Deep Generative Models Probabilistic PCA

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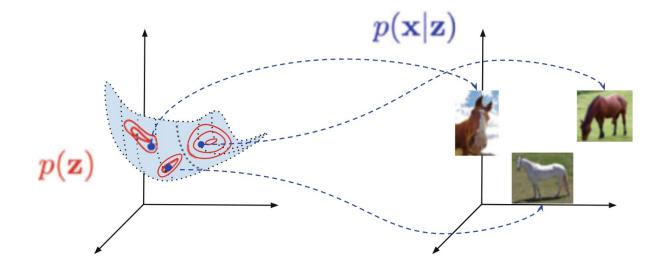
Outline

- Probabilistic PCA: Model
 - Joint, Marginal and Conditional Distributions
 - PPCA as an Encoder-Decoder Architecture
- Linear Algebra Background
 - Singular Value Decomposition (SVD)
 - SVD Properties
- Probabilistic PCA: Learning
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Latent Variable Models

- X = observed variable
- Z = latent variable

- $\mathbf{z} \sim p(\mathbf{z})$
- $x \sim p(x \mid z)$



A latent variable model and a generative process. Note the low-dimensional manifold (here 2D) embedded in the high-dimensional space (here 3D)

Factorization of the joint model

$$p(\mathbf{x}, \mathbf{z}) = p(\mathbf{x} \mid \mathbf{z})p(\mathbf{z})$$

Marginalization of the model

$$p(\mathbf{x}) = \int p(\mathbf{x} \mid \mathbf{z}) p(\mathbf{z}) d\mathbf{z}$$

Probabilistic Principal Component Analysis: Model

• We consider continuous random variables only, i.e.,

$$z \in \mathbb{R}^d$$
 and $x \in \mathbb{R}^D$ with $d \ll D$

• The distribution of z is the standard Gaussian, i.e.,

$$p(\mathbf{z}) = \mathcal{N}(\mathbf{z} \mid \mathbf{0}, \mathbf{I}).$$

• The dependency between z and x is linear, plus Gaussian additive noise:

$$x = Wz + b + \varepsilon$$

• Here $\boldsymbol{\varepsilon} \sim \mathcal{N}(\boldsymbol{\varepsilon} \mid \boldsymbol{0}, \sigma^2 \boldsymbol{I})$ and independent from \boldsymbol{z} .

Probabilistic Principal Component Analysis: Model

PPCA model

$$x = Wz + b + \epsilon$$
, $z \sim \mathcal{N}(z \mid 0, I)$, $\epsilon \sim \mathcal{N}(\epsilon \mid 0, \sigma^2 I)$.

• x is a linear combination of Gaussians, thus $p(x) = \mathcal{N}(x \mid b, WW^{\top} + \sigma^2 I)$ because

$$\mathbb{E}[x] = \mathbb{E}[Wz] + \mathbb{E}[b] + \mathbb{E}[\epsilon] = W\mathbb{E}[z] + b + 0 = b$$

$$\mathbb{V}[x] = \mathbb{V}[Wz + b + \epsilon] = W\mathbb{V}(z)W^{\mathsf{T}} + \mathbb{V}[\epsilon] = WW^{\mathsf{T}} + \sigma^2 I$$

• $x \mid z$ is a constant + a Gaussian, thus $p(x \mid z) = \mathcal{N}(x \mid Wz + b, \sigma^2 I)$ because

$$\mathbb{E}[\mathbf{x} \mid \mathbf{z}] = \mathbf{W}\mathbf{z} + \mathbf{b} + \mathbb{E}[\boldsymbol{\epsilon}] = \mathbf{W}\mathbf{z} + \mathbf{b}$$

$$\mathbb{V}[\mathbf{x} \mid \mathbf{z}] = \mathbb{V}[\boldsymbol{\epsilon}] = \sigma^2 \mathbf{I}$$

