

Date: Tuesday, March 26, 2019

## Hypothesis test continued...

So far:

- 1) State a hypothesis in terms of population metric.

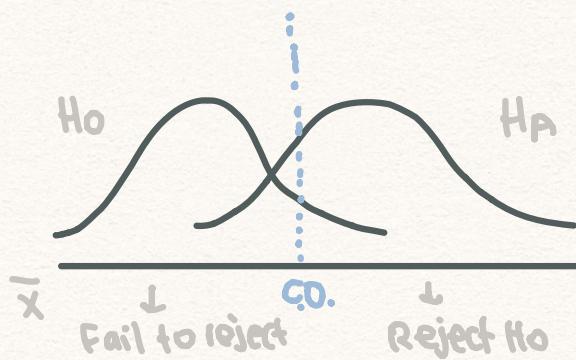
Ex:  $H_0: \mu = 0$  vs  $H_A: \mu > 0$

- 2) Calculate sample statistic that summarizes the metric of interest (Ex.  $\bar{x}$ )

- 3) Define rejection region based on significance.

Chosen by you.  
Represented as  $\alpha$ .

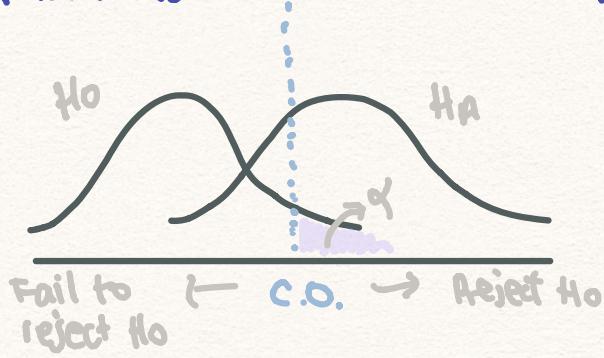
Example:  $p\text{-value} \leq \alpha$



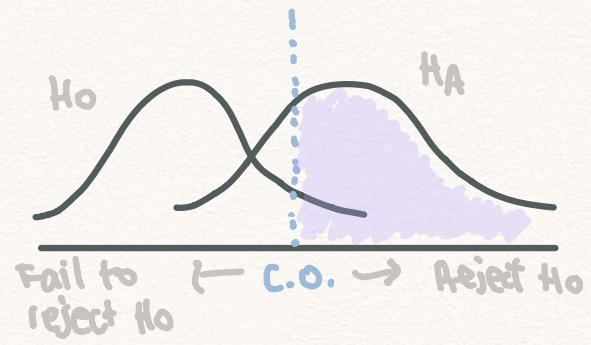
- ↳ We assume that the sample data has a probability distribution under both  $H_0$  and  $H_A$ .
- ↳ Our rejection region depends on a cut-off (C.O.)
- ↳ we reject  $H_0$  if  $\bar{x} > \text{C.O.}$  (Example data)

Not the same thing as  $\alpha$   
 $\alpha$  - p-value  
C.O. - Statistic

What does  $\alpha$  look like?



What does power look like?



$$\alpha = P(\text{Type I error}) \\ = P(\text{reject } H_0 \mid H_0 \text{ true})$$

$$\text{Power} = 1 - \beta \\ = 1 - P(\text{Type II error}) \\ = P(\text{reject } H_0 \mid H_A \text{ true})$$

- \* If we decrease  $\alpha$ , we decrease power
- \* Overall, there is a tradeoff  $\alpha$  and  $\beta$  for a fixed  $n$  and C.O.

For fixed  $n$  and C.O.

- 1)  $\alpha \uparrow \Rightarrow \beta \downarrow$  and power  $\uparrow$
- 2)  $\alpha \downarrow \Rightarrow \beta \uparrow$  and power  $\downarrow$

ALSO, the movement of  $\alpha$  will move the C.O. but which direction depends on the direction in  $H_A$ .

- What does  $n$  do?
- Work with the example of a population mean,  $\mu$ .

↳ Our sample statistic is  $\bar{x} = \frac{1}{n} \sum_{j=1}^n x_j$

By the Central Limit Theorem, we know that  $\bar{x} \sim N\left(\mu, \frac{\sigma^2}{n}\right)$

$$\sigma^2 = \text{var}(x_j)$$

### Example:

$$H_0: \mu = \mu_0 \quad \text{vs} \quad H_A: \mu = \mu_1$$

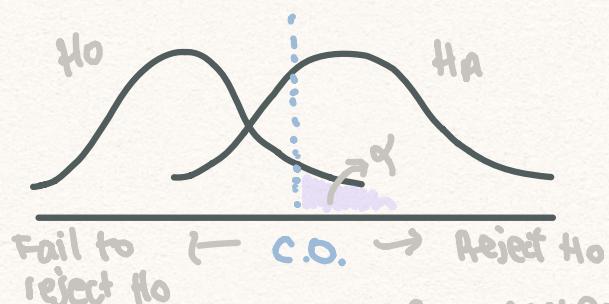
Under  $H_0$ , we have that:

$$\bar{x} \sim N(\mu_0, \frac{\sigma^2}{n})$$

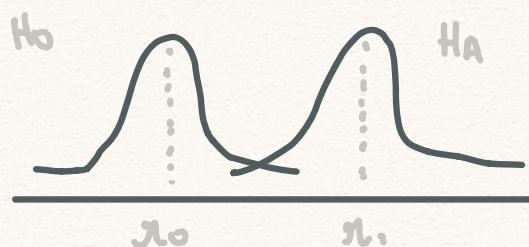
Under  $H_A$ , we have that:

$$\bar{x} \sim N(\mu_1, \frac{\sigma^2}{n})$$

In general,



↓  
→ Power increases because you have more separability  
 $n$  increases  $\Rightarrow$  Spread decreases under both  $H_0$  and  $H_A$ .



