

Local Averaging Methods

Regression: $X_1, \dots, X_n \in [0, 1]^d$
 $Y_1, \dots, Y_n \in \mathbb{R}$

$$\text{mod } \hat{f}(x) = \sum_{i=1}^n w_i(x) Y_i \quad] \begin{matrix} \text{Local} \\ \text{average} \\ \text{of the output} \end{matrix}$$

where $\sum_{i=1}^n w_i(x) = 1$

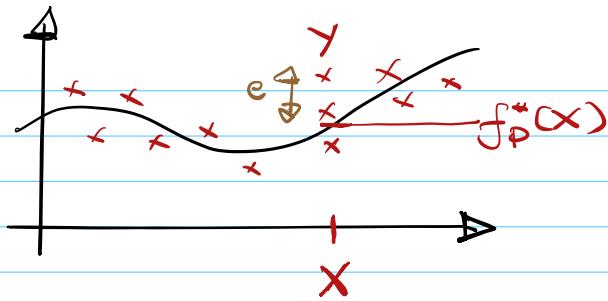
↳ if $\begin{cases} x \text{ close to } X_i : w_i(x) \text{ large} \\ \text{not close } w_i(x) \text{ small.} \end{cases}$

① The Regression problem

$$(X, Y) \sim P$$

$$R_p(f) = E[(f(x) - Y)^2]$$

Bayes estimator: $f_p^*(x) = E_p[Y | X=x]$.



Write $y = f_p^*(x) + e$

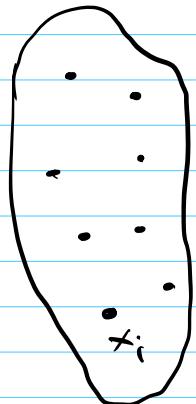
\downarrow
error term = $y - f_p^*(x)$

By definition of f_p^* :

$$E_p[e|x] = E[y|x] - f_p^*(x) = 0.$$

Example: x_1, \dots, x_n street in a city
 y_1, \dots, y_n CO₂ concentration

$\Rightarrow f_p^*(x) =$ Average CO₂ concentration
at street x .



$$y = f_p^*(x) + e$$

Noise magnitude may
depend on the street x .

→ Bayes Risk: $R_p^* = R_p(f_p^*)$

$$= E[(f_p^*(x) - y)^2]$$

$$= E[e^2].$$

What is the excess of risk of a function f ?

$$R_p(f) = \mathbb{E}_p[(f(x) - \gamma)^2]$$

$$\mathbb{E}[(f(x) - \gamma)^2 | x] = \mathbb{E}[(f(x) - f_p^*(x) - e)^2 | x]$$

$$= \underbrace{\mathbb{E}[(f(x) - f_p^*(x))^2 | x]} + e \underbrace{\mathbb{E}[(f(x) - f_p^*(x))e | x]}$$

$$= (f(x) - f_p^*(x))^2 + \mathbb{E}[e^2 | x]$$

$$= (f(x) - f_p^*(x)) \underbrace{\mathbb{E}[e | x]}_{=0} = 0$$

$$= (f(x) - f_p^*(x))^2 + \mathbb{E}[e^2 | x].$$

$$\Rightarrow R_p(f) = \mathbb{E}[\mathbb{E}[(f(x) - f_p^*(x))^2 | x]]$$

$$= \mathbb{E}[(f(x) - f_p^*(x))^2] + \underbrace{\mathbb{E}[\mathbb{E}[e^2 | x]]}_{= \mathbb{E}[e^2]}$$

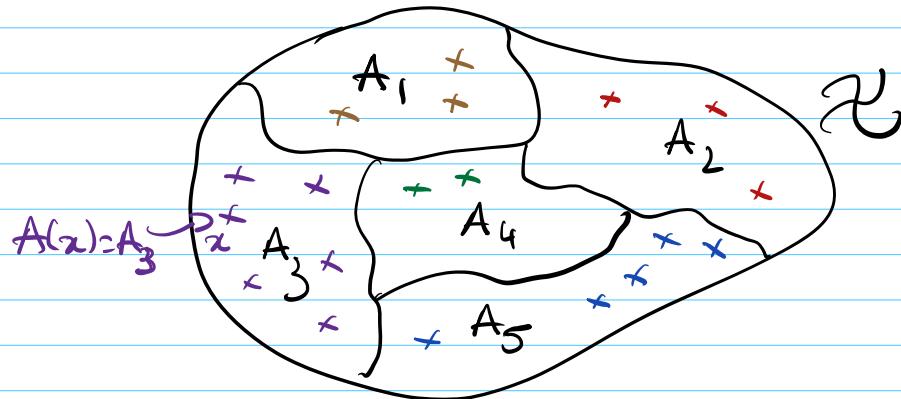
$$= R_p^*$$

\Rightarrow Excess of Risk:

