# GeoModels Tutorial: simulation, estimation and prediction of spatio-temporal data using Gaussian random fields

Moreno Bevilacqua Víctor Morales-Oñate

#### Introduction

In this tutorial we show how to analyze geo-referenced spatio temporal data using Gaussian random fields (RFs) with the R package GeoModels.

We first load the R libraries needed for the analysis and set the name of the model in the GeoModels package:

```
rm(list=ls())
require(GeoModels)
require(fields)
model="Gaussian" # model name in the GeoModels package
set.seed(12)
```

### Simulation of a space-time Gaussian random field

Let us consider a space-time Gaussian RF  $Z = \{Z(s,t), s \in S, t \in B\}$ , where s represents a location in the domain S and t represents a temporal instant the domain S. We assume that Z is stationary with zero mean, unit variance and correlation function given by  $\rho(h, u) = \text{cor}(Z(s+h,t+u),Z(s,t))$ .

Then we consider the RF  $Y = \{Y(s,t), s \in S, t \in T\}$  defined by the location and scale transformation:

$$Y(s,t) = \mu(s,t) + \sigma Z(s,t) \tag{1}$$

where  $\mu(\boldsymbol{s},t) = X(\boldsymbol{s},t)^T \boldsymbol{\beta}$  and  $X(\boldsymbol{s},t)$  is a k-dimensional vector of covariates and  $\boldsymbol{\beta} = (\beta_1, \dots, \beta_k)^T$  is a k-dimensional vector of (unknown) parameters (in this tutorial we fix k = 2). Then  $\mathbb{E}(Y(\boldsymbol{s},t)) = X(\boldsymbol{s},t)^T \boldsymbol{\beta}$ ,  $\operatorname{var}(Y(\boldsymbol{s},t)) = \sigma^2$  and  $\operatorname{cov}(Y(\boldsymbol{s}+\boldsymbol{h},t+u),Y(\boldsymbol{s},t)) = \sigma^2 \rho(\boldsymbol{h},u)$ .

Suppose we want to simulate a realization of Y at  $t_1 = 1, t_2 = 2, ..., t_T = 30$ , T = 30 temporal instants and N = 400 spatial locations uniformly distributed in the unit square. The total number of space-time locations is given by NT = 12000.

We first set the temporal instants and then the spatial coordinates with associated covariates.

```
coordt=1:30  # number of temporal instants
T=length(coordt)
NN=400  # number of spatial locations
```

```
x = runif(NN, 0, 1); y = runif(NN, 0, 1)
coords=cbind(x,y)
X=cbind(rep(1,NN*T),runif(NN*T))
```

We then specify the mean, variance and nugget parameters

```
mean = 0.5; mean1= -0.25
sill=2; nugget=0
```

where mean, mean1 and sill are respectively  $\beta_1$ ,  $\beta_2$  and  $\sigma^2$ .

In this tutorial we assume a simple spatially isotropic and symmetric in time separable Wendland model for the space-time correlation function that is:

$$\rho((\boldsymbol{h}, u); \alpha_s, \alpha_t, \mu_s, \mu_t) = \left(1 - \frac{||\boldsymbol{h}||}{\alpha_s}\right)_+^{\mu_s} \left(1 - \frac{|u|}{\alpha_t}\right)_+^{\mu_t}$$
(2)

This kind of correlation model is compactly supported in space or time that is the correlation function is zero when  $||\boldsymbol{h}|| > \alpha_s$  or  $|u| > \alpha_t$ . This is an interesting feature since algorithms for sparse matrices can be used to speed-up the computations associated with the corresponding correlation matrix.

Then we set the name of the correlation model and the associated parameters:

```
corrmodel="Wend0_Wend0";
scale_s=0.2; scale_t=2
```

where scale\_s and scale\_t corresponds to  $\alpha_s$  and  $\alpha_t$ , the compact supports of the correlation model. We are now ready to simulate, through Cholesky decomposition, the space time Gaussian RF using the function GeoSim:

Note that the option sparse=TRUE allows to consider algorithms for sparse matrices when performing Cholesky decomposition, using in the package spam (Furrer and Sain (2010)). Informations about the sparsity of the covariance matrix can be obtained though the function GeoCovmatrix with the following code

```
cc$nozero
[1] 0.01768
```

This means that (approximatively) 98% of the covariance matrix are zeros *i.e* the matrix is highly sparsed.

