$$E(X) = G'(1) Var(X) = G''(1) + G'(1) - (G'(1))^{2}$$

Bernoulli
$$F(x) = p^x (1-p)^{1-x}$$
 $G(x) = qs^0 + ps^1$

Binomial
$$p(k) = \binom{n}{k} p^k (1-p)^{n-k}$$
 $G(x) = (q+ps)^n$

Geometric
$$p(k) = (1-p)^{k-1} p$$

$$G(x) = \frac{p}{1-qs}$$

Negative Binomial
$$p(k) = \binom{n-1}{k-1} p^r (1-p)^{k-r}$$
 $G(x) = \left(\frac{p}{1-qs}\right)^r$

Poisson
$$p(k) = \frac{\lambda^k e^{-\lambda}}{k!}$$
 $G(x) = e^{\lambda(s-1)}$

Uniform
$$f(x) = \frac{1}{h-a}$$
, $a \le x \le b$

Exponential
$$f(x) = \lambda e^{-\lambda x}, x \ge 0$$

Gamma
$$g(t) = \frac{\lambda^{\alpha} t^{\alpha-1} e^{-\lambda x}}{\Gamma(\alpha)}$$
 for $\alpha = 1$, it tends to exponential distribution

Normal
$$\frac{1}{\sqrt{2\pi\sigma^2}} e^{\frac{-(x-\mu)}{2\sigma^2}}$$

Let us assume that X1, X2, ..., Xn are independent random variables with the common cdf F and density f. Let U denote the maximum of the X_i and V the minimum.

$$F_{U}(u) = [F(u)]^{n}$$
 $f_{U}(u) = nf(u)[F(u)]^{n-1}$

$$F_V(v) = 1 - [1 - F(v)]^n$$
 $f_V(v) = nf(v)[1 - F(v)]^{n-1}$

$$k^{th}$$
 order stats
$$f_k(x) = \frac{n!}{(k-1)! (n-k)!} f(x) F^{k-1}(x) [1 - F(x)]^{n-k}$$

If there are p experiments and the first has n_1 possible outcomes, the second $n_2, ...,$ and the pth n_p possible outcomes, then there are a total of $n_1 \times n_2 \times \cdots \times n_p$ possible outcomes for the p experiments

For a set of size n and a sample of size r, there are n^r different ordered samples with replacement and $n(n-1)(n-2)\cdots(n-r+1)$ different ordered samples without replacement.

The number of unordered samples of r objects selected from n objects without replacement is $\binom{n}{r}$.

The number of ways that n objects can be grouped into r classes with n_i in the i^{th} class is

$$\binom{n}{n_1 \; n_2 \ldots n_r} = \frac{n!}{n_1! \; n_2! \ldots n_r!}$$

$$CDF = \int_{-\infty}^{x} PDF$$
 $CDF = F_{X}(x) = P(X \le x)$

If
$$X \sim N(\mu, \sigma^2)$$
 and $Y = aX + b$, then $Y \sim N(a\mu + b, a^2\sigma^2)$.

Let X have
$$f(x)$$
 and let $Y = g(X)$
$$f_Y(y) = f_X(g^{-1}(y)) \left| \frac{d}{dy} g^{-1}(y) \right|$$

PROPOSITION C Let
$$Z = F(X)$$
; then Z has a uniform distribution on [0,1].

PROPOSITION D Let U be uniform on [0,1], and let
$$X = F - 1(U)$$
. Then the cdf of X is F.

$$E(X) = \sum_{i} x_{i} p(x_{i})$$
 or $E(X) = \int_{-\infty}^{\infty} x f(x) dx$ Suppose that $Y = g(X)$, replace $x = g(x)$

THEOREM A Markov's Inequality

If X is a random variable with $P(X \ge 0) = 1$ and for which E(X) exists, then $P(X \ge t) \le \frac{E(X)}{t}$

If $X_1, ..., X_n$ are jointly distributed random variables with expectations $E(X_i)$ and Y is a linear function

of the
$$X_i, Y = a + \sum_{i=1}^n b_i E(X_i)$$
, then $E(Y) = a + \sum_{i=1}^n b_i E(X_i)$

$$Var(X) = E(X^2) - [E(X)]^2$$
 or $Var(X) = E\{[X - E(X)]^2\}$

If Var(X) exists and Y = a + bX, then $Var(Y) = b^2Var(X)$.

