k(u)} \\ &= p_1^{kf_1(u)}p_2^{kf_2(u)}\dots p_k^{kf_k(u)} \\ &\approx \tilde{p}^k \text{ where } \tilde{p} = p_1^{p_1}p_2^{p_2}\dots p_O^{p_Q}. \end{split}$$

So we have $|T_k| \approx \frac{1}{\tilde{p}^k}$.

Then we have

$$\overline{r} = \frac{\log_2 |T_k|}{k} = -\frac{k \log_2 \tilde{p}}{k} = -\log_2 \tilde{p}.$$

Is that rate good? Can we do better? Can we have a set S with probability ≈ 1 and significantly smaller?

Since $\Pr(\underline{U} \in A_U^k \setminus T_k) \approx 0 \implies \Pr(\underline{U} \in S) \approx \Pr(\underline{U} \in S \cap T_k) \approx \frac{|S|}{|T_k|}$. So when k is large, T_k is the smallest set with large probability. And $R_{AL}^* \approx -\log \tilde{p}$.

How to express \tilde{p} .

$$-\log ilde{p} = -\log \prod_{i=1}^Q p_i^{p_i}$$

$$= -\sum_{i=1}^Q p_i \log p_i =: \mathrm{entropy} = H.$$

Some properties of *H*:

- 1. its unit is bits
- 2. H > 0.
- 3. $H = 0 \iff p_q = 1 \text{ for some } q$.
- 4. $H \leq \log_2 Q$.
- 5. $H = \log_2 Q \iff p_q = \frac{1}{Q}$ for all q.

Identify the set that WLLN says has probability \rightarrow 1. Suppose X_1, X_2, \dots i.i.d. real-valued variables.

$$T = \{\underbrace{x_1 \dots x_n}_{x} \in A_X^N : \frac{1}{N} \sum_{i=1}^N x_i \doteq \overline{x} \pm \varepsilon \}$$

is called a typical set. $\Pr(\underline{X} \in T) \approx 1$ when *N* is large.

Now suppose X_1, X_2, \ldots i.i.d. A_x -valued random variables, function $g: A_x \to \mathbb{R}$. Con-

sider $Y_1, Y_2, ...$ with $Y_i = g(X_i)$. Y_i 's are i.i.d. random variables.

If $\mathbb{E}[g(X)]$ is finite than we can apply WLLN that

$$\Pr\left(\frac{1}{N}\sum_{i=1}^{N}Y_{i} \doteq \mathbb{E}[Y] \pm \varepsilon\right) \to 1 \quad \text{as} \quad N \to \infty.$$

Typical sequences wrt g:

$$T_{x,p_{\lambda},g,\varepsilon}^{N} = \left\{ \underline{x} : \frac{1}{N} \sum_{i=1}^{N} g(x_i) \doteq \overline{g(X)} \pm \varepsilon \right\}.$$

If $\mathbb{E}[g(X)]$ is finite then by WLLN we have

$$\Pr(\underline{X} \in T_q) \to 1 \quad \text{as} \quad N \to \infty.$$

Example 1.1.1 (Indicator function). Suppose $F \subset A_X$, and $g(x) = \begin{cases} 1 & x \in F \\ 0 & x \notin F \end{cases}$. Then $\frac{1}{N} \sum_{i=1}^{N} g(x_i) = f_F(x)$. Now

$$T_q = \{\underline{x} : f_F(x) \doteq \Pr(X \in F) \pm \varepsilon\}.$$

By WLLN,

$$\Pr(\underline{X} \in T_q) \to 1$$
 as $N \to \infty$, $\Longrightarrow \Pr(n_F(\underline{X}) \doteq \mathbb{E}[x] \pm \varepsilon) \to 1$.

Example 1.1.2. $A_x=\mathbb{R}, g(x)=x^2.$ $T_g=\{\underline{x}:\}$

Theorem 1.1.2. Now suppose M functions g_1, g_2, \ldots, g_M . Fix ε . Then

$$T_{g_1,g_2,...,g_M} = \bigcap_{i=1}^{M} T_{g_i}.$$

$$\Pr(\underline{X} \in T_{q_1,q_2,...,q_M}) \to 1 \quad as \quad N \to \infty.$$

Proof.

$$\Pr(\underline{X} \notin T_{g_1,g_2,...,g_M}) = \Pr\left(\underline{X} \in \left(\bigcap_{i=1}^M T_{g_i}\right)^c\right)$$

$$= \Pr\left(\underline{X} \in \left(\bigcup_{i=1}^M T_{g_i}^c\right)\right)$$

$$\leq \sum_{i=1}^M \Pr(\underline{X} \in T_{g_i}^c) \to 0 \quad \text{as} \quad N \to \infty.$$

IMPORTANT APPLICATION

Suppose $A_x=\{a_1,\ldots,a_Q\}$ a finite alphabet with probability p_1,\ldots,p_Q . The $g_q(x)$ be the indicator of a_q . $T_q=\{\underline{x}:f_q(\underline{x})\doteq p_q\pm\varepsilon\}$. And $\tilde{T}=\bigcap_{i=1}^Q T_i=\{\underline{x}:\forall q,f_q(x)\doteq p_q\pm\varepsilon\}$. $\tilde{T}^N_{X,p_X,\varepsilon}$ very typical sequence. We have

$$\Pr(\underline{X} \in \tilde{T}) \to 1$$
 as $N \to \infty$.