Probabilistic Principal Component Analysis: Model

• PPCA model: $\mathbf{x} = \mathbf{W}\mathbf{z} + \mathbf{b} + \boldsymbol{\epsilon}$, $p(\mathbf{z}) = \mathcal{N}(\mathbf{z} \mid \mathbf{0}, \mathbf{I})$, $p(\boldsymbol{\epsilon}) = \mathcal{N}(\boldsymbol{\epsilon} \mid \mathbf{0}, \sigma^2 \mathbf{I})$, $p(\mathbf{x} \mid \mathbf{z}) = \mathcal{N}(\mathbf{x} \mid \mathbf{W}\mathbf{z} + \mathbf{b}, \sigma^2 \mathbf{I}), \qquad p(\mathbf{x}) = \mathcal{N}(\mathbf{x} \mid \mathbf{b}, \mathbf{W}\mathbf{W}^{\mathsf{T}} + \sigma^2 \mathbf{I})$

• Let $M = W^T W + \sigma^2 I$. We can compute the conditional distribution of $(z \mid x)$ as

$$p(z \mid x) = \frac{p(x \mid z)p(z)}{p(x)} \propto e^{-\frac{1}{2\sigma^2}||x - Wz - b||^2} e^{-\frac{1}{2}||z||^2}$$

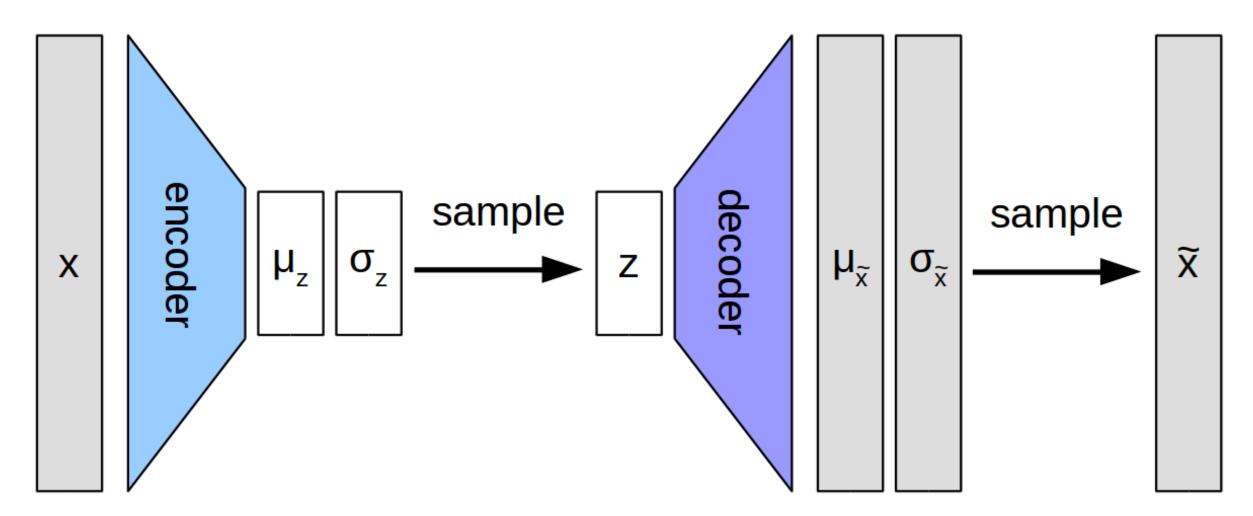
$$p(\boldsymbol{z} \mid \boldsymbol{x}) \propto e^{-\frac{1}{2\sigma^2} (\boldsymbol{z}^T \boldsymbol{W}^T \boldsymbol{W} \boldsymbol{z} - 2\boldsymbol{z}^T \boldsymbol{W}^T (\boldsymbol{x} - \boldsymbol{b}) + \sigma^2 ||\boldsymbol{z}||^2)} \propto e^{-\frac{1}{2\sigma^2} (\boldsymbol{z}^T \boldsymbol{M} \boldsymbol{z} - 2\boldsymbol{z}^T \boldsymbol{W}^T (\boldsymbol{x} - \boldsymbol{b}))}$$

$$p(\boldsymbol{z} \mid \boldsymbol{x}) = \mathcal{N}(\boldsymbol{z} \mid \boldsymbol{M}^{-1}\boldsymbol{W}^{\mathsf{T}}(\boldsymbol{x} - \boldsymbol{b}), \sigma^{2}\boldsymbol{M}^{-1})$$

