

A is idempotent $\rightarrow A^2 = A$

SVD

V = Eigenvector

$L = \text{diag}(l_1, l_2, \dots, l_n)$

$$\rightarrow VLV' = (V^*L^*V') * (V^*L^*V')$$

$$(V^*L^*V') * (V^*L^*V') = V^*L^*(V'V) * L^*V'$$

$(V'V) = \text{Identity matrix}$

$$V^*L^*(V'V) * L^*V'$$

$$\Rightarrow V^*L^*L^*V'$$

$$\Rightarrow V^*L^2^*V'$$

$$V^*L^2^*V' = V^*L^*V' \text{ (remove } V \text{ and } V' \text{ by multiplying } V \text{ and } V' \text{ bc } V^*V' = \text{Identity matrix})$$

$$L^2 = L$$

$$L^2 - L = 0$$

$$\Rightarrow L(L - \text{Identitymatrix})$$

Applied Data Analysis (ADA)

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➤ A preliminary note

- **Please read carefully the slides on the course concept uploaded in RWTHmoodle!**
- The lecture is split into two parts:
 - **Part I: Linear Models** (Cramer)
Lectures from April 13 to May 19
Tutorials: April 23, May 7, 21
 - **Break:** May 24–28 (Pentecost week)
 - **Part II: Generalized Linear Models** (Kateri)
Lectures from June 1 to July 23
Tutorials: June 18, July 2, 16
- Parts I & II will be held as distance teaching course.

Just to let you know who is talking to you on linear models...

▶ Prof. Dr. Erhard Cramer

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▶ In the first part, we consider...

a class of statistical models, so called **linear models**, generated by the equation

$$Y = B\beta + \varepsilon$$

with

- $Y = (Y_1, \dots, Y_n)'$ vector of observations,
- B design matrix,
- β parameter vector
- ε (random) error term (not observable).



Part I: Linear Models

Chapter I.1

Preliminaries

Notation, Linear Algebra & Probability

➤ To be discussed...

- properties of vectors & matrices, rank, trace, singular value decomposition, etc.
- Moore-Penrose general inverse
- Image, kernel, orthogonal projectors
- Random vectors, expectations, covariance matrix
- selected probability distributions on the real line connected to the normal distribution
- non-central χ^2 - and F-distribution

Part I: Linear Models

Chapter I.1

Preliminaries

Linear Algebra

Notation & basic definitions

➤ I.1.1 Notation (vectors and matrices)

➤ \mathbb{R}^p : p -dimensional Euclidean space

➤ $\mathbb{R}^{p \times q}$: set of all $(p \times q)$ -matrices

➤ vectors are written in bold italics: $\mathbf{x} = (x_i)_{1 \leq i \leq p} = \begin{pmatrix} x_1 \\ \vdots \\ x_p \end{pmatrix}$

➤ random vectors are written in capital bold italics: $\mathbf{X} = (X_i)_{1 \leq i \leq p} = \begin{pmatrix} X_1 \\ \vdots \\ X_p \end{pmatrix}$

➤ matrices are written in capitals: $A = (a_{ij})_{1 \leq i \leq p, 1 \leq j \leq q} = \begin{pmatrix} a_{11} & \cdots & a_{1q} \\ \vdots & \cdots & \vdots \\ a_{p1} & \cdots & a_{pq} \end{pmatrix}.$

Notation & basic definitions

➤ I.1.2 Notation (special vectors and matrices)

- $A = \text{diag}(a_1, \dots, a_p)$: diagonal matrix with diagonal elements a_1, \dots, a_p
- $\mathbb{1}_p \in \mathbb{R}^p$: vector of ones, $\mathbf{0} \in \mathbb{R}^p$ zero vector
- $e_{1,p}, \dots, e_{p,p}$: standard basis of \mathbb{R}^p
- $I_p = \text{diag}(1, \dots, 1)$: p -dimensional identity matrix
- $\mathbb{1}_{p \times p} = \mathbb{1}_p \mathbb{1}_p'$: matrix of ones
- $E_p = I_p - \frac{1}{p} \mathbb{1}_{p \times p}$: ortho-projection matrix

