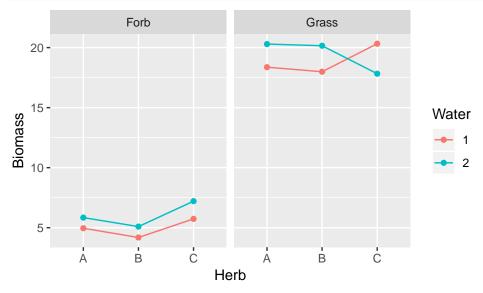
## Assignment8 KEY

40 points total, 2 points per problem part unless otherwise noted.

1. Summary Graph (4pts)



2. Three-way ANOVA table

##

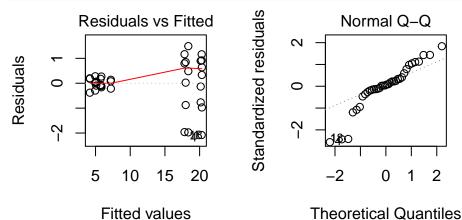
```
options(contrasts=c("contr.sum","contr.poly"))
Model1 <- lm(Biomass ~ Type*Herb*Water, data = InData)
Anova(Model1, type = 3)
## Anova Table (Type III tests)</pre>
```

```
## Response: Biomass
##
                   Sum Sq Df
                               F value
                                          Pr(>F)
                   5473.8
## (Intercept)
                          1 5539.1323 < 2.2e-16 ***
                           1 1698.7721 < 2.2e-16 ***
## Type
                   1678.7
## Herb
                      5.1
                           2
                                2.5644
                                        0.097872
## Water
                      5.9
                                5.9711
                                        0.022271 *
                           1
## Type:Herb
                           2
                                        0.086259 .
                      5.4
                                2.7185
## Type:Water
                      0.7
                           1
                                0.7092
                                        0.408021
## Herb:Water
                      7.9
                           2
                                3.9973
                                        0.031742 *
## Type:Herb:Water
                     13.1
                           2
                                6.6440
                                        0.005055 **
## Residuals
                     23.7 24
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

## 3. Three-way Diagnostic plots

Note: Diagnostic plots not required for full credit, but shown here for completeness. Based on the plot of Resids vs Fitted, there is strong evidence of unequal variance. (Based on the QQplot, there is also some evidence against normality. But this is actually driven by the unequal variance.)

```
par(mfrow=c(1,2))
plot(Model1, which = c(1:2))
```



## 4. Three-way emmeans

```
emout1 <- emmeans(Model1, pairwise ~ Water | Herb*Type)
emout1$contrasts</pre>
```

```
## Herb = A, Type = Forb:
##
   contrast
               estimate
                               SE df t.ratio p.value
##
             -0.8848361 0.8116673 24 -1.090 0.2865
##
## Herb = B, Type = Forb:
##
   contrast
               estimate
                               SE df t.ratio p.value
             -0.9084060 0.8116673 24 -1.119 0.2741
##
##
## Herb = C, Type = Forb:
##
   contrast
               estimate
                               SE df t.ratio p.value
             -1.4730603 0.8116673 24 -1.815 0.0821
##
##
## Herb = A, Type = Grass:
   contrast
               estimate
                               SE df t.ratio p.value
```

```
-1.9244287 0.8116673 24 -2.371 0.0261
##
## Herb = B, Type = Grass:
    contrast
               estimate
                                SE df t.ratio p.value
##
             -2.1667163 0.8116673 24 -2.669 0.0134
##
## Herb = C, Type = Grass:
    contrast
               estimate
                                SE df t.ratio p.value
              2.4991714 0.8116673 24
                                         3.079 0.0051
  5. FORB Two-way ANOVA table
Model2 <- lm(Biomass ~ Herb*Water, data = InData[InData$Type == "Forb",])
Anova(Model2, type =3)
## Anova Table (Type III tests)
##
## Response: Biomass
##
               Sum Sq Df
                             F value
                                         Pr(>F)
## (Intercept) 544.92 1 10432.1562 < 2.2e-16 ***
## Herb
                 10.17
                       2
                             97.3895 3.820e-08 ***
                 5.33 1
                            102.1229 3.197e-07 ***
## Water
## Herb:Water
                  0.33 2
                              3.1846
                                        0.07772 .
## Residuals
                 0.63 12
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
  6. FORB diagnostic plots look much better! Some evidence of unequal variance, but not severe.
par(mfrow = c(1,2))
plot(Model2, which = c(1:2))
                                    Standardized residuals
                                                 Normal Q-Q
          Residuals vs Fitted
             210
                    0
Residuals
                            8
     0.0
                                          0
                0
             190
                 5.5
                      6.5
                                              -2
                                                        0
                                                            1
                                                                 2
            4.5
             Fitted values
                                             Theoretical Quantiles
  7. FORB emmeans #1 (interaction comparisons)
emout2 <- emmeans(Model2, pairwise ~ Water | Herb)</pre>
emout2
## $emmeans
## Herb = A:
   Water
                          SE df lower.CL upper.CL
            emmean
##
          4.955550 0.131953 12 4.668049 5.243051
##
          5.840386 0.131953 12 5.552885 6.127887
##
```

