

The impact of bike-share on real-estate transaction prices in New York City*

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Abstract

Bike-share programs have been introduced in more than two thousand cities around the world, but little is known about their impact on cities. Given their potential to act as a local amenity (providing new transport options to commuters), agents may increase their valuation of property units nearby bike-share stations. In this paper, I test this hypothesis in New York City using the universe of real-estate transactions. I find that transactions within 150 metres of a bike-share station are sold at prices up to 6.1% higher than properties between 150 and 500 metres of the same station, or \$185,755 higher for an average transaction. This result suggests that bike-share is valued by urban dwellers and that it initiated important value creation.

Keywords: Real estate, Urban transportation, Cycling

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1 Introduction

In the past thirty years, cities around the world have invested significantly in policies to encourage cycling. Praised for its potential to reduce traffic congestion, relieve pressure on public transport and curtail air pollution, the implementation of cycling policies has been widespread. On the one hand, cities have improved and expanded their cycling infrastructure (cycling paths, bike parking, etc), making cycling safer and more convenient. On the other hand, they have also sought to make cycling more accessible by providing public bike-share schemes to their inhabitants. More than two thousand such programs now exist around the world, providing an estimated 66 million trips in North America alone in 2021.

The advent of an affordable and practical cycling option thanks to bike-share programs has the potential to significantly change the transport landscape of cities. From previous research, we know that these changes in mobility options may have important impacts on commuting patterns, neighbourhood composition, and real-estate prices [citations]. Despite their popularity in the past two decades, there is a notable scarcity of research on the impact that bike-share programs may have on cities.

In this paper, I test whether the initial roll-out of North America’s largest bike-share scheme in New York City in May 2013 had an impact on real-estate transaction prices. I geocode the universe of real-estate transactions from the New York City Department of Finance and match each of them to a bike-share station opened in the first wave (i.e., between May 2013 and June 2015). I assign transactions within 150 metres of a bike-share station to the treatment group (or ring), and transactions between 150 and 500 metres to the control group. I use a two-ring difference-in-differences approach, comparing transactions in the treatment ring with those in the control ring, before and after the opening of a bike-share station, and including station fixed effects. I find that transaction prices in the treatment ring increased by 6.1% after bike-share was implemented compared to transactions in the control ring, which represents an increase of \$185,755 for an average transaction price.

This study contributes to a large body of research documenting the effects of transportation on real-estate prices, which showed (in the majority of cases) a positive relationship between access to infrastructure and prices. Important contributions for urban rail include Dewees (1976), Baum-Snow and Kahn (2005), Hess and Almeida (2007), Ahlfeldt et al. (2015), Heblich et al. (2020), Zhou, Chen, et al. (2021), and Gupta et al. (2022), and estimate price premiums between 3% and 10% for properties around rail stations. The evidence on high-capacity bus lines is similar: from a 2% to up to a 8% premium for properties around bus rapid transit (BRT) systems (see Munoz-Raskin (2010) and Zhang and Yen (2020)). Highways have also been shown to have a positive effect on nearby residential property prices (see for example Levkovich et al. (2016) and Cohen and Schaffner (2019)).

Recently, the literature started investigating the impact of cycling and cycling in-

frastructure on real-estate prices. A first set of studies (El-Geneidy et al., 2016; Li and Joh, 2017; Qiao et al., 2021) relies on cross-sectional analyses and do not employ causal inference methods. On the other hand, Pelechrinis et al. (2017), Chu et al. (2021), Zhou, Li, et al. (2022), and Shr et al. (2022) employ quasi-experimental methods to recover the causal impact of bike-share on real-estate prices. Pelechrinis et al. (2017) uses aggregated price data at the ZIP code level for the city of Pittsburg, which arguably does not control optimally for the characteristics and location of properties. Looking at free-floating bike-share systems in China, Chu et al. (2021) and Zhou, Li, et al. (2022) find that they decrease the price premium of subway stations. Finally, Shr et al. (2022) investigate the impact of a docked bike-share system on rents in Taiwan’s second largest city, focusing on six months prior and after the deployment. They find that rents increase by 2% for properties with 150 metres of bike-share stations relative to those between 150 and 500 metres. These results suggest that bike-share programs have a positive impact on real-estate prices for units in their vicinity, but say little about their potential long-term effects.

