# Deep Generative Models Probabilistic PCA

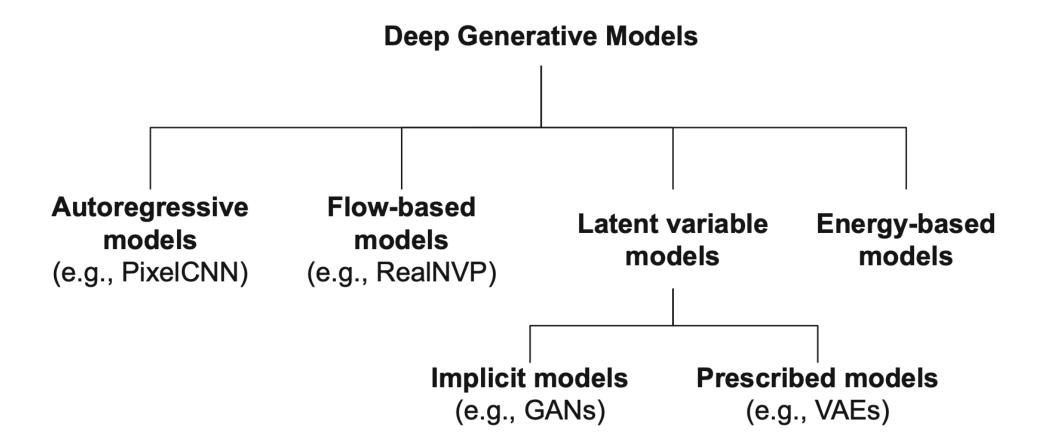
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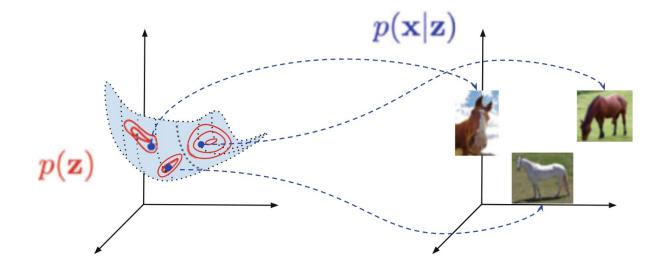
### Taxonomy of Generative Models



#### Latent Variable Models

- X = observed variable
- Z = latent variable

- $\mathbf{z} \sim p(\mathbf{z})$
- $x \sim p(x \mid z)$



A latent variable model and a generative process. Note the low-dimensional manifold (here 2D) embedded in the high-dimensional space (here 3D)

Factorization of the joint model

$$p(\mathbf{x}, \mathbf{z}) = p(\mathbf{x} \mid \mathbf{z})p(\mathbf{z})$$

Marginalization of the model

$$p(\mathbf{x}) = \int p(\mathbf{x} \mid \mathbf{z}) p(\mathbf{z}) d\mathbf{z}$$

#### Probabilistic Principal Component Analysis: Model

• We consider continuous random variables only, i.e.,

$$z \in \mathbb{R}^d$$
 and  $x \in \mathbb{R}^D$  with  $d \ll D$ 

ullet The distribution of  $oldsymbol{z}$  is the standard Gaussian, i.e.,

$$p(\mathbf{z}) = \mathcal{N}(\mathbf{z} \mid \mathbf{0}, \mathbf{I}).$$

• The dependency between z and x is linear and we assume a Gaussian additive noise:

$$x = Wz + b + \varepsilon$$

• Here  $\boldsymbol{\varepsilon} \sim \mathcal{N}(\boldsymbol{\varepsilon} \mid \boldsymbol{0}, \sigma^2 \boldsymbol{I})$  and independent from  $\boldsymbol{z}$ .

#### Probabilistic Principal Component Analysis: Model

PPCA Model

$$x = Wz + b + \epsilon$$
,  $z \sim \mathcal{N}(z \mid 0, I)$ ,  $\epsilon \sim \mathcal{N}(\epsilon \mid 0, \sigma^2 I)$ .