If $\underline{x} \in \tilde{T}$, then $\underline{x} \in \tilde{T}_g$ for any other g. Consider any real-valued g. If $\underline{x} \in \tilde{T}_{\varepsilon}$ then $\underline{x} \in T_{g,\varepsilon c}$ for some c.

$$\frac{1}{N} \sum_{i=1}^{N} g(x_i) = \sum_{q=1}^{Q} \frac{n_q(x)}{N} g(Q_q) = \sum_{q=1}^{Q} (p_q \pm \varepsilon) q(a_q) = \mathbb{E}[g(X)] + \varepsilon \sum_{q=1}^{Q} g(Q_q)$$

 $\Pr(\underline{X} \in \tilde{T}) \to 1 \text{ as } N \to \infty.$

If $x \in \tilde{T}$,

$$p(\underline{x}) = p(x_1)p(x_2) \dots p(x_N)$$

$$= p_1^{n_1(\underline{x})} \dots$$

$$= p_1^{f_1(\underline{x})N} \dots$$

$$\doteq p_1^{(p_1 \pm \varepsilon)N} \dots$$

$$\doteq 2^N \left(\sum_{q=1}^Q p_q \log p_q \pm \varepsilon \sum_{q=1}^Q \log p_q \right) \qquad \dot{=} 2^{-NH \pm N \varepsilon C}$$

Theorem 1.1.3 (Shannon-McMillian Theorem). Suppose $X_1, X_2, ... i.i.d$, $A_x = \{a_1, ..., a_Q\}$ with probability $p_1, ..., p_Q$. Then

1.

$$\Pr(\tilde{X} \in \tilde{T}_{\varepsilon}^N) \to 1 \text{ as } N \to \infty.$$

2. If $\underline{x} \in \tilde{T}_{\varepsilon}^N$, $p(\underline{x}) \doteq 2^{-NH \pm N\varepsilon c}$.

3.
$$\left| \tilde{T}_{\varepsilon}^{N} \right| \doteq \Pr(\underline{X} \in \tilde{T}_{\varepsilon}^{N}) 2^{N(H \pm \varepsilon x)}$$
.

Proof.

1.2 Shannon-McMillian Theorem

Is \tilde{T} essentially the smallest set with probability ≈ 1 ?

Yes. Let $S \in A_x^N$. We have

$$\Pr(\underline{X} \in S = \Pr(X \in S \cap \tilde{T}) + \Pr(X \in S \cap \tilde{T}^c) \ddot{=} |S \cap \tilde{T}| 2^{-NH \pm 2N\varepsilon c} + \Pr(\tilde{T}^c) \to 0 \text{ as } N \to \infty.$$

Theorem 1.2.1. For every $\varepsilon > 0$, there is a sequence $b_{\varepsilon,1}, b_{\varepsilon,2}, \ldots s.t.$ $b_{\varepsilon,N} \to 0$ as $N \to \infty$, $b_{\varepsilon,B} \ge 0$.

For any N and any $S \subset A_X^N$,

$$|S| \ge (\Pr(\underline{X} \in S) - b_{\varepsilon,N}) 2^{NH - N\varepsilon c}.$$

An in hindsight shortcut

Let us directly consider

$$T_{S,\varepsilon}^{N} = \left\{ \underline{x} : p(\underline{x}) \doteq 2^{-N(H \pm \varepsilon)} \right\}$$

$$= \left\{ \underline{x} : -\frac{1}{N} \log p(\underline{x}) \doteq H \pm \varepsilon \right\}$$

$$= \left\{ \underline{x} : -\frac{1}{N} \sum_{i=1}^{N} \log p(x_i) \doteq H \pm \varepsilon \right\}$$

compare $\tilde{T}_{\varepsilon}^{N}$ and $T_{s,\varepsilon}^{N}$.

Claim: $\tilde{T}^N_{\varepsilon} \subset T^N_{s,\varepsilon}$ where $c=-;\sum_{q=1}^Q \log p_q.$

Suppose $\underline{x} \in \tilde{T}_{\varepsilon}^{N}$. Show if it is also in $T_{s,\varepsilon}^{N}$. Check the following $p(x) \doteq 2^{-NH \pm N\varepsilon c}$, $-\log p(x) \doteq NH \pm N\varepsilon c$.