My paper contributes to this literature by using property-level transaction prices to investigate the largest bike-share program in North America over a four year period around the initial launch, which is important for a couple of reasons. First, it remains unclear if a bike-share systems in the largest and densest city in the United States has effects similar to those reported in the previous literature for Asian urban areas. Second, transport habits may be sticky and individuals may take time to adjust to new transport options: as such, one might expect that commuting changes brought by cycling policies take time to materialise into real-estate transaction prices, motivating a long-term analysis. Finally, rental and sale markets might differ in how they respond to changes in cycling policies.

These results are important in several regards. First, they show that cycling policies behave similarly to other transportation policies: they have an impact on real estate markets, even in a city where only about 1% of workers commute to work by bicycle. Second, it highlights the importance of taking into account the distributional impacts of transport policies, as they may change neighbourhood attractiveness and eventually price out some segments of the population. Third, the increase in property value due to bike-share documented here may be partially captured by the city through property taxes, which could go towards financing these investments, and improve policymakers’ and voters’ support for them (Gupta et al., 2022).

2 Data

This section describes the data used in the paper. I begin by describing the outcome data, which originates from the NYC transaction records, the primary operations performed on the data, and the sample creation process. I then transition to describing the bike-share

data, and how I define control and treatment units.

2.1 Real-estate transaction records

To assess the impact of bike share on real-estate prices, I utilise the universe of transaction records collected by the New York City Department of Finance (NYCDF). These records report key variables such as sale price, surface area, tax lot identifiers, and building classes for every real estate transaction in NYC since 2003. I have gathered transaction records spanning from January 2011 to April 2015, and executed a series of cleaning and transformation procedures to ensure the data is ready for analysis.

First, I geocode each transaction using tax lot identifiers. The NYCDF maintains its own property identifier, uniquely locating every piece of property in the city. Using their Digital Tax Map, which associates each property identifier with a geographical polygon, I determine the location of a transaction as the centroid of the polygon.¹

As mentioned above, I concentrate on the first wave of bike-share implementation in NYC (i.e., stations opened between May 2013 and April 2014) to keep the analysis manageable. The estimation strategy is based on a two-ring approach (see the [identification strategy] section), using transactions up to 500 metres away from first-wave bike-share stations, and I retain transactions within that range (significantly reducing computational complexity). Approximately 80 thousand transactions fall within that spatial-temporal range — I outline the precise construction of the treatment variable in subsection 2.3.

Next, I retain only sales with non-zero prices and non-missing surface areas, deflate the sale price to December 2015 levels, and compute the surface area per unit.² I identify price outliers, first using the definition in (Gupta et al., 2022) (sale prices greater than 400 thousand dollars and less than ten million dollars), but also price-per-square-foot outliers (greater than \$50 and less than 20 thousand dollars).³ Finally, I extract building attributes (residential/commercial, elevator, condo, etc: see subsection 2.4 for the complete list of attributes) using the building class category reported for each transaction. I also compute distances to main amenities for each transaction: distance to the nearest subway entrance and bus stop,⁴ and distance to the nearest park.⁵

¹Most polygons are identified using a borough + tax block + tax lot identifier. Condos and other communal property types are uniquely identified at the borough + tax block level only, and their location is therefore the centre of the tax block.

²Surface area is given for the whole building, not the transacted unit: I take the surface area per building unit in order to correctly measure the size of a transacted unit.

³The later outlier definition is based on anecdotal evidence of top and bottom prices in NYC (see for example <https://therealdeal.com/new-york/2018/03/21/these-are-nycs-most-expensive-homes-by-price-per-square-foot/> and <https://www.nyrentownsell.com/blog/best-price-per-sq-foot-in-ny-to-rent-and-buy/>, both accessed 2023-07-13). The range is wide by design, as its goal is to remove to most unlikely prices per square foot.

⁴Historical data on subway entrances and bus stops locations is collected by Baruch College at the City University of New York, and freely available at <https://www.baruch.cuny.edu/confluence/display/geoportal/NYC+Mass+Transit+Spatial+Layers+Archive> (accessed 2023-07-13).

⁵The current location of parks is provided by the New York City Parks at <https://nycopendata>.