$$R_p(f) - R_p^* = \mathbb{E}[(f(x) - f_p^*(x))^2]$$

$$= \int (f(x) - f_p^*(x))^2 dP_x(x)$$

e) Partition estimator:



Partition $A = (A_1, \dots, A_5)$ of a set X

$$\begin{cases} A_j \cap A_{j'} = \emptyset & j \neq j' \\ \bigcup_{j=1}^5 A_j = X \end{cases}$$

$X = [0, 1]^d$

Def: $(x_1, y_1), \dots, (x_n, y_n)$

$$x \in [0, 1]^d$$

$$A(x) = A_j \quad \text{if } x \in A_j.$$

Define

$$w_i(x) = \frac{\mathbb{1}\{x_i \in A(x)\}}{\sum_{i'=1}^n \mathbb{1}\{x_{i'} \in A(x)\}}$$

$$\hat{f}_A(x) = \sum_{i=1}^n w_i(x) y_i$$

Partition estimator
Regressogram

→ let us compute $\hat{f}_k(x)$:

A_1	x_1	A_2	A_3
•	•	•	•
•	•	•	•
•	•	•	•
•	•	•	•

$I_j = \text{indexes } i \text{ such that}$
 $x_i \in A_j$
 $n_j = \text{size of } I_j$.

$$n_1 = 1 \quad n_2 = 2 \quad n_3 = 0$$

$$I_2 = \{2, 4\}$$

→ let $x \in A_j$:

$$\begin{aligned} \hat{f}_k(x) &= \sum_{i=1}^n w_i(x) y_i \\ &= \frac{\sum_{i=1}^n \mathbb{1}\{x_i \in A_j\} y_i}{\sum_{i=1}^n \mathbb{1}\{x_i \in A_j\}}]_{n_j} \end{aligned}$$

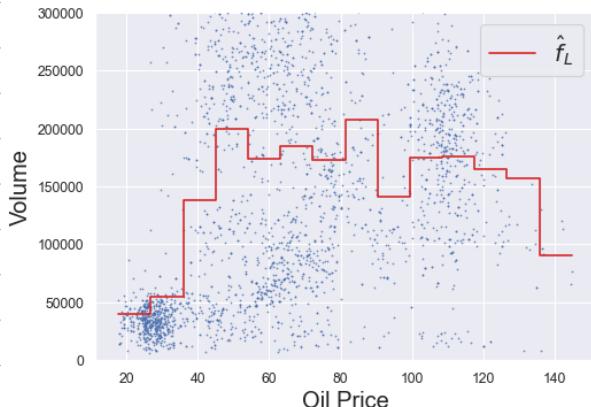
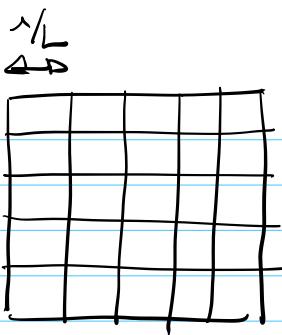
$$\left[\hat{f}_k(x) = \frac{1}{n_j} \sum_{i \in I_j} y_i \right]$$

⋮	x
⋮	A_j

↳ Average of the outputs y_i
 such that $x_i \in A_j$.

Example:

cube partition
 L^d cubes
of side length ϵ/L .



$$d=1$$

Is there a relation between price and volume of oil available on the market?

→ What is the excess of risk of the cube partition estimator?

We assume:

[1] f_p^* is α -Lipschitz:

$$\|f_p^*(x) - f_p^*(x')\| \leq \alpha \|x - x'\|$$

[2] f_p^* is bounded $|f_p^*(x)| \leq \beta$.

[3] Bounded Noise: $|e_i| \leq \sigma$.

