Special topics: Time series analysis, higher-order models,& oscillations,

Optional: latent change score models, accelerated longitudinal designs

Outline

- \triangleright Time series (N = 1, T large)
- \triangleright CARMA(p,q) models (higher order models)
- > Oscillations (and a few words on the virtue of varying time intervals)
- > Latent change score models
- ➤ Accelerated longitudinal designs

- > Sometimes we are interested in just a single individual (patient, politician, athlete, ...)
- > Sometimes we are primarily interested in the (within-person) structure and less on the dynamics/development. For example, the personality structure of an individual.
- ➤ Sometimes we are interested in modeling complex, but systematic development of a single unit over time (e.g., changes in temperature or the activity of the sun)...
- ➤ ...and sometimes such systematic developments are of an oscillating nature (e.g., "emotion as a thermostat"; Chow et al. 2005; psychological regulation; Deboeck & Bergeman; 2013; or circadian rhythms?).

THE THREE BASIC FACTOR-ANALYTIC RESEARCH DESIGNS 501

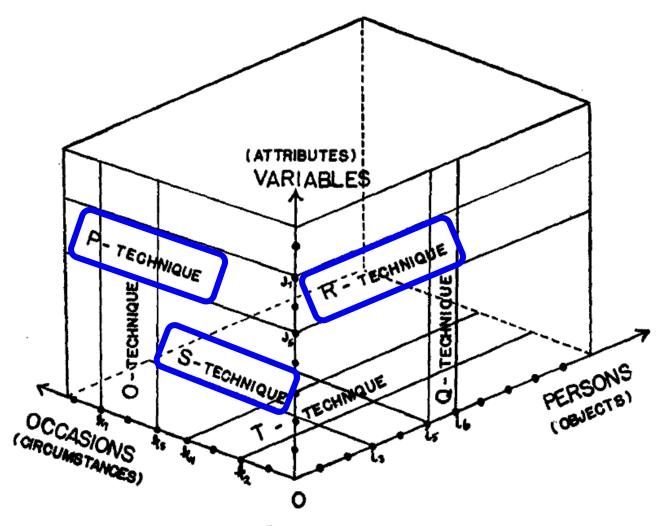


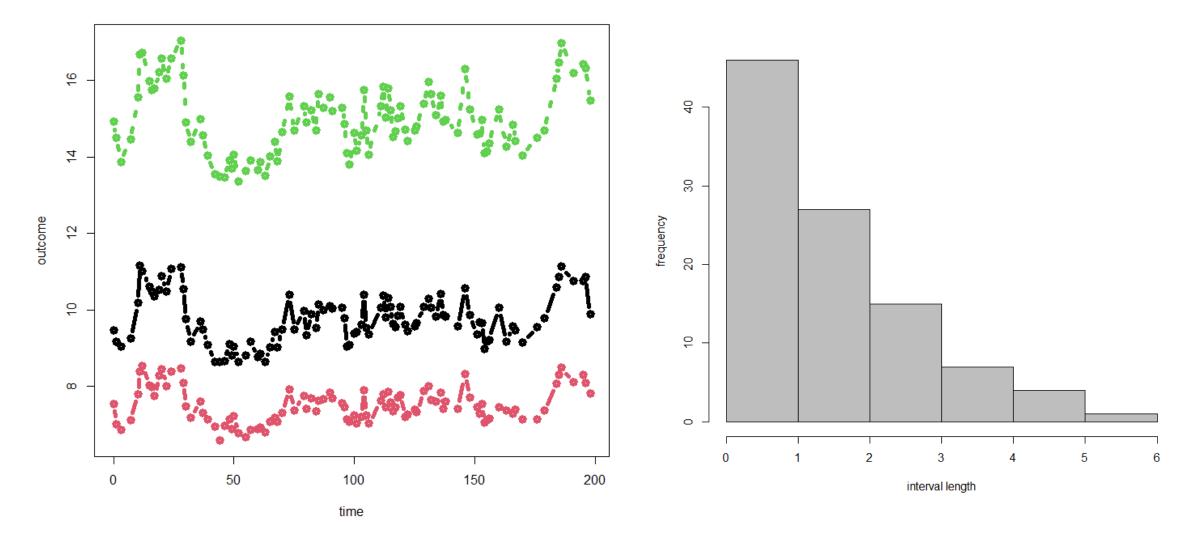
Fig. 1, The Covariation Chart

Original figure from Cattell (1952, p. 501)

Time series (N = 1, T large)

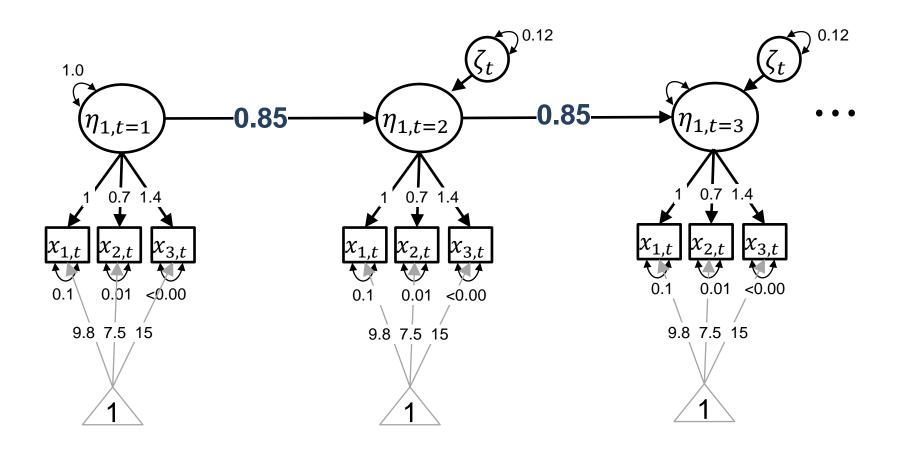
- ➤ Historically, *P*-Technique was proposed to study the factorial structure within a unit/person.
- ➤ P-Technique was criticized, because it does not adequately account for the lagged covariance structure. Dynamic Factor Analysis (DFA; Molenaar, 1985) was proposed as an alternative that overcomes this limitation.
- ➤ Interestingly, in case of equidistant time intervals, *P*-Technique seems to perform reasonably well, despite its known limitations (Molenaar & Nesselroade, 2009).
- ➤ However, by design, neither DFA nor *P*-Technique addresses the problem of unequal time intervals. This is resolved in continuous time modeling.

Dynamic factor analysis (U = 100 measurement occasions; N = 1; 3 manifest variables, differing time intervals)



Time series (N = 1, T large)

Dynamic factor analysis (U = 100 measurement occasions; N = 1; 3 manifest variables, differing time intervals)



Time series (N = 1, T large)

From theory to practice...

```
library(ctsem)
library(ctsemOMX)
data(ctExample3)
#interfacing to OpenMx (wide-data format)
DFAmodel <- ctModel(n.latent = 1, n.manifest = 3, Tpoints = 100,
          TOMEANS = matrix(0,1,1),
          LAMBDA = matrix(c(1, 'lambda2', 'lambda3'), nrow = 3, ncol = 1),
          MANIFESTMEANS = matrix(c('manifestmean1', 'manifestmean2', 'manifestmean3'), nrow = 3, ncol = 1),
          TOVAR = diag(1),
          type = "omx")
DFAmodelfit <- ctFit(ctExample3, ctmodelobj = DFAmodel)</pre>
summary(DFAmodelfit, verbose=T)
plot(DFAmodelfit)
```

From theory to practice...

#Transforming data from wide to long

ctExample3long_intervals <- ctWideToLong(datawide = ctExample3, Tpoints=100, n.manifest=3) #convert wide to long format ctExample3long <- ctDeintervalise(datalong = ctExample3long_intervals, id='id', dT='dT') #convert intervals to absolute time

#interfacing to Stan (long-data format)

DFAmodelv2fit <- ctStanFit(datalong=ctExample3long, ctstanmodel=DFAmodelv2, indvarying=F, optimize = TRUE, nopriors = TRUE) summary(DFAmodelv2fit) plot(DFAmodelv2fit)

- Just like in discrete time series (or in the general cross-lagged panel model), we can move beyond the AR(p = 1) model and move to higher order models of order p (autoregressive part) and q (moving average part).
- ➤ CARMA(*p*,*q*) models can be useful when we are interested in modeling more complex, but systematic, development over time (e.g., changes in temperature or the activity of the sun)...
- ➤ Given the matrix formulation of the ct model (and the corresponding software implementation), higher order models are easily specified by augmenting the latent state vector (i.e.,

Remember:

$$d\mathbf{\eta}(t) = (\mathbf{A}\mathbf{\eta}(t) + \mathbf{b} + \mathbf{M}\mathbf{\chi}(t))dt + \mathbf{G}d\mathbf{W}(t))$$

Symbolically, this can also be written as*

$$\frac{\mathrm{d}\mathbf{\eta}(t)}{\mathrm{d}t} = \mathbf{A}\mathbf{\eta}(t) + \mathbf{b} + \mathbf{M}\mathbf{\chi}(t) + \mathbf{G}\frac{\mathrm{d}\mathbf{W}(t)}{\mathrm{d}t}$$

- ightharpoonup With $\mathbf{y}(t) = \Lambda \mathbf{\eta}(t) + \mathbf{\tau} + \boldsymbol{\varepsilon}(t)$.
- ➤ E.g., for a **uni**variate process, without intercepts and without error variance (to keep it as simple as possible), the measurement model would be:

$$y(t) = 1 \cdot \eta(t) + 0 + 0$$

^{*}This is the notation used in Oud et al. (2018) and elsewhere. Don't get confused.

