

EE901 PROBABILITY AND RANDOM PROCESSES

MODULE 4 EXPECTATION AND MOMENTS

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28

Expectation

29

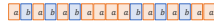
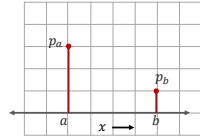
Expectation

- Random variables take different values each time the experiment is performed.
- On the average, what value can we expect?
- We call it mean/average/expected value. How to calculate it?
- Suppose random variable takes only one value, then the mean would be this value only.
- Suppose random variables takes two value a and b , with equal probability. The mean would be $(a + b)/2$.
- What if these values don't have equal chances?

30

Expectation

- Consider a RV X which takes two values a and b with probability p_a and p_b
- If the experiment is repeated N times,
- Approximately
 - $N_a = N p_a$ times, the outcome is a
 - $N_b = N p_b$ times, the outcome is b .



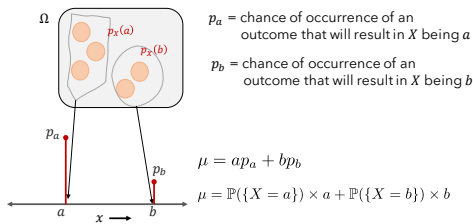
Average is

$$\frac{a + a + \dots + N_a \text{ times } a + b + b + \dots + N_b \text{ times } b}{N}$$

$$= \frac{aN_a + bN_b}{N} = \frac{Np_a a + Np_b b}{N} = ap_a + bp_b$$

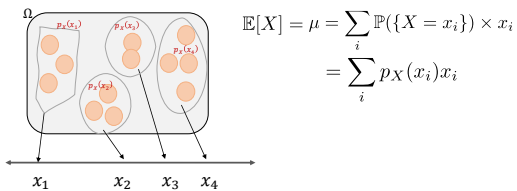
31

Expectation



32

Expectation of Discrete Random Variable



33

Examples

- Bernoulli RV with parameter p

$$\begin{array}{c}
 | \quad | \\
 1-p \quad p \\
 \hline
 0 \quad 1
 \end{array}$$

$$E[X] = 0 \times (1-p) + 1 \times p = p$$

34

Examples

- Dice roll

$$\begin{array}{c}
 | \quad | \quad | \quad | \quad | \quad | \\
 1 \quad 2 \quad 3 \quad 4 \quad 5 \quad 6
 \end{array}
 \quad \frac{1}{6}$$

$$\begin{aligned}
 &1 \times \frac{1}{6} + 2 \times \frac{1}{6} + 3 \times \frac{1}{6} + 4 \times \frac{1}{6} \\
 &\quad + 5 \times \frac{1}{6} + 6 \times \frac{1}{6} \\
 &= \frac{1}{6} [1 + 2 + 3 + 4 + 5 + 6] \\
 &= 3.5
 \end{aligned}$$

35

Examples

- Poisson RV with parameter λ

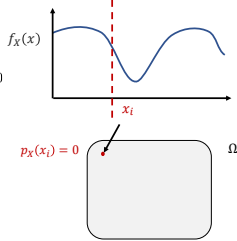
$$\begin{aligned}
 E[X] &= \sum_{x=0}^{\infty} P_X(x) \cdot x \\
 &= \sum_{x=0}^{\infty} \frac{e^{-\lambda} \lambda^x}{x!} \cdot x = e^{-\lambda} \sum_{x=0}^{\infty} \frac{\lambda \cdot \lambda^{x-1}}{(x-1)!} \\
 &= \sum_{k=0}^{\infty} \frac{e^{-\lambda} \cdot \lambda}{(k)!} = e^{-\lambda} \cdot \lambda \cdot e^{\lambda} = \lambda
 \end{aligned}$$

36

Expectation for Continuous Random Variable

$$\mathbb{E}[X] = \sum_i \mathbb{P}(\{X = x_i\})x_i$$

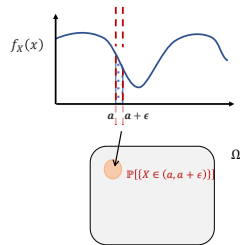
For continuous RV, $\mathbb{P}(\{X = x_i\}) = 0$