THM: We have

$$\mathbb{E}[\mathcal{R}_p(\hat{f}_L) - \mathcal{R}_p^*] \lesssim \frac{\alpha^2}{L^2} + \frac{\sigma^2 L^d}{n}$$
$$L \approx n^{1/d+2} \Rightarrow \mathbb{E}[\mathcal{R}_p(\hat{f}_L) - \mathcal{R}_p^*] \lesssim n^{-\frac{2}{d+2}}$$

worse of dimensionality



→ The theorem gives the order of magnitude of the optimal L.
no Cross-validation.

Proof Sketch

$$x \in A_j$$

$$\begin{aligned} \hat{f}_L(x) - f_p^*(x) &= \frac{1}{n_j} \sum_{i \in I_j} y_i - f_p^*(x) \\ &\quad = f_p^*(x_i) + e_i \\ &= \frac{1}{n_j} \sum_{i \in I_j} (f_p^*(x_i) - f_p^*(x)) + \frac{1}{n_j} \sum_{i \in I_j} e_i \\ &\leq \alpha \|x_i - x\| \\ &\leq \alpha \frac{\sqrt{d}}{L} \end{aligned}$$

$$|\hat{f}_L(x) - f_p^*(x)| \leq \alpha \frac{\sqrt{d}}{L} + \left| \frac{1}{n_j} \sum_{i \in I_j} e_i \right|$$

Conditionally on I_j , this is an average
of n_j r.v.

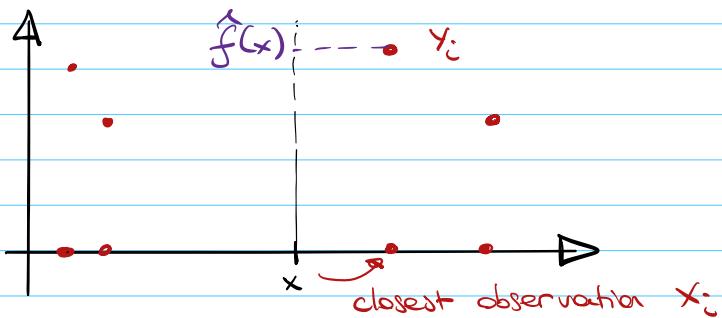
$$\rightarrow E \left[\left(\frac{1}{n_j} \sum_{i \in I_j} e_i \right)^2 \mid I_j \right] = \frac{1}{n_j^2} \sum_{i \in I_j} E[e_i^2 \mid I_j]$$
$$\leq \frac{\sigma^2}{n_j}$$

$n_j \approx n \times L^{-d}$

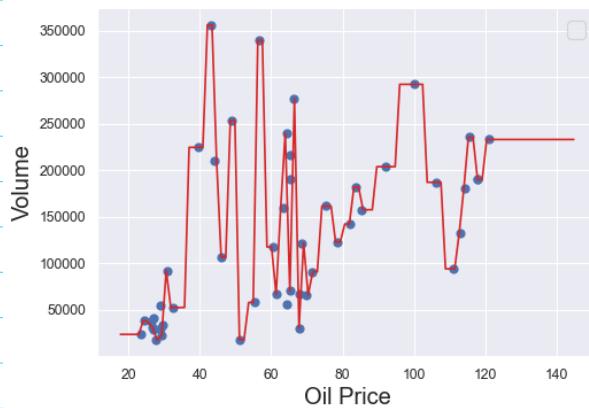
$$\Rightarrow E[\hat{f}_L(x) - f_p^*(x)]^2 \approx \frac{1}{L^d} + \frac{\sigma^2}{n L^d}$$

□

③ Nearest-Neighbor Methods :

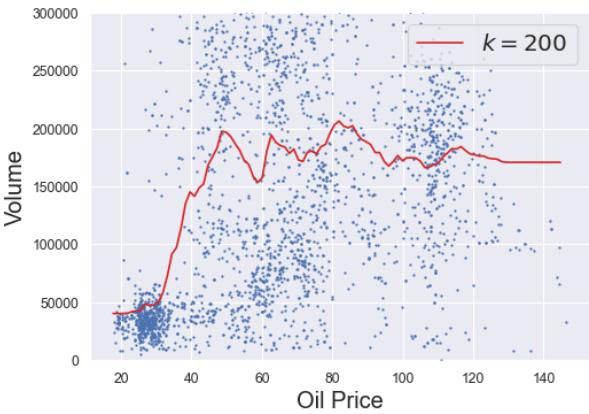


Prediction at x = output y_i of the nearest neighbor x_i .