- \* The distribution under  $H_0$  and  $H_A$  depend on the sample statistic and are specifically the sampling distribution of the statistic given the value of the parameter.
- \* We use Central Limit theorem to approximate these distributions.

- \* In general, when  $n$  increases, we get more separability of  $H_0$  and  $H_A$  due to decreased variance. Thus, power increases!
- \* Note that for a fixed  $\alpha$ , the CO will also change.
- \* The test and calculations that you make depend now on a lot of moving parts:
  - 1) Metric being compared (mean, proportion, count, etc?)
  - 2) Do groups have equal variance?
  - 3) How long do you run the experiment / sample size\*
  - 4) The equality statements under  $H_0$  and  $H_A$ .
- \* The procedure is the same, but we need to think about our choice of  $\alpha$  and what power we'd like to have.
- \* Based on  $\alpha$  and desired power, you can then do sample size calculation to determine your needed  $n$ .  
 ↳ This becomes an important part of planning the experiment.

### Example: Comparing Proportions

Question: Does your ability to pass a test depend on which section you take?

Data:  $y_{ij} = \begin{cases} 1 & \text{, student } i \text{ passes from section } j \\ 0 & \text{, student } i \text{ fails from section } j \end{cases}$   
 $j = \{1, 2\}$

(Section 1 is the cool one!!!)



)

Model: Bernoulli

$y_{ij} \sim \text{Bern}(\pi_j)$

Probability of pairing in session j

$i = 1, \dots, n_j$   
 $j = 1, 2$

Sample statistic:  $\hat{\pi}_j = \frac{1}{n_j} \sum_{i=1}^{n_j} y_{ij}$  } unbiased estimator for  $\pi_j$

By CLT,  $\hat{\pi}_j \sim N(\pi_j, \text{var}(\hat{\pi}_j))$

$$\begin{aligned}\text{var}(\hat{\pi}_j) &= \frac{1}{n_j^2} \sum_{i=1}^{n_j} \text{var}(y_{ij}) \quad \text{indep if indep.} \\ &= \frac{1}{n_j^2} \sum_{i=1}^{n_j} \pi_j (1-\pi_j) \\ &= \frac{1}{n_j^2} n_j (\pi_j (1-\pi_j)) \\ &= \frac{1}{n_j} \pi_j (1-\pi_j)\end{aligned}$$

Hypothesis test:

$H_0: \pi_1 = \pi_2$	$\equiv$	$H_0: \pi_1 - \pi_2 = 0$
$H_A: \pi_1 \neq \pi_2$		equivalent $H_A: \pi_1 - \pi_2 = \delta$ for some $\delta \neq 0$ .

What we actually care about is the difference in population parameters in each section, ie effect size.

$\delta = \text{effect size!}$

What we care about is  $\delta = \pi_1 - \pi_2$   
 $\Rightarrow$  our statistic is  $\hat{\pi}_1 - \hat{\pi}_2$ .

From our previous CLT, we can calculate the distribution for  $\hat{p}_1 - \hat{p}_2$ :

$$\hat{p}_1 - \hat{p}_2 \sim N\left(\pi_1 - \pi_2, \frac{\pi_1(1-\pi_1)}{n_1} + \frac{\pi_2(1-\pi_2)}{n_2}\right)$$

Quick fact:  
 $X_1 \sim N(\mu_1, \sigma_1^2)$ ;  $X_2 \sim N(\mu_2, \sigma_2^2)$   
and  $X_1, X_2$  independent  
 $\Rightarrow X_1 - X_2 \sim N(\mu_1 - \mu_2, \sigma_1^2 + \sigma_2^2)$

$$\hat{z} = \frac{(\hat{p}_1 - \hat{p}_2) - (\pi_1 - \pi_2)}{\sqrt{\frac{\pi_1(1-\pi_1)}{n_1} + \frac{\pi_2(1-\pi_2)}{n_2}}} \sim N(0,1)$$

Sample statistic  $\hat{z}^*$

Note: We have to estimate  $\pi_j$  w/  $p_j$  to calculate  $\hat{z}^*$

↳ We want a cutoff so that

$$P(\text{rejecting } H_0 \mid H_0 \text{ true}) = \alpha$$

$$P(|\hat{z}^*| > c) = \alpha$$

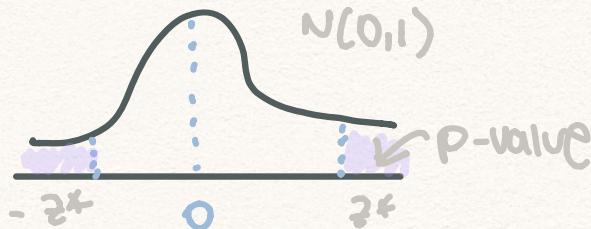
$$P(\hat{z}^* > c \text{ or } \hat{z}^* < -c) = \alpha$$

↓  
We can calculate this  
using quantiles of a  $N(0,1)$

↳ We can also use a p-value:

$P(\text{observing a sample stat more extreme than what we just saw under } H_0)$

$$= P(|z| > z^*) \text{ where } z \sim N(0,1)$$



If p-value  $< \alpha$ , we reject  $H_0$ .

Next time: Calculate power/sample size & multiple comparisons.