EE675A - IITK, 2022-23-II

Lecture: Sample Quiz

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1 Exercise

Let X be a random variable which takes values in [0, 1]. Let $\mathbb{E}[X] = \mu$, and $\hat{\mu}_N$ denotes the sample average obtained by observing N i.i.d samples of X. Suggest a number N such that $\hat{\mu}_N$ does not deviate from μ by more than 0.01 with a very high probability (let us say with 0.99 probability).

Hoeffding's Inequality: It states that the probability of deviation of estimated mean $\hat{\mu}(a)$ from true mean greater than ϵ is upper bound by $2e^{-2\epsilon^2N}$ where N is number of samples

$$P[|\bar{\mu}(a) - \mu(a)| \ge \epsilon] <= 2e^{-2\epsilon^2 N}$$

Given $\hat{\mu}_N$ does not deviate from μ by more than 0.01 with a very high probability (let us say with 0.99 probability)

 $\epsilon = 0.01$ does not deviate with probability 0.99 \implies deviates with probability 0.01

$$P[|\bar{\mu}(a) - \mu(a)| \ge 0.01] \le 0.01$$
$$0.01 \le 2e^{-2\epsilon^2 N}$$
$$0.01 < 2e^{-2(0.01)^2 N}$$

taking log on both sides ans solving further

$$N \ge 24692$$

2 Exercise

Prove that ϵ -greedy algorithm, where ϵ is some fixed constant, incurs a regret that grows linearly in T.

In ϵ - Greedy Algorithm, in every round with probability ϵ an arm is played randomly, and with probability $1 - \epsilon$ the optimal arm is played.

Therefore, $\Delta(a)$ is bounded by,

$$\Delta(a(t)) = \mu(a^*(t)) - \mu(a(t))$$

$$R(t) = \sum_{t \in T} \Delta(a(t))$$

$$E[R(t)] = E[\sum_{t \in T} \Delta(a(t))]$$

$$E[R(t)] = \sum_{t \in T} E[\Delta(a(t))]$$

$$= \sum_{t \in T} (\sum_{a \in A} \Delta(a(t)) P(a(t) = a))$$

$$\hat{\boldsymbol{a}}(t) = \arg\max_{a \in A} \overline{\mu_t}(a)$$

(Note that \hat{a} is the best arm based on sample estimate, whereas a^* is the actual best arm that gives us the maximum reward.)

consider there are k arms

The probability to choose an arm at random is 1/k with ϵ , the probability a random arm is selected which includes the best arm, so the probability of choosing a random arm is ϵ/k

In Greedy case there is only one best arm and it is played with probability 1- ϵ

$$\begin{split} P(a(t) &= \hat{\boldsymbol{a}}) = (1 - \epsilon) + \epsilon/k \\ P(a(t) &= a) = \epsilon/k \\ P(a(t) &= a) \geq \epsilon/k \\ E[R(t)] &= \sum_{t \in T} (\sum_{a \in A} \Delta(a(t)) P(a(t) = a)) \\ E[R(t)] &\geq \sum_{t \in T} (\sum_{a \in A} \Delta(a(t)) \epsilon/k) \\ E[R(t)] &\geq (\epsilon T/k) \sum_{a \in A} \Delta(a(t)) \end{split}$$

 ϵ is constant in this case, the expected regret grows linearly **T**

3 Exercise

In the UCB algorithm discussed in our class, we had

$$UCB_t(a) := \overline{\mu}_t(a) + \sqrt{\frac{2lnT}{n_t(a)}}$$
(2)

To execute this algorithm, we need to know the value of T, *i.e.*, we should know the total number of rounds we are going to play. Let us assume that we do not know the value of T in advance and consider the following variant of UCB.

$$UCB_t(a) := \overline{\mu}_t(a) + \sqrt{\frac{lnt}{n_t(a)}}$$
(3)

Prove that this UCB variant also has a similar regret bound. **HINT**: The proof proceeds in a similar fashion. Just make necessary changes to reflect this new formulation

step-1:

We have $UCB_t(a) \ge UCB_t(a^*)$

calculate the probability that the true mean $\mu(a)$ lies outside the confidence interval that we have at time 't' i.e., \forall actions $a \in A$ and 0 < t < T

$$\mathbb{P}\left(\overline{\mu}_t(a) \notin \left(\overline{\mu}_t(a) - \sqrt{\frac{ln\mathbf{t}}{n_t(a)}}\right), \left(\overline{\mu}_t(a) + \sqrt{\frac{ln\mathbf{t}}{n_t(a)}}\right)\right) \tag{4}$$

Using Hoeffding's Inequality

$$P[|\bar{\mu}(a) - \mu(a)| \ge \epsilon] <= 2e^{-2\epsilon^2 N}$$

considering ϵ as $\sqrt{\frac{lnt}{n_t(a)}}$ we get

$$2e^{-2\epsilon^2 N} = \frac{2}{t^2}$$

step-2:

In the proof discussed in class for UCB, we assumed that for every time $1 \le t \le T$, and for every arm $a \in A$, our confidence intervals are correct with high probability, specifically if we use union bound

 $\mathbb{P}(\text{ our confidence interval is wrong for at least one arm (or) one time slot}) \leq \frac{2}{T^4}Tk$

for a given arm 'a' at a particular round 't' the probability of deviation of true mean out side the bound is $2/\mathbf{T}^4$

Union bound over all time slots 1<t<**T** and all the 'k' arms

Assuming $k \leq T$ which is a reasonable assumption since we should have enough rounds to play each arm at least once