37

Expectation for Continuous Random Variable

$$\mathbb{P}[\{X \in (a, a + \epsilon)\}] = f_X(a) \epsilon$$

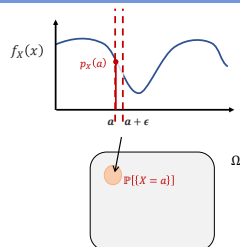


38

Expectation for Continuous Random Variable

$$p_X(a) =$$

$$\mathbb{P}[\{X \in (a, a + \epsilon)\}] = f_X(a) \epsilon$$



39

Expectation for Continuous Random Variable

$$p_X(a_i) =$$

$$\mathbb{P}[\{X \in (a_i, a_i + \epsilon)\}] = f_X(a_i) \epsilon \quad f_X(x)$$

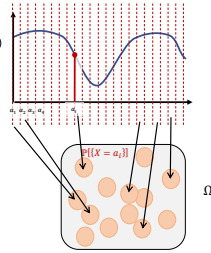
Divide the complete range into ϵ length intervals

$$a_1 = a_1$$

$$a_2 = a_1 + \epsilon$$

$$a_3 = a_2 + \epsilon = a_1 + 2\epsilon$$

$$a_i = a_{i-1} + \epsilon = a_1 + (i-1)\epsilon$$



40

Expectation for Continuous Random Variable

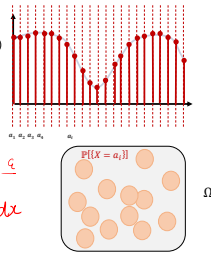
$$p_X(a_i) =$$

$$\mathbb{P}[\{X \in (a_i, a_i + \epsilon)\}] = f_X(a_i) \epsilon \quad f_X(x)$$

$$\mathbb{E}[X] = \sum_i \mathbb{P}(\{X = x_i\}) x_i$$

$$= \sum_i f_X(a_i) \epsilon a_i$$

$$\text{As } \epsilon \rightarrow 0, \quad \mathbb{E}[X] = \int_{-\infty}^{\infty} f_X(a) a \, da = \int_{-\infty}^{\infty} f_X(x) x \, dx$$



41

Examples

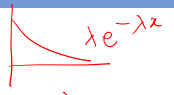
- Uniform random variable (a, b)

$$\begin{aligned} \int_a^b x \cdot f_X(x) \, dx &= \int_a^b x \cdot \frac{1}{b-a} \, dx \\ &= \frac{1}{b-a} \left(\frac{b^2}{2} - \frac{a^2}{2} \right) = \frac{a+b}{2} \end{aligned}$$

42

Examples

- Exponential random variable (λ)

$$\begin{aligned}
 & \int_0^{\infty} x f_X(x) dx \\
 & \Rightarrow \int_0^{\infty} x \cdot \lambda e^{-\lambda x} dx \\
 & = \int_0^{\infty} \frac{t}{\lambda} \cdot \lambda e^{-t} \frac{dt}{\lambda} = \frac{1}{\lambda} \int_0^{\infty} t e^{-t} dt = \frac{1}{\lambda}
 \end{aligned}$$


43

Expectation of Function of RV

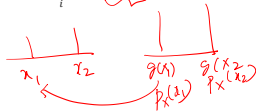
$$\mathbb{E}[X] = \sum_i x_i p_X(x_i)$$

$$\mathbb{E}[X] = \int x f_X(x) dx$$

Let g be some function

$$\mathbb{E}[g(X)] = \sum_i g(x_i) p_X(x_i)$$

$$\mathbb{E}[g(X)] = \int g(x) f_X(x) dx$$



44

Linearity

$$\mathbb{E}[aX + b] = a\mathbb{E}[X] + b$$

$$g(x) = ax + b$$

$$\begin{aligned}
 & \mathbb{E}[g(X)] \\
 & = \sum_{x_i} g(x_i) p_X(x_i) \\
 & = \sum_{x_i} (ax_i + b) p_X(x_i) \\
 & = a \sum_{x_i} x_i p_X(x_i) + b \sum_{x_i} p_X(x_i) \\
 & = a\mathbb{E}[X] + b
 \end{aligned}$$

45

Expectation and CCDF

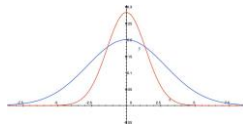
Let X be a RV taking positive values only, then $\mathbb{E}[X] = \int_0^{\infty} \mathbb{P}(\{X > x\}) dx$

$$\begin{aligned} \int_0^{\infty} \underbrace{\mathbb{P}(\{X > x\})}_{\int_x^{\infty} f_X(t) dt} dx &= \int_0^{\infty} \int_x^{\infty} f_X(t) dt dx \\ &\stackrel{\text{Fubini}}{=} \int_0^{\infty} \int_0^t f_X(t) dx dt = \int_0^{\infty} f_X(t) t dt = \mathbb{E}[X] \end{aligned}$$

46

Moments

- The n th moment of a RV X is defined as $\mu_n = \mathbb{E}[X^n]$.
- In particular,
 - $\mathbb{E}[X]$ is the first moment.
 - $\mathbb{E}[X^2]$ is the second moment.
- For some random variables, first moment is enough to fix the distribution.
 - Example: Exponential distribution
- However, for others, first moment is not enough.



