

# Probabilistic forwarding of coded packets on networks

**Vinay Kumar B R**, Navin Kashyap  
Indian Institute of Science

IEEE International Symposium on Information Theory (ISIT) - 2019  
11 Jul, 2019  
Paris, France

## Motivating example

---



## Motivating example

---



**Sensors in a field**

## Motivating example

---



**Sensors in a field**

# Motivating example

---

## Sensors in a field



Update sensing parameters  
among all nodes

## Network of IoT nodes



Over-the-air programming of the  
nodes

# Motivating example

---

## Sensors in a field



Update sensing parameters  
among all nodes

## Network of IoT nodes



Over-the-air programming of the  
nodes

## Broadcast information with the following constraints

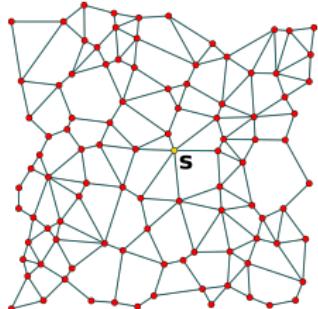
### Nodes:

- Energy constrained
- Limited computational ability
- Limited knowledge of the network

### Algorithm:

- Completely distributed
- Run in finite time

- ▶ Broadcast over an ad-hoc network



A network of nodes  $\mathcal{G} = (V, E)$ .  
 $|V| = N$ .

A source node  $s$  has  $k$  data packets to be broadcast to all the other nodes in an ad-hoc network.

## Flooding

- ▶ Each node forwards every received packet to all its one-hop neighbours.
- ▶ Subsequent receptions of the same packet are neglected.
- ▶ Total number of transmissions  $= kN$ .
- ▶ **Wasteful**<sup>1</sup>: nodes will receive same packet multiple times

---

<sup>1</sup>Y.-C. Tseng, S.-Y. Ni, Y.-S. Chen, and J.-P. Sheu, "The broadcast storm problem in a mobile ad hoc network," *Wireless Networks*, vol. 8, no. 2/3, pp. 153–167, 2002.

- ▶ Retransmission probability  $p$ .
- ▶ Source transmits all  $k$  packets to all its neighbors.
- ▶ Each node, upon receiving packet # $j$  for the first time, forwards it to all its neighbours with probability  $p$ ; does nothing with probability  $1 - p$ .
- ▶ Subsequent receptions of the same packet are neglected.

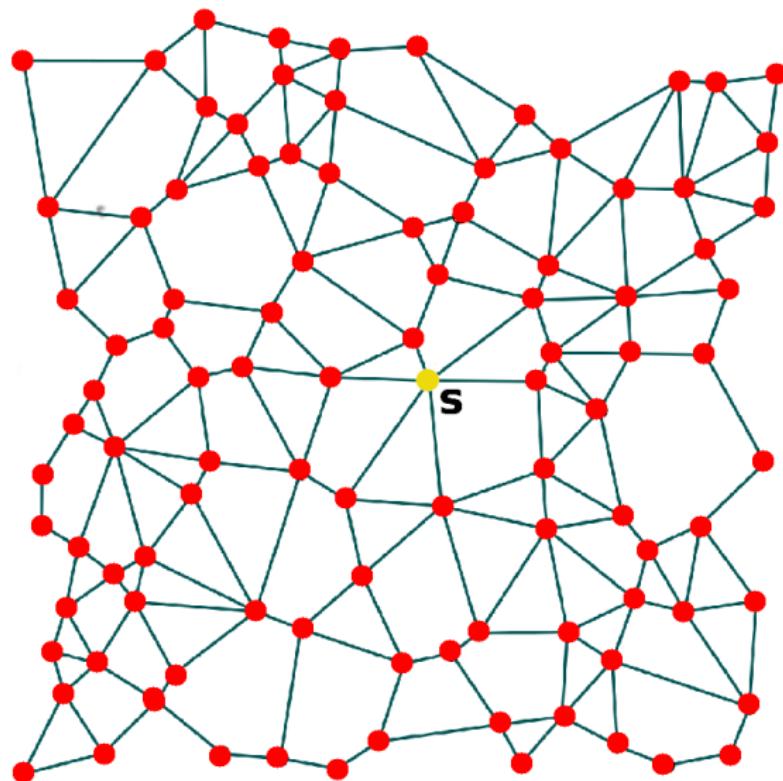
---

<sup>1</sup>Y. Sasson, D. Cavin, and A. Schiper, "Probabilistic broadcast for flooding in wireless mobile ad hoc networks," in Proc. IEEE Wireless Communications and Networking Conf. (WCNC) 2003, vol. 2, March 16–20, 2003, pp. 1124–1130.

<sup>1</sup>Z. J. Haas, J. Y. Halpern, and L. Li, "Gossip-based ad hoc routing," IEEE/ACM Trans. Networking, vol. 14, no. 3, pp. 479–491, 2006.