1 Nearest Neighbor Estimator

Reasonable idea ... but overfitting.



k -Nearest-Neighbor



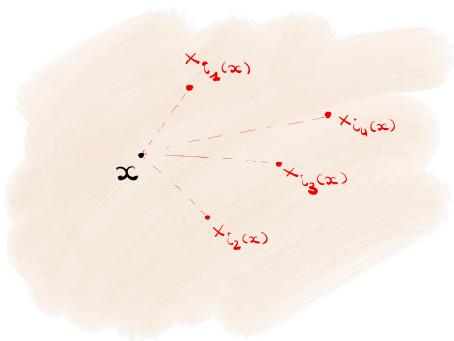
Prediction at x

= average of the
 k outputs y_i
 corresponding to the
 k NN of x .

Def: $(x_1, y_1) \dots (x_n, y_n)$ $k \geq 1$.
 $x \in [0, 1]^d$

so we order the inputs:

$$\|x - X_{i_1(x)}\| \leq \|x - X_{i_2(x)}\| \leq \dots \leq \|x - X_{i_k(x)}\|$$



$\Rightarrow I_k(x) = \{i_1(x), \dots, i_k(x)\}$ the k -Nearest
 -Neighbors

$$w_i(x) = \begin{cases} 1/k & \text{if } i \in I_k(x) \\ 0 & \text{otherwise.} \end{cases}$$

$$\hat{f}_k^{NN}(x) = \sum_{i=1}^n w_i(x) y_i$$

$$= \frac{1}{k} \sum_{i \in I_k(x)} y_i$$

THM: Assume :

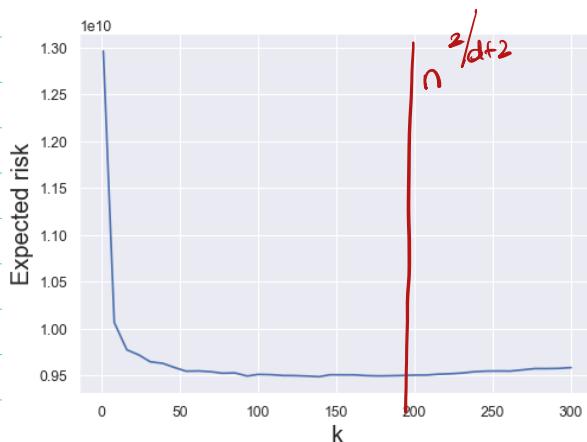
- ① f^* is α -Lipschitz
- ② Noise e is bounded: $|e| \leq \sigma$
- ③ X is uniform on $[0,1]^d$.

Then $E[R_p(\hat{f}_k^{NN}) - R_p^*] \leq \alpha^2 \left(\frac{k}{n}\right)^{2/d} + \frac{\sigma^2}{k}$.

For $k \approx n^{2/d+2}$: we obtain

$$E[R_p(\hat{f}_k^{NN}) - R_p^*] \approx n^{-\frac{2}{d+2}}$$

Curse of dimensionality



Select k with
cross-validation
in practice.

Why does this work?

→ { The k NN of x is close to x .
 f_p^* is Lipschitz }

$$\Rightarrow \|f_p^*(x) - f_p^*(X_{i_k(x)})\| \leq \alpha \|x - X_{i_k(x)}\|$$

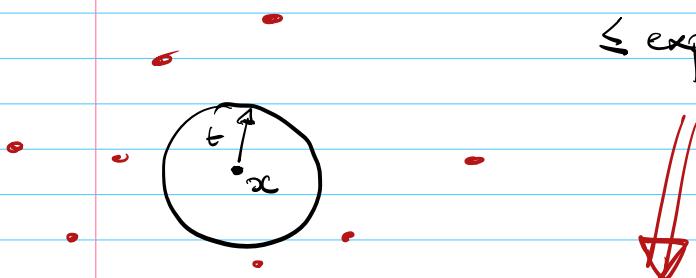
What is the distance between x and its k NN?

