

Fast Mean Estimation with Sub-Gaussian Rates

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1 Introduction

1.1 Goal

To obtain high probability mean estimates when only the existence of the 2^{nd} moment is known. This is also called the *heavy tailed* setting, where higher order moments from the sampling distribution need not exist.

1.2 Existing results

Consider the estimator to be the sample mean $\hat{\theta} = \frac{1}{n} \sum_{i=1}^n X_i$ where $\{X_i\}_{i=1}^n$ are sampled from a distribution P with only finite 2^{nd} moment and mean θ^* . Markov's inequality gives:

$$\Pr(\|\hat{\theta} - \theta^*\|_2 > t) \leq \frac{\mathbb{E}[\|\hat{\theta} - \theta^*\|_2^2]}{t^2}$$

Note that $\hat{\theta} - \theta^* = \frac{1}{n} \sum_{i=1}^n (X_i - \theta^*)$ and hence:

$$\|\hat{\theta} - \theta^*\|_2^2 = \frac{1}{n^2} \sum_{i=1}^n \|X_i - \theta^*\|_2^2 + \frac{1}{n} \sum_{\substack{i,j=1 \\ i \neq j}}^n (X_i - \theta^*)^T (X_j - \theta^*) \Rightarrow \mathbb{E}[\|\hat{\theta} - \theta^*\|_2^2] = \frac{1}{n^2} \sum_{i=1}^n \mathbb{E}[\|X_i - \theta^*\|_2^2]$$

and since $\mathbb{E}[\|X_i - \theta^*\|_2^2] = \mathbb{E}[\text{trace}(X_i - \theta^*)(X_i - \theta^*)^T] = \Sigma$, we get:

$$\mathbb{E}[\|\hat{\theta} - \theta^*\|_2^2] = \frac{\text{trace}(\Sigma)}{n}$$

therefore leading to:

$$\Pr\left(\|\hat{\theta} - \theta^*\|_2 > \sqrt{\frac{\text{trace}(\Sigma)}{n\delta}}\right) \leq \delta$$

which corresponds to: with probability at least $1 - \delta$:

$$\|\hat{\theta} - \theta^*\|_2 \leq \sqrt{\frac{\text{trace}(\Sigma)}{n\delta}}$$

In contrast, when P is Gaussian, we get:

$$\Pr\left(\|\hat{\theta} - \theta^*\|_2 > O\left(\sqrt{\frac{\text{trace}(\Sigma)}{n}} + \sqrt{\frac{\|\Sigma\|_2 \log(1/\delta)}{n}}\right)\right) \leq \delta$$

We will denote $\sqrt{\frac{\text{trace}(\Sigma)}{n}} + \sqrt{\frac{\|\Sigma\|_2 \log(1/\delta)}{n}}$ as $\text{OPT}_{n,\delta,\Sigma}$ as a shorthand.

To show this, consider $Z_i = X_i - \theta^*$ for all $i \in [n]$. Then $\|\hat{\theta} - \theta^*\|_2 = \left\| \frac{1}{n} \sum_{i=1}^n Z_i \right\|_2$, where Z_i s are zero mean Gaussian RVs with covariance Σ . Note that $Z_i = \Sigma^{1/2} Y_i$ for all $i \in [n]$ where Y_i s are standard multivariate Gaussian RVs. Now, we have that:

$$|\|Z_i\| - \|Z'_i\||_2 \leq \|Z_i - Z'_i\|_2 \leq \|\Sigma^{1/2}(Y_i - Y'_i)\|_2 \leq \|\Sigma^{1/2}\|_2 \|Y_i - Y'_i\|_2$$

which shows that $\|Z_i\|$ is a $\|\Sigma^{1/2}\|_2$ -Lipschitz function of Y_i . By a Lipschitz concentration lemma due to Tsirelson, Ibragimov and Sudakov, we have:

$$\Pr \left(\left\| \frac{1}{n} \sum_{i=1}^n Z_i \right\|_2 - \mathbb{E} \left[\left\| \frac{1}{n} \sum_{i=1}^n Z_i \right\|_2 \right] > t \right) \leq \exp \left(-\frac{nt^2}{2\|\Sigma\|_2} \right)$$

leading to:

$$\Pr \left(\left\| \frac{1}{n} \sum_{i=1}^n Z_i \right\|_2 > \mathbb{E} \left[\left\| \frac{1}{n} \sum_{i=1}^n Z_i \right\|_2 \right] + t \right) \leq \exp \left(-\frac{nt^2}{2\|\Sigma\|_2} \right)$$

and with probability at least $1 - \delta$:

$$\begin{aligned} \|\hat{\theta} - \theta^*\|_2 &\leq \mathbb{E}[\|\hat{\theta} - \theta^*\|_2] + \sqrt{\frac{2\|\Sigma\|_2 \log(1/\delta)}{n}} \leq \sqrt{\mathbb{E}[\|\hat{\theta} - \theta^*\|_2^2]} + \sqrt{\frac{2\|\Sigma\|_2 \log(1/\delta)}{n}} \\ &\leq \sqrt{\frac{\text{trace}(\Sigma)}{n}} + \sqrt{\frac{2\|\Sigma\|_2 \log(1/\delta)}{n}} \end{aligned}$$

Lugosi and Mendelson showed that with only bounded 2^{nd} moment, this rate can be achieved, but the estimator proposed is intractable.

2 Main Result

Theorem 2.1. *Let $\{X_i\}_{i=1}^n$ be a set of n i.i.d. random vectors i.e. $X_i \in \mathbb{R}^p$, sampled from a distribution with mean θ^* and covariance Σ . Then Descent – Mean – Estimate with stepsize $\gamma = \frac{1}{20}$ and number of iterations $T = 1000 \frac{\log(\|\theta^*\|_2)}{\epsilon}$ returns a mean estimate $\hat{\theta}_{n,\delta}$ that satisfies with probability at least $1 - \delta$:*

$$\|\hat{\theta}_{n,\delta} - \theta^*\|_2 \leq \max(\epsilon, 480000 \cdot \text{OPT}_{n,\delta,\Sigma})$$

Descent – Mean – Estimate is Algorithm 1 in the main text.