

# ECON 709 - PS 1

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- (1) For two events  $A, B \in S$ , prove that  $A \cup B = (A \cap B) \cup ((A \cap B^C) \cup (B \cap A^C))$ .

Proof: Applying the partition rule and the properties of set operators,

$$\begin{aligned} A \cup B &= ((A \cap B) \cup (A \cap B^C)) \cup B \\ &= ((A \cap B) \cup (A \cap B^C)) \cup ((B \cap A) \cup (B \cap A^C)) \\ &= ((A \cap B) \cup (B \cap A)) \cup ((A \cap B^C) \cup (B \cap A^C)) \\ &= (A \cap B) \cup ((A \cap B^C) \cup (B \cap A^C)) \end{aligned}$$

□

- (2) Prove that  $P(A \cup B) = P(A) + P(B) - P(A \cap B)$ .

Proof: Applying the partition rule and axioms of the probability measure,

$$\begin{aligned} P(A \cup B) &= P(((A \cap B) \cup (A \cap B^C)) \cup ((B \cap A) \cup (B \cap A^C))) \\ &= P((A \cap B) \cup (A \cap B^C) \cup (B \cap A^C)) \\ &= P(A \cap B) + P(A \cap B^C) + P(B \cap A^C) \\ &= P(A \cap B) + P(A \cap B^C) + P(B \cap A^C) + P(A \cap B) - P(A \cap B) \\ &= P((A \cap B) \cup (A \cap B^C)) + P((B \cap A^C) \cup (A \cap B)) - P(A \cap B) \\ &= P(A) + P(B) - P(A \cap B) \end{aligned}$$

□

- (3) Suppose that the unconditional probability of a disease is 0.0025. A screening test for this disease has a detection rate of 0.9, and has a false positive rate of 0.01. Given that the screening test returns positive, what is the conditional probability of having the disease?

Proof: Define  $A$  the event that you have the the disease, so  $P(A) = 0.0025$  and  $P(A^C) = 0.9975$ . Define  $B$  as the event you test positive, so  $P(B|A) = 0.9$  and  $P(B|A^C) = 0.01$ . We want to know the probability you have the disease conditional on a positive test result, or  $P(A|B)$ . Thus, using Bayes Rule:

$$P(A|B) = \frac{P(B|A)P(A)}{P(B|A)P(A) + P(B|A^C)P(A^C)} = \frac{(0.9)(0.0025)}{(0.9)(0.0025) + (0.01)(0.9975)} \approx 0.184$$

□

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- (4) Suppose that a pair of events  $A$  and  $B$  are mutually exclusive, i.e.,  $A \cap B = \emptyset$ , and that  $P(A) > 0$  and  $P(B) > 0$ . Prove that  $A$  and  $B$  are not independent.

Proof: Assume for the sake of a contradiction that  $A$  and  $B$  are mutually exclusive events that independently occur with nonzero probabilities. Since they are independent, we know  $P(A)P(B) = P(A \cap B)$ . Since  $P(B)$  is nonzero,  $P(A) = P(A \cap B)/P(B)$ . Since  $A$  and  $B$  are mutually exclusive,  $P(A) = P(\emptyset)/P(B) = 0/P(B) = 0$ .  $\Rightarrow \Leftarrow$ .  $A$  and  $B$  cannot be independent.  $\square$

- (5) (Conditional Independence) Sometimes, we may also use the concept of conditional independence. The definition is as follows: let  $A, B, C$  be three events with positive probabilities. Then  $A$  and  $B$  are independent given  $C$  if  $P(A \cap B|C) = P(A|C)P(B|C)$ . Consider the experiment of tossing two dice. Let  $A = \{\text{First die is 6}\}$ ,  $B = \{\text{Second die is 6}\}$ , and  $C = \{\text{Both dice are the same}\}$ .

- (a) Show that  $A$  and  $B$  are independent (unconditionally), but  $A$  and  $B$  are dependent given  $C$ .