47

Central Moments

- The n th central moment of a RV X is defined as

$$v_n = \mathbb{E}[(X - \mathbb{E}[X])^n]$$

- The variance for a random variable X

$$\text{Var}(X) = v_2 = \mathbb{E}[(X - \mathbb{E}[X])^2]$$

48

Variance

- Variance represents the spread of a random variable
- It shows the mean value of square deviation of a RV from its mean.
- For a constant random variable, variance is 0.
- Consider the following two RVs

$$p_X(x) = \begin{cases} \frac{1}{4} & \text{if } x = -1 \\ \frac{1}{2} & \text{if } x = 0 \\ \frac{1}{4} & \text{if } x = 1 \end{cases} \quad p_X(x) = \begin{cases} \frac{1}{4} & \text{if } x = -2 \\ \frac{1}{2} & \text{if } x = 0 \\ \frac{1}{4} & \text{if } x = 2 \end{cases}$$

- Both have mean 0. First has $\frac{1}{2}$ variance and second has 2 variance.

49

Properties of Variance

$$\text{Var}(X) = \mathbb{E}[(X - \mathbb{E}[X])^2] = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$$

Proof:

$$\begin{aligned} \text{Var}(X) &= \mathbb{E}[(X - \mathbb{E}[X])^2] \\ &= \mathbb{E}[X^2 + (\mathbb{E}[X])^2 - 2X\mathbb{E}[X]] \\ &= \mathbb{E}[X^2] + (\mathbb{E}[X])^2 - 2\mathbb{E}[X]\mathbb{E}[X] \\ &= \mathbb{E}[X^2] - (\mathbb{E}[X])^2 \end{aligned}$$

$$\begin{aligned} &\mathbb{E}[\underbrace{(\mathbb{E}[X])^2}_{(\mathbb{E}[X])^2}] \\ &\mathbb{E}[X \times \mathbb{E}[X]] \\ &= \mathbb{E}[X] \cdot \mathbb{E}[X] \end{aligned}$$

50

Properties of Variance

- Let $\text{Var}(X) = \sigma^2$. What will the variance of aX be?
- Let

$$\begin{aligned} Y &= aX, \\ \mathbb{E}[Y] &= a\mathbb{E}[X] \\ \text{Var}(Y) &= \mathbb{E}[(Y - \mathbb{E}[Y])^2] = \mathbb{E}[Y^2] - \mathbb{E}[Y]^2 \\ &= \mathbb{E}[a^2X^2] - \mathbb{E}[aX]^2 \\ &= a^2\mathbb{E}[X^2] - a^2\mathbb{E}[X]^2 \\ &= a^2\sigma^2 \end{aligned}$$

51

Properties of Variance

- Let $\text{Var}(X) = \sigma^2$. What will the variance of $X + b$?
- Let

$$\begin{aligned} Y &= X + b, \\ \mathbb{E}[Y] &= \mathbb{E}[X] + b \\ \text{Var}(Y) &= \mathbb{E}[(Y - \mathbb{E}[Y])^2] \end{aligned}$$

52

Properties of Variance

- Let $\text{Var}(X) = \sigma^2$, $\text{Var}(aX + b) = a^2 \text{Var}(X)$

53

Moment Generating Function

54

Moment Generating Function

- The moment generating function (MGF) of a random variable is an alternative characterization in the place of its probability distribution.
- Instead of directly working with PDF/PMF and CDF, many analytical results can be computed easily with MGF.
- If MGFs of two random variables are the same, they have the same distribution.

