Lower Bounds and Hardness Amplification for Learning Shallow Monotone Formulas

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Abstract

Much work has been done on learning various classes of "simple" monotone functions under the uniform distribution. In this paper we give the first unconditional lower bounds for learning problems of this sort by showing that polynomial-time algorithms cannot learn constant-depth monotone Boolean formulas under the uniform distribution in the well-studied Statistical Query model.

Using a recent characterization of Strong Statistical Query learnability due to Feldman [14], we first show that depth-3 monotone formulas of size $n^{o(1)}$ cannot be learned by any polynomial-time Statistical Query algorithm to accuracy $1-1/(\log n)^{\Omega(1)}$. We then build on this result to show that depth-4 monotone formulas of size $n^{o(1)}$ cannot be learned even to a certain $\frac{1}{2} + o(1)$ accuracy in polynomial time. This improved hardness is achieved using a general technique that we introduce for amplifying the hardness of "mildly hard" learning problems in either the PAC or Statistical Query framework. This hardness amplification for learning builds on the ideas in the work of O'Donnell [28] on hardness amplification for approximating functions using small circuits, and is applicable to a number of other contexts.

Finally, we demonstrate that our approach can also be used to reduce the well-known open problem of learning juntas to learning of depth-3 monotone formulas.

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1 Introduction

Motivation. Over the past several decades much work in computational learning theory has focused on developing efficient algorithms for learning monotone Boolean functions under the uniform distribution, see e.g., [1, 3, 9, 16, 18, 21, 27, 29, 30, 34] and other works. An intriguing question, which has driven much of this research and remains open, is whether there is an efficient algorithm to learn *monotone DNF formulas* under the uniform distribution. Such an algorithm A would have the following performance guarantee: for any target function $f: \{0, 1\}^n \to \{0, 1\}$ that is a monotone DNF formula with poly(n) terms, given access to independent uniform random examples (x, f(x)), algorithm A would run in poly $(n, 1/\epsilon)$ time and with high probability output a hypothesis h that disagrees with f on at most an ϵ fraction of inputs from $\{0, 1\}^n$.

Several partial positive results toward learning monotone DNF have been obtained: for constant ϵ , algorithms are known that can learn $2\sqrt{\log n}$ -term monotone DNF [34] and poly(n)-size monotone decision trees [27] in poly(n) time. Partial negative results have also been given: [12] has shown that (under a strong cryptographic hardness assumption) for a sufficiently large absolute constant d there is no poly(n)-time algorithm that can learn depth-d, size- $n^{o(1)}$ Boolean formulas that compute monotone functions to a certain accuracy $\frac{1}{2} + o(1)$. However no hardness results that apply to monotone formulas of small constant depth are known.

In this work we give *unconditional* lower bounds showing that simple monotone functions – computed by monotone Boolean formulas of depth 3 or 4 and size $n^{o(1)}$ – cannot be learned under the uniform distribution in polynomial time. Of course these results are not in the PAC model of learning from random examples (since unconditional lower bounds in this model would prove $P \neq NP$!); our primary lower bounds are for the well-studied *Statistical Query* learning model, which we describe briefly below.

Statistical Query learning. Kearns [20] introduced the *statistical query (SQ)* learning model as a natural variant of the usual PAC learning model. In the SQ model, instead of having access to independent random examples (x, f(x)) drawn from distribution \mathcal{D} , the learner is only allowed to obtain statistical properties of examples. Formally, it has access to a *statistical query oracle* $SQ_{f,\mathcal{D}}$. The oracle $SQ_{f,\mathcal{D}}$ takes as input a *query function* $g: X \times \{+1, -1\} \rightarrow \{+1, -1\}$ and a *tolerance parameter* $\tau \in [0, 1]$ and outputs a value v such that:

$$|v - \mathbf{E}_{\mathcal{D}}[g(x, f(x))]| \le \tau.$$

The learner's goal – to output a hypothesis h such that $\Pr_{x \sim \mathcal{D}}[h(x) \neq f(x)] \leq \epsilon$ – is the same as in PAC learning. A poly $(n, 1/\epsilon)$ -time SQ algorithm is only allowed to make queries in which g can be computed by a poly $(n, 1/\epsilon)$ -size circuit and τ is at most a fixed $1/\operatorname{poly}(n, 1/\epsilon)$ (and of course the algorithm must run for at most poly $(n, 1/\epsilon)$ time steps).

The SQ model is an important and well-studied learning model which has received much research attention in the 15 years since it was introduced. One reason for this intense interest is that any concept class that is efficiently learnable from statistical queries is also efficiently PAC learnable in the presence of random classification noise at any noise rate bounded away from $\frac{1}{2}$ [20]. In fact, since the introduction of the SQ-model virtually all known noise-tolerant learning algorithms have been obtained from (or rephrased as) SQ algorithms¹ [20, 11, 10, 13].

Even more importantly, and quite surprisingly, all known techniques in PAC learning with the exception of Gaussian elimination either fit easily into the SQ model or have SQ analogues. Thus the study of SQ learning is now an integral part of the study of noise-tolerant learning and of PAC learning in general. In

¹One prominent exception is the work of [6], which gives an algorithm for learning parities which is tolerant to random noise, although in a weaker sense than the algorithms derived from statistical queries.

addition, interest in the SQ learning model has been stimulated by recently discovered close connections to privacy-preserving learning [4, 19], evolvability [43, 15, 14] and communication complexity [35].

An important property of the SQ model that we rely on in this work is that it is possible to prove *unconditional* information-theoretic lower bounds on learning in the SQ model (we will say much more about this below). Such lower bounds give a very strong indication that the class is unlikely to be efficiently PAC learnable and even less likely to be PAC learnable in the presence of random classification noise. In particular it also implies that algorithms that rely solely on estimates of Fourier coefficients (which include the best known algorithms for learning monotone functions over the uniform distribution [34, 27]) will not be efficient.

Background on hardness results for SQ learning. In his paper introducing the SQ model [20], Kearns already showed that the class of all parity functions cannot be SQ-learned in polynomial time under the uniform distribution. Soon after this Blum *et al.* [5] characterized the weak learnability of every function class \mathcal{F} in the SQ model in terms of the *statistical query dimension* of \mathcal{F} ; roughly speaking, this is the largest number of functions from \mathcal{F} that are pairwise nearly orthogonal to each other (we give a precise definition in Section 2). The results of [5] imply that if a class \mathcal{F} has SQ-Dimension $n^{\omega(1)}$, then no SQ algorithm can even weakly learn \mathcal{F} to any accuracy $\frac{1}{2} + \frac{1}{\text{poly}(n)}$ in poly(n) time. This bound was already used to give SQ hardness results for weak learning classes such as DNF and decision trees in [5], and more recently for weak-learning intersections of halfspaces in [24]. However, it is well known that the entire class of all monotone Boolean functions over $\{0,1\}^n$ can be weakly learned to accuracy $\frac{1}{2} + \frac{1}{\text{poly}(n)}$ in poly(n) time (an algorithm that achieves optimal accuracy $\frac{1}{2} + \frac{\Theta(\log n)}{\sqrt{n}}$ was recently given in [29]), and indeed the class of all monotone functions can easily be shown to have SQ-dimension O(n). Thus the notion of SQ-dimension alone is not enough to yield SQ lower bounds on learning monotone functions.

Much more recently, Simon [36] introduced a combinatorial parameter of a function class \mathcal{F} called its *strong Statistical Query dimension*, and showed that this parameter at error rate ϵ characterizes the information-theoretic *strong* learnability of \mathcal{F} to accuracy $1 - \epsilon$. (We give a precise definition of the strong SQ-dimension in Section 2.) We use this characterization, which was later strengthened and simplified by Feldman [14] (and independently by Szörényi [37]) to obtain the lower bounds, which we now describe. (Throughout the following description of our results, the underlying distribution is always taken to be uniform over $\{0,1\}^n$.)

Our Results: Unconditional Hardness of Learning Simple Monotone Functions. We give the first strong SQ-dimension lower bound for a class of "simple" monotone functions. More precisely, as our first main result, we show that the class of size- $n^{o(1)}$, depth-3 monotone formulas has strong SQ-dimension $n^{\omega(1)}$ at a certain error rate $1/(\log n)^{\Theta(1)}$. By the results of Simon and Feldman, this implies that such formulas cannot be efficiently learned to accuracy $1 - 1/(\log n)^{\Theta(1)}$ by any polynomial-time SQ learning algorithm. Roughly speaking, our proof works by constructing a class of slice functions of "well-separated" parities over poly $\log(n)$ variables. We show that this class of functions has the combinatorial properties required to satisfy the strong SQ-dimension criterion, and that every function in the class can be computed by a small monotone formula of depth 3.

In addition to this result, we show that a variant of the basic idea of our construction can be used to reduce PAC learning of $\log n$ -juntas (or functions that depend on at most $\log n$ variables) over $k = \log^2 n/\log\log n$ variables to learning of depth-3 monotone functions. Learning of juntas is an important open problem in learning theory [2] for which the best known algorithm achieves only a polynomial factor speed-up over the trivial brute force algorithm [26]. Note that there are $n^{\Omega(\log\log n)}$ juntas in the above class of functions and therefore our reduction implies that strong learning of monotone depth-3 formulas of poly(n)-size (even over

only k variables) would require a major breakthrough on the problem of learning juntas over the uniform distribution.

These results are the first lower bounds for learning monotone depth-3 formulas, and leave only the question of learning monotone DNF formulas (which are of course depth-2 rather than depth-3 monotone formulas) open. However, these results only say that monotone depth-3 formulas cannot be learned to a rather high (1 - o(1)) accuracy. Thus a natural goal is to obtain stronger hardness results which show that simple monotone functions are hard to SQ learn even to coarse accuracy – ideally to some accuracy level $\frac{1}{2} + o(1)$ only slightly better than random guessing. Of course, we might expect that to achieve this we must use somewhat more complicated functions than depth-3 formulas, and this does turn out to be the case – but only a bit more complicated, as we describe below.

We introduce a general method of amplifying the hardness of a class of functions that are "mildly hard to learn" (*i.e.*, hard to learn to high accuracy), to obtain a class of functions that are "very hard to learn" (*i.e.*, hard to learn to accuracy even slightly better than random guessing). We show that our method, which builds on O'Donnell's hardness amplification for approximating Boolean functions using small circuits [28], can be applied both within the uniform-distribution PAC model (Th. 4.3) and within the uniform-distribution Statistical Query model (Th. 4.5). The latter is of course our main interest in this paper, but we feel that the result is of independent interest and therefore present both versions.

We note that while our hardness amplification follows the general approach of O'Donnell, new technical results are required to successfully translate the approach into the SQ setting. We defer the discussion of the proof technique and technical contributions to Section 4.

