

Quasi-Monte Carlo for optimal control and optimal experimental design problems governed by PDEs

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Part I: Quasi-Monte Carlo methods

Lattice rules

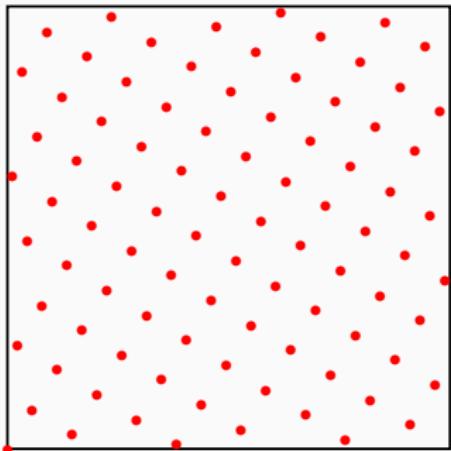
Rank-1 lattice rules

$$Q_{s,n}(f) = \frac{1}{n} \sum_{i=1}^n f(\mathbf{t}_i)$$

have the points

$$\mathbf{t}_i = \text{mod}\left(\frac{i\mathbf{z}}{n}, 1\right), \quad i \in \{1, \dots, n\},$$

where the entire point set is determined by the *generating vector* $\mathbf{z} \in \mathbb{N}^s$, with all components *coprime* to n .



Lattice rule with $\mathbf{z} = (1, 55)$ and $n = 89$
nodes in $[0, 1]^2$

The quality of the lattice rule is determined by the choice of \mathbf{z} .

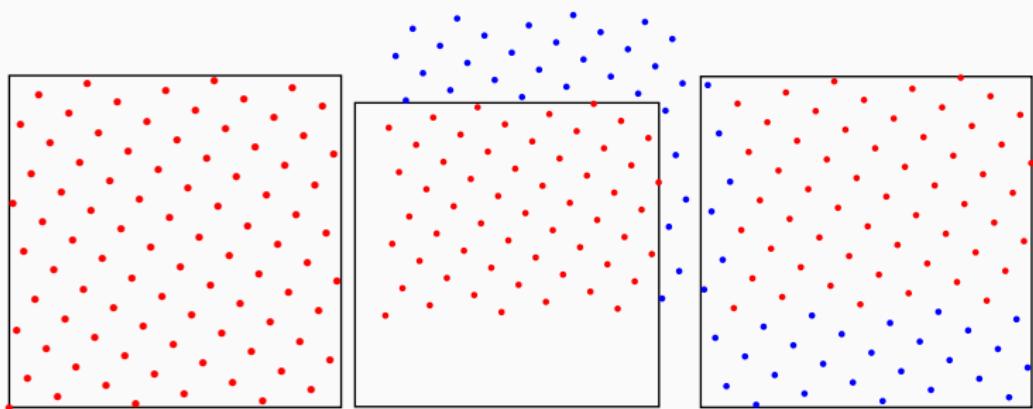
Randomly shifted lattice rules

Shifted rank-1 lattice rules have points

$$\mathbf{t}_i = \text{mod}\left(\frac{i\mathbf{z}}{n} + \Delta, 1\right), \quad i \in \{1, \dots, n\}.$$

$\Delta \in [0, 1)^s$ is the *shift*

Use a number of random shifts for error estimation.



Lattice rule shifted by $\Delta = (0.1, 0.3)$.

Let $\Delta^{(r)}$, $r = 1, \dots, R$, be independent random shifts drawn from $U([0, 1]^s)$ and define

$$Q_{s,n}^{(r)}(f) := \frac{1}{n} \sum_{i=1}^n f(\text{mod}(\mathbf{t}_i + \Delta^{(r)}, 1)). \quad (\text{QMC rule with 1 random shift})$$

Then

$$\overline{Q}_{s,n}(f) = \frac{1}{R} \sum_{r=1}^R Q_{s,n}^{(r)} f \quad (\text{QMC rule with } R \text{ random shifts})$$

is an unbiased estimator of $I_s(f)$.

Let $f: [0, 1]^s \rightarrow \mathbb{R}$ be sufficiently smooth.

Error bound (one random shift):

$$|I_s(f) - Q_{s,n}^{\Delta}(f)| \leq e_{s,n,\gamma}^{\Delta}(z) \|f\|_{\gamma}.$$

R.M.S. error bound (shift-averaged):

$$\sqrt{\mathbb{E}_{\Delta}[|I_s(f) - \bar{Q}_{s,n}(f)|^2]} \leq e_{s,n,\gamma}^{\text{sh}}(z) \|f\|_{\gamma}.$$

We consider weighted Sobolev spaces with dominating mixed smoothness, equipped with norm

$$\|f\|_{\gamma}^2 = \sum_{\mathbf{u} \subseteq \{1:s\}} \frac{1}{\gamma_{\mathbf{u}}} \int_{[0,1]^{|\mathbf{u}|}} \left(\int_{[0,1]^{s-|\mathbf{u}|}} \frac{\partial^{|\mathbf{u}|} f}{\partial \mathbf{y}_{\mathbf{u}}}(\mathbf{y}) d\mathbf{y}_{-\mathbf{u}} \right)^2 d\mathbf{y}_{\mathbf{u}}$$

and (squared) worst case error

$$P(z) := e_{s,n,\gamma}^{\text{sh}}(z)^2 = \frac{1}{n} \sum_{k=0}^{n-1} \sum_{\emptyset \neq \mathbf{u} \subseteq \{1:s\}} \gamma_{\mathbf{u}} \prod_{j \in \mathbf{u}} \omega\left(\left\{ \frac{kz_j}{n} \right\}\right)$$

where $\omega(x) = x^2 - x + \frac{1}{6}$.

