

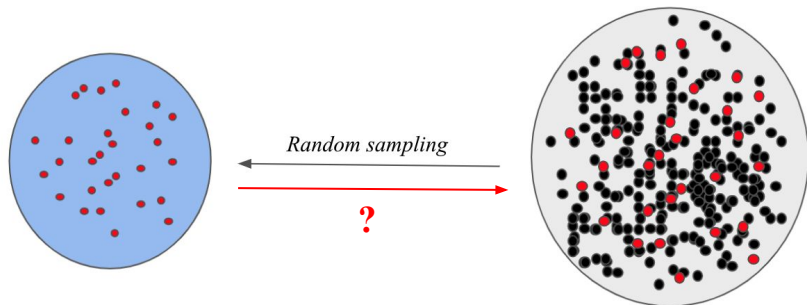
Midterm review

Point estimation & Hypothesis testing

07/15/2021

Population vs. samples

How to learn from samples about $f(x)$?



- **Point estimate:** a particular value $\hat{\theta}$ that best approximates the parameter of interest.
 - MLE; MM estimators;
 - Sufficiency; Estimation error quantification.
- **Interval estimate:** an interval $[\theta - a, \theta + b]$ that would contain the true parameter θ with a certain degree of confidence.
- **Hypothesis testing:** whether to reject a hypothesis.
 - $\theta > a$?
 - $\theta < a$?

Formulate the research problem

\downarrow Population $f(x)$

Make propositions about $f(x)$

\downarrow Parametric $f(x | \theta)$

Survey design & Data collection

\downarrow X_1, \dots, X_n

Parameter estimation of θ

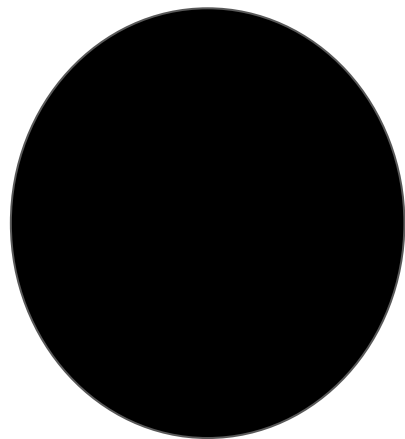
\downarrow Conclusion

- **Regression:** A special parametric model

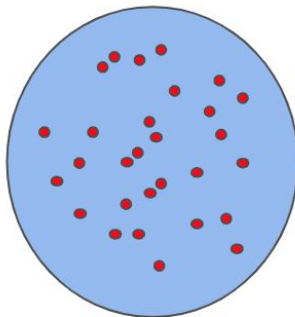
$$f(x | \beta, \sigma^2) = \underbrace{\mathbf{z}\beta}_{\text{Explanatory variables}} + N(0, \sigma^2)$$

IID assumption

If independently drawn, each sample X_i will be identically and independently distributed (i.i.d) under the true $f(x)$.

