

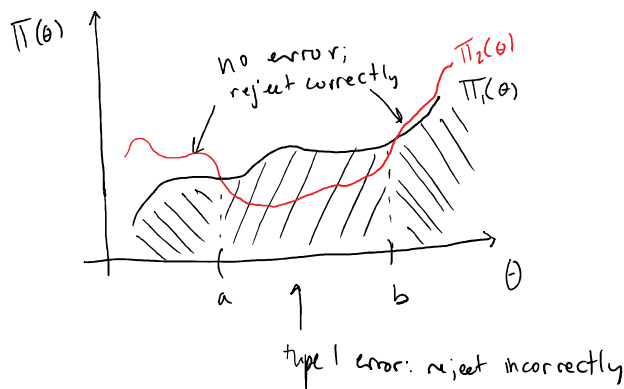
Power function  $\pi(p) = P(\text{reject } H_0; p)$

In simple vs. simple,  $\alpha, 1-\beta$  are only values for type I error rate and power.

In composite vs composite, there are multiple values seen in power curve.

Recall:  $\pi(\theta) = P(\text{reject } H_0; \theta) = \begin{cases} \alpha(\theta) & H_0 \text{ holds} \\ 1-\beta(\theta) & H_1 \text{ holds} \end{cases}$

e.g.  $H_0: \theta \in (a,b)$  vs  $H_1: \theta \notin (a,b)$



Power function is computed according to the "critical region" of the test.

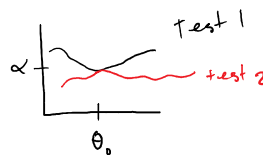
Test 2 is a more powerful & better test.

Comparison of tests via power functions.

Ex:  $H_0: \theta = \theta_0$ ,  $H_1: \theta \neq \theta_0$

\* Test 1 is uniformly

more powerful than 2. So just compare powers, test 1 is better.



both test 1 & test 2 <sup>both of size  $\alpha$</sup>  have same rate of type I error.

Def 12.6 there is a Uniformly Most Powerful (UMP) test.

Section 12.6 Likelihood Ratio tests (LRT).

Recall that NP lemma gives most powerful test when  $H_0, H_1$  simple.

Not necessarily MP test when  $H_0$  or  $H_1$  composite.

let  $X_1, \dots, X_n$  be a RS of size  $n$  from a density  $f(x; \theta)$ ,

let  $\Omega$  = all possible values of  $\theta$ .

suppose want to test  $H_0: \theta \in \omega$ ,  $H_1: \theta \in \omega'$  where  $\omega \subseteq \Omega$  and  $\omega' = \Omega \setminus \omega$ .

Previously,  $|\omega| = |\omega'| = 1$  and looked at ratios of likelihoods, now use max val of Likelihoods.

$$\text{Let } \max L_0 = \max_{\theta \in \omega} \prod_{i=1}^n f(x_i; \theta) = \prod_{i=1}^n f(x_i; \hat{\theta}) \text{ where } \hat{\theta} \text{ is MLE of } \theta \text{ on } \omega$$

$$\max L = \max_{\theta \in \Omega} \prod_{i=1}^n f(x_i; \theta) = \prod_{i=1}^n f(x_i; \hat{\theta}) \text{ where } \hat{\theta} \text{ is on } \Omega$$

$$\Lambda = \frac{\max L_0}{\max L} = \text{likelihood ratio statistic (LRS)}, \quad 0 \leq \Lambda \leq 1.$$

$$\text{under } H_1, \max L_0 \approx \max L \Rightarrow \Lambda \approx 1$$

$$\text{under } H_0, \max L_0 \ll \max L \Rightarrow \Lambda \text{ small}$$

test: reject  $H_0$  if  $\Lambda$  is too small (smaller than  $k \in (0, 1)$ ).

Question: why  $\Lambda$  instead of  $L_1$ ? because then  $k \in (0, 1)$ .

Q: how do we choose  $k$ ?

If  $H_0$  simple, pick so that CR is of size  $\alpha$

If  $H_0$  composite, pick so that type 1 error  $\leq \alpha$  for all  $\theta \in \omega$ , with

\*<sub>1</sub> equality at more  $\theta \in \omega$  to increase the power.

\*<sub>2</sub> equality at at least one  $\theta$  if possible

Note: for testing simple vs simple case  $H_0: \theta = \theta_0$ ,  $H_1: \theta = \theta_1$ ,

$$\max L_0 = L_0(\vec{x})$$

$$\max L = \max \begin{cases} L_0(\vec{x}) \\ L_1(\vec{x}) \end{cases}$$

$$\omega = \{\theta_0\} \quad \omega' = \{\theta_1\} \quad \Omega = \{\theta_0, \theta_1\}$$

So NP-test:  $\frac{L_0(\bar{x})}{L_1(\bar{x})} \leq K$ , LRT:  $\frac{L_0}{\max(L_0, L_1)} \leq \tilde{K}^{(0,1)}$

The two tests are not equivalent: consider  $K=2$ .

Ex:  $X_1, \dots, X_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$ .  $\sigma^2$  known. find CR of LRT for  $H_0: \mu = \mu_0$ ,  $H_1: \mu \neq \mu_0$ .

sol.  $\Omega = \{\mu_0\}$  so  $\hat{\mu} = \mu_0$ ,  $\Omega = \mathbb{R}$  so  $\hat{\mu} = \bar{X}$ .

$$\frac{\max L_0}{\max L} = \exp\left(\frac{1}{2\sigma^2} \sum_{i=1}^n (X_i - \bar{X})^2 - (X_i - \mu_0)^2\right) = \exp\left(-\frac{n}{2\sigma^2} (\bar{X} - \mu_0)^2\right)$$

LRT  $\Rightarrow$  reject  $H_0$  when  $\exp\left(-\frac{n}{2\sigma^2} (\bar{X} - \mu_0)^2\right) \leq K$

$$-\frac{n}{2\sigma^2} (\bar{X} - \mu_0)^2 \leq \log(K)$$

$$(\bar{X} - \mu_0)^2 \geq \frac{\log(K) 2\sigma^2}{-n}$$

$$\bar{X} \geq \underbrace{\sqrt{\frac{\log(K) 2\sigma^2}{-n}} + \mu_0}_{\tilde{K}}$$

$\tilde{K}$  determined s.t. CR has size  $\alpha$ .