PPCA as an Encoder Decoder Architecture

• PPCA model: $\mathbf{x} = \mathbf{W}\mathbf{z} + \mathbf{b} + \boldsymbol{\epsilon}$, $p(\mathbf{z}) = \mathcal{N}(\mathbf{z} \mid \mathbf{0}, \mathbf{I})$, $p(\boldsymbol{\epsilon}) = \mathcal{N}(\boldsymbol{\epsilon} \mid \mathbf{0}, \sigma^2 \mathbf{I})$,

$$p(\boldsymbol{z} \mid \boldsymbol{x}) = \mathcal{N}(\boldsymbol{z} \mid M^{-1}W^{\top}(\boldsymbol{x} - \boldsymbol{b}), \sigma^{-2}M) \qquad p(\boldsymbol{x} \mid \boldsymbol{z}) = \mathcal{N}(\boldsymbol{x} \mid W\boldsymbol{z} + \boldsymbol{b}, \sigma^{2}I)$$



Singular Value Decomposition (SVD)

• A matrix $X \in \mathbb{R}^{m \times n}$ of rank r can be decomposed as

$$X = U\Sigma V^{\top} = \begin{bmatrix} u_1 \cdots u_r u_{r+1} \cdots u_m \end{bmatrix} \begin{bmatrix} \sigma_1 & \cdots & 0 & 0 & \cdots & 0 \\ \vdots & \ddots & \vdots & \vdots & \ddots & \vdots \\ 0 & \cdots & \sigma_r & 0 & \cdots & 0 \\ 0 & \cdots & 0 & 0 & \cdots & 0 \\ \vdots & \ddots & \vdots & \vdots & \ddots & \vdots \\ 0 & \cdots & 0 & 0 & \cdots & 0 \end{bmatrix} [v_1 \cdots v_r v_{r+1} \cdots v_m]^{\top}$$

- $\Sigma \in \mathbb{R}^{m \times n}$ is a "diagonal" matrix with r singular values $\sigma_1 \geq \sigma_2 \geq \cdots \geq \sigma_r > 0$
- $U \in \mathbb{R}^{m \times m}$ and $V \in \mathbb{R}^{n \times n}$ are orthonormal matrices of left/right singular vectors
- Orthonormal means that columns are unit norm and orthogonal to each other:

$$U^{\mathsf{T}}U = UU^{\mathsf{T}} = I_{m \times m}, \qquad V^{\mathsf{T}}V = VV^{\mathsf{T}} = I_{n \times n}$$

Compact Singular Value Decomposition (SVD)

• A matrix $X \in \mathbb{R}^{m \times n}$ of rank r can be decomposed as

• A matrix
$$X \in \mathbb{R}^{m \times n}$$
 of rank r can be decomposed as
$$X = U \Sigma V^{\top} = \begin{bmatrix} u_1 \cdots u_r u_{r+1} \cdots u_m \end{bmatrix} \begin{bmatrix} \sigma_1 & \cdots & 0 & 0 & \cdots & 0 \\ \vdots & \ddots & \vdots & \vdots & \ddots & \vdots \\ 0 & \cdots & \sigma_r & 0 & \cdots & 0 \\ 0 & \cdots & 0 & 0 & \cdots & 0 \\ \vdots & \ddots & \vdots & \vdots & \ddots & \vdots \\ 0 & \cdots & 0 & 0 & \cdots & 0 \end{bmatrix} \begin{bmatrix} v_1 \cdots v_r v_{r+1} \cdots v_m \end{bmatrix}^{\top}$$

$$= U_1 \Sigma_1 V_1^{\top} = \begin{bmatrix} u_1 \cdots u_r \end{bmatrix} \begin{bmatrix} \sigma_1 & \cdots & 0 \\ \vdots & \ddots & \vdots & \vdots \\ 0 & \cdots & \sigma_r \end{bmatrix} \begin{bmatrix} v_1 \cdots v_r \end{bmatrix}^{\top}$$

