

# International Standard

## ISO 29842

# Sensory analysis — Methodology — Balanced incomplete block designs

Analyse sensorielle — Méthodologie — Plans de présentation en blocs incomplets équilibrés

Second edition 2024-06



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Published in Switzerland

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## Foreword

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The procedures used to develop this document and those intended for its further maintenance are described in the ISO/IEC Directives, Part 1. In particular, the different approval criteria needed for the different types of ISO document should be noted. This document was drafted in accordance with the editorial rules of the ISO/IEC Directives, Part 2 (see <a href="https://www.iso.org/directives">www.iso.org/directives</a>).

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This document was prepared by Technical Committee ISO/TC 34, *Food products*, Subcommittee SC 12, *Sensory analysis*.

This second edition cancels and replaces the first edition (ISO 29842:2011), of which it constitutes a minor revision. It also incorporates the Amendment ISO 29842:2011/Amd.1:2015.

The changes are as follows:

- the corrections of the amendment have been included:
- in Table 2, the formulae in the column "Mean square" have been corrected.

Any feedback or questions on this document should be directed to the user's national standards body. A complete listing of these bodies can be found at <a href="https://www.iso.org/members.html">www.iso.org/members.html</a>.

## Sensory analysis — Methodology — Balanced incomplete block designs

## 1 Scope

This document specifies a method for the application of balanced incomplete block designs to sensory descriptive and hedonic tests.

This document is applicable when the number of test samples exceeds the number of evaluations that an assessor can perform reliably in a single session.

This document also specifies the fundamental characteristics of balanced incomplete block designs and establishes guidelines for their application in sensory evaluation.

#### 2 Normative references

The following documents are referred to in the text in such a way that some or all of their content constitutes requirements of this document. For dated references, only the edition cited applies. For undated references, the latest edition of the referenced document (including any amendments) applies.

ISO 3534-1, Statistics — Vocabulary and symbols — Part 1: General statistical terms and terms used in probability

ISO 5492, Sensory analysis — Vocabulary

#### 3 Terms and definitions

For the purposes of this document, the terms and definitions given in ISO 3534-1, ISO 5492 and the following apply.

ISO and IEC maintain terminology databases for use in standardization at the following addresses:

- ISO Online browsing platform: available at <a href="https://www.iso.org/obp">https://www.iso.org/obp</a>
- IEC Electropedia: available at <a href="https://www.electropedia.org/">https://www.electropedia.org/</a>

#### 3.1

#### block design

<sensory analysis> multi-sample serving protocol in which an assessor evaluates all or a subset of the samples in a study

### 3.2

#### repetition

one occurrence of an experimental design

## 4 Specification of balanced incomplete block designs

Balanced incomplete block (BIB) designs apply to sensory tests in which the total number of samples is greater than the number that can be evaluated before sensory and psychological fatigue set in. In BIB designs, each assessor evaluates only a subset of the total number of samples in a single session.

An example of a BIB design is shown in Table 1.

Table 1 — A BIB design with five samples and 10 block/assessors

Block	Test sample								
(assessor)	1	2	3	4	5				
1	×	×	×	_	_				
2	×	×	_	×	_				
3	×	×	_	_	×				
4	×	_	×	×	_				
5	×	_	×	_	×				
6	×	_	_	×	×				
7	_	×	×	×	_				
8	_	×	×	_	×				
9	_	×	_	×	×				
10	_	_	×	×	×				

In a BIB design each assessor evaluates a subset, k, of the total number of samples, t, where k < t. The subset of samples that an assessor evaluates is selected so that, in a single repetition of the BIB design, every sample is evaluated an equal number of times and all possible pairs of two samples are evaluated by an equal number of assessors.

The notation most commonly used in a BIB design is as follows:

- *t* number of test samples;
- k number of samples evaluated by an assessor in a single session (k < t);
- *b* total number of blocks (typically, assessors) in one repetition of the BIB design;
- r number of times each test sample is evaluated in one repetition of the BIB design;
- $-\lambda$  number of times each pair of samples is evaluated by the same assessor;
- *p* number of times the basic BIB design is repeated.

Notationally, each assessor evaluates k of the t samples (k < t). The subset of k samples that an assessor evaluates is selected so that in a single repetition of the BIB design every sample is evaluated an equal number of times and all possible pairs of samples are evaluated by an equal number of assessors. The number of blocks (assessors) required to complete a single repetition of the BIB design is denoted by b. The number of times each sample is evaluated in a single repetition of the BIB design is denoted by r and the number of times every pair of two samples is evaluated together is denoted by  $\lambda$ .

