Fast Mean Estimation with Sub-Gaussian Rates

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1 Introduction

1.1 Goal

To obtain high probability mean estimates when only the existence of the 2^{nd} moment is known. This is also called the *heavy tailed* setting, where higher order moments from the sampling distribution need not exist.

1.2 Existing results

Consider the estimator to be the sample mean $\hat{\theta} = \frac{1}{n} \sum_{i=1}^{n} X_i$ where $\{X_i\}_{i=1}^n$ are sampled from a distribution P with only finite 2^{nd} moment and mean θ^* . Markov's inequality gives:

$$\Pr(||\widehat{\theta} - \theta^{\star}||_2 > t) \le \frac{\mathbb{E}[||\widehat{\theta} - \theta^{\star}||_2^2]}{t^2}$$

Note that $\widehat{\theta} - \theta^* = \frac{1}{n} \sum_{i=1}^n (X_i - \theta^*)$ and hence:

$$||\widehat{\theta} - \theta^{\star}||_{2}^{2} = \frac{1}{n^{2}} \sum_{i=1}^{n} ||X_{i} - \theta^{\star}||_{2}^{2} + \frac{1}{n} \sum_{\substack{i,j=1\\i \neq j}}^{n} (X_{i} - \theta^{\star})^{T} (X_{j} - \theta^{\star}) \Rightarrow \mathbb{E}[||\widehat{\theta} - \theta^{\star}||_{2}^{2}] = \frac{1}{n^{2}} \sum_{i=1}^{n} \mathbb{E}[||X_{i} - \theta^{\star}||_{2}^{2}]$$

and since $\mathbb{E}\left[||X_i - \theta^*||_2^2\right] = \mathbb{E}\left[\operatorname{trace}(X_i - \theta^*)(X_i - \theta^*)^T\right] = \Sigma$, we get:

$$\mathbb{E}[||\widehat{\theta} - \theta^*||_2^2] = \frac{\operatorname{trace}(\Sigma)}{n}$$

therefore leading to:

$$\Pr\left(||\widehat{\theta} - \theta^{\star}||_2 > \sqrt{\frac{\operatorname{trace}(\Sigma)}{n\delta}}\right) \leq \delta$$

which corresponds to: with probability at least $1 - \delta$:

$$||\widehat{\theta} - \theta^*||_2 \le \sqrt{\frac{\operatorname{trace}(\Sigma)}{n\delta}}$$

In contrast, when P is Gaussian, we get:

$$\Pr\left(||\widehat{\theta} - \theta^{\star}||_2 > O\left(\sqrt{\frac{\operatorname{trace}(\Sigma)}{n}} + \sqrt{\frac{||\Sigma||_2 \log(1/\delta)}{n}}\right)\right) \le \delta$$

To show this, consider $Z_i = X_i - \theta^*$ for all $i \in [n]$. Then $||\widehat{\theta} - \theta^*||_2 = \left\|\frac{1}{n}\sum_{i=1}^n Z_i\right\|_2$, where Z_i s are zero mean Gaussian RVs with covariance Σ . Note that $Z_i = \Sigma^{1/2}Y_i$ for all $i \in [n]$ where Y_i s are standard multivariate Gaussian RVs. Now, we have that:

$$|||Z_i - Z_i'||_2 \le ||Z_i - Z_i'||_2 \le ||\Sigma^{1/2}(Y_i - Y_i')||_2 \le ||\Sigma^{1/2}||_2 ||Y_i - Y_i'||_2$$

which shows that $||Z_i||$ is a $||\Sigma^{1/2}||_2$ -Lipschitz function of Y_i . By a Lipschitz concentration lemma due to Tsirelson, Ibragimov and Sudakov, we have:

$$\Pr\left(\left\|\frac{1}{n}\sum_{i=1}^{n}Z_{i}\right\|_{2}-\mathbb{E}\left[\left\|\frac{1}{n}\sum_{i=1}^{n}Z_{i}\right\|_{2}\right]>t\right)\leq\exp\left(-\frac{nt^{2}}{2||\Sigma||_{2}}\right)$$

leading to:

$$\Pr\left(\left\|\frac{1}{n}\sum_{i=1}^{n}Z_{i}\right\|_{2} > \mathbb{E}\left[\left\|\frac{1}{n}\sum_{i=1}^{n}Z_{i}\right\|_{2}\right] + t\right) \leq \exp\left(-\frac{nt^{2}}{2||\Sigma||_{2}}\right)$$

and with probability at least $1 - \delta$:

$$\begin{split} ||\widehat{\theta} - \theta^{\star}||_2 &\leq \mathbb{E}[||\widehat{\theta} - \theta^{\star}||_2] + \sqrt{\frac{2||\Sigma||_2 \log(1/\delta)}{n}} \leq \sqrt{\mathbb{E}[||\widehat{\theta} - \theta^{\star}||_2^2]} + \sqrt{\frac{2||\Sigma||_2 \log(1/\delta)}{n}} \\ &\leq \sqrt{\frac{\operatorname{trace}(\Sigma)}{n}} + \sqrt{\frac{2||\Sigma||_2 \log(1/\delta)}{n}} \end{split}$$

Lugosi and Mendelson showed that with only bounded 2^{nd} moment, this rate can be achieved, but the estimator proposed is intractable.