# Package 'sphunif'

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Type Package

Title Uniformity Tests on the Circle, Sphere, and Hypersphere

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**Description** Implementation of uniformity tests on the circle and (hyper)sphere. The main function of the package is unif test(), which conveniently collects more than 35 tests for assessing uniformity on  $S^{p-1} = \{x \text{ in } R^p : ||x|| = 1\}, p \ge 2.$  The test statistics are implemented in the unif\_stat() function, which allows computing several statistics for different samples within a single call, thus facilitating Monte Carlo experiments. Furthermore, the unif\_stat\_MC() function allows parallelizing them in a simple way. The asymptotic null distributions of the statistics are available through the function unif\_stat\_distr(). The core of 'sphunif' is coded in C++ by relying on the 'Rcpp' package. The package also provides several novel datasets and gives the replicability for the data applications/simulations in García-Portugués et al. (2021) <a href="https://doi:10.1007/978-3-030-69944-4\_12">doi:10.1007/978-3-030-69944-4\_12</a>, García-Portugués et al. (2023) <doi:10.3150/21-BEJ1454>, García-Portugués et al. (2024) <doi:10.48550/arXiv.2108.09874>, and Fernández-de-Marcos and García-Portugués (2024) <doi:10.48550/arXiv.2405.13531>.

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BugReports https://github.com/egarpor/sphunif

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spnunii-package	•	•	•	•	٠	•	٠	•	•	•	 •	•	•	٠	•	•	 •	•	٠	•	 	٠	٠	•	•	•	•	•	•	•	•	3
angles_to_sphere .																					 											4
avail_tests																					 											5
$A_{theta_x \dots \dots$																					 											5
cir_coord_conv																					 											7
cir_gaps																					 											8
cir_stat_Kuiper																					 											9
comets																					 											15
craters																					 											18
F_from_f																					 											19
Gauss_Legen																					 											20
Gegenbauer																																22
harmonics																					 											27
int_sph_MC																					 											28
locdev																																29
planets																					 											33
Pn																																34
proj_unif																																38
Psi																					 											39
p_Kolmogorov																																41
p_sph_stat_Bingham																																48
rhea																					 											52
r_alt																					 											53
r_unif																					 											56
Sobolev																					 											57
Sobolev_coefs																					 											61
sph_stat_Rayleigh .																					 											62
sph_stat_Sobolev .																					 											66
unif_cap																					 											67
unif_stat																																69
unif_stat_distr																					 											71

sphunif-paci	kana	2
spiiuiiii-paci	Kage	-

Index																							92
	venus . wschisq																						
	unif_stat_ unif_test	_																					

sphunif-package

sphunif: Uniformity Tests on the Circle, Sphere, and Hypersphere

## Description

Implementation of uniformity tests on the circle and (hyper)sphere. The main function of the package is  $\operatorname{unif\_test}$ , which conveniently collects more than 35 tests for assessing uniformity on  $S^{p-1} = \{\mathbf{x} \in R^p : ||\mathbf{x}|| = 1\}, p \geq 2$ . The test statistics are implemented in the  $\operatorname{unif\_stat}$  function, which allows computing several statistics for different samples within a single call, thus facilitating Monte Carlo experiments. Furthermore, the  $\operatorname{unif\_stat\_MC}$  function allows parallelizing them in a simple way. The asymptotic null distributions of the statistics are available through the function  $\operatorname{unif\_stat\_distr}$ . The core of  $\operatorname{sphunif-package}$  is coded in C++ by relying on the Rcpp-package. The package also provides several novel datasets and gives the replicability for the data applications/ simulations in García-Portugués et al. (2021) <doi:10.1007/978-3-030-69944-4\_12>, García-Portugués et al. (2023) <doi:10.3150/21-BEJ1454>, García-Portugués et al. (2024) <doi:10.48550/arXiv.2108.09874>, and Fernández-de-Marcos and García-Portugués (2024) <doi:10.48550/arXiv.405.13531>.

# Author(s)

Eduardo García-Portugués and Thomas Verdebout.

## References

Fernández-de-Marcos, A. and García-Portugués, E. (2024) A stereographic test of spherical uniformity. *arXiv*:2405.13531. doi:10.48550/arXiv.2405.13531.

García-Portugués, E. and Verdebout, T. (2018) An overview of uniformity tests on the hypersphere. *arXiv:1804.00286*. doi:10.48550/arXiv.1804.00286.

García-Portugués, E., Navarro-Esteban, P., Cuesta-Albertos, J. A. (2023) On a projection-based class of uniformity tests on the hypersphere. *Bernoulli*, 29(1):181–204. doi:10.3150/21BEJ1454.

García-Portugués, E., Navarro-Esteban, P., and Cuesta-Albertos, J. A. (2021). A Cramér–von Mises test of uniformity on the hypersphere. In Balzano, S., Porzio, G. C., Salvatore, R., Vistocco, D., and Vichi, M. (Eds.), *Statistical Learning and Modeling in Data Analysis*, Studies in Classification, Data Analysis and Knowledge Organization, pp. 107–116. Springer, Cham. doi:10.1007/9783030-699444\_12.

García-Portugués, E., Paindaveine, D., and Verdebout, T. (2024). On a class of Sobolev tests for symmetry of directions, their detection thresholds, and asymptotic powers. *arXiv:2108.09874v2*. doi:10.48550/arXiv.2108.09874.

4 angles\_to\_sphere

angles\_to\_sphere

Conversion between angular and Cartesian coordinates of the (hyper)sphere

## **Description**

```
Transforms the angles (\theta_1, \dots, \theta_{p-1})' in [0, \pi)^{p-2} \times [-\pi, \pi) into the Cartesian coordinates (\cos(x_1), \sin(x_1)\cos(x_2), \dots, \sin(x_1)\cdots\sin(x_{p-2})\cos(x_{p-1}), \sin(x_1)\cdots\sin(x_{p-2})\sin(x_{p-1}))' of S^{p-1}, and vice versa.
```

# Usage

```
angles_to_sphere(theta)
sphere_to_angles(x)
```

# **Arguments**

theta matrix of size c(n, p - 1) with the angles.
 x matrix of size c(n, p) with the Cartesian coordinates. Assumed to be of unit norm by rows.

#### Value

For angles\_to\_sphere, the matrix x. For sphere\_to\_angles, the matrix theta.

avail\_tests 5

avail\_tests

Available circular and (hyper)spherical uniformity tests

# Description

Listing of the tests implemented in the sphunif package.

## Usage

```
avail_cir_tests
avail_sph_tests
```

#### **Format**

An object of class character of length 33.

An object of class character of length 18.

#### Value

A character vector whose elements are valid inputs for the type argument in unif\_test, unif\_stat, unif\_stat\_distr, and unif\_stat\_MC. avail\_cir\_tests provides the available circular tests and avail\_sph\_tests the (hyper)spherical tests.

# **Examples**

```
# Circular tests
avail_cir_tests
# Spherical tests
avail_sph_tests
```

A\_theta\_x

Surface area of the intersection of two hyperspherical caps

# **Description**

Computation of

$$A_x(\theta_{ij}) := \frac{1}{\omega_p} \int_{S^{p-1}} 1_{\{\mathbf{X}_i' \boldsymbol{\gamma} \le x, \mathbf{X}_j' \boldsymbol{\gamma} \le x\}} \, \mathrm{d}\boldsymbol{\gamma},$$

where  $\theta_{ij} := \cos^{-1}(\mathbf{X}_i'\mathbf{X}_j) \in [0, \pi], x \in [-1, 1]$ , and  $\omega_p$  is the surface area of  $S^{p-1}$ .  $A_x(\theta_{ij})$  is the proportion of surface area of  $S^{p-1}$  covered by the intersection of two hyperspherical caps centered at  $\mathbf{X}_i$  and  $\mathbf{X}_j$  and with common solid angle  $\pi - \cos^{-1}(x)$ .

6 A\_theta\_x

### Usage

```
A_{theta}(theta, x, p, N = 160L, as_{matrix} = TRUE)
```

## **Arguments**

theta	vector with values in $[0, \pi]$ .
X	vector with values in $[-1,1]$ .
р	integer giving the dimension of the ambient space $\mathbb{R}^p$ that contains $\mathbb{S}^{p-1}$ .
N	number of points used in the Gauss-Legendre quadrature. Defaults to 160.
as_matrix	return a matrix with the values of $A_x(\theta)$ on the grid formed by theta and x? If FALSE, $A_x(\theta)$ is evaluated on theta and x if they equal in size. Defaults to TRUE.

## **Details**

See García-Portugués et al. (2023) for more details about the  $A_x(\theta)$  function.

## Value

A matrix of size c(length(theta), length(x)) containing the evaluation of  $A_x(\theta)$  if as\_matrix = TRUE. Otherwise, a vector of size c(length(theta) if theta and x equal in size.

#### References

García-Portugués, E., Navarro-Esteban, P., Cuesta-Albertos, J. A. (2023) On a projection-based class of uniformity tests on the hypersphere. *Bernoulli*, 29(1):181–204. doi:10.3150/21BEJ1454.

```
# Plot A_x(theta) for several dimensions and x's
A_lines <- function(x, th = seq(0, pi, 1 = 200)) {
  plot(th, A_theta_x(theta = th, x = x, p = 2), type = "1",
       col = 1, ylim = c(0, 1.25), main = paste("x =", x),
       ylab = expression(A[x](theta)),
       xlab = expression(theta), axes = FALSE)
  axis(1, at = c(0, pi / 4, pi / 2, 3 * pi / 4, pi),
       labels = expression(0, pi / 4, pi / 2, 3 * pi / 4, pi))
  axis(2); box()
  abline(h = c(0, 1), lty = 2)
  lines(th, A_theta_x(theta = th, x = x, p = 3), col = 2)
  lines(th, A_theta_x(theta = th, x = x, p = 4), col = 3)
  lines(th, A_theta_x(theta = th, x = x, p = 5), col = 4)
  legend("top", lwd = 2, legend = paste("p =", 2:5),
         col = 1:4, cex = 0.75, horiz = TRUE)
old_par <- par(mfrow = c(2, 3))
A_{lines}(x = -0.75)
A_{lines}(x = -0.25)
```

cir\_coord\_conv 7

```
A_{lines}(x = 0)
A_{lines}(x = 0.25)
A_{lines}(x = 0.5)
A_{lines}(x = 0.75)
par(old_par)
# As surface of (theta, x) for several dimensions
A_surf <- function(p, x = seq(-1, 1, 1 = 201), th = seq(0, pi, 1 = 201)) {
  col <- c("white", viridisLite::viridis(20))</pre>
  breaks <- c(-1, seq(1e-15, 1, 1 = 21))
  A \leftarrow A_{theta}(theta = th, x = x, p = p)
  image(th, x, A, main = paste("p =", p), col = col, breaks = breaks,
        xlab = expression(theta), axes = FALSE)
  axis(1, at = c(0, pi / 4, pi / 2, 3 * pi / 4, pi),
       labels = expression(0, pi / 4, pi / 2, 3 * pi / 4, pi))
  axis(2); box()
  contour(th, x, A, levels = breaks, add = TRUE)
old_par <- par(mfrow = c(2, 2))
A_surf(p = 2)
A_surf(p = 3)
A_surf(p = 4)
A_surf(p = 5)
par(old_par)
# No matrix return
th <- seq(0, pi, 1 = 5)
x < - seq(-1, 1, 1 = 5)
diag(A_theta_x(theta = th, x = x, p = 2))
A_{theta}(theta = th, x = x, p = 2, as_{matrix} = FALSE)
```

cir\_coord\_conv

Transforming between polar and Cartesian coordinates

# **Description**

Transformation between a matrix Theta containing M circular samples of size n on  $[0, 2\pi)$  and an array X containing the associated Cartesian coordinates on  $S^1 := \{ \mathbf{x} \in R^2 : ||\mathbf{x}|| = 1 \}$ .

# Usage

```
Theta_to_X(Theta)
X_to_Theta(X)
```

8 cir\_gaps

# **Arguments**

Theta a **matrix** of size c(n, M) with M samples of size n of circular data on  $[0, 2\pi)$ . Must not contain NA's.

X an **array** of size c(n, 2, M) containing the Cartesian coordinates of M samples

of size n of directions on  $S^1$ . Must not contain NA's.

#### Value

- Theta\_to\_X: the corresponding X.
- X\_to\_Theta: the corresponding Theta.

# **Examples**

```
# Sample
Theta <- r_unif_cir(n = 10, M = 2)
X <- r_unif_sph(n = 10, p = 2, M = 2)

# Check equality
sum(abs(X - Theta_to_X(X_to_Theta(X))))
sum(abs(Theta - X_to_Theta(Theta_to_X(Theta))))</pre>
```

cir\_gaps

Circular gaps

# Description

Computation of the circular gaps of an angular sample  $\Theta_1, \dots, \Theta_n$  on  $[0, 2\pi)$ , defined as

$$\Theta_{(2)} - \Theta_{(1)}, \dots, \Theta_{(n)} - \Theta_{(n-1)}, 2\pi - \Theta_{(n)} - \Theta_{(1)},$$

where

$$0 \le \Theta_{(1)} \le \Theta_{(2)} \le \ldots \le \Theta_{(n)} \le 2\pi.$$

# Usage

```
cir_gaps(Theta, sorted = FALSE)
```

## **Arguments**

Theta a **matrix** of size c(n, M) with M samples of size n of circular data on  $[0, 2\pi)$ .

Must not contain NA's.

sorted are the columns of Theta sorted increasingly? If TRUE, performance is im-

proved. If FALSE (default), each column of Theta is sorted internally.

## Value

A matrix of size c(n, M) containing the n circular gaps for each of the M circular samples.

# Warning

Be careful on avoiding the next bad usages of cir\_gaps, which will produce spurious results:

- The entries of Theta are *not* in  $[0, 2\pi)$ .
- Theta is *not* sorted increasingly when data\_sorted = TRUE.

# **Examples**

```
Theta <- cbind(c(pi, 0, 3 * pi / 2), c(0, 3 * pi / 2, pi), c(5, 3, 1)) cir_{gaps}(Theta)
```

cir\_stat\_Kuiper

Statistics for testing circular uniformity

# Description

Low-level implementation of several statistics for assessing circular uniformity on  $[0, 2\pi)$  or, equivalently,  $S^1 := \{ \mathbf{x} \in \mathbb{R}^2 : ||\mathbf{x}|| = 1 \}.$ 

## Usage

```
cir_stat_Kuiper(Theta, sorted = FALSE, KS = FALSE, Stephens = FALSE)
cir_stat_Watson(Theta, sorted = FALSE, CvM = FALSE, Stephens = FALSE)
cir_stat_Watson_1976(Theta, sorted = FALSE, minus = FALSE)
cir_stat_Range(Theta, sorted = FALSE, gaps_in_Theta = FALSE,
 max_gap = TRUE)
cir_stat_Rao(Theta, sorted = FALSE, gaps_in_Theta = FALSE)
cir_stat_Greenwood(Theta, sorted = FALSE, gaps_in_Theta = FALSE)
cir_stat_Log_gaps(Theta, sorted = FALSE, gaps_in_Theta = FALSE,
  abs_val = TRUE
cir_stat_Vacancy(Theta, a = 2 * pi, sorted = FALSE,
  gaps_in_Theta = FALSE)
cir_stat_Max_uncover(Theta, a = 2 * pi, sorted = FALSE,
  gaps_in_Theta = FALSE)
cir_stat_Num_uncover(Theta, a = 2 * pi, sorted = FALSE,
  gaps_in_Theta = FALSE, minus_val = TRUE)
cir_stat_Gini(Theta, sorted = FALSE, gaps_in_Theta = FALSE)
```

```
cir_stat_Gini_squared(Theta, sorted = FALSE, gaps_in_Theta = FALSE)
cir_stat_Ajne(Theta, Psi_in_Theta = FALSE)
cir_stat_Rothman(Theta, Psi_in_Theta = FALSE, t = 1/3)
cir_stat_Hodges_Ajne(Theta, asymp_std = FALSE, sorted = FALSE,
  use_Cressie = TRUE)
cir_stat_Cressie(Theta, t = 1/3, sorted = FALSE)
cir_stat_FG01(Theta, sorted = FALSE)
cir_stat_Rayleigh(Theta, m = 1L)
cir_stat_Bingham(Theta)
cir_stat_Hermans_Rasson(Theta, Psi_in_Theta = FALSE)
cir_stat_Gine_Gn(Theta, Psi_in_Theta = FALSE)
cir_stat_Gine_Fn(Theta, Psi_in_Theta = FALSE)
cir_stat_Pycke(Theta, Psi_in_Theta = FALSE)
cir_stat_Pycke_q(Theta, Psi_in_Theta = FALSE, q = 0.5)
cir_stat_Bakshaev(Theta, Psi_in_Theta = FALSE)
cir_stat_Riesz(Theta, Psi_in_Theta = FALSE, s = 1)
cir_stat_PCvM(Theta, Psi_in_Theta = FALSE)
cir_stat_PRt(Theta, Psi_in_Theta = FALSE, t = 1/3)
cir_stat_PAD(Theta, Psi_in_Theta = FALSE, AD = FALSE, sorted = FALSE)
cir_stat_Poisson(Theta, Psi_in_Theta = FALSE, rho = 0.5)
cir_stat_Softmax(Theta, Psi_in_Theta = FALSE, kappa = 1)
cir_stat_CCF09(Theta, dirs, K_CCF09 = 25L, original = FALSE)
```

## **Arguments**

Theta a **matrix** of size c(n, M) with M samples of size n of circular data on  $[0, 2\pi)$ . Must not contain NA's.

sorted	are the columns of Theta sorted increasingly? If TRUE, performance is improved. If FALSE (default), each column of Theta is sorted internally.
KS	compute the Kolmogorov-Smirnov statistic (which is <i>not</i> invariant under origin shifts) instead of the Kuiper statistic? Defaults to FALSE.
Stephens	compute Stephens (1970) modification so that the null distribution of the is less dependent on the sample size? The modification does not alter the test decision.
CvM	compute the Cramér-von Mises statistic (which is <i>not</i> invariant under origin shifts) instead of the Watson statistic? Defaults to FALSE.
minus	compute the invariant $D_n^-$ instead of $D_n^+$ ? Defaults to FALSE.
gaps_in_Theta	does Theta contain the matrix of <i>circular gaps</i> that is obtained with cir_gaps(Theta)? If FALSE (default), the circular gaps are computed internally.
max_gap	compute the maximum gap for the range statistic? If TRUE (default), rejection happens for <i>large</i> values of the statistic, which is consistent with the rest of tests. Otherwise, the minimum gap is computed and rejection happens for <i>low</i> values.
abs_val	return the absolute value of the Darling's log gaps statistic? If TRUE (default), rejection happens for <i>large</i> values of the statistic, which is consistent with the rest of tests. Otherwise, the signed statistic is computed and rejection happens for large <i>absolute</i> values.
a	either:
	• $a_n=a/n$ parameter used in the length of the arcs of the coverage-based tests. Must be positive. Defaults to 2 * pi.
	• $a$ parameter for the Stereo test, a real in $[-1,1]$ . Defaults to 0.
minus_val	return the negative value of the (standardized) number of uncovered spacings? If TRUE (default), rejection happens for <i>large</i> values of the statistic, which is consistent with the rest of tests. Otherwise, rejection happens for <i>low</i> values.
Psi_in_Theta	does Theta contain the shortest angles matrix $\Psi$ that is obtained with Psi_mat(array(Theta, dim = c(n, 1, M)))? If FALSE (default), $\Psi$ is computed internally.
t	t parameter for the Rothman and Cressie tests, a real in $(0,1)$ . Defaults to 1 / 3.
asymp_std	normalize the Hodges-Ajne statistic in terms of its asymptotic distribution? Defaults to FALSE.
use_Cressie	compute the Hodges-Ajne statistic as a particular case of the Cressie statistic? Defaults to TRUE as it is more efficient. If FALSE, the geometric construction in Ajne (1968) is employed.
m	integer $m$ for the $m$ -modal Rayleigh test. Defaults to ${\bf m}$ = 1 (the standard Rayleigh test).
q	q parameter for the Pycke " $q-test$ ", a real in $(0,1).$ Defaults to 1 / 2.
S	s parameter for the $s$ -Riesz test, a real in $(0,2)$ . Defaults to 1.
AD	compute the Anderson-Darling statistic (which is <i>not</i> invariant under origin shifts) instead of the Projected Anderson-Darling statistic? Defaults to FALSE.
rho	$\rho$ parameter for the Poisson test, a real in $[0,1).$ Defaults to 0.5.

kappa	$\kappa$ parameter for the Softmax test, a non-negative real. Defaults to 1.
dirs	a matrix of size $c(n\_proj, 2)$ containing $n\_proj$ random directions (in Cartesian coordinates) on $S^1$ to perform the CCF09 test.
K_CCF09	integer giving the truncation of the series present in the asymptotic distribution of the Kolmogorov-Smirnov statistic. Defaults to 25.
original	return the CCF09 statistic as originally defined? If FALSE (default), a faster and equivalent statistic is computed, and rejection happens for <i>large</i> values of the statistic, which is consistent with the rest of tests. Otherwise, rejection happens for <i>low</i> values.

#### **Details**

Descriptions and references for most of the statistics are available in García-Portugués and Verdebout (2018).

The statistics cir\_stat\_PCvM and cir\_stat\_PRt are provided for the sake of completion, but they equal the more efficiently-implemented statistics 2 \* cir\_stat\_Watson and cir\_stat\_Rothman, respectively.

## Value

A matrix of size c(M, 1) containing the statistics for each of the M samples.

## Warning

Be careful on avoiding the next bad usages of the functions, which will produce spurious results:

- The entries of Theta are *not* in  $[0, 2\pi)$ .
- Theta does *not* contain the circular gaps when gaps\_in\_Theta = TRUE.
- Theta is *not* sorted increasingly when data\_sorted = TRUE.
- Theta does not contain Psi\_mat(array(Theta, dim = c(n, 1, M))) when Psi\_in\_Theta = TRUE.
- The directions in dirs do *not* have unit norm.

## References

García-Portugués, E. and Verdebout, T. (2018) An overview of uniformity tests on the hypersphere. *arXiv:1804.00286*. doi:10.48550/arXiv.1804.00286.

```
## Sample uniform circular data

M <- 2
n <- 100
set.seed(987202226)
Theta <- r_unif_cir(n = n, M = M)

## Tests based on the empirical cumulative distribution function</pre>
```

```
# Kuiper
cir_stat_Kuiper(Theta)
cir_stat_Kuiper(Theta, Stephens = TRUE)
# Watson
cir_stat_Watson(Theta)
cir_stat_Watson(Theta, Stephens = TRUE)
# Watson (1976)
cir_stat_Watson_1976(Theta)
## Partition-based tests
# Ajne
Theta_array <- Theta
dim(Theta_array) <- c(nrow(Theta), 1, ncol(Theta))</pre>
Psi <- Psi_mat(Theta_array)</pre>
cir_stat_Ajne(Theta)
cir_stat_Ajne(Psi, Psi_in_Theta = TRUE)
# Rothman
cir_stat_Rothman(Theta, t = 0.5)
cir_stat_Rothman(Theta)
cir_stat_Rothman(Psi, Psi_in_Theta = TRUE)
# Hodges-Ajne
cir_stat_Hodges_Ajne(Theta)
cir_stat_Hodges_Ajne(Theta, use_Cressie = FALSE)
# Cressie
cir_stat_Cressie(Theta, t = 0.5)
cir_stat_Cressie(Theta)
# FG01
cir_stat_FG01(Theta)
## Spacings-based tests
# Range
cir_stat_Range(Theta)
# Rao
cir_stat_Rao(Theta)
# Greenwood
cir_stat_Greenwood(Theta)
# Log gaps
cir\_stat\_Log\_gaps(Theta)
# Vacancy
cir_stat_Vacancy(Theta)
```

```
# Maximum uncovered spacing
cir_stat_Max_uncover(Theta)
# Number of uncovered spacings
cir_stat_Num_uncover(Theta)
# Gini mean difference
cir_stat_Gini(Theta)
# Gini mean squared difference
cir_stat_Gini_squared(Theta)
## Sobolev tests
# Rayleigh
cir_stat_Rayleigh(Theta)
cir_stat_Rayleigh(Theta, m = 2)
# Bingham
cir_stat_Bingham(Theta)
# Hermans-Rasson
cir_stat_Hermans_Rasson(Theta)
cir_stat_Hermans_Rasson(Psi, Psi_in_Theta = TRUE)
# Gine Fn
cir_stat_Gine_Fn(Theta)
cir_stat_Gine_Fn(Psi, Psi_in_Theta = TRUE)
# Gine Gn
cir_stat_Gine_Gn(Theta)
cir_stat_Gine_Gn(Psi, Psi_in_Theta = TRUE)
# Pycke
cir_stat_Pycke(Theta)
cir_stat_Pycke(Psi, Psi_in_Theta = TRUE)
# Pycke q
cir_stat_Pycke_q(Theta)
cir_stat_Pycke_q(Psi, Psi_in_Theta = TRUE)
# Bakshaev
cir_stat_Bakshaev(Theta)
cir_stat_Bakshaev(Psi, Psi_in_Theta = TRUE)
# Riesz
cir_stat_Riesz(Theta, s = 1)
cir_stat_Riesz(Psi, Psi_in_Theta = TRUE, s = 1)
# Projected Cramér-von Mises
cir_stat_PCvM(Theta)
cir_stat_PCvM(Psi, Psi_in_Theta = TRUE)
```

comets 15

```
# Projected Rothman
cir_stat_PRt(Theta, t = 0.5)
cir_stat_PRt(Theta)
cir_stat_PRt(Psi, Psi_in_Theta = TRUE)
# Projected Anderson-Darling
cir_stat_PAD(Theta)
cir_stat_PAD(Psi, Psi_in_Theta = TRUE)
## Other tests
# CCF09
dirs <- r_unif_sph(n = 3, p = 2, M = 1)[, , 1]
cir_stat_CCF09(Theta, dirs = dirs)
## Connection of Kuiper and Watson statistics with KS and CvM, respectively
# Rotate sample for KS and CvM
alpha <- seq(0, 2 * pi, 1 = 1e4)
KS_alpha <- sapply(alpha, function(a) {</pre>
  cir_stat_Kuiper((Theta[, 2, drop = FALSE] + a) %% (2 * pi), KS = TRUE)
})
CvM_alpha <- sapply(alpha, function(a) {</pre>
  cir_stat_Watson((Theta[, 2, drop = FALSE] + a) %% (2 * pi), CvM = TRUE)
AD_alpha <- sapply(alpha, function(a) {
  cir_stat_PAD((Theta[, 2, drop = FALSE] + a) \% (2 * pi), AD = TRUE)
})
# Kuiper is the maximum rotated KS
plot(alpha, KS_alpha, type = "1")
abline(h = cir_stat_Kuiper(Theta[, 2, drop = FALSE]), col = 2)
points(alpha[which.max(KS_alpha)], max(KS_alpha), col = 2, pch = 16)
# Watson is the minimum rotated CvM
plot(alpha, CvM_alpha, type = "1")
abline(h = cir_stat_Watson(Theta[, 2, drop = FALSE]), col = 2)
points(alpha[which.min(CvM_alpha)], min(CvM_alpha), col = 2, pch = 16)
# Anderson-Darling is the average rotated AD?
plot(alpha, AD_alpha, type = "1")
abline(h = cir_stat_PAD(Theta[, 2, drop = FALSE]), col = 2)
abline(h = mean(AD_alpha), col = 3)
```

comets

Comet orbits

#### **Description**

Comet orbits data from the JPL Small-Body Database Search Engine. The normal vector of a comet orbit represents is a vector on  $S^2$ .

16 comets

# Usage

comets

#### **Format**

A data frame with 3798 rows and 13 variables:

id database ID.

spkid object primary SPK-ID.

full\_name full name/designation following the IUA naming convention.

pdes object primary designation.

**frag** flag indicating if the record is a comet fragment.

diameter diameter from equivalent sphere (in km).

i inclination; the orbit's plane angle with respect to the ecliptic plane, in radians in  $[0, \pi]$ .

om longitude of the ascending node; the counterclockwise angle from the vector pointing to the First Point of Aries and that pointing to the ascending node (the intersection between orbit and ecliptic plane), in radians in  $[0, 2\pi)$ . (Both vectors are heliocentric and within the ecliptic plane.)

**per\_y** sidereal orbital period (in years).

class orbit classification. A factor with levels given below.

- e eccentricity of the orbit.
- a semi-major axis of the orbit (in AU).
- **w** argument of perihelion; the (shortest) angle between the vector pointing to the ascending node and that pointing to the perihelion (nearest orbit point to the Sun), in radians in  $[0, \pi]$ . (Both vectors are heliocentric and within the orbit's plane.)

first\_obs, last\_obs Date of the first and last recorded observations used in the orbit fit.

**ccf09** flag indicating if the comet was considered in the data application in Cuesta-Albertos et al. (2009); see details below.

## **Details**

The normal vector to the ecliptic plane of the comet with inclination i and longitude of the ascending node  $\omega$  is

$$(\sin(i)\sin(\omega), -\sin(i)\cos(\omega), \cos(i))'$$
.

A prograde comet has positive  $\cos(i)$ , negative  $\cos(i)$  represents a retrograde comet.

class has the following levels:

- COM: comet orbit not matching any defined orbit class.
- CTc: Chiron-type comet, as defined by Levison and Duncan (T\_Jupiter > 3; a > a\_Jupiter).
- ETc: Encke-type comet, as defined by Levison and Duncan (T\_Jupiter > 3; a < a\_Jupiter).
- HTC: Halley-type comet, classical definition (20y < P < 200y).
- HYP: comets on hyperbolic orbits.

comets 17

- JFc: Jupiter-family comet, as defined by Levison and Duncan (2 < T\_Jupiter < 3).
- JFC: Jupiter-family comet, classical definition (P < 20y).
- PAR: comets on parabolic orbits.

Hyperbolic and parabolic comets are not periodic; only elliptical comets are periodic.

The ccf09 variable gives the observations considered in Cuesta-Albertos et al. (2009) after fetching in the database in 2007-12-14 for the comets such that !(class %in% c("HYP", "PAR")) & per\_y >= 200. Due to the dynamic nature of the data, more comets were added to the database since 2007 and also some past records were updated.

The script performing the data preprocessing is available at comets.R. The data was retrieved on 2022-05-28. A previous version of this dataset based on the old NASA's JPL Database (accessed on 2020-05-07) is available at comets-old.rda and was obtained with comets-old.R.

#### Source

```
https://ssd.jpl.nasa.gov/tools/sbdb_query.html
```

#### References

Cuesta-Albertos, J. A., Cuevas, A., Fraiman, R. (2009) On projection-based tests for directional and compositional data. *Statistics and Computing*, 19:367–380. doi:10.1007/s1122200890983

```
# Load data
data("comets")
# Add normal vectors
comets$normal <- cbind(sin(comets$i) * sin(comets$om),</pre>
                        -sin(comets$i) * cos(comets$om),
                        cos(comets$i))
# Tests to be performed
type_tests <- c("PCvM", "PAD", "PRt")</pre>
# Excluding the C/1882 R1-X (Great September comet) records with X = B, C, D
comets_ccf09 <- comets[comets$ccf09, ][-c(13:15), ]</pre>
# Sample size
nrow(comets_ccf09)
# Tests for the data in Cuesta-Albertos et al. (2009)
tests_ccf09 <- unif_test(data = comets_ccf09$normal, type = type_tests,</pre>
                          p_value = "asymp")
tests_ccf09
```

18 craters

craters

Craters named by the IUA

## **Description**

*Named* craters of the Solar System by the Gazetteer of Planetary Nomenclature of the International Astronomical Union (IUA).

# Usage

craters

#### **Format**

A data frame with 5235 rows and 7 variables:

ID database ID.

name name of the crater.

target name of the celestial body. A factor with 43 levels, such as "Moon", "Venus", or "Europa".

target\_type type of celestial body. A factor with 3 levels: "Planet", "Moon", "Dwarf planet", or
 "Asteroid".

diameter diameter of the crater (in km).

**theta** longitude angle  $\theta \in [0, 2\pi)$  of the crater center.

**phi** latitude angle  $\phi \in [-\pi/2, \pi/2]$  of the crater center.

#### **Details**

"Craters" are understood in the Gazetteer of Planetary Nomenclature as roughly circular depressions resulting from impact or volcanic activity (the geological origin is unspecified).

Be aware that the dataset only contains *named* craters by the IUA. Therefore, there is likely a **high uniform bias** on the distribution of craters. Presumably the naming process attempts to cover the planet in a somehow uniform fashion (distant craters are more likely to be named than neighboring craters). Also, there are substantially more craters in the listed bodies than those named by the IUA. See venus and rhea for more detailed and specific crater datasets.

The  $(\theta, \phi)$  angles are such their associated planetocentric coordinates are:

$$(\cos(\phi)\cos(\theta),\cos(\phi)\sin(\theta),\sin(\phi))',$$

with (0,0,1)' denoting the north pole.

The script performing the data preprocessing is available at craters.R. The data was retrieved on 2020-05-31.

## Source

https://planetarynames.wr.usgs.gov/AdvancedSearch

F\_from\_f 19

## **Examples**

F\_from\_f

Distribution and quantile functions from angular function

# **Description**

Numerical computation of the distribution function F and the quantile function  $F^{-1}$  for an angular function f in a tangent-normal decomposition.  $F^{-1}(x)$  results from the inversion of

$$F(x) = \int_{-1}^{x} \omega_{p-1} c_f f(z) (1 - z^2)^{(p-3)/2} dz$$

for  $x \in [-1, 1]$ , where  $c_f$  is a normalizing constant and  $\omega_{p-1}$  is the surface area of  $S^{p-2}$ .

# Usage

```
F_{from_f(f, p, Gauss = TRUE, N = 320, K = 1000, tol = 1e-06, ...)}
F_{inv_from_f(f, p, Gauss = TRUE, N = 320, K = 1000, tol = 1e-06, ...)}
```

#### **Arguments**

f	angular function defined on $[-1,1]$ . Must be vectorized.
p	integer giving the dimension of the ambient space $\mathbb{R}^p$ that contains $\mathbb{S}^{p-1}$ .
Gauss	use a Gauss-Legendre quadrature rule to integrate $f$ with N nodes? Otherwise, rely on integrate Defaults to TRUE.
N	number of points used in the Gauss-Legendre quadrature. Defaults to 320.
K	number of equispaced points on $[-1,1]$ used for evaluating ${\cal F}^{-1}$ and then interpolating. Defaults to 1e3.
tol	tolerance passed to uniroot for the inversion of $F$ . Also, passed to integrate's rel.tol and abs.tol if Gauss = FALSE. Defaults to 1e-6.
	further parameters passed to f.

20 Gauss\_Legen

### **Details**

The normalizing constant  $c_f$  is such that F(1) = 1. It does not need to be part of f as it is computed internally.

Interpolation is performed by a monotone cubic spline. Gauss = TRUE yields more accurate results, at expenses of a heavier computation.

If f yields negative values, these are silently truncated to zero.

#### Value

A splinefun object ready to evaluate F or  $F^{-1}$ , as specified.

### **Examples**

Gauss\_Legen

Gauss-Legendre quadrature

# Description

Convenience for computing the nodes  $x_k$  and weights  $w_k$  of the Gauss-Legendre quadrature formula in (a, b):

 $\int_{a}^{b} f(x)w(x) dx \approx \sum_{k=1}^{N} w_{k} f(x_{k}).$ 

Usage

```
Gauss_Legen_nodes(a = -1, b = 1, N = 40L)

Gauss_Legen_weights(a = -1, b = 1, N = 40L)
```

# **Arguments**

Ν

a, b scalars giving the interval (a, b). Defaults to (-1, 1).

number of points used in the Gauss–Legendre quadrature. The following choices are supported: 5, 10, 20, 40, 80, 160, 320, 640, 1280, 2560, and 5120. Defaults to 40.

Gauss\_Legen 21

#### **Details**

For  $C^{\infty}$  functions, Gauss–Legendre quadrature can be very efficient. It is exact for polynomials up to degree 2N-1.

The nodes and weights up to N=80 were retrieved from NIST and have  $10^{-21}$  precision. For N=160 onwards, the nodes and weights were computed with the gauss quad function from the statmod package (Smyth, 1998), and have  $10^{-15}$  precision.

#### Value

A matrix of size c(N, 1) with the nodes  $x_k$  (Gauss\_Legen\_nodes) or the corresponding weights  $w_k$  (Gauss\_Legen\_weights).

#### References

NIST Digital Library of Mathematical Functions. Release 1.0.20 of 2018-09-15. F. W. J. Olver, A. B. Olde Daalhuis, D. W. Lozier, B. I. Schneider, R. F. Boisvert, C. W. Clark, B. R. Miller, and B. V. Saunders, eds. https://dlmf.nist.gov/

Smyth, G. K. (1998). Numerical integration. In: *Encyclopedia of Biostatistics*, P. Armitage and T. Colton (eds.), Wiley, London, pp. 3088-3095.

```
## Integration of a smooth function in (1, 10)
# Weights and nodes for integrating
x_k \leftarrow Gauss\_Legen\_nodes(a = 1, b = 10, N = 40)
w_k \leftarrow Gauss\_Legen\_weights(a = 1, b = 10, N = 40)
# Check quadrature
f \leftarrow function(x) sin(x) * x^2 - log(x + 1)
integrate(f, lower = 1, upper = 10, rel.tol = 1e-12)
sum(w_k * f(x_k))
# Exact for polynomials up to degree 2 * N - 1
f \leftarrow function(x) (((x + 0.5) / 1e3)^5 - ((x - 0.5)/5)^4 +
  ((x - 0.25) / 10)^2 + 1)^20
sum(w_k * f(x_k))
integrate(f, lower = -1, upper = 1, rel.tol = 1e-12)
## Integration on (0, pi)
# Weights and nodes for integrating
th_k \leftarrow Gauss_{Legen_nodes}(a = 0, b = pi, N = 40)
w_k \leftarrow Gauss\_Legen\_weights(a = 0, b = pi, N = 40)
# Check quadrature
p <- 4
psi <- function(th) -sin(th / 2)</pre>
w <- function(th) sin(th)^(p - 2)</pre>
integrate(function(th) psi(th) * w(th), lower = 0, upper = pi,
```

Gegenbauer

Gegenbauer polynomials and coefficients

## **Description**

The Gegenbauer polynomials  $\{C_k^{(\lambda)}(x)\}_{k=0}^\infty$  form a family of orthogonal polynomials on the interval [-1,1] with respect to the weight function  $(1-x^2)^{\lambda-1/2}$ , for  $\lambda>-1/2,\,\lambda\neq 0$ . They usually appear when dealing with functions defined on  $S^{p-1}:=\{\mathbf{x}\in R^p:||\mathbf{x}||=1\}$  with index  $\lambda=p/2-1$ .

The Gegenbauer polynomials are somehow simpler to evaluate for  $x=\cos(\theta)$ , with  $\theta\in[0,\pi]$ . This simplifies also the connection with the Chebyshev polynomials  $\{T_k(x)\}_{k=0}^\infty$ , which admit the explicit expression  $T_k(\cos(\theta))=\cos(k\theta)$ . The Chebyshev polynomials appear as the limit of the Gegenbauer polynomials (divided by  $\lambda$ ) when  $\lambda$  goes to 0, so they can be regarded as the extension by continuity of  $\{C_k^{(p/2-1)}(x)\}_{k=0}^\infty$  to the case p=2.

For a reasonably smooth function  $\psi$  defined on  $[0, \pi]$ ,

$$\psi(\theta) = \sum_{k=0}^{\infty} b_{k,p} C_k^{(p/2-1)}(\cos(\theta)),$$

provided that the coefficients

$$b_{k,p} := \frac{1}{c_{k,p}} \int_0^{\pi} \psi(\theta) C_k^{(p/2-1)}(\cos(\theta)) (\sin(\theta))^{p-2} d\theta$$

are finite, where the normalizing constants are

$$c_{k,p} := \int_0^{\pi} (C_k^{(p/2-1)}(\cos(\theta)))^2 (\sin(\theta))^{p-2} d\theta.$$

The (squared) "Gegenbauer norm" of  $\psi$  is

$$\|\psi\|_{G,p}^2 := \int_0^{\pi} \psi(\theta)^2 C_k^{(p/2-1)}(\cos(\theta))(\sin(\theta))^{p-2} d\theta.$$

The previous expansion can be generalized for a 2-dimensional function  $\psi$  defined on  $[0, \pi] \times [0, \pi]$ :

$$\psi(\theta_1, \theta_2) = \sum_{k=0}^{\infty} \sum_{m=0}^{\infty} b_{k,m,p} C_k^{(p/2-1)}(\cos(\theta_1)) C_k^{(p/2-1)}(\cos(\theta_2)),$$

with coefficients

$$b_{k,m,p} := \frac{1}{c_{k,p}c_{m,p}} \int_0^{\pi} \int_0^{\pi} \psi(\theta_1,\theta_2) C_k^{(p/2-1)}(\cos(\theta_1)) C_k^{(p/2-1)}(\cos(\theta_2)) (\sin(\theta_1))^{p-2} (\sin(\theta_2))^{p-2} d\theta_1 d\theta_2.$$

The (squared) "Gegenbauer norm" of  $\psi$  is

$$\|\psi\|_{G,p}^2 := \int_0^{\pi} \int_0^{\pi} \psi(\theta_1, \theta_2)^2 C_k^{(p/2-1)}(\cos(\theta_1)) C_k^{(p/2-1)}(\cos(\theta_2)) (\sin(\theta_1))^{p-2} (\sin(\theta_2))^{p-2} d\theta_1 d\theta_2.$$

# Usage

```
Gegen_polyn(theta, k, p)

Gegen_coefs(k, p, psi, Gauss = TRUE, N = 320, normalize = TRUE,
    only_const = FALSE, tol = 1e-06, ...)

Gegen_series(theta, coefs, k, p, normalize = TRUE)

Gegen_norm(coefs, k, p, normalize = TRUE, cumulative = FALSE)

Gegen_polyn_2d(theta_1, theta_2, k, m, p)

Gegen_coefs_2d(k, m, p, psi, Gauss = TRUE, N = 320, normalize = TRUE,
    only_const = FALSE, tol = 1e-06, ...)

Gegen_series_2d(theta_1, theta_2, coefs, k, m, p, normalize = TRUE)

Gegen_norm_2d(coefs, k, m, p, normalize = TRUE)
```

# **Arguments**

theta, theta\_1, theta\_2

vectors with values in  $[0, \pi]$ .

k, m vectors with the orders of the Gegenbauer polynomials. Must be integers larger

or equal than 0.

p integer giving the dimension of the ambient space  $R^p$  that contains  $S^{p-1}$ .

psi function defined in  $[0, \pi]$  and whose Gegenbauer coefficients are to be com-

puted. Must be vectorized. For Gegen\_coefs\_2d, it must return a matrix of size

c(length(theta\_1), length(theta\_2)).

Gauss use a Gauss-Legendre quadrature rule of N nodes in the computation of the

Gegenbauer coefficients? Otherwise, call integrate. Defaults to TRUE.

N number of points used in the Gauss-Legendre quadrature for computing the

Gegenbauer coefficients. Defaults to 320.

normalize consider normalized coefficients (divided by  $c_{k,p}$ )? Defaults to TRUE.

only\_const return only the normalizing constants  $c_{k,p}$ ? Defaults to FALSE.

tol tolerance passed to integrate's rel.tol and abs.tol if Gauss = FALSE. De-

faults to 1e-6.

... further arguments to be passed to psi.

coefs for Gegen\_series and Gegen\_norm, a vector of coefficients  $b_{k,p}$  with length

length(k). For Gegen\_series\_2d and Gegen\_norm\_2d, a matrix of coefficients  $b_{k,m,p}$  with size c(length(k), length(m)). The order of the coeffi-

cients is given by k and m.

cumulative return the cumulative norm for increasing truncation of the series? Defaults to

FALSE.

#### **Details**

The Gegen\_polyn function is a wrapper to the functions gegenpoly\_n and gegenpoly\_array in the gsl-package, which they interface the functions defined in the header file gsl\_sf\_gegenbauer.h (documented here) of the GNU Scientific Library.

Note that the function Gegen\_polyn computes the regular unnormalized Gegenbauer polynomials.

For the case p = 2, the Chebyshev polynomials are considered.

#### Value

- Gegen\_polyn: a matrix of size c(length(theta), length(k)) containing the evaluation of the length(k) Gegenbauer polynomials at theta.
- Gegen\_coefs: a vector of size length(k) containing the coefficients  $b_{k,p}$ .
- Gegen\_series: the evaluation of the truncated series expansion, a vector of size length(theta).
- Gegen\_norm: the Gegenbauer norm of the truncated series, a scalar if cumulative = FALSE, otherwise a vector of size length(k).
- Gegen\_polyn\_2d: a 4-dimensional array of size c(length(theta\_1), length(theta\_2), length(k), length(m)) containing the evaluation of the length(k) \* length(m) 2-dimensional Gegenbauer polynomials at the bivariate grid spanned by theta\_1 and theta\_2.
- Gegen\_coefs\_2d: a matrix of size c(length(k), length(m)) containing the coefficients  $b_{k,m,p}$ .
- Gegen\_series\_2d: the evaluation of the truncated series expansion, a matrix of size c(length(theta\_1), length(theta\_2)).
- Gegen\_norm\_2d: the 2-dimensional Gegenbauer norm of the truncated series, a scalar.

# References

Galassi, M., Davies, J., Theiler, J., Gough, B., Jungman, G., Alken, P., Booth, M., and Rossi, F. (2009) *GNU Scientific Library Reference Manual*. Network Theory Ltd. http://www.gnu.org/software/gsl/

NIST Digital Library of Mathematical Functions. Release 1.0.20 of 2018-09-15. F. W. J. Olver, A. B. Olde Daalhuis, D. W. Lozier, B. I. Schneider, R. F. Boisvert, C. W. Clark, B. R. Miller, and B. V. Saunders, eds. https://dlmf.nist.gov/

```
## Representation of Gegenbauer polynomials (Chebyshev polynomials for p = 2)
th <- seq(0, pi, 1 = 500)
k <- 0:3
old_par <- par(mfrow = c(2, 2))
for (p in 2:5) {
  matplot(th, t(Gegen\_polyn(theta = th, k = k, p = p)), lty = 1,
          type = "1", main = substitute(p == d, list(d = p)),
          axes = FALSE, xlab = expression(theta), ylab = "")
  axis(1, at = c(0, pi / 4, pi / 2, 3 * pi / 4, pi),
       labels = expression(0, pi / 4, pi / 2, 3 * pi / 4, pi))
  axis(2); box()
  mtext(text = expression(\{C[k]^{p/2 - 1}\}(cos(theta))), side = 2,
        line = 2, cex = 0.75)
  legend("bottomleft", legend = paste("k =", k), lwd = 2, col = seq_along(k))
}
par(old_par)
## Coefficients and series in p = 2
# Function in [0, pi] to be projected in Chebyshev polynomials
psi <- function(th) -sin(th / 2)</pre>
# Coefficients
p <- 2
k <- 0:4
(coefs <- Gegen_coefs(k = k, p = p, psi = psi))
# Series
plot(th, psi(th), type = "l", axes = FALSE, xlab = expression(theta),
      ylab = "", ylim = c(-1.25, 0))
axis(1, at = c(0, pi / 4, pi / 2, 3 * pi / 4, pi),
     labels = expression(0, pi / 4, pi / 2, 3 * pi / 4, pi))
axis(2); box()
col <- viridisLite::viridis(length(coefs))</pre>
for (i in seq_along(coefs)) {
  lines(th, Gegen_series(theta = th, coefs = coefs[1:(i + 1)], k = 0:i,
                         p = p), col = col[i])
lines(th, psi(th), lwd = 2)
## Coefficients and series in p = 3
# Function in [0, pi] to be projected in Gegenbauer polynomials
psi <- function(th) tan(th / 3)</pre>
# Coefficients
p <- 3
k <- 0:10
(coefs <- Gegen_coefs(k = k, p = p, psi = psi))
```

```
# Series
plot(th, psi(th), type = "1", axes = FALSE, xlab = expression(theta),
      ylab = "", ylim = c(0, 2))
axis(1, at = c(0, pi / 4, pi / 2, 3 * pi / 4, pi),
     labels = expression(0, pi / 4, pi / 2, 3 * pi / 4, pi))
axis(2); box()
col <- viridisLite::viridis(length(coefs))</pre>
for (i in seq_along(coefs)) {
  lines(th, Gegen_series(theta = th, coefs = coefs[1:(i + 1)], k = 0:i,
                         p = p), col = col[i])
lines(th, psi(th), lwd = 2)
## Surface representation
# Surface in [0, pi]^2 to be projected in Gegenbauer polynomials
p <- 3
psi <- function(th_1, th_2) A_theta_x(theta = th_1, x = cos(th_2),
                                       p = p, as_matrix = TRUE)
# Coefficients
k <- 0:20
m <- 0:10
coefs \leftarrow Gegen\_coefs\_2d(k = k, m = m, p = p, psi = psi)
# Series
th <- seq(0, pi, 1 = 100)
col <- viridisLite::viridis(20)</pre>
old_par <- par(mfrow = c(2, 2))
image(th, th, A_theta_x(theta = th, x = cos(th), p = p), axes = FALSE,
      col = col, zlim = c(0, 1), xlab = expression(theta[1]),
      ylab = expression(theta[2]), main = "Original")
axis(1, at = c(0, pi / 4, pi / 2, 3 * pi / 4, pi),
     labels = expression(0, pi / 4, pi / 2, 3 * pi / 4, pi))
axis(2, at = c(0, pi / 4, pi / 2, 3 * pi / 4, pi),
     labels = expression(0, pi / 4, pi / 2, 3 * pi / 4, pi))
box()
for(K in c(5, 10, 20)) {
  A <- Gegen_series_2d(theta_1 = th, theta_2 = th,
                       coefs = coefs[1:(K + 1), ], k = 0:K, m = m, p = p)
  image(th, th, A, axes = FALSE, col = col, zlim = c(0, 1),
        xlab = expression(theta[1]), ylab = expression(theta[2]),
        main = paste(K, "x", m[length(m)], "coefficients"))
  axis(1, at = c(0, pi / 4, pi / 2, 3 * pi / 4, pi),
       labels = expression(0, pi / 4, pi / 2, 3 * pi / 4, pi))
  axis(2, at = c(0, pi / 4, pi / 2, 3 * pi / 4, pi),
       labels = expression(0, pi / 4, pi / 2, 3 * pi / 4, pi))
  box()
}
par(old_par)
```

harmonics 27

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 $(Hyper)spherical\ harmonics$ 

## **Description**

Computation of a certain explicit representation of (hyper)spherical harmonics on  $S^{p-1} := \{ \mathbf{x} \in \mathbb{R}^p : ||\mathbf{x}|| = 1 \}, p \geq 2$ . Details are available in García-Portugués et al. (2024).

### **Usage**

```
g_i_k(x, i = 1, k = 1, m = NULL, show_m = FALSE)
```

# **Arguments**

X	locations in $S^{p-1}$ to evaluate $g_{i,k}$ . Either a matrix of size $c(nx, p)$ or a vector of size $p$ . Normalized internally if required (with a warning message).
i, k	alternative indexing to refer to the i-th (hyper)spherical harmonic of order k. i is a positive integer smaller than d_p_k and k is a non-negative integer.
m	(hyper)spherical harmonic index, as used in Proposition 3.1. The index is computed internally from i and k. Defaults to NULL.
show_m	flag to print $m$ if computed internally when $m = NULL$ .

### **Details**

The implementation uses Proposition 3.1 in García-Portugués et al. (2024), which adapts Theorem 1.5.1 in Dai and Xu (2013) with the correction of typos in the normalizing constant  $h_{\alpha}$  and in the definition of the function  $g_{\alpha}$  of the latter theorem.

# Value

A vector of size nrow(x).

## References

Dai, F. and Xu, Y. (2013). *Approximation Theory and Harmonic Analysis on Spheres and Balls*. Springer, New York. doi:10.1007/9781461466604

García-Portugués, E., Paindaveine, D., and Verdebout, T. (2024). On a class of Sobolev tests for symmetry of directions, their detection thresholds, and asymptotic powers. *arXiv:2108.09874v2*. doi:10.48550/arXiv.2108.09874

```
n <- 3e3
old_par <- par(mfrow = c(2, 3))
k <- 2
for (i in 1:d_p_k(p = 3, k = k)) {
    X <- r_unif_sph(n = n, p = 3, M = 1)[, , 1]</pre>
```

int\_sph\_MC

int\_sph\_MC

Monte Carlo integration of functions on the (hyper)sphere

# **Description**

Monte Carlo approximation of the integral

$$\int_{S^{p-1}} f(x) \, \mathrm{d}x$$

of a function  $f: S^{p-1} \to R$  defined on the (hyper)sphere  $S^{p-1} := \{ \mathbf{x} \in R^p : ||\mathbf{x}|| = 1 \}, p \ge 2.$ 

# Usage

```
int\_sph\_MC(f, p, M = 10000, cores = 1, chunks = ceiling(M/1000), seeds = NULL, ...)
```

# Arguments

f		function to be integrated. Its first argument must be the (hyper)sphere position. Must be vectorized and return a vector of size nrow(x) for a matrix input x. See examples.
р		integer giving the dimension of the ambient space $\mathbb{R}^p$ that contains $\mathbb{S}^{p-1}$ .
М		number of Monte Carlo samples. Defaults to 1e4.
СО	res	number of cores to perform the integration. Defaults to 1.
ch	unks	number of chunks to split the M Monte Carlo samples. Useful for parallelizing the integration in chunks tasks containing ceiling(M / chunks) replications. Useful also for avoiding memory bottlenecks when M is large. Defaults to ceiling(M / 1e3).
se	eds	if provided, a vector of size chunks for fixing the seeds on each of the simulation chunks (useful for reproducing parallel simulations). Specifically, for k in 1:chunks, seeds are set as set.seed(seeds[k], kind = "Mersenne-Twister") in each chunk. Defaults to NULL (no seed setting is done).
		optional arguments to be passed to f or to foreach (for example, .export to

export global variables or other functions to the foreach environment).

#### **Details**

It is possible to have a progress bar if int\_sph\_MC is wrapped with progressr::with\_progress or if progressr::handlers(global = TRUE) is invoked (once) by the user. See the examples below. The progress bar is updated with the number of finished chunks.

## Value

A scalar with the approximate integral.

```
## Sequential simulation
# Vectorized functions to be integrated
x1 \leftarrow function(x) x[, 1]
quad <- function(x, a = 0) a + rowSums(x^4)
# Approximate \inf_{S^{p-1}} x_1 dx = 0
int\_sph\_MC(f = x1, p = 3, M = 1e4, chunks = 2)
# Approximate \inf_{S^{p-1}} (a + \sup_{i \in A} x_i^4) dx
int\_sph\_MC(f = quad, p = 2, M = 1e4, a = 0, chunks = 2)
# Compare with Gauss--Legendre integration on S^2
th_k \leftarrow Gauss\_Legen\_nodes(a = 0, b = 2 * pi, N = 40)
w_k \leftarrow Gauss\_Legen\_weights(a = 0, b = 2 * pi, N = 40)
sum(w_k * quad(cbind(cos(th_k), sin(th_k)), a = 1))
## Parallel simulation with a progress bar
# Define a progress bar
require(progress)
require(progressr)
handlers(handler_progress(
 format = paste("(:spin) [:bar] :percent Iter: :current/:total Rate:",
                  ":tick_rate iter/sec ETA: :eta Elapsed: :elapsedfull"),
 clear = FALSE))
# Call int_sph_MC() within with_progress()
with_progress(int_sph_MC(f = x1, p = 3, cores = 2, M = 1e5, chunks = 100))
# Instead of using with_progress() each time, it is more practical to run
# handlers(global = TRUE)
# once to activate progress bars in your R session
```

# **Description**

Density and random generation for local projected alternatives to uniformity with densities

$$f_{\kappa, \mu}(\mathbf{x}) := \frac{1 - \kappa}{\omega_p} + \kappa f(\mathbf{x}' \mu)$$

where

$$f(z) = \frac{1}{\omega_p} \left\{ 1 + \sum_{k=1}^{\infty} u_{k,p} C_k^{p/2-1}(z) \right\}$$

is the angular function controlling the local alternative in a Gegenbauer series,  $0 \le \kappa \le 1$ ,  $\mu$  is a direction on  $S^{p-1}$ , and  $\omega_p$  is the surface area of  $S^{p-1}$ . The sequence  $\{u_{k,p}\}$  is typically such that  $u_{k,p} = \left(1 + \frac{2k}{p-2}\right)b_{k,p}$  for the Gegenbauer coefficients  $\{b_{k,p}\}$  of the kernel function of a Sobolev statistic (see the transformation between the coefficients  $u_{k,p}$  and  $u_{k,p}$ ).

Also, automatic truncation of the series  $\sum_{k=1}^{\infty} u_{k,p} C_k^{p/2-1}(z)$  according to the proportion of "Gegenbauer norm" explained.

## Usage

```
f_locdev(z, p, uk)
con_f(f, p, N = 320)
d_locdev(x, mu, f, kappa)
r_locdev(n, mu, f, kappa, F_inv = NULL, ...)
cutoff_locdev(p, K_max = 10000, thre = 0.001, type, Rothman_t = 1/3,
    Pycke_q = 0.5, verbose = FALSE, Gauss = TRUE, N = 320, tol = 1e-06)
```

### Arguments

Z	projected evaluation points for $f$ , a vector with entries on $[-1, 1]$ .
р	integer giving the dimension of the ambient space $\mathbb{R}^p$ that contains $\mathbb{S}^{p-1}$ .
uk	coefficients $u_{k,p}$ associated to the indexes 1:length(uk), a vector.
f	angular function defined on $[-1, 1]$ . Must be vectorized.
N	number of points used in the Gauss-Legendre quadrature for computing the Gegenbauer coefficients. Defaults to 320.
X	locations in $S^{p-1}$ to evaluate the density. Either a matrix of size $c(nx, p)$ or a vector of length p. Normalized internally if required (with a warning message).
mu	a unit norm vector of size p giving the axis of rotational symmetry.
kappa	the strength of the local alternative, between $\emptyset$ and 1.
n	sample size, a positive integer.
F_inv	quantile function associated to $f$ . Computed by $F_{inv_from_f}$ if NULL (default).
	further parameters passed to F_inv_from_f.

K_max	integer giving the truncation of the series. Defaults to 1e4.
thre	proportion of norm $not$ explained by the first terms of the truncated series. Defaults to $1e-3$ .
type	name of the Sobolev statistic, using the naming from ${\tt avail\_cir\_tests}$ and ${\tt avail\_sph\_tests}.$
Rothman_t	t parameter for the Rothman test, a real in $(0,1)$ . Defaults to 1 / 3.
Pycke_q	q parameter for the Pycke " $q-test$ ", a real in $(0,1).$ Defaults to 1 / 2.
verbose	output information about the truncation (TRUE or 1) and a diagnostic plot (2)? Defaults to FALSE.
Gauss	use a Gauss-Legendre quadrature rule of N nodes in the computation of the Gegenbauer coefficients? Otherwise, call integrate. Defaults to TRUE.
tol	tolerance passed to $integrate$ 's rel.tol and abs.tol if Gauss = FALSE. Defaults to 1e-6.

#### **Details**

See the definitions of local alternatives in Prentice (1978) and in García-Portugués et al. (2023).

The truncation of  $\sum_{k=1}^{\infty} u_{k,p} C_k^{p/2-1}(z)$  is done to the first K\_max terms and then up to the index such that the first terms leave unexplained the proportion thre of the norm of the whole series. Setting thre = 0 truncates to K\_max terms exactly. If the series only contains odd or even non-zero terms, then only K\_max / 2 addends are *effectively* taken into account in the first truncation.

#### Value

- f\_locdev: angular function evaluated at x, a vector.
- con\_f: normalizing constant  $c_f$  of f, a scalar.
- d\_locdev: density function evaluated at x, a vector.
- r\_locdev: a matrix of size c(n, p) containing a random sample from the density  $f_{\kappa, \mu}$ .
- cutoff\_locdev: vector of coefficients  $\{u_{k,p}\}$  automatically truncated according to K\_max and thre (see details).

#### References

García-Portugués, E., Navarro-Esteban, P., Cuesta-Albertos, J. A. (2023) On a projection-based class of uniformity tests on the hypersphere. *Bernoulli*, 29(1):181–204. doi:10.3150/21BEJ1454.

Prentice, M. J. (1978). On invariant tests of uniformity for directions and orientations. *The Annals of Statistics*, 6(1):169–176. doi:10.1214/aos/1176344075

```
## Local alternatives diagnostics
loc_alt_diagnostic <- function(p, type, thre = 1e-3, K_max = 1e3) {
    # Coefficients of the alternative
    uk <- cutoff_locdev(K_max = K_max, p = p, type = type, thre = thre,</pre>
```

```
N = 640)
 old_par <- par(mfrow = c(2, 2))
 # Construction of f
 z \leftarrow seq(-1, 1, 1 = 1e3)
 f \leftarrow function(z) f_locdev(z = z, p = p, uk = uk)
 plot(z, f(z), type = "l", xlab = expression(z), ylab = expression(f(z)),
       main = paste0("Local alternative f, ", type, ", p = ", p), log = "y")
 # Projected density on [-1, 1]
 f_proj \leftarrow function(z) rotasym::w_p(p = p - 1) * f(z) *
    (1 - z^2)^((p - 3) / 2)
 plot(z, f_proj(z), type = "l", xlab = expression(z),
       ylab = expression(omega[p - 1] * f(z) * (1 - z^2)^{(p - 3) / 2}),
       main = paste0("Projected density, ", type, ", p = ", p), log = "y",
       sub = paste("Integral:", round(con_f(f = f, p = p), 4)))
 # Quantile function for projected density
 mu <- c(rep(0, p - 1), 1)
 F_{inv} \leftarrow F_{inv_from_f(f = f, p = p, K = 5e2)}
 plot(F_inv, xlab = expression(x), ylab = expression(F^{-1}*(x)),
       main = paste0("Quantile function, ", type, ", p = ", p))
 # Sample from the alternative and plot the projected sample
 n <- 5e4
 samp \leftarrow r_locdev(n = n, mu = mu, f = f, kappa = 1, F_inv = F_inv)
 plot(z, f_proj(z), col = 2, type = "l",
       main = paste0("Simulated projected data, ", type, ", p = ", p),
       ylim = c(0, 1.75)
 hist(samp %*% mu, freq = FALSE, breaks = seq(-1, 1, 1 = 50), add = TRUE)
 par(old_par)
## Local alternatives for the PCvM test
loc_alt_diagnostic(p = 2, type = "PCvM")
loc_alt_diagnostic(p = 3, type = "PCvM")
loc_alt_diagnostic(p = 4, type = "PCvM")
loc_alt_diagnostic(p = 5, type = "PCvM")
loc_alt_diagnostic(p = 11, type = "PCvM")
## Local alternatives for the PAD test
loc_alt_diagnostic(p = 2, type = "PAD")
loc_alt_diagnostic(p = 3, type = "PAD")
loc_alt_diagnostic(p = 4, type = "PAD")
loc_alt_diagnostic(p = 5, type = "PAD")
loc_alt_diagnostic(p = 11, type = "PAD")
## Local alternatives for the PRt test
```

}

planets 33

```
loc_alt_diagnostic(p = 2, type = "PRt")
loc_alt_diagnostic(p = 3, type = "PRt")
loc_alt_diagnostic(p = 4, type = "PRt")
loc_alt_diagnostic(p = 5, type = "PRt")
loc_alt_diagnostic(p = 11, type = "PRt")
```

planets

Planet orbits

# Description

Planet orbits data from the JPL Keplerian Elements for Approximate Positions of the Major Planets. The normal vector of a planet orbit represents is a vector on  $S^2$ .

# Usage

planets

## **Format**

A data frame with 9 rows and 3 variables:

planet names of the planets and Pluto.

i inclination; the orbit's plane angle with respect to the ecliptic plane, in radians in  $[0, \pi]$ .

om longitude of the ascending node; the counterclockwise angle from the vector pointing to the First Point of Aries and that pointing to the ascending node (the intersection between orbit and ecliptic plane), in radians in  $[0, 2\pi)$ . (Both vectors are heliocentric and within the ecliptic plane.)

# **Details**

The normal vector to the ecliptic plane of the planet with inclination i and longitude of the ascending node  $\omega$  is

$$(\sin(i)\sin(\omega), -\sin(i)\cos(\omega), \cos(i))'$$
.

The script performing the data preprocessing is available at planets.R. The data was retrieved on 2020-05-16.

#### Source

Table 2a in https://ssd.jpl.nasa.gov/planets/approx\_pos.html

34 Pn

## **Examples**

Pn

Utilities for projected-ecdf statistics of spherical uniformity

# **Description**

Computation of the kernels

$$\psi_p^W(\theta) := \int_{-1}^1 A_x(\theta) \, \mathrm{d}W(F_p(x)),$$

where  $A_x(\theta)$  is the proportion of area surface of  $S^{p-1}$  covered by the intersection of two hyperspherical caps with common solid angle  $\pi - \cos^{-1}(x)$  and centers separated by an angle  $\theta \in [0, \pi]$ ,  $F_p$  is the distribution function of the projected spherical uniform distribution, and W is a measure on [0, 1].

Also, computation of the Gegenbauer coefficients of  $\psi_p^W$ :

$$b_{k,p}^W := \frac{1}{c_{k,p}} \int_0^\pi \psi_p^W(\theta) C_k^{p/2-1}(\cos \theta) \,\mathrm{d}\theta.$$

These coefficients can also be computed via

$$b_{k,p}^{W} = \int_{-1}^{1} a_{k,p}^{x} dW(F_{p}(x))$$

for a certain function  $x \to a_{k,p}^x$ . They serve to define projected alternatives to uniformity.

Pn 35

# Usage

```
psi_Pn(theta, q, type, Rothman_t = 1/3, tilde = FALSE, psi_Gauss = TRUE,
    psi_N = 320, tol = 1e-06)

Gegen_coefs_Pn(k, p, type, Rothman_t = 1/3, Gauss = TRUE, N = 320,
    tol = 1e-06, verbose = FALSE)

akx(x, p, k, sqr = FALSE)

f_locdev_Pn(p, type, K = 1000, N = 320, K_max = 10000, thre = 0.001,
    Rothman_t = 1/3, verbose = FALSE)
```

# Arguments

theta	vector with values in $[0, \pi]$ .
q	integer giving the dimension of the sphere $S^q$ .
type	type of projected-ecdf test statistic. Must be either "PCvM" (Cramér–von Mises), "PAD" (Anderson–Darling), or "PRt" (Rothman).
Rothman_t	t parameter for the Rothman test, a real in $(0,1).$ Defaults to 1 / 3.
tilde	include the constant and bias term? Defaults to FALSE.
psi_Gauss	use a Gauss-Legendre quadrature rule with psi_N nodes in the computation of the kernel function? Defaults to TRUE.
psi_N	number of points used in the Gauss-Legendre quadrature for computing the kernel function. Defaults to 320.
tol	tolerance passed to integrate's rel.tol and abs.tol if Gauss = FALSE. Defaults to 1e-6.
k	vector with the index of coefficients.
р	integer giving the dimension of the ambient space $\mathbb{R}^p$ that contains $\mathbb{S}^{p-1}$ .
Gauss	use a Gauss-Legendre quadrature rule of N nodes in the computation of the Gegenbauer coefficients? Otherwise, call integrate. Defaults to TRUE.
N	number of points used in the Gauss-Legendre quadrature for computing the Gegenbauer coefficients. Defaults to 320.
verbose	flag to print informative messages. Defaults to FALSE.
х	evaluation points for $a_{k,p}^x$ , a vector with values in $[-1,1]$ .
sqr	return the $signed$ square root of $a_{k,p}^x$ ? Defaults to FALSE.
K	number of equispaced points on $\left[-1,1\right]$ used for evaluating $f$ and then interpolating. Defaults to 1e3.
K_max	integer giving the truncation of the series. Defaults to 1e4.
thre	proportion of norm $not$ explained by the first terms of the truncated series. Defaults to $1e-3$ .

36 Pn

## **Details**

The evaluation of  $\psi^W_p$  and  $b^W_{k,p}$  depends on the type of projected-ecdf statistic:

- PCvM: closed-form expressions for  $\psi_p^W$  and  $b_{k,p}^W$  with p=2,3,4, numerical integration required for  $p\geq 5$ .
- PAD: closed-form expressions for  $\psi_2^W$  and  $b_{k,3}^W$ , numerical integration required for  $\psi_p^W$  with  $p \geq 3$  and  $b_{k,p}^W$  with p = 2 and  $p \geq 4$ .
- PRt: closed-form expressions for  $\psi_p^W$  and  $b_{k,p}^W$  for any  $p \geq 2$ .

See García-Portugués et al. (2023) for more details.

#### Value

- psi\_Pn: a vector of size length(theta) with the evaluation of  $\psi$ .
- Gegen\_coefs\_Pn: a vector of size length(k) containing the coefficients  $b_{k,v}^W$ .
- akx: a matrix of size c(length(x), length(k)) containing the coefficients  $a_{k,v}^x$ .
- $f_{locdev_Pn}$ : the projected alternative f as a function ready to be evaluated.

# Author(s)

Eduardo García-Portugués and Paula Navarro-Esteban.

#### References

García-Portugués, E., Navarro-Esteban, P., Cuesta-Albertos, J. A. (2023) On a projection-based class of uniformity tests on the hypersphere. *Bernoulli*, 29(1):181–204. doi:10.3150/21BEJ1454.

```
# Kernels in the projected-ecdf test statistics
k <- 0:10
coefs <- list()</pre>
(coefs$PCvM <- t(sapply(2:5, function(p)</pre>
 Gegen_coefs_Pn(k = k, p = p, type = "PCvM"))))
(coefs$PAD <- t(sapply(2:5, function(p)</pre>
 Gegen_coefs_Pn(k = k, p = p, type = "PAD"))))
(coefs$PRt <- t(sapply(2:5, function(p)</pre>
 Gegen_coefs_Pn(k = k, p = p, type = "PRt"))))
# Gegenbauer expansion
th <- seq(0, pi, length.out = 501)[-501]
old_par <- par(mfrow = c(3, 4))
for (type in c("PCvM", "PAD", "PRt")) {
 for (p in 2:5) {
    plot(th, psi_Pn(theta = th, q = p - 1, type = type), type = "l",
         main = paste0(type, ", p = ", p), xlab = expression(theta),
         ylab = expression(psi(theta)), axes = FALSE, ylim = c(-1.5, 1))
```

Pn 37

```
axis(1, at = c(0, pi / 4, pi / 2, 3 * pi / 4, pi),
         labels = expression(0, pi / 4, pi / 2, 3 * pi / 4, pi))
    axis(2); box()
    lines(th, Gegen_series(theta = th, coefs = coefs[[type]][p - 1, ],
                           k = k, p = p), col = 2)
  }
}
par(old_par)
# Analytical coefficients vs. numerical integration
test_coef <- function(type, p, k = 0:20) {
  plot(k, log1p(abs(Gegen\_coefs\_Pn(k = k, p = p, type = type))),
       ylab = "Coefficients", main = paste0(type, ", p = ", p))
  points(k, log1p(abs(Gegen_coefs(k = k, p = p, psi = psi_Pn, type = type,
                                  q = p - 1)), col = 2)
  legend("topright", legend = c("log(1 + Gegen_coefs_Pn))",
                                "log(1 + Gegen_coefs(psi_Pn))"),
         1wd = 2, col = 1:2)
}
# PCvM statistic
old_par <- par(mfrow = c(2, 2))
for (p in 2:5) test_coef(type = "PCvM", p = p)
par(old_par)
# PAD statistic
old_par <- par(mfrow = c(2, 2))
for (p in 2:5) test_coef(type = "PAD", p = p)
par(old_par)
# PRt statistic
old_par <- par(mfrow = c(2, 2))
for (p in 2:5) test_coef(type = "PRt", p = p)
par(old_par)
# akx
akx(x = seq(-1, 1, 1 = 5), k = 1:4, p = 2)
akx(x = 0, k = 1:4, p = 3)
# PRt alternative to uniformity
z \leftarrow seq(-1, 1, 1 = 1e3)
p <- c(2:5, 10, 15, 17)
col <- viridisLite::viridis(length(p))</pre>
plot(z, f_locdev_Pn(p = p[1], type = "PRt")(z), type = "s",
     col = col[1], ylim = c(0, 0.6), ylab = expression(f[Rt](z)))
for (k in 2:length(p)) {
  lines(z, f_{locdev_{pn}(p = p[k], type = "PRt")(z), type = "s", col = col[k])
legend("topleft", legend = paste("p =", p), col = col, lwd = 2)
```

38 proj\_unif

proj\_unif

Projection of the spherical uniform distribution

## **Description**

Density, distribution, and quantile functions of the projection of the spherical uniform random variable on an arbitrary direction, that is, the random variable  $\gamma'\mathbf{X}$ , where  $\mathbf{X}$  is uniformly distributed on the (hyper)sphere  $S^{p-1} := \{\mathbf{x} \in R^p : ||\mathbf{x}|| = 1\}, p \geq 2$ , and  $\gamma \in S^{p-1}$  is an *arbitrary* projection direction. Note that the distribution is invariant to the choice of  $\gamma$ . Also, efficient simulation of  $\gamma'\mathbf{X}$ .

### Usage

```
d_proj_unif(x, p, log = FALSE)
p_proj_unif(x, p, log = FALSE)
q_proj_unif(u, p)
r_proj_unif(n, p)
```

## **Arguments**

```
x a vector of size nx or a matrix of size c(nx, 1).

p integer giving the dimension of the ambient space \mathbb{R}^p that contains \mathbb{S}^{p-1}.

log compute the logarithm of the density or distribution?

u vector of probabilities.

n sample size.
```

## Value

A matrix of size c(nx, 1) with the evaluation of the density, distribution, or quantile function at x or u. For  $r\_proj\_unif$ , a random vector of size n.

### Author(s)

Eduardo García-Portugués and Paula Navarro-Esteban.

```
# Density function
curve(d_proj_unif(x, p = 2), from = -2, to = 2, n = 2e2, ylim = c(0, 2))
curve(d_proj_unif(x, p = 3), n = 2e2, col = 2, add = TRUE)
curve(d_proj_unif(x, p = 4), n = 2e2, col = 3, add = TRUE)
curve(d_proj_unif(x, p = 5), n = 2e2, col = 4, add = TRUE)
curve(d_proj_unif(x, p = 6), n = 2e2, col = 5, add = TRUE)
```

```
# Distribution function
curve(p_proj_unif(x, p = 2), from = -2, to = 2, n = 2e2, ylim = c(0, 1))
curve(p_proj_unif(x, p = 3), n = 2e2, col = 2, add = TRUE)
curve(p_proj_unif(x, p = 4), n = 2e2, col = 3, add = TRUE)
curve(p_proj_unif(x, p = 5), n = 2e2, col = 4, add = TRUE)
curve(p_proj_unif(x, p = 6), n = 2e2, col = 5, add = TRUE)

# Quantile function
curve(q_proj_unif(u = x, p = 2), from = 0, to = 1, n = 2e2, ylim = c(-1, 1))
curve(q_proj_unif(u = x, p = 3), n = 2e2, col = 2, add = TRUE)
curve(q_proj_unif(u = x, p = 4), n = 2e2, col = 3, add = TRUE)
curve(q_proj_unif(u = x, p = 5), n = 2e2, col = 4, add = TRUE)
curve(q_proj_unif(u = x, p = 6), n = 2e2, col = 5, add = TRUE)

# Sampling
hist(r_proj_unif(n = 1e4, p = 4), freq = FALSE, breaks = 50)
curve(d_proj_unif(x, p = 4), n = 2e2, col = 3, add = TRUE)
```

Psi

Shortest angles matrix

## Description

Efficient computation of the shortest angles matrix  $\Psi$ , defined as

$$\Psi_{ij} := \cos^{-1}(\mathbf{X}_i'\mathbf{X}_i), \quad i, j = 1, \dots, n,$$

for a sample  $X_1, ..., X_n \in S^{p-1} := \{ \mathbf{x} \in R^p : ||\mathbf{x}|| = 1 \}, p \ge 2.$ 

For a circular sample  $\Theta_1, \ldots, \Theta_n \in [0, 2\pi)$ ,  $\Psi$  can be expressed as

$$\Psi_{ij} = \pi - |\pi - |\Theta_i - \Theta_j||, \quad i, j = 1, \dots, n.$$

## Usage

```
Psi_mat(data, ind_tri = 0L, use_ind_tri = FALSE, scalar_prod = FALSE,
    angles_diff = FALSE)

upper_tri_ind(n)
```

#### **Arguments**

data

an array of size c(n, p, M) containing the Cartesian coordinates of M samples of size n of directions on  $S^{p-1}$ . Alternatively if p = 2, an array of size c(n, 1, M) containing the angles on  $[0,2\pi)$  of the M circular samples of size n on  $S^1$ . Must not contain NA's.

ind\_tri

if use\_ind\_tri = TRUE, the vector of 0-based indexes provided by upper\_tri\_ind(n), which allows to extract the upper triangular part of the matrix  $\Psi$ . See the examples.

40 Psi

use_ind_tri	use the already computed vector index ind_tri? If FALSE (default), ind_tri is computed internally.
scalar_prod	return the scalar products $\mathbf{X}_i'\mathbf{X}$ instead of the shortest angles? Only taken into account for data in <i>Cartesian</i> form. Defaults to FALSE.
angles_diff	return the (unwrapped) angles difference $\Theta_i - \Theta_j$ instead of the shortest angles? Only taken into account for data in <i>angular</i> form. Defaults to FALSE.
n	sample size, used to determine the index vector that gives the upper triangular part of $\Psi$ .

#### Value

- Psi\_mat: a matrix of size c(n \* (n 1) / 2, M) containing, for each column, the vector half of Ψ for each of the M samples.
- upper\_tri\_ind: a matrix of size n \* (n 1) / 2 containing the 0-based linear indexes for extracting the upper triangular matrix of a matrix of size c(n, n), diagonal excluded, assuming column-major order.

## Warning

Be careful on avoiding the next bad usages of Psi\_mat, which will produce spurious results:

- The directions in data do *not* have unit norm when Cartesian coordinates are employed.
- The entries of data are *not* in  $[0, 2\pi)$  when polar coordinates are employed.
- ind\_tri is a vector of size n \* (n 1) / 2 that does *not* contain the indexes produced by upper\_tri\_ind(n).

```
# Shortest angles
n <- 5
X \leftarrow r_unif_sph(n = n, p = 2, M = 2)
Theta <- X_to_Theta(X)</pre>
dim(Theta) \leftarrow c(n, 1, 2)
Psi_mat(X)
Psi_mat(Theta)
# Precompute ind_tri
ind_tri <- upper_tri_ind(n)</pre>
Psi_mat(X, ind_tri = ind_tri, use_ind_tri = TRUE)
# Compare with R
A <- acos(tcrossprod(X[, , 1]))
ind <- upper.tri(A)</pre>
A[ind]
# Reconstruct matrix
Psi_vec <- Psi_mat(Theta[, , 1, drop = FALSE])</pre>
Psi <- matrix(0, nrow = n, ncol = n)
Psi[upper.tri(Psi)] <- Psi_vec
Psi <- Psi + t(Psi)
```

p\_Kolmogorov

Asymptotic distributions for circular uniformity statistics

## **Description**

Computation of the asymptotic null distributions of circular uniformity statistics.

## Usage

```
p_Kolmogorov(x, K_Kolmogorov = 25L, alternating = TRUE)
d_Kolmogorov(x, K_Kolmogorov = 25L, alternating = TRUE)
p_{cir_stat_Ajne(x, K_Ajne = 15L)}
d_{cir_stat_Ajne(x, K_Ajne = 15L)}
p_cir_stat_Bingham(x)
d_cir_stat_Bingham(x)
p_cir_stat_Greenwood(x)
d_cir_stat_Greenwood(x)
p_cir_stat_Gini(x)
d_cir_stat_Gini(x)
p_cir_stat_Gini_squared(x)
d_cir_stat_Gini_squared(x)
p_cir_stat_Hodges_Ajne2(x, n, asymp_std = FALSE)
p_cir_stat_Hodges_Ajne(x, n, exact = TRUE, asymp_std = FALSE)
d_cir_stat_Hodges_Ajne(x, n, exact = TRUE, asymp_std = FALSE)
p_cir_stat_Kuiper(x, n, K_Kuiper = 12L, second_term = TRUE,
  Stephens = FALSE)
d_cir_stat_Kuiper(x, n, K_Kuiper = 12L, second_term = TRUE,
  Stephens = FALSE)
p_cir_stat_Log_gaps(x, abs_val = TRUE)
```

```
d_cir_stat_Log_gaps(x, abs_val = TRUE)
p_cir_stat_Max_uncover(x)
d_cir_stat_Max_uncover(x)
p_cir_stat_Num_uncover(x)
d_cir_stat_Num_uncover(x)
p_cir_stat_Pycke(x)
d_cir_stat_Pycke(x)
p_cir_stat_Vacancy(x)
d_cir_stat_Vacancy(x)
p_{cir_stat_Watson}(x, n = 0L, K_Watson = 25L, Stephens = FALSE)
d_{cir_stat_Watson}(x, n = 0L, K_{watson} = 25L, Stephens = FALSE)
p_cir_stat_Watson_1976(x, K_Watson_1976 = 8L, N = 40L)
d_cir_stat_Watson_1976(x, K_Watson_1976 = 8L)
p_cir_stat_Range(x, n, max_gap = TRUE)
d_cir_stat_Range(x, n, max_gap = TRUE)
p_cir_stat_Rao(x)
d_cir_stat_Rao(x)
p_cir_stat_Rayleigh(x)
d_cir_stat_Rayleigh(x)
p_cir_stat_Bakshaev(x, K_max = 1000, thre = 0, method = "I", ...)
d_cir_stat_Bakshaev(x, K_max = 1000, thre = 0, method = "I", ...)
p_cir_stat_Gine_Fn(x, K_max = 1000, thre = 0, method = "I", ...)
d_cir_stat_Gine_Fn(x, K_max = 1000, thre = 0, method = "I", ...)
p_{cir_stat_Gine_Gn(x, K_max = 1000, thre = 0, method = "I", ...)
```

```
d_{cir_stat_Gine_Gn(x, K_max = 1000, thre = 0, method = "I", ...)
p_cir_stat_Hermans_Rasson(x, K_max = 1000, thre = 0, method = "I", ...)
d_cir_stat_Hermans_Rasson(x, K_max = 1000, thre = 0, method = "I", ...)
p_cir_stat_PAD(x, K_max = 1000, thre = 0, method = "I", ...)
d_{cir_stat_PAD}(x, K_{max} = 1000, thre = 0, method = "I", ...)
p_{cir_stat_PCvM}(x, K_max = 1000, thre = 0, method = "I", ...)
d_{cir_stat_PCvM}(x, K_{max} = 1000, thre = 0, method = "I", ...)
p_{cir_stat_PRt(x, t = 1/3, K_max = 1000, thre = 0, method = "I", ...)}
d_{cir}stat_PRt(x, t = 1/3, K_max = 1000, thre = 0, method = "I", ...)
p_cir_stat_Poisson(x, rho = 0.5, K_max = 1000, thre = 0, method = "I",
d_cir_stat_Poisson(x, rho = 0.5, K_max = 1000, thre = 0, method = "I",
p_{cir_stat_Pycke_q(x, q = 0.5, K_max = 1000, thre = 0, method = "I",
  ...)
d_{cir_stat_pycke_q(x, q = 0.5, K_max = 1000, thre = 0, method = "I",
p_cir_stat_Rothman(x, t = 1/3, K_max = 1000, thre = 0, method = "I",
d_{cir}stat_Rothman(x, t = 1/3, K_max = 1000, thre = 0, method = "I",
  ...)
p_{cir_stat_Riesz}(x, s = 1, K_{max} = 1000, thre = 0, method = "I", ...)
d_{cir_stat_Riesz}(x, s = 1, K_{max} = 1000, thre = 0, method = "I", ...)
p_{cir_stat_Sobolev}(x, vk2 = c(0, 0, 1), method = "I", ...)
d_{cir_stat_Sobolev}(x, vk2 = c(0, 0, 1), method = "I", ...)
p_cir_stat_Softmax(x, kappa = 1, K_max = 1000, thre = 0, method = "I",
d_cir_stat_Softmax(x, kappa = 1, K_max = 1000, thre = 0, method = "I",
```

...)

# Arguments

х	a vector of size nx or a matrix of size c(nx, 1).
K_Kolmogorov, k	(_Kuiper, K_Watson, K_Watson_1976, K_Ajne integer giving the truncation of the series present in the null asymptotic distributions. For the Kolmogorov-Smirnov-related series defaults to 25; for the others series defaults to a smaller number.
alternating	use the alternating series expansion for the distribution of the Kolmogorov-Smirnov statistic? Defaults to TRUE.
n	sample size employed for computing the statistic.
asymp_std	compute the distribution associated to the normalized Hodges-Ajne statistic? Defaults to FALSE.
exact	use the exact distribution for the Hodges-Ajne statistic? Defaults to TRUE.
second_term	use the second-order series expansion for the distribution of the Kuiper statistic? Defaults to TRUE.
Stephens	compute Stephens (1970) modification so that the null distribution of the is less dependent on the sample size? The modification does not alter the test decision.
abs_val	compute the distribution associated to the absolute value of the Darling's log gaps statistic? Defaults to TRUE.
N	number of points used in the Gauss-Legendre quadrature. Defaults to 40.
max_gap	compute the distribution associated to the maximum gap for the range statistic? Defaults to TRUE.
K_max	integer giving the truncation of the series that compute the asymptotic p-value of a Sobolev test. Defaults to 1e3.
thre	error threshold for the tail probability given by the the first terms of the truncated series of a Sobolev test. Defaults to $\emptyset$ (no further truncation).
method	method for approximating the density, distribution, or quantile function of the weighted sum of chi squared random variables. Must be "I" (Imhof), "SW" (Satterthwaite–Welch), "HBE" (Hall–Buckley–Eagleson), or "MC" (Monte Carlo; only for distribution or quantile functions). Defaults to "I".
	further parameters passed to p_Sobolev or d_Sobolev (such as x_tail).
t	t parameter for the Rothman and Cressie tests, a real in $(0,1).$ Defaults to 1 / 3.
rho	$\rho$ parameter for the Poisson test, a real in $[0,1).$ Defaults to 0.5.
q	q parameter for the Pycke " $q-test$ ", a real in $(0,1).$ Defaults to 1 / 2.
S	s parameter for the $s$ -Riesz test, a real in $(0,2)$ . Defaults to 1.
vk2	weights for the finite Sobolev test. A non-negative vector or matrix. Defaults to $c(\emptyset,\emptyset,1)$ .
kappa	$\kappa$ parameter for the Softmax test, a non-negative real. Defaults to 1.

#### **Details**

Descriptions and references for most of the tests are available in García-Portugués and Verdebout (2018).

#### Value

A matrix of size c(nx, 1) with the evaluation of the distribution or density function at x.

#### References

García-Portugués, E. and Verdebout, T. (2018) An overview of uniformity tests on the hypersphere. *arXiv*:1804.00286. doi:10.48550/arXiv.1804.00286.

```
# Aine
curve(d\_cir\_stat\_Ajne(x), to = 1.5, n = 2e2, ylim = c(0, 4))
curve(p\_cir\_stat\_Ajne(x), n = 2e2, col = 2, add = TRUE)
# Bakshaev
curve(d_cir_stat_Bakshaev(x, method = "HBE"), to = 6, n = 2e2,
      vlim = c(0, 1))
curve(p_cir_stat_Bakshaev(x, method = "HBE"), n = 2e2, add = TRUE, col = 2)
# Bingham
curve(d\_cir\_stat\_Bingham(x), to = 12, n = 2e2, ylim = c(0, 1))
curve(p_cir_stat_Bingham(x), n = 2e2, col = 2, add = TRUE)
curve(d_cir_stat_Greenwood(x), from = -6, to = 6, n = 2e2, ylim = c(0, 1))
curve(p_cir_stat_Greenwood(x), n = 2e2, col = 2, add = TRUE)
# Hermans-Rasson
curve(p_cir_stat_Hermans_Rasson(x, method = "HBE"), to = 10, n = 2e2,
      ylim = c(0, 1)
curve(d_cir_stat_Hermans_Rasson(x, method = "HBE"), n = 2e2, add = TRUE,
      col = 2)
# Hodges-Ajne
plot(25:45, d_cir_stat_Hodges_Ajne(cbind(25:45), n = 50), type = "h",
     lwd = 2, ylim = c(0, 1)
lines(25:45, p_cir_stat_Hodges_Ajne(cbind(25:45), n = 50), type = "s",
      col = 2)
# Kolmogorov-Smirnov
curve(d_Kolmogorov(x), to = 3, n = 2e2, ylim = c(0, 2))
curve(p_Kolmogorov(x), n = 2e2, col = 2, add = TRUE)
# Kuiper
curve(d_cir_stat_Kuiper(x, n = 50), to = 3, n = 2e2, ylim = c(0, 2))
curve(p_cir_stat_Kuiper(x, n = 50), n = 2e2, col = 2, add = TRUE)
# Kuiper and Watson with Stephens modification
```

```
curve(d_cir_stat_Kuiper(x, n = 8, Stephens = TRUE), to = 2.5, n = 2e2,
      ylim = c(0, 10)
curve(d_cir_stat_Watson(x, n = 8, Stephens = TRUE), n = 2e2, lty = 2,
      add = TRUE)
n \leftarrow c(10, 20, 30, 40, 50, 100, 500)
col <- rainbow(length(n))</pre>
for (i in seq_along(n)) {
  curve(d_cir_stat_Kuiper(x, n = n[i], Stephens = TRUE), n = 2e2,
        col = col[i], add = TRUE)
  curve(d_cir_stat_Watson(x, n = n[i], Stephens = TRUE), n = 2e2,
        col = col[i], lty = 2, add = TRUE)
}
# Maximum uncovered spacing
curve(d_cir_stat_Max\_uncover(x), from = -3, to = 6, n = 2e2, ylim = c(0, 1))
curve(p_cir_stat_Max_uncover(x), n = 2e2, col = 2, add = TRUE)
# Number of uncovered spacing
curve(d\_cir\_stat\_Num\_uncover(x), from = -4, to = 4, n = 2e2, ylim = c(0, 1))
curve(p_cir_stat_Num_uncover(x), n = 2e2, col = 2, add = TRUE)
# Log gaps
curve(d_cir_stat_Log_gaps(x), from = -1, to = 4, n = 2e2, ylim = c(0, 1))
curve(p_cir_stat_Log_gaps(x), n = 2e2, col = 2, add = TRUE)
# Gine Fn
curve(d\_cir\_stat\_Gine\_Fn(x, method = "HBE"), to = 2.5, n = 2e2,
      ylim = c(0, 2))
curve(p_cir_stat_Gine_Fn(x, method = "HBE"), n = 2e2, add = TRUE, col = 2)
# Gine Gn
curve(d\_cir\_stat\_Gine\_Gn(x, method = "HBE"), to = 2.5, n = 2e2,
      ylim = c(0, 2))
curve(p_cir_stat_Gine_Gn(x, method = "HBE"), n = 2e2, add = TRUE, col = 2)
# Gini mean difference
curve(d_cir_stat_Gini(x), from = -4, to = 4, n = 2e2, ylim = c(0, 1))
curve(p_cir_stat_Gini(x), n = 2e2, col = 2, add = TRUE)
# Gini mean squared difference
curve(d_cir_stat_Gini_squared(x), from = -10, to = 10, n = 2e2,
      ylim = c(0, 1))
curve(p\_cir\_stat\_Gini\_squared(x), n = 2e2, col = 2, add = TRUE)
curve(d\_cir\_stat\_PAD(x, method = "HBE"), to = 3, n = 2e2, ylim = c(0, 1.5))
curve(p_cir_stat_PAD(x, method = "HBE"), n = 2e2, add = TRUE, col = 2)
# PCvM
curve(d\_cir\_stat\_PCvM(x, method = "HBE"), to = 4, n = 2e2, ylim = c(0, 2))
curve(p_cir_stat_PCvM(x, method = "HBE"), n = 2e2, add = TRUE, col = 2)
# PRt
```

```
curve(d\_cir\_stat\_PRt(x, method = "HBE"), n = 2e2, ylim = c(0, 5))
curve(p_cir_stat_PRt(x, method = "HBE"), n = 2e2, add = TRUE, col = 2)
# Poisson
curve(d\_cir\_stat\_Poisson(x, method = "HBE"), from = -1, to = 5, n = 2e2,
      ylim = c(0, 2))
curve(p_cir_stat_Poisson(x, method = "HBE"), n = 2e2, col = 2, add = TRUE)
# Pycke
curve(d_cir_stat_Pycke(x), from = -5, to = 10, n = 2e2, ylim = c(0, 1))
curve(p_cir_stat_Pycke(x), n = 2e2, col = 2, add = TRUE)
# Pycke q
curve(d\_cir\_stat\_Pycke\_q(x, method = "HBE"), to = 15, n = 2e2,
      ylim = c(0, 1))
curve(p_cir_stat_Pycke_q(x, method = "HBE"), n = 2e2, add = TRUE, col = 2)
# Range
curve(d\_cir\_stat\_Range(x, n = 50), to = 2, n = 2e2, ylim = c(0, 4))
curve(p\_cir\_stat\_Range(x, n = 50), n = 2e2, col = 2, add = TRUE)
curve(d\_cir\_stat\_Rao(x), from = -6, to = 6, n = 2e2, ylim = c(0, 1))
curve(p\_cir\_stat\_Rao(x), n = 2e2, col = 2, add = TRUE)
# Rayleigh
curve(d\_cir\_stat\_Rayleigh(x), to = 12, n = 2e2, ylim = c(0, 1))
curve(p_cir_stat_Rayleigh(x), n = 2e2, col = 2, add = TRUE)
# Riesz
curve(d_cir_stat_Riesz(x, method = "HBE"), to = 6, n = 2e2,
      ylim = c(0, 1))
curve(p_cir_stat_Riesz(x, method = "HBE"), n = 2e2, add = TRUE, col = 2)
curve(d_cir_stat_Rothman(x, method = "HBE"), n = 2e2, ylim = c(0, 5))
curve(p_cir_stat_Rothman(x, method = "HBE"), n = 2e2, add = TRUE, col = 2)
curve(d\_cir\_stat\_Vacancy(x), from = -4, to = 4, n = 2e2, ylim = c(0, 1))
curve(p\_cir\_stat\_Vacancy(x), n = 2e2, col = 2, add = TRUE)
curve(d\_cir\_stat\_Watson(x), to = 0.5, n = 2e2, ylim = c(0, 15))
curve(p\_cir\_stat\_Watson(x), n = 2e2, col = 2, add = TRUE)
# Watson (1976)
curve(d\_cir\_stat\_Watson\_1976(x), to = 1.5, n = 2e2, ylim = c(0, 3))
curve(p\_cir\_stat\_Watson\_1976(x), n = 2e2, col = 2, add = TRUE)
# Softmax
curve(d\_cir\_stat\_Softmax(x, method = "HBE"), to = 3, n = 2e2, ylim = c(0, 2))
curve(p_cir_stat_Softmax(x, method = "HBE"), n = 2e2, col = 2, add = TRUE)
```

```
# Sobolev vk2 <- c(0.5, 0) curve(d\_cir\_stat\_Sobolev(x = x, vk2 = vk2), to = 3, n = 2e2, ylim = c(0, 2)) curve(p\_cir\_stat\_Sobolev(x = x, vk2 = vk2), n = 2e2, col = 2, add = TRUE)
```

p\_sph\_stat\_Bingham

Asymptotic distributions for spherical uniformity statistics

## **Description**

Computation of the asymptotic null distributions of spherical uniformity statistics.

### Usage

```
p_sph_stat_Bingham(x, p)
d_sph_stat_Bingham(x, p)
p_sph_stat_CJ12(x, regime = 1L, beta = 0)
d_sph_stat_CJ12(x, regime = 3L, beta = 0)
p_sph_stat_Rayleigh(x, p)
d_sph_stat_Rayleigh(x, p)
p_sph_stat_Rayleigh_HD(x, p)
d_sph_stat_Rayleigh_HD(x, p)
p_sph_stat_Ajne(x, p, K_max = 1000, thre = 0, method = "I", ...)
d_{sph_stat_Ajne(x, p, K_{max} = 1000, thre = 0, method = "I", ...)
p_sph_stat_Bakshaev(x, p, K_max = 1000, thre = 0, method = "I", ...)
d_sph_stat_Bakshaev(x, p, K_max = 1000, thre = 0, method = "I", ...)
p_sph_stat_Gine_Fn(x, p, K_max = 1000, thre = 0, method = "I", ...)
d_{sph_stat_{ine}}Fn(x, p, K_{max} = 1000, thre = 0, method = "I", ...)
p_sph_stat_Gine_Gn(x, p, K_max = 1000, thre = 0, method = "I", ...)
d_{sph_stat_Gine_Gn(x, p, K_{max} = 1000, thre = 0, method = "I", ...)
```

```
p_sph_stat_PAD(x, p, K_max = 1000, thre = 0, method = "I", ...)
d_{sph_stat_PAD}(x, p, K_{max} = 1000, thre = 0, method = "I", ...)
p_sph_stat_PCvM(x, p, K_max = 1000, thre = 0, method = "I", ...)
d_{sph_stat_PCvM}(x, p, K_{max} = 1000, thre = 0, method = "I", ...)
p_sph_stat_Poisson(x, p, rho = 0.5, K_max = 1000, thre = 0,
 method = "I", ...)
d_{sph_stat_Poisson}(x, p, rho = 0.5, K_{max} = 1000, thre = 0,
  method = "I", \ldots)
p_sph_stat_PRt(x, p, t = 1/3, K_max = 1000, thre = 0, method = "I",
d_{sph_stat_PRt(x, p, t = 1/3, K_{max} = 1000, thre = 0, method = "I",
  ...)
p_sph_stat_Riesz(x, p, s = 1, K_max = 1000, thre = 0, method = "I",
d_{sph_stat_Riesz}(x, p, s = 1, K_{max} = 1000, thre = 0, method = "I",
p_sph_stat_Sobolev(x, p, vk2 = c(0, 0, 1), method = "I", ...)
d_{sph_stat_sobolev}(x, p, vk2 = c(0, 0, 1), method = "I", ...)
p_sph_stat_Softmax(x, p, kappa = 1, K_max = 1000, thre = 0,
 method = "I", \ldots)
d_{sph_stat_Softmax}(x, p, kappa = 1, K_{max} = 1000, thre = 0,
  method = "I", ...)
p_sph_stat_Stereo(x, p, a = 0, K_max = 1000, method = "I", ...)
d_{sph_stat_Stereo}(x, p, a = 0, K_{max} = 1000, method = "I", ...)
```

## **Arguments**

a vector of size nx or a matrix of size c(nx, 1). Χ integer giving the dimension of the ambient space  $\mathbb{R}^p$  that contains  $\mathbb{S}^{p-1}$ . type of asymptotic regime for the CJ12 test, either 1 (sub-exponential regime), regime 2 (exponential), or 3 (super-exponential; default). beta

 $\beta$  parameter in the exponential regime of the CJ12 test, a non-negative real.

Defaults to 0.

K_max	integer giving the truncation of the series that compute the asymptotic p-value of a Sobolev test. Defaults to 1e3.
thre	error threshold for the tail probability given by the the first terms of the truncated series of a Sobolev test. Defaults to $\emptyset$ (no further truncation).
method	method for approximating the density, distribution, or quantile function of the weighted sum of chi squared random variables. Must be "I" (Imhof), "SW" (Satterthwaite–Welch), "HBE" (Hall–Buckley–Eagleson), or "MC" (Monte Carlo; only for distribution or quantile functions). Defaults to "I".
	further parameters passed to p_Sobolev or d_Sobolev (such as x_tail).
rho	$\rho$ parameter for the Poisson test, a real in $[0,1)$ . Defaults to $\emptyset.5$ .
t	t parameter for the Rothman and Cressie tests, a real in $(0,1).$ Defaults to 1 / 3.
S	s parameter for the $s$ -Riesz test, a real in $(0,2)$ . Defaults to 1.
vk2	weights for the finite Sobolev test. A non-negative vector or matrix. Defaults to $c(0, 0, 1)$ .
kappa	$\kappa$ parameter for the Softmax test, a non-negative real. Defaults to 1.
а	either:
	• $a_n = a/n$ parameter used in the length of the arcs of the coverage-based tests. Must be positive. Defaults to $2 * pi$ .

## **Details**

Descriptions and references on most of the asymptotic distributions are available in García-Portugués and Verdebout (2018).

• a parameter for the Stereo test, a real in [-1, 1]. Defaults to 0.

### Value

- r\_sph\_stat\_\*: a matrix of size c(n, 1) containing the sample.
- $p_{sph_stat_*}$ ,  $d_{sph_stat_*}$ : a matrix of size c(nx, 1) with the evaluation of the distribution or density functions at x.

```
# CJ12
curve(d_{sph_stat_CJ12}(x, regime = 1), from = -10, to = 10, n = 2e2,
      ylim = c(0, 1)
curve(d_sph_stat_CJ12(x, regime = 2, beta = 0.1), n = 2e2, col = 2,
      add = TRUE)
curve(d_sph_stat_CJ12(x, regime = 3), n = 2e2, col = 3, add = TRUE)
curve(p_sph_stat_CJ12(x, regime = 1), n = 2e2, col = 1, add = TRUE)
curve(p\_sph\_stat\_CJ12(x, regime = 2, beta = 0.1), n = 2e2, col = 2,
      add = TRUE)
curve(p\_sph\_stat\_CJ12(x, regime = 3), col = 3, add = TRUE)
# Gine Fn
curve(d_{sph_stat_Gine_Fn(x, p = 3, method = "HBE"), to = 2, n = 2e2,
      ylim = c(0, 2))
curve(p\_sph\_stat\_Gine\_Fn(x, p = 3, method = "HBE"), n = 2e2, col = 2,
      add = TRUE)
# Gine Gn
curve(d_{sph_stat_Gine_Gn(x, p = 3, method = "HBE"), to = 1.5, n = 2e2,
      ylim = c(0, 2.5)
curve(p\_sph\_stat\_Gine\_Gn(x, p = 3, method = "HBE"), n = 2e2, col = 2,
      add = TRUE)
# PAD
curve(d_sph_stat_PAD(x, p = 3, method = "HBE"), to = 3, n = 2e2,
      ylim = c(0, 1.5)
curve(p\_sph\_stat\_PAD(x, p = 3, method = "HBE"), n = 2e2, col = 2,
      add = TRUE)
# PCvM
curve(d_sph_stat_PCvM(x, p = 3, method = "HBE"), to = 0.6, n = 2e2,
      ylim = c(0, 7))
curve(p\_sph\_stat\_PCvM(x, p = 3, method = "HBE"), n = 2e2, col = 2,
      add = TRUE)
# Poisson
curve(d_{sph_stat_Poisson}(x, p = 3, method = "HBE"), to = 2, n = 2e2,
      ylim = c(0, 2))
curve(p_sph_stat_Poisson(x, p = 3, method = "HBE"), n = 2e2, col = 2,
      add = TRUE)
# PRt
curve(d_sph_stat_PRt(x, p = 3, method = "HBE"), n = 2e2, ylim = c(0, 5))
curve(p\_sph\_stat\_PRt(x, p = 3, method = "HBE"), n = 2e2, col = 2, add = TRUE)
# Rayleigh
curve(d_sph_stat_Rayleigh(x, p = 3), to = 15, n = 2e2, ylim = c(0, 1))
curve(p\_sph\_stat\_Rayleigh(x, p = 3), n = 2e2, col = 2, add = TRUE)
# HD-standardized Rayleigh
curve(d_{sph_stat_Rayleigh_HD}(x, p = 3), from = -4, to = 4, n = 2e2,
      ylim = c(0, 1))
```

52 rhea

```
curve(p\_sph\_stat\_Rayleigh\_HD(x, p = 3), n = 2e2, col = 2, add = TRUE)
# Riesz
curve(d_sph_stat_Riesz(x, p = 3, method = "HBE"), n = 2e2, from = 0, to = 5,
      ylim = c(0, 2))
curve(p\_sph\_stat\_Riesz(x, p = 3, method = "HBE"), n = 2e2, col = 2,
      add = TRUE)
# Sobolev
x < - seq(-1, 5, by = 0.05)
vk2 <- diag(rep(0.3, 2))
matplot(x, d_sph_stat_Sobolev(x = x, vk2 = vk2, p = 3), type = "1",
        ylim = c(0, 1), lty = 1)
matlines(x, p\_sph\_stat\_Sobolev(x = x, vk2 = vk2, p = 3), lty = 1)
matlines(x, d_sph_stat_Sobolev(x = x, vk2 = vk2 + 0.01, p = 3), lty = 2)
matlines(x, p\_sph\_stat\_Sobolev(x = x, vk2 = vk2 + 0.01, p = 3), lty = 2)
# Softmax
curve(d_{sph_stat_Softmax}(x, p = 3, method = "HBE"), to = 2, n = 2e2,
      ylim = c(0, 2))
curve(p\_sph\_stat\_Softmax(x, p = 3, method = "HBE"), n = 2e2, col = 2,
      add = TRUE)
# Stereo
curve(d_{sph_stat_Stereo}(x, p = 4, method = "HBE"), from=-5, to = 10, n = 2e2,
      ylim = c(0, 2))
curve(p\_sph\_stat\_Stereo(x, p = 4, method = "HBE"), n = 2e2, col = 2,
      add = TRUE)
```

rhea

Rhea craters from Hirata (2016)

## **Description**

Craters on Rhea from Hirata (2016).

## Usage

rhea

## Format

A data frame with 3596 rows and 4 variables:

```
name name of the crater (if named).
```

diameter diameter of the crater (in km).

**theta** longitude angle  $\theta \in [0, 2\pi)$  of the crater center.

**phi** latitude angle  $\phi \in [-\pi/2, \pi/2]$  of the crater center.

r\_alt 53

## **Details**

The  $(\theta, \phi)$  angles are such their associated planetocentric coordinates are:

```
(\cos(\phi)\cos(\theta),\cos(\phi)\sin(\theta),\sin(\phi))',
```

with (0,0,1)' denoting the north pole.

The script performing the data preprocessing is available at rhea. R.

## Source

```
https://agupubs.onlinelibrary.wiley.com/action/downloadSupplement?doi=10.1002%2F2015JE004940&file=jgre20485-sup-0002-TableS1.txt
```

## References

Hirata, N. (2016) Differential impact cratering of Saturn's satellites by heliocentric impactors. *Journal of Geophysical Research: Planets*, 121:111–117. doi:10.1002/2015JE004940

## **Examples**

r\_alt

Sample non-uniformly distributed spherical data

## **Description**

Simple simulation of prespecified non-uniform spherical distributions: von Mises–Fisher (vMF), Mixture of vMF (MvMF), Angular Central Gaussian (ACG), Small Circle (SC), Watson (W), Cauchy-like (C), Mixture of Cauchy-like (MC), or Uniform distribution with Antipodal-Dependent observations (UAD).

### Usage

```
r_{alt}(n, p, M = 1, alt = "vMF", mu = c(rep(0, p - 1), 1), kappa = 1, nu = 0.5, F_{inv} = NULL, K = 1000, axial_mix = TRUE)
```

54 r\_alt

#### **Arguments**

n	sample size.
p	integer giving the dimension of the ambient space $\mathbb{R}^p$ that contains $\mathbb{S}^{p-1}$ .
М	number of samples of size n. Defaults to 1.
alt	alternative, must be "vMF", "MvMF", "ACG", "SC", "W", "C", "MC", or "UAD". See details below.
mu	location parameter for "vMF", "SC", "W", and "C". Defaults to $c(rep(0, p-1), 1)$ .
kappa	non-negative parameter measuring the strength of the deviation with respect to uniformity (obtained with $\kappa=0$ ).
nu	projection along $e_p$ controlling the modal strip of the small circle distribution. Must be in (-1, 1). Defaults to 0.5.
F_inv	quantile function returned by F_inv_from_f. Used for "SC", "W", and "C". Computed by internally if NULL (default).
K	number of equispaced points on $[-1,1]$ used for evaluating $F^{-1}$ and then interpolating. Defaults to 1e3.
axial_mix	use a mixture of von Mises-Fisher or Cauchy-like that is axial (i.e., symmetrically distributed about the origin)? Defaults to TRUE.

## **Details**

The parameter kappa is used as  $\kappa$  in the following distributions:

- "VMF": von Mises-Fisher distribution with concentration  $\kappa$  and directional mean  $\mu$ .
- "MvMF": equally-weighted mixture of p von Mises-Fisher distributions with common concentration  $\kappa$  and directional means  $\pm \mathbf{e}_1, \ldots, \pm \mathbf{e}_p$  if axial\_mix = TRUE. If axial\_mix = FALSE, then only means with positive signs are considered.
- "ACG": Angular Central Gaussian distribution with diagonal shape matrix with diagonal given by

$$(1,\ldots,1,1+\kappa)/(p+\kappa).$$

- "SC": Small Circle distribution with axis mean  $\mu$  and concentration  $\kappa$  about the projection along the mean,  $\nu$ .
- "W": Watson distribution with axis mean  $\mu$  and concentration  $\kappa$ . The Watson distribution is a particular case of the Bingham distribution.
- "C": Cauchy-like distribution with directional mode  $\mu$  and concentration  $\kappa = \rho/(1-\rho^2)$ . The circular Wrapped Cauchy distribution is a particular case of this Cauchy-like distribution.
- "MC": equally-weighted mixture of p Cauchy-like distributions with common concentration  $\kappa$  and directional means  $\pm \mathbf{e}_1, \ldots, \pm \mathbf{e}_p$  if axial\_mix = TRUE. If axial\_mix = FALSE, then only means with positive signs are considered.

The alternative "UAD" generates a sample formed by  $\lceil n/2 \rceil$  observations drawn uniformly on  $S^{p-1}$  and the remaining observations drawn from a uniform spherical cap distribution of angle  $\pi - \kappa$  about each of the  $\lceil n/2 \rceil$  observations (see unif\_cap). Hence, kappa = 0 corresponds to a spherical cap covering the whole sphere and kappa = pi is a one-point degenerate spherical cap.

Much faster sampling for "SC", "W", "C", and "MC" is achieved providing F\_inv; see examples.

 $r_{-}alt$  55

#### Value

An **array** of size c(n, p, M) with M random samples of size n of non-uniformly-generated directions on  $S^{p-1}$ .

```
## Simulation with p = 2
p <- 2
n <- 50
kappa <- 20
nu < -0.5
angle <- pi / 10
rho <- ((2 * kappa + 1) - sqrt(4 * kappa + 1)) / (2 * kappa)
F_{inv_SC_2} < F_{inv_from_f}(f = function(z) exp(-kappa * (z - nu)^2), p = 2)
F_{inv_W_2} \leftarrow F_{inv_from_f}(f = function(z) exp(kappa * z^2), p = 2)
F_{inv_C_2} \leftarrow F_{inv_from_f} (f = function(z) (1 - rho^2) / 
                              (1 + \text{rho}^2 - 2 * \text{rho} * z)^(p / 2), p = 2)
x1 \leftarrow r_alt(n = n, p = p, alt = "vMF", kappa = kappa)[, , 1]
x2 \leftarrow r_alt(n = n, p = p, alt = "C", F_inv = F_inv_C_2)[, , 1]
x3 \leftarrow r_alt(n = n, p = p, alt = "SC", F_inv = F_inv_SC_2)[, , 1]
x4 \leftarrow r_alt(n = n, p = p, alt = "ACG", kappa = kappa)[, , 1]
x5 \leftarrow r_alt(n = n, p = p, alt = "W", F_inv = F_inv_W_2)[, , 1]
x6 \leftarrow r_alt(n = n, p = p, alt = "MvMF", kappa = kappa)[, , 1]
x7 < -r_{alt}(n = n, p = p, alt = "MC", kappa = kappa)[, , 1]
x8 < -r_{alt}(n = n, p = p, alt = "UAD", kappa = 1 - angle)[, , 1]
r <- runif(n, 0.95, 1.05) # Radius perturbation to improve visualization
plot(r * x1, pch = 16, xlim = c(-1.1, 1.1), ylim = c(-1.1, 1.1), col = 1)
points(r * x2, pch = 16, col = 2)
points(r * x3, pch = 16, col = 3)
plot(r * x4, pch = 16, xlim = c(-1.1, 1.1), ylim = c(-1.1, 1.1), col = 1)
points(r * x5, pch = 16, col = 2)
points(r * x6, pch = 16, col = 3)
points(r * x7, pch = 16, col = 4)
col <- rep(rainbow(n / 2), 2)</pre>
plot(r * x8, pch = 16, xlim = c(-1.1, 1.1), ylim = c(-1.1, 1.1), col = col)
for (i in seq(1, n, by = 2)) lines((r * x8)[i + 0:1, ], col = col[i])
## Simulation with p = 3
n <- 50
p <- 3
kappa <- 20
angle <- pi / 10
nu <- 0.5
rho <- ((2 * kappa + 1) - sqrt(4 * kappa + 1)) / (2 * kappa)
F_{inv_SC_3} \leftarrow F_{inv_from_f}(f = function(z) exp(-kappa * (z - nu)^2), p = 3)
F_{inv_W_3} \leftarrow F_{inv_from_f(f = function(z) exp(kappa * z^2), p = 3)}
F_{inv_C_3} \leftarrow F_{inv_from_f}(f = function(z) (1 - rho^2) / 
                              (1 + \text{rho}^2 - 2 * \text{rho} * z)^(p / 2), p = 3)
x1 \leftarrow r_alt(n = n, p = p, alt = "vMF", kappa = kappa)[, , 1]
x2 \leftarrow r_alt(n = n, p = p, alt = "C", F_inv = F_inv_C_3)[, , 1]
```

56 r\_unif

```
x3 \leftarrow r_alt(n = n, p = p, alt = "SC", F_inv = F_inv_SC_3)[, , 1]
x4 \leftarrow r_alt(n = n, p = p, alt = "ACG", kappa = kappa)[, , 1]
x5 \leftarrow r_alt(n = n, p = p, alt = "W", F_inv = F_inv_W_3)[, , 1]
x6 \leftarrow r_alt(n = n, p = p, alt = "MvMF", kappa = kappa)[, , 1]
x7 \leftarrow r_alt(n = n, p = p, alt = "MC", kappa = kappa)[, , 1]
x8 < -r_alt(n = n, p = p, alt = "UAD", kappa = 1 - angle)[, , 1]
s3d \leftarrow scatterplot3d::scatterplot3d(x1, pch = 16, xlim = c(-1.1, 1.1),
                                      ylim = c(-1.1, 1.1), zlim = c(-1.1, 1.1))
s3dpoints3d(x2, pch = 16, col = 2)
s3d$points3d(x3, pch = 16, col = 3)
s3d \leftarrow scatterplot3d::scatterplot3d(x4, pch = 16, xlim = c(-1.1, 1.1),
                                      ylim = c(-1.1, 1.1), zlim = c(-1.1, 1.1))
s3dpoints3d(x5, pch = 16, col = 2)
s3dpoints3d(x6, pch = 16, col = 3)
s3dpoints3d(x7, pch = 16, col = 4)
col <- rep(rainbow(n / 2), 2)</pre>
s3d \leftarrow scatterplot3d::scatterplot3d(x8, pch = 16, xlim = c(-1.1, 1.1),
                                      ylim = c(-1.1, 1.1), zlim = c(-1.1, 1.1),
                                      color = col)
for (i in seq(1, n, by = 2)) s3dpoints3d(x8[i + 0:1, ], col = col[i],
                                             type = "1")
```

r unif

Sample uniformly distributed circular and spherical data

## **Description**

Simulation of the uniform distribution on  $[0, 2\pi)$  and  $S^{p-1} := \{ \mathbf{x} \in \mathbb{R}^p : ||\mathbf{x}|| = 1 \}, p \geq 2.$ 

## Usage

```
r_unif_cir(n, M = 1L, sorted = FALSE)
r_unif_sph(n, p, M = 1L)
```

## **Arguments**

n sample size. M number of samples of size n. Defaults to 1. sorted return each circular sample sorted? Defaults to FALSE. p integer giving the dimension of the ambient space  $\mathbb{R}^p$  that contains  $\mathbb{S}^{p-1}$ .

## Value

- r\_unif\_cir: a matrix of size c(n, M) with M random samples of size n of uniformly-generated circular data on [0, 2π).
- r\_unif\_sph: an array of size c(n, p, M) with M random samples of size n of uniformly-generated directions on  $S^{p-1}$ .

## **Examples**

```
# A sample on [0, 2*pi)
n <- 5
r_unif_cir(n = n)

# A sample on S^1
p <- 2
samp <- r_unif_sph(n = n, p = p)
samp
rowSums(samp^2)

# A sample on S^2
p <- 3
samp <- r_unif_sph(n = n, p = p)
samp
rowSums(samp^2)</pre>
```

Sobolev

Asymptotic distributions of Sobolev statistics of spherical uniformity

## Description

Approximated density, distribution, and quantile functions for the asymptotic null distributions of Sobolev statistics of uniformity on  $S^{p-1} := \{ \mathbf{x} \in R^p : ||\mathbf{x}|| = 1 \}$ . These asymptotic distributions are infinite weighted sums of (central) chi squared random variables:

$$\sum_{k=1}^{\infty} v_k^2 \chi_{d_{p,k}}^2,$$

where

$$d_{p,k} := \binom{p+k-3}{p-2} + \binom{p+k-2}{p-2}$$

is the dimension of the space of eigenfunctions of the Laplacian on  $S^{p-1}$ ,  $p \ge 2$ , associated to the k-th eigenvalue,  $k \ge 1$ .

## Usage

```
d_p_k(p, k, log = FALSE)
weights_dfs_Sobolev(p, K_max = 1000, thre = 0.001, type, Rothman_t = 1/3,
    Pycke_q = 0.5, Riesz_s = 1, Poisson_rho = 0.5, Softmax_kappa = 1,
    Stereo_a = 0, Sobolev_vk2 = c(0, 0, 1), log = FALSE, verbose = TRUE,
    Gauss = TRUE, N = 320, tol = 1e-06, force_positive = TRUE,
    x_tail = NULL)

d_Sobolev(x, p, type, method = c("I", "SW", "HBE")[1], K_max = 1000,
    thre = 0.001, Rothman_t = 1/3, Pycke_q = 0.5, Riesz_s = 1,
```

```
Poisson_rho = 0.5, Softmax_kappa = 1, Stereo_a = 0,
Sobolev_vk2 = c(0, 0, 1), ncps = 0, verbose = TRUE, N = 320,
x_tail = NULL, ...)

p_Sobolev(x, p, type, method = c("I", "SW", "HBE", "MC")[1], K_max = 1000,
thre = 0.001, Rothman_t = 1/3, Pycke_q = 0.5, Riesz_s = 1,
Poisson_rho = 0.5, Softmax_kappa = 1, Stereo_a = 0,
Sobolev_vk2 = c(0, 0, 1), ncps = 0, verbose = TRUE, N = 320,
x_tail = NULL, ...)

q_Sobolev(u, p, type, method = c("I", "SW", "HBE", "MC")[1], K_max = 1000,
thre = 0.001, Rothman_t = 1/3, Pycke_q = 0.5, Riesz_s = 1,
Poisson_rho = 0.5, Softmax_kappa = 1, Stereo_a = 0,
Sobolev_vk2 = c(0, 0, 1), ncps = 0, verbose = TRUE, N = 320,
x_tail = NULL, ...)
```

#### **Arguments**

p integer giving the dimension of the ambient space  $R^p$  that contains  $S^{p-1}$ .

k sequence of integer indexes.

log compute the logarithm of  $d_{p,k}$ ? Defaults to FALSE.

K\_max integer giving the truncation of the series that compute the asymptotic p-value

of a Sobolev test. Defaults to 1e3.

thre error threshold for the tail probability given by the the first terms of the truncated

series of a Sobolev test. Defaults to 1e-3.

type name of the Sobolev statistic, using the naming from avail\_cir\_tests and

avail\_sph\_tests.

Rothman\_t t parameter for the Rothman test, a real in (0,1). Defaults to 1 / 3. Pycke\_q q parameter for the Pycke "q-test", a real in (0,1). Defaults to 1 / 2.

Riesz\_s s parameter for the s-Riesz test, a real in (0,2). Defaults to 1.

Poisson\_rho  $\rho$  parameter for the Poisson test, a real in [0,1). Defaults to 0.5. Softmax\_kappa  $\kappa$  parameter for the Softmax test, a non-negative real. Defaults to 1.

Stereo\_a a parameter for the Stereo test, a real in [-1, 1]. Defaults to 0.

Sobolev\_vk2 weights for the finite Sobolev test. A non-negative vector or matrix. Defaults to

c(0, 0, 1).

verbose output information about the truncation? Defaults to TRUE.

Gauss use a Gauss-Legendre quadrature rule of N nodes in the computation of the

Gegenbauer coefficients? Otherwise, call integrate. Defaults to TRUE.

N number of points used in the Gauss-Legendre quadrature for computing the

Gegenbauer coefficients. Defaults to 320.

tol tolerance passed to integrate's rel.tol and abs.tol if Gauss = FALSE. De-

faults to 1e-6.

force\_positive set negative weights to zero? Defaults to TRUE.

x_tail	scalar evaluation point for determining the upper tail probability. If NULL, set to the 0.90 quantile of the whole series, computed by the "HBE" approximation.
X	vector of quantiles.
method	method for approximating the density, distribution, or quantile function of the weighted sum of chi squared random variables. Must be "I" (Imhof), "SW" (Satterthwaite–Welch), "HBE" (Hall–Buckley–Eagleson), or "MC" (Monte Carlo; only for distribution or quantile functions). Defaults to "I".
ncps	non-centrality parameters. Either 0 (default) or a vector with the same length as weights.
	further parameters passed to *_wschisq.
u	vector of probabilities.

#### **Details**

The truncation of  $\sum_{k=1}^{\infty} v_k^2 \chi_{d_p,k}^2$  is done to the first K\_max terms and then up to the index such that the first terms explain the tail probability at the x\_tail with an absolute error smaller than thre (see details in cutoff\_wschisq). This automatic truncation takes place when calling \*\_Sobolev. Setting thre = 0 truncates to K\_max terms exactly. If the series only contains odd or even non-zero terms, then only K\_max / 2 addends are *effectively* taken into account in the first truncation.

#### Value

- d\_p\_k: a vector of size length(k) with the evaluation of  $d_{p,k}$ .
- weights\_dfs\_Sobolev: a list with entries weights and dfs, automatically truncated according to K\_max and thre (see details).
- d\_Sobolev: density function evaluated at x, a vector.
- p\_Sobolev: distribution function evaluated at x, a vector.
- q\_Sobolev: quantile function evaluated at u, a vector.

#### Author(s)

Eduardo García-Portugués and Paula Navarro-Esteban.

```
col <- viridisLite::viridis(pmax - 1)</pre>
 curve(p\_Sobolev(x = x, p = 2, type = type, method = "MC", M = M,
                  ...), to = to, n = 2e^2, col = col[pmax - 1],
                  ylab = "Distribution", main = type, ylim = c(0, 1))
 for (p in 3:pmax) {
   curve(p\_Sobolev(x = x, p = p, type = type, method = "MC", M = M,
                    ...), add = TRUE, n = 2e^2, col = col[pmax - p + 1])
 legend("bottomright", legend = paste("p =", 2:pmax), col = rev(col),
         1wd = 2)
}
# Ajne
test_statistic(type = "Ajne")
# Gine_Gn
test_statistic(type = "Gine_Gn", to = 1.5)
# Gine_Fn
test_statistic(type = "Gine_Fn", to = 2)
# Bakshaev
test_statistic(type = "Bakshaev", to = 3)
# Riesz
test_statistic(type = "Riesz", Riesz_s = 0.5, to = 3)
# PCvM
test\_statistic(type = "PCvM", to = 0.6)
# PAD
test_statistic(type = "PAD", to = 3)
test_statistic(type = "PRt", Rothman_t = 0.5)
# Quantiles
p < -c(2, 3, 4, 11)
t(sapply(p, function(p) q\_Sobolev(u = c(0.10, 0.05, 0.01), p = p,
                                  type = "PCvM")))
t(sapply(p, function(p) q\_Sobolev(u = c(0.10, 0.05, 0.01), p = p,
                                  type = "PAD")))
t(sapply(p, function(p) q\_Sobolev(u = c(0.10, 0.05, 0.01), p = p,
                                  type = "PRt")))
# Series truncation for thre = 1e-5
sapply(p, function(p) length(weights_dfs_Sobolev(p = p, type = "PCvM")$dfs))
sapply(p, function(p) length(weights_dfs_Sobolev(p = p, type = "PRt")$dfs))
sapply(p, function(p) length(weights_dfs_Sobolev(p = p, type = "PAD")$dfs))
```

Sobolev\_coefs 61

Sobolev\_coefs

Transformation between different coefficients in Sobolev statistics

## **Description**

Given a Sobolev statistic

$$S_{n,p} = \sum_{i,j=1}^{n} \psi(\cos^{-1}(\mathbf{X}_{i}'\mathbf{X}_{j})),$$

for a sample  $\mathbf{X}_1, \dots, \mathbf{X}_n \in S^{p-1} := \{\mathbf{x} \in R^p : ||\mathbf{x}|| = 1\}, p \geq 2$ , three important sequences are related to  $S_{n,p}$ .

• Gegenbauer coefficients  $\{b_{k,p}\}$  of  $\psi_p$  (see, e.g., the projected-ecdf statistics), given by

$$b_{k,p} := \frac{1}{c_{k,p}} \int_0^{\pi} \psi_p(\theta) C_k^{p/2-1}(\cos \theta) d\theta.$$

• Weights  $\{v_{k,p}^2\}$  of the asymptotic distribution of the Sobolev statistic,  $\sum_{k=1}^\infty v_k^2 \chi_{d_{p,k}}^2$ , given by

$$v_{k,p}^2 = \left(1 + \frac{2k}{p-2}\right)^{-1} b_{k,p}, \quad p \ge 3.$$

• Gegenbauer coefficients  $\{u_{k,p}\}$  of the local projected alternative associated to  $S_{n,p}$ , given by

$$u_{k,p} = \left(1 + \frac{2k}{p-2}\right)v_{k,p}, \quad p \ge 3.$$

For p = 2, the factor (1 + 2k/(p - 2)) is replaced by 2.

## Usage

### Arguments

bk coefficients  $b_{k,p}$  associated to the indexes 1:length(bk), a vector.

p integer giving the dimension of the ambient space  $R^p$  that contains  $S^{p-1}$ .

log	do operations in log scale (log-in, log-out)? Defaults to FALSE.
signs	signs of the coefficients $u_{k,p}$ , a vector of the same size as vk2 or bk, or a scalar. Defaults to 1.
vk2	<b>squared</b> coefficients $v_{k,p}^2$ associated to the indexes 1:length(vk2), a vector.
uk	coefficients $u_{k,p}$ associated to the indexes 1:length(uk), a vector.

#### **Details**

See more details in Prentice (1978) and García-Portugués et al. (2023). The adequate signs of uk for the "PRt" Rothman test can be retrieved with akx and sqr = TRUE, see the examples.

#### Value

The corresponding vectors of coefficients vk2, bk, or uk, depending on the call.

#### References

García-Portugués, E., Navarro-Esteban, P., Cuesta-Albertos, J. A. (2023) On a projection-based class of uniformity tests on the hypersphere. *Bernoulli*, 29(1):181–204. doi:10.3150/21BEJ1454.

Prentice, M. J. (1978). On invariant tests of uniformity for directions and orientations. *The Annals of Statistics*, 6(1):169–176. doi:10.1214/aos/1176344075

## **Examples**

```
# bk, vk2, and uk for the PCvM test in p = 3
(bk <- Gegen_coefs_Pn(k = 1:5, type = "PCvM", p = 3))
(vk2 <- bk_to_vk2(bk = bk, p = 3))
(uk <- bk_to_uk(bk = bk, p = 3))

# vk2 is the same as
weights_dfs_Sobolev(K_max = 10, thre = 0, p = 3, type = "PCvM")$weights

# bk and uk for the Rothman test in p = 3, with adequate signs
t <- 1 / 3
(bk <- Gegen_coefs_Pn(k = 1:5, type = "PRt", p = 3, Rothman_t = t))
(ak <- akx(x = drop(q_proj_unif(t, p = 3)), p = 3, k = 1:5, sqr = TRUE))
(uk <- bk_to_uk(bk = bk, p = 3, signs = ak))</pre>
```

sph\_stat\_Rayleigh

Statistics for testing (hyper)spherical uniformity

## Description

Low-level implementation of several statistics for assessing uniformity on the (hyper)sphere  $S^{p-1} := \{ \mathbf{x} \in \mathbb{R}^p : ||\mathbf{x}|| = 1 \}, p \geq 2.$ 

## Usage

```
sph_stat_Rayleigh(X)
sph_stat_Bingham(X)
sph_stat_Ajne(X, Psi_in_X = FALSE)
sph_stat_Gine_Gn(X, Psi_in_X = FALSE, p = 0L)
sph_stat_Gine_Fn(X, Psi_in_X = FALSE, p = 0L)
sph_stat_Pycke(X, Psi_in_X = FALSE, p = 0L)
sph_stat_Bakshaev(X, Psi_in_X = FALSE, p = 0L)
sph_stat_Riesz(X, Psi_in_X = FALSE, p = 0L, s = 1)
sph_stat_PCvM(X, Psi_in_X = FALSE, p = 0L, N = 160L, L = 1000L)
sph_stat_PRt(X, Psi_in_X = FALSE, p = 0L, t = 1/3, N = 160L,
 L = 1000L)
sph_stat_PAD(X, Psi_in_X = FALSE, p = 0L, N = 160L, L = 1000L)
sph_stat_Poisson(X, Psi_in_X = FALSE, p = 0L, rho = 0.5)
sph_stat_Softmax(X, Psi_in_X = FALSE, p = 0L, kappa = 1)
sph_stat_Stereo(X, Psi_in_X = FALSE, p = 0L, a = 0)
sph_stat_CCF09(X, dirs, K_CCF09 = 25L, original = FALSE)
sph_stat_Rayleigh_HD(X)
sph_stat_CJ12(X, Psi_in_X = FALSE, p = 0L, regime = 3L)
```

## **Arguments**

X	an <b>array</b> of size $c(n, p, M)$ containing the Cartesian coordinates of M samples of size n of directions on $S^{p-1}$ . Must not contain NA's.
Psi_in_X	does X contain the shortest angles matrix $\Psi$ that is obtained with Psi_mat(X)? If FALSE (default), $\Psi$ is computed internally.
р	integer giving the dimension of the ambient space $\mathbb{R}^p$ that contains $\mathbb{S}^{p-1}$ .
S	s parameter for the $s$ -Riesz test, a real in $(0,2)$ . Defaults to 1.
N	number of points used in the Gauss-Legendre quadrature. Defaults to 160.
L	number of discretization points to interpolate angular functions that require evaluating an integral. Defaults to 1e3.

t	t parameter for the Rothman and Cressie tests, a real in $(0,1).$ Defaults to 1 / 3.
rho	$\rho$ parameter for the Poisson test, a real in $[0,1)$ . Defaults to 0.5.
kappa	$\kappa$ parameter for the Softmax test, a non-negative real. Defaults to 1.
a	either:
	• $a_n=a/n$ parameter used in the length of the arcs of the coverage-based tests. Must be positive. Defaults to 2 * pi.
	• a parameter for the Stereo test, a real in $[-1,1]$ . Defaults to 0.
dirs	a matrix of size $c(n_proj, p)$ containing $n_proj$ random directions (in Cartesian coordinates) on $S^{p-1}$ to perform the CCF09 test.
K_CCF09	integer giving the truncation of the series present in the asymptotic distribution of the Kolmogorov-Smirnov statistic. Defaults to 25.
original	return the CCF09 statistic as originally defined? If FALSE (default), a faster and equivalent statistic is computed, and rejection happens for <i>large</i> values of the statistic, which is consistent with the rest of tests. Otherwise, rejection happens for <i>low</i> values.
regime	type of asymptotic regime for the CJ12 test, either 1 (sub-exponential regime), 2 (exponential), or 3 (super-exponential; default).

#### **Details**

Detailed descriptions and references of the statistics are available in García-Portugués and Verdebout (2018).

The Pycke and CJ12 statistics employ the *scalar products* matrix, rather than the shortest angles matrix, when Psi\_in\_X = TRUE. This matrix is obtained by setting scalar\_prod = TRUE in Psi\_mat.

## Value

A matrix of size c(M, 1) containing the statistics for each of the M samples.

## Warning

Be careful on avoiding the next bad usages of the functions, which will produce spurious results:

- The directions in X do *not* have unit norm.
- X does *not* contain Psi\_mat(X) when X\_in\_Theta = TRUE.
- The parameter p does *not* match with the dimension of  $\mathbb{R}^p$ .
- *Not* passing the scalar products matrix to sph\_stat\_CJ12 when Psi\_in\_X = TRUE.
- The directions in dirs do *not* have unit norm.

### References

García-Portugués, E. and Verdebout, T. (2018) An overview of uniformity tests on the hypersphere. *arXiv:1804.00286*. doi:10.48550/arXiv.1804.00286.

```
## Sample uniform spherical data
M < -2
n <- 100
p <- 3
set.seed(123456789)
X \leftarrow r_unif_sph(n = n, p = p, M = M)
## Sobolev tests
# Rayleigh
sph_stat_Rayleigh(X)
# Bingham
sph_stat_Bingham(X)
# Ajne
Psi <- Psi_mat(X)
dim(Psi) <- c(dim(Psi), 1)</pre>
sph_stat_Ajne(X)
sph_stat_Ajne(Psi, Psi_in_X = TRUE)
# Gine Gn
sph_stat_Gine_Gn(X)
sph_stat_Gine_Gn(Psi, Psi_in_X = TRUE, p = p)
# Gine Fn
sph_stat_Gine_Fn(X)
sph_stat_Gine_Fn(Psi, Psi_in_X = TRUE, p = p)
# Pycke
sph_stat_Pycke(X)
sph_stat_Pycke(Psi, Psi_in_X = TRUE, p = p)
# Bakshaev
sph_stat_Bakshaev(X)
sph_stat_Bakshaev(Psi, Psi_in_X = TRUE, p = p)
# Riesz
sph_stat_Riesz(X, s = 1)
sph_stat_Riesz(Psi, Psi_in_X = TRUE, p = p, s = 1)
# Projected Cramér-von Mises
sph_stat_PCvM(X)
sph_stat_PCvM(Psi, Psi_in_X = TRUE, p = p)
# Projected Rothman
sph_stat_PRt(X)
sph_stat_PRt(Psi, Psi_in_X = TRUE, p = p)
# Projected Anderson-Darling
```

sph\_stat\_Sobolev

```
sph_stat_PAD(X)
sph_stat_PAD(Psi, Psi_in_X = TRUE, p = p)
## Other tests
# CCF09
dirs <- r_unif_sph(n = 3, p = p, M = 1)[, , 1]
sph_stat_CCF09(X, dirs = dirs)
## High-dimensional tests
# Rayleigh HD-Standardized
sph_stat_Rayleigh_HD(X)
# CJ12
sph_stat_CJ12(X, regime = 1)
sph_stat_CJ12(Psi, regime = 1, Psi_in_X = TRUE, p = p)
sph_stat_CJ12(X, regime = 2)
sph_stat_CJ12(Psi, regime = 2, Psi_in_X = TRUE, p = p)
sph_stat_CJ12(X, regime = 3)
sph_stat_CJ12(Psi, regime = 3, Psi_in_X = TRUE, p = p)
```

sph\_stat\_Sobolev

Finite Sobolev statistics for testing (hyper)spherical uniformity

## **Description**

Computes the finite Sobolev statistic

$$S_{n,p}(\{b_{k,p}\}_{k=1}^K) = \sum_{i,j=1}^n \sum_{k=1}^K b_{k,p} C_k^{(j)} p/2 - 1)(\cos^{-1}(\mathbf{X}_i'\mathbf{X}_j)),$$

for a sequence  $\{b_{k,p}\}_{k=1}^K$  of non-negative weights. For p=2, the Gegenbauer polynomials are replaced by Chebyshev ones.

## Usage

```
sph_stat_Sobolev(X, Psi_in_X = FALSE, p = 0, vk2 = c(0, 0, 1)) cir_stat_Sobolev(Theta, Psi_in_Theta = FALSE, vk2 = c(0, 0, 1))
```

## **Arguments**

X	an <b>array</b> of size $c(n, p, M)$ containing the Cartesian coordinates of M samples of size n of directions on $S^{p-1}$ . Must not contain NA's.
Psi_in_X	does X contain the shortest angles matrix $\Psi$ that is obtained with Psi_mat(X)? If FALSE (default), $\Psi$ is computed internally.
р	integer giving the dimension of the ambient space $\mathbb{R}^p$ that contains $\mathbb{S}^{p-1}$ .

unif\_cap 67

vk2	weights for the finite Sobolev test. A non-negative vector or matrix. Defaults to $c(\emptyset,\emptyset,1)$ .
Theta	a <b>matrix</b> of size c(n, M) with M samples of size n of circular data on $[0,2\pi)$ . Must not contain NA's.
Psi_in_Theta	does Theta contain the shortest angles matrix $\Psi$ that is obtained with Psi_mat(array(Theta, dim = c(n, 1, M)))? If FALSE (default), $\Psi$ is computed internally.

#### Value

A matrix of size c(M, ncol(vk2)) containing the statistics for each of the M samples.

unif\_cap

Uniform spherical cap distribution

## Description

Density, simulation, and associated functions for a uniform distribution within a spherical cap of angle  $0 \le \alpha \le \pi$  about a direction  $\mu$  on  $S^{p-1} := \{ \mathbf{x} \in R^p : ||\mathbf{x}|| = 1 \}, p \ge 2$ . The density at  $\mathbf{x} \in S^{p-1}$  is given by

$$c_{p,r} 1_{[1-r,1]}(\mathbf{x}'\boldsymbol{\mu})$$
 with  $c_{p,r} := \omega_p [1 - F_p(1-r)],$ 

where  $r = \cos(\alpha)$  is the projected radius of the spherical cap about  $\mu$ ,  $\omega_p$  is the surface area of  $S^{p-1}$ , and  $F_p$  is the projected uniform distribution (see p\_proj\_unif).

The angular function of the uniform cap distribution is  $g(t):=1_{[1-r,1]}(t)$ . The associated projected density is  $\tilde{g}(t):=\omega_{p-1}c_{p,r}(1-t^2)^{(p-3)/2}1_{[1-r,1]}(t)$ .

## Usage

```
d_unif_cap(x, mu, angle = pi/10)

c_unif_cap(p, angle = pi/10)

r_unif_cap(n, mu, angle = pi/10)

p_proj_unif_cap(x, p, angle = pi/10)

q_proj_unif_cap(u, p, angle = pi/10)

d_proj_unif_cap(x, p, angle = pi/10, scaled = TRUE)

r_proj_unif_cap(n, p, angle = pi/10)
```

68 unif\_cap

## **Arguments**

X	locations to evaluate the density or distribution. For d_unif_cap, positions on $S^{p-1}$ given either as a matrix of size c(nx, p) or a vector of length p. Normalized internally if required (with a warning message). For d_unif_proj_cap and p_unif_proj_cap, a vector with values in $[-1,1]$ .
mu	the directional mean $\mu$ of the distribution. A unit-norm vector of length p.
angle	angle $\alpha$ defining the spherical cap about $\mu$ . A scalar in $[0, \pi]$ . Defaults to $\pi/10$ .
р	integer giving the dimension of the ambient space $\mathbb{R}^p$ that contains $\mathbb{S}^{p-1}$ .
n	sample size, a positive integer.
u	vector of probabilities.
scaled	whether to scale the angular function by the normalizing constant. Defaults to TRUE.

#### Value

Depending on the function:

- d\_unif\_cap: a vector of length nx or 1 with the evaluated density at x.
- r\_unif\_cap: a matrix of size c(n, p) with the random sample.
- c\_unif\_cap: the normalizing constant.
- p\_proj\_unif\_cap: a vector of length x with the evaluated distribution function at x.
- q\_proj\_unif\_cap: a vector of length u with the evaluated quantile function at u.
- d\_proj\_unif\_cap: a vector of size nx with the evaluated angular function.
- r\_proj\_unif\_cap: a vector of length n containing simulated values from the cosines density associated to the angular function.

#### Author(s)

Alberto Fernández-de-Marcos and Eduardo García-Portugués.

```
# Simulation and density evaluation for p = 2 mu <- c(0, 1) angle <- pi / 5   
n <- 1e2   
x <- r_unif_cap(n = n, mu = mu, angle = angle)  
col <- viridisLite::viridis(n)  
r_noise <- runif(n, 0.95, 1.05) # Perturbation to improve visualization  
color <- col[rank(d_unif_cap(x = x, mu = mu, angle = angle))]  
plot(r_noise * x, pch = 16, col = color, xlim = c(-1, 1), ylim = c(-1, 1))  
# Simulation and density evaluation for p = 3  
mu <- c(0, 0, 1)  
angle <- pi / 5  
x <- r_unif_cap(n = n, mu = mu, angle = angle)
```

unif\_stat 69

unif\_stat

Circular and (hyper)spherical uniformity statistics

### Description

Implementation of several statistics for assessing uniformity on the (hyper)sphere  $S^{p-1} := \{ \mathbf{x} \in \mathbb{R}^p : ||\mathbf{x}|| = 1 \}, p \geq 2$ , for a sample  $\mathbf{X}_1, \dots, \mathbf{X}_n \in S^{p-1}$ .

unif\_stat receives a (several) sample(s) of directions in *Cartesian coordinates*, except for the circular case (p=2) in which the sample(s) can be  $angles\ \Theta_1,\ldots,\Theta_n\in[0,2\pi)$ .

unif\_stat allows to compute several statistics to several samples within a single call, facilitating thus Monte Carlo experiments.

## Usage

```
unif_stat(data, type = "all", data_sorted = FALSE, CCF09_dirs = NULL,
   CJ12_reg = 3, cov_a = 2 * pi, Cressie_t = 1/3, K_CCF09 = 25,
   Poisson_rho = 0.5, Pycke_q = 0.5, Rayleigh_m = 1, Riesz_s = 1,
   Rothman_t = 1/3, Sobolev_vk2 = c(0, 0, 1), Softmax_kappa = 1,
   Stereo_a = 0)
```

## **Arguments**

data

sample to compute the test statistic. An **array** of size c(n, p, M) containing M samples of size n of directions (in Cartesian coordinates) on  $S^{p-1}$ . Alternatively, a **matrix** of size c(n, M) with the angles on  $[0, 2\pi)$  of the M circular samples of size n on  $S^1$ . Other objects accepted are an array of size c(n, 1, M) or a vector of size n with angular data. Must not contain NA's.

type

type of test to be applied. A character vector containing any of the following types of tests, depending on the dimension p:

- Circular data: any of the names available at object avail\_cir\_tests.
- (Hyper)spherical data: any of the names available at object avail\_sph\_tests.

If type = "all" (default), then type is set as avail\_cir\_tests or avail\_sph\_tests, depending on the value of p.

70 unif\_stat

data_sorted	is the circular data sorted? If TRUE, certain statistics are faster to compute. Defaults to FALSE.
CCF09_dirs	a matrix of size c(n_proj, p) containing n_proj random directions (in Cartesian coordinates) on $S^{p-1}$ to perform the CCF09 test. If NULL (default), a sample of size n_proj = 50 directions is computed internally.
CJ12_reg	type of asymptotic regime for CJ12 test, either 1 (sub-exponential regime), 2 (exponential), or 3 (super-exponential; default).
cov_a	$a_n=a/n$ parameter used in the length of the arcs of the coverage-based tests. Must be positive. Defaults to 2 * pi.
Cressie_t	$t$ parameter for the Cressie test, a real in $(0,1).$ Defaults to 1 $\slash$ 3.
K_CCF09	integer giving the truncation of the series present in the asymptotic distribution of the Kolmogorov-Smirnov statistic. Defaults to 25.
Poisson_rho	$\rho$ parameter for the Poisson test, a real in $[0,1)$ . Defaults to 0.5.
Pycke_q	q parameter for the Pycke " $q-test$ ", a real in $(0,1).$ Defaults to 1 / 2.
Rayleigh_m	integer $m$ for the $m$ -modal Rayleigh test. Defaults to ${\bf m}$ = 1 (the standard Rayleigh test).
Riesz_s	s parameter for the $s$ -Riesz test, a real in $(0,2)$ . Defaults to 1.
Rothman_t	t parameter for the Rothman test, a real in $(0,1)$ . Defaults to 1 / 3.
Sobolev_vk2	weights for the finite Sobolev test. A non-negative vector or matrix. Defaults to $c(\emptyset,\emptyset,1)$ .
Softmax_kappa	$\kappa$ parameter for the Softmax test, a non-negative real. Defaults to 1.
Stereo_a	$a$ parameter for the Stereo test, a real in $[-1,1]$ . Defaults to $\emptyset$ .

## **Details**

Except for CCF09\_dirs, K\_CCF09, and CJ12\_reg, all the test-specific parameters are vectorized. Descriptions and references for most of the statistics are available in García-Portugués and Verdebout (2018).

## Value

A data frame of size c(M, length(type)), with column names given by type, that contains the values of the test statistics.

## References

García-Portugués, E. and Verdebout, T. (2018) An overview of uniformity tests on the hypersphere. *arXiv:1804.00286*. doi:10.48550/arXiv.1804.00286.

```
## Circular data
# Sample
n <- 10</pre>
```

unif\_stat\_distr 71

```
M <- 2
Theta <- r_unif_cir(n = n, M = M)
# Matrix
unif_stat(data = Theta, type = "all")
unif_stat(data = array(Theta, dim = c(n, 1, M)), type = "all")
unif_stat(data = Theta[, 1], type = "all")
## Spherical data
# Circular sample in Cartesian coordinates
n <- 10
M <- 2
X \leftarrow array(dim = c(n, 2, M))
for (i in 1:M) X[, , i] \leftarrow cbind(cos(Theta[, i]), sin(Theta[, i]))
# Array
unif_stat(data = X, type = "all")
# High-dimensional data
X \leftarrow r_unif_sph(n = n, p = 3, M = M)
unif_stat(data = X, type = "all")
## Specific arguments
# Rothman
unif_stat(data = Theta, type = "Rothman", Rothman_t = 0.5)
# CCF09
unif_stat(data = X, type = "CCF09", CCF09_dirs = X[, , 1])
unif_stat(data = X, type = "CCF09", CCF09_dirs = X[, , 1], K_CCF09 = 1)
# CJ12
unif_stat(data = X, type = "CJ12", CJ12_reg = 3)
unif_stat(data = X, type = "CJ12", CJ12_reg = 1)
```

unif\_stat\_distr

Null distributions for circular and (hyper)spherical uniformity statistics

## **Description**

Approximate computation of the null distributions of several statistics for assessing uniformity on the (hyper)sphere  $S^{p-1} := \{ \mathbf{x} \in R^p : ||\mathbf{x}|| = 1 \}, p \geq 2$ . The approximation is done either by means of the asymptotic distribution or by Monte Carlo.

72 unif\_stat\_distr

## Usage

```
unif_stat_distr(x, type, p, n, approx = "asymp", M = 10000,
          stats_MC = NULL, K_max = 10000, method = "I", Stephens = FALSE,
         CCF09_dirs = NULL, CJ12_beta = 0, CJ12_reg = 3, cov_a = 2 * pi,
         Cressie_t = 1/3, K_Ajne = 500, K_CCF09 = 25, K_Kuiper = 25,
         K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} = 25, K_{watson} 
         Rayleigh_m = 1, Riesz_s = 1, Rothman_t = 1/3, Sobolev_vk2 = c(0, 0, 0)
          1), Softmax_kappa = 1, Stereo_a = 0, ...)
```

### **Arguments**

Х

evaluation points for the null distribution(s). Either a vector of size nx, if the evaluation points are common for the tests in type, or a matrix of size c(nx, length(type)) with columns containing the evaluation points for each test. Must not contain NA's.

type

type of test to be applied. A character vector containing any of the following types of tests, depending on the dimension p:

- Circular data: any of the names available at object avail\_cir\_tests.
- (Hyper)spherical data: any of the names available at object avail\_sph\_tests.

If type = "all" (default), then type is set as avail\_cir\_tests or avail\_sph\_tests, depending on the value of p.

integer giving the dimension of the ambient space  $\mathbb{R}^p$  that contains  $\mathbb{S}^{p-1}$ .

n

sample size employed for computing the statistic.

approx

type of approximation to the null distribution, either "asymp" (default) for employing the asymptotic null distribution, if available, or "MC", for employing the Monte Carlo approximation of the exact null distribution.

Μ

number of Monte Carlo replications for approximating the null distribution when approx = "MC". Also, number of Monte Carlo samples for approximating the asymptotic distributions based on weighted sums of chi squared random variables. Defaults to 1e4.

stats\_MC

a data frame of size c(M, length(type)), with column names containing the character vector type, that results from extracting \$stats\_MC from a call to unif\_stat\_MC. If provided, the computation of Monte Carlo statistics when approx = "MC" is skipped. stats\_MC is checked internally to see if it is sorted. Internally computed if NULL (default).

K\_max

integer giving the truncation of the series that compute the asymptotic p-value of a Sobolev test. Defaults to 1e4.

method

method for approximating the density, distribution, or quantile function of the weighted sum of chi squared random variables. Must be "I" (Imhof), "SW" (Satterthwaite–Welch), "HBE" (Hall–Buckley–Eagleson), or "MC" (Monte Carlo; only for distribution or quantile functions). Defaults to "I".

Stephens

compute Stephens (1970) modification so that the null distribution of the is less dependent on the sample size? The modification does not alter the test decision.

unif\_stat\_distr 73

CCF09_dirs	a matrix of size $c(n\_proj, p)$ containing $n\_proj$ random directions (in Cartesian coordinates) on $S^{p-1}$ to perform the CCF09 test. If NULL (default), a sample of size $n\_proj = 50$ directions is computed internally.			
CJ12_beta	$\beta$ parameter in the exponential regime of CJ12 test, a positive real.			
CJ12_reg	type of asymptotic regime for CJ12 test, either 1 (sub-exponential regime), 2 (exponential), or 3 (super-exponential; default).			
cov_a	$a_n=a/n$ parameter used in the length of the arcs of the coverage-based tests. Must be positive. Defaults to 2 * pi.			
Cressie_t	t parameter for the Cressie test, a real in $(0,1)$ . Defaults to 1 / 3.			
K_CCF09	integer giving the truncation of the series present in the asymptotic distribution of the Kolmogorov-Smirnov statistic. Defaults to 25.			
K_Kuiper, K_Watson, K_Watson_1976, K_Ajne				
	integer giving the truncation of the series present in the null asymptotic distributions. For the Kolmogorov-Smirnov-related series defaults to 25.			
Poisson_rho	$\rho$ parameter for the Poisson test, a real in $[0,1)$ . Defaults to 0.5.			
Pycke_q	q parameter for the Pycke " $q-test$ ", a real in $(0,1).$ Defaults to 1 / 2.			
Rayleigh_m	integer $m$ for the $m$ -modal Rayleigh test. Defaults to ${\bf m}$ = 1 (the standard Rayleigh test).			
Riesz_s	s parameter for the $s$ -Riesz test, a real in $(0,2)$ . Defaults to 1.			
Rothman_t	t parameter for the Rothman test, a real in $(0,1)$ . Defaults to 1 / 3.			
Sobolev_vk2	weights for the finite Sobolev test. A non-negative vector or matrix. Defaults to $c(0,0,1)$ .			
Softmax_kappa	$\kappa$ parameter for the Softmax test, a non-negative real. Defaults to 1.			
Stereo_a	$a$ parameter for the Stereo test, a real in $[-1,1]$ . Defaults to $\emptyset$ .			
• • •	if approx = "MC", optional performance parameters to be passed to unif_stat_MC: chunks, cores, and seed.			

# **Details**

When approx = "asymp", statistics that do not have an implemented or known asymptotic are omitted, and a warning is generated.

For Sobolev tests, K\_max = 1e4 produces probabilities uniformly accurate with three digits for the "PCvM", "PAD", and "PRt" tests, for dimensions  $p \leq 11$ . With K\_max = 5e4, these probabilities are uniformly accurate in the fourth digit. With K\_max = 1e3, only two-digit uniform accuracy is obtained. Uniform accuracy deteriorates when p increases, e.g., a digit accuracy is lost when p=51.

Descriptions and references on most of the asymptotic distributions are available in García-Portugués and Verdebout (2018).

# Value

A data frame of size c(nx, length(type)), with column names given by type, that contains the values of the null distributions of the statistics evaluated at x.

74 unif\_stat\_distr

### References

García-Portugués, E. and Verdebout, T. (2018) An overview of uniformity tests on the hypersphere. *arXiv*:1804.00286. doi:10.48550/arXiv.1804.00286.

```
## Asymptotic distribution
# Circular statistics
x < - seq(0, 1, 1 = 5)
unif_stat_distr(x = x, type = "Kuiper", p = 2, n = 10)
unif_stat_distr(x = x, type = c("Ajne", "Kuiper"), p = 2, n = 10)
unif_stat_distr(x = x, type = c("Ajne", "Kuiper"), p = 2, n = 10, K_Ajne = 5)
# All circular statistics
unif_stat_distr(x = x, type = avail_cir_tests, p = 2, n = 10, K_max = 1e3)
# Spherical statistics
unif_stat_distr(x = cbind(x, x + 1), type = c("Rayleigh", "Bingham"),
               p = 3, n = 10
unif_stat_distr(x = cbind(x, x + 1), type = c("Rayleigh", "Bingham"),
               p = 3, n = 10, M = 100)
# All spherical statistics
unif_stat_distr(x = x, type = avail_sph_tests, p = 3, n = 10, K_max = 1e3)
## Monte Carlo distribution
# Circular statistics
x < - seg(0, 5, 1 = 10)
unif_stat_distr(x = x, type = avail_cir_tests, p = 2, n = 10, approx = "MC")
approx = "MC")
# Spherical statistics
unif_stat_distr(x = x, type = avail_sph_tests, p = 3, n = 10,
               approx = "MC")
unif_stat_distr(x = cbind(x, x + 1), type = c("Rayleigh", "Bingham"),
               p = 3, n = 10, approx = "MC")
unif_stat_distr(x = cbind(x, x + 1), type = c("Rayleigh", "Bingham"),
               p = 3, n = 10, approx = "MC")
## Specific arguments
# Rothman
unif_stat_distr(x = x, type = "Rothman", p = 2, n = 10, Rothman_t = 0.5,
               approx = "MC")
# CCF09
dirs <- r_unif_sph(n = 5, p = 3, M = 1)[, , 1]
x < - seq(0, 1, 1 = 10)
```

```
unif_stat_distr(x = x, type = "CCF09", p = 3, n = 10, approx = "MC",
                CCF09_dirs = dirs)
unif_stat_distr(x = x, type = "CCF09", p = 3, n = 10, approx = "MC")
# CJ12
unif_stat_distr(x = x, type = "CJ12", p = 3, n = 100, CJ12_reg = 3)
unif_stat_distr(x = x, type = "CJ12", p = 3, n = 100, CJ12_reg = 2,
               CJ12\_beta = 0.01)
unif_stat_distr(x = x, type = "CJ12", p = 3, n = 100, CJ12_reg = 1)
## Sobolev
x < - seq(0, 1, 1 = 10)
vk2 \leftarrow diag(1, nrow = 3)
unif_stat_distr(x = x, type = "Sobolev", approx = "asymp", p = 3, n = 100,
                Sobolev_vk2 = vk2)
sapply(1:3, function(i)
  unif_stat_distr(x = x, type = "Sobolev", approx = "asymp", p = 3, n = 100,
                  Sobolev_vk2 = vk2[i, ])$Sobolev)
sapply(1:3, function(i)
  unif_stat_distr(x = x, type = "Sobolev", approx = "MC", p = 3, n = 100,
                  Sobolev_vk2 = vk2[i, ], M = 1e3)$Sobolev)
unif_stat_distr(x = x, type = "Sobolev", approx = "MC", p = 3, n = 100,
                Sobolev_vk2 = vk2, M = 1e3)
```

unif\_stat\_MC

Monte Carlo simulation of circular and (hyper)spherical uniformity statistics

# **Description**

Utility for performing Monte Carlo simulation of several statistics for assessing uniformity on the (hyper)sphere  $S^{p-1} := \{ \mathbf{x} \in R^p : ||\mathbf{x}|| = 1 \}, p \geq 2.$ 

unif\_stat\_MC provides a convenient wrapper for parallel evaluation of unif\_stat, the estimation of critical values under the null distribution, and the computation of empirical powers under the alternative.

### Usage

```
unif_stat_MC(n, type = "all", p, M = 10000, r_H1 = NULL,
    crit_val = NULL, alpha = c(0.1, 0.05, 0.01), return_stats = TRUE,
    stats_sorted = FALSE, chunks = ceiling((n * M)/1e+05), cores = 1,
    seeds = NULL, CCF09_dirs = NULL, CJ12_reg = 3, cov_a = 2 * pi,
    Cressie_t = 1/3, K_CCF09 = 25, Poisson_rho = 0.5, Pycke_q = 0.5,
    Rayleigh_m = 1, Riesz_s = 1, Rothman_t = 1/3, Sobolev_vk2 = c(0, 0,
    1), Softmax_kappa = 1, Stereo_a = 0, ...)
```

# **Arguments**

р

n sample size.

type type of test to be applied. A character vector containing any of the following types of tests, depending on the dimension p:

• Circular data: any of the names available at object avail\_cir\_tests.

• (Hyper)spherical data: any of the names available at object avail\_sph\_tests.

If type = "all" (default), then type is set as avail\_cir\_tests or avail\_sph\_tests, depending on the value of p.

integer giving the dimension of the ambient space  $\mathbb{R}^p$  that contains  $\mathbb{S}^{p-1}$ .

M number of Monte Carlo replications. Defaults to 1e4.

r\_H1 if provided, the computation of empirical powers is carried out for the alternative hypothesis sampled with r\_H1. This must be a function with the same arguments

and value as r\_unif\_sph (see examples). Defaults to NULL, indicating that the

critical values are estimated from samples of r\_unif\_sph.

crit\_val if provided, must be the critical values as returned by \$stats\_MC in a call to

unif\_stat\_MC. They are used for computing the empirical powers of the tests present in type. Defaults to NULL, which means that no power computation is

done.

alpha vector with significance levels. Defaults to c(0.10, 0.05, 0.01).

return\_stats return the Monte Carlo statistics? If only the critical values or powers are de-

sired, FALSE saves memory in the returned object. Defaults to TRUE.

stats\_sorted sort the returned Monte Carlo statistics? If TRUE, this is useful for evaluating

faster the empirical cumulative distribution function when approximating the

distribution in unif\_stat\_distr. Defaults to FALSE.

chunks number of chunks to split the M Monte Carlo replications. Useful for paral-

lelizing the simulation study in chunks tasks containing ceiling(M / chunks) replications. Useful also for avoiding memory bottlenecks when M is large. De-

faults to

ceiling((n \* M) / 1e5).

cores number of cores to perform the simulation. Defaults to 1.

seeds if provided, a vector of size chunks for fixing the seeds on each of the simula-

tion chunks (useful for reproducing parallel simulations). Specifically, for k in 1:chunks, seeds are set as set.seed(seeds[k], kind = "Mersenne-Twister")

in each chunk. Defaults to NULL (no seed setting is done).

CCF09\_dirs a matrix of size c(n\_proj, p) containing n\_proj random directions (in Carte-

sian coordinates) on  $S^{p-1}$  to perform the CCF09 test. If NULL (default), a sample

of size n\_proj = 50 directions is computed internally.

CJ12\_reg type of asymptotic regime for CJ12 test, either 1 (sub-exponential regime), 2

(exponential), or 3 (super-exponential; default).

cov\_a  $a_n = a/n$  parameter used in the length of the arcs of the coverage-based tests.

Must be positive. Defaults to 2 \* pi.

Cressie\_t t parameter for the Cressie test, a real in (0,1). Defaults to 1 / 3.

K\_CCF09 integer giving the truncation of the series present in the asymptotic distribution of the Kolmogorov-Smirnov statistic. Defaults to 25.  $\rho$  parameter for the Poisson test, a real in [0, 1). Defaults to 0.5. Poisson\_rho q parameter for the Pycke "q-test", a real in (0, 1). Defaults to 1 / 2. Pycke\_q Rayleigh\_m integer m for the m-modal Rayleigh test. Defaults to m = 1 (the standard Rayleigh Riesz\_s s parameter for the s-Riesz test, a real in (0, 2). Defaults to 1. Rothman\_t t parameter for the Rothman test, a real in (0, 1). Defaults to 1 / 3. Sobolev\_vk2 weights for the finite Sobolev test. A non-negative vector or matrix. Defaults to c(0, 0, 1). $\kappa$  parameter for the Softmax test, a non-negative real. Defaults to 1. Softmax\_kappa a parameter for the Stereo test, a real in [-1, 1]. Defaults to 0. Stereo\_a optional arguments to be passed to the r\_H1 sampler or to foreach (for example, .export to export global variables or other functions to the foreach environment).

### **Details**

It is possible to have a progress bar if unif\_stat\_MC is wrapped with progressr::with\_progress or if progressr::handlers(global = TRUE) is invoked (once) by the user. See the examples below. The progress bar is updated with the number of finished chunks.

All the tests reject for large values of the test statistic (max\_gap = TRUE is assumed for the Range test), so the critical values for the significance levels alpha correspond to the alpha-upper quantiles of the null distribution of the test statistic.

The Monte Carlo simulation for the CCF09 test is made conditionally on the choice of CCF09\_dirs. That is, all the Monte Carlo statistics share the same random directions.

Except for CCF09\_dirs, K\_CCF09, and CJ12\_reg, all the test-specific parameters are vectorized.

# Value

A list with the following entries:

- crit\_val\_MC: a data frame of size c(length(alpha), length(type)), with column names given by type and rows corresponding to the significance levels alpha, that contains the estimated critical values of the tests.
- power\_MC: a data frame of size c(nrow(crit\_val), length(type)), with column names given by type and rows corresponding to the significance levels of crit\_val, that contains the empirical powers of the tests. NA if crit\_val = NULL.
- stats\_MC: a data frame of size c(M, length(type)), with column names given by type, that contains the Monte Carlo statistics.

```
## Critical values
# Single statistic, specific alpha
cir \leftarrow unif_stat_MC(n = 10, M = 1e2, type = "Ajne", p = 2, alpha = 0.15)
summary(cir$stats_MC)
cir$crit_val_MC
# All circular statistics
cir \leftarrow unif_stat_MC(n = 10, M = 1e2, p = 2)
head(cir$stats_MC)
cir$crit_val_MC
# All spherical statistics
sph \leftarrow unif_stat_MC(n = 10, M = 1e2, p = 3)
head(sph$stats_MC)
sph$crit_val_MC
## Using a progress bar
# Define a progress bar
require(progress)
require(progressr)
handlers(handler_progress(
  format = paste("(:spin) [:bar] :percent Iter: :current/:total Rate:",
                 ":tick_rate iter/sec ETA: :eta Elapsed: :elapsedfull"),
  clear = FALSE))
# Call unif_stat_MC() within with_progress()
with_progress(unif_stat_MC(n = 10, M = 1e2, p = 3, chunks = 10))
# With several cores
with_progress(unif_stat_MC(n = 10, M = 1e2, p = 3, chunks = 10, cores = 2))
# Instead of using with_progress() each time, it is more practical to run
# handlers(global = TRUE)
# once to activate progress bars in your R session
## Power computation
# Single statistic
cir_pow \leftarrow unif_stat_MC(n = 10, M = 1e2, type = "Ajne", p = 2,
                        crit_val = cir$crit_val_MC)
cir_pow$crit_val_MC
cir_pow$power_MC
# All circular statistics
cir_pow <- unif_stat_MC(n = 10, M = 1e2, p = 2, crit_val = cir$crit_val_MC)</pre>
cir_pow$crit_val_MC
cir_pow$power_MC
# All spherical statistics
```

```
sph_pow <- unif_stat_MC(n = 10, M = 1e2, p = 3, crit_val = sph$crit_val_MC)</pre>
sph_pow$crit_val_MC
sph_pow$power_MC
## Custom r_H1
# Circular
r_H1 \leftarrow function(n, p, M, 1 = 0.05) {
  stopifnot(p == 2)
  Theta_to_X(matrix(runif(n * M, 0, (2 - 1) * pi), n, M))
dirs <- r_unif_sph(n = 5, p = 2, M = 1)[, , 1]
cir \leftarrow unif_stat_MC(n = 50, M = 1e2, p = 2, CCF09_dirs = dirs)
cir_pow \leftarrow unif_stat_MC(n = 50, M = 1e2, p = 2, r_H1 = r_H1, l = 0.10,
                         crit_val = cir$crit_val_MC, CCF09_dirs = dirs)
cir_pow$crit_val_MC
cir_pow$power_MC
# Spherical
r_H1 \leftarrow function(n, p, M, 1 = 0.5) {
  samp \leftarrow array(dim = c(n, p, M))
  for (j in 1:M) {
    samp[, , j] \leftarrow mvtnorm::rmvnorm(n = n, mean = c(1, rep(0, p - 1)),
                                      sigma = diag(rep(1, p)))
    samp[, , j] <- samp[, , j] / sqrt(rowSums(samp[, , j]^2))</pre>
  }
  return(samp)
dirs <- r_unif_sph(n = 5, p = 3, M = 1)[, , 1]
sph \leftarrow unif_stat_MC(n = 50, M = 1e2, p = 3, CCF09_dirs = dirs)
sph_pow \leftarrow unif_stat_MC(n = 50, M = 1e2, p = 3, r_H1 = r_H1, l = 0.5,
                        crit_val = sph$crit_val_MC, CCF09_dirs = dirs)
sph_pow$power_MC
## Pre-built r_H1
# Circular
dirs <- r_unif_sph(n = 5, p = 2, M = 1)[, , 1]
cir_pow \leftarrow unif_stat_MC(n = 50, M = 1e2, p = 2, r_H1 = r_alt, alt = "vMF",
                          kappa = 1, crit_val = cir$crit_val_MC,
                          CCF09_dirs = dirs)
cir_pow$power_MC
# Spherical
dirs <- r_unif_sph(n = 5, p = 3, M = 1)[, , 1]
sph_pow \leftarrow unif_stat_MC(n = 50, M = 1e2, p = 3, r_H1 = r_alt, alt = "vMF",
                          kappa = 1, crit_val = sph$crit_val_MC,
```

```
CCF09_dirs = dirs)
```

sph\_pow\$power\_MC

unif\_test

Circular and (hyper)spherical uniformity tests

# **Description**

Implementation of several uniformity tests on the (hyper)sphere  $S^{p-1} := \{ \mathbf{x} \in R^p : ||\mathbf{x}|| = 1 \}$ ,  $p \ge 2$ , with calibration either in terms of their asymptotic/exact distributions, if available, or Monte Carlo.

unif\_test receives a sample of directions  $\mathbf{X}_1, \dots, \mathbf{X}_n \in S^{p-1}$  in *Cartesian coordinates*, except for the circular case (p=2) in which the sample can be represented in terms of *angles*  $\Theta_1, \dots, \Theta_n \in [0, 2\pi)$ .

unif\_test allows to perform several tests within a single call, facilitating thus the exploration of a dataset by applying several tests.

# Usage

```
unif_test(data, type = "all", p_value = "asymp", alpha = c(0.1, 0.05, 0.01), M = 10000, stats_MC = NULL, crit_val = NULL, data_sorted = FALSE, K_max = 10000, method = "I", CCF09_dirs = NULL, CJ12_beta = 0, CJ12_reg = 3, cov_a = 2 * pi, Cressie_t = 1/3, K_CCF09 = 25, Poisson_rho = 0.5, Pycke_q = 0.5, Rayleigh_m = 1, Riesz_s = 1, Rothman_t = 1/3, Sobolev_vk2 = c(0, 0, 1), Softmax_kappa = 1, Stereo_a = 0, ...)
```

# **Arguments**

data

sample to perform the test. A matrix of size c(n, p) containing a sample of size n of directions (in Cartesian coordinates) on  $S^{p-1}$ . Alternatively if p = 2, a matrix of size c(n, 1) containing the n angles on  $[0, 2\pi)$  of the circular sample on  $S^1$ . Other objects accepted are an array of size c(n, p, 1) with directions (in Cartesian coordinates), or a vector of size n or an array of size c(n, 1, 1) with angular data. Must not contain NA's.

type

type of test to be applied. A character vector containing any of the following types of tests, depending on the dimension p:

- Circular data: any of the names available at object avail\_cir\_tests.
- (Hyper)spherical data: any of the names available at object avail\_sph\_tests.

If type = "all" (default), then type is set as avail\_cir\_tests or avail\_sph\_tests, depending on the value of p.

p\_value

type of *p*-value computation. Either "MC" for employing the approximation by Monte Carlo of the exact null distribution, "asymp" (default) for the use of the asymptotic/exact null distribution (if available), or "crit\_val" for approximation by means of the table of critical values crit\_val.

Couge

alpha vector with significance levels. Defaults to c(0.10, 0.05, 0.01). number of Monte Carlo replications for approximating the null distribution when approx = "MC". Also, number of Monte Carlo samples for approximating the asymptotic distributions based on weighted sums of chi squared random variables. Defaults to 1e4. stats MC a data frame of size c(M, length(type)), with column names containing the character vector type, that results from extracting \$stats\_MC from a call to unif\_stat\_MC. If provided, the computation of Monte Carlo statistics when approx = "MC" is skipped. stats\_MC is checked internally to see if it is sorted. Internally computed if NULL (default). crit\_val table with critical values for the tests, to be used if p\_value = "crit\_val". A data frame, with column names containing the character vector type and rows corresponding to the significance levels alpha, that results from extracting \$crit\_val\_MC from a call to unif\_stat\_MC. Internally computed if NULL (default). data\_sorted is the circular data sorted? If TRUE, certain statistics are faster to compute. Defaults to FALSE. K max integer giving the truncation of the series that compute the asymptotic p-value of a Sobolev test. Defaults to 1e4. method method for approximating the density, distribution, or quantile function of the weighted sum of chi squared random variables. Must be "I" (Imhof), "SW" (Satterthwaite-Welch), "HBE" (Hall-Buckley-Eagleson), or "MC" (Monte Carlo; only for distribution or quantile functions). Defaults to "I". a matrix of size c(n\_proj, p) containing n\_proj random directions (in Carte-CCF09\_dirs sian coordinates) on  $S^{p-1}$  to perform the CCF09 test. If NULL (default), a sample of size n\_proj = 50 directions is computed internally. CJ12\_beta  $\beta$  parameter in the exponential regime of CJ12 test, a positive real. type of asymptotic regime for CJ12 test, either 1 (sub-exponential regime), 2 CJ12\_reg (exponential), or 3 (super-exponential; default).  $a_n = a/n$  parameter used in the length of the arcs of the coverage-based tests. cov\_a Must be positive. Defaults to 2 \* pi. Cressie\_t t parameter for the Cressie test, a real in (0, 1). Defaults to 1 / 3. K\_CCF09 integer giving the truncation of the series present in the asymptotic distribution of the Kolmogorov-Smirnov statistic. Defaults to 25. Poisson\_rho  $\rho$  parameter for the Poisson test, a real in [0, 1). Defaults to 0.5. q parameter for the Pycke "q-test", a real in (0, 1). Defaults to 1 / 2. Pycke\_q Rayleigh\_m integer m for the m-modal Rayleigh test. Defaults to m = 1 (the standard Rayleigh Riesz\_s s parameter for the s-Riesz test, a real in (0, 2). Defaults to 1. t parameter for the Rothman test, a real in (0,1). Defaults to 1 / 3. Rothman\_t weights for the finite Sobolev test. A non-negative vector or matrix. Defaults to Sobolev\_vk2 c(0, 0, 1).Softmax\_kappa  $\kappa$  parameter for the Softmax test, a non-negative real. Defaults to 1.

Stereo\_a a parameter for the Stereo test, a real in [-1,1]. Defaults to 0. ... If p\_value = "MC" or p\_value = "crit\_val", optional performance parameters to be passed to unif\_stat\_MC: chunks, cores, and seed. If p\_value = "MC", additional parameters to unif\_stat\_distr.

#### **Details**

All the tests reject for large values of the test statistic, so the critical values for the significance levels alpha correspond to the alpha-upper quantiles of the null distribution of the test statistic.

When p\_value = "asymp", tests that do not have an implemented or known asymptotic are omitted, and a warning is generated.

When p\_value = "MC", it is possible to have a progress bar indicating the Monte Carlo simulation progress if unif\_test is wrapped with progressr::with\_progress or if progressr::handlers(global = TRUE) is invoked (once) by the user. See the examples below. The progress bar is updated with the number of finished chunks.

All the statistics are continuous random variables except the Hodges-Ajne statistic ("Hodges\_Ajne"), the Cressie statistic ("Cressie"), and the number of (different) uncovered spacings ("Num\_uncover"). These three statistics are discrete random variables.

The Monte Carlo calibration for the CCF09 test is made conditionally on the choice of CCF09\_dirs. That is, all the Monte Carlo statistics share the same random directions.

Except for CCF09\_dirs, K\_CCF09, and CJ12\_reg, all the test-specific parameters are vectorized.

Descriptions and references for most of the tests are available in García-Portugués and Verdebout (2018).

### Value

If only a **single test** is performed, a list with class htest containing the following components:

- statistic: the value of the test statistic.
- p.value: the p-value of the test. If p\_value = "crit\_val", an NA.
- alternative: a character string describing the alternative hypothesis.
- method: a character string indicating what type of test was performed.
- data.name: a character string giving the name of the data.
- reject: the rejection decision for the levels of significance alpha.
- crit\_val: a vector with the critical values for the significance levels alpha used with p\_value = "MC" or p\_value = "asymp".
- param: parameter(s) used in the test (if any).

If several tests are performed, a type-named list with entries for each test given by the above list.

# References

García-Portugués, E. and Verdebout, T. (2018) An overview of uniformity tests on the hypersphere. *arXiv:1804.00286*. doi:10.48550/arXiv.1804.00286.

```
## Asymptotic distribution
# Circular data
n <- 10
samp_cir <- r_unif_cir(n = n)</pre>
# Matrix
unif_test(data = samp_cir, type = "Ajne", p_value = "asymp")
unif_test(data = samp_cir[, 1], type = "Ajne", p_value = "asymp")
# Array
unif_test(data = array(samp_cir, dim = c(n, 1, 1)), type = "Ajne",
          p_value = "asymp")
# Several tests
unif_test(data = samp_cir, type = avail_cir_tests, p_value = "asymp")
# Spherical data
n <- 10
samp\_sph \leftarrow r\_unif\_sph(n = n, p = 3)
# Array
unif_test(data = samp_sph, type = "Bingham", p_value = "asymp")
unif_test(data = samp_sph[, , 1], type = "Bingham", p_value = "asymp")
# Several tests
unif_test(data = samp_sph, type = avail_sph_tests, p_value = "asymp")
## Monte Carlo
# Circular data
unif_test(data = samp_cir, type = "Ajne", p_value = "MC")
unif_test(data = samp_cir, type = avail_cir_tests, p_value = "MC")
# Spherical data
unif_test(data = samp_sph, type = "Bingham", p_value = "MC")
unif_test(data = samp_sph, type = avail_sph_tests, p_value = "MC")
# Caching stats_MC
stats_MC_cir <- unif_stat_MC(n = nrow(samp_cir), p = 2)$stats_MC</pre>
stats_MC_sph <- unif_stat_MC(n = nrow(samp_sph), p = 3)$stats_MC</pre>
unif_test(data = samp_cir, type = avail_cir_tests,
          p_value = "MC", stats_MC = stats_MC_cir)
unif_test(data = samp_sph, type = avail_sph_tests, p_value = "MC",
          stats_MC = stats_MC_sph)
## Critical values
```

```
# Circular data
unif_test(data = samp_cir, type = avail_cir_tests, p_value = "crit_val")
# Spherical data
unif_test(data = samp_sph, type = avail_sph_tests, p_value = "crit_val")
# Caching crit_val
crit_val_cir <- unif_stat_MC(n = n, p = 2)$crit_val_MC</pre>
crit_val_sph <- unif_stat_MC(n = n, p = 3)$crit_val_MC</pre>
unif_test(data = samp_cir, type = avail_cir_tests,
          p_value = "crit_val", crit_val = crit_val_cir)
unif_test(data = samp_sph, type = avail_sph_tests, p_value = "crit_val",
          crit_val = crit_val_sph)
## Specific arguments
# Rothman
unif_test(data = samp_cir, type = "Rothman", Rothman_t = 0.5)
# CCF09
unif_test(data = samp_sph, type = "CCF09", p_value = "MC",
          CCF09\_dirs = samp\_sph[1:2, , 1])
unif_test(data = samp_sph, type = "CCF09", p_value = "MC",
          CCF09_dirs = samp_sph[3:4, , 1])
## Using a progress bar when p_value = "MC"
# Define a progress bar
require(progress)
require(progressr)
handlers(handler_progress(
  format = paste("(:spin) [:bar] :percent Iter: :current/:total Rate:",
                 ":tick_rate iter/sec ETA: :eta Elapsed: :elapsedfull"),
  clear = FALSE))
# Call unif_test() within with_progress()
with_progress(
  unif_test(data = samp_sph, type = avail_sph_tests, p_value = "MC",
            chunks = 10, M = 1e3)
# With several cores
with_progress(
  unif_test(data = samp_sph, type = avail_sph_tests, p_value = "MC",
            cores = 2, chunks = 10, M = 1e3)
)
# Instead of using with_progress() each time, it is more practical to run
# handlers(global = TRUE)
# once to activate progress bars in your R session
```

venus 85

venus

Venus craters

# **Description**

Craters on Venus from the USGS Astrogeology Science Center.

# Usage

venus

#### **Format**

A data frame with 967 rows and 4 variables:

```
name name of the crater (if named). diameter diameter of the crater (in km). theta longitude angle \theta \in [0,2\pi) of the crater center. phi latitude angle \phi \in [-\pi/2,\pi/2] of the crater center.
```

# **Details**

The  $(\theta, \phi)$  angles are such their associated planetocentric coordinates are:

```
(\cos(\phi)\cos(\theta),\cos(\phi)\sin(\theta),\sin(\phi))',
```

with (0,0,1)' denoting the north pole.

The script performing the data preprocessing is available at venus.R.

# Source

https://astrogeology.usgs.gov/search/map/Venus/venuscraters

wschisq

Weighted sums of non-central chi squared random variables

# **Description**

Approximated density, distribution, and quantile functions for weighted sums of non-central chi squared random variables:

$$Q_K = \sum_{i=1}^K w_i \chi_{d_i}^2(\lambda_i),$$

where  $w_1, \ldots, w_n$  are positive weights,  $d_1, \ldots, d_n$  are positive degrees of freedom, and  $\lambda_1, \ldots, \lambda_n$  are non-negative non-centrality parameters. Also, simulation of  $Q_K$ .

### Usage

```
d_wschisq(x, weights, dfs, ncps = 0, method = c("I", "SW", "HBE")[1],
  exact_chisq = TRUE, imhof_epsabs = 1e-06, imhof_epsrel = 1e-06,
  imhof_limit = 10000, grad_method = "simple",
  grad_method.args = list(eps = 1e-07))

p_wschisq(x, weights, dfs, ncps = 0, method = c("I", "SW", "HBE", "MC")[1],
  exact_chisq = TRUE, imhof_epsabs = 1e-06, imhof_epsrel = 1e-06,
  imhof_limit = 10000, M = 10000, MC_sample = NULL)

q_wschisq(u, weights, dfs, ncps = 0, method = c("I", "SW", "HBE", "MC")[1],
  exact_chisq = TRUE, imhof_epsabs = 1e-06, imhof_epsrel = 1e-06,
  imhof_limit = 10000, nlm_gradtol = 1e-06, nlm_iterlim = 1000,
  M = 10000, MC_sample = NULL)

r_wschisq(n, weights, dfs, ncps = 0)

cutoff_wschisq(thre = 1e-04, weights, dfs, ncps = 0, log = FALSE,
  x_tail = NULL)
```

# Arguments

X	vector of quantiles.
weights	vector with the positive weights of the sum. Must have the same length as dfs.
dfs	vector with the positive degrees of freedom of the chi squared random variables. Must have the same length as weights.
ncps	non-centrality parameters. Either 0 (default) or a vector with the same length as weights.
method	method for approximating the density, distribution, or quantile function. Must be "I" (Imhof), "SW" (Satterthwaite–Welch), "HBE" (Hall–Buckley–Eagleson), or "MC" (Monte Carlo; only for distribution or quantile functions). Defaults to "I".

exact\_chisq if weights and dfs have length one, shall the Chisquare functions be called?

Otherwise, the approximations are computed for this exact case. Defaults to TRUE.

imhof\_epsabs, imhof\_epsrel, imhof\_limit

precision parameters passed to imhof's epsabs, epsrel, and limit, respectively. They default to 1e-6, 1e-6, and 1e4.

grad\_method, grad\_method.args

numerical differentiation parameters passed to grad's method and method.args, respectively. They default to "simple", and list(eps = 1e-7) (better precision than imhof\_epsabs to avoid numerical artifacts).

M number of Monte Carlo samples for approximating the distribution if method =

"MC". Defaults to 1e4.

MC\_sample if provided, it is employed when method = "MC". If not, it is computed internally.

u vector of probabilities.

nlm\_gradtol, nlm\_iterlim

convergence control parameters passed to nlm's gradtol and iterlim, respectively. They default to 1e-6 and 1e3.

n sample size.

thre vector with the error thresholds of the tail probability and mean/variance ex-

plained by the first terms of the series. Defaults to 1e-4. See details.

log are weights and dfs given in log-scale? Defaults to FALSE.

x\_tail scalar evaluation point for determining the upper tail probability. If NULL, set to

the 0.90 quantile of the whole series, computed by the "HBE" approximation.

# Details

Four methods are implemented for approximating the distribution of a weighted sum of chi squared random variables:

- "I": Imhof's approximation (Imhof, 1961) for the evaluation of the distribution function. If this method is selected, the function is simply a wrapper to imhof from the CompQuadForm package (Duchesne and Lafaye De Micheaux, 2010).
- "SW": Satterthwaite–Welch (Satterthwaite, 1946; Welch, 1938) approximation, consisting in matching the first two moments of  $Q_K$  with a gamma distribution.
- "HBE": Hall–Buckley–Eagleson (Hall, 1983; Buckley and Eagleson, 1988) approximation, consisting in matching the first *three* moments of  $Q_K$  with a gamma distribution.
- "MC": Monte Carlo approximation using the empirical cumulative distribution function with M simulated samples.

The Imhof method is exact up to the prescribed numerical accuracy. It is also the most time-consuming method. The density and quantile functions for this approximation are obtained by numerical differentiation and inversion, respectively, of the approximated distribution.

For the methods based on gamma matching, the GammaDist density, distribution, and quantile functions are invoked. The Hall–Buckley–Eagleson approximation tends to overperform the Satterthwaite–Welch approximation.

The Monte Carlo method is relatively inaccurate and slow, but serves as an unbiased reference of the true distribution function. The inversion of the empirical cumulative distribution is done by quantile.

An empirical comparison of these and other approximation methods is given in Bodenham and Adams (2016).

cutoff\_wschisq removes NAs/NaNs in weights or dfs with a message. The threshold thre ensures that the tail probability of the truncated and whole series differ less than thre at x\_tail, or that thre is the proportion of the mean/variance of the whole series that is *not* retained. The (upper) tail probabilities for evaluating truncation are computed using the Hall–Buckley–Eagleson approximation at x\_tail.

#### Value

- d\_wschisq: density function evaluated at x, a vector.
- p\_wschisq: distribution function evaluated at x, a vector.
- q\_wschisq: quantile function evaluated at u, a vector.
- r\_wschisq: a vector of size n containing a random sample.
- cutoff\_wschisq: a data frame with the indexes up to which the truncated series explains the tail probability with absolute error thre, or the proportion of the mean/variance of the whole series that is *not* explained by the truncated series.

# Author(s)

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### References

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Welch, B. L. (1938). The significance of the difference between two means when the population variances are unequal. *Biometrika*, 29(3/4):350–362. doi:10.2307/2332010

```
# Plotting functions for the examples
add_approx_dens <- function(x, dfs, weights, ncps) {</pre>
 lines(x, d_wschisq(x, weights = weights, dfs = dfs, ncps = ncps,
                     method = "SW", exact_chisq = FALSE), col = 3)
 lines(x, d_wschisq(x, weights = weights, dfs = dfs, ncps = ncps,
                     method = "HBE", exact_chisq = FALSE), col = 4)
 lines(x, d_wschisq(x, weights = weights, dfs = dfs, ncps = ncps.)
                     method = "I", exact_chisq = TRUE), col = 2)
 legend("topright", legend = c("True", "SW", "HBE", "I"), lwd = 2,
         col = c(1, 3:4, 2))
}
add_approx_distr <- function(x, dfs, weights, ncps, ...) {</pre>
 lines(x, p_wschisq(x, weights = weights, dfs = dfs, ncps = ncps,
                     method = "SW", exact_chisq = FALSE), col = 3)
 lines(x, p_{wschisq}(x, weights = weights, dfs = dfs, ncps = ncps,
                     method = "HBE", exact_chisq = FALSE), col = 4)
  lines(x, p_wschisq(x, weights = weights, dfs = dfs, ncps = ncps,
                     method = "MC", exact_chisq = FALSE), col = 5,
                     type = "s")
 lines(x, p_wschisq(x, weights = weights, dfs = dfs, ncps = ncps,
                     method = "I", exact_chisq = TRUE), col = 2)
 legend("bottomright", legend = c("True", "SW", "HBE", "MC", "I"), lwd = 2,
         col = c(1, 3:5, 2))
add_approx_quant <- function(u, dfs, weights, ncps, ...) {</pre>
 lines(u, q_wschisq(u, weights = weights, dfs = dfs, ncps = ncps,
                     method = "SW", exact_chisq = FALSE), col = 3)
 lines(u, q_wschisq(u, weights = weights, dfs = dfs, ncps = ncps,
                     method = "HBE", exact_chisq = FALSE), col = 4)
 lines(u, q_wschisq(u, weights = weights, dfs = dfs, ncps = ncps,
                     method = "MC", exact_chisq = FALSE), col = 5,
                     type = "s")
 lines(u, q_wschisq(u, weights = weights, dfs = dfs, ncps = ncps,
                     method = "I", exact_chisq = TRUE), col = 2)
 legend("topleft", legend = c("True", "SW", "HBE", "MC", "I"), lwd = 2,
        col = c(1, 3:5, 2)
}
# Validation plots for density, distribution, and quantile functions
u \leftarrow seq(0.01, 0.99, 1 = 100)
old_par <- par(mfrow = c(1, 3))
# Case 1: 1 * ChiSq_3(0) + 1 * ChiSq_3(0) = ChiSq_6(0)
weights \leftarrow c(1, 1)
dfs <- c(3, 3)
```

```
ncps <- 0
x < - seq(-1, 30, 1 = 100)
main <- expression(1 * chi[3]^2 * (0) + 1 * chi[3]^2 * (0))
plot(x, dchisq(x, df = 6), type = "1", main = main, ylab = "Density")
add_approx_dens(x = x, weights = weights, dfs = dfs, ncps = ncps)
plot(x, pchisq(x, df = 6), type = "l", main = main, ylab = "Distribution")
add_approx_distr(x = x, weights = weights, dfs = dfs, ncps = ncps)
plot(u, qchisq(u, df = 6), type = "l", main = main, ylab = "Quantile")
add_approx_quant(u = u, weights = weights, dfs = dfs, ncps = ncps)
# Case 2: 2 * ChiSq_3(1) + 1 * ChiSq_6(0.5) + 0.5 * ChiSq_12(0.25)
weights <- c(2, 1, 0.5)
dfs <- c(3, 6, 12)
ncps \leftarrow c(1, 0.5, 0.25)
x < - seq(0, 70, 1 = 100)
main <- expression(2 * chi[3]^2 * (1)+ 1 * chi[6]^2 * (0.5) +
                   0.5 * chi[12]^2 * (0.25))
samp \leftarrow r_wschisq(n = 1e4, weights = weights, dfs = dfs, ncps = ncps)
hist(samp, breaks = 50, freq = FALSE, main = main, ylab = "Density",
     xlim = range(x), xlab = "x"); box()
add_approx_dens(x = x, weights = weights, dfs = dfs, ncps = ncps)
plot(x, ecdf(samp)(x), main = main, ylab = "Distribution", type = "s")
add_approx_distr(x = x, weights = weights, dfs = dfs, ncps = ncps)
plot(u, quantile(samp, probs = u), type = "s", main = main,
     ylab = "Quantile")
add_approx_quant(u = u, weights = weights, dfs = dfs, ncps = ncps)
# Case 3: \sum_{k=1}^K k^{-3} * ChiSq_{5k}(1 / k^2)
K <- 1e2
weights<- 1 / (1:K)^3
dfs <- 5 * 1:K
ncps <- 1 / (1:K)<sup>2</sup>
x \leftarrow seq(0, 25, 1 = 100)
main <- substitute(sum(k^{-3}) * chi[5 * k]^2 * (1 / k^2), k == 1, K),
                   list(K = K))
samp <- r_wschisq(n = 1e4, weights = weights, dfs = dfs, ncps = ncps)
hist(samp, breaks = 50, freq = FALSE, main = main, ylab = "Density",
     xlim = range(x), xlab = "x"); box()
add_approx_dens(x = x, weights = weights, dfs = dfs, ncps = ncps)
plot(x, ecdf(samp)(x), main = main, ylab = "Distribution", type = "s")
add_approx_distr(x = x, weights = weights, dfs = dfs, ncps = ncps)
plot(u, quantile(samp, probs = u), type = "s", main = main,
     ylab = "Quantile")
add_approx_quant(u = u, weights = weights, dfs = dfs, ncps = ncps)
par(old_par)
# Cutoffs for infinite series of the last example
K <- 1e7
log_weights < -3 * log(1:K)
\log_{dfs} \leftarrow \log(5) + \log(1:K)
(cutoff <- cutoff_wschisq(thre = 10^(-(1:4)), weights = log_weights,</pre>
                          dfs = log_dfs, log = TRUE))
```

# **Index**

* datasets	cir_stat_Gini (cir_stat_Kuiper), 9
avail_tests, 5	cir_stat_Gini_squared
comets, 15	(cir_stat_Kuiper), 9
craters, 18	<pre>cir_stat_Greenwood(cir_stat_Kuiper), 9</pre>
planets, 33	cir_stat_Hermans_Rasson
rhea, 52	cir_stat_Kuiper), 9
venus, 85	<pre>cir_stat_Hodges_Ajne (cir_stat_Kuiper),</pre>
	9
GaussLegendre quadrature, 23, 30, 35,	cir_stat_Kuiper,9
58	<pre>cir_stat_Log_gaps(cir_stat_Kuiper),9</pre>
	<pre>cir_stat_Max_uncover(cir_stat_Kuiper),</pre>
A_theta_x, 5	9
akx, 62	<pre>cir_stat_Num_uncover(cir_stat_Kuiper),</pre>
akx (Pn), 34	9
angles_to_sphere, 4	<pre>cir_stat_PAD (cir_stat_Kuiper), 9</pre>
angular function, $19$	<pre>cir_stat_PCvM(cir_stat_Kuiper),9</pre>
asymptotic distribution, 61	<pre>cir_stat_Poisson(cir_stat_Kuiper),9</pre>
avail_cir_tests, <i>31</i> , <i>58</i> , <i>69</i> , <i>72</i> , <i>76</i> , <i>80</i>	<pre>cir_stat_PRt (cir_stat_Kuiper), 9</pre>
<pre>avail_cir_tests (avail_tests), 5</pre>	cir_stat_Pycke(cir_stat_Kuiper),9
avail_sph_tests, <i>31</i> , <i>58</i> , <i>69</i> , <i>72</i> , <i>76</i> , <i>80</i>	<pre>cir_stat_Pycke_q(cir_stat_Kuiper),9</pre>
<pre>avail_sph_tests (avail_tests), 5</pre>	<pre>cir_stat_Range(cir_stat_Kuiper), 9</pre>
avail_tests,5	<pre>cir_stat_Rao(cir_stat_Kuiper),9</pre>
	<pre>cir_stat_Rayleigh(cir_stat_Kuiper), 9</pre>
bk_to_uk (Sobolev_coefs), 61	<pre>cir_stat_Riesz(cir_stat_Kuiper), 9</pre>
bk_to_vk2 (Sobolev_coefs), 61	<pre>cir_stat_Rothman(cir_stat_Kuiper),9</pre>
	<pre>cir_stat_Sobolev (sph_stat_Sobolev), 66</pre>
c_unif_cap (unif_cap), 67	<pre>cir_stat_Softmax(cir_stat_Kuiper),9</pre>
Chisquare, 87	<pre>cir_stat_Vacancy(cir_stat_Kuiper),9</pre>
cir_coord_conv, 7	<pre>cir_stat_Watson(cir_stat_Kuiper),9</pre>
cir_gaps, 8, <i>11</i>	<pre>cir_stat_Watson_1976(cir_stat_Kuiper),</pre>
<pre>cir_stat (cir_stat_Kuiper), 9</pre>	9
<pre>cir_stat_Ajne (cir_stat_Kuiper), 9</pre>	comets, 15
cir_stat_Bakshaev(cir_stat_Kuiper),9	con_f (locdev), 29
<pre>cir_stat_Bingham(cir_stat_Kuiper), 9</pre>	craters, 18
<pre>cir_stat_CCF09 (cir_stat_Kuiper), 9</pre>	cutoff_locdev (locdev), 29
<pre>cir_stat_Cressie (cir_stat_Kuiper), 9</pre>	cutoff_wschisq,59
cir_stat_distr(p_Kolmogorov),41	cutoff_wschisq(wschisq), 86
cir_stat_FG01 (cir_stat_Kuiper), 9	
<pre>cir_stat_Gine_Fn (cir_stat_Kuiper), 9</pre>	<pre>d_cir_stat_Ajne(p_Kolmogorov), 41</pre>
<pre>cir_stat_Gine_Gn (cir_stat_Kuiper), 9</pre>	<pre>d_cir_stat_Bakshaev(p_Kolmogorov), 41</pre>

INDEX 93

<pre>d_cir_stat_Bingham (p_Kolmogorov), 41</pre>	d_sph_stat_Gine_Fn
<pre>d_cir_stat_Gine_Fn (p_Kolmogorov), 41</pre>	(p_sph_stat_Bingham), 48
<pre>d_cir_stat_Gine_Gn (p_Kolmogorov), 41</pre>	d_sph_stat_Gine_Gn
<pre>d_cir_stat_Gini (p_Kolmogorov), 41</pre>	(p_sph_stat_Bingham), 48
<pre>d_cir_stat_Gini_squared(p_Kolmogorov),</pre>	<pre>d_sph_stat_PAD (p_sph_stat_Bingham), 48</pre>
41	<pre>d_sph_stat_PCvM(p_sph_stat_Bingham), 48</pre>
<pre>d_cir_stat_Greenwood(p_Kolmogorov), 41</pre>	d_sph_stat_Poisson
d_cir_stat_Hermans_Rasson	(p_sph_stat_Bingham), 48
(p_Kolmogorov), 41	d_sph_stat_PRt (p_sph_stat_Bingham), 48
d_cir_stat_Hodges_Ajne (p_Kolmogorov),	d_sph_stat_Rayleigh
41	(p_sph_stat_Bingham), 48
d_cir_stat_Kuiper(p_Kolmogorov),41	d_sph_stat_Rayleigh_HD
	(p_sph_stat_Bingham), 48
d_cir_stat_Log_gaps (p_Kolmogorov), 41	d_sph_stat_Riesz (p_sph_stat_Bingham),
<pre>d_cir_stat_Max_uncover(p_Kolmogorov),</pre>	
41	48
<pre>d_cir_stat_Num_uncover(p_Kolmogorov),</pre>	d_sph_stat_Sobolev
41	(p_sph_stat_Bingham), 48
<pre>d_cir_stat_PAD (p_Kolmogorov), 41</pre>	d_sph_stat_Softmax
<pre>d_cir_stat_PCvM(p_Kolmogorov), 41</pre>	(p_sph_stat_Bingham), 48
<pre>d_cir_stat_Poisson(p_Kolmogorov), 41</pre>	<pre>d_sph_stat_Stereo(p_sph_stat_Bingham),</pre>
d_cir_stat_PRt (p_Kolmogorov), 41	48
d_cir_stat_Pycke (p_Kolmogorov), 41	d_unif_cap(unif_cap),67
<pre>d_cir_stat_Pycke_q (p_Kolmogorov), 41</pre>	d_wschisq(wschisq), 86
d_cir_stat_Range (p_Kolmogorov), 41	Date, <i>16</i>
d_cir_stat_Rao (p_Kolmogorov), 41	
d_cir_stat_Rayleigh (p_Kolmogorov), 41	F_from_f, 19
	F_inv_from_f, <i>30</i> , <i>54</i>
d_cir_stat_Riesz (p_Kolmogorov), 41	F_inv_from_f(F_from_f), 19
d_cir_stat_Rothman (p_Kolmogorov), 41	f_locdev (locdev), 29
d_cir_stat_Sobolev (p_Kolmogorov), 41	f_locdev_Pn (Pn), 34
<pre>d_cir_stat_Softmax (p_Kolmogorov), 41</pre>	foreach, 28, 77
<pre>d_cir_stat_Vacancy (p_Kolmogorov), 41</pre>	, ,
<pre>d_cir_stat_Watson(p_Kolmogorov), 41</pre>	g_i_k (harmonics), 27
<pre>d_cir_stat_Watson_1976 (p_Kolmogorov),</pre>	GammaDist, 87
41	GaussLegendre quadrature, 19, 35
d_Kolmogorov (p_Kolmogorov), 41	Gauss-Legendre quadrature, 6, 44, 63
d_locdev (locdev), 29	Gauss_Legen, 20
d_p_k, 27	Gauss_Legen_nodes (Gauss_Legen), 20
d_p_k (Sobolev), 57	Gauss_Legen_weights (Gauss_Legen), 20
d_proj_unif(proj_unif), 38	Gegen_coefs (Gegenbauer), 22
<pre>d_proj_unif_cap (unif_cap), 67</pre>	Gegen_coefs_2d (Gegenbauer), 22
d_Sobolev, 44, 50	Gegen_coefs_Pn (Pn), 34
d_Sobolev (Sobolev), 57	Gegen_norm (Gegenbauer), 22
	· · · · · · · · · · · · · · · · · · ·
d_sph_stat_Ajne (p_sph_stat_Bingham), 48	Gegen_norm_2d (Gegenbauer), 22
d_sph_stat_Bakshaev	Gegen_polyn (Gegenbauer), 22
(p_sph_stat_Bingham), 48	Gegen_polyn_2d (Gegenbauer), 22
d_sph_stat_Bingham	Gegen_series (Gegenbauer), 22
(p_sph_stat_Bingham), 48	Gegen_series_2d (Gegenbauer), 22
d_sph_stat_CJ12 (p_sph_stat_Bingham), 48	Gegenbauer, 22

94 INDEX

Gegenbauer coefficients, 34,61	<pre>p_cir_stat_Riesz(p_Kolmogorov), 41</pre>
Gegenbauer series, $30$	<pre>p_cir_stat_Rothman(p_Kolmogorov), 41</pre>
gegenpoly_array,24	<pre>p_cir_stat_Sobolev (p_Kolmogorov), 41</pre>
gegenpoly_n, 24	<pre>p_cir_stat_Softmax (p_Kolmogorov), 41</pre>
grad, <i>87</i>	<pre>p_cir_stat_Vacancy (p_Kolmogorov), 41</pre>
	<pre>p_cir_stat_Watson(p_Kolmogorov), 41</pre>
harmonics, 27	p_cir_stat_Watson_1976 (p_Kolmogorov),
	41
imhof, 87	p_Kolmogorov, 41
int_sph_MC, 28	p_proj_unif, 67
integrate, 19, 23, 31, 35, 58	p_proj_unif(proj_unif), 38
intersection of two hyperspherical	p_proj_unif_cap (unif_cap), 67
caps, <i>34</i>	p_Sobolev, 44, 50
	p_Sobolev (Sobolev), 57
local projected alternative, 61	p_sph_stat_Ajne (p_sph_stat_Bingham), 48
locdev, 29	p_sph_stat_Bakshaev
	(p_sph_stat_Bingham), 48
nlm, 87	p_sph_stat_Bingham, 48
	p_sph_stat_CJ12 (p_sph_stat_Bingham), 48
<pre>p_cir_stat_Ajne (p_Kolmogorov), 41</pre>	p_sph_stat_Gine_Fn
<pre>p_cir_stat_Bakshaev(p_Kolmogorov), 41</pre>	
p_cir_stat_Bingham(p_Kolmogorov),41	(p_sph_stat_Bingham), 48
<pre>p_cir_stat_Gine_Fn (p_Kolmogorov), 41</pre>	p_sph_stat_Gine_Gn
<pre>p_cir_stat_Gine_Gn (p_Kolmogorov), 41</pre>	(p_sph_stat_Bingham), 48
<pre>p_cir_stat_Gini (p_Kolmogorov), 41</pre>	p_sph_stat_PAD (p_sph_stat_Bingham), 48
<pre>p_cir_stat_Gini_squared(p_Kolmogorov),</pre>	p_sph_stat_PCvM(p_sph_stat_Bingham), 48
41	p_sph_stat_Poisson
<pre>p_cir_stat_Greenwood(p_Kolmogorov), 41</pre>	(p_sph_stat_Bingham), 48
p_cir_stat_Hermans_Rasson	p_sph_stat_PRt (p_sph_stat_Bingham), 48
(p_Kolmogorov), 41	p_sph_stat_Rayleigh
<pre>p_cir_stat_Hodges_Ajne (p_Kolmogorov),</pre>	(p_sph_stat_Bingham), 48
41	p_sph_stat_Rayleigh_HD
<pre>p_cir_stat_Hodges_Ajne2 (p_Kolmogorov),</pre>	(p_sph_stat_Bingham), 48
41	<pre>p_sph_stat_Riesz (p_sph_stat_Bingham),</pre>
<pre>p_cir_stat_Kuiper (p_Kolmogorov), 41</pre>	48
<pre>p_cir_stat_Log_gaps (p_Kolmogorov), 41</pre>	p_sph_stat_Sobolev
<pre>p_cir_stat_Max_uncover (p_Kolmogorov),</pre>	(p_sph_stat_Bingham), 48
41	p_sph_stat_Softmax
<pre>p_cir_stat_Num_uncover (p_Kolmogorov),</pre>	(p_sph_stat_Bingham), 48
41	<pre>p_sph_stat_Stereo(p_sph_stat_Bingham),</pre>
p_cir_stat_PAD(p_Kolmogorov),41	48
<pre>p_cir_stat_PCvM(p_Kolmogorov), 41</pre>	p_wschisq(wschisq), 86
<pre>p_cir_stat_Poisson(p_Kolmogorov), 41</pre>	planets, 33
<pre>p_cir_stat_PRt (p_Kolmogorov), 41</pre>	Pn, 34
<pre>p_cir_stat_Pycke (p_Kolmogorov), 41</pre>	progressr::with_progress, 29,77,82
<pre>p_cir_stat_Pycke_q(p_Kolmogorov), 41</pre>	proj_unif,38
<pre>p_cir_stat_Range (p_Kolmogorov), 41</pre>	projected alternatives to uniformity,
p_cir_stat_Rao(p_Kolmogorov),41	34
p cir stat Ravleigh (p Kolmogorov), 41	projected spherical uniform

INDEX 95

distribution, 34 projected-ecdf statistics, 61 Psi, 39 Psi_mat, 11, 12, 63, 64, 66, 67 Psi_mat (Psi), 39 psi_Pn (Pn), 34	<pre>sph_stat_Stereo(sph_stat_Rayleigh), 62 sphere_to_angles(angles_to_sphere), 4 sphunif, 5 sphunif(sphunif-package), 3 sphunif-package, 3 splinefun, 20</pre>
<pre>psi_Pn (Pn), 34  q_proj_unif (proj_unif), 38 q_proj_unif_cap (unif_cap), 67 q_Sobolev (Sobolev), 57 q_wschisq (wschisq), 86 quantile, 88  r_alt, 53 r_locdev (locdev), 29 r_proj_unif (proj_unif), 38 r_proj_unif_cap (unif_cap), 67 r_unif, 56 r_unif_cap (unif_cap), 67 r_unif_sph, 76 r_unif_sph, 76 r_unif_sph (r_unif), 56 r_wschisq (wschisq), 86 rhea, 18, 52 Rothman test, 62  Sobolev_coefs, 61 sph_stat (sph_stat_Rayleigh), 62 sph_stat_Bingham (sph_stat_Rayleigh), 62 sph_stat_Bingham (sph_stat_Rayleigh), 62 sph_stat_CJ12 (sph_stat_Rayleigh), 62 sph_stat_Gine_Fn (sph_stat_Rayleigh), 62 sph_stat_Gine_Fn (sph_stat_Rayleigh), 62 sph_stat_Gine_Fn (sph_stat_Rayleigh), 62 sph_stat_Gine_Gn (sph_stat_Rayleigh), 62 sph_stat_PAD (sph_stat_Rayleigh), 62 sph_stat_POvM (sph_stat_Rayleigh), 62 sph_stat_Poisson (sph_stat_Rayleigh), 62 sph_stat_PCvM (sph_stat_Rayleigh), 62 sph_stat_PRt (sph_stat_Rayleigh), 62 sph_stat_PRt (sph_stat_Rayleigh), 62 sph_stat_Pycke (sph_stat_Rayleigh), 62 sph_stat_Rayleigh, 62 sph_stat_Rayleigh_HD</pre>	
(sph_stat_Rayleigh), 62	
<pre>sph_stat_Riesz (sph_stat_Rayleigh), 62</pre>	
sph_stat_Sobolev, 66	
<pre>sph_stat_Softmax (sph_stat_Rayleigh), 62</pre>	