- $\|x\| = \sqrt{\sum_{i=1}^p x_i^2}$ denotes the (Euclidean) norm of a vector $x \in \mathbb{R}^p$.
- $\text{rank}(A)$ denotes the rank of a matrix A .
- $\det(A)$ denotes the determinant of a squared matrix A .
- $\text{trace}(A)$ denotes the trace of a squared matrix $A \in \mathbb{R}^{p \times p}$, i.e., $\text{trace}(A) = \sum_{i=1}^p a_{ii}$

- The transpose of a matrix A is denoted by A' .
- The inverse of a matrix $A \in \mathbb{R}^{p \times p}$ is denoted by A^{-1} (provided it exists), i.e., $AA^{-1} = A^{-1}A = I_p$.

Notation & basic definitions

➤ I.1.3 Definition

- A matrix $A \in \mathbb{R}^{p \times p}$ is called symmetric if $A = A'$.
- A matrix $A \in \mathbb{R}^{p \times p}$ is called an orthogonal matrix if $AA' = A'A = I_p$.
- A matrix $A \in \mathbb{R}^{p \times p}$ is called positive (non-negative) definite if $A = A'$ and

$$\mathbf{x}'A\mathbf{x} > (\geq) 0 \quad \forall \mathbf{x} \in \mathbb{R}^p \setminus \{\mathbf{0}\}.$$

For short, we write $A > 0$ or $A \geq 0$, respectively.

- $\mathbb{R}_{>0}^{p \times p}$: set of all positive definite $(p \times p)$ -matrices
- $\mathbb{R}_{\geq 0}^{p \times p}$: set of all non-negative definite $(p \times p)$ -matrices

Some linear algebra

► I.1.4 Lemma

Let $A, C \in \mathbb{R}^{p \times p}$ with $\det(AC) \neq 0$ and $B \in \mathbb{R}^{k \times p}$, $1 \leq k \leq p$. Then:

- ① $(C')^{-1} = (C^{-1})'$
- ② $(AC)' = C'A'$
- ③ $(AC)^{-1} = C^{-1}A^{-1}$
- ④ $\text{rank}(B') = \text{rank}(B)$
- ⑤ $\text{rank}(BC) = \text{rank}(B)$
- ⑥ For $D \in \mathbb{R}^{p \times k}$, we have $\text{rank}(BD) = \text{rank}(DB)$.
- ⑦ $\text{rank}(BB') = \text{rank}(B'B) = \text{rank}(B)$
- ⑧ $\det(AC) = \det(CA) = \det(A) \cdot \det(C)$ for all $A, C \in \mathbb{R}^{p \times p}$
- ⑨ $\text{trace}(A + C) = \text{trace}(A) + \text{trace}(C)$
- ⑩ $\text{trace}(BD) = \text{trace}(DB)$ for all $D \in \mathbb{R}^{p \times k}$

Singular Value Decomposition (SVD)

► I.1.5 Theorem

Let $\Sigma \in \mathbb{R}_{\geq 0}^{p \times p}$. Then, the **singular value decomposition** (eigen decomposition) of Σ is given by

$$\Sigma = V \Lambda V', \quad V' = V^{\wedge(-1)}$$

where $\lambda_1 \geq \dots \geq \lambda_p \geq 0$ denote the eigenvalues and $\mathbf{v}_1, \dots, \mathbf{v}_p$ the corresponding (orthonormal) eigenvectors of Σ . Further, $\Lambda = \text{diag}(\lambda_1, \dots, \lambda_p)$ and $V = [\mathbf{v}_1 \mid \dots \mid \mathbf{v}_p]$ with $V'V = VV' = I_p$.

Furthermore, with the definitions $\Lambda^{1/2} = \text{diag}(\sqrt{\lambda_1}, \dots, \sqrt{\lambda_p})$ and $\Sigma^{1/2} = V \Lambda^{1/2} V'$, we have

► $\Sigma = \Sigma^{1/2} \Sigma^{1/2}$ and $(\Sigma^{1/2})' = \Sigma^{1/2}$.