$$-\log p(\underline{x}) = -\log \prod_{i=1}^{N} p(x_i)$$

$$= -\log \prod_{q=1}^{Q} p_q^{n_q(x)}$$

$$= -\log \prod_{q=1}^{Q} p_q^{Nf_q(x)}$$

$$\doteq -\log \prod_{q=1}^{Q} p_q^{N(p_q \pm \varepsilon)}$$

$$\doteq -\sum_{q=1}^{Q} N(p_q \pm \varepsilon) \log p_a$$

$$\doteq NH \pm N\varepsilon \sum_{q=1}^{C} \log p_k$$

$$\doteq NH \pm N\varepsilon c.$$

Extreme example:

$$\begin{split} A_x &= \{0,1\}, p_0 = p_1 = \frac{1}{2}. \ H = 1. \\ p(\underline{x}) &= 2^{-N}. \\ T_{s,\varepsilon}^N &= \left\{\underline{x}: p(\underline{x}) = 2^{-N(H\pm\varepsilon)} = 2^{-N}\right\} = A_X^N. \\ \tilde{T}_{\varepsilon}^N &= \left\{\underline{x}: n_1(\underline{x} \doteq N\left(\frac{1}{2} + \varepsilon\right))\right\}. \\ |T_{s,\varepsilon}^N| &\coloneqq 2^{N(H\pm\varepsilon)}, |\tilde{T}_{\varepsilon}^N| &\coloneqq 2^{N(H\pm2\varepsilon c)}. \end{split}$$

 T_s is called probability typical. \tilde{T} is called frequency typical.

Example
$$A_x = \{0,1\}$$
, $p_1 = \frac{1}{4}$, $p_0 = \frac{3}{4}$. $\tilde{T}_{\varepsilon}^N = \{\underline{x}: f_1(\underline{x}) \doteq \frac{1}{4} + \varepsilon\}$. $T_{s,\varepsilon}^T = \{\underline{x}: f_1(\underline{x}) = \frac{1}{4} \pm N\varepsilon\log\frac{1-p_1}{p_1}\}$

Typical sequences for an infinite alphabet

There are two cases: A_x is countably infinite / random variables are continuous

In the first case, frequency typical approach doesn't work. Probabilistic typical approach works just as is. $H = -\sum_{q=1}^{\infty} p_q \log p_q$ can be infinite.

Let $S_{\delta,N}=$ size of the smallest set of N sequences form A_x with probability at least $1-\delta$. Then for any $0<\delta<1$ and any h, $\frac{S_{\delta,N}}{2^Nh}\to\infty$ as $N\to\infty$.

1.3 Fixed Length to Variable Length (FVB) Lossless Source codes

Recall that FFB perfectly lossless has $R_{PL}^* = \log_2 |A_x|$, and FFB almost lossless has $R_{AL}^* = H$.

FVB perfectly lossless $R_{VL}^* \le \log_2 |A_x|$.

Suppose we have a source with $A_x = \{a, b, c, d\}$ with probability $\{\frac{1}{2}, \frac{1}{4}, \frac{1}{8}, \frac{1}{8}\}$.

p(u)	u	code1	code2	code3	code4	code5	code6
$\frac{1}{2}$	a	00	0	0	0	0	0
$\frac{1}{4}$	b	01	10	10	10	1	01
$\frac{1}{8}$	c	10	110	10	11	01	011
$\frac{1}{8}$	d	11	111	11	111	10	0111
	Rate	2	1.75	1.5	1.625	1.25	1.875

We can see that code 3-5 are all bad.

Code 6 has an advantage that you know 0 represents the start of a codeword. We will see later why (Example 1.4.2).

FVB source code is characterized by

- source length k
- codebook of binary codewords $C = \{\underline{v}_1, \underline{v}_2, \dots, \underline{v}_{Q^K}\}$, $Q = |A_U|$.
- encoding rule $\alpha: A_U^K \to C$
- decoding rule $\beta: C \to A_U^K$.

The encoder operates in block fashion. The decoder does not.

Distinguish codes that look like code2 and codes that look like code6.

Definition 1.3.1. A codebook *C* is *prefix-free* if no codeword is the prefix of another.

A prefix-free code is called a prefix code. We will stick to prefix codes until states otherwise. (instantaneously decodable)

We like to draw binary tree diagrams of code.

Code 1:

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A prefix is perfectly lossless if and only if α is 1-to-1. The rate: $\overline{r}(c) = \frac{\overline{L}}{K} = \frac{1}{K} \sum_{\underline{u}} p(\underline{u}) L(\underline{u})$

(length of codeword assigned to \underline{u})

 $R_{VL}^*(k) = \min \{ \overline{r}(c) : c \text{ is perfectly lossless FVB with source length } k \}$.

 $R_{VL}^* = \inf\left\{\overline{r}(c) : c \text{ is PL FVB prefix code with any source length}\right\} = \inf_K R_{VL}^*(k).$

How does one design a prefix code to have small or smallest rate?

Focus first k = 1. Shannon's idea: $L_q \approx -\log_2 p_q$.

$$\sum_{q=1}^{Q} p_q L_q \approx -\sum_{q=1}^{Q} p_q \log p_q = H.$$

Question. Is there a prefix code with $L_q \approx -\log p_q$ for $q=1,2,\ldots,Q$? Could there be prefix codes with even smaller rate?

