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Chapter 1

Projection-Based Evaluation of Tukey Depth from Weighted Student Differences

1.1 Definition and Notation

In [1], the distribution of a weighted difference of two independent Student variables is derived in the univariate setting ($d = 1$). For a given angle θ and independent Student variables t_1 and t_2 with degrees of freedom ν_1 and ν_2 , respectively, one considers the random variable

$$v = t_1 \sin \theta - t_2 \cos \theta.$$

Although the original derivation employs the difference, the substitution of a plus sign is justified by the symmetry of the Student distribution. Moreover, this modification allows the extension of the range of θ from the usual $[0, \pi/2]$ to the full circle $[0, 2\pi)$ without affecting the final density.

Following the approach in [1], the cumulative distribution function (CDF) of v is given by

$$F_v(v_0) = \int_0^1 \Pr\{T_{\nu_1+\nu_2} \leq v_0 \varphi(u)\} \frac{u^{\nu_1/2-1}(1-u)^{\nu_2/2-1}}{B(\nu_1/2, \nu_2/2)} du,$$

where

$$\varphi(u) = \sqrt{\frac{\sin^2 \theta}{u} + \frac{\cos^2 \theta}{1-u}},$$

and $T_{\nu_1+\nu_2}$ denotes a Student variable with $\nu_1 + \nu_2$ degrees of freedom. Note that the squared trigonometric terms eliminate any sign ambiguity, ensuring the expression remains valid for any $\theta \in [0, 2\pi)$.

1.2 The Projection Method and Computation of Tukey Depth

In the multivariate setting, the classical definition of the Tukey (halfspace) depth of a point $x \in \mathbb{R}^p$ with respect to a probability measure F is given by

$$\text{TD}(x) = \inf_{u \in S^{p-1}} F\left(\{y \in \mathbb{R}^p : \langle y, u \rangle \geq \langle x, u \rangle\}\right),$$

i.e. the smallest probability mass in any closed halfspace that contains x .

Our investigation begins by studying the distribution of the weighted sum

$$v = t_1 \sin \theta + t_2 \cos \theta,$$

whose CDF is expressed as

$$F_v(v_0) = \int_0^1 \Pr\{T_{\nu_1+\nu_2} \leq v_0 \varphi(u)\} \frac{u^{\nu_1/2-1}(1-u)^{\nu_2/2-1}}{B(\nu_1/2, \nu_2/2)} du,$$

with

$$\varphi(u) = \sqrt{\frac{\sin^2 \theta}{u} + \frac{\cos^2 \theta}{1-u}},$$

and where $T_{\nu_1+\nu_2}$ is a Student variable with $\nu_1+\nu_2$ degrees of freedom. Notice that by considering the complementary probability $1 - F_v(v_0)$, we obtain the mass contained in the halfspace corresponding to the projection direction determined by θ .

To connect this univariate result to the multivariate Tukey depth, note that if we were to compute the depth by considering all possible projection directions, we would have

$$\text{TD}(x) = \inf_{\theta \in [0, 2\pi)} \{1 - F_v(v_0(\theta))\},$$

where $v_0(\theta)$ denotes the threshold in the projected one-dimensional space corresponding to the halfspace boundary for direction θ .

Under the assumption of elliptical symmetry, it can be shown that the infimum is achieved when the projection is aligned with the point x . That is, by choosing the optimal projection direction

$$u = \frac{x}{\|x\|},$$

the scalar projection of x becomes exactly $\|x\|$. Hence, the Tukey depth simplifies to

$$\text{TD}(x) = 1 - F_v(\|x\|) = 1 - \mathbb{E}\left[G(\|x\| \varphi(U))\right],$$

where

$$G(t) = \Pr\{T_{\nu_1+\nu_2} \leq t\},$$

and U is a Beta random variable with density

$$f_U(u) = \frac{u^{\nu_1/2-1}(1-u)^{\nu_2/2-1}}{B(\nu_1/2, \nu_2/2)}.$$

This formulation connects our analysis of the weighted sum v to the classical Tukey depth definition by demonstrating that, under optimal projection, the centrality of x is measured by the probability mass in the halfspace determined by $x/\|x\|$. Consequently, the computation of $\text{TD}(x)$ reduces to evaluating a univariate expectation.