Proof: Define  $(x, y)$  where  $x = \{1, 2, 3, 4, 5, 6\}$  equals the number rolled on the first die and  $y = \{1, 2, 3, 4, 5, 6\}$  equals the number rolled on the second die. Thus,

$$\begin{aligned} A &= \{(6, 1), (6, 2), (6, 3), (6, 4), (6, 5), (6, 6)\} \\ B &= \{(1, 6), (2, 6), (3, 6), (4, 6), (5, 6), (6, 6)\} \\ C &= \{(1, 1), (2, 2), (3, 3), (4, 4), (5, 5), (6, 6)\} \\ A \cap B &= A \cap C = B \cap C = A \cap B \cap C = \{(6, 6)\} \end{aligned}$$

If the dice are fair, each  $(x, y)$  has probability of  $1/36$ , so

$$\begin{aligned} P(A) &= P(B) = P(C) = 6/36 = 1/6 \\ P(A \cap B) &= P(A \cap C) = P(B \cap C) = P(A \cap B \cap C) = 1/36 \end{aligned}$$

Since  $P(A \cap B) = 1/36 = (1/6)(1/6) = P(A)P(B)$ ,  $A$  and  $B$  are independent.

$$\begin{aligned} P(A|C) &= P(A \cap C)/P(C) = (1/36)/(1/6) = 1/6 \\ P(B|C) &= P(B \cap C)/P(C) = (1/36)/(1/6) = 1/6 \\ P(A|C)P(B|C) &= (1/6)(1/6) = 1/36 \\ P(A \cap B|C) &= P(A \cap B \cap C)/P(C) = (1/36)/(1/6) = 1/6 \end{aligned}$$

Since  $P(A \cap B|C) \neq P(A|C)P(B|C)$ ,  $A$  and  $B$  are dependent given  $C$ .  $\square$

- (b) Consider the following experiment: let there be two urns, one with 9 black balls and 1 white balls and the other with 1 black ball and 9 white balls. First randomly (with equal probability) select one urn. Then take two draws with replacement from the selected urn. Let  $A$  and  $B$  be drawing a black ball in the first and the second draw, respectively, and let  $C$  be the event urn 1 is selected. Show that  $A$  and  $B$  are not independent, but are conditionally independent given  $C$ .

Proof: First, notice that since we are drawing with replacement, the ball drawn first does not affect the probability of which color of the ball drawn second, so  $P(A|C) = P(B|C) = 9/10$  and  $P(A \cap B|C) = P(A|C)P(B|C)$ . Thus, conditional on  $C$ ,  $A$  and  $B$  are independent. Now, let's find  $P(A)$ ,  $P(B)$ , and  $P(A \cap B)$  using the partition rule and the definition of conditional probability:

$$\begin{aligned}
 P(A) &= P((A \cap C) \cup (A \cap C^C)) \\
 &= P(A \cap C) + P(A \cap C^C) \\
 &= P(A|C)P(C) + P(A|C^C)P(C^C) \\
 &= (9/10)(1/2) + (1/10)(1/2) \\
 &= 9/20 + 1/20 \\
 &= 1/2
 \end{aligned}$$

$$\begin{aligned}
 P(B) &= P((B \cap C) \cup (B \cap C^C)) \\
 &= P(B \cap C) + P(B \cap C^C) \\
 &= P(B|C)P(C) + P(B|C^C)P(C^C) \\
 &= (9/10)(1/2) + (1/10)(1/2) \\
 &= 9/20 + 1/20 \\
 &= 1/2
 \end{aligned}$$

$$\begin{aligned}
 P(A \cap B) &= P(((A \cap B) \cap C) \cup ((A \cap B) \cap C^C)) \\
 &= P((A \cap B) \cap C) + P((A \cap B) \cap C^C) \\
 &= P(A \cap B|C)P(C) + P(A \cap B|C^C)P(C^C) \\
 &= P(A|C)P(B|C)P(C) + P(A|C^C)P(B|C^C)P(C^C) \\
 &= (9/10)(9/10)(1/2) + (1/10)(1/10)(1/2) \\
 &= 82/200
 \end{aligned}$$

Since  $P(A)P(B) \neq P(A \cap B)$ ,  $A$  and  $B$  are not independent.  $\square$

- (6) A CDF  $F_X$  is stochastically greater than a CDF  $F_Y$  if  $F_X(t) \leq F_Y(t)$  for all  $t$  and  $F_X(t) < F_Y(t)$  for some  $t$ . Prove that if  $X \sim F_X$  and  $Y \sim F_Y$ , then

$$P(X > t) \geq P(Y > t) \text{ for every } t,$$

and

$$P(X > t) > P(Y > t) \text{ for some } t,$$

that is,  $X$  tends to be bigger than  $Y$ .