$\alpha = 1$

$$P(\|x - X_{i_1(x)}\| \geq t) = P(\text{No } \|x - X_i\| \geq t)$$

$$= (1 - P(B(x, t)))^n$$

$$\leq \exp(-n \underbrace{P(B(x, t))}_{\approx t^d})$$



$$\left[E[\|x - X_{i_1(x)}\|^2] \lesssim n^{-2/d} \right]$$

General fact: $E[z^2] = E\left[\int_0^{+\infty} \mathbb{1}\{t \leq z^2\} dt\right]$

$$= \int_0^{+\infty} P(z^2 \geq t) dt$$

$$= 2 \int_0^{+\infty} u P(z \geq u) du$$

Proof Sketch :

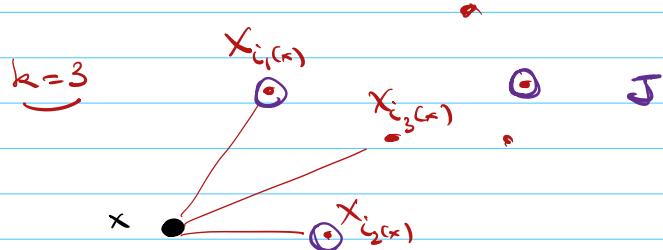
$$\left| \hat{f}_k^{NN}(x) - f_p^*(x) \right|^2 = \left| \frac{1}{k} \sum_{i \in I_k(x)} y_i - f_p^*(x) \right|^2$$

$$\lesssim \left(\frac{1}{k} \sum_{\{i \in I_k(x)\}} \underbrace{\left| f_p^+(x_i) - f_p^+(x) \right|}_{{\leq \alpha \|x - x_i\|}} \right)^2 + \left| \frac{1}{k} \sum_{i \in I_k(x)} e_i \right|^2$$

average of

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$$\leq \frac{\alpha^2}{k} \sum_{i \in \mathbb{A}_k(x)} \|x - X_i\|^2 \quad \Rightarrow \quad \mathcal{R} \geq \frac{\sigma^2}{k}.$$



If J is any set of k index:

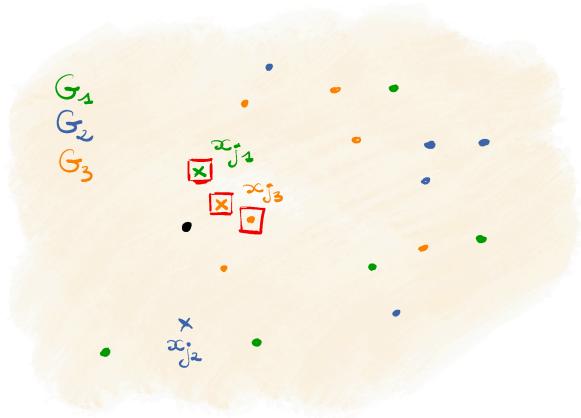
$$\frac{1}{k} \sum_{i \in I_k} \|x - x_i\|^2 \leq \frac{1}{k} \sum_{i \in J} \|x - x_i\|^2$$

→ We split the observations in k groups:

$j_1 = \text{index of}$
 closest to
 $x \text{ in } G_1$

$j_k = \text{index of}$
 $\text{closest to } x$
in G_k

$$\rightarrow J = \{j_1, \dots, j_k\}$$



$$\begin{aligned}
 \mathbb{E} \left[\frac{1}{k} \sum_{i \in \Omega_k} \|x - X_i\|^2 \right] &\leq \underbrace{\mathbb{E} \left[\frac{1}{k} \sum_{m=1}^k \|x - X_{jm}^*\|^2 \right]}_{= \frac{1}{k} \sum_{m=1}^k \mathbb{E} [\|x - X_{jm}^*\|^2]} \\
 &\quad \text{closest to } x \text{ in the group } G_m. \\
 \Rightarrow \mathbb{E}[\|x\|^2] &\leq \frac{1}{(n/k)^{2/d}}
 \end{aligned}$$

$\lesssim \left(\frac{k}{n}\right)^{2/d}.$

Conclusion:

$$\mathbb{E} \left[Q_p(\hat{f}_k^{NN}) - Q_p^* \right] \lesssim \alpha^2 \left(\frac{k}{n} \right)^{2/d} + \frac{\sigma^2}{k}.$$

Conclusion :

- * Two local averaging methods:
Partition estimators + kNN
- * Easy to compute and interpret.
- * Depends on one parameter that needs to be tuned.
- * Suffers from the curse of dimensionality.
 \hookrightarrow Rate of CV $n^{-\frac{2}{d+2}}$

Other local averaging method:

Nadaraya - Watson : see Lecture Notes.