The entire BIB design may need to be repeated several times in order to achieve an adequate level of precision for the study. The number of repetitions of the basic BIB design is denoted by p. The total number of blocks (typically assessors) is then p\*b and the total of evaluations per sample is then p\*r. The total number of times each pair of samples is seen together is  $p*\lambda$ .

The constant values of r and  $\lambda$  for all samples in the BIB design imparts important statistical properties to data collected from the design. The constant value of r ensures that the mean values of all of the samples are estimated with equal precision. The constant value of  $\lambda$  ensures that all pair-wise comparisons between any two samples are equally sensitive.

## 5 Data analysis

#### 5.1 General

Two types of data can be collected using balanced incomplete block designs. Ratings data, or scores, are obtained when assessors use a scale to report the perceived intensities of the attributes or impressions they are evaluating. Rank data are obtained when assessors order the samples from lowest to highest (or vice

versa) relative to the attribute they are evaluating. Different data analysis methods are used for ratings and rank data.

### 5.2 Analysis of variance for rating data

Analysis of variance (ANOVA) is used to analyse ratings data obtained from the BIB design. The sources of variability accounted for in the ANOVA model for the BIB design are the same as those accounted for in a randomized (complete) block design. In both cases, the total variability is partitioned into the separate effects of blocks (typically assessors), treatments (typically samples) and errors. Because each assessor evaluates only a subset of the total number of test samples, more complicated formulae are required to calculate the ANOVA sum-of-squares for the BIB design than for the randomized (complete) block design. The sensory analyst shall ensure that the program used to perform the analysis is capable of handling BIB designs. In many statistical computer packages, the ANOVA procedure applies only to complete designs, i.e. studies in which every assessor evaluates all of the test samples. For incomplete designs, such as BIB designs, the general linear model (GLM) procedure or a mixed model procedure is required.

The form of the ANOVA used to analyse BIB data depends on how the design is administered.

Where the experiment is of the form of the example in <u>Table 1</u>, with a single repetition of the design, the ANOVA table is as shown in <u>Table 2</u>.

Source of variation Degrees of freedom Sum of squares F Mean square (DF) (SS) (MS)  $v_{\rm T} = t^*r - \overline{1}$  $S_{\mathrm{T}}$ Total  $S_{\mathrm{B}}$  $v_{\rm B} = b - 1$ Assessors Samples (adjusted for assessors)  $S_{\rm S}$  $MS_S = S_S/\nu_S$  $MS_{\rm S}/MS_{\rm E}$  $v_{\rm E} = t^*r - t - b + 1$  $S_{\rm E}$  $MS_{\rm E} = S_{\rm E}/\nu_{\rm E}$ 

Table 2 — ANOVA table for balanced incomplete block design (single repetition)

If the F-statistic in Table 2 exceeds the upper- $\alpha$  critical value of an F with the corresponding degrees of freedom, then the null hypothesis assumption of equivalent mean ratings is rejected. If the F-statistic is significant, a multiple comparison procedure, such as Fisher's LSD (least significant difference), L, shall be applied to determine which samples are significantly different from one another. The equation for Fisher's LSD, L, appropriate for a single repetition of this BIB design is:

$$L = t_{\alpha/2, v_E} \sqrt{\frac{2MS_E}{r}} \sqrt{\frac{k(t-1)}{(k-1)t}}$$

where

*t*, *k* and *r* are as defined in Clause 4;

*MS*<sub>E</sub> is the mean square for error from the ANOVA table;

 $v_{\rm E}$  is the number of degrees of freedom for error from the ANOVA table;

 $t_{\alpha/2,\nu_{\rm E}}$  is the upper  $\alpha/2$  critical value of Student's *t*-distribution with  $\nu_{\rm E}$  degrees of freedom.

The same value of  $\alpha$  shall be used for assessing the significance of the *F*-statistic and in Fisher's LSD, *L*.

The BIB design shall be repeated p times to achieve an adequate level of precision from the study. If the total number of blocks is too large for each assessor to evaluate all of them, each of the p\*b assessors shall evaluate only one block of k samples. Within each block, the order in which the k samples are evaluated shall be done at random. The ANOVA table for this design is presented in Table 3.