- \triangleright To incorporate higher order processes, we simply expand the state vector $\eta(t)$, without changing the measurements (because the data do not change).
- > E.g., to include a higher order term in a univariate model we write:

$$y(t) = (1 \quad 0) \begin{pmatrix} \eta_1 \\ \eta_2 \end{pmatrix} (t)$$

Note that η_2 is a latent variable without a reflective indicator. Only η_1 is connected to data.

Plugging this into the general equation (and ignoring the "irrelevant" parts), yields:

$$d \begin{pmatrix} \eta_1 \\ \eta_2 \end{pmatrix} (t) = \left(\mathbf{A} \begin{pmatrix} \eta_1 \\ \eta_2 \end{pmatrix} (t) \right) dt + \mathbf{G} d\mathbf{W}(t)$$

To incorporate a higher order effect (second order derivative) we define $\eta_2 = \frac{d\eta_1}{dt}$, which yields

$$\frac{\mathrm{d} \binom{\eta_1}{\eta_2}(t)}{\mathrm{d} t} = \frac{\mathrm{d} \binom{\frac{1}{d\eta_1}}{\mathrm{d} t}(t)}{\mathrm{d} t} = \binom{\frac{d\eta_1}{\mathrm{d} t}}{\frac{d^2\eta_1}{\mathrm{d} t^2}}$$
 on the left-hand side of the equation.

➤ With this "dependent vector" we can now set up a second order model:

$$\begin{pmatrix} \frac{d\eta_1}{dt} \\ \frac{d^2\eta_1}{dt^2} \end{pmatrix} = \begin{pmatrix} 0 & 1 \\ a_0 & a_1 \end{pmatrix} \begin{pmatrix} \frac{\eta_1}{d\eta_1} \\ \frac{dt}{dt} \end{pmatrix}$$
A matrix

Without matrix notation:

$$\frac{\mathrm{d}\eta_1}{\mathrm{d}t} = 1\frac{\mathrm{d}\eta_1}{\mathrm{d}t}$$

$$\frac{\mathrm{d}^2\eta_1}{\mathrm{d}t^2} = a_0\eta_1 + a_1\frac{\mathrm{d}\eta_1}{\mathrm{d}t}$$
 second order drift effect

Note that we expressed the univariate second order differential equation as a first order differential equation in matrix notation.

$$\begin{pmatrix} \frac{d\eta_1}{dt} \\ \frac{d^2\eta_1}{dt^2} \end{pmatrix} = \begin{pmatrix} 0 & 1 \\ a_0 & a_1 \end{pmatrix} \begin{pmatrix} \frac{\eta_1}{d\eta_1} \\ \frac{d}{dt} \end{pmatrix} + \begin{pmatrix} 0 & 0 \\ 0 & G_0 \end{pmatrix} \frac{d\mathbf{W}(t)}{dt}$$

This is a special case of general equation we are familiar with

$$\frac{\mathrm{d}\mathbf{\eta}(t)}{\mathrm{d}t} = \mathbf{A}\mathbf{\eta}(t) + \mathbf{0} + \mathbf{0} + \mathbf{G}\frac{\mathrm{d}\mathbf{W}(t)}{\mathrm{d}t}$$

> Thus, the solution of the (matrix) stochastic differential equation remains unchanged by incorporating higher order effects.

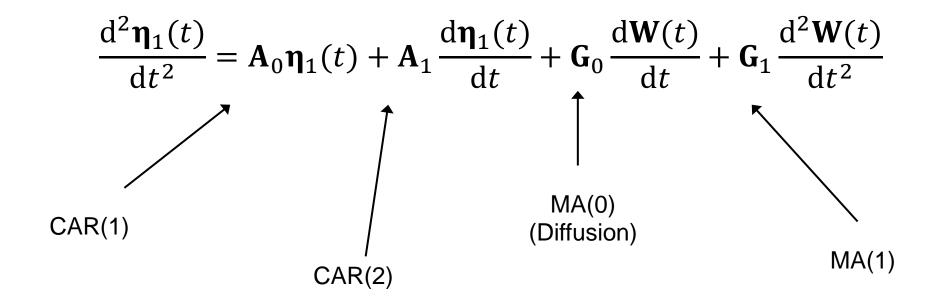
- The approach can be generalized to any order p, any order q and can be easily extended from univariate to multivariate systems (possibly with a measurement model).
- ➤ The math gets a bit tedious, though. For details, see Oud et al. (2018). Switching gears, we can for example define a CARMA(2,1) model as:

with
$$\frac{\mathrm{d}\eta(t)}{\mathrm{d}t} = \mathbf{A}\eta(t) + \mathbf{0} + \mathbf{0} + \mathbf{G}\frac{\mathrm{d}\mathbf{W}(t)}{\mathrm{d}t}$$

$$\mathbf{\eta} = \begin{pmatrix} \mathbf{\eta}_1 \\ \mathrm{d}\mathbf{\eta}_1 \\ \mathrm{d}t \end{pmatrix} \qquad \mathbf{A} = \begin{pmatrix} \mathbf{0} & \mathbf{1} \\ \mathbf{A}_0 & \mathbf{A}_1 \end{pmatrix} \qquad \mathbf{G} = \begin{pmatrix} \mathbf{0} & \mathbf{0} \\ \mathbf{G}_0 & \mathbf{G}_1 \end{pmatrix}$$

yields
$$\begin{pmatrix} \frac{\mathrm{d} \eta_1}{\mathrm{d} t} \\ \frac{\mathrm{d}^2 \eta_1}{\mathrm{d} t^2} \end{pmatrix} = \begin{pmatrix} \mathbf{0} & \mathbf{1} \\ \mathbf{A}_0 & \mathbf{A}_1 \end{pmatrix} \begin{pmatrix} \eta_1 \\ \frac{\mathrm{d} \eta_1}{\mathrm{d} t} \end{pmatrix} + \begin{pmatrix} \mathbf{0} & \mathbf{0} \\ \mathbf{G}_0 & \mathbf{G}_1 \end{pmatrix} \begin{pmatrix} \frac{\mathrm{d} \mathbf{W}}{\mathrm{d} t} \\ \frac{\mathrm{d}^2 \mathbf{W}}{\mathrm{d} t^2} \end{pmatrix}$$

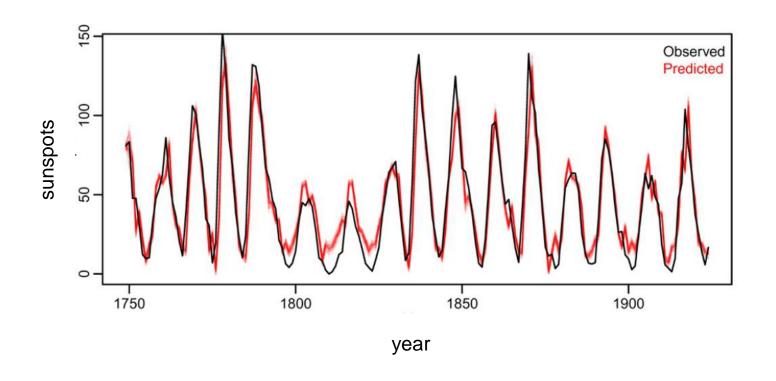
Thus, the CARMA(2,1) model is



Note that this is only one possible parametrization. If there is no measurement model (as in this example) another way to include MA(q) effects is via the loading matrix.

CARMA(p, q) models

- > From theory to practice
- ➤ A famous example of a real-world process that is well captured by a CARMA(2,1) model is the sunspot activity.



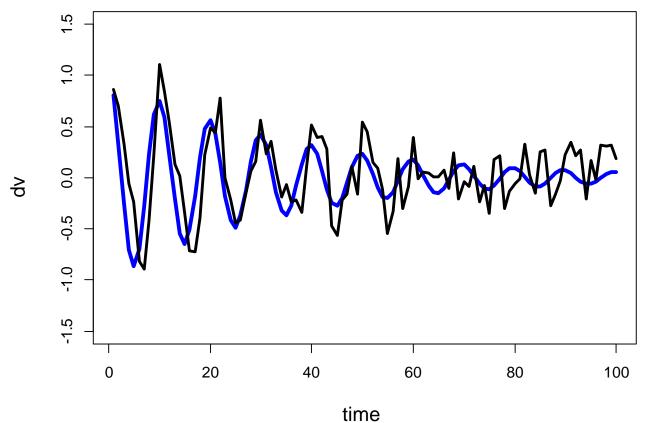
CARMA(p, q) models (sunspot data)

Let's try this:

```
library(ctsem)
sunspots<-sunspot.year #get data
sunspots<-sunspots[50: (length(sunspots) - (1988-1924))]
id <- 1
time <- 1749:1924
datalong <- cbind(id, time, sunspots)
#setup model
ssmodel <- ctModel(type='stanct', n.latent=2, n.manifest=1,
          manifestNames='sunspots',
          latentNames=c('ss level', 'ss velocity'),
          LAMBDA=matrix(c(1, 'ma1'), nrow=1, ncol=2),
          DRIFT=matrix(c(0, 'a21', 1, 'a22'), nrow=2, ncol=2),
          MANIFESTMEANS=matrix(c('m1'), nrow=1, ncol=1),
          CINT=matrix(c(0, 0), nrow=2, ncol=1),
          TOVAR=matrix(c(1,0,0,1), nrow=2, ncol=2), #Because single subject
          DIFFUSION=matrix(c(0, 0, 0, "diffusion"), ncol=2, nrow=2))
ssmodel$pars$indvarying<-FALSE #Because single subject
ssmodel$pars$transform[14]<- '(param)*5+44 ' #Because not mean centered
ssmodel$pars$transform[4]<-'log(exp(-param*1.5)+1)' #To avoid multimodality
ssfit <- ctStanFit(datalong, ssmodel, iter=1000, chains=4)</pre>
summary(ssfit, parmatrices = TRUE)
plot(ssfit)
```

Oscillations (and a few words on the virtue of varying time intervals)

- As apparent from the sunspot example, continuous time models are able to capture rhythmic / oscillating patterns.
- This includes standard oscillations, such as damped and/or coupled linear oscillators.



