

## Computer Lab 1

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You are recommended to use R for solving the labs.

You work and submit your labs in pairs, but both of you should contribute equally and understand all parts of your solutions.

**It is not allowed to share exact solutions** with other student pairs.

The submitted lab reports will be verified through URKUND and indications of plagiarism will be investigated by the Disciplinary Board.

Submit your solutions via LISAM, no later than April 18 at 23:30.

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### 1. Bernoulli ... again.

Let  $y_1, \dots, y_n | \theta \sim \text{Bern}(\theta)$ , and assume that you have obtained a sample with  $s = 5$  successes in  $n = 20$  trials. Assume a  $\text{Beta}(\alpha_0, \beta_0)$  prior for  $\theta$  and let  $\alpha_0 = \beta_0 = 2$ .

- (a) Draw random numbers from the posterior  $\theta | y \sim \text{Beta}(\alpha_0 + s, \beta_0 + f)$ ,  $y = (y_1, \dots, y_n)$ , and verify graphically that the posterior mean and standard deviation converges to the true values as the number of random draws grows large.
- (b) Use simulation (`nDraws = 10000`) to compute the posterior probability  $\Pr(\theta > 0.3 | y)$  and compare with the exact value [Hint: `pbeta()`].
- (c) Compute the posterior distribution of the log-odds  $\phi = \log \frac{\theta}{1-\theta}$  by simulation (`nDraws = 10000`). [Hint: `hist()` and `density()` might come in handy]

### 2. Log-normal distribution and the Gini coefficient.

Assume that you have asked 10 randomly selected persons about their monthly income (in thousands Swedish Krona) and obtained the following ten observations: 44, 25, 45, 52, 30, 63, 19, 50, 34 and 67. A common model for non-negative continuous variables is the log-normal distribution. The log-normal distribution  $\log \mathcal{N}(\mu, \sigma^2)$  has density function

$$p(y | \mu, \sigma^2) = \frac{1}{y \cdot \sqrt{2\pi\sigma^2}} \exp \left[ -\frac{1}{2\sigma^2} (\log y - \mu)^2 \right],$$

for  $y > 0$ ,  $\mu > 0$  and  $\sigma^2 > 0$ . The log-normal distribution is related to the normal distribution as follows: if  $y \sim \log \mathcal{N}(\mu, \sigma^2)$  then  $\log y \sim \mathcal{N}(\mu, \sigma^2)$ . Let  $y_1, \dots, y_n | \mu, \sigma^2 \stackrel{iid}{\sim} \log \mathcal{N}(\mu, \sigma^2)$ , where  $\mu = 3.7$  is assumed to be known but  $\sigma^2$  is unknown with non-informative prior  $p(\sigma^2) \propto 1/\sigma^2$ . The posterior for  $\sigma^2$  is the  $Inv - \chi^2(n, \tau^2)$  distribution, where

$$\tau^2 = \frac{\sum_{i=1}^n (\log y_i - \mu)^2}{n}.$$

- (a) Simulate 10,000 draws from the posterior of  $\sigma^2$  (assuming  $\mu = 3.7$ ) and compare with the theoretical  $Inv - \chi^2(n, \tau^2)$  posterior distribution.
  - (b) The most common measure of income inequality is the Gini coefficient,  $G$ , where  $0 \leq G \leq 1$ .  $G = 0$  means a completely equal income distribution, whereas  $G = 1$  means complete income inequality. See Wikipedia for more information. It can be shown that  $G = 2\Phi(\sigma/\sqrt{2}) - 1$  when incomes follow a  $\log \mathcal{N}(\mu, \sigma^2)$  distribution.  $\Phi(z)$  is the cumulative distribution function (CDF) for the standard normal distribution with mean zero and unit variance. Use the posterior draws in a) to compute the posterior distribution of the Gini coefficient  $G$  for the current data set.
  - (c) Use the posterior draws from b) to compute a 90% equal tail credible interval for  $G$ . A 90% equal tail interval  $(a, b)$  cuts off 5% percent of the posterior probability mass to the left of  $a$ , and 5% to the right of  $b$ . Also, do a kernel density estimate of the posterior of  $G$  using the `density` function in R with default settings, and use that kernel density estimate to compute a 90% Highest Posterior Density interval for  $G$ . Compare the two intervals.
3. *Bayesian inference for the concentration parameter in the von Mises distribution.*  
 This exercise is concerned with directional data. The point is to show you that the posterior distribution for somewhat weird models can be obtained by plotting it over a grid of values. The data points are observed wind directions at a given location on ten different days. The data are recorded in degrees:

$$(40, 303, 326, 285, 296, 314, 20, 308, 299, 296),$$

where North is located at zero degrees (see Figure 1 on the next page, where the angles are measured clockwise). To fit with Wikipedias description of probability distributions for circular data we convert the data into radians  $-\pi \leq y \leq \pi$ . The 10 observations in radians are

$$(-2.44, 2.14, 2.54, 1.83, 2.02, 2.33, -2.79, 2.23, 2.07, 2.02).$$

Assume that these data points are independent observations following the von Mises distribution

$$p(y|\mu, \kappa) = \frac{\exp[\kappa \cdot \cos(y - \mu)]}{2\pi I_0(\kappa)}, \quad -\pi \leq y \leq \pi,$$

where  $I_0(\kappa)$  is the modified Bessel function of the first kind of order zero [see `?besselI` in R]. The parameter  $\mu$  ( $-\pi \leq \mu \leq \pi$ ) is the mean direction and  $\kappa > 0$  is called the concentration parameter. Large  $\kappa$  gives a small variance around  $\mu$ , and vice versa. Assume that  $\mu$  is known to be 2.39. Let  $\kappa \sim \text{Exponential}(\lambda = 1)$  a priori, where  $\lambda$  is the rate parameter of the exponential distribution (so that the mean is  $1/\lambda$ ).

- (a) Plot the posterior distribution of  $\kappa$  for the wind direction data over a fine grid of  $\kappa$  values.
- (b) Find the (approximate) posterior mode of  $\kappa$  from the information in a).

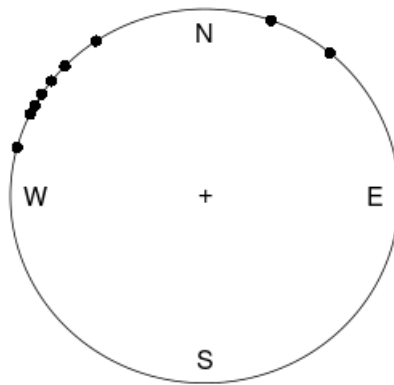


Figure 1: The wind direction data. Angles are measured clock-wise starting from North.

MAY BAYES BE WITH YOU!