## Table 3 — ANOVA table for balanced incomplete block design

(p repetitions performed by p\*b assessors each evaluating a single block of k samples)

Source of variation	Degrees of freedom (DF)	Sum of squares (SS)	Mean square (MS)	F
Total	$v_{\mathrm{T}} = t^* p^* r - 1$	ST		
Blocks (assessors)	$v_{\rm B} = p^*b - 1$	SB		
Samples (adjusted for assessors)	$v_{\rm S} = t - 1$	SS	$MS_{\rm S} = S_{\rm S}/\nu_{\rm S}$	$MS_{\rm S}/MS_{\rm E}$
Error	$v_{\rm E} = t^*p^*r - t - p^*b + 1$	SE	$MS_{\rm E} = S_{\rm E}/\nu_{\rm E}$	

If the *F*-statistic in <u>Table 3</u> exceeds the critical value of an *F* with the corresponding degrees of freedom, then the null hypothesis assumption of equivalent mean ratings is rejected. If the *F*-statistic is significant, a multiple comparison procedure, such as Fisher's LSD, *L*, shall be applied to determine which samples are significantly different from one another. The equation for Fisher's LSD, *L*, appropriate for a BIB design is:

$$L = t_{\alpha/2, v_E} \sqrt{\frac{2MS_E}{pr}} \sqrt{\frac{k(t-1)}{(k-1)t}}$$

where

t, k, p and r are as defined in Clause 4;

 $MS_{\rm E}$  is the mean square for error from the ANOVA table;

 $v_{\rm E}$  is the number of degrees of freedom for error from the ANOVA table;

 $t_{\alpha/2\nu_{\rm F}}$  is the upper  $\alpha/2$  critical value of Student's *t*-distribution with  $\nu_{\rm E}$  degrees of freedom.

The same value of  $\alpha$  shall be used for assessing the significance of the *F*-statistic and in Fisher's LSD, *L*.

If each assessor evaluates all b blocks in the BIB design, then the "assessor effect" and the "assessor-by-sample" interaction can be partitioned out of the total variability (see <u>Table 4</u>). This approach is especially applicable when the total number of blocks in one repetition of the BIB design is small (e.g.  $b \le 6$ ). The order in which the blocks are presented to the assessor shall be done at random. Within each block, the order in which the samples are evaluated shall be done at random. In either approach, the variability that arises from the assessors is accounted for and the interaction between assessors and samples replaces the error term that was used in <u>Tables 2</u> and <u>3</u>.

Table 4 — ANOVA table for balanced incomplete block design

(p repetitions performed by p assessors each evaluating b blocks of k samples)

Source of variation	Degrees of freedom (DF)	Sum of squares (SS)	Mean square (MS)	F
Total	$v_{\mathrm{T}} = t^* p^* r - 1$	$S_{\mathrm{T}}$		
Assessor	$v_{\rm P} = p - 1$	$S_{ m P}$		
Blocks (within assessor)	$v_{B(P)} = p^*(b-1)$	$S_{B(P)}$		
Samples (adjusted for assessor)	$v_{\rm S} = t - 1$	$S_{ m S}$	$MS_{\rm S} = S_{\rm S}/\nu_{\rm S}$	$MS_{S}/MS_{A*S}$
Assessor*samples	$v_{A*S} = (p-1)(t-1)$	$S_{A*S}$	$MS_{A*S} = S_{A*S}/\nu_{A*S}$	
Residual	$v_{\rm E} = p^*(t^*r - t - b + 1)$	$S_{\mathrm{E}}$	$MS_{\rm E} = S_{\rm E}/\nu_{\rm E}$	

If the *F*-statistic in <u>Table 4</u> exceeds the critical value of an *F* with the corresponding degrees of freedom, then the null hypothesis assumption of equivalent mean ratings is rejected. If the *F*-statistic is significant, a multiple comparison procedure, such as Fisher's LSD, *L*, shall be applied to determine which samples are significantly different from one another. The equation for Fisher's LSD, *L*, appropriate for a BIB design is:

$$L = t_{\alpha/2, v_{\text{A*S}}} \sqrt{\frac{2MS_{\text{A*S}}}{pr}} \sqrt{\frac{k(t-1)}{(k-1)t}}$$

where

t, k, p and r are as defined in Clause 4;

*MSA*\*S is the mean square for the assessor\*sample interaction from the ANOVA table;

 $v_{A*S}$  is the number of degrees of freedom for the assessor\*sample interaction from the

ANOVA table;

 $t_{\alpha/2,\nu_{A*S}}$  is the upper  $\alpha/2$  critical value of Student's *t*-distribution with  $\nu_{A*S}$  degrees of freedom.

The same value of  $\alpha$  shall be used for assessing the significance of the *F*-statistic and in Fisher's LSD, *L*.