The analysis presented later will concentrate on residential units excluding price and price-per-square-foot outliers, but results including outliers and commercial units are reported in the appendix.

2.2 Bike-share data

This paper estimates the impact of the first wave of bike-share stations on real-estate prices. I use the universe of bike-share trips⁶ to identify the opening (and occasionally closing) date of each station. The first wave is defined by the first spatial extent of the system, i.e. the initial area of the city that the bike-share system covered. The bike-share system in NYC was launched in May 2013 and the subsequent spatial expansion took place in July 2015. During that period, stations opened in three different months: May and June 2013, and March 2014. Upon visual inspection, it appears that some stations were closed and others opened within the same calendar month, and very near to each other. Since the estimation strategy relies on bike-share station fixed effects (see the section 3), it is critical to identify the correct set of stations, so I match those that opened and closed within a month and within 50 metres of each other as the same station.

2.3 Treatment construction

As detailed later in the paper, the estimation strategy compares real-estate transactions close to a bike-share station (within 150 metres) to those further away (between 150 and 500 metres from the station), before and after the opening of the station. Each transaction thus has to be matched with one (or more, depending on the case) bike-share station. In this subsection, I detail the steps and decisions made in matching transactions to stations.

I start by computing, for each transaction, all the bike-share stations within 500 metres. Each transaction may be matched to multiple stations: indeed, it is not uncommon for a transaction to fall within multiple 500-metre rings around bike-share stations, with a median of ten stations matched and a maximum of 17. At this stage, every row in the dataset is a transaction-station pair, with as many rows per transaction as it matches bike-share stations, and includes a measure of distance to the station (between zero and 500 metres by construction). However, not all matches are valid for estimation: indeed, a transaction cannot be treated by one station and act as a control for another. I therefore select station matches according to the following algorithm:

Case 1 The transaction matches only one station: keep that match.

socrata.com/Recreation/Parks-Properties/enfh-gkve (accessed 2023-07-13). I was not able to find historical locations of parks — the assumption is that the location of new parks is not correlated with the deployment of bike-share stations, which is plausible.

⁶Available on the bike-share provider’s website: <https://citibikenyc.com/system-data> (accessed 2023-07-13).

Case 2.1 The transaction matches multiple stations, and all are further than 150 metres (i.e., the transaction is always a control): keep all matches. This allows the transaction to act as a control for multiple bike-share stations.

Case 2.2 The transaction matches multiple stations, and all are within 150 metres (i.e., the transaction is always treated): keep the earliest and closest matched station (in that order). If the transaction is impacted by treatment, the first station within 150 metres likely started affecting its value first, and it is probable (if the stations within 150 metres opened at the same time) that the closest one has the most impact.

Case 2.3 The transaction matches multiple stations, some of them within 150 metres, some of them between 150 and 500 metres: keep the station (1) within 150 metres, (2) opened earliest, and (3) closest (in that order). Once treated, a station should not be considered a control (it would violate SUTVA), and the station matches between 150 and 500 metres are discarded. If there are multiple stations within 150 metres, the same criteria as in Case 2.2 are used.⁷

By allowing for multiple matches as described above (and after cleaning transactions as outlined earlier in subsection 2.1), I end up with about 20 thousand transaction-station pairs [check]. Transactions which are within 150 metres of their matched stations are coded as within the treatment ring (those who are between 150 and 500 metres are coded as belonging to the control ring), and transactions taking place after the opening of their matched station are coded as post-period (those before as pre-period). Figure 1 illustrates how treatment and control status is attributed, and how transactions may act as controls for multiple bike-share stations. Figure 2 provides an overview of treated and control areas and transactions. As detailed below, the treatment effect will be identified by the interaction between the treatment ring and post-period dummies.

⁷Note that cases 2.2 and 2.3 imply that I do not exploit the potential cumulative effect of multiple bike-share stations in the vicinity. While potentially important, accounting for repeated or cumulative treatment is not straightforward in practice.



Figure 1: Treatment construction illustration. *Notes:* The green symbols represent treated transactions (i.e., within the 150-metre, yellow ring), while the red symbols are controls (500-metre, grey ring). Upward triangles are transactions matched to bike-share station 54, while downward triangles are transactions matched to station 21. As described in the algorithm, some stations are used as controls by both stations: those are represented by a diamond shape.

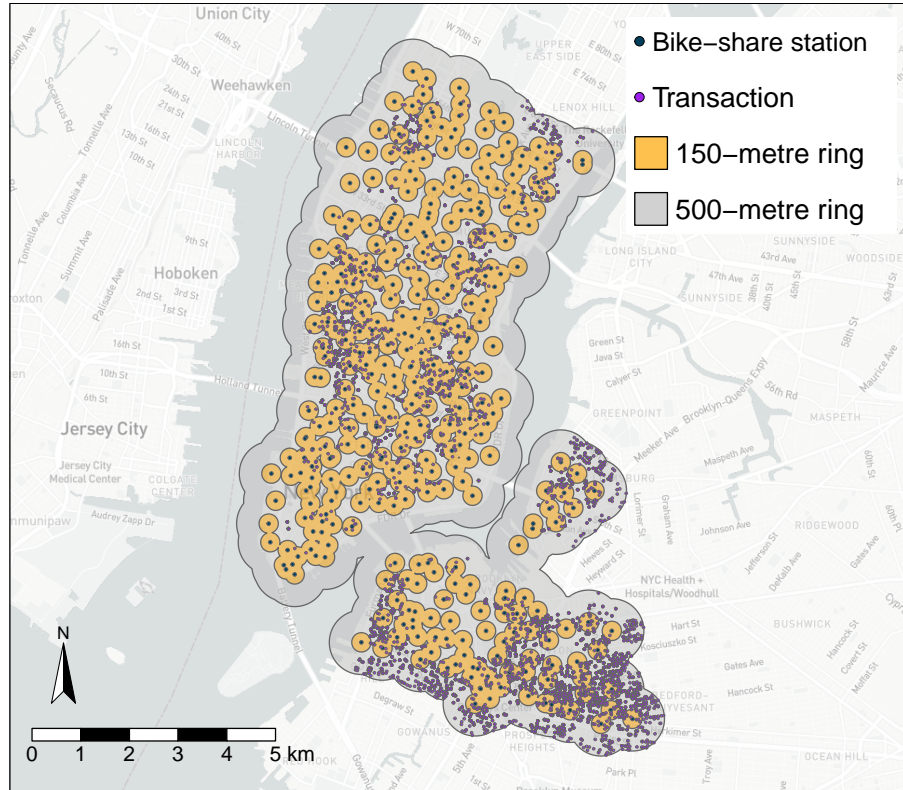


Figure 2: Overview of the study area, treatment and control areas, and location of transactions. *Notes:* Transactions are represented by purple points. The yellow area represents the treatment rings (i.e., within 150 metres of a bike-share station), while the grey area represents the controls rings (i.e., between 150 and 500 metres from a bike-share station). Note that for clarity, the overlapping areas were merged into a single area: in practice, each stations has its own individual treatment and control ring (see figure B.1 for a detailed map).

2.4 Descriptive statistics

Here I report summary statistics for the variables used in the estimation, as well as balance tables by treated vs control rings before the opening of bike-share.

Table 1: Summary statistics, numeric variables

	Mean	SD	Min	Median	Max	Miss
Sale price (2015 \$)	1,183,105	1,455,369	4e+05	673,843.3	9,983,196	0
Log sale price (2015 \$)	13.64	0.7	12.9	13.42	16.12	0
Sale price per sqft (2015 \$)	1,171.24	1,948.77	51.11	563.3	19,905.42	0
Residential units (count)	3.95	17.11	0	2	1,681	0
Commercial units (count)	0.13	0.53	0	0	15	0
Total units (count)	4.08	17.23	0	2	1,684	0
Built surface (sqft)	4,073.58	14,206.07	0	2,400	1,231,250	0
Land surface (sqft)	3,258.42	4,123.63	0	2,500	382,704	0
Final surface (sqft)	4,094.56	14,207.76	300	2,400	1,231,250	0
Surface per unit (sqft)	1,348.77	760.42	132	1,176	31,494	0
Building age	76.32	31.25	0	84	217	8
Year built	1,936.65	31.23	1,798	1,930	2,015	0
Distance to bus stop (m)	2,348.91	3,816.24	6.72	299.55	21,983.22	0
Distance to subway entrance (m)	2,213.29	3,738.27	5.47	626.58	22,533.2	0
Distance to bike-share station (m)	308.82	129.51	2.09	327.57	499.93	65795
Distance to park (m)	459.12	309.77	0	399.12	2,308.86	0
Sale quarter	10.3	5.08	1	11	18	0

