

Likelihood Ratio test

$$\Lambda = \frac{\max L_0}{\max L}, \quad 0 \leq \Lambda \leq 1.$$

$$(1+x)^n \geq 1+nx$$

$$\frac{1+x}{n} \geq \frac{1+x}{n} \geq \sqrt[n]{1+x}$$

$$\Rightarrow 1+x \geq \sqrt[n]{n+nx}$$

$$\Rightarrow (1+x)^n \geq n+nx \geq 1+nx$$

test: reject  $H_0$  if  $\Lambda \leq k$  for some  $k \in (0, 1)$ .

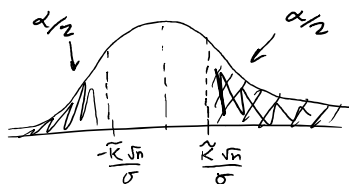
↳ simplify ratio, get equivalent eqn.

Q: how to find  $\tilde{k}$ ? (refer back to last time).

$\bar{X} \sim N(\mu_0, \frac{\sigma^2}{n})$  since under  $H_0$ ,  $\mu = \mu_0$  and we're trying to find  $\alpha$ .

$$\Rightarrow P(|\bar{X} - \mu_0| > \tilde{k}) = \alpha.$$

$$\Rightarrow P\left(\left|\frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}}\right| > \frac{\tilde{k}}{\sigma/\sqrt{n}}\right) = \alpha$$



So reject  $H_0$  if

$$|\bar{X} - \mu_0| \geq \frac{z_{\alpha/2} \sigma}{\sqrt{n}} \quad (\text{two-sided test}).$$

$$z_{\alpha/2} = \frac{\tilde{k}/\sigma}{\sigma/\sqrt{n}}$$

$$\Rightarrow \tilde{k} = \frac{z_{\alpha/2} \sigma}{\sqrt{n}}$$

Q: In this example, RS is from a normal distribution. so we can get distribution of  $\bar{X}$ . What if we don't have a normal dist?

↳ Try to prove this theorem to test math skills.

Thm 12.2 (Wilk's Thm).

For large  $n$ , the distribution of  $-2\log(\Lambda)$  approaches  $\chi^2_1$  distribution under general conditions.

LRT Approximation: Reject  $H_0$  if  $-2\log(\Lambda) \geq \chi^2_{\alpha, 1}$

$$\text{Check with our example: } -2\log(\Lambda) = -2\left(\frac{n}{2\sigma^2}(\bar{X} - \mu_0)^2\right) = \frac{n}{\sigma^2}(\bar{X} - \mu_0)^2 = \left(\frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}}\right)^2$$

$$\text{and } (\chi^2_{1(0.1)})^2 = \chi^2.$$

$$(2\sigma^2(X-\mu_0)^2)^{-1/2} = \frac{1}{\sigma^2} (X-\mu_0)^2 = \left( \frac{\hat{\sigma}^2 - \mu_0}{\sigma/\sqrt{n}} \right)$$

and  $(N(0,1))^2 = \chi_1^2$ .

Problem 12.22: Let  $X_i$ 's be iid  $N(\mu, \sigma^2)$  Both  $\mu$  and  $\sigma^2$  unknown.

Find LRT of  $H_0: \mu = \mu_0$  vs  $H_1: \mu \neq \mu_0$  with level  $\alpha$ .

Sol.  $\Omega = \mathbb{R} \times (0, \infty)$ .  $\omega = \{(\mu, \sigma^2): \sigma^2 > 0\}$ .

(i) find MLEs:

under  $\Omega$ ,  $\hat{\mu} = \bar{X}$   $\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2$

under  $\omega$ ,  $\hat{\mu} = \mu_0$   $\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \mu_0)^2$

$$\Lambda = \frac{\max_{\omega} L_0}{\max_{\Omega} L} = \frac{\left(\frac{1}{\hat{\sigma}^2}\right)^{n/2} \exp\left\{-\frac{1}{2\hat{\sigma}^2} \sum_{i=1}^n (X_i - \mu_0)^2\right\}}{\left(\frac{1}{\hat{\sigma}^2}\right)^{n/2} \exp\left\{-\frac{1}{2\hat{\sigma}^2} \sum_{i=1}^n (X_i - \bar{X})^2\right\}} \leq K.$$

$$\left(\frac{\hat{\sigma}^2}{\hat{\sigma}^2}\right)^{-n/2} = \left(\frac{\sum_{i=1}^n (X_i - \mu_0)^2}{\sum_{i=1}^n (X_i - \bar{X})^2}\right)^{-n/2}$$

$$\begin{aligned} \sum_{i=1}^n (X_i - \mu_0)^2 &= \sum_{i=1}^n ((X_i - \bar{X}) + (\bar{X} - \mu_0))^2 = \sum_{i=1}^n (X_i - \bar{X})^2 + 2 \underbrace{\left(\sum_{i=1}^n (X_i - \bar{X})\right)}_0 \cdot (\bar{X} - \mu_0) + \sum_{i=1}^n (\bar{X} - \mu_0)^2 \\ &= \sum_{i=1}^n (X_i - \bar{X})^2 + n(\bar{X} - \mu_0)^2 \end{aligned}$$

And the  $\frac{\exp\{\dots\}}{\exp\{\dots\}}$  part is 1 since there is cancellation.

$$\text{So } \Lambda < K \Leftrightarrow \left[ \frac{\sum_{i=1}^n (X_i - \bar{X})^2 + n(\bar{X} - \mu_0)^2}{\sum_{i=1}^n (X_i - \bar{X})^2} \right]^{-n/2} \leq K.$$

$$\Leftrightarrow 1 + \frac{n(\bar{X} - \mu_0)^2}{\sum_{i=1}^n (X_i - \bar{X})^2} \geq K^{-2/n}$$

$$\Leftrightarrow \frac{n(\bar{X} - \mu_0)^2}{\sum (X_i - \bar{X})^2} \geq K^{-2/n} - 1$$

$$\Leftrightarrow \frac{(\bar{X} - \mu_0)^2}{\frac{1}{n} \left( \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 \right)} \geq (n-1) (K^{-2/n} - 1)$$

$$\Leftrightarrow \underbrace{\left| \frac{\bar{X} - \mu_0}{S/\sqrt{n}} \right|}_{t_{n-1}} \geq \underbrace{\sqrt{(n-1)(K^{-2/n} - 1)}}_{\tilde{K}}$$

(2) find  $\tilde{K}$  So  $\tilde{K}$  should be  $t_{n-1, \frac{\alpha}{2}}$ .

So reject  $H_0$  when  $|\bar{X} - \mu_0| \geq t_{n-1, \frac{\alpha}{2}} \frac{S}{\sqrt{n}}$