55

MGF

- The MGF of a random variable is given as:

$$M_X(t) = \mathbb{E}[e^{tX}] \approx \mathbb{E}[e^{t^X}]$$

- For DRV

$$M_X(t) = \mathbb{E}[e^{tX}] = \sum_x e^{tx} p_X(x)$$

- For CRV

$$M_X(t) = \mathbb{E}[e^{tX}] = \int_{\mathbb{R}} e^{tx} f_X(x) dx$$

56

MGF of Exponential RV

- Example: Let X is an exponential random variable with parameter λ . The density function of X is given as

$$M_X(t) = \int_{\mathbb{R}} e^{tx} f_X(x) dx$$

$$= \int_0^{\infty} e^{tx} \lambda e^{-\lambda x} dx = \frac{1}{1 - \frac{t}{\lambda}}$$

$$f_X(x) = \lambda e^{-\lambda x} 1(x \geq 0)$$

- If we know that X is a random variable whose $M_X(t) = \frac{1}{1 - \frac{t}{2}}$, then we can say that X is an exponential random variable with parameter 2.

57

MGF of Few Distributions

Bernoulli (p): $M_X(t) = 1 - p + pe^t$

Uniform (a, b):
 $M_X(t) = \frac{\exp(tb) - \exp(ta)}{t(b-a)}$

58

Moments and MGF

- As its name implies, the moment generating function can be used to compute a distribution's moments.
- The n th moment of a RV X is defined as

$$\mu_n = \mathbb{E}[X^n].$$

- The n th moment is equal to the n th derivative of the moment-generating function, evaluated at 0:

$$\mu_n = \underline{M_X^{(n)}}(0). \quad \overset{(1)}{f} \quad \overset{(2)}{f}$$

59

Moments and MGF

$$M_X(t) = \mathbb{E}[e^{tx}] = \int_{\mathbb{R}} e^{tx} f_X(x) dx$$

- Differentiating with respect to t ,

$$\begin{aligned} M_X^{(1)}(t) &= \frac{d}{dt} \mathbb{E}[e^{tx}] = \frac{d}{dt} \int_{\mathbb{R}} e^{tx} f_X(x) dx \\ &= \int_{\mathbb{R}} \frac{d}{dt} e^{tx} f_X(x) dx \\ &= \int_{\mathbb{R}} x e^{tx} f_X(x) dx \end{aligned}$$

- Put $t = 0$

$$M_X^{(1)}(0) = \int_{\mathbb{R}} x f_X(x) dx = \mathbb{E}[X]$$

60

Moments and MGF

- $M_X^{(1)}(t) = \int_{\mathbb{R}} x e^{tx} f_X(x) dx$
- $M_X^{(2)}(t) = \frac{d}{dt} \int_{\mathbb{R}} x e^{tx} f_X(x) dx$

$$= \int_{\mathbb{R}} x \frac{d}{dt} e^{tx} f_X(x) dx$$

$$= \int_{\mathbb{R}} x^2 e^{tx} f_X(x) dx$$
- Put $t = 0$

$$M_X^{(2)}(0) = \int_{\mathbb{R}} x^2 f_X(x) dx = \mathbb{E}[X^2]$$
- Similarly, other moments can be derived as $\mathbb{E}[X^n] = M_X^{(n)}(0)$

61

Moments and MGF

Bernoulli (p): $M_X(t) = 1 - p + pe^t$

$$\frac{d}{dt} M_X(t) = pe^t \rightarrow p \quad \mathbb{E}[X]$$

$$\frac{d^2}{dt^2} M_X(t) = pe^t \rightarrow p \quad \mathbb{E}[X^2]$$

$$\text{Var} = p - p^2 = p(1-p)$$

62

Moments and MGF

- MGF can be written in terms of moments.

$$M_X(t) = \int_{-\infty}^{\infty} e^{tx} f_X(x) dx$$

63

Moments and MGF

- MGF can be written in terms of moments.

$$M_X(t) = \int_{-\infty}^{\infty} e^{tx} f_X(x) dx$$

64

Moments and MGF

- MGF can be written in terms of moments.

$$M_X(t) = \int_{-\infty}^{\infty} \left(1 + tx + \frac{t^2 x^2}{2!} + \cdots + \frac{t^n x^n}{n!} + \cdots \right) f_X(x) dx$$

65

Moments and MGF

- MGF can be written in terms of moments.

$$\begin{aligned} M_X(t) &= \int_{-\infty}^{\infty} \left(1 + tx + \frac{t^2 x^2}{2!} + \cdots + \frac{t^n x^n}{n!} + \cdots \right) f_X(x) dx \\ &= 1 + tm_1 + \frac{t^2 m_2}{2!} + \cdots + \frac{t^n m_n}{n!} + \cdots, \end{aligned}$$

66