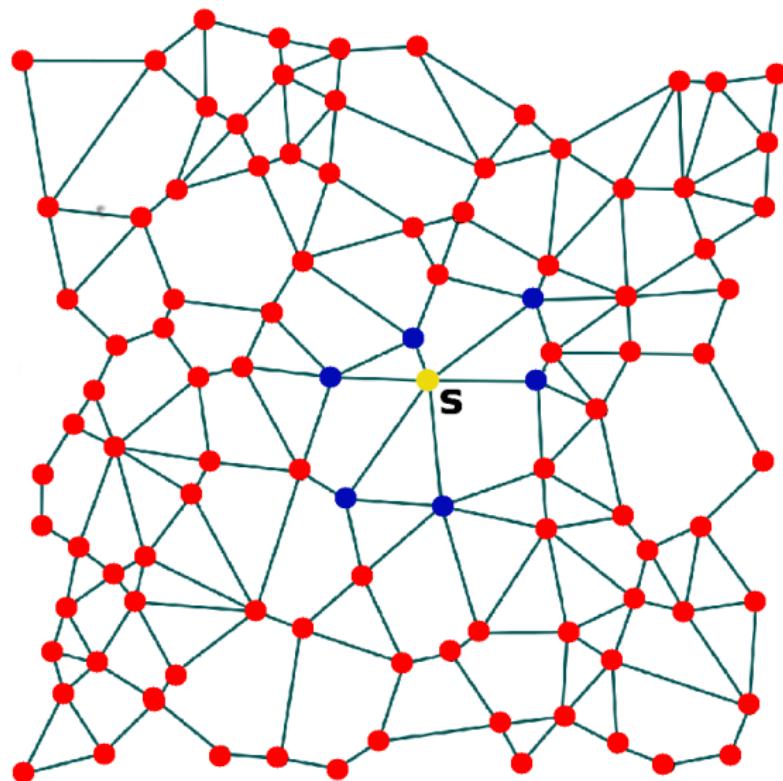
# Probabilistic forwarding

---



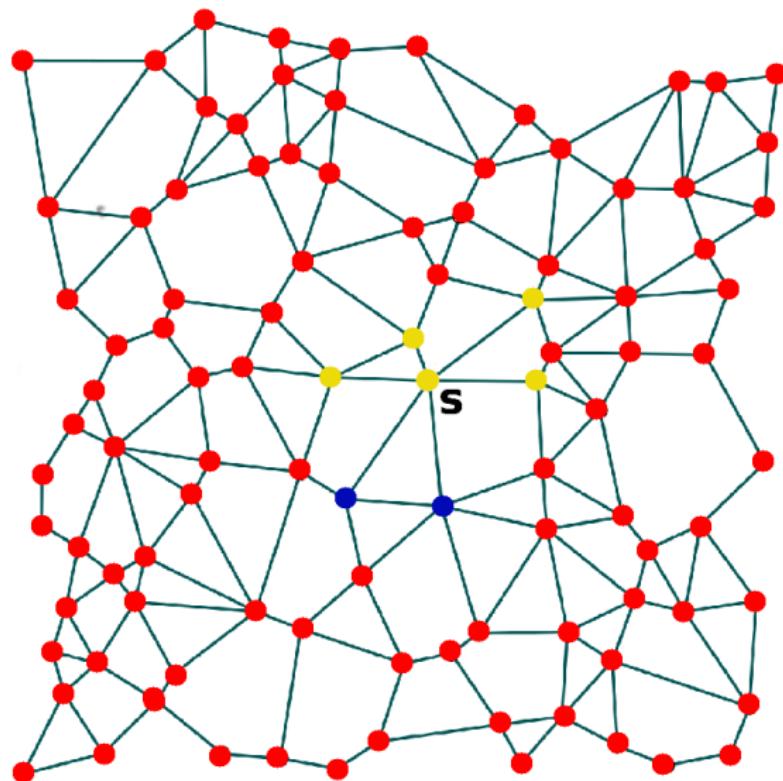
# Probabilistic forwarding

---



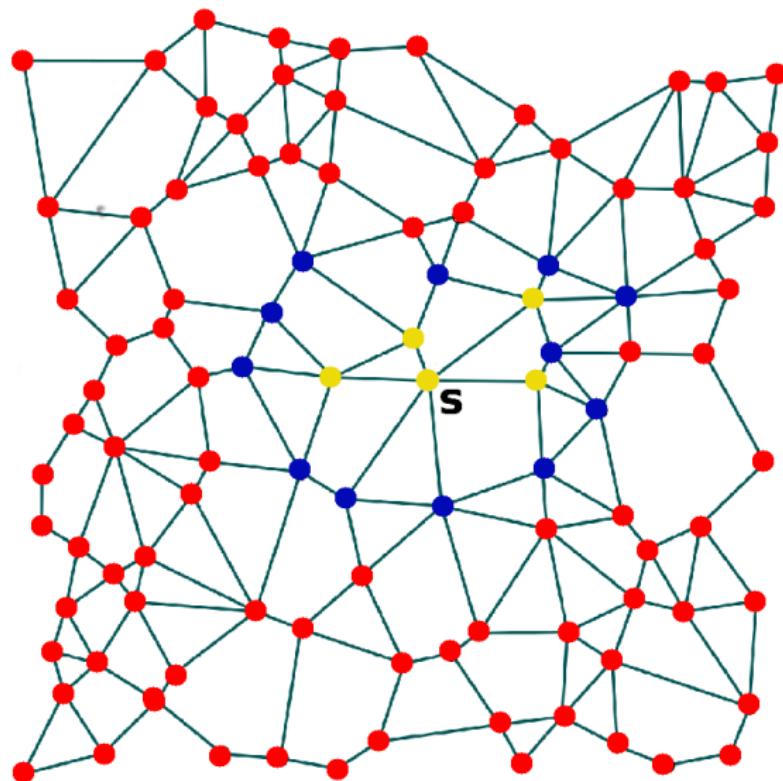
# Probabilistic forwarding

---



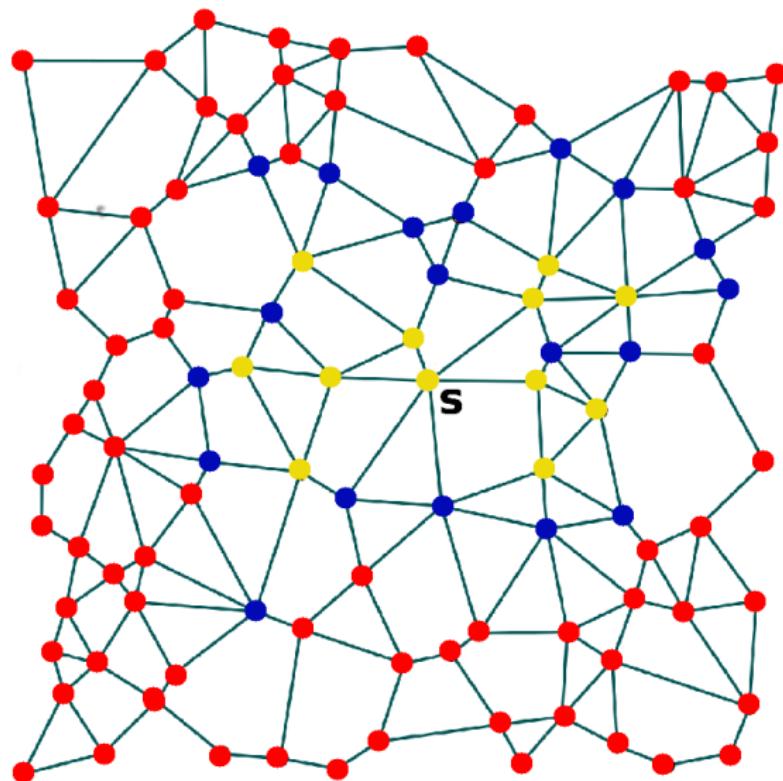
# Probabilistic forwarding

---



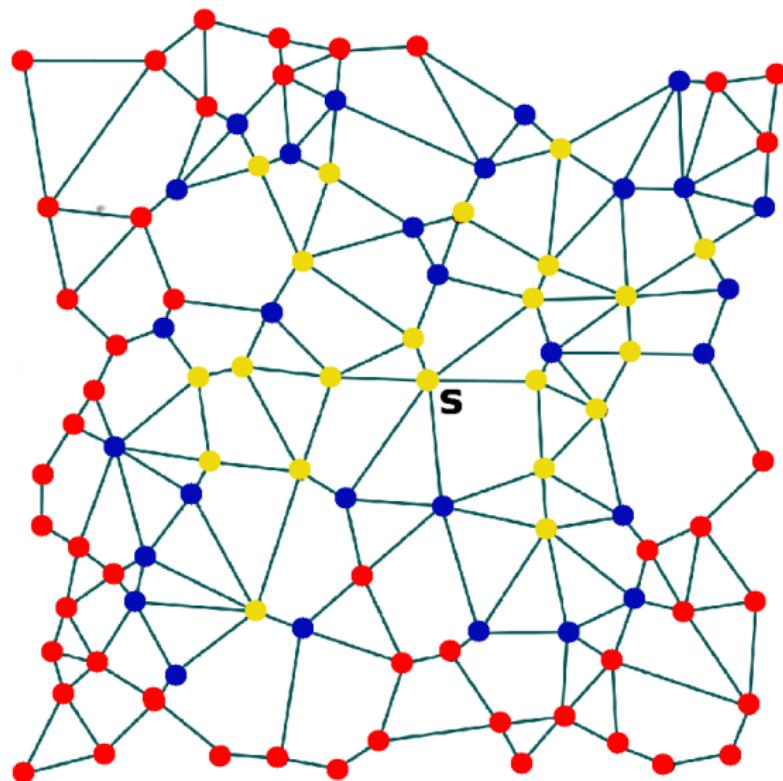
# Probabilistic forwarding

---



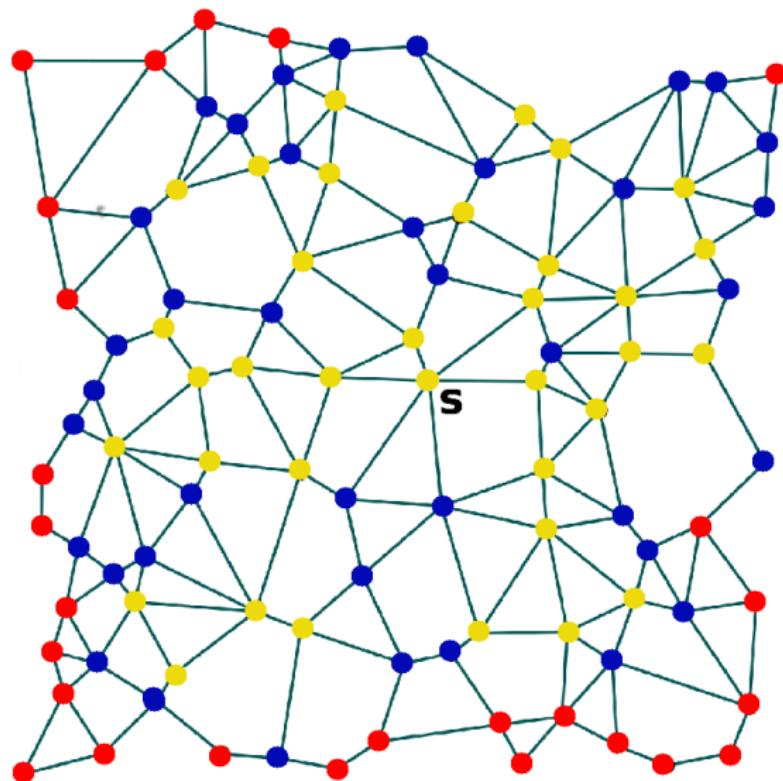
# Probabilistic forwarding

---



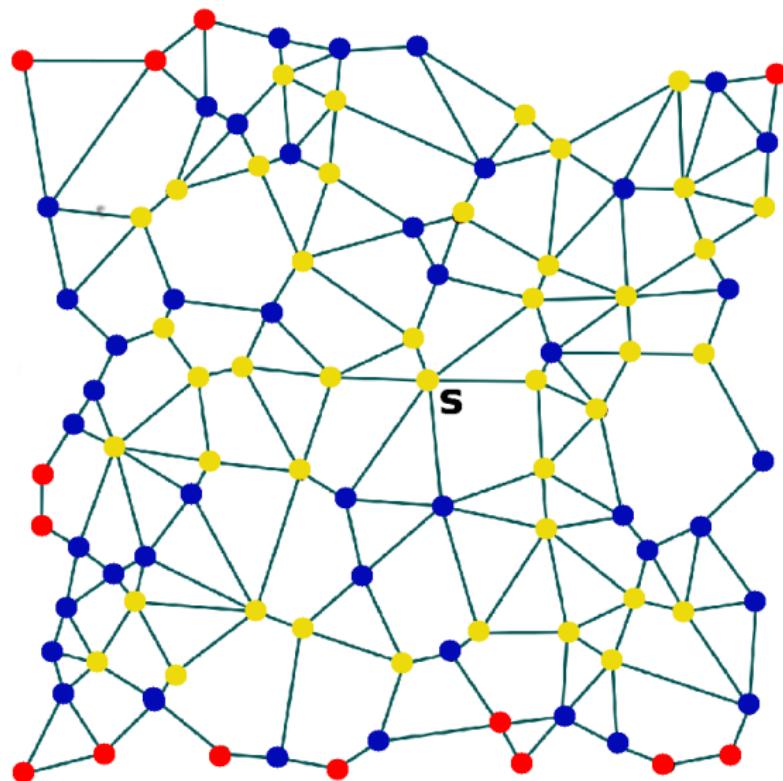
# Probabilistic forwarding

---



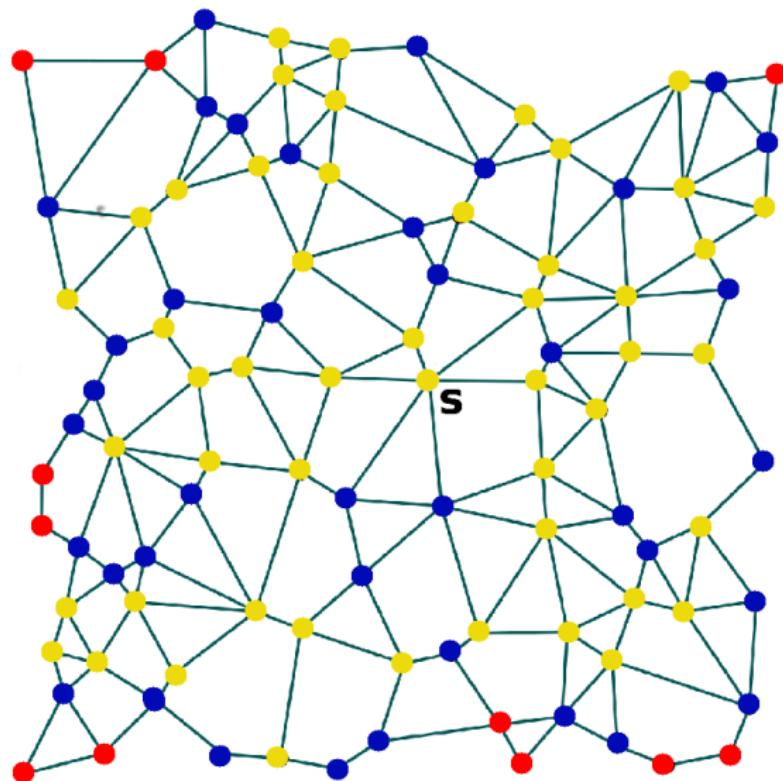
# Probabilistic forwarding

---



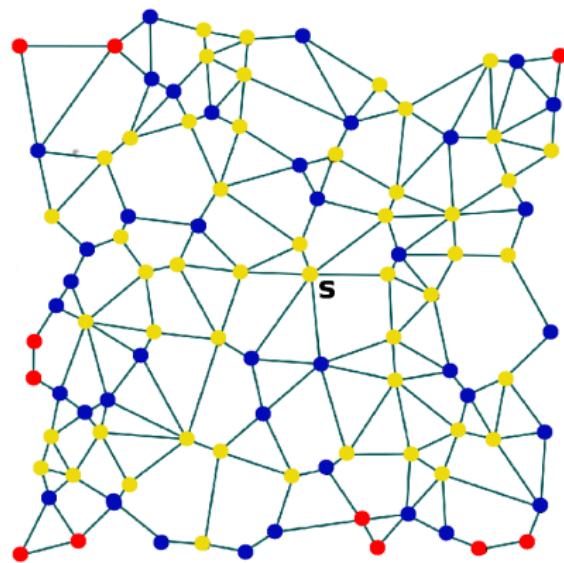
# Probabilistic forwarding

---



# Probabilistic forwarding

---



Lesser number of  
transmissions ( $kNp$ )  
compared to flooding  
( $kN$ ).

BUT

Information lost even if a  
single packet out of the  $k$   
packets is not received.

# Introducing Coded Packets

---

## Coding:

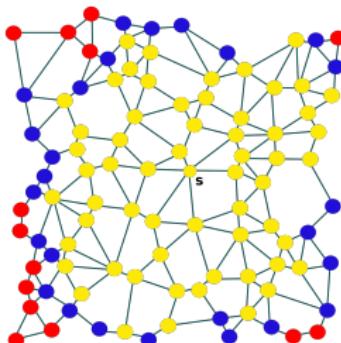
- ▶ The source node encodes the  $k$  data packets into  $n$  coded packets using a Maximum Distance Separable (MDS) code.