Theorem 1.3.1 (Kraft inequality theorem). *There is a binary prefix code with length* L_1, L_2, \ldots, L_Q *iff the Kraft sum*

$$\sum_{q=1}^{Q} 2^{-L_q} \le 1.$$

Proof. Suppose $\underline{v}_1, \dots, \underline{v}_Q$ is a prefix code with length L_1, \dots, L_Q . Let $L_{\max} = \max_q L_q$.

From the tree, the number of sequences of length L_{\max} prefixed by any codeword, is $\sum_{q=1}^Q 2^{L_{\max}-L_q} \leq 2^{L_{\max}} \implies \sum_{q=1}^Q 2^{-L_q} \leq 1$. So the Kraft inequality holds.

Now suppose

$$L_q = \left[-\log_2 p_q \right], q = 1, \dots, Q. \tag{1.3.1}$$

Is there a code with these lengths? Check Kraft.

$$\sum_{q=1}^{Q} 2^{-L_q} = \sum_{q=1}^{Q} 2^{-\lceil -\log p_q \rceil}$$

$$\leq \sum_{q=1}^{Q} 2^{-(-\log p_q)}$$

$$\leq \sum_{q=1}^{Q} 2^{\log p_q}$$

$$\leq \sum_{q=1}^{Q} p_q = 1.$$

So the Kraft inequality holds. \exists a prefix code with length L_1, \ldots, L_Q given by (1.3.1), called Shannon-Fano code.

Now the question is how good is this Shannon-Fano Code?

For the Shannon-Fano code, the rate (average length) is

$$\overline{L}_{SF} = \sum_{q=1}^{Q} p_q \left[-\log p_q \right].$$

We have the following bounds:

$$H = \sum_{q=1}^{Q} p_q(-\log p_q) \le \overline{L}_{SF} < \sum_{q=1}^{Q} p_q(-\log p_q + 1) = H + 1.$$

Question. Can we do better now?

We will show that $\overline{L} \ge H$ for any prefix code.

Let C be a prefix code with length L_1, \ldots, L_Q . Take the difference $\overline{L} - H = \sum_{q=1}^Q p_q L_q + \sum_{q=1}^Q p_q \log p_q$.

$$\begin{split} \overline{L} - H &= \sum_{q=1}^{Q} p_q L_q + \sum_{q=1}^{Q} p_q \log p_q \\ &= -\sum_{q} p_q \log \frac{2^{-L_q}}{p_q} \\ &= -\sum_{q} p_q \ln \frac{2^{-L_q}}{p_q} \frac{1}{\ln(2)} \\ &\geq -\sum_{q} p_q \left(\frac{2^{-L_q}}{p_q} - 1 \right) \frac{1}{\ln(2)} \\ &\geq -\frac{1}{\ln(2)} \sum_{q} 2^{-L_q} + \sum_{q} p_q \frac{1}{\ln(2)} = \frac{1}{\ln 2} (1 - 1) = 0. \end{split}$$

In homework we will that that \overline{L} can get very close to H+1.

Now allow $k \geq 1$. WE have a $C = \{\underline{v}_1, \dots, \underline{v}_{Q^k}\}$ of length L_1, \dots, L_{Q^k} . We want small

$$\overline{r}(c) = \frac{\overline{L}}{K} = \frac{\sum_{u} p(\underline{u}) L(\underline{u})}{k}.$$

Shannon-Fano code achieve that

$$H^k \le \overline{L}_{SF} < H^k + 1 \implies \frac{H^k}{k} \le \overline{r}_{SF} = \frac{\overline{L}_{SF}}{k} < \frac{H^k}{k} + \frac{1}{k}.$$

Since $H^k = kH$ we have

$$H \le \overline{r}_{SF} < H^k + \frac{1}{k}.$$

Similarly we have for any prefix code, we have

$$\overline{r} = \frac{\overline{L}}{K} \ge \frac{H^k}{k} = H.$$

This leads to a new coding theorem.

Theorem 1.3.2. Given i.i.d. source U with alpha A_U and entropy H. We have

- 1. For every k, $R_{VL}^* \le R_{VL}^*(k) < H + \frac{1}{k}$.
- 2. For every k, $R_{VL}^*(k) \ge R_{VL}^* \ge H$.

Combined we have

$$\forall k \in \mathbb{Z}_{>0}, H \le R_{VL}^*(k) < H + \frac{1}{k}$$

and

$$R_{VL}^* = H.$$

1.4 Huffman's Code Design

Given p_1, \ldots, p_Q , it finds a prefix code with smallest \overline{L} .

Algorithm 1.4.1 Huffman Code

Input: Alphabet probability $\{p_i|i=1,\ldots,Q\}$, WLOG assume $p_1\geq p_2\geq \ldots \geq p_Q$.

Output: FVB Codebook for alphabet $\{a_i | i = 1, ..., Q\}$.