► $\Sigma^{1/2}$ is non-negative definite.

👉 $\Sigma^{1/2}$ is called the **root of Σ** .

► Notice that, for a regular matrix Σ , $(\Sigma^{1/2})^{-1} = \Sigma^{-1/2}$ where $\Sigma^{-1/2} = V \Lambda^{-1/2} V'$ and $\Lambda^{-1/2} = \text{diag}(\sqrt{\lambda_1^{-1}}, \dots, \sqrt{\lambda_p^{-1}})$.

Moore-Penrose (general) inverse of a matrix

► I.1.6 Theorem

The Moore-Penrose (general) inverse of a matrix A is denoted by A^+ , i.e., A^+ is the unique matrix satisfying the four equations:

$$AA^+A = A, \quad A^+AA^+ = A^+, \quad (AA^+)' = AA^+, \quad (A^+A)' = A^+A.$$

It has the following properties:

- $(A^+)^+ = A$
- $(A^+)' = (A')^+$
- $A^+ = A^+(A^+)'A'$
- $A = AA'(A^+)'$
- $A' = A'AA^+$
- $A' = A^+AA'$
- $(AA')^+ = (A')^+A^+$
- $A^+ = (A'A)^+A'$
- $A^+ = A'(AA')^+$
- If $A \in \mathbb{R}^{p \times q}$ exhibits the SVD $A = U\Lambda V'$, then A^+ has the SVD $A^+ = V\Lambda^+U'$ where Λ^+ is the Moore-Penrose inverse of the matrix Λ .
- If $A \in \mathbb{R}^{p \times p}$ is a regular matrix then $A^+ = A^{-1}$.

Image, Kernel, Orthogonal Projectors

► I.1.7 Definition

- For a matrix $A \in \mathbb{R}^{p \times q}$, let
 - $\text{Ker}(A) = \{\mathbf{x} \in \mathbb{R}^q \mid A\mathbf{x} = \mathbf{0}\}$ be the kernel (null space) of A .
 - $\text{Im}(A) = \{A\mathbf{x} \mid \mathbf{x} \in \mathbb{R}^q\}$ be the image of A .
- For a linear subspace $\mathcal{A} \subseteq \mathbb{R}^p$, $\mathcal{A}^\perp = \{\mathbf{y} \in \mathbb{R}^p \mid \mathbf{x}'\mathbf{y} = 0 \text{ for all } \mathbf{x} \in \mathcal{A}\}$ denotes the corresponding orthogonal space.
- Let $\mathcal{A}, \mathcal{B} \subseteq \mathbb{R}^p$ be linear subspaces with $\mathcal{A} \cap \mathcal{B} = \{\mathbf{0}\}$. Then, $\mathcal{A} \oplus \mathcal{B} = \{\mathbf{x} + \mathbf{y} \mid \mathbf{x} \in \mathcal{A}, \mathbf{y} \in \mathcal{B}\}$ is called the direct sum of \mathcal{A}, \mathcal{B} .

Notice that $\text{Ker}(A) \subseteq \mathbb{R}^q$ and $\text{Im}(A) \subseteq \mathbb{R}^p$ are linear subspaces.

► I.1.8 Definition

A matrix $Q \in \mathbb{R}^{p \times p}$ is called

- idempotent if $Q^2 = Q$
- orthogonal projector on a linear subspace $\mathcal{A} \subseteq \mathbb{R}^p$ if
 - for any $\mathbf{x} \in \mathcal{A}$: $Q\mathbf{x} = \mathbf{x}$
 - for any $\mathbf{y} \in \mathcal{A}^\perp$: $Q\mathbf{y} = \mathbf{0}$

Properties of Orthogonal Projectors and Moore-Penrose inverse

► I.1.9 Lemma

- Orthogonal projectors on a linear subspace $\{0\} \neq \mathcal{A} \subseteq \mathbb{R}^p$ are unique.
- $Q \in \mathbb{R}^{p \times p}$ is an orthogonal projector (on $\text{Im}(Q)$) iff $Q^2 = Q$ and $Q' = Q$.