#### Estimation of a Gaussian space-time random field

Given a space-time realization  $\{Y(\mathbf{s}_i, t_l), l = 1...T, i = 1,...N\}$ , let  $f_Y(y_{il}, y_{jk})$  the Gaussian density of a pair of observations  $Y(\mathbf{s}_i, t_l)$  and  $Y(\mathbf{s}_j, t_k)$ . Then, the pairwise likelihood function is defined as (Bevilacqua and Gaetan, 2015):

$$pl(\boldsymbol{\theta}) = \sum_{i,j,l,k \in D} log(f_Y(y_{il}, y_{jk})) w_{ijlk}$$
(3)

where

$$D = \begin{cases} l = 1 \dots T, & i = 1, \dots, N, & k = l, \dots, T \\ j = i + 1, \dots, N & \text{if} \quad l = k \\ j = 1, \dots, N & \text{if} \quad l > k \end{cases}$$

and  $w_{ijlk}$  are non-negative weights, not depending on  $\boldsymbol{\theta}$ , specified as:

$$w_{ijlk}(m, d_t) = \begin{cases} 1 & \mathbf{s}_i \in N_m(\mathbf{s}_j), |t_l - t_k| < d_t \\ 0 & \text{otherwise} \end{cases}$$
 (4)

and in this case  $\boldsymbol{\theta} = (\mu, \sigma^2, \alpha_s, \alpha_t)^T$ . The pairwise likelihood estimator  $\hat{\boldsymbol{\theta}}_{pl}$  is obtained maximizing (3) with respect to  $\boldsymbol{\theta}$ . In the GeoModels package we can choose the fixed parameters and the parameters that must be estimated. Pairwise likelihood estimation is performed with the function GeoFit using the BFGS method:

Note that the option neighb=5 and maxtime=1 set the (arbitrary) compact supports of the weight function (4) i.e. m = 5 and  $d_t = 1$ .

The object fit include informations about the pairwise likelihood estimation

```
fit
Maximum Composite-Likelihood Fitting of Gaussian Random Fields
Setting: Marginal Composite-Likelihood
Model: Gaussian
Distance: Eucl
Type of the likelihood objects: Pairwise
Covariance model: Wend0_Wend0
Optimizer: BFGS
Number of spatial coordinates: 400
Number of dependent temporal realisations: 30
Type of the random field: univariate
Number of estimated parameters: 5
Type of convergence: Successful
Maximum log-Composite-Likelihood value: -610875.85
Estimated parameters:
        mean1 scale_s scale_t
  mean
                               sill
0.5232 -0.2841 0.2010
                      2.0634
                              1.9699
```

## Checking model assumptions

Given the estimation of the mean regression and sill parameters, the estimated residuals

$$\widehat{Z}(\boldsymbol{s}_i,t_l) = \frac{Y(\boldsymbol{s}_i,t_l) - X(\boldsymbol{s}_i,t_l)^T \widehat{\boldsymbol{\beta}}}{(\widehat{\sigma}^2)^{\frac{1}{2}}} \quad i = 1,\dots,N \quad l = 1,\dots,T$$

can be viewed as a realization of a zero mean stationary Gaussian RF with correlation function  $\rho(\boldsymbol{h}, u)$ . The residuals can be computed using the GeoResiduals function:

```
res=GeoResiduals(fit) # computing residuals
```

Then the marginal distribution assumption on the residuals can be graphically checked for instance with a qq-plot (Figure 1, left part) using the function GeoQQ

```
### checking model assumptions: marginal distribution

GeoQQ(res)
```

The correlation model assumption can be checked comparing the empirical and the estimated space-time semivariogram functions using the GeoVariogram and GeoCovariogram functions (Figure 1, right part):