- ▶ MDS code ensures that reception of any  $k$  of the  $n$  coded packets by any node, suffices to recover the original  $k$  data packets.

## Probabilistic forwarding of coded packets:

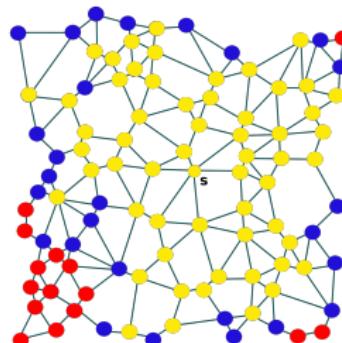
- ▶ Source node transmits all  $n$  coded packets to its one-hop neighbours.
- ▶ Thereafter, the probabilistic forwarding protocol takes over. Each packet is forwarded independently of other packets and other nodes.
- ▶ Nodes which receive at least  $k$  out of  $n$  packets are termed **successful receivers**
- ▶ A **near-broadcast** is when the expected fraction of **successful receivers** is  $\geq 1 - \delta$ .

# Illustration

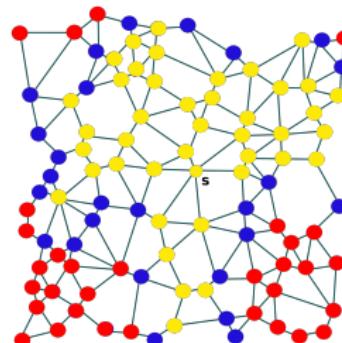
---



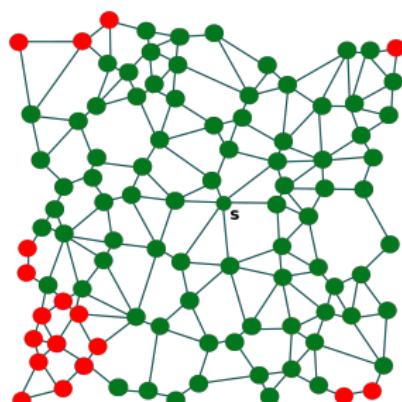
Packet 1



Packet 2



Packet 3



Successful receivers

- ▶  $k = 2, n = 3$
- ▶ ● - Received and transmitted the packet.
- ▶ ● - Received but did not transmit the packet.
- ▶ ● - Did not receive the packet.
- ▶ ● - Successful receivers

# Formal Problem Statement

---

Given:

- ▶ a connected graph  $\mathcal{G}$  with  $N$  nodes