CBC algorithm (Sloan, Kuo, Joe 2002)

The idea of the *component-by-component* (CBC) algorithm is to find a good generating vector $\mathbf{z} = (z_1, \dots, z_s)$ by proceeding as follows:

1. Set $z_1 = 1$;
 2. With z_1 fixed, choose z_2 to minimize error criterion $P(z_1, z_2)$;
 3. With z_1 and z_2 fixed, choose z_3 to minimize error criterion
 $P(z_1, z_2, z_3)$
- ⋮

Efficient implementation using FFT (QMC4PDE, QMCPy, etc.) if weights have certain structure (e.g., POD weights).

Theorem (CBC error bound)

The generating vector $\mathbf{z} \in \mathbb{N}^s$ constructed by the CBC algorithm, minimizing the squared shift-averaged worst-case error $[e_{s,n,\gamma}^{\text{sh}}(\mathbf{z})]^2$ for the weighted unanchored Sobolev space in each step, satisfies

$$[e_{s,n,\gamma}^{\text{sh}}(\mathbf{z})]^2 \leq \left(\frac{1}{\varphi(n)} \sum_{\emptyset \neq u \subseteq \{1:s\}} \gamma_u^\lambda \left(\frac{2\zeta(2\lambda)}{(2\pi^2)^\lambda} \right)^{|u|} \right)^{1/\lambda} \quad \text{for all } \lambda \in (1/2, 1],$$

where $\zeta(x) := \sum_{k=1}^{\infty} k^{-x}$ denotes the Riemann zeta function for $x > 1$.

Remarks:

- Optimal rate of convergence $\mathcal{O}(n^{-1+\delta})$ in weighted Sobolev spaces, independently of s under an appropriate condition on the weights.
- Cost of algorithm for POD weights is $\mathcal{O}(s n \log n + s^2 n)$ using FFT.

Significance: Suppose that $f \in H_{s,\gamma}$ for all $\gamma = (\gamma_u)_{u \subseteq \{1:s\}}$. Then for any given sequence of weights γ , we can use the CBC algorithm to obtain a generating vector satisfying the error bound

$$\sqrt{\mathbb{E}_{\Delta} |I_s f - Q_{n,s}^{\Delta} f|^2} \leq \left(\frac{1}{\varphi(n)} \sum_{\emptyset \neq u \subseteq \{1:s\}} \gamma_u^\lambda \left(\frac{2\zeta(2\lambda)}{(2\pi^2)^\lambda} \right)^{|u|} \right)^{1/(2\lambda)} \|f\|_{s,\gamma} \quad (1)$$

for all $\lambda \in (1/2, 1]$. We can use the following strategy:

- For a given integrand f , estimate the norm $\|f\|_{s,\gamma}$.
- Find weights γ which *minimize* the error bound (1).
- Using the optimized weights γ as input, use the CBC algorithm to find a generating vector which *satisfies* the error bound (1).

Part II: Optimal control problems under uncertainty

Optimal control

- Goal:

- steer u toward target \mathbf{g}

$$\min_u \left(\frac{\alpha_1}{2} \int_0^T \|u(\cdot, t) - \mathbf{g}(\cdot, t)\|^2 dt + \frac{\alpha_2}{2} \|u(\cdot, T) - \mathbf{g}(\cdot, T)\|^2 \right)$$

- control u through the source \mathbf{z} via the PDE

$$\partial_t u - \nabla \cdot (\mathbf{a} \nabla u) = \mathbf{z} \quad \text{in } D \times (0, T)$$

- Resulting optimal control problem $\min_z J(z)$, where

$$J(z) = \mathcal{R} \left(\frac{\alpha_1}{2} \int_0^T \|u(\cdot, t) - \mathbf{g}(\cdot, t)\|^2 dt + \frac{\alpha_2}{2} \|u(\cdot, T) - \mathbf{g}(\cdot, T)\|^2 \right) + \frac{\alpha_3}{2} \int_0^T \|\mathbf{z}(\cdot, t)\|^2 dt$$

subject to

$$\partial_t u - \nabla \cdot (\mathbf{a} \nabla u) = \mathbf{z} \quad \text{in } D \times (0, T)$$

+ initial-boundary values at $t = 0$ and on ∂D

- How do we handle uncertainty in $a\mathbf{y}(\mathbf{x})$, parameterized by $\mathbf{y} \in U$?

- \mathcal{R} risk measure

- Robust optimization: $\mathcal{R}(\cdot) = \int_U \cdot d\mathbf{y}$ (risk neutral)

- Entropic risk measure: $\mathcal{R}(\cdot) = \frac{1}{\theta} \ln \left(\int_U \exp(\theta \cdot) d\mathbf{y} \right)$ (risk averse)

Optimal control problems

[Borzì, Schulz, Schillings, Winckel (2010); Ali, Ullmann, Hinze (2017)]

- **Mean-based control:** replace a by $\int_U a^y \, dy$
- **Pathwise control:** for several y , find

$$\begin{aligned} z^*(y) = \arg \min & \left(\frac{\alpha_1}{2} \int_0^T \|u^y(\cdot, t) - g(\cdot, t)\|^2 dt + \frac{\alpha_2}{2} \|u^y(\cdot, T) - g(\cdot, T)\|^2 \right) \\ & + \frac{\alpha_3}{2} \int_0^T \|z(\cdot, t)\|^2 dt \end{aligned}$$

and compute the statistics of $z^*(y)$, e.g., $\mathbb{E}[z^*(\cdot)]$

- **Stochastic control:**

$$J(z) = \mathcal{R} \left(\frac{\alpha_1}{2} \int_0^T \|u^y(\cdot, t) - g(\cdot, t)\|^2 dt + \frac{\alpha_2}{2} \|u^y(\cdot, T) - g(\cdot, T)\|^2 + \frac{\alpha_3}{2} \int_0^T \|z^y(\cdot, t)\|^2 dt \right)$$

[Kouri et al. (2013); Kunoth, Schwab (2013); Ali, Ullmann, Hinze (2017); Chen, Ghattas (2018)]