THEOREM C Chebyshev's Inequality

Let X be a random variable with mean μ and variance σ^2 . Then, for any t > 0,

$$P(|X - \mu| > t) \le \frac{\sigma^2}{t^2}$$

COROLLARY A If
$$Var(X) = 0$$
, then $P(X = \mu) = 1$.

A Model for Measurement Error

$$X = x_0 + \beta + \varepsilon$$
 $E(X) = x_0 + \beta$ and $Var(X) = \sigma^2$ β is bias

$$MSE = E[(X - x_0)^2] = \beta^2 + \sigma^2$$

$$Cov(X,Y) = E(XY) - E(X)E(Y)$$

$$Cov(a + X, Y) = Cov(X, Y)$$

$$Cov(aX, bY) = abCov(X, Y)$$

$$Cov(X,Y + Z) = Cov(X,Y) + Cov(X,Z)$$

$$Cov(aW + bX, cY + dZ) = acCov(W, Y) + bcCov(X, Y) + adCov(W, Z) + bdCov(X, Z)$$

$$U = a + \sum_{i=1}^{n} b_i X_i$$
 and $V = c + \sum_{j=1}^{n} d_j Y_j$, then $Cov(U, V) = \sum_{\substack{1 \le i \le n \\ 1 \le j \le n}} b_i d_j Cov(U, V)$

$$Var\left(a + \sum_{i=1}^{n} b_i X_i\right) = \sum_{\substack{1 \le i \le n \\ 1 \le i \le n}} b_i b_j Cov(X_i, X_j)$$

$$Var\left(\sum_{i=1}^{n} X_i\right) = \sum_{i=1}^{n} Var(X_i)$$
 if X_i are independent

Correlation
$$\rho = \frac{Cov(X,Y)}{\sqrt{Var(X)Var(Y)}}$$

 $-1 \le \rho \le 1$. Furthermore, $\rho = \pm 1$ if and only if P(Y = a + bX) = 1 for some constants a and b.

$$\binom{n}{k} = \binom{n-1}{r-1} + \binom{n-1}{r} \\ |A_1 \cup A_2 \cup \dots \cup A_p| = \sum_{1 \le i \le p} |A_i| - \sum_{1 \le i_1 < i_2 \le p} |A_{i_1} \cap A_{i_2}| +$$

$$\sum_{1\leq i_1< i_2< i_3\leq p} \left|A_{i_1}\cap A_{i_2}\cap A_{i_3}\right|-\cdots+(-1)^{p-1}\left|A_1\cap A_2\cap\cdots\cap A_p\right|,$$

Inclusion Exclusion

Derangement
$$n! \left(1 - \frac{1}{1!} + \frac{1}{2!} - \frac{1}{2!} + \dots + (-1)^n \frac{1}{n!}\right) = n! \sum_{r=0}^n (-1)^r \frac{1}{n!}$$

Law of large numbers:
$$-\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$$
 $E[\bar{X}_n] = \mu$ $Var[\bar{X}_n] = \frac{\sigma^2}{n}$
$$\hat{I}(f) = \frac{1}{n} \sum_{i=1}^n f(X_i)$$

Central limit theorem: - MGF =
$$\int_{-\infty}^{\infty} e^{tx} f(x) dx$$
 $P(a > Y > b) = \emptyset\left(\frac{Y - b}{\sigma}\right) - \emptyset\left(\frac{Y - a}{\sigma}\right)$
 $E[Y] = nE[X]$ $Var[Y] = n^2 Var(X)$

Continuity Correction: -
$$P(A) = P\left(l - \frac{1}{2} \le Y \le u + \frac{1}{2}\right)$$