- ▶ number of data packets,  $k$
- ▶ number of coded packets,  $n$
- ▶  $\delta$  close to 0.
- ▶ retransmission probability  $p$

Define

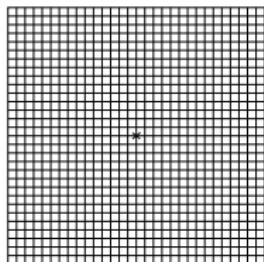
- ▶  $R_{k,n}$  = # nodes that receive at least  $k$  out of  $n$  coded packets

Want to find

- ▶  $p_{k,n,\delta}$  = minimum  $p$  such that  $\mathbb{E}_p\left[\frac{1}{N} R_{k,n}\right] \geq 1 - \delta$
- ▶  $T_{k,n,\delta} = \mathbb{E}_{p_{k,n,\delta}}[\text{total } \# \text{ transmissions over all } N \text{ nodes of } \mathcal{G}]$

On what graphs is coding along with probabilistic forwarding beneficial?

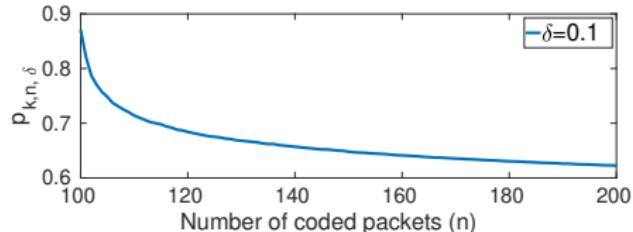
# On well-connected graphs



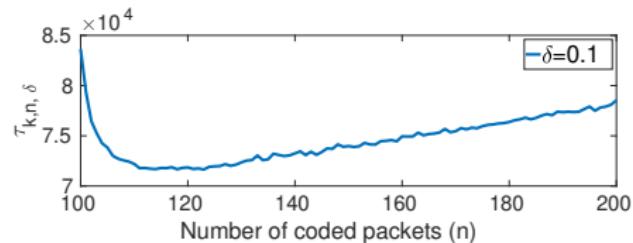
31 × 31 grid

1.  $p_{k,n,\delta}$  decreases (to 0) as  $n$  increases; for fixed  $k$  and  $\delta$ .
2.  $\tau_{k,n,\delta} = \sum_{i=1}^n T_i$ , where  $T_i$  is the expected number of transmissions of packet  $i$ .
  - ▶ Each term  $T_i$  decreases as  $n$  increases since  $p_{k,n,\delta}$  decreases.
  - ▶ The number of terms in the above equation increases with  $n$ .

