

## Segment 2: Randomized Studies

### Section 03: Blocked and Paired Designs

# Motivating Idea: Reducing “Unlucky” Assignments

- ▶ Randomized assignment will give us the “right” answer *on average*
  - ▶ Because potential outcomes will be *balanced* between groups, on average
  - ▶ Covariates, too
- ▶ Does not guarantee the estimate for any particular random assignment will be close to the truth
  - ▶ “Unlucky” assignment
- ▶ Reducing the chances of “unlucky” randomizations → decrease the variance of effect estimates
- ▶ The comparison between Bernoulli randomization and the Completely Randomized Design is one example

# Example: Hypothetical Dietary Experiment

## Observed Data

**Table:** Observed Data from the Hypothetical Dietary Experiment, **Idealized Assignment**

Unit, $i$	Female, $x_{1i}$	Age, $x_{2i}$	Treatment $Z_i$	Potential $Y_i^c$	Potential $Y_i^t$
Audrey	1	40	0	140	135
Anna	1	40	1	140	135
Bob	0	50	0	150	140
Bill	0	50	1	150	140
Caitlin	1	60	0	160	155
Cara	1	60	1	160	155
Dave	0	70	0	170	160
Doug	0	70	1	170	160

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## Observed Data

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**Question:** Is there anything we can do in our randomized study design to reduce the chance of this type of unlucky imbalance?

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**Key Idea:** If there are certain types of units with the same potential outcomes, randomize (possibly differently) within these types of units (i.e., within *blocks*)

- ▶ We will never know the potential outcomes
- ⇒ determine *randomization blocks* based on observed covariates,  $\mathbf{X}$ 
  - ▶ E.g., if we expect sex and age to be important in determining blood pressure, consider sex/age strata as *blocking factors* to ensure equal numbers in each treatment group
  - ▶ Guarantee that all randomizations will be *balanced* with respect to these factors
- ▶ Result will be a randomization distribution that *only* considers randomizations that are balanced with respect to these factors
- ⇒ Fewer “unlucky” randomizations, less uncertainty in the inference

# Block Randomized Experiment

- ▶ Divide units into  $J$  blocks based on covariates
  - ▶  $B_i = B(\mathbf{X}_i) \in \{1, 2, \dots, J\}$
- ▶ Conduct a completely randomized experiment in each block:

$$Pr(\mathbf{Z}|\mathbf{X}, \mathbf{Y}^c, \mathbf{Y}^t) = \prod_{j=1}^J \binom{n(j)}{n_t(j)}^{-1}$$

- ▶ Set of possible  $\mathbf{Z}$  is:  $\{\mathbf{Z}; \sum_{i:B_i=j}^n Z_i = n_t(j) \text{ for } j = 1, 2, \dots, J\}$
- ▶ Randomization depends on  $\mathbf{X}$  through definition of blocks

# Ignorability and Unconfoundedness

For a block-randomized experiment, the assignment mechanism:

$$Pr(\mathbf{Z}|\mathbf{X}, \mathbf{Y}^c, \mathbf{Y}^t) = \prod_{j=1}^J \binom{n(j)}{n_t(j)}^{-1}$$

Depends (overall) on:

- ▶ Unit characteristics,  $\mathbf{X}$
- ▶ Potential outcomes
- ⇒ Not unconfounded!

Block-randomized design is only unconfounded *within* blocks  
(just a completely-randomized design within each block)

Formalized with the assumption of *conditional ignorability*



# Implications of (Conditional) Ignorability

$$Z \perp\!\!\!\perp Y^c, Y^t | \mathbf{X}$$

Within blocks (or strata) defined by  $\mathbf{X}$ ,

- ▶ Observed covariates provide information about treatment assignment
- ▶ Within values of  $\mathbf{X}$ , potential outcomes do not
- ▶ Within values of  $\mathbf{X}$ , comparison of observed outcomes in treated/control groups will provide an **unbiased estimate** of the conditional average causal effect,  $\tau_{CATE|X}$
- ▶ An unbiased estimate of  $\tau_{SATE}$  can be obtained by averaging block-specific estimates

**Key Implication:** The analysis must “adjust” for  $X$  in order to provide unbiased estimation of  $\tau_{PATE}$

- ▶ A weighted average of block-specific estimates
- ▶ A regression that adjusts for indicator of block

# Example: Hypothetical Dietary Experiment

Now with more units!