```
## Herb = B:
## Water emmean
                        SE df lower.CL upper.CL
         4.183644 0.131953 12 3.896143 4.471145
         5.092050 0.131953 12 4.804549 5.379551
##
## Herb = C:
                        SE df lower.CL upper.CL
## Water emmean
         5.734037 0.131953 12 5.446536 6.021538
## 1
##
         7.207097 0.131953 12 6.919597 7.494598
##
## Confidence level used: 0.95
##
## $contrasts
## Herb = A:
## contrast
             estimate
                              SE df t.ratio p.value
##
            -0.8848361 0.1866097 12 -4.742 0.0005
##
## Herb = B:
## contrast estimate
                              SE df t.ratio p.value
## 1 - 2 -0.9084060 0.1866097 12 -4.868 0.0004
##
## Herb = C:
## contrast estimate
                              SE df t.ratio p.value
          -1.4730603 0.1866097 12 -7.894 <.0001
  8. FORB LSD #1 (interaction comparisons)
n = 3 #for each treatment combination mean
qt(0.975, df = 12)*sqrt(2*(0.63/12)/3)
## [1] 0.4076186
qt(0.975, df = 12)*0.187
## [1] 0.407438
#or peeling MSResid out of ANOVA table with R
#numerator is SSResid, denominator is df resid from ANOVA table
MSResid=(Anova(Model2, type = 3)[[1]][5])/(Anova(Model2, type = 3)[[2]][5])
MSResid
## [1] 0.05223477
qt(.975,12)*sqrt(2*MSResid/n)
## [1] 0.4065876
  9. FORB emmeans #2 (main effect comparison)
emout3 <- emmeans(Model2, pairwise ~ Water)</pre>
## NOTE: Results may be misleading due to involvement in interactions
emout3$contrasts
## contrast estimate
                             SE df t.ratio p.value
## 1 - 2 -1.088768 0.1077392 12 -10.106 <.0001
## Results are averaged over the levels of: Herb
```

```
10. FORB LSD #2 (main effect comparison)
```

```
n=9 #because we are comparing Water across 3 levels of Herb (each with 3)
qt(0.975, df = 12)*sqrt(2*(0.63/12)/9)

## [1] 0.2353387
qt(0.975, df = 12)*0.108

## [1] 0.2353118
#or this method using same MSResid from ANOVA
qt(.975,12)*sqrt(2*MSResid/n)
```

## [1] 0.2347435

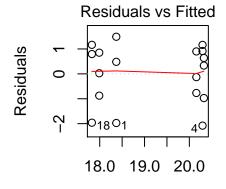
- 11. The power is higher for the main effect comparison (#10) because the LSD (ME) is smaller.
- 12. GRASS Two-way ANOVA table

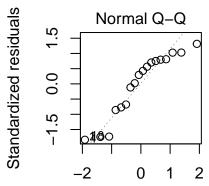
```
Model3 <- lm(Biomass ~ Herb*Water, data = InData[InData$Type == "Grass",])
Anova(Model3, type =3)</pre>
```

```
## Anova Table (Type III tests)
##
## Response: Biomass
##
               Sum Sq Df
                                      Pr(>F)
                           F value
## (Intercept) 6607.6
                      1 3433.9972 4.029e-16 ***
## Herb
                 0.3 2
                                      0.9333
                            0.0694
## Water
                 1.3 1
                            0.6586
                                      0.4329
## Herb:Water
                 20.7 2
                            5.3786
                                      0.0215 *
## Residuals
                 23.1 12
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

13. GRASS diagnostic plots look much better! Some evidence of skew (non-normality), but not severe.

```
par(mfrow = c(1,2))
plot(Model3, which = c(1:2))
```





Fitted values

Theoretical Quantiles

14. GRASS emmeans (interaction comparisons)

```
emout4 <- emmeans(Model3, pairwise ~ Water | Herb)
emout4$contrasts</pre>
```

## Herb = A:

```
estimate
                             SE df t.ratio p.value
   contrast
##
   1 - 2
             -1.924429 1.132601 12 -1.699 0.1150
##
## Herb = B:
##
   contrast
              estimate
                             SE df t.ratio p.value
   1 - 2
             -2.166716 1.132601 12 -1.913 0.0799
##
## Herb = C:
##
   contrast
              estimate
                             SE df t.ratio p.value
   1 - 2
              2.499171 1.132601 12
                                     2.207 0.0476
```

- 15. No, because there is evidence of an interaction between Water and Herb. In particular, the estimated difference in mean response for Water 1 vs 2 changes sign when comparing across Herb. When there is evidence of an interaction, it does not make sense to look at main effects.
- 16. (4pts) The estimated differences are the SAME. The SEs are DIFFERENT (Combined SE = 0.812, Forb SE = 0.187, Grass SE = 1.13). The df is DIFFERENT (combined df = 24, split df = 12).
- 17. (1) The combined analysis showed strong evidence of unequal variance. The model assumptions are better satisfied using the split analyses. (2) The combined analysis showed evidence of a three-way interaction.
- 18. (1) When we split the analysis we reduce the df (hence reducing power). (2) When we split the analysis we cannot get direct comparisons of Forb versus Grass.