- **Deterministic control:**

$$J(z) = \mathcal{R} \left(\frac{\alpha_1}{2} \int_0^T \|u^y(\cdot, t) - g(\cdot, t)\|^2 dt + \frac{\alpha_2}{2} \|u^y(\cdot, T) - g(\cdot, T)\|^2 \right) + \frac{\alpha_3}{2} \int_0^T \|z(\cdot, t)\|^2 dt$$

- Taylor approximations [Chen, Villa, Ghattas (2019)], Sparse grids [Kouri (2014); Kouri Surowiec (2016)], Multilevel Monte Carlo [Van Barel, Vandewalle (2019)]

Problem formulation

- Optimal control problem $\min_{z \in \mathcal{Z}} J(z)$, where

$$J(z) = \mathcal{R} \left(\frac{\alpha_1}{2} \int_0^T \|u^y(\cdot, t) - g(\cdot, t)\|^2 dt + \frac{\alpha_2}{2} \|u^y(\cdot, T) - g(\cdot, T)\|^2 \right) + \frac{\alpha_3}{2} \int_0^T \|z(\cdot, t)\|^2 dt$$

subject to

$$\begin{aligned} \partial_t u^y(x, t) - \nabla \cdot (a^y(x) \nabla u^y(x, t)) &= z(x, t) & x \in D, t \in (0, T), y \in U, \\ u^y(x, t) &= 0 & x \in \partial D, t \in (0, T), y \in U, \\ u^y(x, 0) &= u_0(x) & x \in D, y \in U, \end{aligned}$$

- Look for optimal control z in feasible set \mathcal{Z} , where we consider

$$\mathcal{Z} := L^2(D \times (0, T)) \quad (\text{unconstrained})$$

$$\mathcal{Z} := \{z \in L^2(D \times (0, T)) \mid z_{\min}(x, t) \leq z(x, t) \leq z_{\max}(x, t) \text{ a.e.}\} \quad (\text{constrained})$$

- "Uniform" or "affine" model:

$$a^y(x) = \bar{a}(x) + \sum_{j=1}^{\infty} y_j \psi_j(x), \quad x \in D, y \in U := [-\frac{1}{2}, \frac{1}{2}]^{\mathbb{N}}$$

with assumptions

- $0 < a_{\min} \leq a^y(x) \leq a_{\max} < \infty$
- $\sum_{j=1}^{\infty} \|\psi_j\|_{L^\infty}^p < \infty$ for some $p \in (0, 1)$
- $\|\psi_1\|_{L^\infty} \geq \|\psi_2\|_{L^\infty} \geq \dots$

Reduced formulation (robust optimization)

- Optimal control problem $\min_{z \in \mathcal{Z}} J(z)$, with

$$J(z) = \mathcal{R} \left(\underbrace{\frac{\alpha_1}{2} \int_0^T \|u^y(\cdot, t) - g(\cdot, t)\|^2 dt + \frac{\alpha_2}{2} \|u^y(\cdot, T) - g(\cdot, T)\|^2}_{=: \Phi(y)} \right) + \frac{\alpha_3}{2} \int_0^T \|z(\cdot, t)\|^2 dt$$

- u^y is the “state” corresponding to the control z
- \mathcal{Z} is the feasible set, which is bounded, closed, convex, nonempty
- Theorem:** There exists a unique optimal solution z^* .
- The Fréchet derivative of J is given by

$$J'(z) = \mathcal{R}(q) + \alpha_3 z \quad (\text{linear risk measure } \mathcal{R})$$

$$J'(z) = \frac{1}{\int_U \exp(\theta \Phi(y)) dy} \int_U \exp(\theta \Phi(y)) q^y dy + \alpha_3 z \quad (\text{entropic risk measure})$$

where q is the adjoint state.

Theorem: Let \mathcal{R} be the expected value or the entropic risk measure. A control $z = z^*$ is the unique minimizer of the problem if and only if it satisfies the KKT system

$$\left\{ \begin{array}{ll} \partial_t u^y(x, t) - \nabla \cdot (a^y(x) \nabla u^y(x, t)) = z(x) & x \in D, \quad t \in (0, T), \\ u^y(x, t) = 0 & x \in \partial D, \quad t \in (0, T), \\ u^y(x, 0) = u_0(x) & x \in D, \\ -\partial_t q^y(x, t) - \nabla \cdot (a^y(x) \nabla q^y(x, t)) \\ \qquad \qquad \qquad = \alpha_1(u^y(x, t) - g(x, t)) & x \in D, \quad t \in (0, T), \\ q^y(x, t) = 0 & x \in \partial D, \quad t \in (0, T), \\ q^y(x, T) = \alpha_2(u^y(x, T) - g(x, T)) & x \in D, \\ J'(z^*) - \mu_a + \mu_b = 0 & x \in D, \quad t \in (0, T), \\ z_{\min}(x, t) \leq z(x, t) \leq z_{\max}(x, t) & x \in D, \quad t \in (0, T) \\ \mu_a(x, t) \geq 0, \quad \mu_b(x, t) \geq 0 & x \in D, \quad t \in (0, T) \\ (z(x, t) - z_{\min}(x, t))\mu_a(x, t) \\ \qquad \qquad \qquad = (z_{\max}(x, t) - z(x, t))\mu_b(x, t) = 0 & x \in D, \quad t \in (0, T), \end{array} \right.$$

for all $y \in U$. Can be solved, e.g., using projected gradient descent.

Discretization of the problem (robust risk measure $\mathcal{R} = \int_U \cdot \mathrm{d}\mathbf{y}$)

- Optimal control problem:

$$J(z) = \int_U \left(\frac{\alpha_1}{2} \int_0^T \|u^{\textcolor{red}{y}}(\cdot, t) - g(\cdot, t)\|^2 \mathrm{d}t + \frac{\alpha_2}{2} \|u^{\textcolor{red}{y}}(\cdot, T) - g(\cdot, T)\|^2 \right) \mathrm{d}\mathbf{y} + \frac{\alpha_3}{2} \int_0^T \|z(\cdot, t)\|^2 \mathrm{d}t$$
$$J'(z) = \int_U q^{\textcolor{red}{y}} \mathrm{d}\mathbf{y} + \alpha_3 z$$

- Discretized problem:

- Dimension truncation to s terms
- Quasi-Monte Carlo quadrature with n points

$$J_{s,n}(z) = \frac{1}{n} \sum_{i=1}^n \left(\frac{\alpha_1}{2} \int_0^T \|u_s^{\textcolor{blue}{y}^{(i)}}(\cdot, t) - g(\cdot, t)\|^2 \mathrm{d}t \right.$$
$$\left. + \frac{\alpha_2}{2} \|u_s^{\textcolor{blue}{y}^{(i)}}(\cdot, T) - g(\cdot, T)\|^2 \right) + \frac{\alpha_3}{2} \int_0^T \|z(\cdot, t)\|^2 \mathrm{d}t$$

$$J'_{s,n}(z) = \frac{1}{n} \sum_{i=1}^n q_s^{\textcolor{blue}{y}^{(i)}} + \alpha_3 z$$

- Positive (equal) weights of QMC preserve convexity of the problem

Discretization of the problem (entropic risk measure)

- Optimal control problem:

$$J(z) = \frac{1}{\theta} \ln \left(\int_U \exp(\theta \Phi(\mathbf{y})) q^{\mathbf{y}} d\mathbf{y} \right) + \frac{\alpha_3}{2} \int_0^T \|z(\cdot, t)\|^2 dt$$

$$J'(z) = \frac{\int_U \exp(\theta \Phi(\mathbf{y})) q^{\mathbf{y}} d\mathbf{y}}{\int_U \exp(\theta \Phi(\mathbf{y})) d\mathbf{y}} + \alpha_3 z$$

$$\Phi(\mathbf{y}) = \frac{\alpha_1}{2} \int_0^T \|u^{\mathbf{y}}(\cdot, t) - g(\cdot, t)\|^2 dt + \frac{\alpha_2}{2} \|u^{\mathbf{y}}(\cdot, T) - g(\cdot, T)\|^2$$

- Discretized problem:

- Dimension truncation to s terms
- Quasi-Monte Carlo quadrature with n points

$$J_{s,n}(z) = \frac{1}{\theta} \ln \left(\frac{1}{n} \sum_{i=1}^n \exp(\theta \Phi_s(\mathbf{y}^{(i)})) q_s^{\mathbf{y}^{(i)}} \right) + \frac{\alpha_3}{2} \int_0^T \|z(\cdot, t)\|^2 dt$$

$$J'_{s,n}(z) = \frac{\sum_{i=1}^n \exp(\theta \Phi_s(\mathbf{y}^{(i)})) q_s^{\mathbf{y}^{(i)}}}{\sum_{i=1}^n \exp(\theta \Phi_s(\mathbf{y}^{(i)}))} + \alpha_3 z$$

Error analysis

- **Theorem (robust risk measure)** A randomly shifted lattice rule can be constructed using the fast CBC algorithm using POD weights such that

$$\sqrt{\mathbb{E}_\Delta \|z^* - z_{s,n}^*\|_{L^2(D \times (0,T))}^2} = \mathcal{O}(s^{-2/p+1} + n^{-\min\{1/p-1/2, 1-\delta\}}),$$

where the implied coefficient is independent of s .

- **Theorem (entropic risk measure)** A randomly shifted lattice rule can be constructed using the fast CBC algorithm using POD weights such that

$$\sqrt{\mathbb{E}_\Delta \|z^* - z_{s,n}^*\|_{L^2(D \times (0,T))}^2} = \mathcal{O}(s^{-2/p+1} + n^{-\min\{1/p-1/2, 1-\delta\}}),$$

where the implied coefficient is independent of s .

Numerical experiments

Consider the coupled PDE system

$$\begin{cases} \partial_t u^y(\mathbf{x}, t) - \nabla \cdot (a^y(\mathbf{x}) \nabla u^y(\mathbf{x}, t)) = z(\mathbf{x}) & \mathbf{x} \in D, t \in (0, T), \\ u^y(\mathbf{x}, t) = 0 & \mathbf{x} \in \partial D, t \in (0, T), \\ u^y(\mathbf{x}, 0) = u_0(\mathbf{x}) & \mathbf{x} \in D, \\ -\partial_t q^y(\mathbf{x}, t) - \nabla \cdot (a^y(\mathbf{x}) \nabla q^y(\mathbf{x}, t)) \\ = \alpha_1(u^y(\mathbf{x}, t) - g(\mathbf{x})) & \mathbf{x} \in D, t \in (0, T), \\ q^y(\mathbf{x}, t) = 0 & \mathbf{x} \in \partial D, t \in (0, T) \\ q^y(\mathbf{x}, T) = \alpha_2(u^y(\mathbf{x}, T) - g(\mathbf{x})) & \mathbf{x} \in D, t \in (0, T), \end{cases}$$

for $\mathbf{y} \in [-\frac{1}{2}, \frac{1}{2}]^s$ in $D = (0, 1)^2$, equipped with

$$a(\mathbf{x}, \mathbf{y}) = 1 + c \sum_{j=1}^{\infty} y_j \underbrace{\frac{1}{(k_j^2 + \ell_j^2)^{\vartheta}} \sin(k_j \pi x_1) \sin(\ell_j \pi x_2)}_{=\psi_j},$$

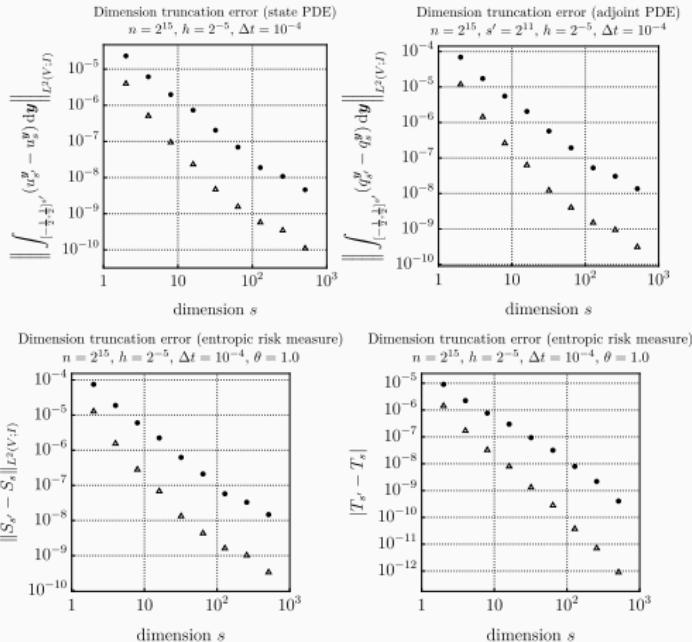
where $(k_j, \ell_j)_{j \geq 1}$ are ordered s.t. $(\|\psi_j\|_{L^\infty(D)})_{j \geq 1}$ is non-increasing.

Weyl asymptotics: $\|\psi_j\|_{L^\infty(D)} \sim j^{-\vartheta}$ as $j \rightarrow \infty$. We fix $z(\mathbf{x}) = x_2$.

Numerical experiments

Dimension truncation error

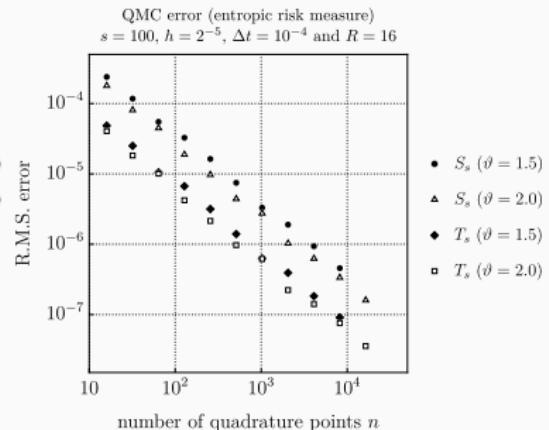
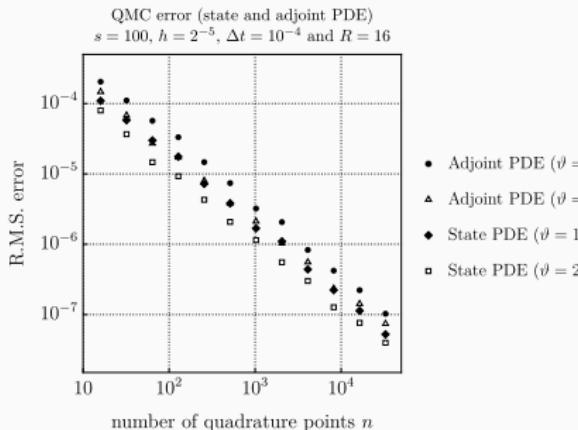
We use $s = 2^{11}$ as the reference solution.



Numerical experiments

QMC error

We use $R = 16$ random shifts Δ drawn from $U([0, 1]^s)$.



LHS: QMC errors corresponding to $\int_{[-\frac{1}{2}, \frac{1}{2}]^{100}} u^y dy$ (state PDE) and $\int_{[-\frac{1}{2}, \frac{1}{2}]^{100}} q^y dy$ (adjoint PDE).

RHS: QMC errors corresponding to $S_{100} = \int_{[-\frac{1}{2}, \frac{1}{2}]^{100}} \exp(\Phi_s(y)) q^y dy$ and

$$T_{100} = \int_{[-\frac{1}{2}, \frac{1}{2}]^{100}} \exp(\Phi_s(y)) dy$$

Numerical experiments

Optimal control problem – unconstrained gradient descent $\min_{z \in L^2(D \times (0, T))} J(z)$,

$$J(z, u) = \mathcal{R}_y \left(\frac{\alpha_1}{2} \int_0^T \|u^y(\cdot, t) - g\|_{L^2(D \times (0, T))}^2 dt + \frac{\alpha_2}{2} \|u^y(\cdot, T) - g(\cdot, T)\|_{L^2(D)}^2 \right) + \frac{\alpha_3}{2} \|z\|_{L^2(D \times (0, T))}^2$$

We fix $\vartheta = 1.2$, $h = 2^{-4}$, $\Delta t = 10^{-2}$, $n = 2^{10}$, $s = 100$, $T = 1.0$, $u_0(x) = \sin(2\pi x_1) \sin(2\pi x_2)$, $\alpha_1 = 1$, $\alpha_2 = \alpha_3 = 10^{-4}$, and $\theta = 10^5$. As the diffusion coefficient, we take

$$a(x, y) = 100 + 55 \sum_{j=1}^{100} y_j j^{-1.2} \sin(j\pi x_1) \sin(j\pi x_2).$$

As the target, we choose

$$g(x, t) := \chi_{\|x - (c_1(t), c_2(t))\|_\infty \leq \frac{1}{10}}(x) g_1(x, t) + \chi_{\|x + (c_1(t), c_2(t)) - (1, 1)\|_\infty \leq \frac{1}{10}}(x) g_2(x, t),$$

where $g_1(x, t) := (4\sqrt{10})^4 (x_1 - c_1(t) - \frac{1}{10})(x_2 - c_2(t) - \frac{1}{10})(x_1 - c_1(t) + \frac{1}{10})(x_2 - c_2(t) + \frac{1}{10})$,
 $g_2(x, t) := (4\sqrt{10})^4 (x_1 + c_1(t) - \frac{11}{10})(x_2 + c_2(t) - \frac{11}{10})(x_1 + c_1(t) - \frac{9}{10})(x_2 + c_2(t) - \frac{9}{10})$,
 $c_1(t) := \frac{1}{2} + \frac{1}{4}(1 - t^{10}) \cos(4\pi t^2)$, and $c_2(t) := \frac{1}{2} + \frac{1}{4}(1 - t^{10}) \sin(4\pi t^2)$.