1.3 Numerical Evaluation and Monte Carlo Approximation

The integral representation of $F_v(v_0)$ can be recast in expectation form. Define

$$G(t) = \Pr\{T_{\nu_1+\nu_2} \leq t\},$$

and let U be a Beta random variable with density

$$f_U(u) = \frac{u^{\nu_1/2-1}(1-u)^{\nu_2/2-1}}{B(\nu_1/2, \nu_2/2)}, \quad 0 \leq u \leq 1.$$

Then the CDF can be written as

$$F_v(v_0) = \mathbb{E}\left[G(v_0 \varphi(U))\right].$$

In practice, this one-dimensional integral is computed either by numerical quadrature or by Monte Carlo simulation. In the quadrature approach, one employs standard routines to evaluate

$$\int_0^1 G(v_0 \varphi(u)) f_U(u) du,$$

while in a Monte Carlo approximation one would generate N independent samples u_1, \dots, u_N from the Beta distribution and approximate the expectation by

$$\widehat{F}_v(v_0) = \frac{1}{N} \sum_{i=1}^N G(v_0 \varphi(u_i)).$$

This reduction to a one-dimensional expectation is particularly attractive in view of its computational tractability.

1.4 Visualization

For the purpose of illustration, we now specialize to the case where $\nu_1 = \nu_2 = v$. Under this simplification, the density of U becomes

$$f_U(u) = \frac{u^{v/2-1}(1-u)^{v/2-1}}{B(v/2, v/2)},$$

and the depth formula reduces to

$$\text{TD}(x) = 1 - \int_0^1 G(\|x\| \varphi(u)) f_U(u) du,$$

with the optimal projection determined by $x/\|x\|$.

To provide a concrete visual interpretation, consider contour plots of the Tukey depth in \mathbb{R}^2 . In these plots, the depth at a point x is computed as above, with the scalar projection given by its Euclidean norm. Figure 1.1 shows three panels corresponding to different values of the degrees-of-freedom parameter v : $v = 1$, $v = 5$, and $v = 100$. In each case the depth is calculated over a grid of points, and the contour plots illustrate how the depth function changes with v and provide insight into the centrality of points in the plane as measured by Tukey depth.

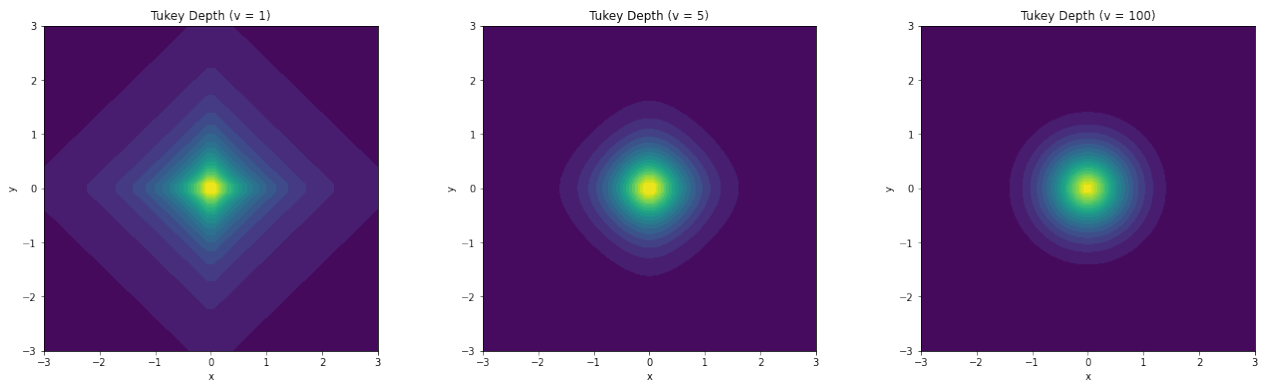


Figure 1.1: Contour plots of the Tukey Depth for three values of v .

Bibliography

- [1] Harold Ruben. On the distribution of the weighted difference of two independent student variables. *Journal of the Royal Statistical Society. Series B (Methodological)*, 22(1):188–196, 1960.