$k = 100$  packets and  $\delta = 0.1$



Minimum forwarding probability

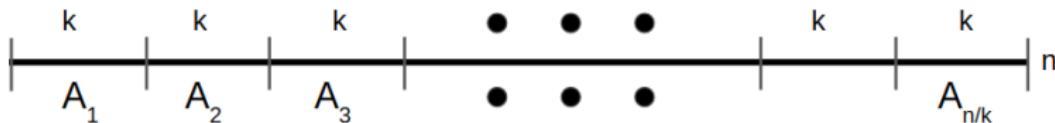


Expected total number of transmissions

## Why $p_{k,n,\delta} \downarrow 0$ ?

$R_{k,n}$  = # nodes that receive at least  $k$  out of  $n$  coded packets.

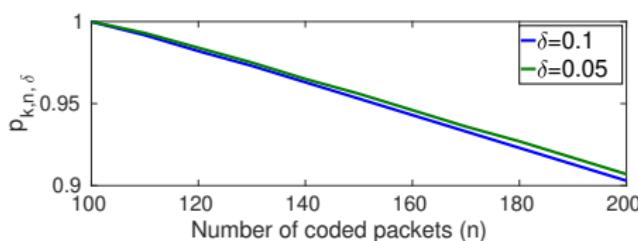
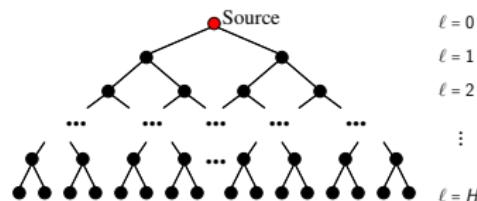
$$p_{k,n,\delta} = \inf\{p \mid \mathbb{E}_p[\frac{1}{N} R_{k,n}] \geq 1 - \delta\}$$



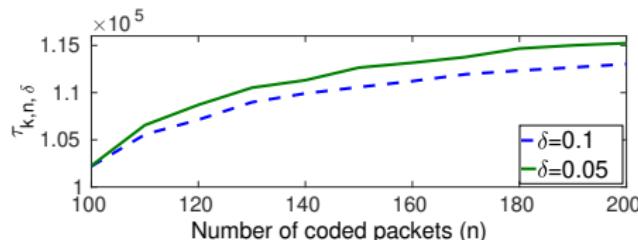
- For  $j = 1, 2, \dots, \lfloor \frac{n}{k} \rfloor$ , let  $A_j$  be the event that the  $j$ th set of  $k$  coded packets is received by at least  $1 - \delta/2$  fraction of the nodes.
- The events  $A_j$  are mutually independent and have the same probability of occurrence.
- For any  $p > 0$ , we have  $P(A_j)$  being small but strictly positive.
- Hence,  $P(\text{at least one } A_j \text{ occurs}) \geq 1 - \delta/2$  for all sufficiently large  $n$ , so that  $\frac{1}{N} R_{k,n} \geq 1 - \delta/2$  with probability at least  $1 - \delta/2$ .
- Thus, for any  $p > 0$ , we have  $p_{k,n,\delta} \leq p$  for sufficiently large  $n$ .

# On trees

Rooted binary tree of height  $H = 10$  with  $k = 100$  and  $\delta = 0.1$



Minimum forwarding probability



Expected total number of transmissions

- ▶ Large fraction of nodes on the leaves
- ▶ Unique path from the source to any node on the tree
- ▶ It can be shown that

$$p_{k,n,\delta} \approx c \left( \frac{k}{n} \right)^{\frac{1}{H-1}}$$

for some constant  $c$   
dependent on  $H$  and  $\delta$ .

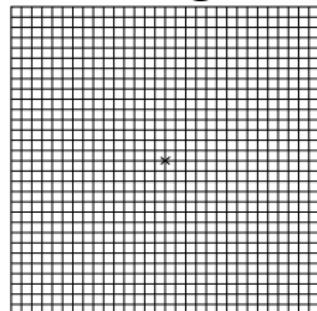
- ▶ This gives

$$\tau_{k,n,\delta} \approx n \frac{(2c)^H \left( \frac{k}{n} \right)^{\frac{H}{H-1}} - 1}{2c \left( \frac{k}{n} \right)^{\frac{1}{H-1}} - 1}$$

which is increasing in  $n$ .

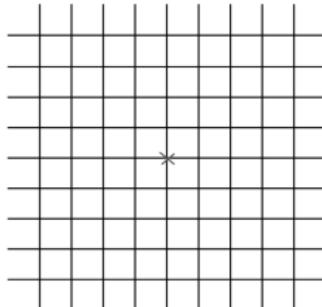
# On the grid

Probabilistic forwarding on  
the  $m \times m$  grid,  $\Gamma_m$

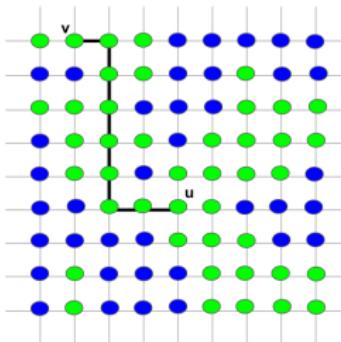


$m \rightarrow \infty$

Probabilistic forwarding on  
the  $\mathbb{Z}^2$  lattice



Site percolation on the  $\mathbb{Z}^2$   
lattice

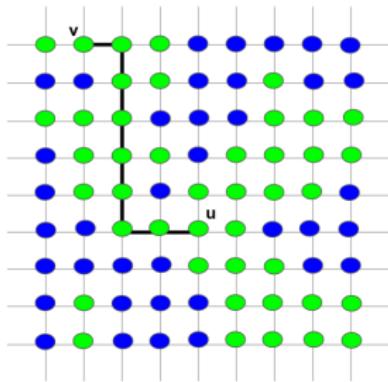


We will see that the site percolation process on  $\mathbb{Z}^2$  is a faithful model for the probabilistic forwarding mechanism on  $\mathbb{Z}^2$ .

# Site percolation on $\mathbb{Z}^2$ - Definitions

---

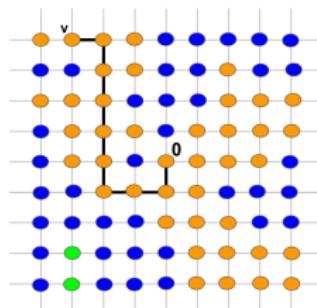
- ▶ Associate each vertex (site)  $u$  of  $\mathbb{Z}^2$  with a Bernoulli( $p$ ) random variable  $X_u$ . The vertex is **open** if  $X_u = 1$ ; else **closed**.
- ▶ Two open sites  $u$  and  $v$  are said to be **connected by an open path** ( $u \longleftrightarrow v$ ), if there is a path of open sites from  $u$  to  $v$ .