► I.1.10 Theorem

Let $A \in \mathbb{R}^{p \times q}$ with Moore-Penrose inverse A^+ and define $P_1 = I_q - A^+A$, $P_2 = I_p - AA^+$. Then:

- P_1 and P_2 are orthogonal projectors, respectively, that is, $P_i^2 = P_i$, $P_i' = P_i$, $i = 1, 2$.
- $Q = AA^+ = A(A'A)^+A'$ is the (unique) orthogonal projector on $\text{Im}(A)$.
- $A^+A = A'(AA')^+A$ is the (unique) orthogonal projector on $\text{Im}(A')$.
- $\text{Ker}(A) = \text{Im}(P_1)$, $\text{Im}(A) = \text{Ker}(P_2)$.
- $\text{Ker}(A^+) = \text{Im}(P_2)$, $\text{Im}(A^+) = \text{Ker}(P_1)$
- $\text{Im}(A) = \text{Ker}(A^+)^\perp$, $\text{Im}(A^+) = \text{Ker}(A)^\perp$,
- $\text{Ker}(A) \oplus \text{Im}(A^+) = \mathbb{R}^q$, $\text{Ker}(A^+) \oplus \text{Im}(A) = \mathbb{R}^p$

Part I: Linear Models

Chapter I.1

Preliminaries

Probability

Expectations of random vectors and random matrices

► I.1.11 Definition (expectation of random vectors and random matrices)

- ① The **expectation of a random vector** $\mathbf{X} = (X_1, \dots, X_p)'$ is defined by the vector of means, that is,

$$E\mathbf{X} = \begin{pmatrix} EX_1 \\ \vdots \\ EX_p \end{pmatrix};$$

subsequently, we use the notation $\boldsymbol{\mu} = E\mathbf{X}$;

- ② The **expectation of a random matrix** $\mathcal{X} = (X_{ij})_{1 \leq i \leq p, 1 \leq j \leq q}$ is defined by the matrix of means, that is,

$$E\mathcal{X} = \begin{pmatrix} EX_{11} & \cdots & EX_{1q} \\ \vdots & \ddots & \vdots \\ EX_{p1} & \cdots & EX_{pq} \end{pmatrix}.$$

In the following, all expectations are supposed to exist.

► **I.1.12 Lemma**

- ① Let $\mathbf{X} = (X_1, \dots, X_p)'$ be a p -dimensional random vector and $A \in \mathbb{R}^{k \times p}$, $\mathbf{b} \in \mathbb{R}^k$. Then:

$$E(A\mathbf{X} + \mathbf{b}) = AE(\mathbf{X}) + \mathbf{b}.$$

- ② Let $\mathbf{Z}_1, \dots, \mathbf{Z}_n$ be p -dimensional random vectors and $A_1, \dots, A_n \in \mathbb{R}^{k \times p}$. Then:

$$E\left(\sum_{j=1}^n A_j \mathbf{Z}_j\right) = \sum_{j=1}^n A_j E(\mathbf{Z}_j) \in \mathbb{R}^k.$$

► I.1.13 Definition (variance-covariance matrix)

Let $\mathbf{X} = (X_1, \dots, X_p)'$, $\mathbf{Y} = (Y_1, \dots, Y_q)'$ be random vectors. Then, the **covariance matrix** of \mathbf{X} and \mathbf{Y} is defined by

$$\text{Cov}(\mathbf{X}, \mathbf{Y}) = \begin{pmatrix} \text{Cov}(X_1, Y_1) & \cdots & \text{Cov}(X_1, Y_q) \\ \vdots & \ddots & \vdots \\ \text{Cov}(X_p, Y_1) & \cdots & \text{Cov}(X_p, Y_q) \end{pmatrix}.$$

The **variance-covariance matrix** of \mathbf{X} is defined by $\Sigma = \text{Cov}(\mathbf{X}) = \text{Cov}(\mathbf{X}, \mathbf{X})$.

