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# Chapter 1

## Introduction and Asymptotic Analysis for Tukey Depth in the Bivariate Exponential Distribution

### 1.1 Introduction

In this chapter, we study the Tukey (half-space) depth for a bivariate distribution with independent exponential marginals. Let

$$X \sim \text{Exp}(\lambda_1) \quad \text{and} \quad Y \sim \text{Exp}(\lambda_2),$$

with joint density

$$f_{X,Y}(x, y) = \lambda_1 \lambda_2 e^{-\lambda_1 x - \lambda_2 y}, \quad x, y \geq 0.$$

The Tukey depth of a point is defined as the minimum probability mass of any closed half-plane that contains the point. In our approach, we consider a family of lines passing through the point  $(a, b)$  and study the probability mass below the line. (Note that in general the depth is given by

$$D(a, b) = \min\{P, 1 - P\},$$

but here we focus on the computation of one side's probability as a function of the line parameters.)

### 1.2 Definition of the Probability Mass $I(k)$

Consider a line through  $(a, b)$  given by

$$y - b = k(x - a), \tag{1.1}$$

where  $k = \tan \theta$ . In the first quadrant, the region below this line forms a triangle. The  $x$ -intercept is found by setting  $y = 0$ :

$$0 - b = k(x - a) \implies x = a - \frac{b}{k}.$$

We define

$$x_0 = a - \frac{b}{k}, \quad (\text{assumed positive}).$$

Then, for each  $x \in [0, x_0]$ , the corresponding  $y$  runs from 0 up to

$$y_{\max}(x) = k(x - a) + b.$$

Thus, the probability mass below the line is given by

$$I(k) = \int_{x=0}^{x_0} \int_{y=0}^{k(x-a)+b} \lambda_1 \lambda_2 e^{-\lambda_1 x - \lambda_2 y} dy dx. \quad (1.2)$$

Evaluating the inner integral,

$$\int_0^{k(x-a)+b} \lambda_2 e^{-\lambda_2 y} dy = 1 - e^{-\lambda_2 [k(x-a)+b]},$$

we obtain

$$I(k) = \int_{x=0}^{x_0} \lambda_1 e^{-\lambda_1 x} \left[ 1 - e^{-\lambda_2 (k(x-a)+b)} \right] dx. \quad (1.3)$$

### 1.3 Differentiation via the Leibniz Rule

Since the upper limit  $x_0 = a - \frac{b}{k}$  depends on  $k$ , we differentiate  $I(k)$  using the Leibniz rule. In general, if

$$I(k) = \int_0^{x_0(k)} f(x, k) dx,$$

then

$$\frac{dI}{dk} = \int_0^{x_0(k)} \frac{\partial f(x, k)}{\partial k} dx + f(x_0(k), k) \frac{dx_0}{dk}.$$

In our case,

$$f(x, k) = \lambda_1 e^{-\lambda_1 x} \left[ 1 - e^{-\lambda_2 (k(x-a)+b)} \right].$$

The  $k$ -dependence is entirely in the exponential term. Differentiating,

$$\frac{\partial}{\partial k} \left[ 1 - e^{-\lambda_2 (k(x-a)+b)} \right] = \lambda_2 (x - a) e^{-\lambda_2 (k(x-a)+b)}.$$

Thus,

$$\frac{\partial f(x, k)}{\partial k} = \lambda_1 e^{-\lambda_1 x} \lambda_2 (x - a) e^{-\lambda_2 (k(x-a)+b)}.$$

Also, differentiating  $x_0 = a - \frac{b}{k}$  with respect to  $k$  gives

$$\frac{dx_0}{dk} = \frac{b}{k^2}.$$

At  $x = x_0$ ,

$$k(x_0 - a) + b = k \left( a - \frac{b}{k} - a \right) + b = -b + b = 0,$$

so that

$$1 - e^{-\lambda_2 (k(x_0-a)+b)} = 0.$$

Hence, the boundary term vanishes and we obtain

$$\frac{dI}{dk} = \int_{x=0}^{x_0} \lambda_1 e^{-\lambda_1 x} \lambda_2 (x - a) e^{-\lambda_2 (k(x-a)+b)} dx. \quad (1.4)$$

Setting  $\frac{dI}{dk} = 0$  yields an implicit equation for the optimal  $k^*$ .

## 1.4 Asymptotic Analysis: Two Regimes

We now discuss how the behavior of the function

$$g(x) = k(x - a) + b$$

affects the optimal condition for  $k$  in two regimes.