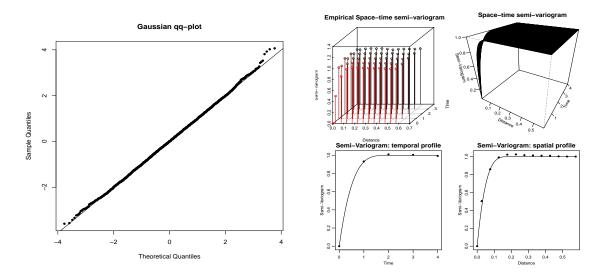


Figure 1: Left: QQ-plot for the residuals of the space-time Gaussian RF. Right: space-time empirical vs estimated semi-variogram function for the residuals

## Prediction of space-time Gaussian random fields

For a given space time location  $(s_0, t_0)$  with associated covariates  $X(s_0, t_0)$ , the optimal prediction of Gaussian RF is computed as:

$$\widehat{Y}(\boldsymbol{s}_0, t_0) = X(\boldsymbol{s}_0, t_0)^T \widehat{\boldsymbol{\beta}} + \sum_{l=1}^T \sum_{i=1}^N \lambda_{l,i} [Y(\boldsymbol{s}_i, t_l) - X(\boldsymbol{s}_i, t_l)^T \widehat{\boldsymbol{\beta}}]$$
 (5)

where the vector of weights  $\lambda = (\lambda_{1,1}, \dots, \lambda_{T,N})'$  is given by  $\lambda = R^{-1}c$  and

• 
$$c = (cor(Y(s_0, t_0), Y(s_1, t_1)), \dots, cor(Y(s_0, t_0), Y(s_N, t_T)))^T$$
.

•  $R = [[\operatorname{cor}(Y(s_i, t_l), Y(s_j, t_k)]_{l,k=1}^T]]_{i,j=1}^N$  is the correlation matrix.

Kriging can be performed using the **GeoKrig** function. We just need to specify the spatial locations and temporal instants to predict. In this example we consider a spatial regular grid and one temporal instant:

```
xx=seq(0,1,0.03)
loc_to_pred=as.matrix(expand.grid(xx,xx))  # locations to predict
n_loc=nrow(loc_to_pred)
times=c(2)  #time to predict
```

In addition, we need to specify the associated covariates:

```
Xloc=cbind(rep(1,n_loc),runif(n_loc))
```

Then the optimal linear prediction (5), using the estimated parameters, can be performed using the GeoKrig function:

```
pr = GeoKrig(fit,loc=loc_to_pred,time=times,Xloc=Xloc,sparse=TRUE,mse=TRUE)
```

A kriging map for the temporal instant t = 2 with associate mean square error (Figure 2) can be obtained with the following code:

Local spacetime kriging can be an alternative in order to speedup the computation of the kriging predictions using a fixed space time neighboord. It can be performed using the GeoKrigloc function that can be parallelized:

```
pr1 = GeoKrigloc(fit,loc=loc_to_pred,time=times,Xloc=Xloc,sparse=TRUE,
mse=TRUE,neighb=150,maxtime=2,parallel=TRUE)
```

It can be appreciated that classical and local kriging predictions, in this example, are pretty similar

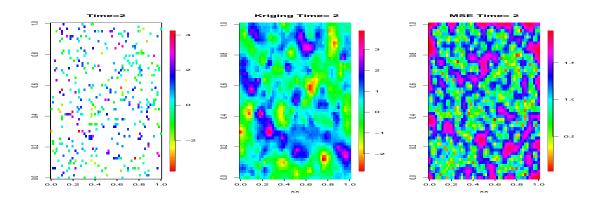


Figure 2: From left to right: observed spatial data at time t=2, associated kriging map and mean square error map.

```
summary(c(pr$pred-prloc$pred))
Min. 1st Qu. Median Mean 3rd Qu. Max.
-1.167e-04 -5.526e-07 6.700e-10 9.572e-07 1.157e-06 2.031e-04
```

# References

Bevilacqua, M. and C. Gaetan (2015). Comparing composite likelihood methods based on pairs for spatial Gaussian random fields. *Statistics and Computing* 25, 877–892.

Furrer, R. and S. R. Sain (2010). spam: a sparse matrix R package with emphasis on mcmc methods for Gaussian Markov random rields. *Journal of Statistical Software 36*, 1–25.