Chi Square Distribution: - $\chi_n = U_1 + U_2 + \dots + U_n$ where $U = Z^2$, Z is Standard Normal Distribution

It is gamma distribution
$$\frac{\lambda^{\alpha} t^{\alpha-1} e^{-\lambda x}}{\Gamma(\alpha)}$$
 with $\alpha = \frac{n}{2}$ and $\lambda = \frac{1}{2}$ for $DOF = n$ $MGF = (1 - 2t)^{-\frac{n}{2}}$

t Distribution:
$$-\frac{Z}{\sqrt{\frac{U}{n}}} \qquad f(x) = \frac{\frac{\Gamma(n+1)}{2}}{\sqrt{n} \pi \Gamma(\frac{n}{2})} \left(1 + \frac{x^2}{n}\right)^{-\frac{n+1}{2}} \qquad f(x) = f(-x)$$

Sample Mean
$$\bar{X} = \sum_{i=1}^{n} X_i$$
 Sample Variance $s^2 = \frac{1}{n-1} \sum_{i=1}^{n} (X_i - \bar{X})^2$ or $s^2 = \frac{1}{n-1} \sum_{i=1}^{n} (X_i^2 - n \bar{X}^2)$

$$\frac{(n-1)s^2}{\sigma^2} \longrightarrow chi \ square \ with \ (n-1) \ DOF \qquad \frac{\bar{X}-\mu}{\frac{S}{\sqrt{n}}} \longrightarrow t \ distribution \ with \ t_{n-1}$$

For Simple Random Sampling

Population Parameter	Estimate	Variance of Estimate	Estimated Variance
μ	$\overline{X} = \frac{1}{n} \sum_{i=1}^{n} X_i$	$\sigma_{\overline{X}}^2 = \frac{\sigma^2}{n} \left(\frac{N-n}{N-1} \right)$	$s_{\overline{X}}^2 = \frac{s^2}{n} \left(1 - \frac{n}{N} \right)$
p	$\hat{p} = \text{sample proportion}$	$\sigma_{\hat{p}}^2 = \frac{p(1-p)}{n} \left(\frac{N-n}{N-1} \right)$	$s_{\hat{p}}^2 = \frac{\hat{p}(1-\hat{p})}{n-1} \left(1 - \frac{n}{N}\right)$
τ	$T = N\overline{X}$	$\sigma_T^2 = N^2 \sigma_{\overline{X}}^2$	$s_T^2 = N^2 s_{\overline{X}}^2$
σ^2	$\left(1-\frac{1}{N}\right)s^2$		

where
$$s^2 = \frac{1}{n-1} \sum_{i=1}^{n} (X_i - \overline{X})^2$$
.

$$\mu = \frac{1}{N} \sum_{i=1}^{n} x_i$$
 $\tau = N\mu$ $\sigma^2 = s^2$ $Var(\bar{X})$ with replacement $= \frac{\sigma^2}{n}$

For Stratified Random Sampling

L stratas have pop elements $l=N_l$, $N=N_1+N_2+\cdots+N_l$, population mean and variance μ_l and σ_l^2

$$W_l = \frac{N_l}{N}$$
 , $x_{il} \rightarrow i^{th} pop in l^{th} stratum$, $\mu = \sum_{l=1}^{L} W_l \mu_l$

Sample mean $\bar{X}_l = \frac{1}{n_l} \sum_{i=1}^l X_{il}$ Startified estimate $\bar{X}_s = \sum_{l=1}^L W_l \, \bar{X}_l$ Population total $T_s = N \, \bar{X}_s$

Variance of startified mean $Var(\bar{X}_s) = \sum_{l=1}^L W_l^2 \left(\frac{1}{n_l}\right) \left(1 - \frac{n_l - 1}{N_l - 1}\right) \sigma_l^2$

$$s_l^2 = \frac{1}{n_l - 1} \ \sum_{i=1}^{n_l} (X_{il} - \bar{X}_l)^2 \quad s_{\bar{X}_s}^2 = \sum_{l=1}^L W_l^2 \ \left(\frac{1}{n_l}\right) \left(1 - \frac{n_l}{N_l}\right) \, s_l^2$$