In large parts the following content is based on:

Technically, when the eigenvalues of the drift matrix are complex, the process is oscillating.

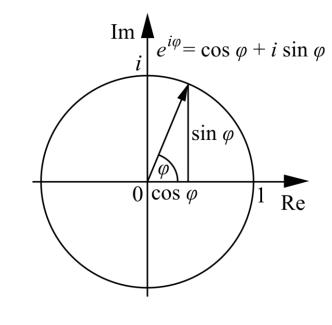
The eigenvalues of the drift matrix

$$\mathbf{A} = \begin{pmatrix} 0 & 1 \\ -\omega_0^2 & -\gamma \end{pmatrix}$$

are

$$\lambda_{1,2}=-rac{\gamma}{2}\pm i\sqrt{-rac{\gamma^2}{4}+\omega_0^2}$$
 with $i=\sqrt{-1}$ yielding

$$i = \sqrt{-1}$$



$$e^{\mathbf{A}\cdot\Delta t_{i}} = e^{-\frac{\gamma}{2}\Delta t_{i}} \times \begin{bmatrix} \frac{\gamma}{2\omega}\sin(\omega\Delta t_{i}) + \cos(\omega\Delta t_{i}) & \frac{1}{\omega}\sin(\omega\Delta t_{i}) \\ -\frac{\omega_{0}^{2}}{\omega}\sin(\omega\Delta t_{i}) & \cos(\omega\Delta t_{i}) - \frac{\gamma}{2\omega}\sin(\omega\Delta t_{i}) \end{bmatrix}$$

Just like before, the "trick" is to define a second order differential equation as a first order differential equation by expanding the state vector

$$\begin{pmatrix} \frac{d\eta(t)}{dt} \\ \frac{d^2\eta(t)}{dt^2} \end{pmatrix} = \begin{pmatrix} 0 & 1 \\ -\omega_0^2 & -\gamma \end{pmatrix} \begin{pmatrix} \eta(t) \\ \frac{d\eta(t)}{dt} \end{pmatrix} + \begin{pmatrix} 0 \\ b \end{pmatrix} + \begin{pmatrix} 0 & 0 \\ 0 & g \end{pmatrix} \begin{pmatrix} \frac{dW_1(t)}{dt} \\ \frac{dW_2(t)}{dt} \end{pmatrix}$$
$$= \frac{d\eta(t)}{dt} = \mathbf{A}\eta(t) + \mathbf{b} + \mathbf{G} \frac{d\mathbf{W}(t)}{dt}.$$

with the first order derivative as a latent variable without an underlying manifest variable.

$$y(t) = (1 \quad 0) \left(\frac{\eta(t)}{d\eta(t)} \right)$$

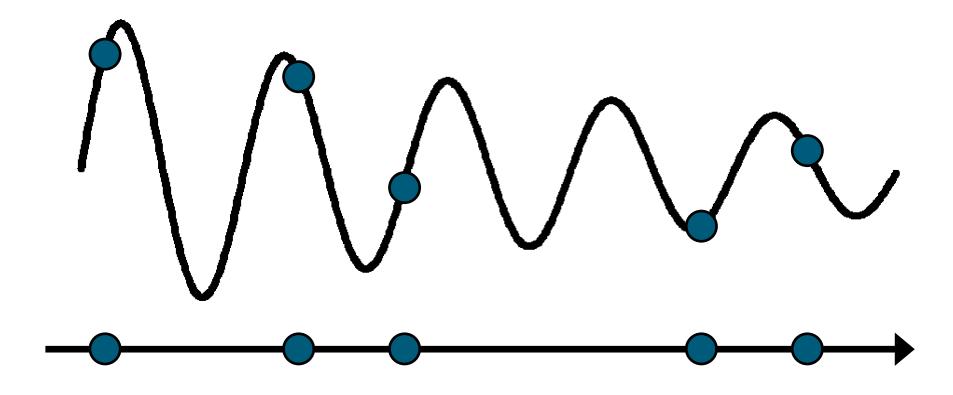
Let's try this:

```
library(ctsem)
data(Oscillating)
#interfacing to OpenMx
inits <- c(-39, -.3, 1.01, 10.01, .1, 10.01, 0.05, .9, 0)
names(inits) <- c("crosseffect", "autoeffect", "diffusion",
          "T0var11", "T0var21", "T0var22", "m1", "m2", 'manifestmean')
oscillatingm <- ctModel(n.latent = 2, n.manifest = 1, Tpoints = 11,
              MANIFESTVAR = matrix(c(0), nrow = 1, ncol = 1),
             LAMBDA = matrix(c(1, 0), nrow = 1, ncol = 2),
              TOMEANS = matrix(c('m1', 'm2'), nrow = 2, ncol = 1),
             TOVAR = matrix(c("TOvar11", "TOvar21", 0, "TOvar22"), nrow = 2, ncol = 2),
              DRIFT = matrix(c(0, "crosseffect", 1, "autoeffect"), nrow = 2, ncol = 2),
             CINT = matrix(0, ncol = 1, nrow = 2),
              MANIFESTMEANS = matrix('manifestmean', nrow = 1, ncol = 1),
              DIFFUSION = matrix(c(0, 0, 0, "diffusion"), nrow = 2, ncol = 2),
              startValues=inits, type="omx")
oscillating fit <- ctFit(Oscillating, oscillatingm)
summary(oscillating fit, verbose=T)
plot(oscillating fit)
```

Let's try this:

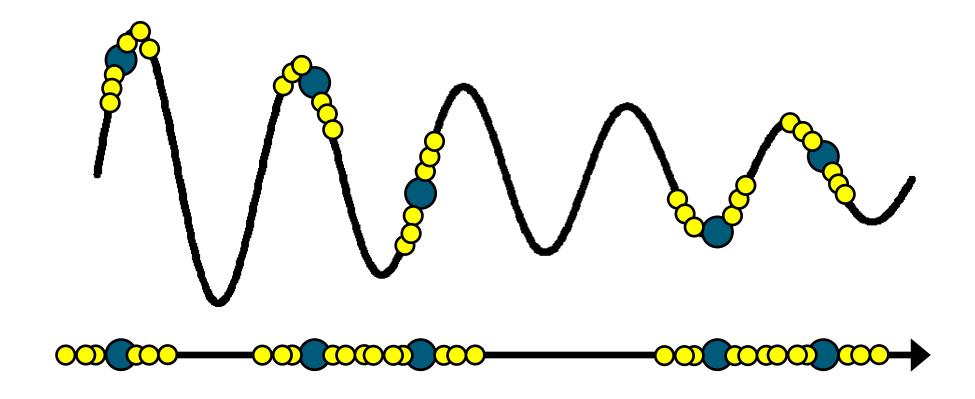
```
#interfacing to Stan (long-data format)
oscillationlong intervals <- ctWideToLong(datawide = Oscillating, Tpoints=11, n.manifest=1) #convert wide
to long format
oscillationlong <- ctDeintervalise(datalong = oscillationlong intervals, id='id', dT='dT') #convert intervals
to absolute time
#hist(oscillationlong[,2], ylab = "frequency", xlab="interval length", main="", col="gray")
oscillatingmodel stan <-ctModel(n.latent = 2, n.manifest=1, Tpoints=11,
                  MANIFESTVAR=matrix(c(0), nrow=1, ncol=1),
                  LAMBDA=matrix(c(1, 0), nrow=1, ncol=2),
                  DRIFT=matrix(c(0, "cross", 1, "auto"), nrow=2, ncol=2),
                  CINT=matrix(c(0,0), ncol=1, nrow=2, ),
                  DIFFUSION=matrix(c(0, 0, 0, "diffusion22"), nrow=2, ncol=2),
                  #startValues = inits,
                  type="stanct")
oscillatingmodel stan$pars$indvarying <- FALSE # no individual differences
oscillatingfit stan <-ctStanFit(oscillationlong, oscillatingmodel stan)
summary(oscillatingfit stan, verbose=T)
plot(oscillatingfit stan)
```

Unequal time intervals may help to detect oscillations with high(er) frequency, even if the sampling in time is low(er).



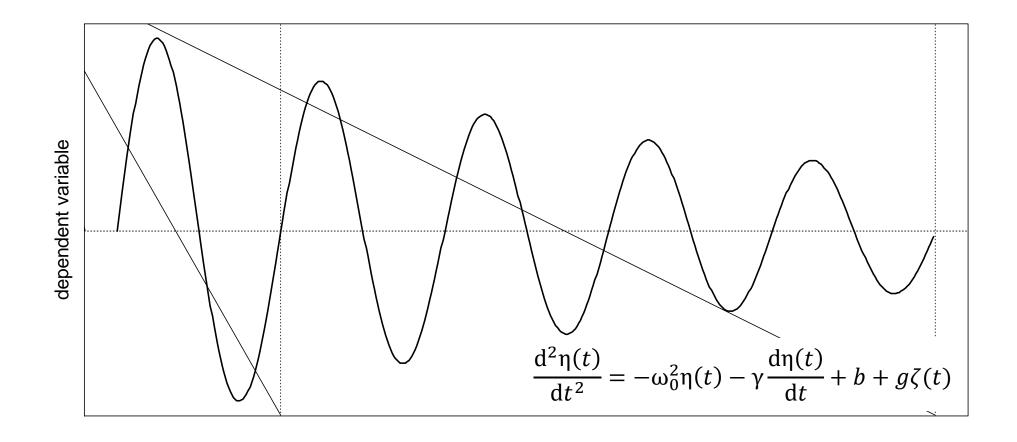
Time

Unequal time intervals may help to detect oscillations with high(er) frequency, even if the sampling in time is low(er).

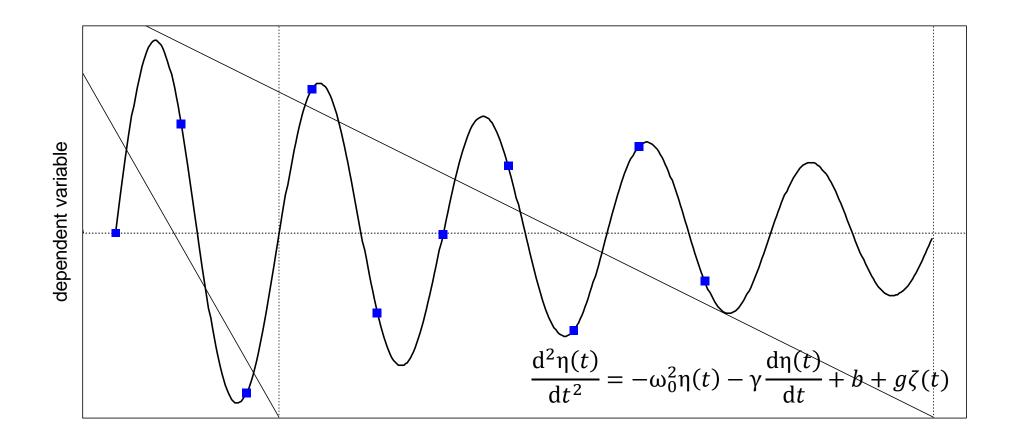


Time

To support this claim, consider the following situation: $\omega_0^2 = 39.48 \, [(2\pi)^2]$ and $\gamma = 0.5$:



Generate data for N = 200 and T = 11 with $\Delta t = 0.4$:

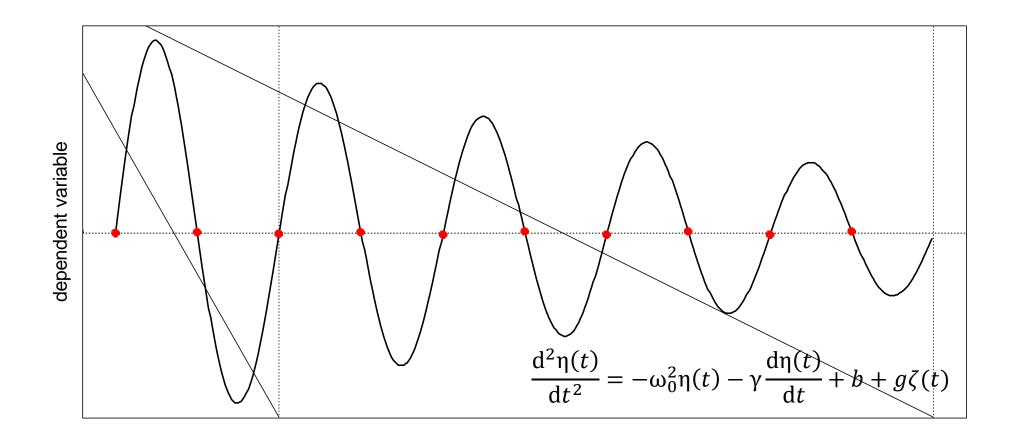


Fixed intervals:

Δt	0.4
------------	-----

Valid solutions	100%
$Mean(\omega_0^2)$	39.521
$SD(\omega_0^2)$	0.088
Mean(γ)	0.501
$SD(\gamma)$	0.012

Generate data for N = 200 and T = 11 with $\Delta t = 0.5$:



Fixed intervals:

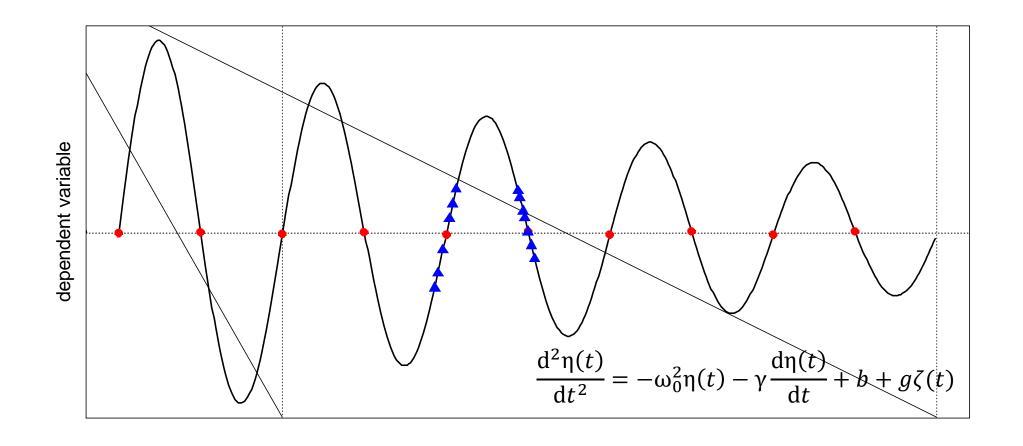
Δt 0.4 0.5

Valid solutions	100%	20% ^{a)}
$Mean(\omega_0^2)$	39.521	39.43
$SD(\omega_0^2)$	0.088	0.389
Mean(γ)	0.501	0.517
$SD(\gamma)$	0.012	0.068

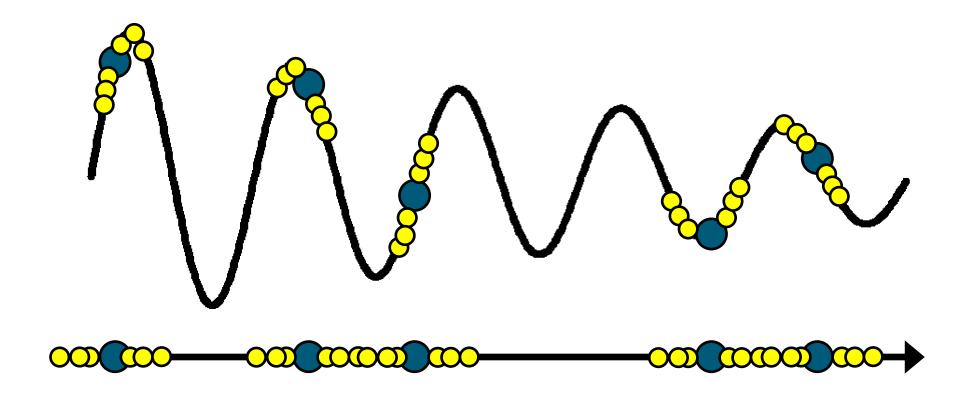
Fixed intervals:

Δt	0.4	0.5	1.0	1.0 no dampening	1.0 inferior SV
			NS-criterion not met		
Valid solutions	100%	20% a)	23% b)	19% ^{c)}	49% ^{d)}
$Mean(\omega_0^2)$	39.521	39.43	39.433	39.787	230.14
$SD(\omega_0^2)$	0.088	0.389	0.409	0.292	630.547
Mean(γ)	0.501	0.517	0.511	0	886.22
$SD(\gamma)$	0.012	0.068	0.04	0.024	2381.932

Generate data for N = 200 and T = 11 with $\Delta t_{ij} \sim N(0.4 \text{ or } 0.5 \text{ or } 1.0; 1/12)$:



Individually varying intervals:



Time

Individually varying intervals:

Valid solutions	100%
$Mean(\omega_0^2)$	39.529
$SD(\omega_0^2)$	0.091
Mean(γ)	0.501
SD(γ)	0.013

Individually varying intervals:

$$\Delta t$$
 $N\left(0.4, \frac{1}{12}\right) N\left(0.5, \frac{1}{12}\right)$

Valid solutions	100%	100%	
$Mean(\omega_0^2)$	39.529	39.558	
$SD(\omega_0^2)$	0.091	0.097	
Mean(γ)	0.501	0.502	
SD(γ)	0.013	0.012	

Individually varying intervals:

Δt	$N\left(0.4, \frac{1}{12}\right)$	$N\left(0.5, \frac{1}{12}\right)$	$N\left(1,\frac{1}{2}\right)$	$N\left(1,\frac{1}{2}\right)$ no dampening	$N\left(1, \frac{1}{2}\right)$ inferior SV
			NS-criterion not met		
Valid solutions	100%	100%	99% ^{e)}	56% ^{f)}	7% g)
$\overline{\text{Mean}(\omega_0^2)}$	39.529	39.558	39.969	39.993	39.956
$SD(\omega_0^2)$	0.091	0.097	0.091	0.038	0.109
Mean(γ)	0.501	0.502	0.51	0	0.512
SD(γ)	0.013	0.012	0.012	0.006	0.011

- Latent Change Score (LCS) models are rather popular in some fields of psychology
- LCS are particularly useful for pre-post data (two wave) data
- For more waves, standard models, such as the
 - Proportional Change Score Model, or
 - Dual Change Score Model
 - are (usually) just restricted approximations to true ct models
- In simple cases (of equal time intervals) ct and LCS models yield identical results, for more complex situations ct models are superior
- ...what follows is a quick demonstration to support this argument.

 \triangleright Latent change score model for T=2

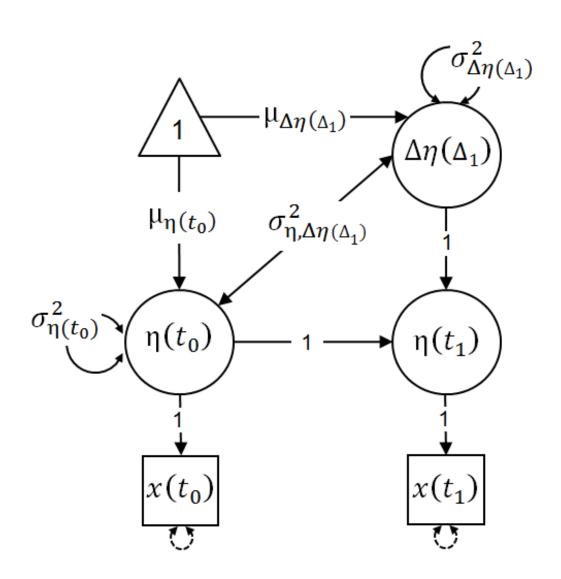
$$\Delta \eta(t_u - t_{u-1}) = \eta(t_u) - \eta(t_{u-1}) = \Delta \eta(\Delta_u)$$

- > Advantage: Separation of true variance and error variance
- For $\Delta_u = t_u t_{u-1} = 1$, the (latent) difference between two time points is equal to the rate of change

$$\frac{\Delta \eta(\Delta_u)}{\Delta_u}$$

So that

$$\eta(t_u) = 1 \cdot \eta(t_{u-1}) + 1 \cdot \Delta \eta$$



- ➤ For T > 2 we can not only separate true from error variance but begin to impose a structure among the latent variables
- > The proportional change score (PCS) model is one example

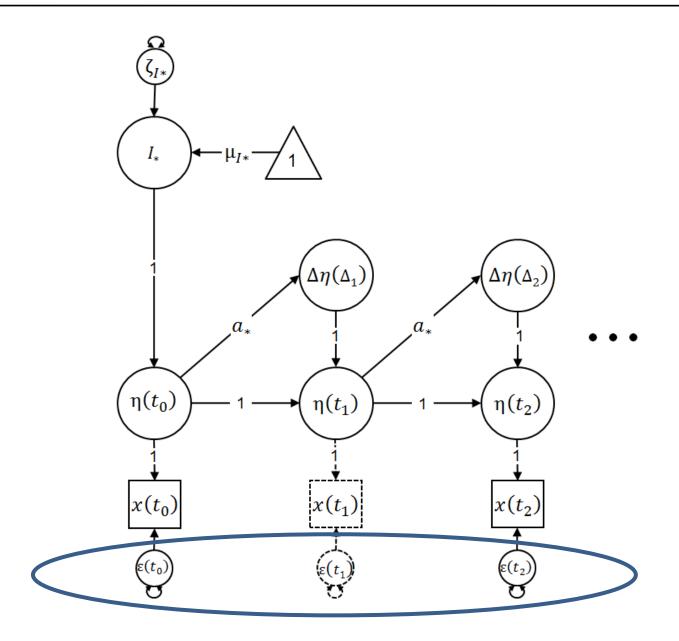
$$\Delta \eta = \beta \cdot \eta(t_{u-1})$$

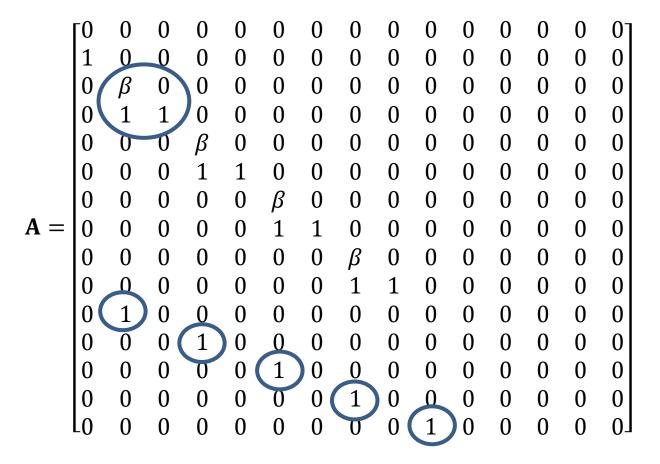
> The dual change score (DCS) model adds an additional "slope"

$$\Delta \eta = \beta \cdot \eta(t_{u-1}) + S_*$$

> Further distinction between measurement and dynamic error yields:

$$\Delta \mathbf{\eta} = \mathbf{A}_* \cdot \mathbf{\eta}(t_{u-1}) + \mathbf{S}_* + \mathbf{\zeta}_*$$





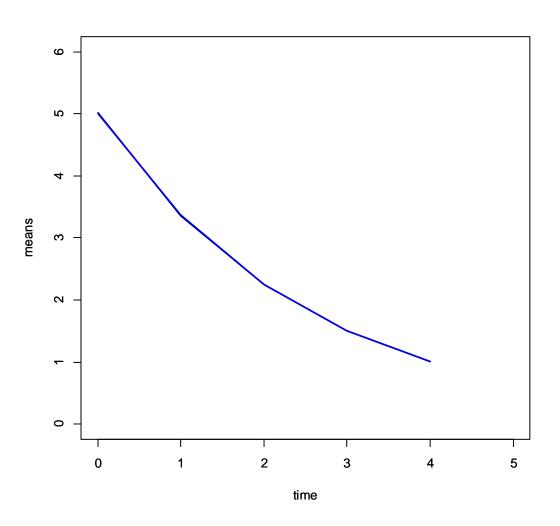
```
R script (with OpenMx and Lavaan Code):

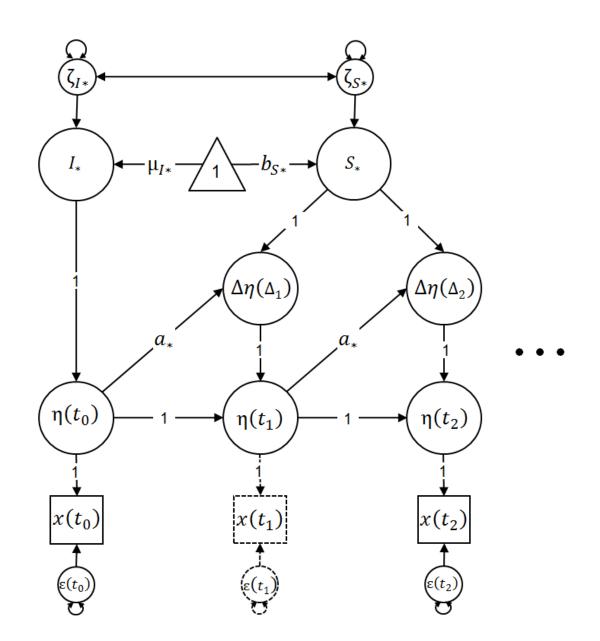
1_PCS_model_OpenMx_Lavaan.R

Data:

1_data_PCSmodel.dat
```

```
free parameters:
    name matrix row col
                          Estimate
                                     Std.Error A
                      2 -0.3304290 0.001382494
                        1.0028719 0.049993041
   var1
                11 11 0.1985027 0.004438655
3 var(e)
                         5.0146682 0.034204305
  meta1
Model Statistics:
                                 Degrees of Freedom | Fit (-21nL units)
                  Parameters
      Model:
                                                                  8406.41
                                               4996
                                               4980
   Saturated:
                          20
                                                                       NA
Independence:
                          10
                                               4990
                                                                       NA
Number of observations/statistics: 1000/5000
Information Criteria:
         df Penalty
                        Parameters Penalty |
                                               Sample-Size Adjusted
AIC:
           -1585.59
                                  8414.410
                                                           8414.450
          -26104.74
                                  8434.041
                                                           8421.337
BIC:
```







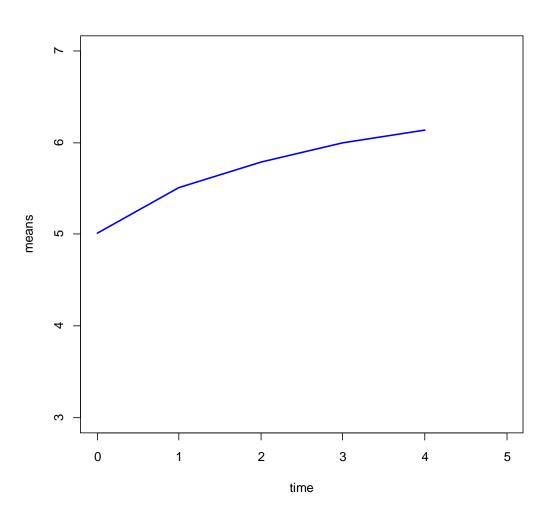
```
R script (with OpenMx and Lavaan Code):

2_DCS_model_OpenMx_Lavaan.R

Data:

2_data_DCSmodel.dat
```

```
free parameters:
                           Estimate
     name matrix row col
                                      Std.Error A
                      3 -0.33426495 0.006656825
2 var(S)
                     1 1.73772734 0.083646333
                2 1 0.05610959 0.043449833
3 cov(1S)
              S 2 2 0.91148597 0.047891566
     var1
              S 12 12 0.20355572 0.005255795
  var(e)
                 1 S 2.14360572 0.056047387
                      I 5.01742658 0.032716777
   meta1
Model Statistics:
                                Degrees of Freedom | Fit (-21nL units)
                  Parameters
       Model:
                                              4993
                                                                12957.27
   Saturated:
                          20
                                              4980
                                                                      NA
Independence:
                                              4990
                                                                      NA
Number of observations/statistics: 1000/5000
Information Criteria:
        df Penalty
                       Parameters Penalty
                                              Sample-Size Adjusted
AIC:
           2971.272
                                 12971.27
                                                          12971.38
         -21533.150
                                                          12983.39
BIC:
                                 13005.63
```