- 1: **function** HUFFMAN($P_Q = \{p_i | i = 1, ..., Q\}$)
- 2: **if** Q = 2 **then return** $\{0, 1\}$
- 3: end if
- 4: $p'_{Q-1} \leftarrow p_{Q-1} + p_Q$
- 5: $P_{Q-1} \leftarrow (P_Q \setminus \{p_{Q-1} + p_Q\}) \cup \{p_{Q-1}'\}$
- 6: $c_{Q-1} \leftarrow \text{HUFFMAN}(P_{Q-1}) =: \{\underline{v}_1, \dots, \underline{v}_{Q-1}\}$
- 7: $c_Q \leftarrow \{\underline{v}_1, \dots, \underline{v}_{Q-2}, \underline{v}_{Q-1}0, \underline{v}_{Q-1}1\}$
- 8: end function

Proposition 1.4.1. If c_{Q-1} is optimal for P_{Q-1} then c_Q is optimal for P_Q .

Example 1.4.1.

We found that

Buffering Yiwei Fu

But there is a tighter upper bound

$$\overline{L}^* \le \begin{cases} H + p_{\text{max}} & p_{\text{max}} < \frac{1}{2} \\ H + p_{\text{max}} + 0.086 & p_{\text{max}} \ge \frac{1}{2}. \end{cases}$$

Hence

$$\mathscr{R}_{VL}^{*}(k) \leq \begin{cases} H + \frac{p_{\max}^{k}}{k} & (p_{\max})^{k} < \frac{1}{2} \\ H + \frac{p_{\max}^{k}}{k} + \frac{0.086}{k} & (p_{\max})^{k} \geq \frac{1}{2}. \end{cases}$$

Up till now we've only focused on i.i.d RV's. Now suppose RV's are dependent, then

$$\frac{H^k}{k} < \frac{kH}{k}$$
.

For a stationary random process,

$$\frac{H^k}{k} \searrow H_{\infty}.$$

For example, English has

$$H^1 \approx 4.08, H_\infty \approx 1.$$

The bits produced by a good lossless source code ($\overline{r} \approx H$) are approximately i.i.d. equiprobable.

Synchronization and transmission entropy

Example 1.4.2. Suppose $\{01, 001, 101, 110\}$ for $\{a, b, c, d\}$.

$$\underline{u} = ddddddddd\dots \implies \underline{z} = 110110110110110110\dots$$
 (if one leading 1 is missing) $\underline{z}' = 101101101101\dots \implies \hat{u} = cccccc\dots$

Now if $\{1,01,001,000\}$ for $\{a,b,c,d\}$. Then the same problem will not happen.

1.5 Buffering

Suppose the source is outputting at R symbols per second. The encoder would have $R\overline{r}$ bits per second.

Buffer overflow happens when a long sequence of low probability symbols are encoded. Buffer underflow happens when a long sequence of high probability symbols are encoded. Buffer will be empty. Include an additional codeword in codebook called a "flag". Insert this codeword when buffer becomes empty.

We focused prefix codes. There are some non prefix codes that can be decoded losslessly.

Definition 1.5.1. A code is *separable* if any finite sequence of codewords is different from any other finite sequence of codewords.

Remark. Prefix codes are separable. And determining if a non prefix code is separable is not easy.

Could prefix codes have smaller *I*?

McMillian's theorem says that Kraft inequality holds for separable codes. If you have a separable code whose codeword satisfy Kraft, then there is a prefix code with same lengths.

Lossless coding for source with infinite alphabet.

Suppose $A_n = \{1, 2, 3, \ldots\}.$

- 1. FFB codes can't have finite rate if perfectly lossless
- 2. AL FFB then SM theorem
- 3. FVB. Current approach is based on Kraft inequality. It still holds for infinite case. (See Apppendix). But Huffman's optimal design does not apply.

Other forms of variable length lossless source codes.

- 1. Run-length coding
- 2. Dictionary coding

1.6 Universal Source Coding

Suppose you are to encode 10^6 symbols from alphabet $A=\{a,b,c,d\}$. We can calculate $n_a(\underline{u}), n_b(\underline{u}), n_c(\underline{u}), n_d(\underline{u})$ and similarly, frequencies. Then we can apply Huffman or Shannon-Fano code.

Chapter 2

Entropy

The star of this chapter is

$$\ln x \le x - 1$$

2.1 Entropy

Entropy

$$H := -\sum_{x} p(x) \log p(x)$$

is a measure of randomness or uncertainty.

$$H_q := -\sum_x p(x) \log_q p(x), \quad H_q = H_r \frac{1}{\log_r q}.$$

$$H(X) = \mathbb{E}[\log p_X(X)]$$

 $\textit{Remark.} \qquad 1. \ \ H(X) \geq 0 \ \text{and} \ \ H(X) = 0 \iff p(x) = 1 \ \text{for some} \ x.$

- 2. $H(X) = \infty$ if X is continuous or has continuous component.
- 3. $H(X,Y) \ge H(X)$
- 4. H(X,Y) = H(X) + H(Y) if X and Y are independent.
- 5. $H(X_1, X_2, \dots, X_n) = H(X_{\sigma(1)}, \dots, X_{\sigma(n)})$ for any $\sigma \in S_n$.