Neyman Allocation $n_l=n \; \frac{W_l \, \sigma_l}{\sum_{k=1}^L W_k \, \sigma_k}$ proportional allocation $n_l=n \, \frac{N_l}{N}$

$$Var(\bar{X}_{sp}) = \frac{1}{n} \left(\sum_{l=1}^{L} W_l \sigma_l^2 \right)^2$$

$$Var(\bar{X}_{so}) = \frac{1}{n} \sum_{l=1}^{L} W_l \sigma$$

$$Var(\bar{X}_{sp}) = \frac{1}{n} \left(\sum_{l=1}^{L} W_l \, \sigma_l^2 \right)^2 \qquad Var(\bar{X}_{so}) = \frac{1}{n} \sum_{l=1}^{L} W_l \, \sigma_l^2 \qquad Var(\bar{X}_{sp}) - Var(\bar{X}_{so}) = \sum_{l=1}^{L} W_l \, (\sigma_l - \bar{\sigma})^2$$

$$MLE\ l(\theta) = \sum_{i=1}^{n} \log(f(X_i|\theta))$$

MLE $l(\theta) = \sum_{i=1}^{n} \log(f(X_i|\theta))$ MLE multinomial Cell Probabilities $f(x_i \mid p_m) = \frac{n!}{\sum x_m!} \sum_{i=1}^{n} (p_m)^{x_m}$

Confidence intervals for MLE

$$P\left(\bar{X} - \frac{S}{\sqrt{n}} t_{n-1} \left(\frac{\alpha}{2}\right) \le \mu \le \bar{X} + \frac{S}{\sqrt{n}} t_{n-1} \left(\frac{\alpha}{2}\right)\right) = 1 - \alpha$$

$$P\left(\frac{n\,\widehat{\sigma}^2}{\chi^2_{n-1}\left(\frac{\alpha}{2}\right)} \leq \sigma^2 \leq \frac{n\,\widehat{\sigma}^2}{\chi^2_{n-1}\left(\frac{1-\alpha}{2}\right)}\right) = 1 - \alpha$$

Hypothesis Testing

Null Hypothesis (H_0) Alternate Hypothesis (H_A)

 $P(reject H_0 | H_0) = P(X > c | H_0) \longrightarrow Type I \ error(\alpha) \ Significance \ Level = \alpha$

 $P(accept H_0 | H_1) = p(X \le c | H_1) \rightarrow Type II \ error(\beta) \ Power = 1 - \beta$

$$Z(\alpha) = \frac{x_0 - \mu_0}{\frac{\sigma}{\sqrt{n}}} \qquad Z(\alpha) = \frac{\bar{X} - \mu_0}{\frac{S}{\sqrt{n}}}$$

$$Z(\alpha) = \frac{\bar{X} - \mu_0}{\frac{S}{\sqrt{n}}}$$

Likelihood Ratio Tests

$$\widehat{p}_i = \frac{x_i}{n}$$
 $\Lambda = \prod_{i=1}^m$

$$\widehat{p_i} = \frac{x_i}{n} \qquad \Lambda = \prod_{i=1}^m \left(\frac{P_i(\widehat{\theta})}{\widehat{P_i}}\right)^{x_i} \qquad x_i = observed \ count \ in \ m \ cells, \ \widehat{\theta} = MLE \ of \ \theta$$

 $p_i(\hat{\theta}) = corresponding probabilities$

$$-2\log\Lambda = 2\sum_{i=1}^{m} O_{i}\log\left(\frac{O_{i}}{E_{i}}\right)$$

$$O_i = n \, \widehat{p}_i$$
 and $E_i = n \, p_i \, (\widehat{\theta})$

Pearson's chi – square statistic $\chi^2 = \sum_{n \in \mathbb{N}} \frac{\left[x_i - n p_i(\hat{\theta})\right]^2}{n n (\hat{\theta})}$