Unconstrained optimization

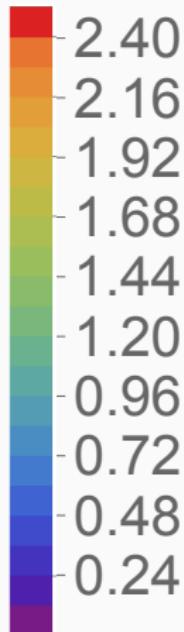


Figure 3: The reconstructed optimal control z^* (without box constraints).

Reference

-  P. A. Guth, K., F. Y. Kuo, C. Schillings, I. H. Sloan. Parabolic PDE-constrained optimal control under uncertainty with entropic risk measure using quasi-Monte Carlo integration. Preprint 2022, arXiv:2208.02767 [math.NA].

Part III: Toward Bayesian optimal experimental design

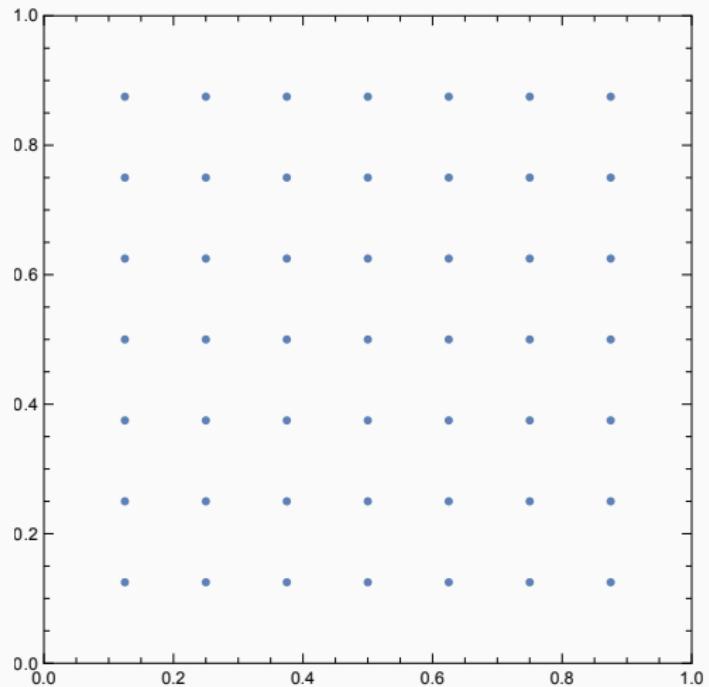
Let $G: \Theta \times \Xi \rightarrow \mathbb{R}^k$ be a forward mapping depending on a true parameter $\theta \in \Theta$ and a design parameter $\xi \in \Xi$.

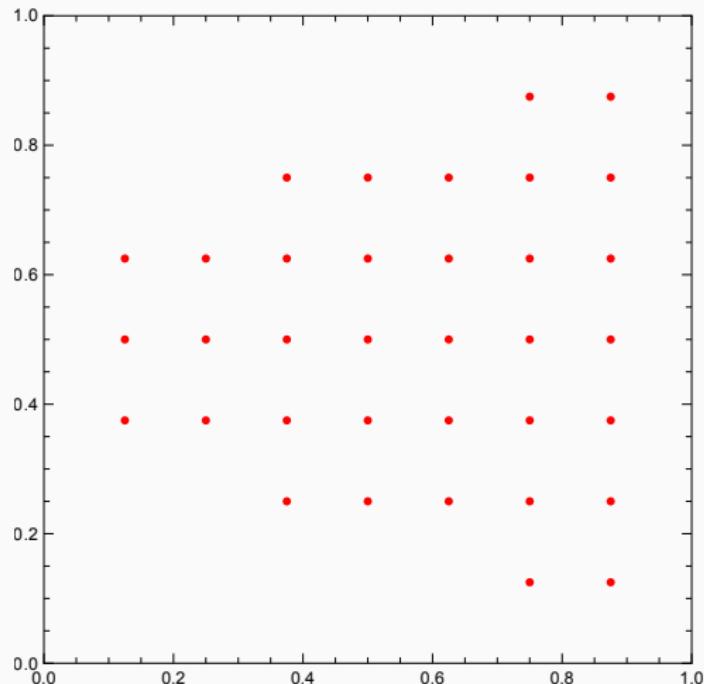
Measurement model:

$$\mathbf{y} = G(\theta, \xi) + \boldsymbol{\eta},$$

where $\mathbf{y} \in \mathbb{R}^k$ is the measurement data and $\boldsymbol{\eta} \in \mathbb{R}^k$ is Gaussian noise such that $\boldsymbol{\eta} \sim \mathcal{N}(0, \Gamma)$ with positive definite covariance matrix $\Gamma \in \mathbb{R}^{k \times k}$.

Goal in Bayesian optimal experimental design: Recover the design parameter ξ for the Bayesian inference of θ , which we model as a random variable endowed with prior distribution $\pi(\theta)$.