Table: Observed Data from the Hypothetical Dietary Experiment

Unit, $i$	Female, $x_{1i}$	Age, $x_{2i}$	Treatment $Z_i$	Potential $Y_i^c$	Potential $Y_i^t$
Audrey	1	40	0	140	135
Abigail	1	40	0	140	135
Arielle	1	40	0	140	135
Anna	1	40	1	140	135
Bob	0	50	0	150	140
Bill	0	50	0	150	140
Burt	0	50	0	150	140
Brad	0	50	1	150	140
Caitlin	1	60	0	160	155
Cara	1	60	1	160	155
Cassie	1	60	1	160	155
Cindy	1	60	1	160	155
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Doug	0	70	1	170	160
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# Example: Hypothetical Dietary Experiment

Now with more units! And blocks!

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# Need to Account for Design in the Analysis

If observations were randomized within blocks, *we need to take this into account when we analyze the results of the trial*

- ▶ Calculate treatment effects within each block
  - ▶ Randomized within block  $\Rightarrow$  unbiased for treatment effect in that block
- ▶ Average block-specific estimates to estimates
  - ▶ Represent the effect in the sample (or population)
  - ▶  $\tau_{SATE}$  or  $\tau_{PATE}$

# Examples of Blocking Factors

Blocking factors work best when they arise “naturally”

- ▶ Randomize students within blocks defined by school
  - ▶ Avoid problems where school membership is *imbalanced* across treated/control students
- ▶ Randomize products within blocks defined by product class
  - ▶ Avoid problems where a treatment group of products is dominated by products in a certain class
- ▶ Randomize surgery recovery programs within blocks defined by the type of surgery
  - ▶ Make sure there are equal numbers of pancreatic, liver, stomach cancer patients in the two treatment groups

**Bottom Line:** Define blocks based on covariates believed to be predictive of the outcome, but balance against practical constraints

# Matched Pairs Design

**Basic Idea:** push the idea of blocking to the extreme  $\Rightarrow$  have as many blocks as units in the treatment group

- ▶ Special case of randomized block design
- ▶ Units are arranged into pairs based on having closely-matched characteristics,  $\mathbf{X}$ 
  - ▶ 2-unit blocks
- ▶ Randomize exactly one unit per pair to receive treatment
- ▶ **Practical limitation** is the ability to find units that are well matched based on the entire  $\mathbf{X}$
- ▶ Can arise naturally when pairs present themselves “naturally”
  - ▶ E.g., twins, family members

# Paired Randomized Experiment

Same as a blocked experiment where there are 2 units within each stratum and one is randomized to treatment:

$$Pr(\mathbf{Z}|\mathbf{X}, \mathbf{Y}^c, \mathbf{Y}^t) = 2^{n/2}$$

- ▶ Set of possible  $\mathbf{Z}$  is:  $\{\mathbf{Z}; \sum_{i:B_i=j}^n Z_i = 1 \text{ for } j = 1, 2, \dots, n/2\}$
- ▶ Extreme stratified/block experiment
- ▶ Randomization depends on  $\mathbf{X}$  through definition of pairs

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# Analysis of Matched Pair Designs

- ▶ Ignorability:  $Z \perp\!\!\!\perp Y^c, Y^t | \mathbf{X}$
- ▶ Equal probability of treatment in each block  $\Rightarrow$  adjustment for pair not required for unbiasedness
- ▶ For efficiency gains, analysis must account for pairs
- ▶ Analysis:
  - ▶ Simple difference in means:
  - ▶ Average of within-pair differences: