Expectation and Variance of a Random Variable

Recall: The Discrete Case

• If X is a discrete random variable having a probability mass function p(x), then the expected value of X is defined by

$$E[X] = \sum_{x:p(x)>0} xp(x)$$

The expected value of X is a weighted average of the possible values that X can take on; each value being weighted by the probability that X assumes that value.

Example: Toss a die

$$p(1) = p(2) = p(3) = p(4) = p(5) = p(6) = \frac{1}{6}$$

$$E[X] = 1 \times \frac{1}{6} + 2 \times \frac{1}{6} + 3 \times \frac{1}{6} + 4 \times \frac{1}{6} + 5 \times \frac{1}{6} + 6 \times \frac{1}{6}$$
= 3.5

Recall: The Discrete Case

• If X is a discrete random variable having a probability mass function p(x), then the variance of X is defined by

$$Var[X] = E[(X - \mu)^2], \mu = E[X]$$

The variance of *X* measures the expected square of the deviation of *X* from its expected value.

$$Var(X) = \sigma^{2} = E[(X - \mu)^{2}]$$

$$= E[X^{2} - 2.X.\mu + \mu^{2}]$$

$$= E[X^{2}] - 2\mu E[X] + (E[X])^{2}$$

$$= E[X^{2}] - 2E[X]E[X] + (E[X])^{2}$$

$$= E[X^{2}] - (E[X])^{2}$$

Recall: The Discrete Case

ullet If X is a discrete random variable having a probability mass function p(x), then the variance of X is defined by

$$Var[X] = E[(X - E[X])^2] = E[X^2] - (E[X])^2$$

Example: Toss a die

• Example: Toss a die
$$E[X^2] = 1 \times \frac{1}{6} + 2^2 \times \frac{1}{6} + 3^2 \times \frac{1}{6} + 4^2 \times \frac{1}{6} + 5^2 \times \frac{1}{6} + 6^2 \times \frac{1}{6}$$
$$E[X^2] = \frac{91}{6}$$

$$Var[X] = \frac{91}{6} - \frac{49}{4} = \frac{182 - 147}{12} = \frac{35}{12} \approx 2.9167$$

Recall: Bernoulli Random Variable

• If X is a Bernoulli random variable with parameter p, i.e., $p(\mathbf{0}) = \mathbf{1} - p$, $p(\mathbf{1}) = p$

$$E[X] = \mathbf{0}(\mathbf{1} - p) + \mathbf{1}(p) = p$$

$$E[X^2] = 0^2(1-p) + 1^2(p) = p$$

$$Var[X] = p - p^2 = p(1 - p)$$

Recall: Binomial Random Variable

ullet If X is binomially distributed with parameters n and p

Refer slides on Discrete random variables posted earlier

• Var(X) = np(1-p) [Proof Next Slide]

with
$$E[X] = \sum_{i=0}^{n} ip(i)$$

$$= \sum_{i=0}^{n} i \binom{n}{i} p^{i} (1-p)^{n-i}$$

$$= \sum_{i=1}^{n} \frac{in!}{(n-i)! i!} p^{i} (1-p)^{n-i}$$

$$= \sum_{i=1}^{n} \frac{n!}{(n-i)! (i-1)!} p^{i} (1-p)^{n-i}$$

$$= np \sum_{i=1}^{n} \frac{(n-1)!}{(n-i)! (i-1)!} p^{i-1} (1-p)^{n-i}$$

$$= np \sum_{k=0}^{n-1} \binom{n-1}{k} p^{k} (1-p)^{n-1-k}$$

$$= np [p + (1-p)]^{n-1}$$

$$= np$$

Recall: Variance of a Binomial Random Variable

$$\begin{split} \mathsf{E}\left(X^{2}\right) &= \sum_{k\geq 0}^{n} k^{2} \binom{n}{k} p^{k} q^{n-k} \\ &= \sum_{k=0}^{n} kn \binom{n-1}{k-1} p^{k} q^{n-k} \\ &= np \sum_{k=1}^{n} k \binom{n-1}{k-1} p^{k-1} q^{(n-1)-(k-1)} \\ &= np \sum_{j=0}^{m} (j+1) \binom{m}{j} p^{j} q^{m-j} \\ &= np \left(\sum_{j=0}^{m} j \binom{m}{j} p^{j} q^{m-j} + \sum_{j=0}^{m} \binom{m}{j} p^{j} q^{m-j}\right) \\ &= np \left(\sum_{j=0}^{m} m \binom{m-1}{j-1} p^{j} q^{m-j} + \sum_{j=0}^{m} \binom{m}{j} p^{j} q^{m-j}\right) \\ &= np \left((n-1) p \sum_{j=1}^{m} \binom{m-1}{j-1} p^{j-1} q^{(m-1)-(j-1)} + \sum_{j=0}^{m} \binom{m}{j} p^{j} q^{m-j}\right) \\ &= np \left((n-1) p(p+q)^{m-1} + (p+q)^{m}\right) \\ &= np \left((n-1) p+1\right) \\ &= n^{2} p^{2} + np \left(1-p\right) \end{split}$$

Definition of Binomial Distribution: p+q=1

Factors of Binomial Coefficient:
$$k \binom{n}{k} = n \binom{n-1}{k-1}$$

Change of limit: term is zero when k-1=0

putting
$$j=k-1, m=n-1$$

splitting sum up into two

Binomial Theorem

as p+q=1

Factors of Binomial Coefficient:
$$jinom{m}{j}=minom{m-1}{j-1}$$

Change of limit: term is zero when j-1=0

$$\mathsf{var}\,(X) \ = \ \mathsf{E}\,\big(X^2\big) - (\mathsf{E}\,(X))^2$$

$$= \ np\,(1-p) + n^2p^2 - (np)^2$$

by algebra
$$= np(1-p)$$

• If X is Poisson random variable with parameters λ

•
$$Var(X) = \lambda$$

$$E[X^2] = E[X(X-1) + X] = E[X(X-1)] + \lambda$$

$$E[X(X-1)] = \sum_{x=0}^{\infty} (x)(x-1)e^{-\lambda} \frac{\lambda^x}{x!}$$

$$\stackrel{\bullet}{E}[X(X-1)] = \sum_{x=0}^{\infty} (x)(x-1)e^{-\lambda} \frac{\lambda^x}{x!}$$

$$E[X(X-1)] = \sum_{x=2}^{\infty} (x)(x-1)e^{-\lambda} \frac{\lambda^x}{x!}$$

$$E[X(X-1)] = \sum_{x=2}^{\infty} e^{-\lambda} \frac{\lambda^x}{(x-2)!}$$

$$\stackrel{\bullet}{E}[X(X-1)] = \sum_{x=2}^{\infty} e^{-\lambda} \frac{\lambda^x}{(x-2)!}$$

$$E[X(X-1)] = \lambda^{2}e^{-\lambda} \sum_{x=2}^{\infty} \frac{\lambda^{x-2}}{(x-2)!}$$

$$E[X(X-1)] = \lambda^2 e^{-\lambda} e^{\lambda}$$

$$E[X^2] = \lambda^2 + \lambda$$

$$Var[X] = \lambda^2 + \lambda - \lambda^2 = \lambda$$

$$p(X = x|p,r) = {x-1 \choose r-1} p^r (1-p)^{x-r}, x = r, r+1, ...$$

$$E[X] = \sum_{r=r}^{\infty} x {x-1 \choose r-1} p^r (1-p)^{x-r}$$

$$E[X] = \sum_{r=r}^{\infty} x \frac{(x-1)!}{(x-r)! (r-1)!} p^r (1-p)^{x-r}$$

Let,
$$k = x - r$$

$$\overset{\bullet}{E}[X] = \sum_{r=r}^{\infty} \frac{x!}{(x-r)! (r-1)!} p^{r} (1-p)^{x-r}$$

Let,
$$k = x - r$$

$$E[X] = \sum_{k=0}^{\infty} \frac{(k+r)!}{k! (r-1)!} p^{r} (1-p)^{k}$$

$$E[X] = r \sum_{k=0}^{\infty} \frac{(k+r)!}{k! \, r!} p^{r} (1-p)^{k}$$

$$E[X] = r p^r \sum_{k=0}^{\infty} {k+r \choose k} (1-p)^k \qquad {-\beta-1 \choose k} = (-1)^k {k+\beta \choose k}.$$

$$E[X] = rp^{r} \sum_{k=0}^{\infty} {\binom{-(r+1)}{k}} (-1)^{k} (1-p)^{k}$$

$$\overset{\bullet}{E}[X] = r p^r \sum_{k=0}^{\infty} {\binom{-(r+1)}{k}} (-1)^k (1-p)^k$$

$$E[X] = rp^{r}(1-(1-p))^{-(r+1)}$$

$$E[X] = rp^r(p)^{-(r+1)} = \frac{r}{p}$$

$$Var[X] = \frac{r(1-p)}{p^2}$$

Recall: Geometric Random Variable

 If X a geometric random variable having parameter p.