A measure of the information gain for a given design ξ and data y is given by the Kullback–Leibler divergence

$$D_{\text{KL}}(\pi(\cdot|y, \xi) \| \pi(\cdot)) := \int_{\Theta} \log \left(\frac{\pi(\theta|y, \xi)}{\pi(\theta)} \right) \pi(\theta|y, \xi) d\theta. \quad (2)$$

A Bayesian optimal design ξ^* maximizing the expected utility (2) over the design space Ξ with respect to the data y and model parameters θ is given by

$$\xi^* = \arg \max_{\xi \in \Xi} \underbrace{\int_Y \int_{\Theta} \log \left(\frac{\pi(\theta|y, \xi)}{\pi(\theta)} \right) \pi(\theta|y, \xi) \pi(y|\xi) d\theta dy}_{=: \text{EIG}}$$

where $\pi(\theta|y, \xi)$ corresponds to the posterior distribution of the parameter θ and $\pi(y|\xi) = \int \pi(y|\theta, \xi) \pi(\theta) d\theta$ is the marginal distribution of the data y .

The posterior is given by Bayes' theorem

$$\pi(\boldsymbol{\theta}|\mathbf{y}, \boldsymbol{\xi}) = \frac{\pi(\mathbf{y}|\boldsymbol{\theta}, \boldsymbol{\xi})\pi(\boldsymbol{\theta})}{\pi(\mathbf{y}|\boldsymbol{\xi})},$$

which means that the expected utility can be written as

$$\begin{aligned}\text{EIG} &= \int_Y \int_{\Theta} \log \left(\frac{\pi(\boldsymbol{\theta}|\mathbf{y}, \boldsymbol{\xi})}{\pi(\boldsymbol{\theta})} \right) \pi(\boldsymbol{\theta}|\mathbf{y}, \boldsymbol{\xi}) d\boldsymbol{\theta} \pi(\mathbf{y}|\boldsymbol{\xi}) d\mathbf{y} \\ &= \int_{\Theta} \left[\int_Y \log \left(\frac{\pi(\mathbf{y}|\boldsymbol{\theta}, \boldsymbol{\xi})}{\pi(\mathbf{y}|\boldsymbol{\xi})} \right) \pi(\mathbf{y}|\boldsymbol{\theta}, \boldsymbol{\xi}) d\mathbf{y} \right] \pi(\boldsymbol{\theta}) d\boldsymbol{\theta}.\end{aligned}$$

Approaches taken in the literature:

- Double-loop Monte Carlo (Beck, Mansour, Espath, Long, Tempone)
- MCLA (Beck, Mansour, Espath, Long, Tempone)
- DLMCIS (Beck, Mansour, Espath, Long, Tempone)

Why QMC might work

Assume, e.g., the following:

- The forward model $\mathbf{y} = G(\boldsymbol{\theta}, \boldsymbol{\xi}) + \boldsymbol{\eta}$ satisfies

$$|\partial_{\boldsymbol{\theta}}^{\boldsymbol{\nu}} G_j(\boldsymbol{\theta}, \boldsymbol{\xi})| \leq C_0 |\boldsymbol{\nu}|! \mathbf{b}^{\boldsymbol{\nu}}, \quad j \in \{1, \dots, k\},$$

where $\mathbf{b} := (b_j)_{j \geq 1} \in \ell^p$ for some $p \in (0, 1)$ and $C_0 > 0$ is independent of $\boldsymbol{\xi}$.

- $\Theta = [-\frac{1}{2}, \frac{1}{2}]^s$ and $\pi(\boldsymbol{\theta}) = 1$ for $\boldsymbol{\theta} \in \Theta$ and 0 otherwise.
- We have the likelihood

$$\pi(\mathbf{y} | \boldsymbol{\theta}, \boldsymbol{\xi}) = C e^{-\frac{1}{2} \|\mathbf{y} - G(\boldsymbol{\theta}, \boldsymbol{\xi})\|_{\Gamma^{-1}}^2}, \quad C = \frac{1}{(2\pi)^{k/2} \sqrt{\det \Gamma}}.$$

Under these conditions, it is easy to see that

$$\begin{aligned} \text{EIG} &= \int_Y \int_{\Theta} \log \left(\frac{\pi(\mathbf{y} | \boldsymbol{\theta}, \boldsymbol{\xi})}{\pi(\mathbf{y} | \boldsymbol{\xi})} \right) \pi(\mathbf{y} | \boldsymbol{\theta}, \boldsymbol{\xi}) \pi(\boldsymbol{\theta}) d\boldsymbol{\theta} d\mathbf{y} \\ &= \log C - 1 - \int_Y \log \left(\int_{\Theta} C e^{-\frac{1}{2} \|\mathbf{y} - G(\boldsymbol{\theta}, \boldsymbol{\xi})\|_{\Gamma^{-1}}^2} d\boldsymbol{\theta} \right) \int_{\Theta} C e^{-\frac{1}{2} \|\mathbf{y} - G(\boldsymbol{\theta}, \boldsymbol{\xi})\|_{\Gamma^{-1}}^2} d\boldsymbol{\theta} d\mathbf{y}. \end{aligned}$$

Consider

$$\int_Y \log \left(\int_{\Theta} C e^{-\frac{1}{2} \|y - G(\theta, \xi)\|_{r-1}^2} d\theta \right) \int_{\Theta} C e^{-\frac{1}{2} \|y - G(\theta, \xi)\|_{r-1}^2} d\theta dy \quad (3)$$

Observations:

- Parametric regularity of the integrand

$$\int_{\Theta} C e^{-\frac{1}{2} \|y - G(\theta, \xi)\|_{r-1}^2} d\theta$$

well-understood (as long as the parametric regularity of G can be quantified); straightforward modification of Herrmann–Keller–Schwab (2021).