Using this hardness amplification for SQ learning together with our first main result, we obtain our second main result: we show that the class of size- $n^{o(1)}$, depth-4 monotone formulas cannot be SQ-learned even to $\frac{1}{2} + 2^{-(\log n)^{\gamma}}$ accuracy in poly(n) time for any $\gamma < 1/2$. We are able to increase the depth by only one (from 3 to 4) by a careful construction of the combining function in our hardness amplification framework; we use a depth-2 combining function due to Talagrand and the complement of the "tribes" function which have useful extremal noise stability properties as shown by Mossel and O'Donnell [28, 25].

Relation to previous work. To the best of our knowledge even the "mild hardness" result that we prove for depth-3 monotone formulas is the first unconditional negative result known for learning a class of polynomial-time computable monotone functions in the uniform-distribution SQ model. We note that the strong $\frac{1}{2} + o(1)$ hardness that we establish for depth-4 monotone formulas is provably near-optimal, since as mentioned earlier the class of *all* monotone functions over $\{0,1\}^n$ can be learned to accuracy $\frac{1}{2} + \frac{\Theta(\log n)}{\sqrt{n}}$ in polynomial time [29].

While the recent work in [12] also gave negative results for learning constant-depth monotone formulas, those results are different from ours in significant ways. [12] used a strong cryptographic hardness assumption – that Blum integers are $2^{n^{\epsilon}}$ -hard to factor on average for some fixed $\epsilon > 0$ – to show that for some sufficiently large absolute constant d, the class of monotone functions computed by size- $n^{o(1)}$, depth-d formulas cannot be PAC learned, under the uniform distribution, to a certain accuracy $\frac{1}{2} + o(1)$. In contrast, our main hardness result applies to the more restricted classes of size- $n^{o(1)}$, depth-3 and 4 *monotone* formulas, and gives *unconditional* hardness for polynomial-time algorithms in the Statistical Query model.

Our reduction from learning juntas can be thought of as giving a hardness result based on a relatively strong computational assumption and hence has the same flavor as the result in [12]. In addition to better depth, our reduction is substantially simpler and more direct than the reduction from factoring Blum integers.

Finally, we remark here that our hardness amplification method for PAC and SQ learning may be viewed as a significant strengthening and generalization of some earlier results. Boneh and Lipton [7] described a form of uniform-distribution hardness amplification for PAC learning based on the XOR lemma; our PAC

hardness amplification generalizes their result and extends to SQ learning. More recently, [12] used elements of O'Donnell's technique to amplify information-theoretic hardness of learning. Specifically, the "mildly hard" class of functions \mathcal{F} used in [12] consists of all functions of the form $\operatorname{slice}(f)$, where f may be any Boolean function and $\operatorname{slice}(f)$ is the function which agrees with Majority everywhere except the middle layer of the Boolean hypercube. An easy argument shows that \mathcal{F} is a class of monotone functions that is hard to learn to accuracy $1 - \Theta(1)/\sqrt{n}$. Using the fact that a random function in \mathcal{F} is trivial to predict off of the middle layer and is totally random on the middle layer, expected bias analysis from [28] is used in [12] to derive information-theoretic hardness of learning the combined function class \mathcal{F}^g . In contrast, the hardness amplification results of this paper amplify computational hardness, do not assume any particular structure of the base class \mathcal{F} (only that it is "mildly hard to learn") and also apply to learning in the SQ model.

Organization. Section 2 gives background on Statistical Query learning, the SQ-dimension, and the strong SQ-dimension. In Section 3 we describe our class of depth-3 monotone formulas and show that it is "mildly" hard to learn in the SQ model by giving a superpolynomial lower bound on its strong SQ-dimension. We give the reduction from learning juntas in Section 3.3. Section 4 presents our general hardness amplification results for the uniform-distribution PAC model and the uniform-distribution Statistical Query learning model. We apply our hardness amplification technique from Section 4 to obtain our second main result, strong SQ hardness for depth-4 formulas, in Section 5.

2 The Statistical Query Model, SQ-Dimension, and Strong SQ-Dimension

Recall the definition of Statistical Query learning from Section 1. Blum *et al.* [5] introduced the notion of the SQ-dimension of function class \mathcal{F} under distribution \mathcal{D} , and showed that it characterizes the weak-learnability of \mathcal{F} under \mathcal{D} in the SQ model. Bshouty and Feldman [8] and Yang [44] later generalized and sharpened the Blum *et al.* result. We will use Yang's version here extended to sets of arbitrary real-valued functions. We write " $\langle f, g \rangle_{\mathcal{D}}$ " to denote $\mathbf{E}_{x \sim \mathcal{D}}[f(x)g(x)]$ and " $||f||_{\mathcal{D}}$ " to denote $(\langle f, f \rangle_{\mathcal{D}})^{1/2}$.

Definition 2.1 Given a set C of real-valued functions, the SQ-dimension of C with respect to D (written SQ-DIM(C, D)) is the largest number d such that $\exists \{f_1, \ldots, f_d\} \subseteq C$ with the property that $\forall i \neq j$,

$$\left| \langle f_i, f_j \rangle_{\mathcal{D}} \right| \le \frac{1}{d}. \tag{1}$$

When \mathcal{D} is the uniform distribution we simply write SQ-DIM(\mathcal{C}). We refer to the LHS of Equation (1) as the *correlation* between f_i and f_j under \mathcal{D} .

Intuitively, this condition says that C contains d "nearly-uncorrelated" functions. It is easy to see that if C is a concept class with SQ-DIM $(C,\mathcal{D})=d$ then C can be weakly learned with respect to \mathcal{D} to accuracy $\frac{1}{2}+\frac{\Theta(1)}{d}$ using d Statistical Queries with tolerance $\frac{\Theta(1)}{d}$; simply ask for the correlation between the unknown target function f and each function in the set $\{f_1,\ldots,f_d\}$. Since the set is maximal, the target function must have correlation at least 1/d with at least one of the functions.

Blum *et al.* showed that the other direction is true as well; if *C* is efficiently weakly learnable, then *C* must have small SQ-dimension.

Theorem 2.2 ([5] Theorem 12) Given a concept class C and a distribution \mathcal{D} , let $SQ\text{-}DIM(C,\mathcal{D})=d$. Then if the tolerance τ of each query is always at least $1/d^{1/3}$, at least $\frac{1}{2}d^{1/3}-1$ queries are required to learn C with advantage $1/d^3$.

As an example, the class PAR of all parity functions over n variables has $SQ\text{-DIM}(PAR) = 2^n$, and thus any SQ algorithm for learning parities over the uniform distribution \mathcal{U} to accuracy $\frac{1}{2} + \frac{1}{2^{O(n)}}$ requires exponential time.

2.1 The Strong SQ-Dimension

The statistical query dimension only characterizes the weak SQ-learnability of a class and is not sufficient to characterize its strong SQ-learnability. The first characterization of strong SQ learning was given by Simon [36], but for our application a subsequent accuracy-preserving characterization by Feldman will be more convenient to use [14].

Let \mathcal{F}_1^{∞} denote the set of all functions from $\{0,1\}^n \to [-1,1]$, *i.e.*, all functions with L_{∞} -norm bounded by 1. For a Boolean function f, we define $B_{\mathcal{D}}(f,\epsilon)$ to be $\{g:\{0,1\}^n \to \{-1,1\}: \mathbf{Pr}_{\mathcal{D}}[g \neq f] \leq \epsilon\}$, *i.e.*, the ϵ -ball around f, and $\mathbf{0}$ to be the constant 0 function. The sign function is defined as $\mathrm{sign}(z) = 1$ for $z \geq 0$, $\mathrm{sign}(z) = -1$ for z < 0. Finally, for a set of real-valued functions C, let $C - g = \{f - g: f \in C\}$.

Definition 2.3 Given a concept class C and $\epsilon > 0$, the strong SQ-dimension of C with respect to \mathcal{D} is defined to be:

$$\mathit{SQ\text{-}SDIM}(C,\mathcal{D},\epsilon) = \sup_{g \in \mathcal{F}_1^\infty} \mathit{SQ\text{-}DIM}((C \setminus B_{\mathcal{D}}(\mathsf{sign}(g),\epsilon)) - g,\mathcal{D}).$$

Just as for the weak SQ-dimension, the strong SQ-dimension completely characterizes the strong SQ-learnability of a concept class.

Theorem 2.4 ([14]) Let C be a concept class over $\{0,1\}^n$, \mathcal{D} be a probability distribution over $\{0,1\}^n$ and $\epsilon > 0$. If there exists a polynomial $p(\cdot,\cdot)$ such that C is SQ learnable over \mathcal{D} to accuracy ϵ from $p(n,1/\epsilon)$ queries of tolerance $1/p(n,1/\epsilon)$ then SQ-SDIM(C, \mathcal{D} , $\epsilon + 1/p(n,1/\epsilon)$) $\leq p'(n,1/\epsilon)$ for some polynomial $p'(\cdot,\cdot)$. Further, if SQ-SDIM(C, \mathcal{D} , ϵ) $\leq q(n,1/\epsilon)$ for some polynomial $q(\cdot,\cdot)$ then C is SQ learnable over \mathcal{D} to accuracy ϵ from $q'(n,1/\epsilon)$ queries of tolerance $1/q'(n,1/\epsilon)$ for some polynomial $q'(\cdot,\cdot)$.

Armed with Definition 2.3 and Theorem 2.4, we can show that a concept class C is not polynomial-time learnable to high accuracy by choosing a suitable $\epsilon = \Omega(1/\operatorname{poly}(n))$ and a suitable function $g \in \mathcal{F}_1^{\infty}$ and proving that $\operatorname{SQ-DIM}((C \setminus B_{\mathcal{U}}(\operatorname{sign}(g), 2\epsilon)) - g) = n^{\omega(1)}$ (we can assume without loss of generality that ϵ upper bounds the tolerance of an SQ algorithm). We do just this, for a class of depth-3 monotone formulas, in the next section.

3 Lower Bounds for Depth-3 Monotone Formulas

In this section we describe our lower bound for SQ learning of depth-3 monotone formulas and the reduction from learning juntas.

3.1 Strong SQ lower bound

We start be showing a family of monotone functions that cannot be strong SQ-learned in polynomial time under the uniform distribution. The high-level idea is that we embed a family of non-monotone functions with high SQ-dimension – a family of parity functions – into the middle level of the *k*-dimensional Boolean cube, and thus obtain a class of monotone functions with high strong SQ-dimension.

A *k-variable slice function for* f, where f is a real-valued function over $\{0,1\}^k$, is denoted slice f. For f is a real-value of slice f is 1 if f has more than f ones, f ones, f if f has fewer than f ones, and f if f has exactly f ones. The functions we consider will only be defined over the first f out of f variables. Throughout the rest of this section, without loss of generality, we will always assume that f is even.

Theorem 3.1 Let \mathcal{P} be the class of 2^{k-1} parity functions $\chi: \{0,1\}^k \to \{+1,-1\}$ over an odd number of the first k variables. Let \mathcal{M} be the class of corresponding k-variable slice functions $\operatorname{slice}_{\chi}$ for $\chi \in \mathcal{P}$. Let $k = \log^{2-\beta}(n)$ for β any absolute constant in (0,1). Then for every $\epsilon = o(1/\sqrt{k})$, we have SQ-SDIM $(\mathcal{M}, \epsilon) = n^{\Theta(\log^{1-\beta}n)}$, and every function in \mathcal{M} is balanced.

Proof: We first show that every function $\operatorname{slice}_{\chi} \in \mathcal{M}$ is balanced, i.e. outputs +1 and -1 with equal probability. As k is even, the number of inputs with greater than k/2 ones is the same as the number of inputs with fewer than k/2 ones. As for the middle layer, given an input with exactly k/2 ones on which χ outputs +1, flipping all the bits gives another point with exactly k/2 ones on which χ outputs -1 (as χ is a parity on an odd number of bits). Thus every $\operatorname{slice}_{\chi} \in \mathcal{M}$ is balanced on the middle layer and thus is balanced overall.

Let $g = \text{slice}_0$. We will show that SQ-DIM($\mathcal{M} \setminus B_{\mathcal{U}}(\text{sign}(g), 2\epsilon) - g) = n^{\omega(1)}$. By Stirling's approximation, the middle layer of the k-dimensional hypercube is a $\lambda_k = \binom{k}{k/2}/2^k = \Theta(1/\sqrt{k})$ fraction of the 2^k points. Thus for $\epsilon = o(1/\sqrt{k})$ we have that \mathcal{M} is disjoint from $B_{\mathcal{U}}(\text{sign}(g), 2\epsilon) = \emptyset$ (since sign(g) equals +1 everywhere on the middle layer and every function in \mathcal{M} is balanced on the middle layer), and it is enough to lower-bound SQ-DIM($\mathcal{M} - g$) in order to lower-bound the strong SQ-dimension of \mathcal{M} .

The functions in $\mathcal{M} - g$ have a nice structure as they output 0 everywhere except the middle layer of $\{0,1\}^k$, where they output ± 1 . Thus, the correlation between any two functions in $\mathcal{M} - g$ depends only on the values on the middle slice. Let $\chi_A, \chi_B \in \mathcal{P}$ be the parity functions over the sets of variables $A, B \subseteq [k]$. Recalling Equation (1),

$$|\langle \operatorname{slice}_{\chi_A} - g, \operatorname{slice}_{\chi_B} - g \rangle_{\mathcal{U}}| = |\mathbf{E}_{\mathcal{U}}[\mathbf{1}_{|x|=k/2} \cdot \chi_A \cdot \chi_B]| = \mathbf{E}_{\mathcal{U}}[\mathbf{1}_{|x|=k/2} \cdot \chi_{A \oplus B}] = \widehat{\mathbf{1}_{|x|=k/2}}(A \oplus B)$$

where $A \oplus B$ denotes the symmetric difference between the sets A and B, $\mathbf{1}_{|x|=k/2}$ is the indicator function of the middle slice, and $\widehat{h}(A)$ is the Fourier coefficient of h with index χ_A . In other words, the correlation between $(\operatorname{slice}_{\chi_A} - g)$ and $(\operatorname{slice}_{\chi_B} - g)$ is exactly the Fourier coefficient of $\mathbf{1}_{|x|=k/2}$ with index $A \oplus B$. Let $s = |A \oplus B|$. By symmetry, all $\binom{k}{s}$ of the degree-s Fourier coefficients of $\mathbf{1}_{|x|=k/2}$ are the same, and since by Parseval's identity the squares of all the Fourier coefficients sum to $\mathbf{E}_{\mathcal{U}}[\mathbf{1}^2_{|x|=k/2}] = \lambda_k$, we have $|\langle \operatorname{slice}_{\chi_A} - g, \operatorname{slice}_{\chi_B} - g \rangle_{\mathcal{U}}| \leq \sqrt{\lambda_k/\binom{k}{s}} \leq \binom{k}{s}^{-1/2}$.

It remains to identify a large collection of these slice functions such that the pairwise correlations are small. This can be done easily by picking any $\chi_A \in \mathcal{P}$, removing all $\chi_B \in \mathcal{P}$ such that $|A \oplus B| \notin [k/3, 2k/3]$, and repeating this process. Since each removal step removes at most a $\frac{1}{2^{\Theta(k)}}$ fraction of all 2^{k-1} elements of \mathcal{P} , in this fashion we can construct a set S of size $2^{\Theta(k)}$. Every pair of parities in S has symmetric difference s for some $s \in [k/3, 2k/3]$, and for such an s we have $\binom{k}{s} = 2^{\Theta(k)}$. Thus the set $\{\text{slice}_{\chi} - g\}_{\chi \in S}$ is a collection of $2^{\Theta(k)} = n^{\Theta(\log^{1-\beta} n)}$ functions in $\mathcal{M} - g$ whose pairwise correlations are each at most $1/2^{\Theta(k)} = 1/n^{\Theta(\log^{1-\beta} n)}$, and thus the SQ-dimension of $\mathcal{M} - g$ is at least $n^{\Theta(\log^{1-\beta} n)}$.

3.2 The depth-3 construction

It remains to show that every function in \mathcal{M} has a depth-3 monotone formula.

Theorem 3.2 Let χ be any parity function over some subset of the variables x_1, \ldots, x_k where $k = \log^{2-\beta}(n)$ for β any absolute constant in (0,1). Then the k-variable slice function slice χ is computed by an $n^{o(1)}$ -size, depth-3 monotone formula.

Proof: Let Th_j^k be the k-variable threshold function that outputs TRUE if at least j of the k inputs are set to 1, and FALSE otherwise. The threshold function Th_j^k can be computed by a monotone formula of size $n^{o(1)}$ and depth 3 using the construction of Klawe *et al.* [22].

Let χ be a parity function on j out of the first k variables. For $x \in \{0, 1\}^k$ let x^1 refer to the j variables of χ and x^2 refer to the remaining k - j variables. We claim that

$$\operatorname{slice}_{\chi}(x) = \bigvee_{\text{odd } i < j} [\operatorname{Th}_{i}^{j}(x^{1}) \wedge \operatorname{Th}_{k/2-i}^{k-j}(x^{2})] .$$

To see this, note that if an input x has fewer than k/2 ones, then there can be no i such that $\operatorname{Th}_i^j(x^1)$ and $\operatorname{Th}_{k/2-i}^{k-j}(x^2)$ both hold, so this function outputs FALSE as it should. If x has more than k/2 ones, some ℓ of them are in x^1 , and at least $k/2 - \ell + 1$ of them are in x^2 . If ℓ is odd then $i = \ell$ makes the OR output TRUE, and if ℓ is even then $i = (\ell - 1)$ makes the OR output TRUE. Finally, if x has exactly x0 ones, and an odd number of them are in x1, the formula is satisfied; if an even number of them are in x1, the formula is not satisfied.

Each Th_i^j and $\operatorname{Th}_{k/2-i}^{k-j}$ can be computed by a $n^{o(1)}$ -size, depth-3 monotone formula with an OR on top [22]. Using the distributive law we can convert $\operatorname{Th}_i^j(x^1) \wedge \operatorname{Th}_{k/2-i}^{k-j}(x^2)$ to also be a $n^{o(1)}$ -size, depth-3 monotone formula with an OR on top. This OR can be collapsed with the top $\lceil j/2 \rceil$ -wise OR, yielding a $n^{o(1)}$ -size, depth-3 monotone formula for slice_Y.

We have thus established:

Theorem 3.3 For some $\epsilon = 1/(\log n)^{\Theta(1)}$, the class of $n^{o(1)}$ -size, depth-3 monotone formulas has Strong SO-Dimension $n^{\omega(1)}$.

As an immediate corollary, by Theorem 2.4 we get:

Corollary 3.4 The class of $n^{o(1)}$ -size, depth-3 monotone formulas is not SQ-learnable to some accuracy $1 - 1/(\log n)^{\Theta(1)}$ in poly(n) time.

3.3 Reduction from learning juntas

We now show that ideas from the proof of Theorem 3.2 can also be used to reduce learning of other non-monotone function classes to learning of shallow monotone formulas. Namely, we give the following reduction from learning of juntas over $k = \log^2 n / \log \log n$ variables to learning of depth-3 monotone formulas. The *i*th variable of a Boolean function f is said to be *relevant* if there exist inputs x and y in $\{0, 1\}^n$ that differ only on the *i*th coordinate such that $f(x) \neq f(y)$. A j-junta is precisely a function that has at most j relevant variables.

Theorem 3.5 Let A be a uniform distribution PAC learning algorithm that learns the class of poly(n)-size depth-3 monotone formulas to accuracy ϵ in time polynomial in n and $1/\epsilon$. Then there exists a uniform-distribution PAC learning algorithm C that exactly learns the class of log(n)-juntas where the relevant variables are chosen from the first $k = \log^2(n)/\log\log(n)$ variables in time polynomial in n.

The proof of this theorem has two components. The first one is to show that the slice function of any DNF formula of polynomial size on $k = \log^2(n)/\log\log(n)$ variable can be expressed as depth-3 monotone formula of polynomial size. This allows us to express juntas on up to $\log(n)$ variables using a polynomial-size depth-3 monotone formula. The second component is the algorithm that reconstructs a junta function exactly given the slice function of the junta. This is possible since for every combination of values of the relevant variables, the points in the middle slice that have this combination of values constitute an inverse polynomial fraction of the domain.

We now provide the proof of the first component.

Lemma 3.6 Let $f = T_1, ..., T_t$ be a DNF formula over some subset of the variables $x_1, ..., x_k$ where $t = n^{O(1)}$ and $k = \log^2(n)/\log\log(n)$. Then the k-variable slice function slice_f is computable by an $n^{O(1)}$ -size, depth-3 monotone formula.

Proof: As in the proof of Theorem 3.2 let Th_j^k be the k-variable threshold function that outputs TRUE if at least j of the k inputs are set to 1, and FALSE otherwise. The threshold function Th_j^k can be computed by a monotone formula of size $n^{O(1)}$ and depth 3 using the construction of Klawe *et al.* [22].

Given a term T_i of f, let A_i be the set of unnegated variables in T_i and let B_i be the set of negated variables in T_i . Let $x^{\bar{B}_i}$ refer to x restricted to variables outside of B_i .

We claim that

$$\operatorname{slice}_f(x) = \operatorname{Th}_{k/2+1}^k(x) \vee \bigvee_{i=1}^t \left(\operatorname{Th}_{k/2}^{k-|B_i|}(x^{\bar{B}_i}) \wedge \bigwedge_{i \in A_i} x_i \right),$$

where each $\operatorname{Th}_{k/2}^{k-|B_i|}$ is TRUE whenever there are at least k/2 ones in the variables outside of B_i .

If an input x has more than k/2 ones, the first threshold function will output TRUE, and if x has fewer than k/2 ones, every threshold function $(\operatorname{Th}_{k/2+1}^k$ and $\operatorname{Th}_{k/2}^{k-|B_i|}$ for all i) will output FALSE. If x has exactly k/2 ones $\operatorname{Th}_{k/2}^{k-|B_i|}$ will output 1 only if all the variables in B_i are 0. Thus, each $\operatorname{Th}_{k/2}^{k-|B_i|}(x) \wedge \bigwedge_{j \in A_i} x_j$ correctly computes T_i on the middle slice.

The functions $\operatorname{Th}_{k/2+1}^k$ and each $\operatorname{Th}_{k/2}^{k-|B_i|}$ can be computed by an $n^{O(1)}$ -size, depth-3 monotone formula with an OR on top [22]. Using the distributive law we can convert $\operatorname{Th}_{k/2}^{k-|B_i|}(x) \wedge \bigwedge_{j \in A_i} x_j$ to also be a $n^{O(1)}$ -size, depth-3 monotone formula with an OR on top. This OR can be collapsed with the (t+1)-wise OR yielding a $n^{O(1)}$ -size, depth-3 monotone formula for slice f.

We are now ready to prove Th. 3.5 by describing the second component.

Proof: Any *j*-junta can be represented by a DNF formula with 2^j terms, and hence $n^{O(1)}$ -size DNF formulas over $k = \log^2(n)/\log\log(n)$ can represent any $\log(n)$ -junta f over the first k variables. Thus, by Lemma 3.6, slice f can be computed by an $n^{O(1)}$ -size, depth-3 monotone formula.

Let $j = \log(n)$ (for simplicity we ignore the insignificant rounding errors for j and k). Given examples from a j-junta f, the algorithm C will generate examples for slice f and run the learning algorithm f0 with f0. Let f1 be the f2-accurate hypothesis returned by algorithm f3.

Let $z\{0,1\}^j$ be a setting of the j relevant variables of f, and let i be the number of ones in z. Then the fraction of the points in the middle slice that equals to z when restricted to the relevant variables is

$${k-j \choose k/2-i} / {k \choose k/2} \ge {k-j \choose k/2-j} / {k \choose k/2} = \frac{(k/2)\cdots(k/2-j+1)}{k\cdots(k-j+1)} > \left(\frac{k/2-j+1}{k}\right)^j \ge 2^{-j}(1-\frac{2j}{k})^j$$

$$\ge \frac{e^{-4j^2/k}}{2^j} = \Omega(\frac{1}{n\log n})$$
(2)

By Stirling's approximation the weight of the whole middle slice is $\Omega(1/\sqrt{k}) = \Omega(1/\log n)$ and therefore the points in the middle equal to z on the relevant variables comprise $\Omega(\frac{1}{n\log^2 n}) > 2/n^{3/2}$ fraction of the domain (for sufficiently large n). The same is also true for every every setting of j+1 variables since by equation (2), $\frac{e^{-4(j+1)^2/k}}{2^{j+1}} = \Omega(\frac{1}{n\log n})$.

Thus, by our choice of $\epsilon = 1/n^3$, the hypothesis h output by A, will output a correct answer with probability at least $1 - \delta$ for $\delta = 1/(2n^{3/2})$ when restricted to points in the middle slice that match the setting z of the relevant variables.

To find the relevant variables of the junta, C will do the following for all k(k-1)/2 unordered pairs of variables $\{x_a, x_b\}$. It will choose a random x in the middle slice where $x_a = 0$ and $x_b = 1$ or $x_a = 1$ and $x_b = 0$, and then compare the output of h(x) and h(x') where x' is x with the variables x_a and x_b flipped. It is important to note that x' remains in the middle slice. We call pairs that contain at least one relevant variable "candidates".

If x_a is relevant then x will match the setting z for which x_a changes the value when x_a is flipped with probability at least $p' \cdot 4/n^{3/2}$. This is true since there are two restriction of j+1 variables that will be included in the part of the middle slice from which we sampled randomly, each occurring with probability at least $2/n^{3/2}$. Hence if x_a is relevant and x_b is not, then $h(x) \neq h(x')$ with probability at least $p'(1-2\delta) \geq 2/n^{3/2}$. If both x_a and x_b are not relevant to the junta, then the probability that $h(x) \neq h(x') < 2\delta = 1/n^{3/2}$. Thus, given poly(n) examples we can identify all the candidate pairs. A relevant variable x_a is occurs in at least k-j candidate pairs whereas an irrelevant one can only appear in at most j < k-j candidate pairs. Hence C can identify all the relevant variables for the junta. Finally, C can try all $2^j = n$ different settings of the relevant variables and find the corresponding value of the target junta by taking the majority of values of h on random points in the middle slice with the corresponding setting of the relevant variables. Our analysis of the error on such restrictions implies that the correct truth table of the junta will be obtained with high probability. It is easy to verify that the running time of C is polynomial in n.

We note that Theorem 3.5 is incomparable to Corollary 3.4. It is easy to translate Theorem 3.5 into the SQ model. As a result we would obtain a superpolynomial lower bound for strong SQ learning monotone depth-3 formulas. This is true since a junta can be a parity function and there are at least $\binom{\log^2 n/\log\log n}{\log n} = n^{\Omega(\log\log n)}$ different parities in the above class of juntas. However both the lower bound and the accuracy parameter in Corollary 3.4 are substantially better. The better accuracy parameter is required for hardness amplification using a single additional level.

In the next section we introduce hardness amplification machinery that will enable us to extend the SQ learning hardness result to accuracy $\frac{1}{2} + o(1)$ (for depth-4 formulas).

4 Hardness Amplification for Uniform Distribution Learning

In [28] O'Donnell developed a general technique for hardness amplification. His approach, which may be viewed as a generalization of Yao's XOR lemma, gives a bound on the hardness of $g \otimes f = g(f(x_1), \ldots, f(x_k))$ where f is a "mildly" hard function and g is an arbitrary k-bit combining function.

At a high-level O'Donnell's proof has three components. The first component shows the existence of a circuit weakly approximating f over any δ -fraction of the domain whenever there exists a circuit for $g \otimes f$ that outperforms the expected bias of g (see definition below). The second part of O'Donnell's proof uses Impagliazzo's hard-core lemma [17] to obtain a δ -approximating circuit for f given circuits that weakly approximate f over any δ -fraction of the domain. The third component is the construction of combining functions that have the desired expected bias.

The first of the two primary obstacles in translating the result to learning framework is that the first component uses non-uniform advice that depends on f. This advice is, in general, not available to the learning algorithm². A substantial effort was devoted to obtaining (more) uniform versions of O'Donnell's result, most notably by Trevisan [39, 40]. Both of his reductions are uniform and do not use access to f but neither is sufficient for our purposes. The first reduction only amplifies to accuracy 3/4 + o(1) [39] and the second reduction uses a specific combining function of non-constant circuit depth [40]. At the same time a learning algorithm has a form of access to f (either random examples or statistical queries) and hence hardness amplification need not be independent of f (or "black-box"). Indeed, it is not hard to show that Trevisan's simpler and more uniform version of the first component [39] can be simulated using random examples of f in place of non-uniform advice [41]. However, it is unclear if this approach can be used with access only to statistical queries. To solve this problem we show a uniform and general version of the first component, namely an algorithm that given a circuit for $g \otimes f$ produces a short list of circuits that, with significant probability, contains a circuit weakly approximating f over any δ -fraction of the domain chosen in advance (Lem. 4.2). The algorithm does not use access to f but can use statistical queries to find the weakly approximating circuit among the candidate circuits.

The second obstacle is fact that in order to obtain a circuit for $g \otimes f$ a learning algorithm needs to simulate a statistical query oracle for $g \otimes f$ using a statistical query oracle for f (when learning from random examples simulation of random examples of $g \otimes f$ is trivial). We show that this is possible by giving a procedure that uses a function ψ that approximates f in place of f to answer statistical queries for $g \otimes f$. To create such ψ we use a form of gradient descent to f in which the equivalent of the gradient can be generated whenever ψ cannot be used in place of f to answer a statistical query for $g \otimes f$. The number of steps of the gradient descent is bounded and therefore this method produces correct answers to statistical queries for $g \otimes f$.

We replace the second component (the hard-core lemma) with "smooth boosting," a technique from computational learning theory which is known to be analogous to hard-core set constructions [23].

Finally, for the third component we need to show that appropriate hardness amplification can be obtained by using only one additional level of depth. By combining balanced Talagrand CNF and the complement of the "tribes" DNF with carefully chosen parameters and using analysis from [28, 25], we demonstrate hardness amplification from $1 - \log^{-\alpha} n$ accuracy to $1/2 + 2^{-\log^{\beta} n}$ accuracy using a small monotone CNF as a combining function, where α and β are positive constants (Lem. 5.1).

Notation and Terminology. For g a k-variable Boolean function and f an n-variable Boolean function, we write $g \otimes f$ to denote the nk-variable function $g(f(x_1), \ldots, f(x_k))$. For \mathcal{F} a class of n-variable functions and g a fixed k-variable combining function, we write \mathcal{F}^g to denote the class $\{g \otimes f : f \in \mathcal{F}\}$.

 $^{^2}$ To avoid a potential source of confusion we remark that while our SQ learning lower bounds are information-theoretic and hence allow obtaining an SQ learning algorithm that uses non-uniform advice, such advice cannot depend on f.

Let P_{δ}^k denote the distribution of random restrictions ρ on k coordinates, in which each coordinate is mapped independently to \star with probability δ , to 0 with probability $(1 - \delta)/2$, and to 1 with probability $(1 - \delta)/2$. We write h_{ρ} for the function given by applying restriction ρ to the function h. For a k-variable ± 1 -valued function h we write bias(h) to denote max $\{\Pr[h = -1], \Pr[h = 1]\}$. The *expected bias of h at \delta* is $\exp \operatorname{Bias}_{\delta}(h) = \mathbb{E}_{\rho}[\operatorname{bias}(h_{\rho})]$, where ρ is a random restriction from P_{δ}^k .

4.1 Hardness Amplification in the PAC Setting

The most significant use of non-uniformity in the first component of O'Donnell's proof is the lemma asserting that if one can predict a Boolean function on the hypercube noticeably better than the function's bias then there exist two adjacent points of the hypercube on which predictions are noticeably different [28]. We start by showing an average-case version of this lemma by proving that predictions need to be different on average over all edges of the hypercube.

