R Script

```
Part 1
```

```
library(cobalt)
data(lalonde, package = "MatchIt")
install.packages("causaldata")
library(causaldata)
l1 <- lalonde
love.plot(treat ~ age + educ + race + married + nodegree + re74 + re75, data = l1, stars="std")
mod_la1 <- lm(re78 ~ treat, data = l1)
mod_la2 <- lm(re78 ~ treat + age + educ + race + married +
        nodegree + re74 + re75, data = l1)
library(texreg)
screenreg(list(mod_la1, mod_la2))
library(causaldata)
library(cem)
library(MatchIt)
# Breaks for educ, age, and race using cutpoints code
table(I1$educ)
table(I1$age)
educut <- c(0, 6.5, 8.5, 12.5, 16)
agecut <- c(16, 22, 28, 34, 40, 46, 55)
# New dataset with matched pairs
mat1 <- cem(treatment = "treat", data = I1, drop = "re78",
      cutpoints = list(educ = educut, age = agecut))
mat1
#
      G0 G1
       429 185
#All
#Matched 116 96
#Unmatched 313 89
# Estimated effect
m_ate <- att(mat1, re78 ~ treat, data = 11)
summary(m_ate)
```

```
Covariate Balance
      age*
     educ*
 race_black
race_hispan
                                               Sample
 race_white
                                                Unadjusted
   married
  nodegree
     re74*
     re75*
               -0.4
                         0.0
                   Mean Differences
```

Assignment 2

Linear regression model estimated on matched data only

```
Coefficients:
```

```
Estimate Std. Error t value p-value
(Intercept) 4910.55 644.58 7.6183 8.667e-13 ***

treat 1638.10 957.87 1.7101 0.08872 .
---

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1

m_ate2 <- att(mat1, re78 ~ treat + age + educ + race + married + nodegree + re74 + re75, data = l1)

summary(m_ate2)
```

Treatment effect estimation for data:

```
G0 G1
All 429 185
Matched 116 96
Unmatched 313 89
```

match.l1

Linear regression model estimated on matched data only

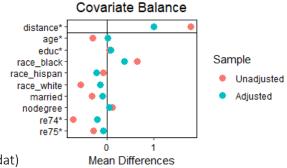
```
Coefficients:
      Estimate Std. Error t value p-value
      (Intercept) -8458.88346 5173.37751 -1.6351 0.10359
              1778.06402 944.91127 1.8817 0.06131.
      age
               101.09146 71.67635 1.4104 0.15996
               881.15966 388.83549 2.2662 0.02450 *
      educ
      racehispan 666.44976 2732.38587 0.2439 0.80755
      racewhite 600.60117 1655.58616 0.3628 0.71715
      married -1807.37453 2370.67861 -0.7624 0.44672
      nodegree 1808.08457 1530.99012 1.1810 0.23900
      re74
              re75
                Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1
# Genetic Matching (w/out and w/ covariates)
install.packages("rgenoud")
set.seed(123)
match.l1 <- matchit(treat ~ age + educ + race + married +
          nodegree + re74 + re75,
         data = I1, method = "genetic",
         replace = FALSE, pop.size = 50, print = 0) #, caliper = 0.4)
```

A matchit object

- method: 1:1 genetic matching without replacement
- distance: Propensity score
 - estimated with logistic regression
- number of obs.: 614 (original), 370 (matched)
- target estimand: ATT
- covariates: age, educ, race, married, nodegree, re74, re75

love.plot(match.l1, stars ="std") #to see adjusted covariate balance match_dat <- match.data(match.l1)</pre>

w/out covariates
mod_la_match1 <- lm(re78 ~ treat, data = match_dat)
screenreg(mod_la_match1)</pre>



Propensity Score Matching

library(tidyverse) library(haven)

Call:

glm(formula = treat \sim age + educ + race + married + nodegree + re74 + re75, family = binomial(link = "logit"), data = |1)

Deviance Residuals:

Min 1Q Median 3Q Max -1.7645 -0.4736 -0.2862 0.7508 2.7169

Coefficients:

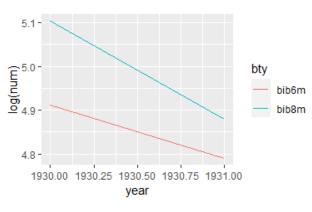
Estimate Std. Error z value Pr(>|z|)
(Intercept) -1.663e+00 9.709e-01 -1.713 0.08668 .
age 1.578e-02 1.358e-02 1.162 0.24521
educ 1.613e-01 6.513e-02 2.477 0.01325 *
racehispan -2.082e+00 3.672e-01 -5.669 1.44e-08 ***
racewhite -3.065e+00 2.865e-01 -10.699 < 2e-16 ***
married -8.321e-01 2.903e-01 -2.866 0.00415 **
nodegree 7.073e-01 3.377e-01 2.095 0.03620 *
re74 -7.178e-05 2.875e-05 -2.497 0.01253 *

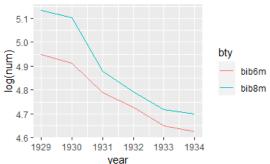
Assignment 2

```
5.345e-05 4.635e-05 1.153 0.24884
       re75
       Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
|1control <- |1 %>%
 mutate(pscore = logit_l1$fitted.values)
# mean pscore
pscore_control <- l1control %>%
 filter(treat == 0) %>%
 pull(pscore) %>%
 mean()
pscore_treated <- |1control %>%
 filter(treat == 1) %>%
 pull(pscore) %>%
 mean()
summary(logit_l1)
m.nn \leftarrow matchit(treat age + educ + race + nodegree + re74 + re75, data = 11,
        method= "nearest", ratio = 1)
# estimating treatment effect using PSM
mdata = match.data(m.nn)
names(mdata)
match.data <- mdata
avg.income78.treated=weighted.mean(mdata$re78[mdata$treat == 1],
                  mdata$weights[mdata$treat==1])
avg.income78.control=weighted.mean(mdata$re78[mdata$treat==0],
                  mdata$weights[mdata$treat==0])
avg.income78.treated
avg.income78.control
avg.income78.treated- avg.income78.control
# estimating effect using regression --> get same result as method above of +278
lm_treat1 <- lm(re78 ~ treat, data = match.data)</pre>
screenreg(lm_treat1)
```

```
Model 1
_____
(Intercept) 6070.34 ***
      (531.30)
treat
         278.80
      (751.38)
R^2
          0.00
Adj. R^2
           -0.00
Num. obs. 370
_____
*** p < 0.001; ** p < 0.01; * p < 0.05
#Installing packages
install.packages("tidyverse")
library(tidyverse)
install.packages("tidyr")
library(tidyr)
install.packages("lubridate")
library(lubridate)
install.packages("dplyr")
library(dplyr)
install.packages("ggplot2")
library(ggplot2)
library(readr)
banks <- read_csv("C:/Users/Lg/Downloads/banks.csv")
View(banks)
#creating variables
date = banks$date
day = banks$day
month = banks$month
weekday = banks$weekday
year = banks$year
bib6 = banks$bib6
bio6 = banks$bio6
bib8 = banks$bib8
bio8 = banks$bio8
#Calculate the mean number of banks in business each year in the 6th and 8th districts
mean(bib6, trim = 0, na.rm = FALSE) #118.6193
mean(bio6, trim = 0, na.rm = FALSE) #117.1283
mean(bib8, trim = 0, na.rm = FALSE) #133.0367
mean(bio8, trim = 0, na.rm = FALSE) #131.1416
```

Assignment 2





Compared to the graph for years 1930 and 1931, the graph for 1929 - 1934 shows a decline in descension. The banks in the 8th district after 1931 are closing at a slower rate than what it was prior to the implementation of the treatment. Post treatment shows similar trends for the 6th and 8th district. The DiD for the two districts is (120 - 136) - (132 - 165) = -16 - (-33) = 17. In this final graph by using the DiD, we can say that it withstands the parallel trend assumption.