## Site percolation on $\mathbb{Z}^2$ - Transmitters

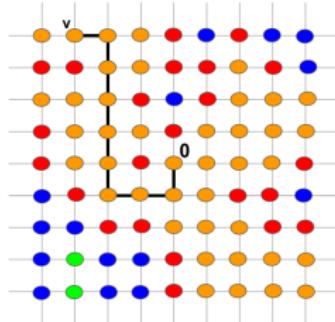
---

- ▶ Associate each vertex (site)  $u$  of  $\mathbb{Z}^2$  with a Bernoulli( $p$ ) random variable  $X_u$ . The vertex is **open** if  $X_u = 1$ ; else **closed**.
- ▶ Two open sites  $u$  and  $v$  are said to be **connected by an open path** ( $u \leftrightarrow v$ ), if there is a path of open sites from  $u$  to  $v$ .
- ▶ Probabilistic forwarding of a single packet over  $\mathbb{Z}^2$  is modelled by site percolation on  $\mathbb{Z}^2$  conditioned on the origin **0** being open.
- ▶ Nodes transmitting the  **$j$ th packet** (for fixed  $j \in [n]$ ) may be viewed as open sites of a site percolation process which are connected to the origin by an open path. Call this cluster of nodes as  $C_j^o$ .
- ▶ The total number of transmissions is simply  $\sum_{j=1}^n |C_j^o|$ .



## Site percolation on $\mathbb{Z}^2$ – Receivers

- ▶ The boundary,  $\partial C_j^o$ , is the set of all closed sites which are adjacent to a site in  $C_j^o$ .
- ▶ The set  $C_j^{o+} := C_j^o \cup \partial C_j^o$  is called the extended cluster of the origin.



For site percolation on  $\mathbb{Z}^2$ , there exists  $p_c \in (0, 1)$  s.t. for  $p > p_c$ ,

- ▶ There exists a unique infinite open cluster (IOC),  $C$ , almost surely.  $p_c \approx 0.59$  for site percolation.
- ▶ Hence, there also exists a unique infinite extended cluster (IEC),  $C^+ = C \cup \partial C$ , a.s..
- ▶  $\theta(p) :=$  percolation probability, i.e.,  $\mathbb{P}(\mathbf{0} \in C)$   
 $\theta^+(p) := \mathbb{P}(\mathbf{0} \in C^+)$

**Lemma:**  $\theta^+(p) = \frac{\theta(p)}{p}$ .

**Proof:**  $\{\mathbf{0} \in C\} = \{\mathbf{0} \in C^+ \text{ and } \mathbf{0} \text{ is open}\}$

# Ergodic theorems

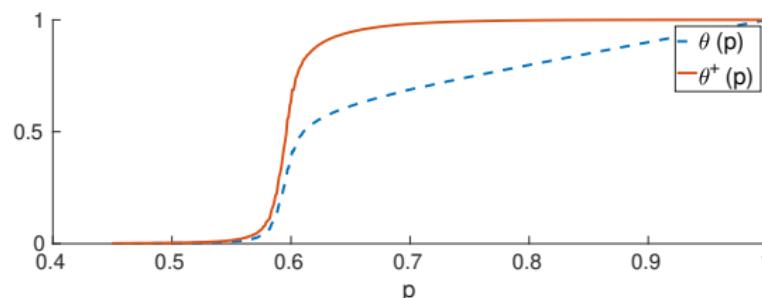
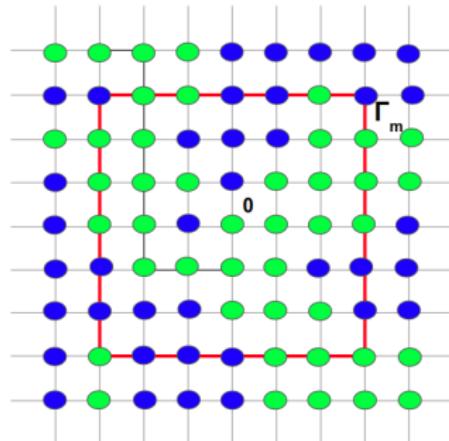
For  $p > p_c$ , if  $C$  is the unique IOC and  $C^+$  is the unique IEC, then,

$$\lim_{m \rightarrow \infty} \frac{|C \cap \Gamma_m|}{m^2} = \theta(p) \quad a.s.$$

Similarly,

$$\lim_{m \rightarrow \infty} \frac{|C^+ \cap \Gamma_m|}{m^2} = \theta^+(p) \quad a.s.$$

Using DCT, expected values also converge.



## Site percolation and probabilistic forwarding

---

- ▶ Prob. forwarding of a single packet over  $\mathbb{Z}^2$  is modelled by site percolation on  $\mathbb{Z}^2$  conditioned on the origin  $\mathbf{0}$  being open.
- ▶  $n$  pkts  $\leftrightarrow$   $n$  independent site percolations with  $\mathbf{0}$  open in all.
- ▶  $\mathcal{R}_{k,n}^{pr.fwd}(\Gamma_m) := \{\text{sites in } \Gamma_m \text{ that receive at least } k \text{ out of } n \text{ pkts}\}$
- ▶ We are interested in finding

$$p_{k,n,\delta} = \min \left\{ p \mid \mathbb{E}_p \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\Gamma_m)| \right] \geq 1 - \delta \right\}$$

# Site percolation and probabilistic forwarding

- ▶ Prob. forwarding of a single packet over  $\mathbb{Z}^2$  is modelled by site percolation on  $\mathbb{Z}^2$  conditioned on the origin  $\mathbf{0}$  being open.
- ▶  $n$  pkts  $\leftrightarrow$   $n$  independent site percolations with  $\mathbf{0}$  open in all.
- ▶  $\mathcal{R}_{k,n}^{pr.fwd}(\Gamma_m) := \{\text{sites in } \Gamma_m \text{ that receive at least } k \text{ out of } n \text{ pkts}\}$
- ▶ We are interested in finding

$$p_{k,n,\delta} = \min \left\{ p \mid \mathbb{E}_p \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\Gamma_m)| \right] \geq 1 - \delta \right\}$$

## Theorem

For  $p > p_c$ , we have

$$\lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\Gamma_m)| \right] = \sum_{t=k}^n \sum_{j=k}^t \binom{n}{t} \binom{t}{j} (\theta^+(p))^{t+j} (1 - \theta^+(p))^{n-j}.$$

# Total transmissions

Consider the transmissions of a single packet on  $\mathbb{Z}^2$

- ▶  $C^\circ :=$  open cluster containing  $\mathbf{0}$  (set of transmitters)
- ▶ The expected number of transmissions on a large grid,  $\Gamma_m$ , is given by  $\mathbb{E}[|C^\circ \cap \Gamma_m| \mid \mathbf{0} \text{ is open}]$ .