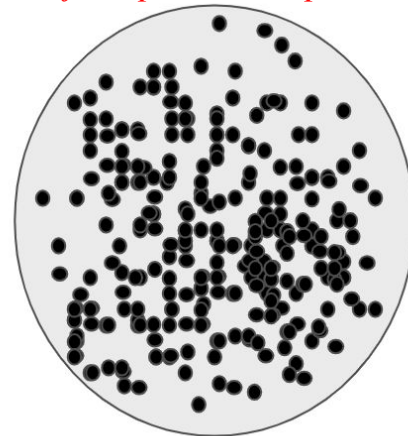


Uncountable

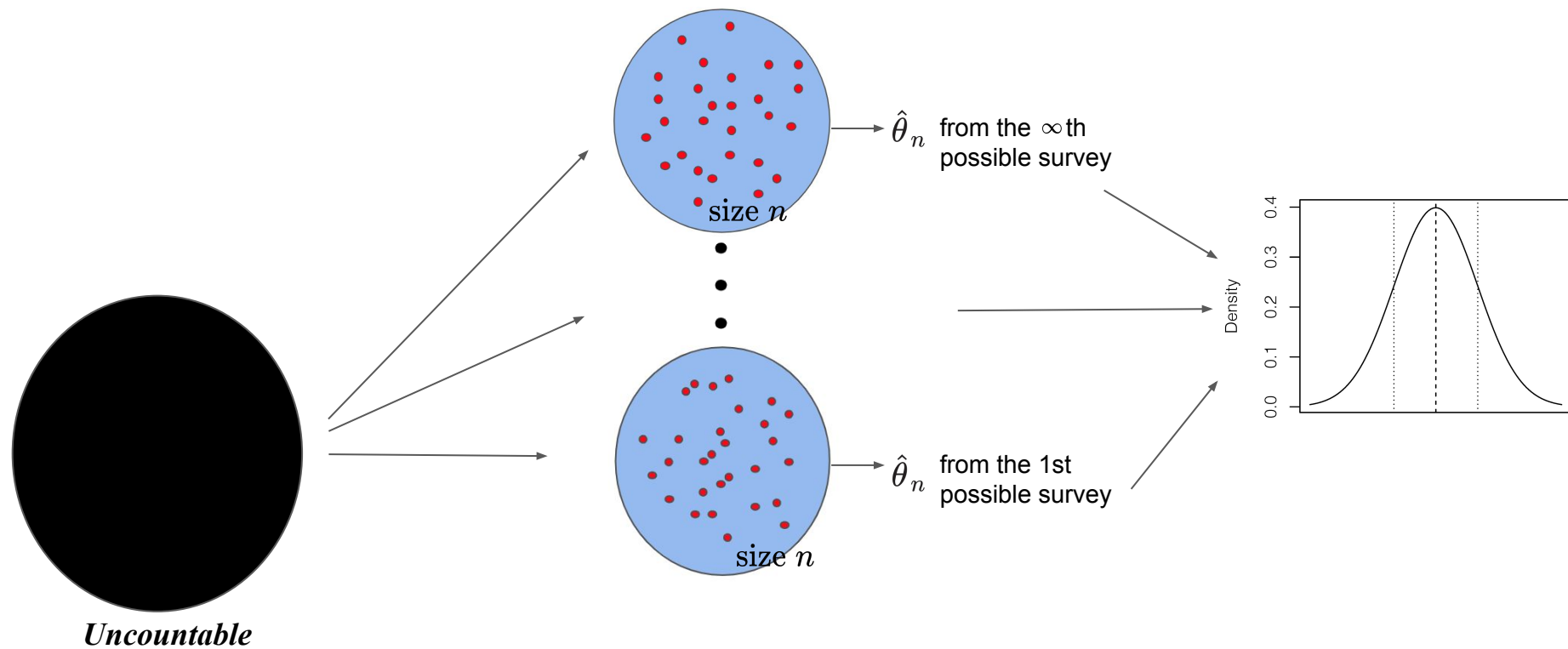


A sample of size n :
 X_1, \dots, X_n

For finite population, I.I.D can **only** be assumed if sampled **with** replacement.



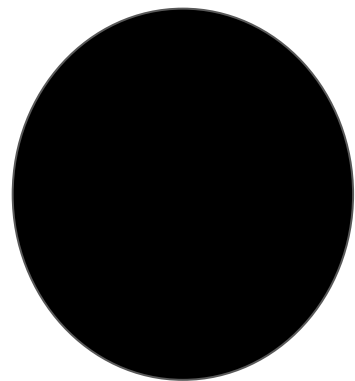
Sampling distribution



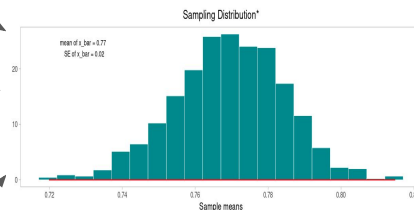
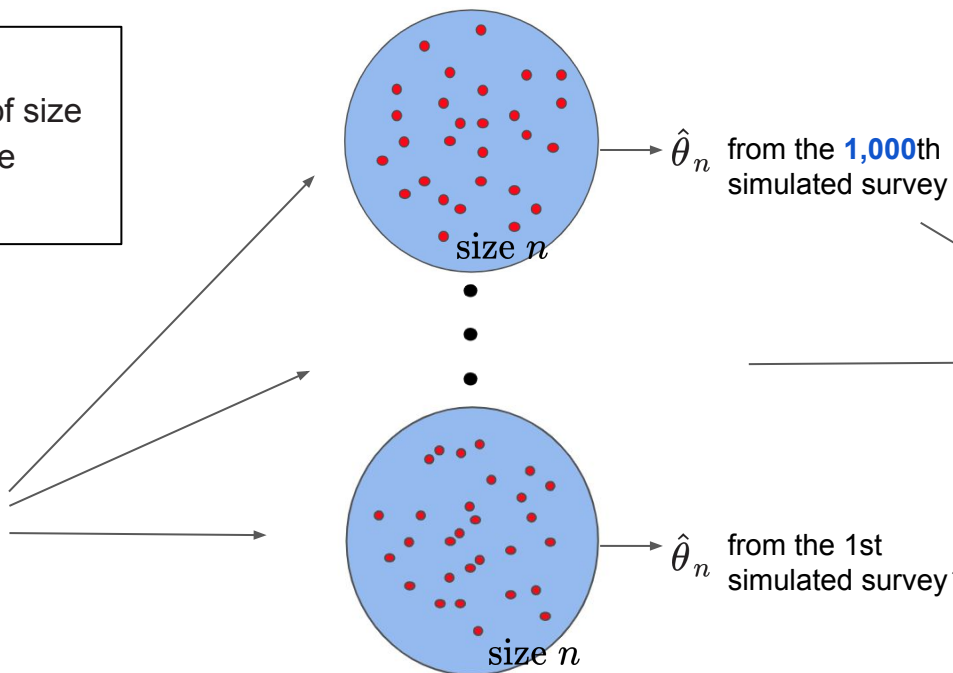
Bootstrapping