## 5.3 Friedman's sum rank analysis for rank data<sup>1)</sup>

A Friedman-type statistic shall be applied to rank data arising from a BIB design. Friedman's test statistic, *F*test, is given by:

$$F_{\text{test}} = \frac{12}{p\lambda t(k+1)} \sum_{j=1}^{t} R_j^2 - \frac{3(k+1)pr^2}{\lambda}$$

where  $t, k, r, \lambda$  and p are as defined above and Rj is the rank sum of the jth sample (Reference [8]). Tables of critical values of  $F_{\text{test}}$  are available for selected combinations of t=3 ... 6, k=2 ... 5, and p=1 ... 7 (Reference [9]). However, in most sensory studies, the total number of blocks exceeds the values in the tables. For these situations, the test procedure is to reject the assumption of equivalency among the samples if the value of  $F_{\text{test}}$  exceeds the upper  $\alpha$  critical value of a  $\chi^2$ -statistic with (t-1) degrees of freedom.

If the  $\chi^2$ -statistic is significant, then a multiple comparison procedure shall be performed to determine which samples differ significantly from one another. The equation for Fisher's LSD, L, appropriate for a BIB design is:

$$L = z_{\alpha/2} \sqrt{\frac{p(k+1)(rk-r+\lambda)}{6}}$$

where

p, k, r and  $\lambda$  are as defined in Clause 4;

 $z_{\alpha/2}$  is the upper- $\alpha/2$  critical value of the standard normal distribution.

The same value of  $\alpha$  shall be used for assessing the significance of the  $F_{\text{test}}$ -statistic and in Fisher's LSD, L.

<sup>1)</sup> There are several statistical methods for analysing rank data that are obtained from a BIB design. Interested readers are encouraged to review the statistical literature on the topic. Friedman's method has been chosen for detailed discussion because it is statistically powerful and computationally convenient.

## 6 Application in sensory evaluation

The number of samples that are evaluated by an assessor in a single session, k, shall not exceed the assessor's ability to provide reliable ratings of the samples. The value of k depends on several factors including the overall intensity of the sensory characteristics of the samples, the degree of carryover effects (e.g. lingering aftertastes), and the number of attributes the assessor is rating. The sensory analyst shall limit the number of samples an assessor evaluates in a single session to avoid both sensory and psychological fatigue.

Sensory and psychological fatigue limit the number of evaluations an assessor can perform reliably. However, the sensory analyst shall make every attempt to have each assessor perform as many evaluations as possible in order to counteract the influence of context effects. Since each assessor evaluates only a subset of the total number of samples, the assessor is not exposed to the full range of variability present in the samples. This can lead the assessors to exaggerate the differences in the ratings that they assign to the products. The larger the proportion of samples the assessors evaluate in a single session, the greater the range of sensory differences they are exposed to and, therefore, the lesser the chance that context effects bias the results of the study.

The total number of evaluations of each sample,  $r^*p$ , shall be determined based on the level of sensitivity required of the test. The same criteria used to decide on the total number of evaluations in a complete block design shall be applied. For example, if it is standard practice to perform 12 evaluations per sample in a complete block study with a descriptive analysis panel and the number of replications in one repetition of the BIB design is r = 3, then the BIB design shall be replicated p = 4 times in order to achieve the required total number of evaluations.

## Annex A

(informative)

## Catalogue of incomplete block designs

The following BIB designs apply for  $t = 3 \dots 10$  and  $k = 2 \dots t - 1$  or  $k = 2 \dots 6$ , whichever is lesser. It is possible to construct BIB designs for any values of t and t by collecting all possible combinations of t items in groups of size t. However, it is sometimes possible to construct BIB designs that have fewer blocks than are required to form all possible combinations. For any value of t and t, the catalogue presents the BIB design with the smallest number of blocks.

a)	2 of 3 BIB:	All possible pairs of three samples	$(t = 3, k = 2, b = 3, r = 2, \lambda = 1).$
b)	2 of 4 BIB:	All possible pairs of four samples	$(t = 4, k = 2, b = 6, r = 3, \lambda = 1).$
c)	3 of 4 BIB:	All possible triads of four samples	$(t = 4, k = 3, b = 4, r = 3, \lambda = 2).$
d)	2 of 5 BIB:	All possible pairs of five samples	$(t = 5, k = 2, b = 10, r = 4, \lambda = 1).$
e)	3 of 5 BIB:	All possible triads of five samples	$(t = 5, k = 3, b = 10, r = 6, \lambda = 3).$
f)	4 of 5 BIB:	All possible four-tuples of five samples	$(t = 5, k = 4, b = 5, r = 4, \lambda = 3).$
g)	2 of 6 BIB:	All possible pairs of six samples	$(t = 6, k = 2, b = 15, r = 5, \lambda = 1).$
h)	3 of 6 BIB:	10 Triads of six samples	$(t = 6, k = 3, b = 10, r = 5, \lambda = 2).$

See Table A.1.