► I.1.14 Remark

Defining the random matrix $\mathcal{C}_{\mathbf{X}, \mathbf{Y}} = (\mathbf{X} - E(\mathbf{X}))(\mathbf{Y} - E(\mathbf{Y}))'$, we get

- ① $\text{Cov}(\mathbf{X}, \mathbf{Y}) = E\mathcal{C}_{\mathbf{X}, \mathbf{Y}}$
- ② $\text{Cov}(\mathbf{X}) = \text{Cov}(\mathbf{X}, \mathbf{X}) = E\mathcal{C}_{\mathbf{X}, \mathbf{X}}$
- ③ $\text{Cov}(\mathbf{X}, \mathbf{Y}) = E(\mathbf{X}\mathbf{Y}') - E\mathbf{X} \cdot E\mathbf{Y}'$
- ④ $\text{Cov}(\mathbf{X}) = E(\mathbf{X}\mathbf{X}') - E\mathbf{X} \cdot E\mathbf{X}'$

Covariance matrices are always non-negative definite, that is, $\text{Cov}(\mathbf{X}) \geq 0$.

► I.1.15 Notation (block matrix)

A matrix $A \in \mathbb{R}^{(p+q) \times (k+r)}$ can be written as a **block matrix**

$$A = \begin{bmatrix} A_{11} & A_{12} \\ A_{21} & A_{22} \end{bmatrix} \quad \text{with } A_{11} \in \mathbb{R}^{p \times k}, A_{12} \in \mathbb{R}^{p \times r}, A_{21} \in \mathbb{R}^{q \times k}, A_{22} \in \mathbb{R}^{q \times r}.$$

► I.1.16 Lemma

With the notation from Definition I.1.13, we get for $A \in \mathbb{R}^{k \times p}$, $B \in \mathbb{R}^{r \times q}$, $\mathbf{b} \in \mathbb{R}^k$, $\mathbf{c} \in \mathbb{R}^r$:

① $\text{Cov}(A\mathbf{X} + \mathbf{b}, B\mathbf{Y} + \mathbf{c}) = A \text{Cov}(\mathbf{X}, \mathbf{Y}) B'$,

② $\text{Cov}(A\mathbf{X} + \mathbf{b}) = A \text{Cov}(\mathbf{X}) A'$,

③ $\text{Cov} \begin{pmatrix} \mathbf{X} \\ \mathbf{Y} \end{pmatrix} = \begin{bmatrix} \text{Cov}(\mathbf{X}) & \text{Cov}(\mathbf{X}, \mathbf{Y}) \\ \text{Cov}(\mathbf{Y}, \mathbf{X}) & \text{Cov}(\mathbf{Y}) \end{bmatrix}$,

④ $\text{Cov}(\mathbf{X}, \mathbf{Y}) = \text{Cov}(\mathbf{Y}, \mathbf{X})'$.

Using Lemma I.1.16, we can write with $\Sigma_{\mathbf{X}\mathbf{Y}} = \text{Cov}(\mathbf{X}, \mathbf{Y})$:

$$\Sigma_{\begin{pmatrix} \mathbf{X} \\ \mathbf{Y} \end{pmatrix}} = \begin{bmatrix} \Sigma_{\mathbf{X}\mathbf{X}} & \Sigma_{\mathbf{X}\mathbf{Y}} \\ \Sigma_{\mathbf{Y}\mathbf{X}} & \Sigma_{\mathbf{Y}\mathbf{Y}} \end{bmatrix} \stackrel{(*)}{=} \begin{bmatrix} \Sigma_{\mathbf{X}\mathbf{X}} & \Sigma_{\mathbf{X}\mathbf{Y}} \\ \Sigma'_{\mathbf{X}\mathbf{Y}} & \Sigma_{\mathbf{Y}\mathbf{Y}} \end{bmatrix}.$$

