## ECON 710A - Problem Set 4

Alex von Hafften\*

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1. Let X be generated by the following random coefficients discrete choice model  $X = 1\{-U_0 + ZU_1 > 0\}$  where  $U = (U_0, U_1)'$  is independent of Z and  $Z \in \{0, 1\}$ . Provide conditions on U such that Pr(Defying) = 0 and Pr(Complying) > 0.

 $\begin{array}{lll} Pr(Defying) = 0 \text{ iff } X_U(1) = 0 & \Longrightarrow & X_U(0) = 0 \text{ and } X_U(0) = 1 & \Longrightarrow & X_U(1) = 1. & X_U(1) = 0 & \Longrightarrow \\ X_U(0) = 0 \text{ iff } -U_0 + (1)U_1 = -U_0 + U_1 < 0 & \Longrightarrow & -U_0 + (0)U_1 = -U_0 < 0. & X_U(0) = 1 & \Longrightarrow & X_U(1) = 1 \text{ iff } \\ -U_0 + (0)U_1 = -U_0 > 0 & \Longrightarrow & -U_0 + (1)U_1 = -U_0 + U_1 > 0. & \text{Thus, } U_1 \geq 0. \end{array}$ 

 $Pr(Complying) > 0 \iff Pr(X_U(1) = 1 \text{ and } X_U(0) = 0) > 0.$  Since  $U_1 \ge 0$ , this implies that  $U_1 > U_0 \ge 0$ .

- 2. Let  $\{Y_t\}_{t=1}^T$  be generated by the following MA(q) model, i.e.,  $Y_t = \mu + \varepsilon_t + \theta_1 \varepsilon_{t-1} + \dots + \theta_q \varepsilon_{t-q}$  where  $\{\varepsilon_t\}_{t=0}^T$  are i.i.d. random variables with mean zero and variance  $\sigma^2$ .
- (i) Find the autocovariance function  $\gamma(k)$ .

For k = 0:

$$\begin{split} \gamma(0) &= Cov(Y_t, Y_t) \\ &= Var(Y_t) \\ &= Var(\mu + \varepsilon_t + \theta_1 \varepsilon_{t-1} + \ldots + \theta_q \varepsilon_{t-q}) \\ &= Var(\varepsilon_t) + \theta_1^2 Var(\varepsilon_{t-1} + \ldots + \theta_q^2 Var(\varepsilon_{t-q}) \\ &= \sigma^2 + \theta_1^2 \sigma^2 + \ldots + \theta_q^2 \sigma^2 \\ &= \sigma^2 (1 + \theta_1^2 + \ldots + \theta_q^2) \end{split}$$

For k = 1:

$$\begin{split} \gamma(1) &= Cov(Y_t, Y_{t+1}) \\ &= Cov(\mu + \varepsilon_t + \theta_1 \varepsilon_{t-1} + \ldots + \theta_q \varepsilon_{t-q}, \mu + \varepsilon_{t+1} + \theta_1 \varepsilon_{t+1-1} + \ldots + \theta_q \varepsilon_{t+1-q}) \\ &= Cov(\varepsilon_t + \theta_1 \varepsilon_{t-1} + \ldots + \theta_q \varepsilon_{t-q}, \varepsilon_{t+1} + \theta_1 \varepsilon_t + \ldots + \theta_q \varepsilon_{t+1-q}) \\ &= Cov(\varepsilon_t + \theta_1 \varepsilon_{t-1} + \ldots + \theta_{q-1} \varepsilon_{t+1-q}, \theta_1 \varepsilon_t + \ldots + \theta_q \varepsilon_{t+1-q}) \\ &= \theta_1 Var(\varepsilon_t) + \theta_1 \theta_2 Var(\varepsilon_{t-1}) + \ldots + \theta_{q-1} \theta_q Var(\varepsilon_{t+1-q}) \\ &= \sigma^2(\theta_1 + \theta_1 \theta_2 + \ldots + \theta_{q-1} \theta_q) \end{split}$$

<sup>\*</sup>I worked on this problem set with a study group of Michael Nattinger, Andrew Smith, and Ryan Mather. I also discussed problems with Sarah Bass, Emily Case, Danny Edgel, and Katherine Kwok.

