

# Applications of Chatterjee's Correlation in MCMC

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## Definition

A Markov chain is a discrete time stochastic process  $X_1, X_2, \dots$  such that the next state of the process depends solely on the present state, i.e.

$$\Pr(X_{n+1}|X_1, \dots, X_n) = \Pr(X_{n+1}|X_n).$$

- A Markov chain can be specified by two things
  - 1 The initial distribution, i.e. the marginal of  $X_1$ .
  - 2 The transition probabilities, i.e. the conditional distribution of  $X_{n+1}|X_n$



# Markov chain Monte Carlo

## Definition

A Markov chain is said to be **stationary** if the marginal of  $X_n$  is independent of  $n$ . This invariant distribution is called the stationary distribution.

## Definition

A Markov chain is **ergodic** if the distribution of  $X_n$  converges to the invariant distribution.

## Definition

A stationary Markov chain is **time-reversible** w.r.t. the stationary distribution if  $X_n$  and  $X_{n+1}$  are exchangeable.



# Pearson's Correlation Coefficient

- Pearson's Correlation Coefficient measures the linear correlation between two sets of data.

## Definition

Given random variables  $X, Y$ , Pearson Correlation  $\rho$  is defined as

$$\rho = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X)} \cdot \sqrt{\text{Var}(Y)}}.$$

- Pearson Correlation is very powerful in detecting monotone relations and has a well developed asymptotic theory.



# Problems with Pearson Correlation

- There are some problems with the Pearson Correlation.
- We would like the correlation to be close to its maximum value iff one variable is a function of the other. In Pearson's case, it is equal to  $\pm 1$  iff the variables are linearly dependent.
- We would also like the correlation to be 0 iff the variables are independent. If the variables are independent, Pearson correlation is indeed 0, but the converse is not always true.



# Chatterjee's Correlation Coefficient

- In order to solve these problems, Chatterjee came up with a new measure of dependence in [?]. It overcomes the above mentioned drawbacks, and has a computationally efficient and consistent estimator.

## Definition

Given random variables  $X, Y$ , where  $Y$  is not a constant, Chatterjee correlation  $\xi$  is defined as

$$\xi(X, Y) = \frac{\int \text{Var}(\mathbb{E}(1_{\{Y \geq t\}} | X)) d\mu(t)}{\int \text{Var}(1_{\{Y \geq t\}}) d\mu(t)},$$

where  $\mu$  is the law of  $Y$ .



# Consistent Estimator of $\xi$

- Let  $\{(X_i, Y_i)\}_{i=1}^n$  be i.i.d. pairs following the same distribution as  $(X, Y)$ . Rearrange the data as  $(X_{(1)}, Y_{(1)}), \dots, (X_{(n)}, Y_{(n)})$  such that  $X_{(1)} < \dots < X_{(n)}$ . Let  $r_i$  be the rank of  $Y_{(i)}$ , i.e. the number of  $j$  such that  $Y_{(j)} \leq Y_{(i)}$ . Then the correlation coefficient  $\xi_n$  is defined to be

$$\xi_n(X, Y) := 1 - \frac{3 \sum_{i=1}^{n-1} |r_{i+1} - r_i|}{n^2 - 1}.$$

## Theorem

*If  $Y$  is not almost surely a constant, then as  $n \rightarrow \infty$ ,  $\xi_n(X, Y)$  converges almost surely to  $\xi(X, Y)$ .*





# Properties of $\xi$

- $\xi(X, Y) \in [0, 1]$
- $\xi(X, Y) = 0$  if and only if  $X$  and  $Y$  are independent.
- $\xi(X, Y) = 1$  if and only if at least one of  $X$  and  $Y$  is a measurable function of the other.
- $\xi$  is not symmetric in  $X, Y$ . This is intentional and useful as we might want to study if  $Y$  is a measurable function of  $X$ , or  $X$  is a measurable function of  $Y$ . To get a symmetric coefficient, it suffices to consider  $\max(\xi(X, Y), \xi(Y, X))$ .
- $\xi_n$  is based on ranks, and for the same reason, it can be computed in  $O(n \log n)$ .



# Chatterjee's Autocorrelation Function

- We present a new autocorrelation function (ACF) using Chatterjee's correlation coefficient.

## Definition

Let  $X_1, X_2, \dots$  be a stationary, time homogeneous Markov chain with stationary distribution  $\pi$ . We define the new lag- $k$  ACF as follows

$$\gamma'_{k,n} := \xi(X_n, X_{n+k}).$$



$\gamma'_{k,n}$  is independent of  $n$

### Theorem

$\gamma_k = \text{Cov}(X_n, X_{n+k})$  is independent of  $n$ .

### Theorem

$\gamma'_{k,n} = \xi(X_n, X_{n+k})$  is independent of  $n$ , where  $n$  and  $k$  are in  $\mathbb{N}$ .

- From now on, we'll denote  $\gamma'_{k,n}$  by  $\gamma'_k$



## Theorem

$\xi(X_n, X_{n+k}) = \xi(X_{n+k}, X_n)$  for time reversal Markov chains for any  $n, k \in \mathbb{N}$ .



# Convergence of $\gamma'_k$

## Theorem

$\lim_{n \rightarrow \infty} \xi(X_1, X_n) = 0$  for an Ergodic Markov chain.



# Proof of consistency of estimator

- Chatterjee presented the proof of consistency of the estimator in [?], where the samples drawn from  $(X, Y)$  are i.i.d.
- In our case of stationary Markov chains, we have correlated but identically distributed draws.
- We aim to prove the consistency of the estimator for our case as well.



# Proof of consistency of estimator

## Theorem

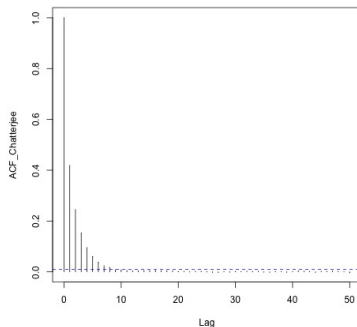
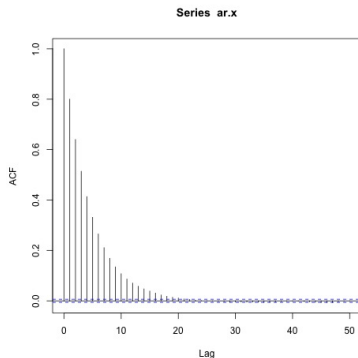
*Let  $X_1, X_2, \dots$  be a stationary time-homogeneous Markov chain with stationary distribution  $\mu$ . Then  $\xi_n(X, Y)$  estimated using the draws from the Markov chain converge to  $\xi(X, Y)$  as  $n \rightarrow \infty$ , where  $X, Y$  are any two time points in the chain.*

- We presented some ideas and a pathway for proving it but could not complete it and is left as future work.



# Simulations

- AR(1) process with  $\rho = 0.8$ .





# Simulations

- Metropolis-Hastings algorithm with initial distribution as  $\text{Exp}(0.01)$  and target distribution  $\mathcal{N}(0, 1)$ .