- $U_1 \in \mathbb{R}^{m \times r}$ and $V \in \mathbb{R}^{n \times r}$ are orthonormal matrices of left/right singular vectors
- $\Sigma \in \mathbb{R}^{r \times r}$ is a diagonal matrix with r singular values $\sigma_1 \geq \sigma_2 \geq \cdots \geq \sigma_r > 0$

SVD Properties

We can use any partition of the SVD of a matrix to expand it as

$$X = U\Sigma V^{\mathsf{T}} = \begin{bmatrix} U_1 & U_2 \end{bmatrix} \begin{bmatrix} \Sigma_1 & 0 \\ 0 & \Sigma_2 \end{bmatrix} \begin{bmatrix} V_1 & V_2 \end{bmatrix}^{\mathsf{T}} = U_1\Sigma_1V_1^{\mathsf{T}} + U_2\Sigma_2V_2^{\mathsf{T}}$$

- Example: $I = U_1 U_1^{\mathsf{T}} + U_2 U_2^{\mathsf{T}}$, where $U_1^{\mathsf{T}} U_1 = I$, $U_1^{\mathsf{T}} U_2 = 0$, and $U_2^{\mathsf{T}} U_2 = I$
- If X is invertible, we can use its SVD to compute the inverse of a matrix

$$X = U\Sigma V^{\top} \Rightarrow X^{-1} = V\Sigma^{-1}U^{\top}$$

• We can use any partition of the SVD of a matrix to compute its inverse as

$$X = U_1 \Sigma_1 V_1^{\top} + U_2 \Sigma_2 V_2^{\top} \Rightarrow X^{-1} = V_1 \Sigma_1^{-1} U_1^{\top} + V_2 \Sigma_2^{-1} U_2^{\top}$$

SVD Properties

• We can use the SVD of a square matrix to compute its trace and determinant

$$X = U\Sigma V^{\top} \Rightarrow \operatorname{trace}(X) = \operatorname{trace}(\Sigma) = \sum_{i} \sigma_{i}$$
 and $\det(X) = \det(\Sigma) = \prod_{i} \sigma_{i}$

We can use any partition of the SVD to compute its trace and determinant

$$X = U_1 \Sigma_1 V_1^{\mathsf{T}} + U_2 \Sigma_2 V_2^{\mathsf{T}} \Rightarrow \operatorname{trace}(X) = \operatorname{trace}(\Sigma_1) + \operatorname{trace}(\Sigma_2)$$
$$X = U_1 \Sigma_1 V_1^{\mathsf{T}} + U_2 \Sigma_2 V_2^{\mathsf{T}} \Rightarrow \det(X) = \det(\Sigma_1) \det(\Sigma_2)$$

Example

$$X = \begin{bmatrix} I_m & 0 \\ 0 & \sigma^2 I_n \end{bmatrix} \Rightarrow \operatorname{trace}(X) = m + \sigma^2 n \quad \text{and} \quad \det(X) = \sigma^{2n}$$

• Recall the ML estimators of the parameters of a Gaussian $\mathcal{N}(x \mid \mu, \Sigma)$ are

$$\mu_N = \frac{1}{N} \sum_{i=1}^N x_i$$
, $\Sigma_N = \frac{1}{N} \sum_{i=1}^N (x_i - \mu_N) (x_i - \mu_N)^T$