• x is a linear combination of Gaussians, thus  $p(x) = \mathcal{N}(x \mid b, WW^{\top} + \sigma^2 I)$  because

$$\mathbb{E}[x] = \mathbb{E}[Wz] + \mathbb{E}[b] + \mathbb{E}[\epsilon] = W\mathbb{E}[z] + b + 0 = b$$

$$\mathbb{V}[x] = \mathbb{V}[Wz + b + \epsilon] = W\mathbb{V}(z)W^{\mathsf{T}} + \mathbb{V}[\epsilon] = WW^{\mathsf{T}} + \sigma^2 I$$

•  $x \mid z$  is a constant + a Gaussian, thus  $p(x \mid z) = \mathcal{N}(x \mid Wz + b, \sigma^2 I)$  because

$$\mathbb{E}[\mathbf{x} \mid \mathbf{z}] = \mathbf{W}\mathbf{z} + \mathbf{b} + \mathbb{E}[\boldsymbol{\epsilon}] = \mathbf{W}\mathbf{z} + \mathbf{b}$$

$$\mathbb{V}[\mathbf{x} \mid \mathbf{z}] = \mathbb{V}[\boldsymbol{\epsilon}] = \sigma^2 \mathbf{I}$$

# Probabilistic Principal Component Analysis: Model

• PPCA model:  $\mathbf{x} = \mathbf{W}\mathbf{z} + \mathbf{b} + \boldsymbol{\epsilon}$ ,  $p(\mathbf{z}) = \mathcal{N}(\mathbf{z} \mid \mathbf{0}, \mathbf{I})$ ,  $p(\boldsymbol{\epsilon}) = \mathcal{N}(\boldsymbol{\epsilon} \mid \mathbf{0}, \sigma^2 \mathbf{I})$ ,  $p(\mathbf{x} \mid \mathbf{z}) = \mathcal{N}(\mathbf{x} \mid \mathbf{W}\mathbf{z} + \mathbf{b}, \sigma^2 \mathbf{I}), \qquad p(\mathbf{x}) = \mathcal{N}(\mathbf{x} \mid \mathbf{b}, \mathbf{W}\mathbf{W}^{\mathsf{T}} + \sigma^2 \mathbf{I})$ 

• Let  $M = W^T W + \sigma^2 I$ . We can compute the conditional distribution of  $(z \mid x)$  as

$$p(z \mid x) = \frac{p(x \mid z)p(z)}{p(x)} \propto e^{-\frac{1}{2\sigma^2}||x - Wz - b||^2} e^{-\frac{1}{2}||z||^2}$$

$$p(\boldsymbol{z} \mid \boldsymbol{x}) \propto e^{-\frac{1}{2\sigma^2} (\boldsymbol{z}^T \boldsymbol{W}^T \boldsymbol{W} \boldsymbol{z} - 2\boldsymbol{z}^T \boldsymbol{W}^T (\boldsymbol{x} - \boldsymbol{b}) + \sigma^2 ||\boldsymbol{z}||^2)} \propto e^{-\frac{1}{2\sigma^2} (\boldsymbol{z}^T \boldsymbol{M} \boldsymbol{z} - 2\boldsymbol{z}^T \boldsymbol{W}^T (\boldsymbol{x} - \boldsymbol{b}))}$$

$$p(\boldsymbol{z} \mid \boldsymbol{x}) = \mathcal{N}(\boldsymbol{z} \mid \boldsymbol{M}^{-1}\boldsymbol{W}^{\mathsf{T}}(\boldsymbol{x} - \boldsymbol{b}), \sigma^{2}\boldsymbol{M}^{-1})$$

• Recall the ML estimators of the parameters of a Gaussian  $\mathcal{N}(x \mid \mu, \Sigma)$  are

$$\mu_N = \frac{1}{N} \sum_{i=1}^N x_i$$
,  $\Sigma_N = \frac{1}{N} \sum_{i=1}^N (x_i - \mu_N) (x_i - \mu_N)^T$ 

• For PPCA we need to estimate the parameters of a Gaussian with structured covariance  $\Sigma = WW^T + \sigma^2 I$ . The estimate of the mean is the same as before  $\mu = \mu_N$ . To estimate W, we need to maximize the log-likelihood w.r.t.  $(W, \sigma)$ 

$$\ell = -\frac{N}{2}\log(\det(\mathbf{\Sigma})) - \frac{N}{2}\operatorname{trace}(\mathbf{\Sigma}^{-1}\mathbf{\Sigma}_N)$$

ullet Taking derivatives w.r.t.  $oldsymbol{W}$  we get

$$\frac{\partial \ell}{\partial \mathbf{W}} = \frac{\partial \ell}{\partial \mathbf{\Sigma}} \frac{\partial \mathbf{\Sigma}}{\partial \mathbf{W}} = -\frac{N}{2} (\mathbf{\Sigma}^{-1} - \mathbf{\Sigma}^{-1} \mathbf{\Sigma}_N \mathbf{\Sigma}^{-1}) 2\mathbf{W} = 0 \implies \mathbf{\Sigma}_N \mathbf{\Sigma}^{-1} \mathbf{W} = \mathbf{W}$$