Lemma 4.1 Given two functions $h: \{0,1\}^k \to \{-1,1\}$ and $p: \{0,1\}^k \to [0,1]$, suppose that

$$\frac{1}{2^k} \left(\sum_{x: h(x)=1} p(x) + \sum_{x: h(x)=-1} (1 - p(x)) \right) \ge \text{bias}(h) + \epsilon . \tag{3}$$

Then $\mathbf{E}_{(x,y)}[|p(x) - p(y)|] \ge 4\epsilon^2/k$ where (x,y) is a randomly and uniformly chosen edge in the Boolean hypercube $\{0,1\}^k$.

Proof: Let us assume without loss of generality that h is biased towards 1. By the Poincaré inequality over the discrete cube we know that for any function p over $\{0, 1\}^k$:

$$\mathbf{Var}[p] = \mathbf{E}[p^2] - \mathbf{E}[p]^2 \le \frac{k}{4} \mathbf{E}_{(x,y)} [(p(x) - p(y))^2].$$

The range of p is [0, 1], so $\mathbf{E}_{(x,y)}[|p(x) - p(y)|] \ge \mathbf{E}_{(x,y)}[(p(x) - p(y))^2] \ge 4\mathbf{Var}[p]/k$. It is now sufficient to prove that $\mathbf{Var}[p] \ge \epsilon^2$.

Let $b := bias(h) = Pr[h = 1] \ge 1/2$. We can rewrite Equation 3 as

$$b + \epsilon \le \frac{1}{2^k} \left(\sum_{x:h(x)=1} p(x) + \sum_{x:h(x)=-1} (1 - p(x)) \right)$$

= $\mathbf{E}[h(x)(p(x) - \mathbf{E}[p(x)])] + \mathbf{E}[p(x)]b + (1 - \mathbf{E}[p(x)])(1 - b).$

As $b \ge 1/2$, $b\mathbf{E}[p] + (1-b)(1-\mathbf{E}[p]) < b$, and thus $\mathbf{E}[h(x)(p(x)-\mathbf{E}[p(x)])] \ge \epsilon$. Because $h(x) \in \{-1,1\}$ we obtain $\mathbf{E}[|p-\mathbf{E}[p]|] \ge \epsilon$. Using the Cauchy-Schwarz inequality, we get $\mathbf{Var}[p] = \mathbf{E}[(p-\mathbf{E}[p])^2] \ge \mathbf{E}[|p-\mathbf{E}[p]|]^2 \ge \epsilon^2$.

Suppose we are given a circuit C that approximates $g \otimes f$ sufficiently well that it outperforms the expected bias of g. Roughly speaking, the following lemma shows that for any large enough set S, from C we can extract a circuit C' that weakly approximates f over the inputs in S.

Lemma 4.2 There is a randomized algorithm Extract with the following property: For any:

1. Parameters $0 < \epsilon \le 1/2$, $0 < \eta < 1$, subset $S \subseteq \{0,1\}^n$ such that $|S| = \eta 2^n$, Boolean function g over $\{0,1\}^k$, and

2. Boolean function f such that $bias(f) \le 1/2 + \epsilon/(8k)$ and $bias(f|_S) \le 1/2 + \epsilon^2/(4k)$, given a circuit C over $\{0,1\}^{k \times n}$ s.t.

$$\Pr_{\mathcal{U}^k}[C = g \otimes f] = \Pr_{(x_1, \dots, x_k) \in \{0,1\}^{k \times n}}[C(x_1, \dots, x_k) = g(f(x_1), \dots, f(x_k))] \ge \operatorname{ExpBias}_n(g) + \epsilon,$$

the algorithm Extract returns an n-input circuit C' such that with probability at least ϵ^2/k (over the randomness of Extract) we have $\mathbf{Pr}_{x \in S}[C'(x) = f(x)] \ge 1/2 + \epsilon^2/(2k)$. The algorithm Extract runs in time O(nk + |C|) and the circuit C' is of size at most |C|.

Proof: The algorithm Extract is very simple: it chooses $i \in [k]$ and $x_1, x_2, ..., x_{i-1}, x_{i+1}, ..., x_k \in \{0, 1\}^n$ randomly and uniformly. With probability 1/2 Extract outputs the circuit C_1 such that $C_1(x_i) = C(x_1, x_2, ..., x_{i-1}, x_i, x_{i+1}, ..., x_k)$, and with probability 1/2 it outputs the circuit $-C_1$.

We first assume for simplicity that f is unbiased on $\{0,1\}^n$ and S, and consequently also on $\{0,1\}^n \setminus S$. For a given restriction $\rho \in P_\eta^k$, we say that an input (x_1,\ldots,x_k) matches ρ if for all i we have $x_i \in S \Rightarrow \rho(i) = \star$, and $x_i \notin S \Rightarrow \rho(i) = f(x_i)$. Since we assumed that f is unbiased, the probability that x matches ρ is exactly the probability of ρ in P_η^k . Let c_ρ be the probability that C correctly computes $g \otimes f$ conditioned on the event that (x_1,\ldots,x_k) matches ρ . Then, by definition, $\mathbf{E}_\rho[c_\rho - \mathrm{bias}(g_\rho)] = \mathbf{Pr}_{\mathcal{U}^k}[C = g \otimes f] - \mathrm{ExpBias}_\eta(g) \geq \epsilon$. Let adv_ρ denote $\mathrm{max}\{0,c_\rho - \mathrm{bias}(g_\rho)\}$. Clearly, $\mathbf{E}_\rho[\mathrm{adv}_\rho] \geq \epsilon$. For a given ρ let $I_\rho = \{i \mid \rho(i) = \star\}$ and let $k_\rho = |I_\rho|$. For $y \in \{-1,1\}^{k_\rho}$, let $p_\rho(y)$ be the probability that C(x) = 1 conditioned on the event that (x_1,\ldots,x_k) matches ρ and $f(x_i) = y_i$ for each i such that $\rho(i) = \star$ (we refer to this condition as x matching (ρ,y)). We index the bits of y with subscripts from I_ρ . Note that under our assumption of f being unbiased on f0 the probability of f1 matching f2.

Then it follows that:

$$c_{\rho} = 2^{-k_{\rho}} \left(\sum_{y: g_{\rho}(y)=1} p_{\rho}(y) + \sum_{y: g_{\rho}(y)=-1} (1 - p_{\rho}(y)) \right).$$

If $\operatorname{adv}_{\rho} > 0$ then $c_{\rho} = \operatorname{bias}(g_{\rho}) + \operatorname{adv}_{\rho}$, and we can invoke Lemma 4.1 on $g_{\rho}(y)$ and $p_{\rho}(y)$. By this lemma, the expected value of $|p_{\rho}(z) - p_{\rho}(z^{i})|$ for a random edge (z, z^{i}) in $\{-1, 1\}^{k_{\rho}}$ is at least $\operatorname{4adv}_{\rho}^{2}/k_{\rho} \geq \operatorname{4adv}_{\rho}^{2}/k$. (Here z^{i} is z with the bit z_{i} flipped, where i is an element of I_{ρ} .) Therefore we have

$$\mathbf{E}_{\rho,i\in I_{\rho},z\in\{-1,1\}^{k_{\rho}}}[|p_{\rho}(z)-p_{\rho}(z^{i})|] \ge \mathbf{E}_{\rho,i\in I_{\rho},z\in\{-1,1\}^{k_{\rho}}}[4\mathrm{adv}_{\rho}^{2}/k] = 4\mathbf{E}_{\rho}[\mathrm{adv}_{\rho}^{2}]/k \ge 4\epsilon^{2}/k, \tag{4}$$

where we used $\mathbf{E}_{\rho}[\operatorname{adv}_{\rho}] \ge \epsilon$ and the Cauchy-Schwartz inequality to obtain the last inequality. We now observe that $|p_{\rho}(z) - p_{\rho}(z^{i})|$ is the expected correlation of C(x) and $f(x_{i})$ when conditioned on x matching either (ρ, z) or (ρ, z^{i}) . Specifically,

$$\begin{aligned} & \left| \mathbf{E}_{\mathcal{U}^k} [C(x) \cdot f(x_i) \mid x \text{ matches } (\rho, z) \text{ or } (\rho, z^i)] \right| \\ & = \left| \mathbf{E}_{\mathcal{U}^k} [C(x) \mid x \text{ matches } (\rho, z)] - \mathbf{E}_{\mathcal{U}^k} [C(x) \mid x \text{ matches } (\rho, z^i)] \right| \\ & = 2 \cdot |p_{\rho}(z) - p_{\rho}(z^i)|. \end{aligned}$$

We denote by $x^{(i)} = x_1, \dots, x_{i-1}, x_{i+1}, \dots, x_k$ and let z' be equal to z without the i-th bit. With this notation

$$\mathbf{E}_{\mathcal{U}^k}[C(x) \cdot f(x_i) \mid x \text{ matches } (\rho, z) \text{ or } (\rho, z^i)] = \mathbf{E}_{\mathcal{U}^{k-1}} \left[\mathbf{E}_{x_i \in S}[C(x) \cdot f(x_i)] \mid x^{(i)} \text{ matches } (\rho, z') \right]$$

and hence

$$\mathbf{E}_{\mathcal{U}^{k-1}}\left[\left|\mathbf{E}_{x_i \in S}\left[C(x) \cdot f(x_i)\right]\right| \mid x^{(i)} \text{ matches } (\rho, z')\right] \ge 2 \cdot |p_{\rho}(z) - p_{\rho}(z^i)|.$$

By combining this with equation (4) we obtain:

$$\mathbf{E}_{\rho,i\in I_{\rho},z\in\{-1,1\}^{k\rho}}\left[\mathbf{E}_{\mathcal{U}^{k-1}}\left[\left|\mathbf{E}_{x_{i}\in S}\left[C(x)\cdot f(x_{i})\right]\right|\mid x^{(i)} \text{ matches } (\rho,z')\right]\right] \geq 8\epsilon^{2}/k ,$$

which is equivalent to

$$\mathbf{E}_{i \in [k], x^{(i)} \sim \mathcal{U}^{k-1}} \left[\left| \mathbf{E}_{x_i \in S} \left[C(x) \cdot f(x_i) \right] \right| \right] \ge 8\epsilon^2 / k.$$

This implies that for randomly chosen i and $x^{(i)}$, with probability at least $4\epsilon^2/k$, we have that $C_1(x_i) = C(x_1, x_2, \dots, x_{i-1}, x_i, x_{i+1}, \dots, x_k)$ satisfies $\left|\mathbf{E}_{x_i \in S}[C_1(x_i) \cdot f(x_i)]\right| \ge 4\epsilon^2/k$. Hence either $\mathbf{Pr}_S[C_1 = f] \ge 1/2 + 2\epsilon^2/k$ or $\mathbf{Pr}_S[-C_1 = f] \ge 1/2 + 2\epsilon^2/k$. This means that with probability at least $2\epsilon^2/k$ the output of our algorithm satisfies $\mathbf{Pr}_S[C' = f] \ge 1/2 + 2\epsilon^2/k$. Thus we have proved the lemma under the additional assumption that f is unbiased on $\{0, 1\}^n$ and on S.