Table A.1 — 10 Triads of six samples

Block			San	ıple		
Біоск	1	2	3	4	5	6
1	×	×			×	
2	×	×				×
3	×		×	×		
4	×		×			×
5	×			×	×	
6		×	×	×		
7		×	×		×	
8		×		×		×
9			×		×	×
10				×	×	×

i)	4 of 6 BIB:	All possible four-tuples of six samples	$(t = 6, k = 4, b = 15, r = 10, \lambda = 6).$
j)	5 of 6 BIB:	All possible five-tuples of six samples	$(t = 6, k = 5, b = 6, r = 5, \lambda = 4).$
k)	2 of 7 BIB:	All possible pairs of seven samples	$(t = 7, k = 2, b = 21, r = 6, \lambda = 1).$
l)	3 of 7 BIB:	Seven triads of seven samples	$(t = 7, k = 3, b = 7, r = 3, \lambda = 1).$

## See <u>Table A.2</u>.

Table A.2 — Seven triads of seven samples

Block		Sample								
Diock	1	2	3	4	5	6	7			
1	×	×			×					
2	×		×			×				
3	×			×			×			
4		×	×				×			
5		×		×		×				
6			×	×	×					
7					×	×	×			

m) 4 of 7 BIB:

14 Four-tuples of seven samples

 $(t = 7, k = 4, b = 14, r = 8, \lambda = 4).$ 

See <u>Table A.3</u>.

Table A.3 — 14 Four-tuples of seven samples

Dlask		Sample									
Block	1	2	3	4	5	6	7				
1	×	×	×		×						
2	×	×	×				×				
3	×	×		×		×					
4	×	×				×	×				
5	×		×	×	×						
6	×		×	×		×					
7	×			×	×		×				
8	×				×	×	×				
9		×	×	×			×				
10		×	×		×	×					
11		×		×	×	×					
12		×		×	×		×				
13			×	×		×	×				
14			×		×	×	×				

n) 5 of 7 BIB: All possible five-tuples of seven samples

 $(t = 7, k = 5, b = 21, r = 15, \lambda = 10).$ 

o) 6 of 7 BIB: All possible six-tuples of seven samples

 $(t = 7, k = 6, b = 7, r = 6, \lambda = 5).$ 

p) 2 of 8 BIB: All possible pairs of eight samples

 $(t = 8, k = 2, b = 28, r = 7, \lambda = 1).$ 

q) 3 of 8 BIB: All possible triads of eight samples

 $(t = 8, k = 3, b = 56, r = 21, \lambda = 6).$ 

r) 4 of 8 BIB: 14 Four-tuples of eight samples

 $(t = 8, k = 4, b = 14, r = 7, \lambda = 3).$ 

See <u>Table A.4</u>.

Table A.4 — 14 Four-tuples of eight samples

Dl. d		Sample									
Block	1	2	3	4	5	6	7	8			
1	×	×	×	×							
2	×	×			×	×					
3	×	×					×	×			
4	×		×		×		×				
5	×		×			×		×			
6	×			×	×			×			
7	×			×		×	×				
8		×	×		×			×			
9		×	×			×	×				
10		×		×	×		×				
11		×		×		×		×			
12			×	×	×	×					
13			×	×			×	×			
14					×	×	×	×			

s) 5 of 8 BIB: All possible five-tuples of eight samples  $(t = 8, k = 5, b = 56, r = 35, \lambda = 20)$ .

t) 6 of 8 BIB: All possible six-tuples of eight samples  $(t = 8, k = 6, b = 28, r = 21, \lambda = 15)$ .

u) 2 of 9 BIB: All possible pairs of nine samples  $(t = 9, k = 2, b = 36, r = 8, \lambda = 1)$ .

v) 3 of 9 BIB: 12 Triads of nine samples  $(t = 9, k = 3, b = 12, r = 4, \lambda = 1)$ .

See <u>Table A.5</u>.

Table A.5 — 12 Triads of nine samples

Dlask	Sample									
Block	1	2	3	4	5	6	7	8	9	
1	×	×	×							
2	×			×			×			
3	×				×				×	
4	×					×		×		
5		×		×					×	
6		×			×			×		
7		×				×	×			
8			×	×				×		
9			×		×		×			
10			×			×			×	
11				×	×	×				
12							×	×	×	

w) 4 of 9 BIB: 18 Four-tuples of nine samples  $(t = 9, k = 4, b = 18, r = 8, \lambda = 3)$ .