For general k:

$$\gamma(k) = \begin{cases} \sigma^2(1 + \theta_1^2 + \dots + \theta_q^2) & \text{if } k = 0\\ \sigma^2(\theta_k + \theta_1\theta_{k+1} + \theta_2\theta_{k+1} + \dots + \theta_{q-k}\theta_q) & \text{if } 0 < k \le q\\ 0, & \text{if } k > q \end{cases}$$

(ii) Suppose that q=1 and find the autocorrelation function,  $\rho(k)=\frac{\gamma(k)}{\gamma(0)}.$  If q=1:

$$Y_t = \mu + \varepsilon_t + \theta_1 \varepsilon_{t-1}$$

 $\gamma(k)$  simplies to

$$\gamma(k) = \begin{cases} \sigma^2(1 + \theta_1^2), & \text{if } k = 0\\ \sigma^2\theta_1, & \text{if } k = 1\\ 0, & \text{if } k > 1 \end{cases}$$

The autocorrelation function:

$$\rho(k) = \frac{\gamma(k)}{\gamma(0)}$$

$$= \begin{cases} \frac{\sigma^2(1+\theta_1^2)}{\sigma^2(1+\theta_1^2)}, & \text{if } k = 0\\ \frac{\sigma^2\theta_1}{\sigma^2(1+\theta_1^2)}, & \text{if } k = 1\\ \frac{0}{\sigma^2(1+\theta_1)}, & \text{if } k > 1 \end{cases}$$

$$= \begin{cases} 1, & \text{if } k = 0\\ \frac{\theta_1}{1+\theta_1^2}, & \text{if } k = 1\\ 0, & \text{if } k > 1 \end{cases}$$

(iii) Is  $\theta_1$  identified from the autocorrelation function?

No. First, notice that  $\theta_1$  only appears in autocorrelation function when k = 1. If  $\theta_1 = x$ ,  $\theta_1 = 1/x$  yields the same value from the autocorrelation function:

$$\rho(1|\theta_1 = x) = \frac{x}{1+x^2}$$

$$\rho(1|\theta_1 = x^{-1}) = \frac{x^{-1}}{1+(x^{-1})^2}$$

$$= \frac{x^{-1}}{1+x^{-2}} \frac{x^2}{x^2}$$

$$= \frac{x}{1+x^2}$$

(iv) Suppose  $\theta_1 \in [-1, 1]$ . Does you answer to (iii) change?

Yes,  $\theta_1$  is identified because if  $\theta_1 = x \in [-1, 1] \implies 1/x \notin [-1, 1]$ .

- 3. Consider an ARMA(1,1) model:  $Y_t = \alpha_0 + Y_{t-1}\rho + U_t$  and  $U_t = \varepsilon_t + \theta\varepsilon_{t-1}$  for all t = 1, ..., T;  $Y_0 = \mu + \varepsilon_0 + \nu$  where  $|\rho| < 1$ ,  $|\theta| \le 1$ ,  $\varepsilon_0, ..., \varepsilon_T$  are idd  $N(0, \sigma^2)$  and independent of  $\nu \sim N(0, \tau)$ .
- (i) Find  $\mu$  and  $\tau$  (as functions of  $\alpha_0, \rho, \theta$ , and/or  $\sigma^2$ ) such that  $E[Y_t]$  and  $Var(Y_t)$  does not depend on t.

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(ii) For the  $\mu$  and  $\tau$  found above, you may use without proof that  $\{Y_t\}_{t=1}^T$  is covariance stationary. Under what conditions on  $\alpha_0, \rho, \theta$ , and/or  $\sigma^2$  is  $(1, Y_{t-2})$  a valid instrument for  $(1, Y_{t-1})$ .

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