Divergence is a measure of dissimilarity of two probability distribution.

Definition 2.1.1. Suppose p and q are probability mass functions. The *divergence* from p

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to q is

$$D(p||q) = \sum_{x} p(x) \log \frac{p(x)}{q(x)}$$

Remark. 1. $p = q \implies D(p||q) = 0$

2. It is *not symmetric*. $D(p||q) \neq D(q||p)$. You can make symmetric by taking the sum, but then it is not nicely related to information theory.

What if p(x) = 0 for some x? We take $0 \log \frac{0}{q(x)} = 0$. So if $\exists x \ s.t. \ q(x) = 0$ and $p(x) \neq 0$. Then $D(p||q) = \infty$.

When alphabet A_x is infinite, D(p||q) can be ∞ even when p(x) > 0 and q(x) > 0 for all $x \in A_x$.

Is $\sum_{x} p(x) \log \frac{p(x)}{q(x)}$ is always defined? Write

$$\sum_{x} p(x) \log \frac{p(x)}{q(x)} = \sum_{x, p(x) > q(x)} \log \frac{p(x)}{q(x)} + \sum_{x, p(x) < q(x)} \log \frac{p(x)}{q(x)}$$

We will show later that the second term is never $-\infty$, so it is always well-defined.

Proposition 2.1.1 (Divergence inequality). *For any* p, q,

$$D(p||q) \ge 0, D(p||q) = 0 \iff p = q.$$

Proof.

$$D(p||q) = \sum_{x} p(x) \log \frac{p(x)}{q(x)}$$

$$= \sum_{x} p(x) \ln \frac{p(x)}{q(x)} \frac{1}{\ln 2}$$

$$= -\sum_{x} p(x) \ln \frac{q(x)}{p(x)} \frac{1}{\ln 2}$$

$$\leq -\sum_{x} p(x) \frac{q(x) - p(x)}{p(x)} \frac{1}{\ln 2}$$

$$\leq -\left(\sum_{x} p(x) - q(x)\right) \frac{1}{\ln 2} = 0$$

For first equality, \iff is clear. Now suppose D(p||q) = 0. Then

$$\ln \frac{q(x)}{p(x)} = \frac{q(x)}{p(x)} - 1 \implies p(x) = q(x) \text{ for all } x \text{ with } p(x) > 0.$$

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Let's rewrite the divergence inequality a little bit.

$$0 \le D(p||q) = \sum_{x} p(x) \log \frac{p(x)}{q(x)} = -H(x) - \sum_{x} p(x) \log(q_x)$$

$$H(x) \le -\sum_{x} p(x) \log(q_x)$$

with $=\iff p=q$. $-\sum_x p(x)\log(q_x)$ is called cross entropy.

Definition 2.1.2. The *cross entropy* of p with respect to q is

$$H_c(p,q) := -\sum_x p(x) \log q(x).$$

Cross-entropy inequality: for any p, q

$$H_p(X) \leq H_c(p,q)$$

with $= \iff p = q$.

Remark.

$$D(p||q) = H_c(p,q) - H_p(X) \iff H_c(p,q) = H_p(X) + D(p||q).$$

Definition 2.1.3. Variation distance

$$V(p,q) = \sum_{x} |p(x) - q(x)|$$

How does D(p||q) compare to V(p,q)?

Proposition 2.1.2 (Pinskev's inequality).

$$V(p,q) \le \sqrt{(2\ln 2)D(p\|q)}.$$

So small $D(p||q) \implies$ small V(p,q) even when $|A_x| = \infty$. On the other hand the converse is not true.

Lemma 2.1.1. If there exists $0 < \delta < 1$ such that $\frac{|p(x) - q(x)|}{p(x)} \le \delta$ for all x s.t. p(x) > 0 then

$$D(p||q) \le \frac{\delta}{1-\delta} \frac{1}{\ln 2}$$

If $D(p||q) \approx 0$ then $p \approx q$, meaning $V(p,q) \approx 0$.

If p, q are percentage wise close, then $D(p||q) \approx 0$.

Log-sum inequality

Suppose $u_1, u_2, \ldots, u_n, v_1, \ldots, v_n$ nonnegative. Then

$$\sum_{i} u_i \log \frac{u_i}{v_i} \ge \left(\sum_{i} u_i\right) \log \frac{\sum_{i} u_i}{\sum_{i} v_i}.$$

This is a generalization of divergence inequality.