See <u>Table A.6</u>.

Table A.6 — 18 Four-tuples of nine samples

DI I		Sample										
Block	1	2	3	4	5	6	7	8	9			
1	×	×	×	×								
2	×	×		×					×			
3	×	×			×		×					
4	×		×			×		×				
5	×		×					×	×			
6	×			×		×	×					
7	×				×	×			×			
8	×				×		×	×				
9		×	×		×	×						
10		×	×			×	×					
11		×				×		×	×			
12		×		×	×			×				
13		×					×	×	×			
14			×	×	×			×				
15			×	×			×		×			
16			×		×		×		×			
17				×	×	×			×			
18				×		×	×	×				

x) 5 of 9 BIB: 18 Five-tuples of nine samples

 $(t = 9, k = 5, b = 18, r = 10, \lambda = 5).$ 

See <u>Table A.7</u>.

Table A.7 — 18 Five-tuples of nine samples

pll		Sample												
Block	1	2	3	4	5	6	7	8	9					
1	×	×	×		×				×					
2	×	×	×				×	×						
3	×	×		×		×		×						
4	×	×			×	×		×						
5	×	×				×	×		×					
6	×		×	×	×	×								
7	×		×	×	×		×							
8	×		×			×	×		×					
9	×			×	×			×	×					
10	×			×			×	×	×					
11		×	×	×		×			×					
12		×	×	×			×	×						
13		×	×		×			×	×					
14		×		×	×	×	×							
15		×		×	×		×		×					

Table A.7 (continued)

Dlogle		Sample												
Block	1	2	3	4	5	6	7	8	9					
16			×	×		×		×	×					
17			×		×	×	×	×						
18					×	×	×	×	×					

y) 6 of 9 BIB: 12 Six-tuples of nine samples  $(t = 9, k = 6, b = 12, r = 8, \lambda = 5)$ .

See <u>Table A.8</u>.

Table A.8 — 12 Six-tuples of nine samples

Dla ala		Sample												
Block	1	2	3	4	5	6	7	8	9					
1	×	×		×	×		×	×						
2	×	×			×	×	×		×					
3	×	×	×	×	×	×								
4	×	×	×				×	×	×					
5	×	×		×		×		×	×					
6	×		×	×	×			×	×					
7	×		×	×		×	×		×					
8	×		×		×	×	×	×						
9		×	×	×	×		×		×					
10		×	×	×		×	×	×						
11		×	×		×	×		×	×					
12				×	×	×	×	×	×					

z) 2 of 10 BIB: All possible pairs of 10 samples  $(t = 10, k = 2, b = 45, r = 9, \lambda = 1)$ .

aa) 3 of 10 BIB: 30 Triads of 10 samples  $(t = 10, k = 3, b = 30, r = 9, \lambda = 2)$ .

See <u>Table A.9</u>.

Table A.9 — 30 Triads of 10 samples

Block		Sample											
BIOCK	1	2	3	4	5	6	7	8	9	10			
1	×	×	×										
2	×	×		×									
3	×		×		×								
4	×			×		×							
5	×				×		×						
6	×					×		×					
7	×						×		×				
8	×							×		×			
9	×								×	×			

Table A.9 (continued)

		Sample												
Block	1	2	3	4	5	6	7	8	9	10				
10		×	×			×								
11		×		×						×				
12		×			×			×						
13		×			×				×					
14		×				×	×							
15		×					×		×					
16		×						×		×				
17			×	×			×							
18			×	×				×						
19			×		×	×								
20			×				×			×				
21			×					×	×					
22			×						×	×				
23				×	×				×					
24				×	×					×				
25				×		×			×					
26				×			×	×						
27					×	×				×				
28					×		×	×						
29						×	×			×				
30						×		×	×					

bb) 4 of 10 BIB:

15 Four-tuples of 10 samples

 $(t = 10, k = 4, b = 15, r = 6, \lambda = 2).$ 

See <u>Table A.10</u>.