We now show that a similar but slightly weaker bound can obtained even when f is slightly biased. Let $X' = \{0,1\}^n \cup Z$ for a set Z of size $\epsilon 2^k/(4k)$. Let f' be a function over X' such that $f'|_{\{0,1\}^n} = f$ and $f'|_Z$ is defined in any way that makes f' balanced over X'. This is possible since bias $(f) \le 1/2 + \epsilon/(8k)$. Further let S' be $S \cup S_Z$ where $S_Z \subseteq X'$ is a set (disjoint from S) of size at most $|S| \epsilon^2/(2k)$ defined so as to make f' balanced on S'. This is possible since f' is balanced on X' and bias $f'|_S \le 1/2 + \epsilon^2/(4k)$. Let C_a be the circuit over X'^k that equals to C on $\{0,1\}^{nk}$ and 1 everywhere else. We now apply our argument to C_a and f'.

First,

$$\Pr_{\mathcal{U}'^k}[C_a = g \otimes f'] \geq \Pr_{\mathcal{U}'^k}[x \in \{0, 1\}^{nk}] \cdot \Pr_{\mathcal{U}^k}[C = g \otimes f] \geq (1 - \epsilon/(4k))^k (\operatorname{ExpBias}_{\eta}(g) + \epsilon)$$

$$\geq \operatorname{ExpBias}_{\eta}(g) + 3\epsilon/4 \geq \operatorname{ExpBias}_{\eta'}(g) + 3\epsilon/4 ,$$

where \mathcal{U}' denotes the uniform distribution over X' and $\eta' = |S'|/|X'| \leq \frac{1+\epsilon^2/(2k)}{1+\epsilon/(4k)}\eta \leq \eta$. Note that here we used the fact that for every function g, ExpBias $_{\eta}(g)$ is a monotone function of η [28]. Therefore Extract applied to C_a over X'^k gives us a circuit C'_a such that with probability at least $2(3\epsilon/4)^2/k$, we have $\Pr_{S'}[C'_a = f'] \geq 1/2 + 2(3\epsilon/4)^2/k$. There are at most $|S|\epsilon^2/(2k)$ points in $S' \setminus S$ and therefore $\Pr_{S}[C'_a|_S = f] \geq 1/2 + 2(3\epsilon/4)^2/k - \epsilon^2/(2k) > 1/2 + \epsilon^2/(2k)$. Finally, let \mathcal{D}' denote the distribution over $C'_a|_S$ that is produced when Extract is applied to C_a over X' conditioned on C'_a not hardwiring any x_j to an element of Z. We observe that \mathcal{D}' is exactly the same as the distribution that Extract produces when applied to C over $\{0,1\}^n$. However, if $\Pr_{S}[C'_a|_S = f] > 1/2 + \epsilon^2/(2k)$ then C'_a cannot have any x_j hardwired to an element of Z, since, by definition of C_a , hardwiring x_j to an element of Z turns the circuit into the constant-1 function, whereas we know that $\Pr_{S}[1 = f] \leq \operatorname{bias}(f|_S) < 1/2 + \epsilon^2/(2k)$. We can therefore conclude that Extract applied to C over $\{0,1\}^n$ will output C' that with probability at least $2(3\epsilon/4)^2/k > \epsilon^2/k$ (over the randomness of Extract) satisfies $\Pr_{S}[C' = f] \geq 1/2 + \epsilon^2/(2k)$.

As we will see below, two key properties of this lemma are that Extract is efficient and is oblivious of both f and S. The second property is crucial for hardness amplification in the SQ model. In the second part of the proof, we show how an algorithm A that learns the combined class \mathcal{F}^g to moderate accuracy can be used to obtain an algorithm B that learns the original class \mathcal{F} to high accuracy. This is exactly the well-studied "weak learning \implies strong learning" paradigm of boosting in computational learning theory (see [31, 32] for introductions to boosting). Roughly speaking, boosting algorithms are automatic procedures that can be used to convert any weak learning algorithm (that only achieves low accuracy slightly better than 1/2) into a strong learning algorithm (that achieves high accuracy close to 1). Boosting algorithms work by repeatedly running the weak learning algorithm under a sequence of carefully chosen probability

distributions $\mathcal{D}_1, \mathcal{D}_2, \ldots$, obtaining weak hypotheses h_1, h_2, \ldots If each hypothesis h_i has non-negligible accuracy under the distribution \mathcal{D}_i that was used to generate it, then the boosting guarantee ensures that the final hypothesis h (which combines h_1, h_2, \ldots) has high accuracy under the original distribution.

Since we require the set |S| to be "large" (recall the statement of Lemma 4.2), we will need to use a so-called "smooth boosting algorithm" such as the algorithm of [33]. A $1/\delta$ -smooth boosting algorithm is a boosting algorithm with the following property: if the original distribution is uniform over a finite domain X (as is the case for us here), then in learning to final accuracy δ , every distribution \mathcal{D}_i that the smooth boosting algorithm constructs will be " $1/\delta$ -smooth," meaning that it puts probability weight at most $\frac{1}{\delta} \cdot \frac{1}{|X|}$ on any example $x \in X$. Such $1/\delta$ -smooth distributions correspond naturally to large sets S (of size $\delta 2^n$) in Lemma 4.2.

So at a high level, we use a smooth boosting algorithm, and for each smooth distribution that it constructs we use Extract several times to generate a set of candidate weak hypotheses (recall that Extract constructs a "good" C' only with some nonnegligible probability). These hypotheses are then tested using uniform examples (filtered according to the current smooth distribution; since the distribution is smooth this does not incur much overhead), and we identify one which has the required nonnegligible accuracy. The boosting guarantee ensures that the combined hypothesis has accuracy $1-\delta$ under the original (uniform) distribution.

Having sketched the intuition for the second stage, we now state the main hardness amplification theorem for PAC learning.

Theorem 4.3 Let \mathcal{F} be a class of functions such that for every $f \in \mathcal{F}$, bias $(f) \leq 1/2 + \epsilon/(8k)$. Let A be a uniform distribution PAC learning algorithm that learns \mathcal{F}^g to accuracy $\text{ExpBias}_{\delta}(g) + \epsilon$. There exists a uniform-distribution PAC learning algorithm B that learns \mathcal{F} to accuracy $1 - \delta$ in time $O(T_1 \cdot T_2 \cdot poly(n, k, 1/\epsilon, 1/\delta))$ where T_1 is the time required to evaluate g and T_2 is the running time of A.

Proof: Let f denote the unknown target function. We will first simulate A to obtain a circuit C that is $(\text{ExpBias}_{\delta}(g) + \epsilon)$ -close to $g \otimes f$. To generate a random example of $g \otimes f$ we simply draw k random examples of $f: (x_1, \ell_1), \ldots, (x_k, \ell_k)$ and give the example $((x_1, \ldots, x_k), g(\ell_1, \ldots, \ell_k))$ to A.

We now use C to produce weak hypotheses on distributions produced by the $1/\delta$ -smooth boosting algorithm of Servedio [33] (here δ refers to the desired accuracy parameter).

Let $D_t(x)$ denote the distribution obtained at step t of boosting and let h_1, \ldots, h_{t-1} be the hypotheses obtained in the previous stages of boosting. Let $M = 2^n L_{\infty}(D_t)$ and let S_t denote a set obtained by including each point $x \in \{0, 1\}^n$ randomly with probability $D_t(x)/M$. As it is easy to see (e.g., [17]), for any function h fixed independently of the random choices that determine S_t , with probability at least $1 - 2^{-n/2}$ (over the choice of S_t) $|\Pr_{D_t}[h = f] - \Pr_{S_t}[h = f]| \le 2^{-n/2}$. Therefore for our purposes we can treat $\Pr_{S_t}[h = f]$ as equal to $\Pr_{D_t}[h = f]$.

If bias $(f|_{S_t}) \ge 1/2 + \epsilon^2/(4k)$ then $\Pr_{S_t}[f=b] \ge 1/2 + \epsilon^2/(4k)$ for $b \in \{-1,1\}$. Otherwise, by Lemma 4.2, with probability at least ϵ^2/k the algorithm Extract returns a circuit C_1 such that $\Pr_{S_t}[C_1=f] \ge 1/2 + \epsilon^2/(2k)$. As it is easy to see from the analysis of [33], the value $D_t(x)/M$ equals $\mu_t(f(x), h_1(x), \dots, h_{t-1}(x))$ for a fixed function μ_t defined by the boosting algorithm. This allows the learning algorithm to generate random examples from $D_t(x)$ by filtering random and uniform examples using μ_t . In particular, we can estimate $\Pr_{D_t}[h=f]$ to accuracy $\epsilon^2/(12k)$ and confidence 1/2 using $\tilde{O}(k^2/\epsilon^4\delta)$ random and uniform examples in order to test whether either -1, 1 or C' give a good weak hypothesis (the $1/\delta$ factor in the number of examples suffices because we are using a $1/\delta$ -smooth boosting algorithm). By repeating the execution of Extract a total of $O(\epsilon^{-2} \cdot k \log(k/\epsilon\delta))$ times we can ensure that with probability at least 2/3 this weak learning step is successful in all $O(k^2/(\epsilon^4\delta))$ boosting stages that the booster of [33] requires. This implies

that the boosting algorithm produces a $(1 - \delta)$ -accurate hypothesis with probability at least 2/3. It is easy to verify that the running time of this algorithm is as claimed.

Remark 4.4 This hardness amplification also applies to algorithms using membership queries since membership queries to $g \otimes f$ can be easily simulated using membership queries to f.

4.2 Hardness amplification in the Statistical Query setting

We now establish the SQ version of hardness amplification.

Theorem 4.5 Let \mathcal{F} be a class of functions such that for every $f \in \mathcal{F}$, bias $(f) \leq 1/2 + \epsilon/(8k)$. Let A be a uniform-distribution SQ-learning algorithm that learns \mathcal{F}^g to accuracy $\operatorname{ExpBias}_{\delta}(g) + \epsilon$ using queries of tolerance τ . There exists a uniform-distribution SQ learning algorithm B that learns \mathcal{F} to accuracy $1 - \delta$ in time $O(T_1 \cdot T_2 \cdot \operatorname{poly}(n, k, 1/\epsilon, 1/\delta))$ using queries of tolerance $\Omega(\delta \cdot \min\{\tau/k, \epsilon^2/k\})$, where T_1 is the time required to evaluate g and T_2 is the running time of A.