2.2 Basic Properties of Entropy

Proposition 2.2.1.

$$H(X^N) \le \sum_{i=1}^N H(X_i)$$

 $with = \iff X_1, \dots, X_N \text{ are independent.}$

Proof.

$$H(X^N) \le H_L(p,q)$$

where p is probability mass function of X_1, \ldots, X_N . Choose

$$q(x_1,\ldots,x_n)=p(x_1)p(x_2)\ldots p(x_n).$$

Since we are dealing with discrete random variables, it is useful to think about probability mass function as a set of probabilities $\{p_1, p_2, \ldots\}$. Write

$$H(p_1, p_2, \ldots) = -\sum_i p_i \log p_i$$

Let $p'_i = p_i + p_j$, replace p_i, p_j with p'_i and leave all others the same. We have

$$-p_i \log p - i - p_i \log p_i \ge -(p_i + p_i) \log(p_i + p_i)$$

i.e. entropy decreases when two probabilities are merged.

Proposition 2.2.2. *If* X *is* Q-ary with $Q < \infty$ then

$$H(X) \le \log_2 Q$$

Proof. Let $q(x) = \frac{1}{Q}$.

$$\begin{split} H(X) &\leq H_c(p,q) = -\sum_x p(x) \log p(x) \\ &= \sum_x p(x) \log_2 Q = \log_2 Q. \end{split}$$

Proposition 2.2.3. Suppose Y = g(X). Then

$$H(Y) = H(g(X)) \le H(X)$$

 $and = \iff g \text{ is one-to-one (probabilistically)}.$

Proof.

$$H(Y) = \sum_{y} p(y) \log p(y)$$

where $p(y) = \sum_{x,g(x)=y} p(x)$.

$$p_i = q^i(1-q), i = 0, 1, 2, \ldots$$
, then

$$H(X) = -\sum_{i=0}^{\infty} p_i \log p_i$$

$$= -\sum_{i=0}^{\infty} q^i (1-q) \log q^i (1-q)$$

$$= -\sum_{i=0}^{\infty} q^i (1-q) (i \log q + \log(1-q))$$

$$= -\log q \sum_{i=0}^{\infty} q^i (1-q)i - \sum_{i=0}^{\infty} q^i (1-q) \log(1-q)$$

$$= -\log q \cdot q (1-q) \frac{d}{dq} \sum_{i=0}^{\infty} q^i - \log(1-q)$$

$$= -\log q \cdot q (1-q) \frac{1}{(1-q)^2} - \log(1-q)$$

$$= -\log q \cdot \frac{q}{1-q} - \log(1-q) = \frac{-q \log(q) - (1-q) \log(1-q)}{1-q}$$

$$= H(Q) \frac{1}{1-q} < \infty.$$

 $p_i = \frac{\alpha}{i(\ln i)^2}, i = 2, 3, \ldots$, then

$$H(X) = -\sum_{i=2}^{\infty} \frac{\alpha}{i(\ln i)^2} \log \left(\frac{\alpha}{i(\ln i)^2}\right)$$

2.3 Conditional Entropy

$$H(X \mid Y) = \sum_{x,y} p(x,y) \log p(x|y) \ge 0$$

with equality iff X is a function of Y.

$$H(X \mid Y) \le H(X)$$

with equality iff X, Y are independent.

chain rule:

$$H(X,Y) = H(X) + H(Y \mid X) \implies H(X,Y) \ge H(X)$$

Convexity Yiwei Fu

Conditional lossless source coding

2.4 Convexity

Goal: entropy is a concave (convex \cap).

Extended definition of entropy

Definition 2.4.1.

$$\overline{H}(x) = \sup_{\text{finite quantizers } Q} H(Q(x))$$

where finite quantizer is a function $Q: A \to B, |\{Q(x): x \in A\}| < \infty$.

This gives normal definition for discrete random variables and ∞ for continuous and mixed random variables

Chapter 3

Information

3.1 Information

Not a good question: How much information is there in *X*? Better questions: The information in *X* about random variable *Y*?

Definition 3.1.1. The (mutual) information given by *X* about *Y* is defined as

$$I(X;Y) := H(Y) - H(Y \mid X)$$

Definition 3.1.2. Y = outcome of a fair 6-sided die. X = oddity of the outcome. We have

$$H(Y) = \log_2 6, H(Y \mid X) = \log_2 3 \implies I(X; Y) = \log_2 6 - \log_2 3 = \log_2 2 = 1.$$

Lemma 3.1.1. Suppose X, Y are discrete random variables. Then

1. $I(X;Y) \ge 0$, = 0 iff X, Y independent.

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2. We have alternate formulas

$$\begin{split} I(X;Y) &= -\sum_{y} p(y) \log p(y) + \sum_{x,y} p(x,y) \log p(y|x) \\ &= -\sum_{x,y} p(x,y) \log p(y) + \sum_{x,y} p(x,y) \log p(y|x) \\ &= \sum_{x,y} p(x,y) \log \frac{p(y|x)}{p(y)} \\ &= \sum_{x,y} p(x,y) \log \frac{p(x,y)}{p(x)p(y)} \end{split}$$

This shows that I is symmetric: I(X;Y) = I(Y;X).

3. From above,

$$I(X;Y) = -\sum p(x,y)\log p(x) - \sum p(x,y)\log p(y) + \sum p(x,y)\log p(x,y)$$

= $H(X) + H(Y) - H(X,Y)$.