Proof: If  $X \sim F_X$  and  $Y \sim F_Y$ , then, for every  $t$ ,

$$\begin{aligned} F_X(t) &= P(X \leq t) = 1 - P(X > t) \\ F_Y(t) &= P(Y \leq t) = 1 - P(Y > t) \end{aligned}$$

Furthermore, for every  $t$ , we know that

$$\begin{aligned} F_X(t) &\leq F_Y(t) \\ 1 - P(X > t) &\leq 1 - P(Y > t) \\ P(X > t) &\geq P(Y > t). \end{aligned}$$

In addition, we know for some  $t$ ,

$$\begin{aligned} F_X(t) &< F_Y(t) \\ 1 - P(X > t) &< 1 - P(Y > t) \\ P(X > t) &> P(Y > t). \end{aligned}$$

□

(7) Show that the function  $F_X(x) = \begin{cases} 0, & x < 0 \\ 1 - \exp(-x), & x \geq 0 \end{cases}$  is a CDF, and find  $f_X(x)$  and  $F_X^{-1}(y)$ .

Proof: CDFs have three unique properties: (1)  $\lim_{x \rightarrow \infty} F_X(x) = 1$ ,  $\lim_{x \rightarrow -\infty} F_X(x) = 0$ , (2)  $F_X(x)$  is non-decreasing, (3),  $F_X(x)$  is right-continuous, that is  $\lim_{x \rightarrow x_0^+} F_X(x) = F_X(x_0)$ .

For (1), since  $\lim_{x \rightarrow \infty} \exp(-x) = 0$ ,  $\lim_{x \rightarrow \infty} (1 - \exp(-x)) = 1$  and  $\lim_{x \rightarrow -\infty} (0) = 0$ .

For (2), we show that  $x_1 \leq x_2 \implies F_X(x_1) \leq F_X(x_2)$ . If  $x_1 < 0$ , then  $F_X(x_1) = 0$ . Since  $F_X(x_2) \geq 0$ , then  $F_X(x_1) = 0 \leq F_X(x_2)$ . If  $x_1 \geq 0$ , then  $F_X(x_1) = 1 - \exp(-x_1)$  and  $F_X(x_2) = 1 - \exp(-x_2)$ .  $x_1 \leq x_2 \implies 1 - \exp(-x_1) \leq 1 - \exp(-x_2) \implies F_X(x_1) \leq F_X(x_2)$ . Thus,  $F_X$  is nondecreasing.

For (3),  $F_X(x) = 0$  is continuous for  $x < 0$  and  $x > 0$  because both a constant function and  $1 - \exp(-x)$  is continuous. Finally, at  $x = 0$ ,  $\lim_{x \rightarrow 0^+} F_X(x) = \lim_{x \rightarrow 0^+} (1 - \exp(-x)) = 1 - \exp(-0) = 0$ .  $\square$

To find the PDF, take the derivative of the CDF:

$$\begin{aligned} f_X(x) &= \frac{d}{dx} F_X(x) \\ &= \begin{cases} \frac{d}{dx}(0), & x < 0 \\ \frac{d}{dx}(1 - \exp(-x)), & x \geq 0 \end{cases} \\ &= \begin{cases} (0), & x < 0 \\ 0 - \exp(-x)(-1), & x \geq 0 \end{cases} \\ &= \begin{cases} 0, & x < 0 \\ \exp(-x), & x \geq 0 \end{cases} \end{aligned}$$

Notice that  $F_X^{-1}(y)$  is defined over  $y \in [0, 1)$  because of the properties of CDFs (i.e.,  $G(x) \in [0, 1] \forall x \in \mathbb{R}$  for any CDF  $G$ ) and  $F_X(x)$  asymptotically converges to 1 as  $x \rightarrow \infty$ . So, for  $x \geq 0$ ,

$$\begin{aligned} y &= 1 - \exp(-x) \\ -\ln(1 - y) &= x \end{aligned}$$

Thus,

$$F_X^{-1}(y) = -\ln(1 - y), y \in [0, 1)$$

Therefore,  $F_X^{-1}(F_X(x)) = -\ln(1 - (1 - \exp(-x))) = x$  for  $x \geq 0$ .