Proof: The main challenge in translating the result to SQ-learning is to simulate SQs for $g \otimes f$ using SQs for f. Given the circuit C we can proceed exactly as in the proof of Theorem 4.3 but use SQs of tolerance $\Omega(\delta \epsilon^2/k)$ to estimate the bias of f on S_t or to test whether the output of Extract is a weak hypothesis.

We now describe how to simulate statistical queries to $g \otimes f$. The distribution is known to be uniform therefore it is sufficient to answer only correlational statistical queries of A [8], namely, it is sufficient to be able to estimate $\mathbf{E}_{\mathcal{U}^k}[\phi \cdot (g \otimes f)]$ within $\tau/2$, where ϕ is a Boolean function over $\{0,1\}^{kn}$. To estimate $\mathbf{E}_{\mathcal{U}^k}[\phi \cdot (g \otimes f)]$ we plan to use random sampling with an approximation to f used in place of f. We refer to the approximation as $\psi_r(x)$. However, before doing so, we first test whether ψ_r is suitable to be used as a replacement. In the main technical claim we prove that if $\psi_r(x)$ cannot be used to replace f then we can find a way to update ψ_r to ψ_{r+1} which is closer to f than ψ_r in L_2 distance. The number of such updates will be bounded and therefore we will eventually obtain ψ_r that can be used in place of f. Formally, let $\psi_0(x) \equiv 0$ and for a function $\psi_r(x) \in \mathcal{F}_1^{\infty}$ we denote by $\Psi_r(x)$ the random $\{-1,1\}$ variable with expectation $\psi_r(x)$. We also denote by $g \otimes \Psi_r$ the random variable obtained by applying g to k evaluations of Ψ_r .

Lemma 4.6 For $i \in [k]$ and $y = y_1, \dots, y_{i-1}, y_{i+1}, \dots, y_k \in \{0, 1\}^n$ and any function ϕ over $\{0, 1\}^{kn}$, we denote $\phi_{i,y}(x_i) = \phi(y_1, y_2, y_{i-1}, x_i, y_{i+1}, \dots, y_k)$. Let $\lambda = |\mathbf{E}_{\mathcal{U}^k}[\phi \cdot g \otimes f] - \mathbf{E}_{\mathcal{U}^k,\Psi}[\phi \cdot g \otimes \Psi]|$. Then for randomly and uniformly chosen i, y, with probability at least $\lambda/(4k)$, $|\mathbf{E}_{\mathcal{U}}[\phi_{i,y}(x) \cdot f(x)] - \mathbf{E}_{\mathcal{U}}[\phi_{i,y}(x) \cdot \psi(x)]| \ge \lambda/(2k)$.

Proof: First we denote by $g \otimes f^{i,\Psi}$ the *i*-th hybrid between $g \otimes f$ and $g \otimes \Psi$. Namely, the randomized function $g \otimes f^{i,\Psi}(x) = g(f(x_1), \dots, f(x_i), \Psi(x_{i+1}), \dots, \Psi(x_k))$. Now, $g \otimes f^{k,\Psi} = g \otimes f$ and $g \otimes f^{0,\Psi} = g \otimes \Psi$. Hence we can write,

$$\left|\mathbf{E}_{\mathcal{U}^{k}}[\phi \cdot g \otimes f] - \mathbf{E}_{\mathcal{U}^{k},\Psi}[\phi \cdot g \otimes \Psi]\right| = k \cdot \left|\mathbf{E}_{i \in [k]}\left[\mathbf{E}_{\mathcal{U}^{k},\Psi}[\phi \cdot g \otimes f^{i,\Psi}] - \mathbf{E}_{\mathcal{U}^{k},\Psi}[\phi \cdot g \otimes f^{i-1,\Psi}]\right]\right|$$

We now split the random and uniform choice over $\{0,1\}^{kn}$ into choosing $y=y_1,\ldots,y_{i-1},y_{i+1},\ldots,y_k\in\{0,1\}^n$ and $x_i\in\{0,1\}^n$ randomly and uniformly.

$$\left| \mathbf{E}_{i \in [k], y} \left[\mathbf{E}_{\mathcal{U}, \Psi} [(\phi \cdot g \otimes f^{i, \Psi})_{i, y}(x_i)] - \mathbf{E}_{\mathcal{U}, \Psi} [(\phi \cdot g \otimes f^{i-1, \Psi})_{i, y}(x_i)] \right] \right| \ge \lambda/k . \tag{5}$$

We claim that

$$\mathbf{E}_{\mathcal{U},\Psi}[(\phi \cdot g \otimes f^{i,\Psi})_{i,y}(x_i)] = \mathbf{E}_{\mathcal{U}}\left[\phi_{i,y} \cdot \mathbf{E}_{\Psi}[(g \otimes f^{i,\Psi})_{i,y}(x_i)]\right] = \mathbf{E}_{\mathcal{U}}[\phi_{i,y} \cdot (\alpha_{i,y}f(x_i) + \beta_{i,y})].$$

Here $\alpha_{i,y}$ and $\beta_{i,y}$ are constants in [-1, 1].

To see this assume for simplicity that Ψ is deterministic. Then

$$(g \otimes f^{i,\Psi})_{i,y}(x_i) = g(f(y_1), \dots, f(y_{k-1}), f(x_i), f(y_{k-1}), \dots, f(y_k)).$$

All the variables of g are fixed except for the i-th and therefore this restriction of g equals $1, -1, f(x_i)$ or $-f(x_i)$. This corresponds to $\alpha_{i,y}, \beta_{i,y} \in \{-1, 0, 1\}$ and exactly one of them is non-zero. For randomized Ψ we obtain a fixed convex combination of the deterministic cases that can be represented by $\alpha_{i,y}, \beta_{i,y} \in [-1, 1]$. Similarly,

$$\mathbf{E}_{\mathcal{U},\Psi}[(\phi \cdot g \otimes f^{i-1,\Psi})_{i,y}(x_i)] = \mathbf{E}_{\mathcal{U}}[\phi_{i,y} \cdot (\alpha_{i,y}\psi(x_i) + \beta_{i,y})].$$

By substituting this into equation (5), we obtain

$$\left| \alpha_{i,y} \cdot \mathbf{E}_{i \in [k],y} \left[\mathbf{E}_{\mathcal{U}} [\phi_{i,y} \cdot f(x_i)] - \mathbf{E}_{\mathcal{U}} [\phi_{i,y} \cdot \psi(x_i)] \right] \right| \ge \lambda/k .$$

By the averaging argument, we obtain that with probability at least $\lambda/(4k)$ over the choice of i and y, $|\mathbb{E}_{\mathcal{U}}[\phi_{i,y} \cdot f(x_i)] - \mathbb{E}_{\mathcal{U}}[\phi_{i,y} \cdot \psi(x_i)]| \ge \lambda/(2k)$.

If $|\mathbf{E}_{\mathcal{U}^k}[\phi \cdot g \otimes f] - \mathbf{E}_{\mathcal{U}^k,\Psi_r}[\phi \cdot g \otimes \Psi_r]| \geq \tau/3$ then with probability at least $\tau/(12k)$ for a randomly chosen $\phi_{i,y}$, $|\mathbf{E}_{\mathcal{U}}[\phi_{i,y} \cdot f(x_i)] - \mathbf{E}_{\mathcal{U}}[\phi_{i,y} \cdot \psi(x_i)]| \geq \tau/(6k)$. Let $\tau' = \tau/(6k)$. Now we sample $\phi_{i,y}$ and test if $|\mathbf{E}_{\mathcal{U}}[\phi_{i,y} \cdot f(x_i)] - \mathbf{E}_{\mathcal{U}}[\phi_{i,y} \cdot \psi(x_i)]| \leq 2\tau'/3$ using a single SQ of tolerance $\tau'/6$ and an estimate of $\mathbf{E}_{\mathcal{U}}[\phi_{i,y} \cdot \psi(x_i)]$ within $\tau'/6$ obtained using random sampling. It is easy to see that by repeating this procedure $O(k \log (1/\Delta)/\tau)$ times and using $O(k \log (1/\Delta)/\tau)$ random samples we can ensure that with probability at least $1 - \Delta$ some $\phi_{i',y'}$ will pass the test whenever $|\mathbf{E}_{\mathcal{U}^k}[\phi \cdot g \otimes f] - \mathbf{E}_{\mathcal{U}^k,\Psi_r}[\phi \cdot g \otimes \Psi_r]| \geq \tau/3$ and also that $|\mathbf{E}_{\mathcal{U}}[\phi_{i',y'} \cdot f(x_{i'})] - \mathbf{E}_{\mathcal{U}}[\phi_{i',y'} \cdot \psi(x_{i'})]| \geq 2\tau'/3 - \tau'/3 = \tau'/3$ whenever $\phi_{i',y'}$ passes the test.

If the test was not passed then we estimate $\mathbf{E}_{\mathcal{U}^k,\Psi}[\phi \cdot g \otimes \Psi]$ within $\tau/6$ using random sampling and return the estimate as the answer to the query. Using $O(k \log (1/\Delta)/\tau)$ random samples we can ensure that with probability $1 - 2\Delta$ the returned estimate is $\tau/2$ close to $\mathbf{E}_{\mathcal{U}^k}[\phi \cdot g \otimes f]$.

Otherwise, we use such $\phi_{i',y'}$ to update ψ_r using the idea from Feldman's strong SQ characterization [14, Th.3.5]: $\psi_{r+1} = P_1(\psi_r + (\tau'/3) \cdot \phi_{i',y'})$. Here $P_1(a)$ is the function that equals a when $a \in [-1, 1]$ and equals sign(a) otherwise.

As is proved in [14, Cl.3.6], $|\mathbf{E}_{\mathcal{U}}[\phi_{i',y'} \cdot (f(x_{i'}) - \psi(x_{i'})]| \ge \tau'/3$ implies that $\mathbf{E}_{\mathcal{U}}[(f - \psi_{r+1})^2] \le \mathbf{E}_{\mathcal{U}}[(f - \psi_r)^2] - (\tau'/3)^2$. Therefore at most $O(k^2/\tau^2)$ such updates are possible giving an upper bound on the additional time required to produce the desired estimates to all the SQs of A (a similar bound can also be obtained using a different update method from [42]). For an appropriate $\Delta = \text{poly}(k, 1/\tau)$ we can make sure that the success probability is at least 2/3.

5 Amplified Hardness for SQ Learning of Depth-4 Monotone Formulas

We begin this section by showing how a refinement of the constructions and analysis from [28, 25] can be used to obtain a small monotone CNF with low expected bias. Specifically, we prove the following lemma.