• For PPCA we need to estimate the parameters of a Gaussian with structured covariance $\Sigma = WW^T + \sigma^2 I$. The estimate of the mean is the same as before $\mu = \mu_N$. To estimate W, we need to maximize the log-likelihood w.r.t. (W, σ)

$$\ell = -\frac{N}{2}\log(\det(\mathbf{\Sigma})) - \frac{N}{2}\operatorname{trace}(\mathbf{\Sigma}^{-1}\mathbf{\Sigma}_N)$$

ullet Taking derivatives w.r.t. $oldsymbol{W}$ we get

$$\frac{\partial \ell}{\partial \mathbf{W}} = \frac{\partial \ell}{\partial \mathbf{\Sigma}} \frac{\partial \mathbf{\Sigma}}{\partial \mathbf{W}} = -\frac{N}{2} (\mathbf{\Sigma}^{-1} - \mathbf{\Sigma}^{-1} \mathbf{\Sigma}_N \mathbf{\Sigma}^{-1}) 2\mathbf{W} = 0 \implies \mathbf{\Sigma}_N \mathbf{\Sigma}^{-1} \mathbf{W} = \mathbf{W}$$

We thus need to solve the nonlinear equations

$$\Sigma_N \Sigma^{-1} W = W$$
 and $\Sigma = W W^T + \sigma^2 I$

- A trivial solution is $(W, \sigma) = (0, 0)$, but this is a minimum of the log-likelihood.
- Another solution is $\Sigma = \Sigma_N$, but this would require the structure of the sample covariance Σ_N to match the structure of $\Sigma = WW^T + \sigma^2 I$, i.e., the smallest eigenvalues would need to be all equal to each other and equal to σ^2 .
- Alternatively, let

$$\boldsymbol{W} = \begin{bmatrix} \boldsymbol{Z}_1 & \boldsymbol{Z}_2 \end{bmatrix} \begin{bmatrix} \boldsymbol{\Gamma}_1 & \boldsymbol{0} \\ \boldsymbol{0} & \boldsymbol{0} \end{bmatrix} [\boldsymbol{V}_1 & \boldsymbol{V}_2]^T = \boldsymbol{Z}_1 \boldsymbol{\Gamma}_1 \boldsymbol{V}_1^T$$

Then

$$\Sigma = WW^{T} + \sigma^{2}I = Z_{1}\Gamma_{1}^{2}Z_{1}^{T} + \sigma^{2}(Z_{1}Z_{1}^{T} + Z_{2}Z_{2}^{T}) = Z_{1}(\Gamma_{1}^{2} + \sigma^{2}I)Z_{1}^{T} + \sigma^{2}Z_{2}Z_{2}^{T}$$

$$\Sigma^{-1}W = (Z_{1}(\Gamma_{1}^{2} + \sigma^{2}I)^{-1}Z_{1}^{T} + \sigma^{-2}Z_{2}Z_{2}^{T})Z_{1}\Gamma_{1}V_{1}^{T} = Z_{1}(\Gamma_{1}^{2} + \sigma^{2}I)^{-1}\Gamma_{1}V_{1}^{T}$$

Therefore,

$$\Sigma_{N}\Sigma^{-1}W = W \Rightarrow \Sigma_{N}Z_{1}(\Gamma_{1}^{2} + \sigma^{2}I)^{-1}\Gamma_{1}V_{1}^{T} = Z_{1}\Gamma_{1}V_{1}^{T} \Rightarrow$$

$$\Sigma_{N}Z_{1}(\Gamma_{1}^{2} + \sigma^{2}I)^{-1} = Z_{1} \Rightarrow \Sigma_{N}Z_{1} = Z_{1}(\Gamma_{1}^{2} + \sigma^{2}I) \Rightarrow \Sigma_{N}Z_{i} = (\gamma_{i}^{2} + \sigma^{2})Z_{i}$$

- In other words, z_i is an eigenvector of Σ_N with eigenvalue $\gamma_i^2 + \sigma^2$.
- Thus, if $\Sigma_N = [U_1 \ U_2] \begin{bmatrix} \Lambda_1 & \mathbf{0} \\ \mathbf{0} & \Lambda_2 \end{bmatrix} [U_1 \ U_2]^T$, then $Z_1 = U_1$, $\Gamma_1^2 + \sigma^2 I = \Lambda_1$
 - In other words, $m{U}_1$ is a matrix whose d columns correspond to d singular vectors of $m{\Sigma}_N$
- Therefore, $\boldsymbol{W} = \boldsymbol{Z}_1 \boldsymbol{\Gamma}_1 \boldsymbol{V}_1^T = \boldsymbol{U}_1 (\boldsymbol{\Lambda}_1 \sigma^2 \boldsymbol{I})^{1/2} \boldsymbol{V}_1^T$
- Having "almost" found W (we don't know which d columns), we now turn to finding σ .