Dual change score model with measurement error and dynamic error for T = 5

```
R script (with OpenMx and Lavaan Code):

3_DCS_model_OpenMx_Lavaan.R

Data:

3_data_DCSmodel.dat
```

➤ Dual change score model with measurement error <u>and</u> dynamic error for T = 5

```
free parameters:
                            Estimate Std.Error A
     name matrix row col
                       3 -0.31873809 0.02453309
  var(S)
                          1.68148341 0.16702023
                  2 1 0.06780329 0.05194576
3 cov(1S)
     var1
                          0.93882555 0.08146233
   var(z)
                          2.01040764 0.12547098
   var(e)
                 12 12 0.20174170 0.06882644
                          2.02171459 0.14367425
       mS
                          5.01101731 0.03371202
    meta1
Model Statistics:
                                 Degrees of Freedom
                                                        Fit (-21nL units)
                  Parameters
       Model:
                                               4992
                                                                 19082.67
                          20
                                               4980
   Saturated:
                                                                       NA
Independence:
                                               4990
                                                                       NA
Number of observations/statistics: 1000/5000
Information Criteria:
         df Penalty
                        Parameters Penalty
                                               Sample-Size Adjusted
           9098.666
                                  19098.67
                                                           19098.81
AIC:
         -15400.848
                                  19137.93
                                                           19112.52
BIC:
```

Dual change score model reformulated as a latent AR model

For $\Delta \eta_i(\Delta_{j,i}) = \eta_i(t_{j,i}) - \eta_i(t_{j-1,i})$, the dual change score model is:

$$\mathbf{\eta}_i(t_{j,i}) = (\mathbf{A}_* \cdot \Delta_{j,i} + \mathbf{I}) \cdot \mathbf{\eta}_i(t_{j-1,i}) + \Delta_{j,i} \cdot \mathbf{S}_{*i} + \mathbf{\zeta}_{*i},$$

which corresponds to a vector autoregressive and cross-lagged (ARCL) panel model of the general form

$$\mathbf{\eta}_i(t_{j,i}) = \mathbf{A}(\Delta_{j,i}) \cdot \mathbf{\eta}_i(t_{j-1,i}) + \mathbf{S}_i(\Delta_{j,i}) + \mathbf{\zeta}_i(\Delta_{j,i}),$$

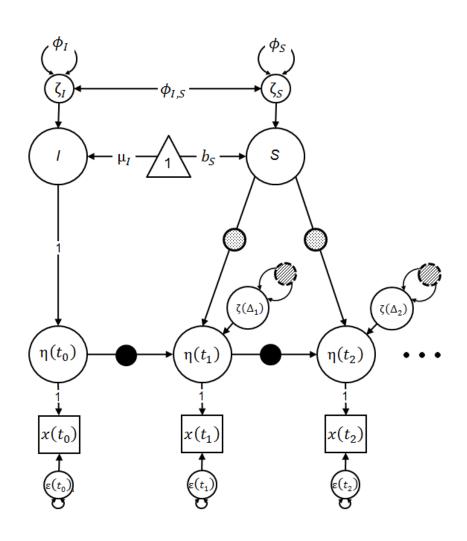
with

$$\mathbf{A}(\Delta_{j,i}) = (\mathbf{A}_* \cdot \Delta_{j,i} + \mathbf{I}),$$

$$\mathbf{S}_{i}(\Delta_{j,i}) = \Delta_{j,i} \cdot \mathbf{S}_{*i}$$
, and

$$\zeta_i(\Delta_{j,i}) = \zeta_{*i}.$$

Dual change score model reformulated as a latent AR model



Latent Change Score Model:

$$= a(\Delta_{j,i}) = a_* \cdot \Delta_{j,i} + 1$$

$$\bigcirc$$
 = $\Delta_{j,i}$

$$= q(\Delta_{j,i}) = q_*$$

We have already seen how to formulate a continuous time model as a latent AR model

$$\mathbf{\eta}_i(t_{j,i}) = \mathbf{e}^{\mathbf{A} \cdot \Delta_{j,i}} \mathbf{\eta}_i(t_{0,i}) + \mathbf{A}^{-1} \left[\mathbf{e}^{\mathbf{A} \cdot \Delta_{j,i}} - \mathbf{I} \right] \mathbf{S}_i + \int_{t_0}^{t_{j,i}} \mathbf{e}^{\mathbf{A} \cdot (t_{j,i} - s)} \mathbf{G} d\mathbf{W}_i(s)$$

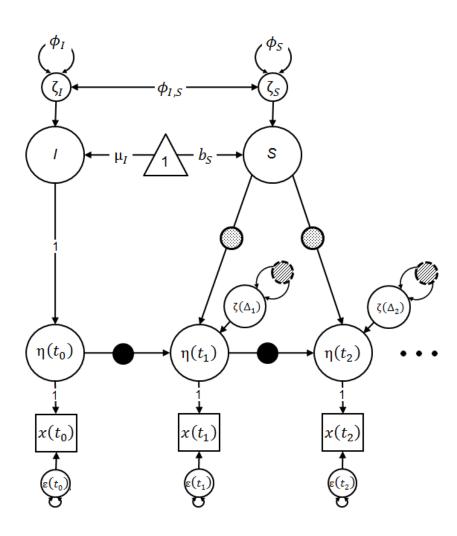
This corresponds to a vector autoregressive and cross-lagged panel model with

$$\mathbf{A}(\Delta_{j,i}) = \mathbf{e}^{\mathbf{A} \cdot \Delta_{j,i}},$$

$$\mathbf{S}_i(\Delta_{j,i}) = \mathbf{A}^{-1} [\mathbf{e}^{\mathbf{A}\cdot\Delta_{j,i}} - \mathbf{I}]\mathbf{S}_i$$
, and

$$\zeta_i(\Delta_{j,i}) = \int_{t_0}^{t_{j,i}} e^{\mathbf{A} \cdot (t_{j,i} - s)} \mathbf{G} d\mathbf{W}_i(s)$$
.

Continuous time model reformulated as a latent AR model



Continuous Time Model:

$$= a(\Delta_{j,i}) = e^{a \cdot \Delta_j}$$

$$= a^{-1}(e^{a\cdot\Delta_{j,i}}-1)$$

$$= q(\Delta_{j,i}) = (a+a)^{-1} (e^{(a+a)\cdot \Delta_{j,i}} - 1) \cdot q$$

> Thus, there is a direct relationship between latent change score and continuous time models:

Parameter	Latent Change Score Model	Continuous Time Model
$\mathbf{A}(\mathbf{\Delta}_{j,i})$	$= (\mathbf{A}_* \cdot \Delta_{j,i} + \mathbf{I})$	$=\mathrm{e}^{\mathbf{A}\cdot\Delta_{j,i}}$
$\mathbf{S}_i(\mathbf{\Delta}_{j,i})$	$=\Delta_{j,i}\cdot\mathbf{S}_{*i},$	$= \mathbf{A}^{-1} \big[e^{\mathbf{A} \cdot \Delta_{j,i}} - \mathbf{I} \big] \mathbf{S}_i$
$oldsymbol{\zeta}_iig(oldsymbol{\Delta}_{j,i}ig)$	$= \zeta_{*i}$	$= \int_{t_0}^{t_{j,i}} e^{\mathbf{A} \cdot (t_{j,i} - s)} \mathbf{G} d\mathbf{W}_i(s)$

➤ The relation between LCS and CT models is best understood when putting the matrix exponential constraint of the CT model in power series expansion:

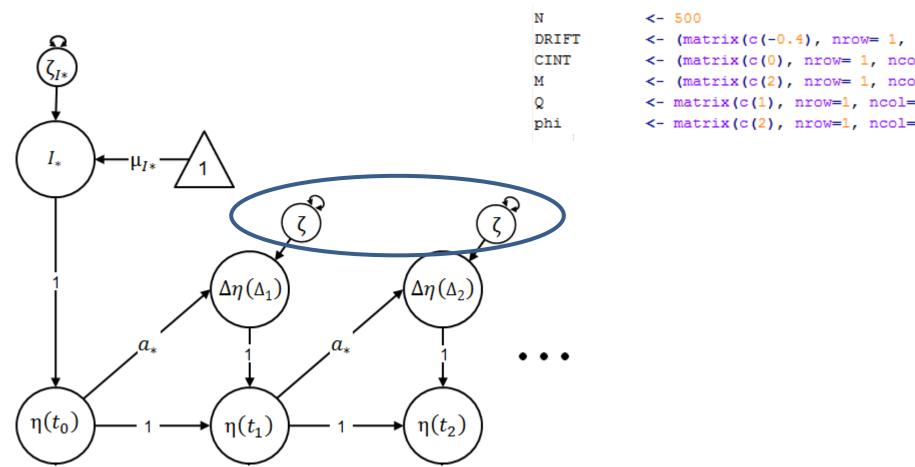
$$\mathbf{A}(\Delta_{j,i}) = e^{\mathbf{A}\cdot\Delta_{j,i}} = (\mathbf{I} + \mathbf{A}\cdot\Delta_{j,i}) + \frac{1}{2!}\mathbf{A}^2\cdot\Delta_{j,i}^2 + \frac{1}{3!}\mathbf{A}^3\cdot\Delta_{j,i}^3\cdots$$

From theory to practice...

...let's have another example:

- 1. Generate data according to a known dynamic model with equal time intervals (T = 5)
- 2. Fit a latent change score model to the data
- 3. Replicate results using a CT model
- Generate data according to a known dynamic model with individually varying time intervals
- Fit a CT model to the data
- 6. Try to fit a latent change score model to the data

1. Generate data according to a known dynamic model with equal time intervals (T = 5)



2. Fit a latent change score model to the data

```
R script (with OpenMx and Lavaan Code):

4a_PCS_model_OpenMx_Lavaan.R

Data:

4_sim_univariate_delta_1.dat
```

2. Fit a latent change score model to the data

```
free parameters:
    name matrix row col
                         Estimate Std.Error A
                     2 -0.3368296 0.01020483
             S 1 1 1.9845720 0.12551366
   var1
             S 3 0.6729965 0.02128202
3 var(z)
4 meta1
                     I 1.9915945 0.06300088
Model Statistics:
                                Degrees of Freedom | Fit (-21nL units)
                 Parameters
      Model:
                                                               6645.364
                                              2496
  Saturated:
                         20
                                              2480
                                                                     NA
Independence:
                         10
                                              2490
                                                                     NA
Number of observations/statistics: 500/2500
Information Criteria:
                                             Sample-Size Adjusted
        df Penalty
                       Parameters Penalty
          1653.364
AIC:
                                 6653.364
                                                         6653.445
BIC:
         -8866.298
                                 6670.222
                                                         6657,526
```

3. Replicate results using a CT model

```
R script:

4b_CT_model_OpenMx.R

Data:

4_sim_univariate_delta_1.dat
```

3. Replicate results using a CT model

Remember the relationship between the continuous and discrete time parameters:

$$\exp(-0.4107233) = 0.6631704 = 1-0.3368297$$
 and $((a+a)^{**}-1)^{*}(\exp(a+a)-1)^{*}(g^{**}2) = 0.6729962$



4. Generate data according to a known dynamic model with individually varying time intervals. We use the same model as before.

$$\Delta_{u,i} \sim \text{U}(0,4]$$

for all *u* and *i*.

5. Fit a CT model to the data

```
R script:

5a_CT_model_OpenMx.R

Data:

5_sim_univariate_deltavarying.dat
```

5. Fit a CT model to the data

6. Try to fit a latent change score model to the data

```
R script (with OpenMx and Lavaan Code):

5b_PCS_model_OpenMx_Lavaan.R

Data:

5_sim_univariate_deltavarying.dat
```

6. Try to fit a latent change score model to the data

```
free parameters:
    name matrix row col
                           Estimate Std.Error A
                       2 -0.6226777 0.01420925
2 var1
3 var(z)
              S 1 1 2.0789260 0.13148419
S 3 3 1.1104069 0.03511424
                          2.0305927 0.06448176
 meta1
Model Statistics:
                                   Degrees of Freedom |
                   Parameters |
                                                           Fit (-2]nL units)
                                                                     7670.072
                                                  2496
       Model:
   Saturated:
                                                  2480
                           20
Independence:
                           10
                                                  2490
                                                                           NA
Number of observations/statistics: 500/2500
Information Criteria:
         df Penalty |
                         Parameters Penalty |
                                                 Sample-Size Adjusted
AIC:
           2678.072
                                    7678.072
                                                               7678, 152
          -7841.590
                                    7694.930
                                                               7682,234
BIC:
```

Also:

 $\exp(-0.4083978) = 0.6647144$ IS NOT IDENTICAL TO

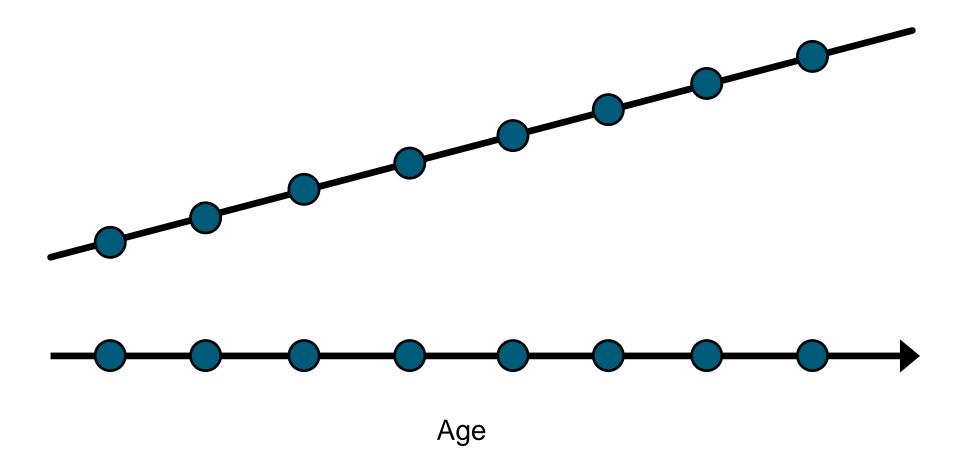
1-0.6226777 = 0.3773223



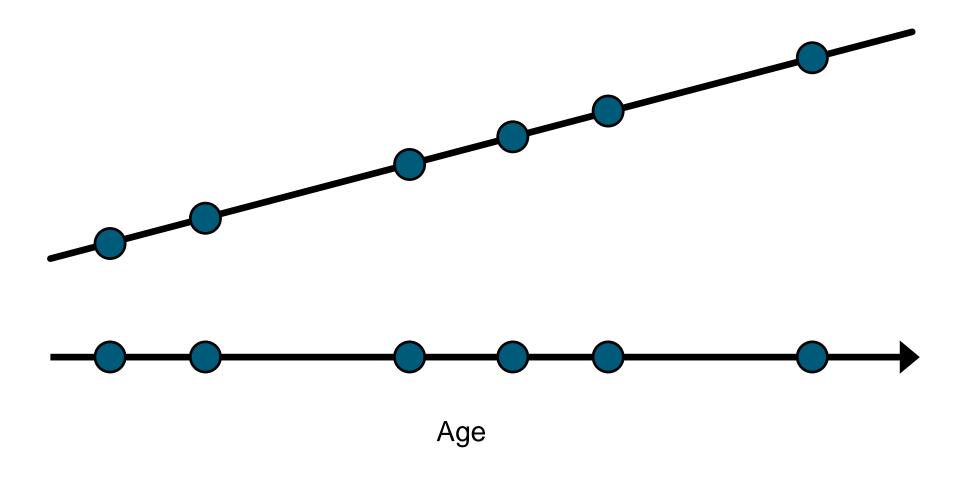
Summary Latent Change Score models

- > As dynamic models, LCS models are powerful models for the analysis of change.
- ightharpoonup For $\Delta t = 1$, the rate of change equals the latent difference.
- > LCS models can be reformulated as (multivariate) autoregressive models.
- > CT models manifest themselves as autoregressive models in discrete time.
- \triangleright LCS models are special cases of the more general class of CT models. In case of $\Delta t = 1$, discrete time parameter estimates of CT and LCS models are "equivalent". For unequal Δt , CT models should be used.

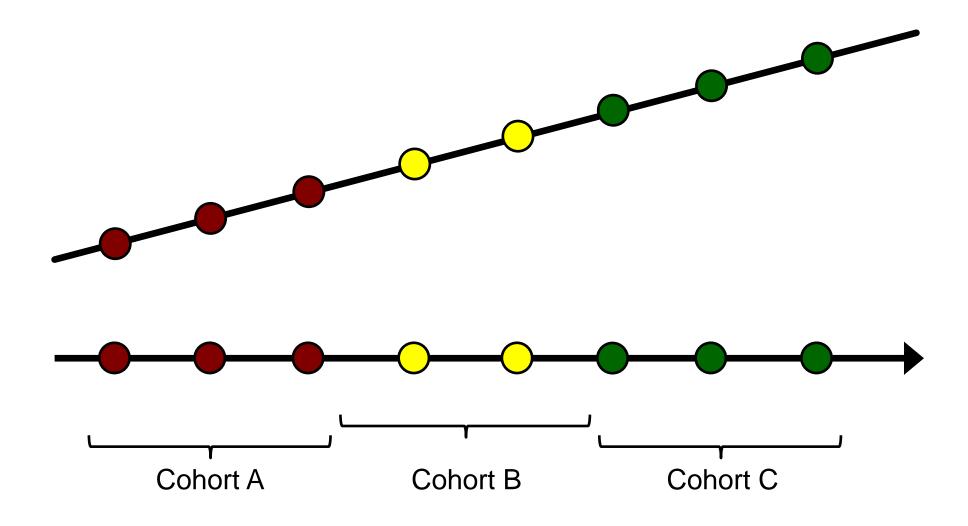
Single cohort equally spaced:



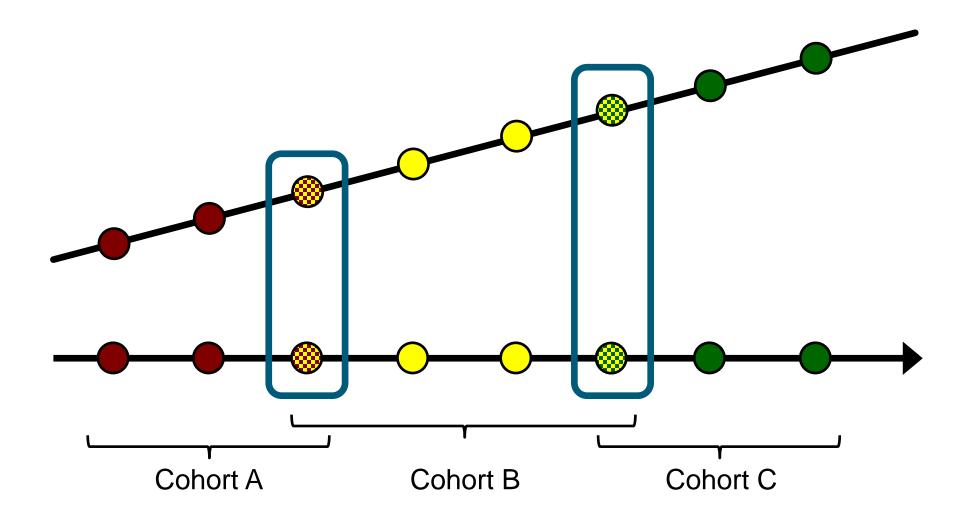
Single cohort unequally spaced:



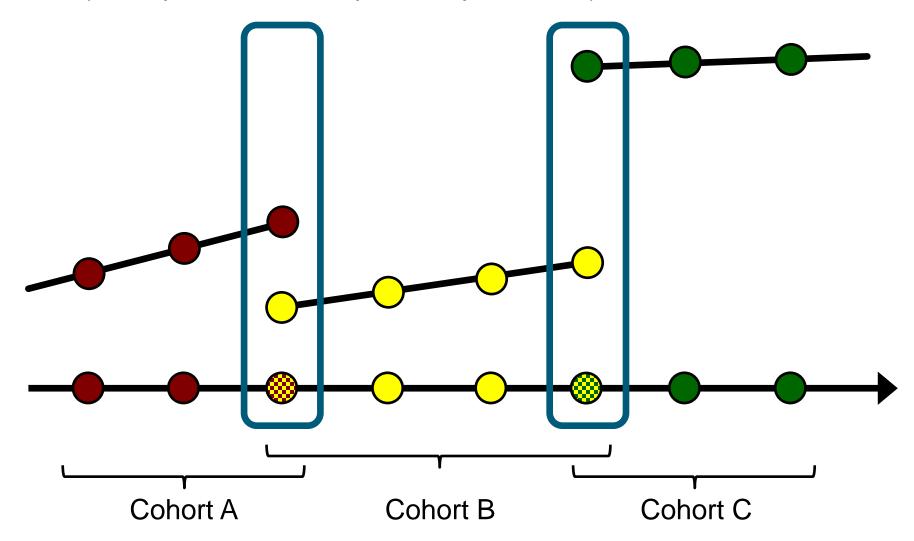
Multiple cohorts (equivalent developmental process):



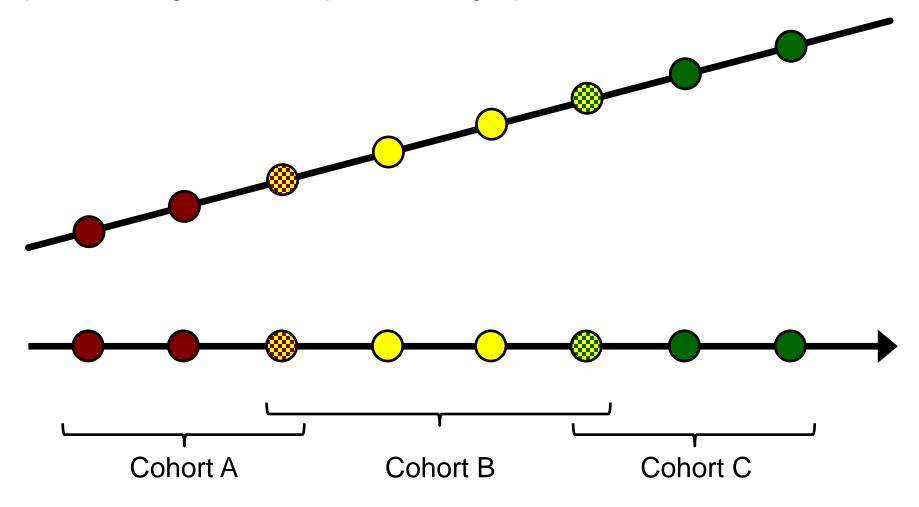
Multiple cohorts (testing for equivalence):

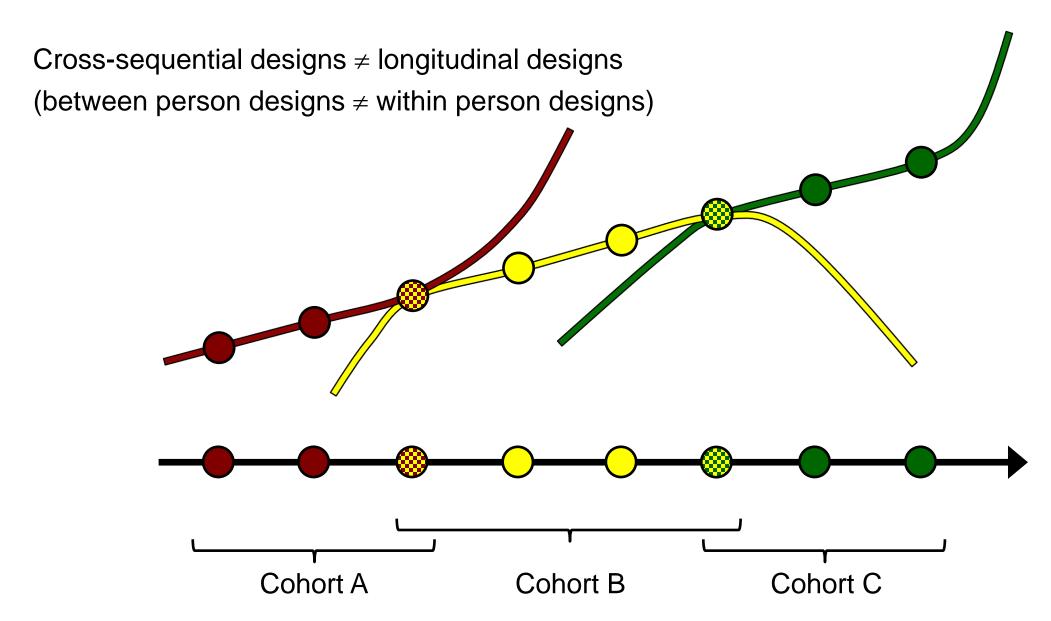


Multiple cohorts (nonequivalent developmental processes):



Cross-sequential designs ≠ longitudinal designs (between person designs ≠ within person designs)





Renewed interest in accelerated longitudinal designs (in continuous time), e.g.,

Cáncer, P. F., Estrada, E., & Ferrer, E. (2023). A Dynamic Approach to Control for Cohort Differences in Maturation Speed Using Accelerated Longitudinal Designs. *Structural Equation Modeling: A Multidisciplinary Journal*, 1-17. https://doi.org/10.1080/10705511.2022.2163647

Estrada, E., Bunge, S. A., & Ferrer, E. (2023). Controlling for cohort effects in accelerated longitudinal designs using continuous- and discrete-time dynamic models. *Psychological Methods*, *28*(2), 359–378. https://doi.org/10.1037/met0000427

Estrada, E., & Ferrer, E. (2019). Studying developmental processes in accelerated cohort-sequential designs with discrete- and continuous-time latent change score models. *Psychol Methods*, *24*(6), 708-734. https://doi.org/10.1037/met0000215

Let's implement a continuous time cross-sequential design using *ctsem* along a blog-post by Charlie.

For scripts and details, see:

https://cdriver.netlify.app/post/accelerated/

09a_CT_cross_sequential.R

For an R-script using the multi-group functionality in ctsemOMX see :

09b_CT_cross_sequential.R

References

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- Estrada, E., & Ferrer, E. (2019). Studying developmental processes in accelerated cohort-sequential designs with discrete- and continuous-time latent change score models. *Psychol Methods*, *24*(6), 708-734. https://doi.org/10.1037/met0000215
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- Molenaar, P. C. M., & Nesselroade, J. R. (2009). The recoverability of P-technique factor analysis. Multivariate Behavioral Research, 44(1), 130-141. https://doi.org/10.1080/00273170802620204
- Oud, J. H. L., Voelkle, M. C., & Driver, C. C. (2018). SEM Based CARMA Time Series Modeling for Arbitrary *N*. Multivariate Behavioral Research, 53(1), 36-56. https://doi.org/10.1080/00273171.2017.1383224
- Voelkle, M. C., & Oud, J. H. L. (2013). Continuous time modelling with individually varying time intervals for oscillating and non-oscillating processes. British Journal of Mathematical and Statistical Psychology, 103-126. https://doi.org/10.1111/j.2044-8317.2012.02043.x

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Study Questions

Question 1:

Compare P-technique and Dynamic Factor Analysis for time series data from a single subject. Why might one be preferred over the other, and under what conditions?

Question 2:

Explain the structure and utility of CARMA(p, q) models. What psychological or real-world phenomena might benefit from emplying CARMA(2,1) models?

Question 3:

Describe a damped linear oscillator in the context of ctsem. How are they detected and what psychological constructs might show this behavior?

Study Questions

Question 4:

Compare latent change score (LCS) and continuous-time models. Under what conditions do they yield the same results, and when should a CT model be preferred?

Question 2:

Explain the structure of accelerated longitudinal designs. How can ctsem be used to disentangle cohort effects from developmental processes?