$$E[X] = \sum_{n=1}^{\infty} np(1-p)^{n-1}$$

• Setting, q=p-1

$$E[X] = p \sum_{n=1}^{\infty} nq^{n-1}$$

Recall: Geometric Random Variable

$$E[X] = p \sum_{n=1}^{\infty} nq^{n-1}$$

$$E[X] = p \sum_{n=1}^{\infty} \frac{d}{dq} q^{n}$$

$$E[X] = p \frac{d}{dq} \left(\sum_{n=1}^{\infty} q^n \right)$$

Recall: Geometric Random Variable

$$\stackrel{\bullet}{E}[X] = p \frac{d}{dq} \left(\sum_{n=1}^{\infty} q^n \right) \qquad \frac{dy}{dx} = \frac{d}{dx} \left(\frac{u}{v} \right)
E[X] = p \frac{d}{dq} \left(\frac{q}{1-q} \right) \qquad \frac{dy}{dx} = \frac{v \frac{d}{dx}(u) - u \frac{d}{dx}(v)}{v^2}$$

$$E[X] = p\left(\frac{(1-q)(1)-q(-1)}{(1-q)^2}\right) = p\left(\frac{1}{(1-q)^2}\right)$$

$$E[X] = \frac{1}{p}$$

$$Var[X] = \frac{1-p}{p^2}$$

Variable

► PMF

•
$$P(X = q) = \frac{\binom{K}{q}\binom{N-K}{t-q}}{\binom{N}{t}}$$

$$E[X] = \sum_{q=0}^{K} q \frac{\binom{K}{q} \binom{N-K}{t-q}}{\binom{N}{t}}$$

$$E[X] = \sum_{q=1}^{K} q \frac{\binom{K}{q} \binom{N-K}{t-q}}{\binom{N}{t}}$$

Variable

$$E[X] = \sum_{q=1}^{K} q \frac{\binom{K}{q} \binom{N-K}{t-q}}{\binom{N}{t}}$$

$$E[X] = \sum_{q=1}^{K} q \frac{\binom{K}{q} \binom{N-K}{t-q}}{\frac{N}{t} \binom{N-1}{t-1}}$$

$$E[X] = \sum_{q=1}^{K} \frac{K {K-1 \choose q-1} {N-K \choose t-q}}{\frac{N}{t} {N-1 \choose t-1}}$$

Variable

$$E[X] = \sum_{q=1}^{K} \frac{K {K-1 \choose q-1} {N-K \choose t-q}}{\frac{N}{t} {N-1 \choose t-1}}$$

$$E[X] = \frac{Kt}{N} \sum_{q=1}^{K} \frac{\binom{K-1}{q-1} \binom{N-K}{t-q}}{\binom{N-1}{t-1}}$$

Let,
$$j = q - 1$$

Variable

$$E[X] = \frac{Kt}{N} \sum_{q=1}^{K} \frac{\binom{K-1}{q-1} \binom{N-K}{t-q}}{\binom{N-1}{t-1}}$$

Let,
$$j = q - 1$$

$$E[X] = \frac{Kt}{N} \sum_{j=0}^{K-1} \frac{\binom{K-1}{j} \binom{N-K}{t-j-1}}{\binom{N-1}{t-1}}$$

Variable

$$E[X] = \frac{Kt}{N} \sum_{j=0}^{K-1} \frac{\binom{K-1}{j} \binom{N-K}{t-j-1}}{\binom{N-1}{t-1}}$$

Probability that when you draw t-1 balls from an urn that contains K-1 red and N-k white, you will get exactly j red balls.

$$E[X] = \frac{Kt}{N}$$

1

$$Var[X] = \frac{tK}{N} \left(\frac{N-K}{N}\right) \left(\frac{N-t}{N-1}\right)$$

Continuous Random Variable

Continuous Random

• If X is a continuous random variable having a probability density function f(x), then the expected value of X is defined by

$$E[X] = \int_{-\infty}^{\infty} x f(x) dx$$

Variance

$$Var[X] = \int_{-\infty}^{\infty} (x - E[X])^2 f(x) dx = E[X^2] - E[X]$$

Recall: Uniform Random Variable

• Expectation of a random variable uniformly distributed over (α, β)

$$E[X] = \int_{\alpha}^{\beta} \frac{x}{\beta - \alpha} dx$$

$$= \left[\frac{x^2}{2(\beta - \alpha)} \right]_{\alpha}^{\beta}$$

$$= \frac{\beta^2 - \alpha^2}{2(\beta - \alpha)}$$

$$= \frac{\beta + \alpha}{2}$$

$$f(x) = \begin{cases} \frac{1}{\beta - \alpha} & for \beta \le x \le \alpha \\ 0 & elsewhere \end{cases}$$

Recall: Uniform Random Variable

• Expectation of a random variable uniformly distributed over (α, β)

$$Var[X] = \int_{\alpha}^{\beta} \frac{x^2}{\beta - \alpha} dx - \left(\frac{\beta + \alpha}{2}\right)^2$$

$$Var[X] = \left[\frac{x^3}{3(\beta - \alpha)}\right]_{\alpha}^{\beta} - \left(\frac{\beta + \alpha}{2}\right)^2$$

$$Var[X] = \frac{\beta^2 - \alpha\beta + \alpha^2}{3} - \left(\frac{\beta + \alpha}{2}\right)^2$$

Recall: Uniform Random Variable

$$Var[X] = \frac{\beta^2 - \alpha\beta + \alpha^2}{3} - \left(\frac{\beta + \alpha}{2}\right)^2$$

$$Var[X] = \frac{\beta^2 - \alpha\beta + \alpha^2}{3} - \frac{\beta^2 + \alpha^2 + 2\alpha\beta}{4}$$

$$Var[X] = \frac{4\beta^2 - 4\alpha\beta + 4\alpha^2 - 3\beta^2 - 3\alpha^2 - 6\alpha\beta}{12}$$

$$Var[X] = \frac{\beta^2 - 2\alpha\beta + \alpha^2}{12} = \frac{(\beta - \alpha)^2}{12}$$

Recall: Exponential Random Variable

• Let X be exponentially distributed with parameter λ.

$$E[X] = \int_0^\infty x \lambda e^{-\lambda x} dx$$

• Integrating by parts $(u = x, dv = \lambda e^{-\lambda x})$ yields

$$E[X] = \left[-xe^{-\lambda x}\right]_0^{\infty} + \int_0^{\infty} e^{-\lambda x} dx$$

$$E[X] = 0 - \left[\frac{e^{-\lambda x}}{\lambda}\right]_0^{\infty} = \frac{1}{\lambda}$$

$$f(x) = \begin{cases} \lambda e^{-\lambda x} & \text{if } x \ge 0 \\ 0 & \text{if } x < 0 \end{cases}$$

$$Var[X] = \frac{1}{\lambda^2}$$

Recall: Normal Random Variable

$$\stackrel{\bullet}{E}[X] = \frac{1}{\sqrt{2\pi\sigma}} \int_{-\infty}^{\infty} xe^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

$$f(x) = \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, -\infty < x < \infty$$

• Writing x as $(x - \mu) + \mu$ yields

$$E[X] = \frac{1}{\sqrt{2\pi\sigma}} \int_{-\infty}^{\infty} (x - \mu) e^{-\frac{(x - \mu)^2}{2\sigma^2}} dx + \mu \frac{1}{\sqrt{2\pi\sigma}} \int_{-\infty}^{\infty} e^{-\frac{(x - \mu)^2}{2\sigma^2}} dx$$

• Letting $y = x - \mu$

$$E[X] = \frac{1}{\sqrt{2\pi\sigma}} \int_{-\infty}^{\infty} y e^{-\frac{y^2}{2\sigma^2}} dy + \mu \frac{1}{\sqrt{2\pi\sigma}} \int_{-\infty}^{\infty} f(x) dx$$

Recall: Normal Random Variable

$$\stackrel{\bullet}{E}[X] = \frac{1}{\sqrt{2\pi\sigma}} \int_{-\infty}^{\infty} y e^{-\frac{y^2}{2\sigma^2}} dy + \mu \frac{1}{\sqrt{2\pi\sigma}} \int_{-\infty}^{\infty} f(x) dx$$

$$E[X] = 0 + \mu \frac{1}{\sqrt{2\pi\sigma}} \int_{-\infty}^{\infty} f(x) dx = \mu$$

$$Var[X] = \sigma^2$$

Recall: Gamma Random Variable

$$\bullet E[X] = \frac{\alpha}{\lambda}$$

•
$$Var[X] = \frac{\alpha}{\lambda^2}$$

$$f(x) = \begin{cases} \frac{\lambda e^{-\lambda x} (\lambda x)^{\alpha - 1}}{\Gamma(\alpha)} & \text{if } x \ge 0\\ 0 & x < 0 \end{cases}$$

Reference

- Lecture notes on Probability Theory by Phanuel Mariano
- Introduction to Probability Models, Sheldon M. Ross, Tenth Edition