Recall the log-likelihood

$$\ell = -\frac{N}{2}\log(\det(\mathbf{\Sigma})) - \frac{N}{2}\operatorname{trace}(\mathbf{\Sigma}^{-1}\mathbf{\Sigma}_N)$$

• We have $\mathbf{\Sigma}=m{W}m{W}^T+\sigma^2I$ and $m{W}=m{U}_1m{\Gamma}_1m{V}_1^T$. Thus, $m{W}m{W}^T=m{U}_1m{\Gamma}_1^2m{U}_1^T$ and

$$\Sigma = U_1(\Gamma_1^2 + \sigma^2 I)U_1^T + \sigma^2 U_2 U_2^T = U_1 \Lambda_1 U_1^T + \sigma^2 U_2 U_2^T$$

$$\mathbf{\Sigma}^{-1}\mathbf{\Sigma}_{N} = (\mathbf{U}_{1}\mathbf{\Lambda}_{1}^{-1}\mathbf{U}_{1}^{T} + \sigma^{-2}\mathbf{U}_{2}\mathbf{U}_{2}^{T})(\mathbf{U}_{1}\mathbf{\Lambda}_{1}\mathbf{U}_{1}^{T} + \mathbf{U}_{2}\mathbf{\Lambda}_{2}\mathbf{U}_{2}^{T}) = \mathbf{U}_{1}\mathbf{U}_{1}^{T} + \sigma^{-2}\mathbf{U}_{2}\mathbf{\Lambda}_{2}\mathbf{U}_{2}^{T}$$

• Substituting into the log-likelihood, we get

$$\ell = -\frac{N}{2}\log(\det(\mathbf{\Lambda}_1)\sigma^{2(D-d)}) - \frac{N}{2}(d + \sigma^{-2}\operatorname{trace}(\mathbf{\Lambda}_2))$$

- Taking the derivative yields $\frac{\partial \ell}{\partial \sigma^2} = -\frac{N}{2} \left(\frac{D-d}{\sigma^2} \frac{\operatorname{trace}(\Lambda_2)}{\sigma^4} \right) = 0 \Longrightarrow \sigma^2 = \frac{\operatorname{trace}(\Lambda_2)}{D-d}$
- ullet Final piece: we have not shown that $oldsymbol{\Lambda}_1$ corresponds to $oldsymbol{top}$ d eigenvalues of $oldsymbol{\Sigma}_N$
 - Exercise 2.13 in GPCA textbook
- Theorem. The ML estimates for the parameters of the PPCA model b, W, and σ can be obtained from the ML estimates of the mean and covariance of the data, μ_N and Σ_N , respectively, as

$$\boldsymbol{b} = \boldsymbol{\mu}_N$$
, $\boldsymbol{W} = \boldsymbol{U}_1 (\boldsymbol{\Lambda}_1 - \sigma^2 I)^{1/2} \boldsymbol{R}$ and $\sigma^2 = \frac{1}{D-d} \sum_{i=d+1}^{D} \lambda_i$

• where U_1 is the matrix with the top d eigenvectors of Σ_N , Λ_1 is the matrix with the corresponding top d eigenvalues, $R \in \mathbb{R}^{d \times d}$ is an arbitrary orthogonal matrix, and λ_i is the ith largest eigenvalue of Σ_N .

Application of PPCA to Generating Face Images



Fig. 2.2 Face images of subject 20 under 10 different illumination conditions in the extended Yale B data set. All images are frontal faces cropped to size 192×168 .