4. We can view information as an expectation:

$$I(X;Y) = \mathbb{E}\left[\log\frac{p(Y\mid X)}{p(Y)}\right] = \mathbb{E}\left[\log\frac{p(X\mid Y)}{p(X)}\right] = \mathbb{E}\left[\log\frac{p(X,Y)}{p(X)p(Y)}\right].$$

5. We can view information with respect to divergence:

$$I(X,Y) = D(p_{XY} || p_X p_Y).$$

Remark. Alternate notation $I_{\phi}(X,Y), I_{X;Y}(p), I(p)$.

What happens if p(x,y)=0, or p(x)=0, or p(y)=0? p(x)=0 or $p(y)=0 \implies p(x,y)=0$.

Remark. I(X;Y) is possible. Suppose $H(Y)=\infty$ and Y is a function of X. Then $I(X;Y)=H(Y)-H(Y\mid X)=\infty-0=\infty$.

Information for more variables

$$I(X, Y; V, W, Z) = H(X, Y) - H(X, Y \mid V, W, Z)$$

Relations between information entropy

1.
$$I(X;Y) = H(X) - H(X \mid Y) = H(Y) - H(Y \mid X)$$
.

- 2. $I(X;Y) \leq H(X)$, = iff X is a function of Y.
- 3. I(X; X) = H(X).
- 4. I(X; g(X)) = H(g(X)).

3.2 Conditional Information

Recall that we have two concepts of conditional entropy.

$$H(X \mid Y = y), H(X \mid Y) = \sum_{y} p(y)H(X \mid Y = y)$$

We are going to use the same approach for conditional information.

Definition 3.2.1. Suppose X, Y, Z are two discrete random variables,

$$I(X;Y|Z=z) = \sum_{x,y} p(x,y) \log \frac{p(xy \mid Z=z)}{p(x|Z=z)p(y|Z=z)}.$$

$$I(X;Y\mid Z) = \sum_{z} p(z)I(X;Y\mid Z=z).$$

Lemma 3.2.1. 1. $I(X; Y \mid Z = z) \ge 0$

- 2. $I(X;Y \mid Z) \ge 0$, = 0 iff X, Y are conditionally independent given Z.
- 3. $I(X; Y \mid Z) = H(X \mid Z) H(X \mid Y, Z)$.
- 4. $I(X; Y, Z) = I(X; Z) + I(X; Y \mid Z)$

chain rule of conditional information

$$I(X;YZ\mid U) = I(X;Z\mid U) + I(X;Y\mid Z,U).$$

3.3 Cryptography From Information Perspective

$$A_X = A_K \text{ and } |A_X| = |A_K| = 2^N. \ p_K(k) = 2^{-N}, k \in A_K.$$

If $|K| < |A_X|$ then the crypto system is not perfect.

Fix $x \in |A_X|$. For each y we have P(Y = y | X = x) = P(Y = y) > 0. Therefore for each y, there must be some key $k \in K$ such that $y = e_K(x)$. It follows that $|K| \ge |Y|$. The encryption is injective giving $|Y| \ge |A_X|$.

Chapter 4

Continuous Random Variables

4.1

Definition 4.1.1.

We assume alphabet is \mathbb{R} , X is absolutely continuous

Example 4.1.1. 1. Uniform

$$p(x) = \begin{cases} \frac{1}{b-a} & a \le x \le b\\ 0 & \text{otherwise} \end{cases}$$

2. Gaussian

$$p(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-m)^2}{2\sigma^2}}$$

3. Laplacian

$$p(x) = \frac{1}{\sqrt{2}\sigma} e^{-\frac{\sqrt{2}}{\sigma}|x|}$$

4. Exponential

$$p(x) = \begin{cases} \frac{\sqrt{2}}{\sigma} e^{-\frac{\sqrt{2}}{\sigma}|x|} & x \ge 0\\ 0 & x < 0 \end{cases}$$

Definition 4.1.2. The *support* of a random variable X or of its probability distribution is defined by

$$S := \{x : \Pr(X \doteq x \pm \varepsilon) > 0, \forall \varepsilon > 0\}.$$

Definition 4.1.3. Conditional probability

$$\Pr(F \mid X = x) = \frac{\Pr(F, X = x)}{\Pr(X = x)}.$$

When X is continuous,

$$\Pr(F \mid X = x) = \lim_{\delta \to 0} \Pr(F \mid X \doteq x \pm \delta)$$

Suppose Y = 3X

$$1=\Pr(Y=3\mid X=1)=\lim_{\delta=0}\frac{\Pr(Y=3\mid X\doteq 1\pm \delta)}{\Pr(X\doteq 1\pm \delta)}=0, \text{ contradition}$$

So we use

$$\Pr(Y \in F \mid X = x) = \lim_{\varepsilon \to 0} \lim_{\delta \to 0} \Pr(Y \in F_{\delta} \mid X \doteq x \pm \varepsilon)$$

where

$$F_{\delta} = \{y \mid ||y - y'|| \text{ for some } y' \in F\}.$$

Generalized sum