We thus need to solve the nonlinear equations

$$\Sigma_N \Sigma^{-1} W = W$$
 and  $\Sigma = W W^T + \sigma^2 I$ 

- A trivial solution is W = 0, but this is a minimum of the log-likelihood.
- Another solution is  $\Sigma = \Sigma_N$ , but this would require the structure of the sample covariance  $\Sigma_N$  to match the structure of  $\Sigma = WW^T + \sigma^2 I$ , i.e., the smallest eigenvalues would need to be all equal to each other and equal to  $\sigma^2$ .
- Alternatively, let

$$W = \begin{bmatrix} \boldsymbol{Z}_1 & \boldsymbol{Z}_2 \end{bmatrix} \begin{bmatrix} \boldsymbol{\Gamma}_1 & \boldsymbol{0} \\ \boldsymbol{0} & \boldsymbol{0} \end{bmatrix} [\boldsymbol{V}_1 & \boldsymbol{V}_2]^T = \boldsymbol{Z}_1 \boldsymbol{\Gamma}_1 \boldsymbol{V}_1^T$$

Then

$$\Sigma = WW^T + \sigma^2 I = Z_1 \Gamma_1^2 Z_1^T + \sigma^2 (Z_1 Z_1^T + Z_2 Z_2^T) = Z_1 (\Gamma_1^2 + \sigma^2 I) Z_1^T + \sigma^2 Z_2 Z_2^T$$

$$\Sigma^{-1} W = (Z_1 (\Gamma_1^2 + \sigma^2 I)^{-1} Z_1^T + \sigma^{-2} Z_2 Z_2^T) Z_1 \Gamma_1 V_1^T = Z_1 (\Gamma_1^2 + \sigma^2 I)^{-1} \Gamma_1 V_1^T$$

Therefore,

$$\Sigma_{N}\Sigma^{-1}W = W \Rightarrow \Sigma_{N}Z_{1}(\Gamma_{1}^{2} + \sigma^{2}I)^{-1}\Gamma_{1}V_{1}^{T} = Z_{1}\Gamma_{1}V_{1}^{T} \Rightarrow$$

$$\Sigma_{N}Z_{1}(\Gamma_{1}^{2} + \sigma^{2}I)^{-1} = Z_{1} \Rightarrow \Sigma_{N}Z_{1} = Z_{1}(\Gamma_{1}^{2} + \sigma^{2}I) \Rightarrow \Sigma_{N}Z_{i} = (\gamma_{i}^{2} + \sigma^{2})Z_{i}$$

- In other words,  $z_i$  is an eigenvector of  $\Sigma_N$  with eigenvalue  $\gamma_i^2 + \sigma^2$ .
- Thus if  $\Sigma_N = [U_1 \ U_2] \begin{bmatrix} \Lambda_1 & \mathbf{0} \\ \mathbf{0} & \Lambda_2 \end{bmatrix} [U_1 \ U_2]^T$ , then  $Z_1 = U_1$ ,  $\Gamma_1^2 + \sigma^2 I = \Lambda_1$ .
  - In other words,  $m{U}_1$  is a matrix whose d columns correspond to d singular vectors of  $m{\Sigma}_N$
- Therefore,  $\boldsymbol{W} = \boldsymbol{Z}_1 \boldsymbol{\Gamma}_1 \boldsymbol{V}_1^T = \boldsymbol{U}_1 (\boldsymbol{\Lambda}_1 \sigma^2 \boldsymbol{I})^{1/2} \boldsymbol{V}_1^T$
- Having "almost" found W (we don't know which d columns), we now turn to finding  $\sigma$ .

Recall the log-likelihood

$$\ell = -\frac{N}{2}\log(\det(\mathbf{\Sigma})) - \frac{N}{2}\operatorname{trace}(\mathbf{\Sigma}^{-1}\mathbf{\Sigma}_N)$$

• We have  $\mathbf{\Sigma}=m{W}m{W}^T+\sigma^2I$  and  $m{W}=m{U}_1m{\Gamma}_1m{V}_1^T$ . Thus,  $m{W}m{W}^T=m{U}_1m{\Gamma}_1^2m{U}_1^T$  and

$$\Sigma = U_1(\Gamma_1^2 + \sigma^2 I)U_1^T + \sigma^2 U_2 U_2^T = U_1 \Lambda_1 U_1^T + \sigma^2 U_2 U_2^T$$

$$\mathbf{\Sigma}^{-1}\mathbf{\Sigma}_{N} = (\mathbf{U}_{1}\mathbf{\Lambda}_{1}^{-1}\mathbf{U}_{1}^{T} + \sigma^{-2}\mathbf{U}_{2}\mathbf{U}_{2}^{T})(\mathbf{U}_{1}\mathbf{\Lambda}_{1}\mathbf{U}_{1}^{T} + \mathbf{U}_{2}\mathbf{\Lambda}_{2}\mathbf{U}_{2}^{T}) = \mathbf{U}_{1}\mathbf{U}_{1}^{T} + \sigma^{-2}\mathbf{U}_{2}\mathbf{\Lambda}_{2}\mathbf{U}_{2}^{T}$$