P(confidence interval getting voilated at least in one slot or for one arm) $\leq \mathcal{O}\left(\frac{1}{T^2}\right)$

For UCB=
$$\sqrt{\frac{logt}{n_t(a)}}$$
:

If we do a similar analysis for $\sqrt{\frac{logt}{n_t(a)}}$ version of UCB Using Hoeffding's inequality

$$\mathbf{P}(\text{voilation for arm 'a' at round 't'}) \le \frac{2}{T^2} \forall a \in \mathcal{A}$$
 (6)

 $\mathbf{P}(\text{ voilation for at least on arm at least on o=round}) \leq \sum_{t=1}^{T} \sum_{a \in A} \frac{2}{t^2}$

$$\mathbf{P} = K \sum_{t=1}^{T} \frac{2}{t^2} > K \tag{7}$$

$$\mathbf{as} T \to \infty$$

Conclusion: This bound is greater than 1, which is useless, we cant hopw our confidence intervals are correct all the times

Step-3: For the UCB version done in class we have shown that

i)A bad arm cannot be played too many times.ii)If any arm is played a lot of times, its not a very bad arm.

similarly, " If a suboptimal arm is played sufficient number of times then the probability of playing it further is very less

$$\mathbb{P}\left(\frac{a(t+1)=a}{n_t(a) \ge \frac{4lnt}{(\triangle(a))^2}}\right) \le \frac{4}{t^2} \tag{8}$$

Note: using the following properties 1. If an arm has to be played at time 't',

$$UCB_t(a) \ge UCB_t(a^*)$$
 (9)

$$\bar{\mu}_t(a) + \epsilon_t(a) \ge \bar{\mu}_t(a^*) + \epsilon_t(a^*) \tag{10}$$

$$\bar{\mu}_t(a^*) - \bar{\mu}_t(a) \le \epsilon_t(a) - \epsilon_t(a^*) \tag{11}$$

Also we have,

$$\Delta(a) = \mu(a^*) - \mu(a) \tag{12}$$

$$\Delta(a) \le \bar{\mu}_t(a^*) + \epsilon_t(a^*) - (\bar{\mu}_t(a) - \epsilon_t(a)) \tag{13}$$

$$\Delta(a) \le 2\epsilon_t(a) \tag{14}$$

$$\Delta(a) \le 2\sqrt{\frac{\log t}{n_t(a)}}\tag{15}$$

$$\boldsymbol{n}_t(a) \ge \frac{4\log t}{(\Delta(a))^2} \tag{16}$$

 $\mathbb{P}(\text{confidence interval voilation for arm a, at t}) \leq \frac{2}{t^2} \text{using this property for both a, a*}$

P(confidence interval going wrong either for a or a^*) $\leq \frac{4}{t^2}$

$$\left| \left(\left| \mu(a) - \overline{\mu}_t(a) \right) \right| \sqrt{\frac{\ln t}{n_t(a)}} \right) \le \frac{2}{t^2}$$
(17)

(18)

using for both a and a^* we get,

P(confidence interval going wrong for either a or
$$a^*$$
) $\leq \frac{4}{t^2}$ (19)

$$P(\text{confidence interval correct in both cases}) \le 1 - \frac{4}{t^2}$$
 (20)

(21)

we can say that

$$\mathbb{P}\left(\frac{a(t+1)=a}{n_t(a) \ge \frac{4lnt}{(\triangle(a))^2}}\right) \le \frac{4}{t^2}$$
(22)

Step 4:
$$E[n_t(a)] \le \frac{4lnT}{(\triangle(a))^2} + 8$$

This is nothing but "A bad arm in not played many times". Express

$$E[n_t(a)] = 1 + E\left[\sum_{t=K+1}^{T} 1_{(a(t)=a)}\right]$$

every arm is played once in first K rounds so 1

Divide the second term into two parts

$$\mathbf{E}\left[\sum_{t=K+1}^{T} 1_{a(t)=a,n_{t}(a) \leq \frac{4lnt}{(\triangle(a))^{2}}}\right] + \mathbf{E}\left[\sum_{t=K+1}^{T} 1_{a(t)=a,n_{t}(a) \geq \frac{4lnt}{(\triangle(a))^{2}}}\right]$$

$$\mathbf{E}\left[\sum_{t=K+1}^{T} 1_{a(t)=a,n_t(a) \leq \frac{4lnt}{(\triangle(a))^2}}\right]$$
the total contribution of this term is $\frac{4lnT}{((a))^2}$ (23)

$$\begin{split} \mathbf{E} \Big[\sum_{t=K+1}^{T} \mathbf{1}_{a(t)=a,n_t(a) \geq \frac{4lnt}{(\triangle(a))^2}} \Big] \\ \sum_{t=k+1}^{T} \mathbb{P} \left(\frac{a(t)}{n_t(a) \geq \frac{4lnt}{(\triangle(a))^2}} \right) \mathbb{P} \left(n_t(a) \geq \frac{4lnt}{(\triangle(a))^2} \right) \\ \sum_{t=k+1}^{T} \frac{4}{t^2} = 8 \end{split}$$

step 5: calculate $\mathbf{E}\left[R(T;a)\right]$ using $\mathbf{E}[n_t(a)]$ from step 4 use $\mathbf{E}[R(T)] = \sum_a \mathbf{E}[R(T;a)]$

$$\sum_{a} \mathbf{E}[R(T;a)] = \sum_{a} \mathbf{E}[n_t(a)] \triangle(a)$$
(24)

$$\sum_{a} \mathbf{E}[R(T;a)] \le \sum_{a} \frac{4\log t}{(\Delta(a))} + 8(\Delta(a)) \tag{25}$$