Table 2: Balance table treated vs control ring, numeric variables, pre-treatment period

	Control ring 0 (N=4014)		Treated ring 1 (N=679)		Diff. in Means	p-value
	Mean	Std. Dev.	Mean	Std. Dev.		
Sale price (2015 \$)	2,992,981.78	2,349,601.20	3,319,213.45	2,414,041.15	326,231.66***	0.00
Log sale price (2015 \$)	14.57	0.87	14.70	0.85	0.13***	0.00
Sale price per sqft (2015 \$)	3,503.77	3,632.46	3,778.58	3,972.99	274.81*	0.09
Residential units (count)	10.98	63.92	8.13	23.91	-2.85**	0.04
Commercial units (count)	0.47	0.89	0.54	1.05	0.07*	0.08
Total units (count)	11.45	64.13	8.68	24.54	-2.78**	0.04
Built surface (sqft)	9,283.55	48,620.97	7,600.34	24,341.75	-1,683.21	0.16
Land surface (sqft)	2,809.18	10,872.52	2,260.12	2,132.94	-549.06***	0.00
Final surface (sqft)	9,360.91	48,620.97	7,607.12	24,340.01	-1,753.78	0.15
Surface per unit (sqft)	1,220.12	888.96	1,270.99	917.76	50.87	0.18
Building age	98.81	25.19	101.41	23.73	2.60***	0.01
Year built	1,913.01	25.12	1,910.42	23.68	-2.59***	0.01
Distance to bus stop (m)	103.89	61.26	93.18	49.87	-10.70***	0.00
Distance to subway entrance (m)	291.07	159.62	295.86	201.74	4.79	0.56
Distance to bike-share station (m)	346.00	100.78	98.95	34.37	-247.04***	0.00
Distance to park (m)	334.56	213.42	309.88	205.70	-24.67***	0.00
Sale quarter	5.78	2.73	5.77	2.70	-0.01	0.92

Notes: Significance codes: *. 0.1, **. 0.05, ***: 0.01.

Table 3: Balance table treated vs control ring, categorical variables, whole sample period

		Control ring 0 (N=7503)		Treated ring 1 (N=1386)		Total	
		N	Pct.	N	Pct.	N	Pct.
Post-period	0	4014	53.5	679	49.0	4693	6.3
	1	3489	46.5	707	51.0	4196	5.6
Treated (treatment ring \times post)	0	7503	100.0	679	49.0	8182	11.0
	1	0	0.0	707	51.0	707	0.9
Elevator	0	7312	97.5	1329	95.9	73696	98.7
	1	191	2.5	57	4.1	988	1.3
Walkup	0	5000	66.6	886	63.9	64588	86.5
	1	2503	33.4	500	36.1	10096	13.5
Condo	0	7453	99.3	1380	99.6	74615	99.9
	1	50	0.7	6	0.4	69	0.1
Coop	0	7383	98.4	1370	98.8	74503	99.8
	1	120	1.6	16	1.2	181	0.2
Rental	0	4183	55.8	682	49.2	61580	82.5
	1	3320	44.2	704	50.8	13104	17.6

Table 4: Balance table treated vs control ring, residential building class categories, whole sample period

Building class category	Control ring 0 (N=7503)		Treated ring 1 (N=1386)		Total	
	N	Pct.	N	Pct.	N	Pct.
01 One Family Dwellings	833	11.1	164	11.8	25530	34.2
02 Two Family Dwellings	2018	26.9	320	23.1	27599	37.0
03 Three Family Dwellings	1203	16.0	180	13.0	8259	11.1
07 Rentals - Walkup Apartments	2480	33.1	495	35.7	10042	13.4
08 Rentals - Elevator Apartments	125	1.7	47	3.4	896	1.2
09 Coops - Walkup Apartments	23	0.3	5	0.4	54	0.1
10 Coops - Elevator Apartments	57	0.8	8	0.6	81	0.1
11a Condo-Rentals	1	0.0	1	0.1	12	0.0
13 Condos - Elevator Apartments	9	0.1	2	0.1	11	0.0
14 Rentals - 4-10 Unit	714	9.5	161	11.6	2154	2.9
17 Condo Coops	40	0.5	3	0.2	46	0.1