Table A.10 — 15 Four-tuples of 10 samples

Dlask			,		San	nple				1
Block	1	2	3	4	5	6	7	8	9	10
1	×	×	×	×						
2	×	×			×	×				
3	×		×				×	×		
4	×			×					×	×
5	×				×		×		×	
6	×					×		×		×
7		×	×			×			×	
8		×		×			×			×
9		×			×			×		×
10		×					×	×	×	
11			×	×	×			×		
12			×		×				×	×

Table A.10 (continued)

Block		Sample											
	1	2	3	4	5	6	7	8	9	10			
13			×			×	×			×			
14				×	×	×	×						
15				×		×		×	×				

cc) 5 of 10 BIB: 18 Five-tuples of 10 samples  $(t = 10, k = 5, b = 18, r = 9, \lambda = 4)$ .

See <u>Table A.11</u>.

Table A.11 — 18 Five-tuples of 10 samples

Dlook		Sample												
Block	1	2	3	4	5	6	7	8	9	10				
1	×	×	×	×	×									
2	×	×	×			×	×							
3	×	×		×		×			×					
4	×	×			×		×	×						
5	×		×			×		×	×					
6	×		×				×	×		×				
7	×			×	×	×				×				
8	×			×				×	×	×				
9	×				×		×		×	×				
10		×	×	×				×		×				
11		×	×		×				×	×				
12		×		×			×	×	×					
13		×			×	×		×		×				
14		×				×	×		×	×				
15			×	×	×		×		×					
16			×	×		×	×			×				
17			×		×	×		×	×					
18				×	×	×	×	×						

dd) 6 of 10 BIB: 15 Six-tuples of 10 samples  $(t = 10, k = 6, b = 15, r = 9, \lambda = 5)$ .

See <u>Table A.12</u>.

Table A.12 — 15 Six-tuples of 10 samples

Dla ala					San	nple				
Block	1	2	3	4	5	6	7	8	9	10
1	×	×	×		×		×			×
2	×	×	×					×	×	×
3	×	×		×	×			×	×	
4	×	×		×		×	×	×		
5	×	×				×	×		×	×
6	×		×	×	×	×				×
7	×		×	×		×	×		×	
8	×		×		×	×		×	×	
9	×			×	×		×	×		×
10		×	×	×	×		×		×	
11		×	×	×		×		×		×
12		×	×		×	×	×	×		
13		×		×	×	×			×	×
14			×	×			×	×	×	×
15					×	×	×	×	×	×

## Annex B

(informative)

## Example of balanced incomplete block design with ratings data

## **B.1** Problem and situation

The quality control manager of a mustard factory routinely screens samples of finished product to be added to the pool of reference samples. New reference samples are needed at regular intervals, as the older samples will have changed with time and are no longer appropriate. The procedure is also used to eliminate from the pool any current reference samples that may have deteriorated.

## **B.2** Test design

Samples from six lots are evaluated for overall off-flavour. Evaluations are performed by 15 trained assessors using a 10 point category scale from 0, representing no off-flavour, to 9, indicating extreme off-flavour. The assessors evaluate four of the six samples. The four samples that each assessor evaluates are determined by the BIB design presented in <a href="Table B.1">Table B.1</a>. Each of the 15 assessors is randomly assigned one block of four samples from the design. The order of presentation of the samples within each block is randomized.

**Table B.1** — BIB design for ratings data t = 6, k = 4, r = 10, b = 15,  $\lambda = 6$ 

Block	Sample									
(assessor)	1	2	3	4	5	6				
1	×	×	×	×						
2	×			×	×	×				
3		×	×		×	×				
4	×	×	×		×					
5	×	×		×		×				
6			×	×	×	×				
7	×	×	×			×				
8	×		×	×	×					
9		×		×	×	×				
10	×	×		×	×					
11	×		×		×	×				
12		×	×	×		×				
13	×	×			×	×				
14	×		×	×		×				
15		×	×	×	×					

#### **B.3** Results

The off-flavour intensity ratings are presented in <u>Table B.2</u>. The data are analysed by a program capable of performing a BIB design ANOVA. The resulting ANOVA table is presented in <u>Table B.3</u>. The F-statistic for samples is highly significant (p < 0.000 1), indicating that there are perceptible differences in off-flavour intensity among the samples. An LSD multiple comparison procedure is applied to the average ratings of the samples to determine which samples have significantly different off-flavour intensities. The results in

<u>Table B.4</u> reveal that sample 1 has significantly higher off-flavour than all of the other samples. There are no significant differences among the remaining samples.