Application of PPCA to Generating Face Images

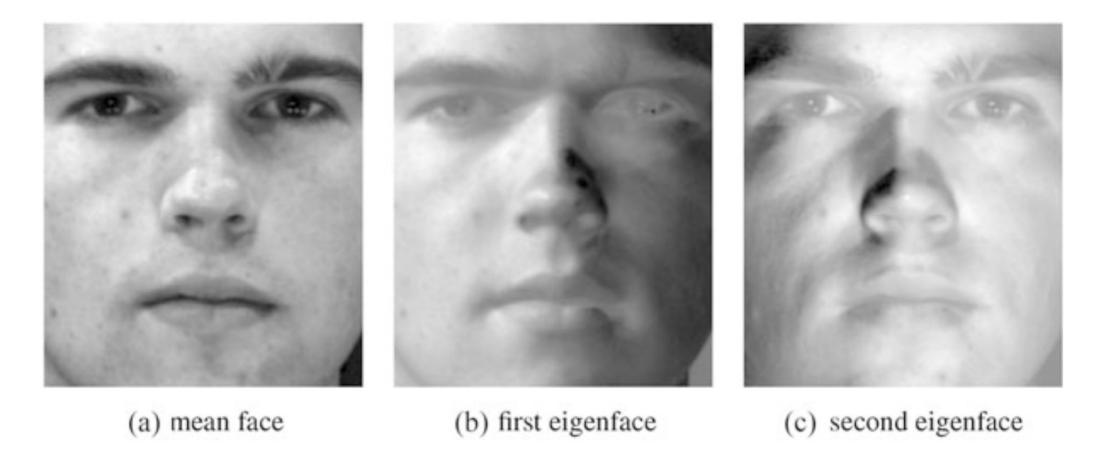


Fig. 2.5 Mean face and the first two eigenfaces by applying PPCA to the ten images in Figure 2.2.

Application of PPCA to Generating Face Images

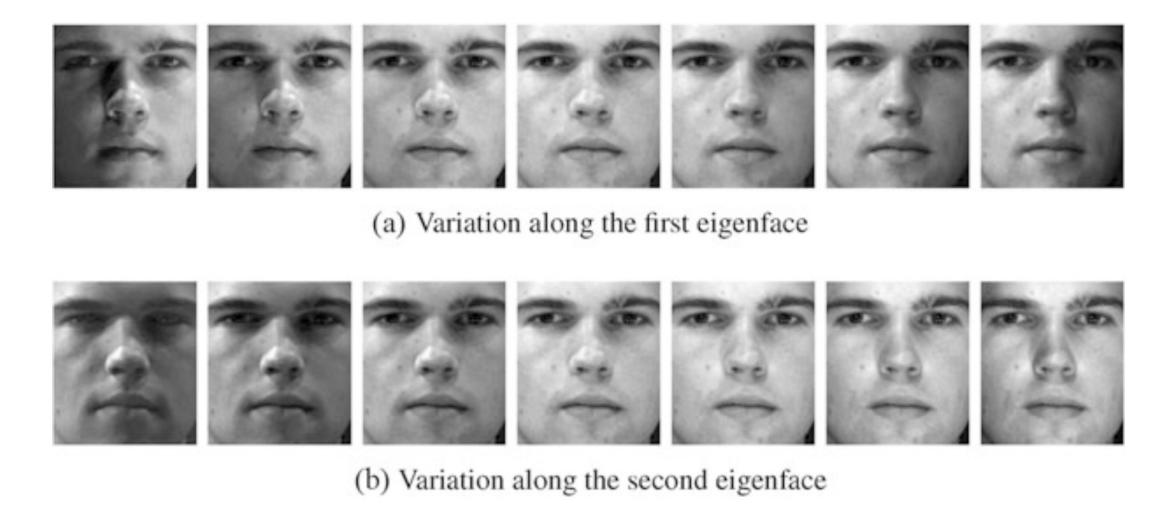


Fig. 2.6 Variation of the face images along the two eigenfaces given by PPCA. Each row plots $\mu + y_i u_i$ for $y_i = -1$: $\frac{1}{3}$: 1, i = 1, 2.