2.5 Descriptive evidence

Prior to presenting the empirical strategy and the results of the statistical analysis, I present in this section descriptive evidence that may point towards bike-share having an impact on property prices. In figure 3, I plot the results of a local polynomial regression of sale prices on periods to treatment (in months), reminiscent of an event study. Prior to treatment, treated and control transaction show relatively similar price trends, while after treatment the treated group displays a relatively higher level. While not a statistical test, figure 3 provides motivation for the analysis to follow.

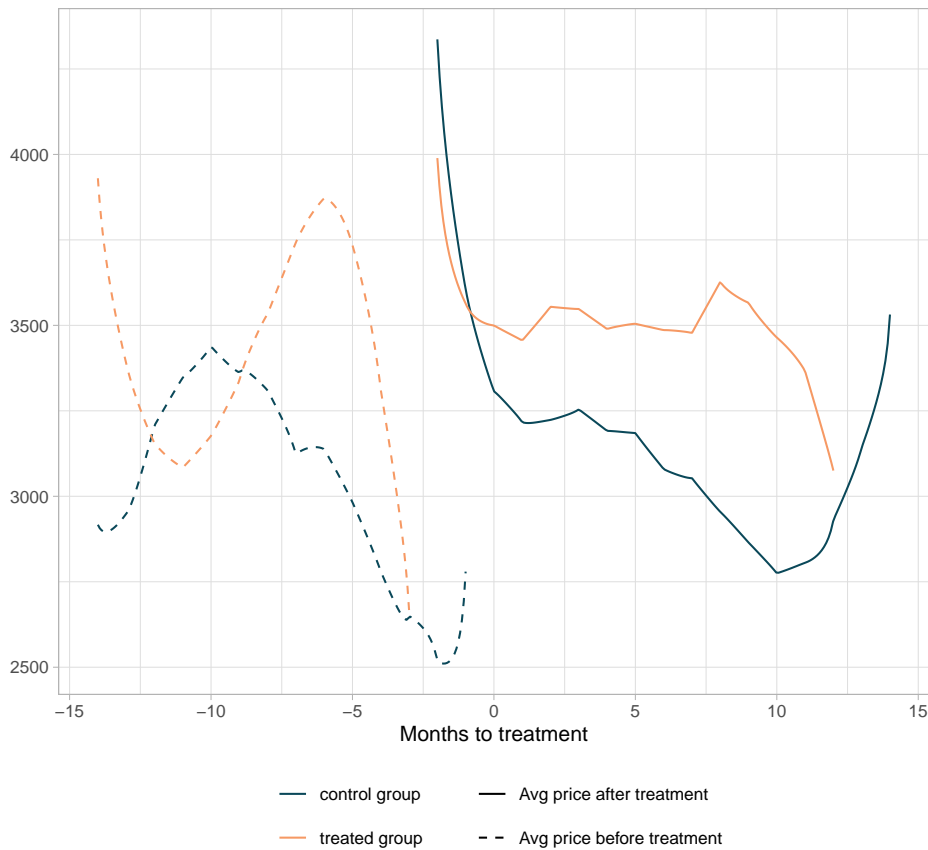


Figure 3: Local polynomial regression of sales prices on periods to treatment. *Notes:* Periods to treatment are given in months. The dashed lines are pre-treatment, and solid lines post-treatment results of a local polynomial regression of sales prices on periods to treatment. Yellow lines identify the treated transactions (i.e., within 150 metres of a bike-share station), blue ones the control group (i.e., between 150 and 500 metres of a bike-share station). The local polynomial regressions use an Epanechnikov kernel and three-month bandwidth.

3 Empirical strategy

Identifying the impact of bike-share on real-estate transaction prices is lined with several empirical challenges. First, transaction prices are determined by a multitude of well-

known factors (size, unit attributes, distance to amenities, etc.). Second, other dynamics might be at play that could determine the evolution of real-estate prices concurrently to bike-share stations. In this section, I discuss these challenges and how I address them with my estimation strategy.