Part 3

```
install.packages("reshape2")
library(reshape2)
install.packages("tidyr")
library(tidyr)
install.packages("magrittr")
library(magrittr)
library(foreign)
library(tidyverse)
traffic1 long <- reshape(data = traffic1, idvar = "state",</pre>
              varying = c("admn90", "admn85", "open90", "open85", "dthrte90", "dthrte85", "speed90",
"speed85", "cdthrte", "cadmn", "copen", "cspeed"),
              v.name = "value",
              time=c("admn90", "admn85", "open90", "open85", "dthrte90", "dthrte85", "speed90",
"speed85", "cdthrte", "cadmn", "copen", "cspeed"),
              new.row.names = 1:1000,
              direction= c("long"))
```

```
traffic1_long1 = gather(traffic1, key = "state", value = "admn90", "admn85", "open90", "open85",
"dthrte90", "dthrte85", "speed90", "speed85", "cdthrte", "cadmn", "copen", "cspeed", na.rm = FALSE,
convert = FALSE, factor_key = FALSE)
#Estimate the difference-in-differences estimate - treatment effect, including interaction term
install.packages("texreg")
library(texreg)
#OLS w/ deaths and open container laws in 1985 and 1990
mod85 open <- lm(dthrte85~open85, data=traffic1)
screenreg(mod85_open)
mod90 open <- lm(dthrte90~open90, data=traffic1)
screenreg(mod90_open)
screenreg(list(mod85_open, mod90_open))
#DiD for open models
DD model open <- lm(cdthrte ~ copen, data = traffic1)
screenreg(DD_model_open)
#OLS w/ deaths and admin laws in 1985 and 1990
mod85_admn <- lm(dthrte85~admn85, data=traffic1)
mod90_admn <- lm(dthrte90~admn90, data=traffic1)
screenreg(list(mod85_admn, mod90_admn))
#DiD for admn models
       _____
             Model 1 Model 2
       _____
       (Intercept) 2.61 *** 2.11 ***
                   (0.11) (0.11)
       admn85
                  0.23
                     (0.17)
       admn90
                       0.07
                   (0.15)
       R^2
               0.03 0.00
       Adj. R^2 0.02
                       -0.02
```

Num. obs. 51

51 _____ *** p < 0.001; ** p < 0.01; * p < 0.05

Assignment 2

Model 1 is the OLS regression with deaths in 1985. Model 2 is the regression for 1990. There is a positive .23 unit increase in traffic deaths whereas, in model 2 there is a 0.07 unit increase. .23 - .07 = .16.

DD_model_admn <- lm(cdthrte ~ cadmn, data = traffic1)

screenreg(list(mod85 admn, mod90 admn, DD model admn))

```
_____
    Model 1 Model 2 Model 3
(Intercept) 2.61 *** 2.11 *** -0.52 ***
    (0.11) (0.11) (0.05)
admn85
      0.23
    (0.17)
admn90
          0.07
        (0.15)
cadmn
              -0.18
           (0.12)
R^2 0.03 0.00 0.04
Adj. R^2 0.02 -0.02 0.02
Num. obs. 51
                51
            51
*** p < 0.001; ** p < 0.01; * p < 0.05
```

#Estimate ATE for both laws

DD_model_both <- Im(cdthrte ~ copen + cadmn, data = traffic1) screenreg(DD_model_both)

Question 1:

Some of the issues that arise when using this cross-sectional data is that states vary in culture, attitude, location, etc, which may have unobserved omitted variables that cannot be controlled for. We also cannot derive counterfactual outcomes from single cross-sectional data as it creates randomization uncertainty. Cross sectional data cannot control for time invariant unobserved heterogeneity like panel and pooled data.

Question 2:

Difference-in-differences would be the method of choice for having data with two periods. This method allows us to utilize the data at two points to establish trends prior to instilling the treatment as a baseline to understand how the two groups relate to each other. Then once the treatment has been applied, we can predict what the counterfactual would have been in the instance that the treatment was not applied, and the difference between the counterfactual and the outcome is the treatment effect. Furthermore, difference-in-differences does not require that the states be alike, due to being able to establish the trends from the two-period data. Controlling for covariates before and after the treatment to avoid bias as a change in one covariate might have a significant effect on the treatment or control group.

Question 5:

We cannot completely prove the parallel assumption only provide the evidence to support it. Pretrends are the typical way, looking at trends in treatment and control group in the years before the treatment. But more should be done in terms of arguing why it should hold at the time of treatment. Some would show that after the implementation of the treatment, the trend became closer to the control group.