## Proposition

For  $p > p_c$ , we have

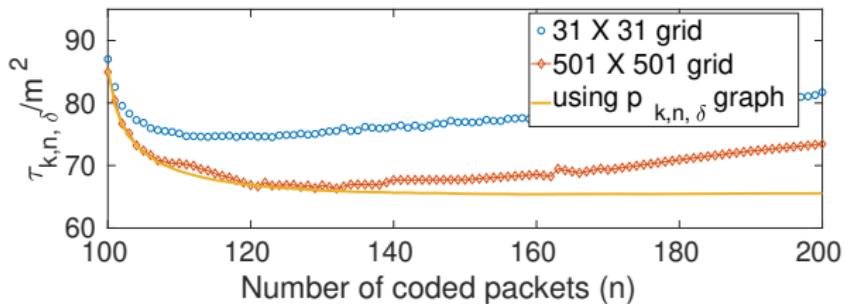
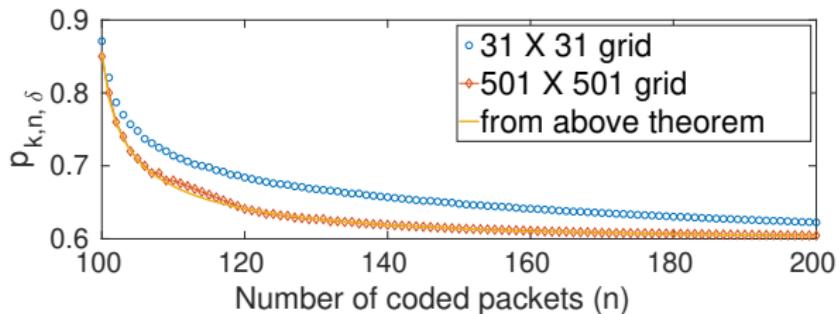
$$\lim_{m \rightarrow \infty} \frac{1}{m^2} \mathbb{E}[|C^\circ \cap \Gamma_m| \mid \mathbf{0} \text{ is open}] = \frac{\theta(p)^2}{p}.$$

- ▶ Hence, for  $n$  coded packets, with each packet being transmitted with probability  $p_{k,n,\delta} > p_c$ , we obtain

$$\tau_{k,n,\delta} \approx nm^2 \frac{\theta^2(p_{k,n,\delta})}{p_{k,n,\delta}}.$$

## Comparison with simulation results

---



# Proof sketch

---

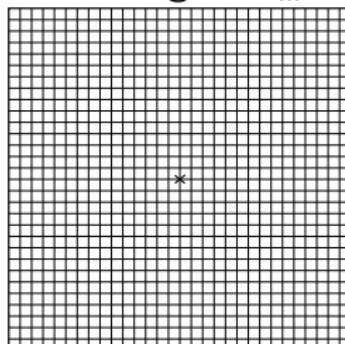
We want to show,

$$\lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\Gamma_m)| \right] = \sum_{t=k}^n \sum_{j=k}^t \binom{n}{t} \binom{t}{j} (\theta^+(p))^{t+j} (1 - \theta^+(p))^{n-j}.$$

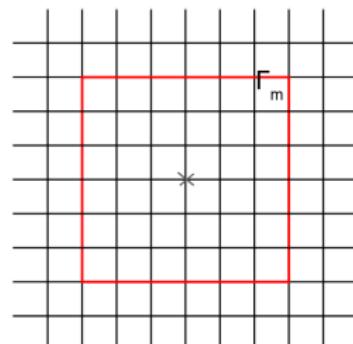
Step 1:  $\mathcal{R}_{k,n}^{pr.fwd}(\mathbb{Z}^2) := \{\text{sites that receive at least } k \text{ out of } n \text{ pkts on } \mathbb{Z}^2\}$

$$\lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\Gamma_m)| \right] = \lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\mathbb{Z}^2) \cap \Gamma_m| \right]$$

**Probabilistic forwarding on the  
 $m \times m$  grid,  $\Gamma_m$**



**Probabilistic forwarding on the  
 $\mathbb{Z}^2$  lattice**



# Proof sketch

---

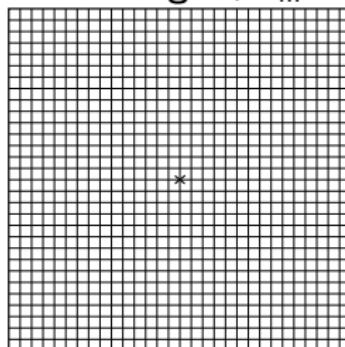
We want to show,

$$\lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\Gamma_m)| \right] = \sum_{t=k}^n \sum_{j=k}^t \binom{n}{t} \binom{t}{j} (\theta^+(p))^{t+j} (1 - \theta^+(p))^{n-j}.$$

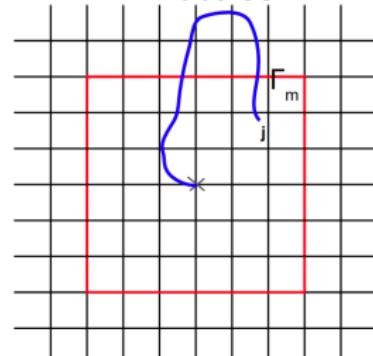
Step 1:  $\mathcal{R}_{k,n}^{pr.fwd}(\mathbb{Z}^2) := \{\text{sites that receive at least } k \text{ out of } n \text{ pkts on } \mathbb{Z}^2\}$

$$\lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\Gamma_m)| \right] = \lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\mathbb{Z}^2) \cap \Gamma_m| \right]$$

**Probabilistic forwarding on the  
 $m \times m$  grid,  $\Gamma_m$**



**Probabilistic forwarding on the  
 $\mathbb{Z}^2$  lattice**



## Proof sketch

---

We want to show,

$$\lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\Gamma_m)| \right] = \sum_{t=k}^n \sum_{j=k}^t \binom{n}{t} \binom{t}{j} (\theta^+(p))^{t+j} (1 - \theta^+(p))^{n-j}.$$

Step 1:  $\mathcal{R}_{k,n}^{pr.fwd}(\mathbb{Z}^2) := \{\text{sites that receive at least } k \text{ out of } n \text{ pkts on } \mathbb{Z}^2\}$

$$\lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\Gamma_m)| \right] = \lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\mathbb{Z}^2) \cap \Gamma_m| \right]$$

Step 2: Let  $\mathcal{R}_{k,n}^{perc}(\mathbb{Z}^2) := \{\text{sites that are in at least } k \text{ out of } n C^+ \text{s}\}$

Using ergodicity, we show that for  $p > p_c$ ,

$$\lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{|\mathcal{R}_{k,n}^{perc}(\mathbb{Z}^2) \cap \Gamma_m|}{m^2} \right] = \underbrace{\sum_{j=k}^n \binom{n}{j} (\theta^+(p))^j (1 - \theta^+(p))^{n-j}}_{\equiv \theta_{k,n}^+(p)}.$$