Lemma 5.1 For every $0 < \gamma < 1/2$, there exists a circuit $C_{k,m}$ over k variables such that:

$$\operatorname{ExpBias}_{1/\sqrt{m}}(C_{k,m}) \leq \frac{1}{2} + 2^{-(\log n)^{\gamma}},$$

where $k = 2^{(\log n)^{\alpha}}$ and $m = \log^{2-\beta}(n)$ for $\gamma < \alpha < \beta/2 < 1/2$, and $C_{k,m}$ is computable by a monotone CNF of size $n^{o(1)}$.

Before we prove Lemma 5.1 we review some of the constructions and tools from [28, 25].

As described in [28], instead of calculating the expected bias of a function directly, one can use *noise* stability to estimate the expected bias. For $x \in \{0, 1\}^n$, let $N_{\delta}(x)$ be a random variable in $\{0, 1\}^n$ obtained by flipping each bit of x independently with probability δ .

Definition 5.2 *The* noise stability *of* f *at noise rate* δ *is:*

NoiseStab_{$$\delta$$}(f) = $\mathbf{E}[f(x) \cdot f(N_{\delta}(x))],$

where the probability is taken over the noise and a uniform choice of x.

The following two facts about the noise stability will be useful in proving Lemma 5.1.

Fact 5.3 *If f is a balanced function and g is any function, then*

NoiseStab_{$$\delta$$}($g \otimes f$) = NoiseStab _{$\frac{1-\text{NoiseStab}_{\delta}(f)}{2}$} (g).

The noise stability of a function provides a bound on its expected bias.

Lemma 5.4 ([28]) ExpBias<sub>2
$$\delta$$</sub> $(f) \le \frac{1}{2} + \frac{1}{2} \sqrt{\text{NoiseStab}_{\delta}(f)}$.

The first class of functions we consider are the Talagrand functions. Talagrand's function is a randomly constructed CNF formula on n inputs [38]. Specifically, it has $2^{\sqrt{n}}$ clauses of size \sqrt{n} , where the \sqrt{n} variables of each clause are selected independently and uniformly at random from [n]. We write \mathcal{T}_n to denote this probability distribution over n-variable CNFs. Mossel and O'Donnell showed that these functions are somewhat sensitive to very small amounts of noise. They prove:

Lemma 5.5 ([25]) For infinitely many values of n, there is a function T_n in the support of \mathcal{T}_n such that NoiseStab_{1/ \sqrt{n}} $(T_n) \le 1 - \Omega(1)$.

The functions of Lemma 5.5 are not necessarily balanced. We show that there is a balanced version of T_n that can be computed by a slightly larger CNF formula while preserving its noise stability.

Lemma 5.6 Let f be a monotone function over n variables represented by a size-s CNF, and let $0 < \eta < 1/2$ be such that $\Pr[f(x) \neq f(N_{\eta}(x))] \geq \delta$. Then there exists a balanced monotone function g over n + 2 variables represented by a size-(s + n) CNF formula such that $\Pr[g(x) \neq g(N_{\eta}(x))] \geq \delta/16$.

Proof: Our proof is a simple refinement of the proof from [25] that shows the existence of a balanced version of T_n (not necessarily computable by a "small" CNF).

The new function g will be defined on x and on two extra variables z_1 and z_2 . Let $m_a(x)$ be the monotone function that checks if the number represented by x in binary is at least the number represented by a. Let $g_a(zx) = (f(x) \lor z_1) \land (m_a(x) \lor z_2)$. Note that $g_a(01x) = f(x)$, $g_a(10x) = m_a(x)$, and g(11x) = TRUE. The function g_{0^n} is biased towards TRUE and g_{1^n} is biased towards FALSE and for consecutive a and a', g_a differs from $g_{a'}$ on exactly one point. Therefore there exists a value of a such that g_a is perfectly balanced. The function $m_a(x)$ can be represented as a short CNF as follows $\land_{i:a_i=1}(x_i \lor_{j:j< i,a_j=0} x_j)$. Thus $g_a(x)$ has at most n more clauses than f.

Finally, we have

$$\Pr[g(zx) \neq g(N_{\eta}(zx))] \ge \Pr[z_1 = 0 \land z_2 = 1 \land N_{\eta}(z) = 01] \cdot \Pr[f(x) \neq f(N_{\eta}(x))] \ge \delta/16,$$

and thus the noise stability of the function is preserved.

Applying Lemma 5.6 to the appropriate Talagrand function chosen from T_{n-2} gives us the following corollary.

Corollary 5.7 There exists an infinite family of balanced monotone functions $\operatorname{Tal}_n: \{0,1\}^n \to \{+1,-1\}$ with $\operatorname{NoiseStab}_{n^{-1/2}}(\operatorname{Tal}_n) \le 1 - \Omega(1)$, that can be represented by a $(2^{\sqrt{n}} + n)$ -size monotone CNF.

The second function we consider is the "tribes" function, which is sensitive to moderate noise. Given a positive integer b, let $n = n_b$ be the smallest positive integral multiple of b such that $(1 - 2^{-b})^{n/b} \le 1/2$, so n is roughly $(\ln 2)b2^b$ and b equals $\log n - \log \ln n + o(1)$. The "tribes" function on n variables is the read-once n/b-term monotone b-DNF

$$(x_1 \wedge \cdots \wedge x_b) \vee (x_{b+1} \wedge \cdots \wedge x_{2b}) \vee \cdots \vee (x_{n-b+1} \wedge \cdots \wedge x_n).$$

Lemma 5.8 ([25, 28]) Let Tribes_n: $\{0,1\}^n \to \{+1,-1\}$ denote the n-variable tribes function. For every constant $0 < \delta < 1$, we have

NoiseStab_{$$\delta$$}[Tribes _{n}] = $O(n^{-c_{\delta}})$

where $c_{\delta} > 0$ is a constant depending only on δ .

We will use the function Tribes $_n^{\dagger}$ which is the monotone complement of Tribes (*i.e.*, \neg Tribes($\neg x$)) which is given by a n/b-clause b-CNF and has the same noise stability as Tribes $_n$ by Boolean duality.

Now we will combine the tribes function with the balanced Talagrand function to obtain a function that is sensitive to very small amounts of noise, using the following lemma to bound its noise stability.

Lemma 5.9 For infinitely many values of k, m with m < k, the function $C_{k,m} := \text{Tribes}_{k/m}^{\dagger} \otimes \text{Tal}_m$ on k variables has:

NoiseStab<sub>1/
$$\sqrt{m}$$</sub>($C_{k,m}$) $\leq O\left(\left(\frac{m}{k}\right)^{c}\right)$

where c > 0 is some absolute constant.

Proof: By Corollary 5.7 we know that Tal_m is balanced so we can use Fact 5.3 to combine the noise stabilities of Tal_m and $\operatorname{Tribes}_{k/m}^{\dagger}$ (Lemma 5.8).

The function $C_{k,m}$ can in fact be computed by a small monotone CNF formula.

Lemma 5.10 Let $C_{k,m} := \text{Tribes}_{k/m}^{\dagger} \otimes \text{Tal}_m$. The function $C_{k,m}$ can be computed by a monotone CNF formula of size $\frac{k}{mb} \cdot (2^{\sqrt{m}} + m)^b$), where $b = \log(k/m) - \log \ln(k/m) + o(1)$.

Proof: $C_{k,m}$ can be represented as a depth-4 monotone formula by combining the k/(mb)-clause monotone b-CNF formula for the Tribes $_{k/m}^{\dagger}$ function with the $(2^{\sqrt{m}} + m)$ -clause monotone CNF for the Tal $_m$ function. Using the distributive rule we can rewrite the b-wise ORs at the bottom of the Tribes $_{k/m}^{\dagger}$ function with the $(2^{\sqrt{m}} + m)$ -wise ANDs of the Tal $_m$ functions to get $(2^{\sqrt{m}} + m)^b$ -wise ANDs followed by b-wise ORs. Now the two adjacent levels of ANDs and the two adjacent levels of ORs can be collapsed to obtain a size- $(\frac{k}{mb} \cdot (2^{\sqrt{m}} + m)^b)$, monotone CNF formula.

We are now ready to prove Lemma 5.1.

Proof: Let $C_{k,m}$ be the function from Lemma 5.9. Combining Lemmas 5.9 and 5.4, we can bound the expected bias by

$$\frac{1}{2} + O\left(\left(\frac{m}{k}\right)^{c}\right) = \frac{1}{2} + O\left(\log^{c(2-\beta) + \alpha}(n) \cdot 2^{-c(\log n)^{\alpha}}\right) = \frac{1}{2} + O(2^{-(\log n)^{\gamma}}),$$

recalling that $\gamma < \alpha < \beta/2 < 1/2$.

By Lemma 5.10, we know that $C_{k,m}$ can be computed by a monotone CNF formula of size $\frac{k}{mb} \cdot (2^{\sqrt{m}} + m)^b$, where $b = \log(k/m) - \log \ln(k/m) + o(1)$. Therefore, for our setting of k and m, the CNF formula will have size at most $2^{(\log n)^{1-\beta/2+\alpha} + (\log n)^{\alpha}} = n^{o(1)}$.

Coupled with our hardness result for depth-3 monotone formulas, Lemma 5.1 gives the claimed lower bound for depth-4 monotone formulas.

Theorem 5.11 For every $0 < \gamma < 1/2$, the class of $n^{o(1)}$ -size, depth-4 monotone formulas is not SQ-learnable to accuracy $\frac{1}{2} + 2^{-(\log n)^{\gamma}}$ in poly(n) time.

Proof: Let \mathcal{M} be the class from Theorem 3.1 that is defined over $\log^{2-\beta}(n)$ variables, let $C_{k,m}$ be the function from Lemma 5.9 defined over $k = 2^{(\log n)^{\alpha}}$ for $m = \log^{2-\beta}(n)$, and let $\gamma < \alpha < \beta/2 < 1/2$, $c < \beta/2$.

Let $\delta = 1/\log(n)^{1-c}$. As $\delta = o(1/\sqrt{m}) = o(1/\log^{1-\beta/2}(n))$, by Theorem 2.4 and Theorem 3.1, there is no poly(*n*)-time algorithm that can SQ-learn the class \mathcal{M} (of *m*-variable functions) to accuracy $1 - \delta$. Every function in \mathcal{M} has bias 1/2. Thus, by Theorem 4.5 and Lemma 5.1, there is no poly(*n*)-time algorithm that can SQ-learn the class $\mathcal{M}^{C_{k,m}}$ to accuracy

$$\operatorname{ExpBias}_{\delta}(C_{k,m}) \leq \frac{1}{2} + O(2^{-(\log n)^{\gamma}}).$$

Combining the size- $n^{o(1)}$, depth-2 monotone formulas constructed in Lemma 5.1 with ORs on the bottom layer with k copies of the size- $n^{o(1)}$, depth-3 monotone formulas constructed in Theorem 3.2 with ORs on the top layer gives us a $n^{o(1)}$ -size, depth-4 monotone formula as required.

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