Table B.2 — Off-flavour intensity ratings of six mustards

			San	nple		
Assessor	1	2	3	4	5	6
1	6	1	1	2		
2	6			1	3	3
3		4	2		5	2
4	7	2	3		2	
5	3	5		1		1
6			1	1	3	2
7	7	4	4			3
8	2		1	1	1	
9		2		2	2	3
10	4	2		2	5	
11	5		3		1	1
12		3	2	1		2
13	4	2			1	1
14	5		2	2		1
15		2	4	5	3	

Table B.3 — Balanced incomplete block ANOVA table of ratings data: off-flavour intensity of six mustards

Source of variation	Sum of squares (SS)	Degrees of freedom (DF)	Mean square (MS)	F	р
Total	156,60	59			
Assessors (blocks)	38,58	14			
Samples (treatments, adjusted for blocks)	64,08	5	12,82	9,60	< 0,000 1
Error	53,42	40	1,34		

Table B.4 — Adjusted means of off-flavour intensities of six mustards

Sample		1	2	3	4	5	6	
Adjusted mean		5,0 A	2,5 B	2,2 B	2,0 B	2,6 B	1,9 B	
NOTE Means with no letters in common are significantly different at the 5 % level of significance ( $L = 1,1$ ).								

## **Annex C**

(informative)

## Example of balanced incomplete block design with rank data

## C.1 Problem and situation

A pepper sauce manufacturer wants to assess the spiciness of 15 varieties of peppers. The manufacturer wants to specify a limited number of varieties to use in its products.

## C.2 Test design

Because of the high spiciness levels of the peppers, a balanced incomplete design is chosen. Each assessor evaluates k = 3 of the t = 15 samples. The basic BIB design consists of b = 35 blocks, in which each sample is evaluated r = 7 times and every pair of samples is evaluated together once ( $\lambda = 1$ ). The basic design is repeated p = 3 times to obtain a total of p\*r = 21 evaluations per sample.

A randomly selected group of 105 assessors was recruited to participate in the test. The assessors screened for their ability to distinguish different levels of spiciness and were instructed on the use of the ranking method. Because of their lack of extensive training in sensory testing, the assessors were instructed to rank the three samples as: 1 = most spicy; 2 = medium spicy; and 3 = least spicy.

## **C.3** Analysis of results

To make the results easier to analyse, the rank data from the study are arranged as shown in Table C.1. The rank sum for a given pepper variety is simply the sum of all the numbers in the column corresponding to that variety. The value of Friedman's test statistic,  $F_{\text{test}}$ , is computed to determine if there are any differences in spiciness among the varieties. The value of  $F_{\text{test}}$  = 68,53 exceeds the upper 5 % critical value of a  $\chi^2$  with (t-1) = 14 degrees of freedom ( $\chi^2_{14,0,05}$  = 23,69), and it is concluded that there are indeed significant differences in the dataset. A 95 % LSD multiple comparison statistic for rank data is computed to determine which of the varieties are significantly different (L = 11). The results of the multiple comparisons are presented in Table C.2.

Table C.1 — Results obtained in BIB example with rank data: spiciness of peppers

Block	Sample or variety a														
(assessor)	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15
1	1	2	3												
2				3			1			2					
3					1					3					2
4						3					2		1		
5							3		1					2	
6	3			2	1										
7		2						3		1					
8			3										2	1	
9						3			2						1
10							2				1	3			
a Responses:	1 = mo:	st spicy	; 2 = me	dium sp	oicy; 3 =	least s	picy.								

**Table C.1** (continued)

Block	Sample or variety a														
(assessor)	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15
101	2													3	1
102		1		3		2									
103			1					2			3				
104					1				2			3			
105							3			2			1		
Rank sum	35	45	54	43	28	37	55	42	37	50	49	50	34	42	29
a Responses	: 1 = mo	st spicy	; 2 = me	dium sr	oicy; 3 =	least sı	oicy.		•						

Table C.2 — Summary of results and statistical analysis of the BIB rank data: spiciness of peppers

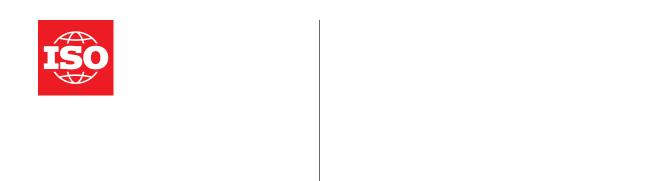
Sample or variety	Rank sum				ı
5	28	a			
15	29	a			
13	34	a	b		
1	35	a	b		
6	37	a	b		
9	37	a	b		
14	42		b	С	
8	42		b	С	
4	43		b	С	
2	45			С	d
11	49			С	d
10	50			С	d
12	50			С	d
3	54				d
7	55				d

NOTE Rank sums with no letters in common are significantly different at the 5 % significance level (Lrank = 11).

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<sup>2)</sup> Withdrawn.



ICS 67.240

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