Probability distributions on \mathbb{R}

► I.1.17 Remark (density functions of distributions on \mathbb{R})

- Normal distribution $N(\mu, \sigma^2)$:

$$f(x) = \varphi_{\mu, \sigma^2}(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{(x-\mu)^2}{2\sigma^2}\right\}, \quad x \in \mathbb{R}$$

- χ^2 -distribution $\chi^2(p)$ with $p \in \mathbb{N}$ degrees of freedom:

$$f(x) = \frac{1}{2^{p/2}\Gamma(p/2)} x^{p/2-1} e^{-x/2}, \quad x > 0$$

- t-distribution $t(p)$ with $p \in \mathbb{N}$ degrees of freedom:

$$f(x) = \frac{\Gamma(\frac{n+1}{2})}{\sqrt{p\pi}\Gamma(\frac{p}{2})} \left(1 + \frac{x^2}{p}\right)^{-(p+1)/2}, \quad x \in \mathbb{R}$$

- F-distribution $F(p, q)$ with $p \in \mathbb{N}$ numerator and $q \in \mathbb{N}$ denominator degrees of freedom:

$$f(x) = \frac{\Gamma(\frac{p+q}{2})}{\Gamma(\frac{p}{2})\Gamma(\frac{q}{2})} \left(\frac{p}{q}\right)^{n/2} \frac{x^{p/2-1}}{(1 + \frac{p}{q}x)^{\frac{p+q}{2}}}, \quad x > 0$$

Connections of probability distributions

> I.1.18 Notation

The notation $X_1, \dots, X_k \stackrel{\text{iid}}{\sim} P$ means that the random variables X_1, \dots, X_k are independent and identically distributed (iid) with $X_1 \sim P$.

The same notation is used for samples of random vectors.

> I.1.19 Proposition

- ① Let $X \sim N(0, 1)$ and $\mu \in \mathbb{R}, \sigma > 0$. Then, $\mu + \sigma X \sim N(\mu, \sigma^2)$.
- ② Let $X_1, \dots, X_p \stackrel{\text{iid}}{\sim} N(0, 1)$. Then, $\sum_{i=1}^p X_i^2 \sim \chi^2(p)$.
- ③ Let $X \sim \chi^2(p)$ and $Z \sim \chi^2(q)$ be independent random variables. Then, $X + Z \sim \chi^2(p + q)$.
- ④ Let $X \sim N(0, 1)$ and $Z \sim \chi^2(p)$ be independent random variables. Then, $\frac{X}{\sqrt{\frac{1}{p}Z}} \sim t(p)$.
- ⑤ Let $X \sim \chi^2(p)$ and $Z \sim \chi^2(q)$ be independent random variables. Then, $\frac{\frac{1}{p}X}{\frac{1}{q}Z} \sim F(p, q)$.

Non-central χ^2 - and F-distribution

► I.1.20 Remark

- Given independent random variables X_1, \dots, X_p with $X_i \sim N(\mu_i, 1)$, $\mu_i \in \mathbb{R}$, $1 \leq i \leq p$, the distribution of

$$\sum_{i=1}^p X_i^2$$

is called non-central χ^2 -distribution $\chi^2(p, \delta)$ with $p \in \mathbb{N}$ degrees of freedom and non-centrality parameter $\delta = \frac{1}{2} \sum_{i=1}^p \mu_i^2 \geq 0$.

Clearly, $\chi^2(p) = \chi^2(p, 0)$.

- Let $X \sim \chi^2(p, \delta)$ and $Z \sim \chi^2(q)$ be independent random variables. Then,

$$\frac{\frac{1}{p}X}{\frac{1}{q}Z} \sim F(p, q, \delta),$$

that is, the ratio has a non-central F-distribution $F(p, q, \delta)$ with $p \in \mathbb{N}$ numerator and $q \in \mathbb{N}$ denominator degrees of freedom and non-centrality parameter $\delta \geq 0$.

Clearly, $F(p, q) = F(p, q, 0)$.