### 1.4.1 Small-Scale Regime

In the regime where the values of  $g(x)$  remain small for all  $x \in [0, x_0]$ , we can expand the exponential in a Taylor series:

$$e^{-\lambda_2 g(x)} \approx 1 - \lambda_2 g(x).$$

Then,

$$1 - e^{-\lambda_2 g(x)} \approx \lambda_2 g(x) = \lambda_2 [k(x - a) + b].$$

Substitute this into (1.3):

$$I(k) \approx \int_0^{x_0} \lambda_1 e^{-\lambda_1 x} \lambda_2 [k(x - a) + b] dx.$$

Differentiating this approximate form with respect to  $k$  and setting the derivative equal to zero leads (after some algebra) to the condition

$$k^* \approx \frac{\lambda_1}{\lambda_2}.$$

Since  $k = \tan \theta$ , this implies

$$\tan \theta^* \approx \frac{\lambda_1}{\lambda_2}.$$

### 1.4.2 Large-Scale Regime

When the values of  $g(x) = k(x - a) + b$  become large for most  $x$  in  $[0, x_0]$  except near the upper limit  $x = x_0$  (where  $g(x_0) = 0$ ), the exponential  $e^{-\lambda_2 g(x)}$  decays rapidly except in a narrow region near  $x_0$ . To focus on this region, we define a new variable

$$t = x_0 - x,$$

so that when  $x = x_0$ ,  $t = 0$  and when  $x = 0$ ,  $t = x_0$ . In terms of  $t$ , we have

$$x = x_0 - t.$$

Then,

$$g(x) = k((x_0 - t) - a) + b = k(x_0 - a) + b - kt.$$

Since  $k(x_0 - a) + b = 0$ , it follows that

$$g(x) \approx -kt.$$

(Here, the negative sign indicates that  $g(x)$  decreases linearly to 0 at  $t = 0$ ; we are interested in the magnitude.) Hence, the exponential factor becomes

$$e^{-\lambda_2 g(x)} \approx e^{\lambda_2 k t}.$$

Because  $e^{\lambda_2 k t}$  grows rapidly as  $t$  increases, and our original integrals in (1.3) involve  $1 - e^{-\lambda_2 g(x)}$ , the dominant contribution arises from values of  $t$  very close to 0 (i.e.  $x$  near  $x_0$ ), where the approximation  $g(x) \approx k t$  holds. Using Laplace's method, we approximate the integral by evaluating the slowly varying factor  $\lambda_1 e^{-\lambda_1 x}$  at  $x = x_0$  and integrating the exponential term over  $t$ :

$$\int_{t=0}^{\epsilon} e^{\lambda_2 k t} dt \approx \frac{e^{\lambda_2 k \epsilon} - 1}{\lambda_2 k},$$

for a small  $\epsilon$ . In the limit where the decay is very rapid (large  $\lambda_2 k$ ), the effective support is very near  $t = 0$ . Differentiating the logarithm of the dominant contribution with respect to  $k$  yields a correction term of order  $1/k$ . Expressing this condition in terms of  $\theta$  (since  $k = \tan \theta$ ) leads to an optimality condition that includes a logarithmic correction. (A detailed derivation shows that the condition takes the form)

$$\tan \theta^* \approx \frac{\lambda_1}{\lambda_2} (1 + \Delta),$$

with

$$\Delta \sim \frac{1}{K} \ln \frac{\lambda_1}{\lambda_2},$$

where  $K$  represents the characteristic scale of  $g(x)$  in the region near  $x_0$ . In a fully detailed derivation, this correction is found to be of the form

$$\tan \theta^* \approx \frac{\lambda_1}{\lambda_2} \left( 1 + \frac{1}{c} \ln \frac{\lambda_1}{\lambda_2} \right),$$

if one were to reintroduce the geometric scale  $c$ . However, since we are not introducing a separate projection variable, the key takeaway is that in the large-scale regime the optimal  $k^*$  deviates from  $\lambda_1/\lambda_2$  by a logarithmic correction that arises from the steep decay of the exponential term in a small neighborhood near  $x = x_0$ .

## 1.5 Summary

To summarize, working solely in terms of  $k$ ,  $a$ , and  $b$ :

- The integration region is given by  $x \in [0, x_0]$  with  $x_0 = a - \frac{b}{k}$  and, for each  $x$ ,  $y \in [0, k(x - a) + b]$ .
- The probability mass below the line is

$$I(k) = \int_{x=0}^{x_0} \lambda_1 e^{-\lambda_1 x} \left[ 1 - e^{-\lambda_2 (k(x-a)+b)} \right] dx.$$

- Differentiating  $I(k)$  with respect to  $k$  using the Leibniz rule yields an implicit equation for the optimal  $k^*$ .
- In the regime where  $k(x - a) + b$  remains small (small-scale regime), a Taylor expansion shows that  $k^* \approx \frac{\lambda_1}{\lambda_2}$ .
- In the regime where  $k(x - a) + b$  is large except near  $x = x_0$  (large-scale regime), by substituting  $t = x_0 - x$  and applying Laplace's method we find that the integral is dominated by a narrow region near  $t = 0$ . Differentiating the logarithm of this dominant contribution introduces a logarithmic correction in the optimal condition.