Simulation:

Draw as many samples of size n as possible, and plot the histogram of all $\hat{\theta}_n$.



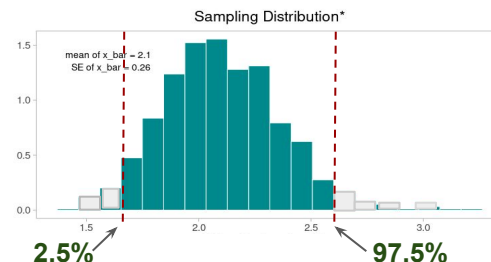
Uncountable



Sampling distribution

- \bar{X}_n and $\hat{\sigma}_n^2$ for population mean and variance;
- MM estimators;
- Maximum likelihood estimators.

1. *Exact sampling distribution;*
*e.g. MLE for $U(-\theta, \theta)$,
 \bar{X}_n for $\text{Gamma}(\alpha, \beta)$ and $N(\mu, \sigma^2)$
 $\hat{\sigma}_n^2$ for $N(\mu, \sigma^2)$*
2. *Asymptotic normality;*
 $\bar{X}_n, g(\bar{X}_n), \text{MLE}$
3. *Parametric or non-parametric bootstrapping*
(Page 10, Lecture 3 & Lab 3).

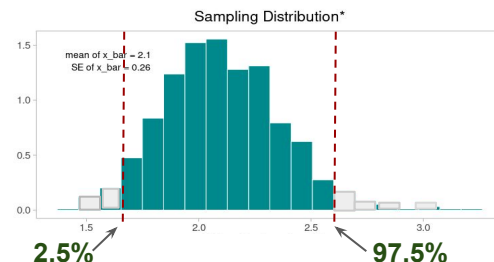


Confidence interval

$$\hat{\theta}_n \pm \text{Distr}_{\alpha/2} * \text{SE}(\hat{\theta}_n)$$

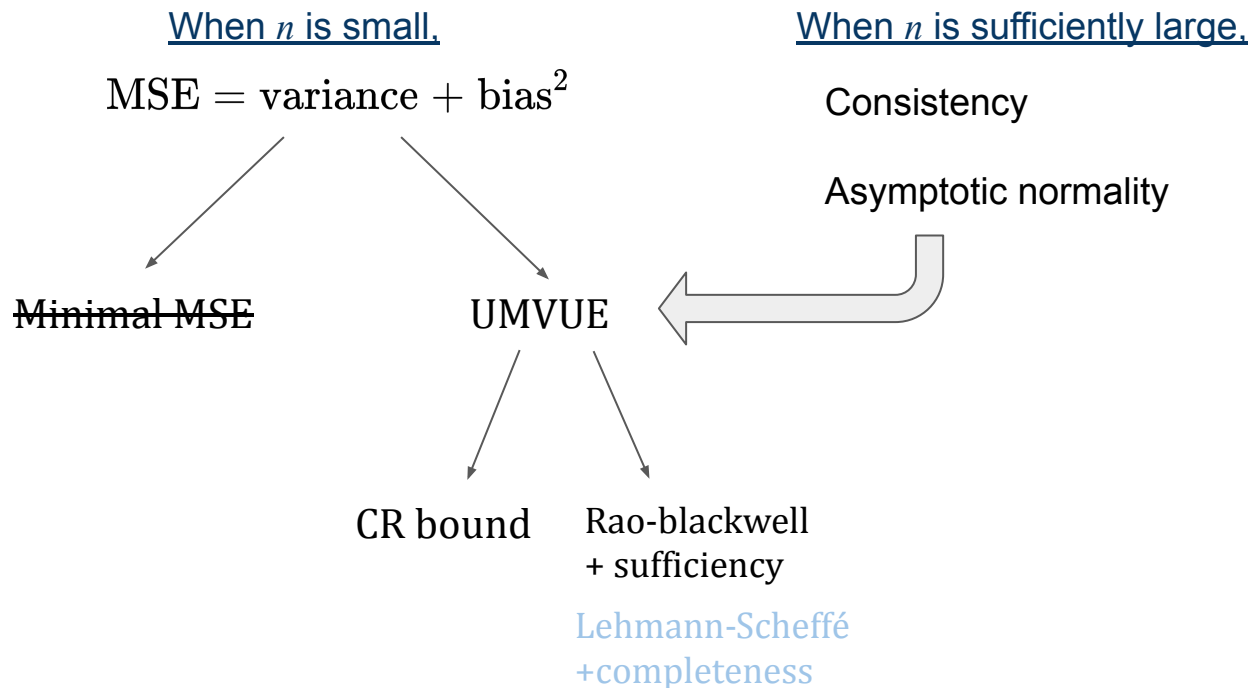
The confidence levels represents *theoretical long-run frequency* (i.e., the proportion) of confidence intervals that contain the true θ .

