

Neyman-Pearson Lemma: only for simple-vs-simple tests.

Finish Example from last time:

$$L_0 = \begin{cases} 0.25 & x=0 \\ 0.5 & x=1 \\ 0.25 & x=2 \end{cases}$$

$$L_1 = \begin{cases} 0.0625 & x=0 \\ 0.375 & x=1 \\ 0.5625 & x=2 \end{cases}$$

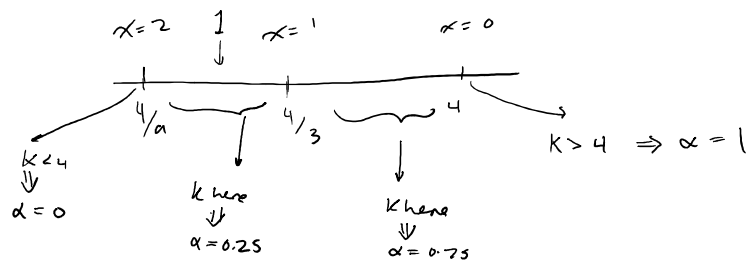
$$\frac{L_0}{L_1} = \begin{cases} 4 & x=0 \\ \frac{4}{3} & x=1 \\ \frac{4}{3} & x=2 \end{cases}$$

find K s.t. $\frac{L_0}{L_1} \leq K$ inside C w/size $\leq \alpha$.

for $\alpha = 0.25$, $K = 1$ works.

for $\alpha = 0.5$, $K = 1$ works.

for $\alpha = 0.05$ there is no real good K .



Remark: For a Discrete distribution, the possible levels ^{α} of the test are not completely under our control (e.g. no test can have $\alpha = 0.5$ using only likelihood ratio). Can use randomness to remedy this.

Ex:^(NPL) Let $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Exp}(\theta)$. Hyp: $H_0: \theta = \theta_0$, $H_1: \theta = \theta_1$ or $\theta_1 > \theta_0$. Use NPL.

$$\frac{L_0}{L_1} = \frac{\prod_{i=1}^n \frac{1}{\theta_0} \exp(-\frac{x_i}{\theta_0})}{\prod_{i=1}^n \frac{1}{\theta_1} \exp(-\frac{x_i}{\theta_1})} = \left(\frac{\theta_1}{\theta_0}\right)^n \exp\left(n\left(\frac{1}{\theta_1} \sum_{i=1}^n x_i - \frac{1}{\theta_0} \sum_{i=1}^n x_i\right)\right) \leq K.$$

$$n \log\left(\frac{\theta_1}{\theta_0}\right) + n(\frac{1}{\theta_1} - \frac{1}{\theta_0}) \sum_{i=1}^n x_i \leq \log(\kappa)$$

See ex 7.14 p 221.

$$\Rightarrow \sum_{i=1}^n x_i \geq \tilde{K} \quad \text{want} \quad \mathbb{P}(\sum_{i=1}^n x_i \geq \tilde{K}; \theta = \theta_0) = \alpha. \quad \text{note } \sum x_i \sim \text{Gamma}(n, \theta_0) \quad \text{Use a computer to figure this } \tilde{K} \text{ out.}$$

Remarks: Consider a test (i.e. critical region C)

When we reject H_0 (i.e. sample in C)

① $\mathbb{P}(H_1 \text{ true}) = 1 - \beta$? NO. $\mathbb{P}(H_0 \text{ false}) = 1 - \beta$ but $H_0 \text{ false} \not\Rightarrow H_1 \text{ true}$.

② $\mathbb{P}(H_0 \text{ true}) = \alpha$? NO, don't know $\mathbb{P}(H_0 \text{ true} | \text{reject } H_0)$.

The test & result make no references to $\mathbb{P}(H_0 \text{ true})$ or $\mathbb{P}(H_1 \text{ true})$.

§12.5 power of a test.

So far, we've only done simple vs simple. in general, we might want simple vs composite (eg $H_0: \theta = \theta_0$ and $H_1: \theta > \theta_0$). to evaluate these tests, use "power function"

Def: Power function of a test of H_0 against H_1 is given by

$$\pi(\theta) = \begin{cases} \alpha(\theta) & \text{for value of } \theta \text{ under } H_0 \\ 1 - \beta(\theta) & \text{for value of } \theta \text{ under } H_1 \end{cases}$$

$$\pi(\theta) = \mathbb{P}(\text{reject } H_0; \theta).$$

Ex Wicking Example: we tested $H_0: p = 0.1$ vs $H_1: p = 0.4$.

Now consider test $H_0: p \leq 0.1$ vs $H_1: p > 0.1$.

\bar{X} = # successes in 20 trials $\sim \text{Bin}(20, p)$.

We reject H_0 if $X > 5$.

$$\pi(p) = \mathbb{P}(\text{reject } H_0; p) = \mathbb{P}(X > 5; p) = 1 - \mathbb{P}(X \leq 5; p)$$

$$= 1 - \sum_{i=0}^5 \mathbb{P}(X=i; p)$$

$$= 1 - \sum_{i=0}^s P(X=i|P)$$

P	0.05	0.10	0.15	...	0.50
$\pi(P)$	0.0003	0.0125	0.0693	...	0.9793
error type	type 1	type 1	no error	...	no error