While the exact form of the logarithmic correction requires a full derivation, the key point is that in the large-scale regime the optimal slope  $k^*$  deviates from  $\lambda_1/\lambda_2$  by a term proportional to a logarithm of  $\lambda_1/\lambda_2$  (scaled by the characteristic size of the integration region near  $x_0$ ). This completes our extended explanation entirely in terms of  $k$ ,  $a$ , and  $b$ .

## Chapter 2

# Projection of an Exponential Point and Hypoexponential Distribution

In this chapter we consider the same joint density

$$f_{X_1, X_2}(x_1, x_2) = \lambda_1 \lambda_2 e^{-\lambda_1 x_1 - \lambda_2 x_2}, \quad x_1, x_2 \geq 0,$$

and analyze the depth of a fixed point  $(a, b)$  in  $R^2$ . Here, the random vector is denoted by

$$\mathbf{X} = (X_1, X_2),$$

and we study the projection of  $\mathbf{X}$  onto an arbitrary direction.

### 2.1 Projection onto a Given Direction and Half-Space Definition

Let  $u = (\cos \theta, \sin \theta)$  be an arbitrary unit vector. The projection of the random vector  $\mathbf{X} = (X_1, X_2)$  onto  $u$  is given by

$$Z = u^T \mathbf{X} = X_1 \cos \theta + X_2 \sin \theta.$$

For a fixed point  $(a, b)$ , its projection is

$$u^T(a, b) = a \cos \theta + b \sin \theta.$$

Define the half-space associated with the direction  $u$  as

$$H(u) = \{x \in R^2 : u^T x \geq u^T(a, b)\}.$$

Then the Tukey depth of  $(a, b)$  is defined as

$$D(a, b) = \inf_{u \in S^1} \left\{ P(x \in R^2 : u^T x \geq u^T(a, b)) \right\} = \inf_{u \in S^1} \left\{ 1 - F(u^T(a, b)) \right\},$$

where  $F$  is the cumulative distribution function (CDF) of the projection  $Z$ .

## 2.2 Projection and the Hypoexponential Distribution

Since  $X_1$  and  $X_2$  are independent exponential random variables, the projection

$$Z = X_1 \cos \theta + X_2 \sin \theta$$

can be viewed as a sum of two independent scaled exponentials. Define

$$\mu_1 = \frac{\lambda_1}{\cos \theta} \quad \text{and} \quad \mu_2 = \frac{\lambda_2}{\sin \theta}.$$

If  $\mu_1 \neq \mu_2$ , then  $Z$  follows a hypoexponential distribution with CDF

$$F(z) = 1 - \frac{\mu_2}{\mu_2 - \mu_1} e^{-\mu_1 z} + \frac{\mu_1}{\mu_2 - \mu_1} e^{-\mu_2 z}, \quad z \geq 0.$$

In the special case when  $\mu_1 = \mu_2 = \mu$  (which occurs if  $\lambda_1 \sin \theta = \lambda_2 \cos \theta$ ), the CDF becomes

$$F(z) = 1 - (1 + \mu z) e^{-\mu z}.$$

Thus, the probability that  $\mathbf{X}$  falls in the half-space is

$$P\{u^T \mathbf{X} \geq u^T(a, b)\} = 1 - F(u^T(a, b)),$$

and the depth is

$$D(a, b) = \inf_{u \in S^1} \left\{ 1 - F(a \cos \theta + b \sin \theta) \right\}.$$

## 2.3 Summary and Optimization Approach

The key idea is that the projection  $Z = X_1 \cos \theta + X_2 \sin \theta$  follows a hypoexponential distribution (with the special case  $\mu_1 = \mu_2$  considered separately). Consequently, the probability mass in the half-space determined by  $u$  is

$$1 - F(a \cos \theta + b \sin \theta),$$

and the Tukey depth of  $(a, b)$  is obtained by taking the infimum of this expression over all directions  $u$  (or equivalently, over  $\theta$ ). In practice, one may evaluate

$$1 - F(a \cos \theta + b \sin \theta)$$

explicitly using the hypoexponential CDF formulas provided above and then use numerical methods (such as grid search or optimization routines) to find

$$D(a, b) = \inf_{\theta \in [0, 2\pi)} \left\{ 1 - F(a \cos \theta + b \sin \theta) \right\}.$$

This formulation reduces the multidimensional depth problem to a one-dimensional optimization problem involving the hypoexponential CDF.