The primary obstacle to determining the causal relationship between bike share and real-estate transaction prices is the non-random placement of bike share stations throughout the city. This deliberate selection process is logical from a city planning perspective, as it aims to ensure the success of the bike share program by targeting areas with a significant transportation market. In the case of NYC, this resulted in the initial deployment of bike share stations in Manhattan south of 60th Street and downtown Brooklyn.

The second challenge lies in the vast and diverse nature of NYC, which is continually evolving. Each year, numerous policies are implemented that may impact real-estate transaction prices, potentially confounding the estimation of the bike share's impact.

I address these challenges using a two-ring difference-in-differences (TRDD) strategy. The empirical strategy compares transactions taking place close to bike-share stations (inner or treatment ring) to those further away (outer or control ring), before and after the first deployment of bike-share. The two-ring construction identifies treatment (ring from zero to 150 metres around the station) and control (ring from 150 to 500 metres) transactions for each bike-share station, and estimates the difference between treatment and control transactions after bike-share implementation, net of station and time fixed effects.

The TRDD addresses the concerns around the identification by carefully selecting a credible control group of transactions (the outer ring). By focusing on potential impacts at a very small geographical scale, this identification strategy gives more ex-ante credibility to the parallel trends assumption required by difference-in-differences: everything else equal, it is reasonable to assume that transactions in the control and treatment ring follow similar sale-price trends in the absence of bikes share. I will provide evidence of this in later sections using an event study design.

The second main assumption of difference-in-differences is that no other concomitant policy that could have affected sale prices was enacted at the same time and place where bike-share was rolled out. I address this assumption by narrowing down the sample period to four and a half years (January 2011 to June 2015), which should limit the number of potential policies that might affect the estimates. Moreover, to be a significant threat to identification, other policies would have to exactly follow the spatial and temporal pattern of bike-share stations, which I control for to the best of my abilities.

To explore the impact of bike-share on sale price, I specify the following model, which uses the universe of transactions and controls for within bike-share-station sample area (or ring; less than 500 metres away) and within bike-share-station treatment area (or ring;

less than 150 metres away). This specification is formalised in the following equation:

$$\begin{aligned} \ln(P_{ijct}) = & \delta_{<150} D_{ij}^{<150} \times Post_{ijt} + \lambda_{<150} D_{ij}^{<150} \\ & + \delta_{<500} D_{ij}^{<500} \times Post_{ijt} + \lambda_{<500} D_{ij}^{<500} + \beta' X_{it} + \gamma' Z_{it} \\ & + \kappa_j + \varphi_c + \tau_t + \kappa_j \times t + \varepsilon_{ijct}, \end{aligned} \quad (1)$$

where P_{ijct} is the real sale price (base December 2015) of transaction i , matched to bike-share station j , at time t ; $D_{ij}^{<150}$ is an indicator variable which is equal to one if unit i is within 150 metres of station j (i.e., in the treatment ring), and zero otherwise (i.e., in the control ring); $D_{ij}^{<500}$ is an indicator variable similar to $D_{ij}^{<150}$ but for transactions within 500 metres of a bike-share station (i.e., the sample ring); $Post_{ijt}$ is a dummy variable indicating whether a unit is transacted after the opening of bike-share station j ; X_{it} is a vector of unit attributes (elevator, building age, etc); Z_{it} is a vector of distances to nearby (dis)amenities (subway station, bus stop); κ_j and τ_t are station and year-month fixed effects, respectively; $\kappa_j \times t$ are station-specific linear time trends (some specifications); ε_{ijct} standard errors clustered at station level.

The econometric specification used to estimate the TRDD (i.e., when restricting the sample to transactions within 500 metres of bike-share stations) is described by this second equation:

$$\begin{aligned} \ln(P_{ijct}) = & \delta_{<150} D_{ij}^{<150} \times Post_{ijt} + \lambda_{<150} D_{ij}^{<150} + \beta' X_{it} + \gamma' Z_{it} \\ & + \kappa_j + \varphi_c + \tau_t + \kappa_j \times t + \varepsilon_{ijct}, \end{aligned} \quad (2)$$

The coefficient of interest in both models 1 and 2 is $\delta_{<150}$, which represents the average treatment effect of bike-share on the treated transactions. Concretely, it is the average change in log sales for a transaction within the treatment ring (i.e., within 150 metres) of a bike-share station after the opening of the station. Alternatively, I replace the treatment ring dummy $D_{ij}^{<150}$ by a continuous measure of distance (in hundreds of metres) to the matched bike-share station D_{ij} . The coefficient in this case reports, for a transaction, the average effect (in percent) of being 100 metres further away from its matched bike-share station.