- If we replace $\int_{\Theta} C e^{-\frac{1}{2} \|y - G(\theta, \xi)\|_{r-1}^2} d\theta$ by a QMC approximation

$$\int_{\Theta} C e^{-\frac{1}{2} \|y - G(\theta, \xi)\|_{r-1}^2} d\theta \approx \frac{C}{n} \sum_{i=1}^n e^{-\frac{1}{2} \|y - G(\theta_i, \xi)\|_{r-1}^2}$$

then (3) can be approximated by

$$\int_Y \log \left(\frac{C}{n} \sum_{i=1}^n e^{-\frac{1}{2} \|y - G(\theta_i, \xi)\|_{r-1}^2} \right) \frac{C}{n} \sum_{i=1}^n e^{-\frac{1}{2} \|y - G(\theta_i, \xi)\|_{r-1}^2} dy.$$

Lemma

Under the previous conditions, we have the a priori bound for the higher-order derivatives

$$\begin{aligned} & \partial_{\mathbf{y}}^{\boldsymbol{\nu}} \log \left(C \frac{1}{n} \sum_{j=1}^n e^{-\frac{1}{2} \| \mathbf{y} - G(\boldsymbol{\theta}_j, \boldsymbol{\xi}) \|_{\Gamma^{-1}}^2} \right) \\ & \leq 1.1^{k|\boldsymbol{\nu}|} k^{|\boldsymbol{\nu}|} \mu_{\min}^{-|\boldsymbol{\nu}|/2} \sum_{\lambda=1}^{|\boldsymbol{\nu}|} (\lambda-1)! \left(\sum_{j=1}^n e^{-\frac{1}{2} \| \mathbf{y} - G(\boldsymbol{\theta}_j, \boldsymbol{\xi}) \|_{\Gamma^{-1}}^2} \right)^{-\lambda} S(|\boldsymbol{\nu}|, \lambda) \end{aligned}$$

for all $\mathbf{y} \in \mathbb{R}^k =: Y$ and $\boldsymbol{\nu} \in \mathbb{N}_0^k$. Here μ_{\min} is a lower bound for the largest eigenvalue of Γ and $S(n, k)$ denotes the Stirling number of the second kind.

Numerical example

Let $D = (0, 1)^2$ and consider the parametric PDE

$$\begin{cases} -\nabla \cdot (a(\mathbf{x}, \boldsymbol{\theta}) \nabla u(\mathbf{x}, \boldsymbol{\theta})) = x_1, & \mathbf{x} \in D, \boldsymbol{\theta} \in \Theta, \\ u(\mathbf{x}, \boldsymbol{\theta}) = 0, & \mathbf{x} \in \partial D, \boldsymbol{\theta} \in \Theta \end{cases}$$

equipped with an uncertain diffusion coefficient

$$a(\mathbf{x}, \boldsymbol{\theta}) = 1 + \sum_{j=1}^s j^{-2} \theta_j \sin(j\pi x_1) \sin(j\pi x_2), \quad \mathbf{x} \in D, \boldsymbol{\theta} \in \Theta.$$

Experiment 1: QMC convergence of the inner integral

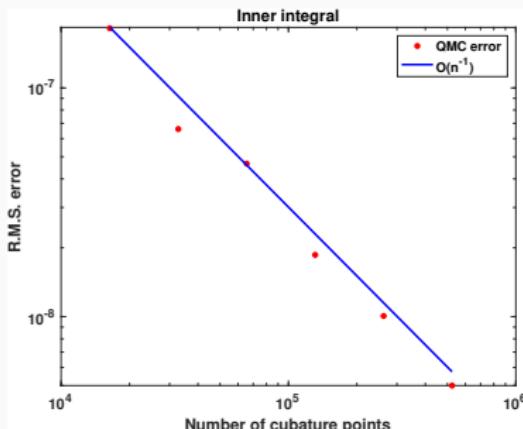
First we consider the QMC convergence of the “inner integral”

$$\int_{\Theta} C e^{-\frac{1}{2} \|\mathbf{y} - G(\boldsymbol{\theta}, \xi)\|_{\Gamma^{-1}}^2} d\boldsymbol{\theta}.$$

Set $\xi := (\frac{5}{8}, \frac{7}{8})$, $\Gamma := 0.2^2 I$, and with $k = 1$ consider

$$G(\xi, \theta) = u(\xi, \theta).$$

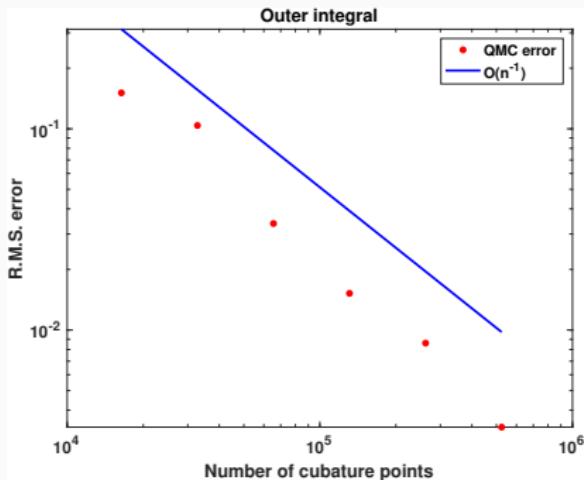
Finally, we use $\mathbf{y} = y = 0.396586$, and $s = 100$. We use an off-the-shelf cubature rule to compute QMC approximations of this integral.



Experiment 2: QMC convergence of the outer integral

$$\int_Y \log \left(\frac{C}{n} \sum_{i=1}^n e^{-\frac{1}{2} \|y - G(\theta_i, \xi)\|_{\Gamma^{-1}}^2} \right) \frac{C}{n} \sum_{i=1}^n e^{-\frac{1}{2} \|y - G(\theta_i, \xi)\|_{\Gamma^{-1}}^2} dy.$$

Fix QMC discretization of the inner integral ($n = 32$). Also, take $Y = [-1/2, 1/2]^{10}$ and fix design ξ s.t. the observation operator corresponds to $k = 10$ point evaluations within the computational domain.



- Use Laplace approximation for the integral over the data?
- In addition to discrete optimal design problems, continuous design problems? Consider the input current as a (continuous) design variable for EIT?