

$$T = \frac{\hat{\beta} - \beta}{\hat{\sigma}} \sqrt{\frac{(n-2) S_{xx}}{n}} \sim t_{n-2}$$

$$CI \text{ for } \beta: \left( \hat{\beta} \pm t_{\frac{\alpha}{2}, n-2} \hat{\sigma} \sqrt{\frac{n}{(n-2) S_{xx}}} \right)$$

Ex: construct a 95% CI for  $\beta$  given data from last time.

$$\text{Sol: } 3.471 \pm 2.306 \cdot 4.720 \cdot \sqrt{\frac{16}{8 \cdot 376}} = (2.84, 4.10)$$

### §14.5 Normal Correlation Analysis (Google: Gaussian Copula)

Last time we required  $w(y|x)$  is normal (and  $x$  is observed already)

Now we drop assumption  $x$  is fixed.  $(X, Y) \sim$  Bivariate normal dist:

$$f(x, y) = \frac{1}{2\pi \sigma_1 \sigma_2 \sqrt{1-\rho^2}} \exp \left[ \frac{-1}{2(1-\rho^2)} \left( \left( \frac{x-\mu_1}{\sigma_1} \right)^2 - 2\rho \left( \frac{x-\mu_1}{\sigma_1} \right) \left( \frac{y-\mu_2}{\sigma_2} \right) + \left( \frac{y-\mu_2}{\sigma_2} \right)^2 \right) \right]$$

(Note:  $X \sim N(\mu_1, \sigma_1^2)$ ,  $Y \sim N(\mu_2, \sigma_2^2)$ )

Given paired data set  $\{(x_i, y_i) : i=1, \dots, n\}$ ,

$$L = \prod_{i=1}^n f(x_i, y_i) = \left( \frac{1}{2\pi \sigma_1 \sigma_2 \sqrt{1-\rho^2}} \right)^n \exp \left\{ \frac{-1}{2(1-\rho^2)} \left[ \sum_{i=1}^n \left( \frac{x_i - \mu_1}{\sigma_1} \right)^2 - 2\rho \sum_{i=1}^n \left( \frac{x_i - \mu_1}{\sigma_1} \right) \left( \frac{y_i - \mu_2}{\sigma_2} \right) + \sum_{i=1}^n \left( \frac{y_i - \mu_2}{\sigma_2} \right)^2 \right] \right\}$$

$$\Rightarrow \begin{aligned} \hat{\mu}_1 &= \bar{x} & \hat{\sigma}_1 &= \sqrt{\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2} \\ \hat{\mu}_2 &= \bar{y} & \hat{\sigma}_2 &= \sqrt{\frac{1}{n} \sum_{i=1}^n (y_i - \bar{y})^2} \end{aligned}, \quad \hat{\rho} = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\hat{\sigma}_1 \hat{\sigma}_2} = \frac{S_{xy}}{\sqrt{S_{xx} S_{yy}}}$$

→ also called sample correlation coefficient, usually denoted  $r$ .

### Remarks:

①  $\rho$  measures strength of the "linear" relationship between  $X$  and  $Y$ .

We often are interested in tests concerning  $\rho$

i.e.  $H_0: \rho = 0$  vs.  $H_1: \rho \neq 0$ .

② When  $X=x$  is given, the conditional variance of  $Y$  has the following formula:

$$\sigma_{Y|X}^2 = \sigma_2^2(1-\rho^2) \quad (\text{Thm 6.9})$$

$\rightarrow Y$  completely determined by  $X$ .

thus,  $\rho = \pm 1 \Rightarrow \sigma_{Y|X}^2 = 0 \Rightarrow Y = \alpha + \beta X$  for some  $\alpha, \beta \in \mathbb{R}$ .

③  $\sigma_{Y|X}^2$  is a function of  $\sigma_2$  and  $\rho$

thus by invariance prop. of MLE,

$$\hat{\sigma}_{Y|X}^2 = \hat{\sigma}_2^2(1-\hat{\rho}^2)$$

$$\Rightarrow r^2 = \hat{\rho}^2 = \frac{\hat{\sigma}_2^2 - \hat{\sigma}_{Y|X}^2}{\hat{\sigma}_2^2} \leftarrow \begin{array}{l} \text{Cov. variance of } Y_s \text{ for fixed values of } x, \\ \text{measures total variance of } Y_s \end{array}$$

thus,  $100r^2$  is percentage of total variation of the  $Y_s$  that is accounted for by the relationship with  $X$ .

④ let  $R$  be the RV whose value is  $r$

one can show that  $\frac{1}{2} \log \left( \frac{1+R}{1-R} \right) \rightsquigarrow N \left( \frac{1}{2} \log \left( \frac{1+\rho}{1-\rho} \right), \frac{1}{n-3} \right)$

$$\Rightarrow Z = \frac{\sqrt{n-3}}{2} \left[ \log \left( \frac{1+R}{1-R} \right) - \log \left( \frac{1+\rho}{1-\rho} \right) \right] \sim N(0,1)$$

Which can be used as a test stat for hyp. test or CI for  $\rho$ .