• Substituting into the log-likelihood, we get

$$\ell = -\frac{N}{2}\log(\det(\mathbf{\Lambda}_1)\sigma^{2(D-d)}) - \frac{N}{2}(d + \sigma^{-2}\operatorname{trace}(\mathbf{\Lambda}_2))$$

- Taking the derivative yields  $\frac{\partial \ell}{\partial \sigma^2} = -\frac{N}{2} \left( \frac{D-d}{\sigma^2} \frac{\operatorname{trace}(\Lambda_2)}{\sigma^4} \right) = 0 \Longrightarrow \sigma^2 = \frac{\operatorname{trace}(\Lambda_2)}{D-d}$
- ullet Final piece: we have not shown  $oldsymbol{\Lambda}_1$  corresponds to  $oldsymbol{top}$  d eigenvalues of  $oldsymbol{\Sigma}_N$ 
  - Exercise 2.13 in GPCA textbook

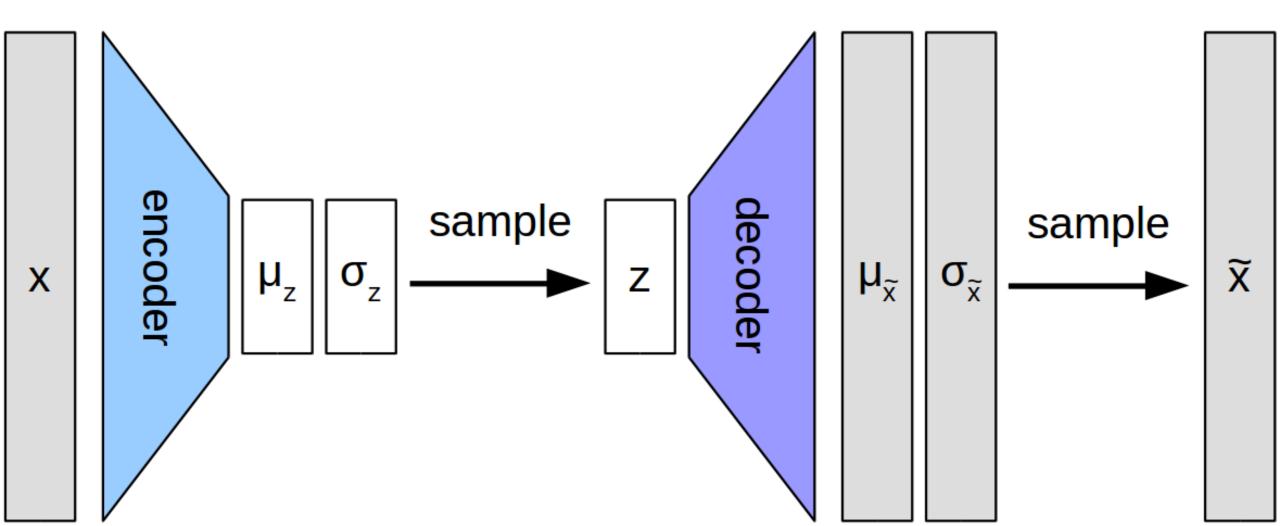
• Theorem. The ML estimates for the parameters of the PPCA model b, W, and  $\sigma$  can be obtained from the ML estimates of the mean and covariance of the data,  $\mu_N$  and  $\Sigma_N$ , respectively, as

$$\boldsymbol{b} = \boldsymbol{\mu}_N$$
,  $\boldsymbol{W} = \boldsymbol{U}_1 (\boldsymbol{\Lambda}_1 - \sigma^2 I)^{1/2} R$  and  $\sigma^2 = \frac{1}{D-d} \sum_{i=d+1}^{D} \lambda_i$ 

• where  $U_1$  is the matrix with the top d eigenvectors of  $\Sigma_N$ ,  $\Lambda_1$  is the matrix with the corresponding top d eigenvalues,  $R \in \mathbb{R}^{d \times d}$  is an arbitrary orthogonal matrix, and  $\lambda_i$  is the ith largest eigenvalue of  $\Sigma_N$ .