1. *Exact sampling distribution;*
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2. *Asymptotic normality;*
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Evaluating estimators

- \bar{X}_n and $\hat{\sigma}_n^2$
- MM
- MLE

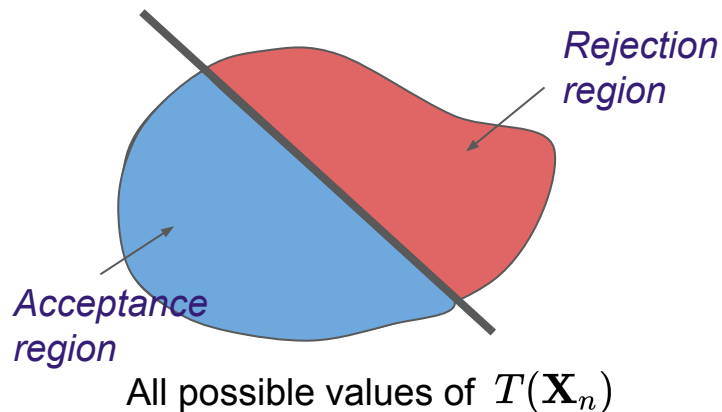


Hypothesis testing

$$H_0 : \theta \in \Theta_0 \quad \leftrightarrow \quad H_1 : \theta \in \Theta_1$$

$R = \{\text{Unlikely } T(\mathbf{X}_n) \text{ values under } H_0\}$

$$\xleftarrow{\text{formulate}} \text{LRT} \quad R = \{\lambda(\mathbf{X}_n) \leq c\}$$



Strategy: Maximizing power among tests with $\sup_{\theta \in \Theta_0} \beta(\theta) \leq \alpha$.

	H_0 is true	H_1 is true
Reject H_0	Type I error	Correct decision
Fail to reject H_0	Correct decision	Type II error

UMP tests

$$H_0 : \theta = \theta_0 \quad \leftrightarrow \quad H_1 : \theta = \theta_1 \quad (\theta_1 > \theta_0).$$



$$H_0 : \theta = \theta_0 \quad \leftrightarrow \quad H_1 : \theta > \theta_0.$$



Karlin-Rubin

$$H_0 : \theta \leq \theta_0 \quad \leftrightarrow \quad H_1 : \theta > \theta_0.$$

$$H_0 : \theta = \theta_0 \quad \leftrightarrow \quad H_1 : \theta = \theta_1 \quad (\theta_1 < \theta_0).$$



$$H_0 : \theta = \theta_0 \quad \leftrightarrow \quad H_1 : \theta < \theta_0.$$



Karlin-Rubin

$$H_0 : \theta \geq \theta_0 \quad \leftrightarrow \quad H_1 : \theta < \theta_0.$$

LRT is uniformly most powerful!

Perform HT

Perform hypothesis testing with real data:

- Specify significance level α and calculate the rejection region;
- Calculate p -value and compare it with α .

p -value. Under H_0 , the probability of observing a result *at least as extreme as* the test statistic.

Duality between CIs and HTs

$(1 - \alpha) \times 100\%$ exact CI for μ :

$$P \left(\bar{X}_n - \frac{S}{\sqrt{n}} t_{n-1}(\alpha/2) \leq \mu \leq \bar{X}_n + \frac{S}{\sqrt{n}} t_{n-1}(\alpha/2) \right) = 1 - \alpha$$

LRT with significance level α :

$$H_0 : \mu = \mu_0 \quad \text{versus} \quad H_1 : \mu \neq \mu_0. \quad \Longleftrightarrow \quad R = \left\{ \left| \frac{\sqrt{n}(\bar{X}_n - \mu_0)}{S} \right| \geq t_{n-1}(\alpha/2) \right\}$$