# Proof sketch

---

We want to show,

$$\lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\Gamma_m)| \right] = \sum_{t=k}^n \sum_{j=k}^t \binom{n}{t} \binom{t}{j} (\theta^+(p))^{t+j} (1 - \theta^+(p))^{n-j}.$$

Step 1:  $\mathcal{R}_{k,n}^{pr.fwd}(\mathbb{Z}^2) := \{\text{sites that receive at least } k \text{ out of } n \text{ pkts on } \mathbb{Z}^2\}$

$$\lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\Gamma_m)| \right] = \lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\mathbb{Z}^2) \cap \Gamma_m| \right]$$

Step 2: Let  $\mathcal{R}_{k,n}^{perc}(\mathbb{Z}^2) := \{\text{sites that are in at least } k \text{ out of } n C^+ \text{s}\}$

Using ergodicity, we show that for  $p > p_c$ ,

$$\lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{|\mathcal{R}_{k,n}^{perc}(\mathbb{Z}^2) \cap \Gamma_m|}{m^2} \right] = \underbrace{\sum_{j=k}^n \binom{n}{j} (\theta^+(p))^j (1 - \theta^+(p))^{n-j}}_{\equiv \theta_{k,n}^+(p)}.$$

Recall:

- ▶  $\theta^+(p) = \mathbb{P}(\mathbf{0} \in C^+)$
- ▶  $\lim_{m \rightarrow \infty} \frac{|C^+ \cap \Gamma_m|}{m^2} = \theta^+(p) \quad a.s.$

# Proof sketch

---

We want to show,

$$\lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\Gamma_m)| \right] = \sum_{t=k}^n \sum_{j=k}^t \binom{n}{t} \binom{t}{j} (\theta^+(p))^{t+j} (1 - \theta^+(p))^{n-j}.$$

Step 1:  $\mathcal{R}_{k,n}^{pr.fwd}(\mathbb{Z}^2) := \{\text{sites that receive at least } k \text{ out of } n \text{ pkts on } \mathbb{Z}^2\}$

$$\lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\Gamma_m)| \right] = \lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\mathbb{Z}^2) \cap \Gamma_m| \right]$$

Step 2: Let  $\mathcal{R}_{k,n}^{perc}(\mathbb{Z}^2) := \{\text{sites that are in at least } k \text{ out of } n C^+ \text{s}\}$

Using ergodicity, we show that for  $p > p_c$ ,

$$\lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{|\mathcal{R}_{k,n}^{perc}(\mathbb{Z}^2) \cap \Gamma_m|}{m^2} \right] = \underbrace{\sum_{j=k}^n \binom{n}{j} (\theta^+(p))^j (1 - \theta^+(p))^{n-j}}_{\equiv \theta_{k,n}^+(p)}.$$

Step 3: Carefully relate  $\mathcal{R}_{k,n}^{pr.fwd}(\mathbb{Z}^2)$  and  $\mathcal{R}_{k,n}^{perc}(\mathbb{Z}^2)$  to obtain,

$$\lim_{m \rightarrow \infty} \mathbb{E} \left[ \frac{1}{m^2} |\mathcal{R}_{k,n}^{pr.fwd}(\mathbb{Z}^2) \cap \Gamma_m| \right] = \sum_{t=k}^n \sum_{\substack{T \subseteq [n]: \\ |T|=t}} \theta_{k,t}^+(p) (\theta^+(p))^t (1 - \theta^+(p))^{n-t}.$$

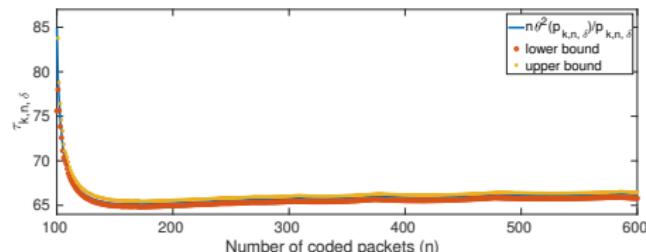
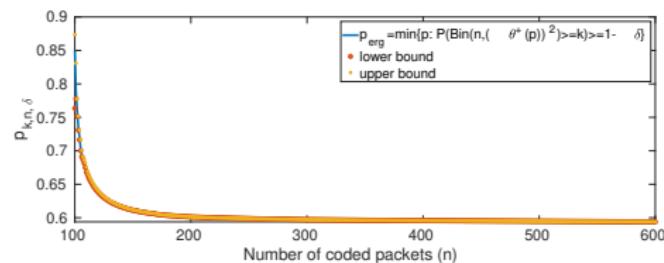
# Some bounds

It can be shown that

$$\sum_{t=k}^n \sum_{j=k}^t \binom{n}{t} \binom{t}{j} (\theta^+(p))^{t+j} (1 - \theta^+(p))^{n-j} = \Pr(Y \geq k)$$

where  $Y \sim \text{Bin}(n, (\theta^+(p))^2)$ . So, for large grids, we have

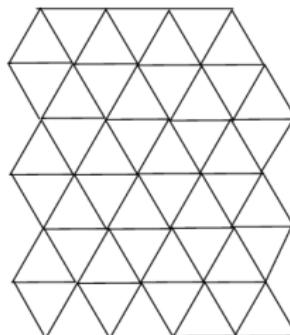
$$p_{k,n,\delta} = \inf\{p : \Pr(Y \geq k) \geq 1 - \delta\} \text{ and } \tau_{k,n,\delta} \approx nm^2 \frac{\theta^2(p_{k,n,\delta})}{p_{k,n,\delta}}.$$



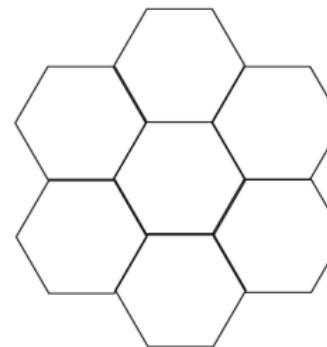
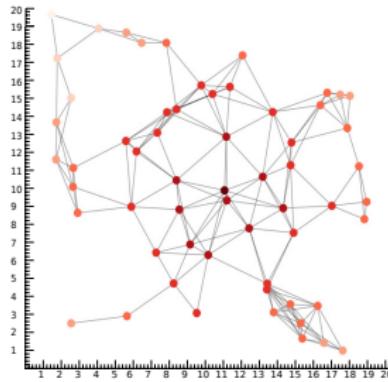
## Other graphs

---

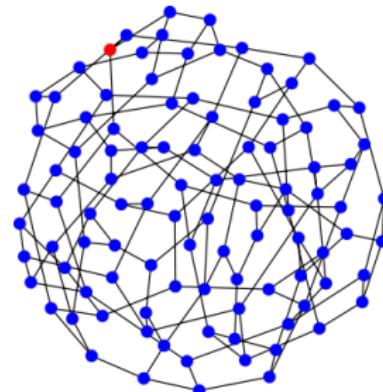
Lattices- triangular, hexagonal.



Random geometric graphs



Random regular graphs



Thank you  
😊