#### PPCA as an Encoder Decoder Architecture

• 
$$p(z|x) = \mathcal{N}(z \mid M^{-1}W^{T}(x-b), \sigma^{-2}M)$$
  $p(x \mid z) = \mathcal{N}(x \mid Wz + b, \sigma^{2}I)$ 



#### Application of PPCA to Generating Face Images



Fig. 2.2 Face images of subject 20 under 10 different illumination conditions in the extended Yale B data set. All images are frontal faces cropped to size  $192 \times 168$ .

#### Application of PPCA to Generating Face Images

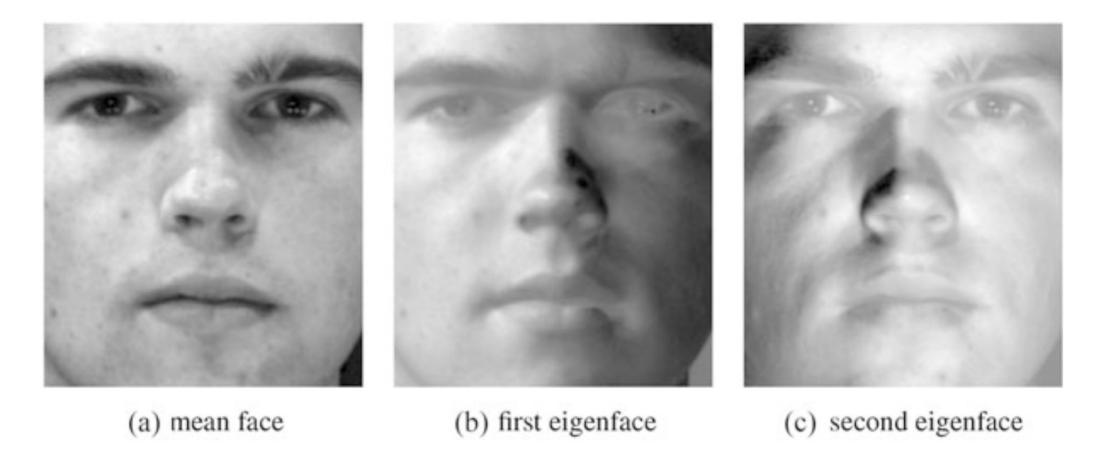


Fig. 2.5 Mean face and the first two eigenfaces by applying PPCA to the ten images in Figure 2.2.

### Application of PPCA to Generating Face Images

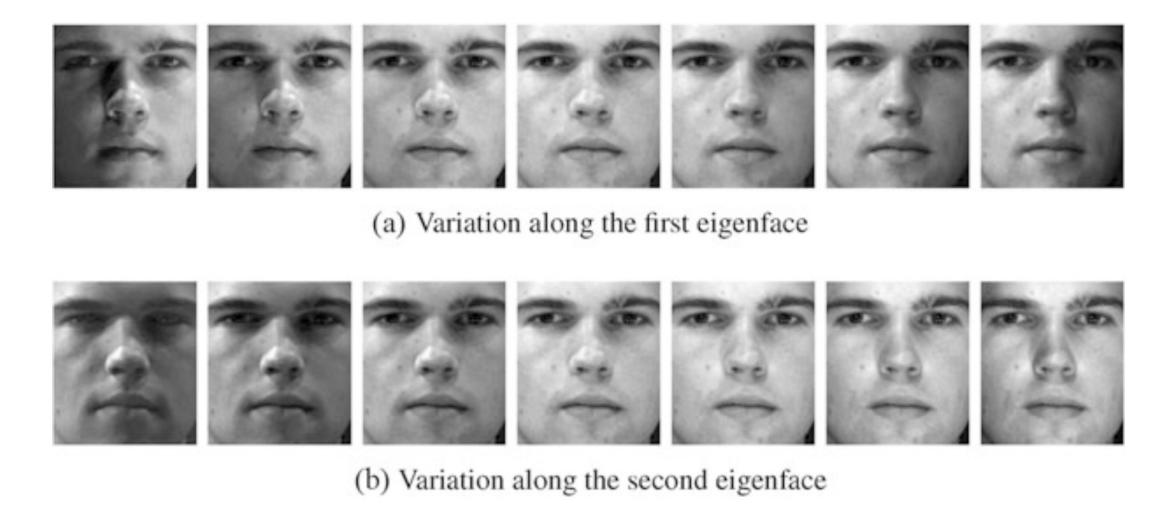


Fig. 2.6 Variation of the face images along the two eigenfaces given by PPCA. Each row plots  $\mu + y_i u_i$  for  $y_i = -1$ :  $\frac{1}{3}$ : 1, i = 1, 2.