Finally, I also run a dynamic TWFE model in order to investigate the dynamic effect of bike-share with respect to the timing of treatment. The dynamic DD specification, also known as event study, plots the treatment effect for all periods. The dynamic specification also allows us to test for differential pretrends between groups: by plotting the difference between treatment and control in the pre-treatment period, we will be able to evaluate the validity of the parallel trends assumption.

The specification for the dynamic DD is given in equation 3:

$$\begin{aligned} \ln(P_{ijct}) = & \sum_{k=-6}^{-2} \beta_k \cdot D_{ij}^{<150} + \sum_{k=0}^6 \beta_k \cdot D_{ij}^{<150} + \beta' X_{it} + \gamma' Z_{it} \\ & + \kappa_j + \varphi_c + \tau_t + \kappa_j \times t + \varepsilon_{ijct}, \end{aligned} \quad (3)$$

where k denotes the relative time to the first year of treatment, the other terms being the same as in the previous specification. The coefficients of interest are β_k , which are then plotted against relative time. In this setting, the reference period is relative time $k = -1$, therefore the plotted β_k s denote the relative difference between treatment and control groups compared to the period right before treatment.

4 Results

In this section, I first concentrate on residential properties and present the results of a hedonic model, regressing the sale price of transactions on unit attributes and distances to (dis)amenities. I then explore the potential treatment effects of bike-share stations on residential transaction prices. In the second subsection, I test whether bike-share stations had an impact on the sale prices of commercial properties, and check the potential heterogeneity of these effect across types of commercial properties. Finally, I discuss the results in light of previous research.

4.1 Residential properties

4.1.1 Hedonic model

I briefly present the results of the hedonic models in this subsection. This main goal of this analysis is to validate the selection of the sample of interest, the construction of the variables (unit characteristics and amenities), and the overall soundness of the real-estate data set. Moreover, showing that prices react in a sensible manner to unit attributes and amenities strengthens the credibility of the analysis of bike-share treatment effects.

The results of the hedonic model are displayed in Table 5. Columns 1 to 3 use the universe of transactions across NYC during the sample period (2011-01-01 to Jun 2015-06-30), while columns 4 to 6 restrict the sample to transaction that took place within the sample area of bike-share stations (500 metres). The depend variable for all specifications is the log sale price,⁸ which is regressed on units' surface area, building age, and distances to bus stops, subway entrances, and the nearest park. From column 1 to 3 and 4 to 6, I add fixed-effects that control for temporal, spatial and building-class unobservable characteristics. Standard errors are clusters at the Neighbourhood Tabulation Areas

⁸Prices are deflated to June 2015 levels using the Consumer Price Index for New York-Newark-Jersey City from the Bureau of Labor Statistics, available at <https://www.bls.gov/regions/northeast/data/xg-tables/ro2xgcpiny.htm> (accessed 2023-07-26).

(NTA) level for columns 1 to 3, and at the bike-share station level for columns 4 to 6.

Table 5: Hedonic model of residential transactions' sale prices

	Log sale price (2015 \$)					
	(1)	(2)	(3)	(4)	(5)	(6)
Surface per unit (100s sqft)	0.0098*** (0.0021)	0.0076*** (0.0011)	0.0205*** (0.0012)	0.0090*** (0.0015)	0.0090*** (0.0008)	0.0242*** (0.0011)
Building age (10s years)	0.0413*** (0.0082)	-0.0146*** (0.0025)	-0.0020 (0.0020)	0.0155*** (0.0054)	-0.0019 (0.0035)	-0.0013 (0.0025)
Distance to bus stop (100s m)	-0.0024* (0.0013)	0.0042** (0.0021)	0.0024 (0.0018)	0.0694*** (0.0182)	0.0853*** (0.0161)	0.0882*** (0.0151)
Distance to subway (100s m)	-0.0015 (0.0013)	-0.0066*** (0.0021)	-0.0033* (0.0018)	-0.1061*** (0.0161)	-0.0375*** (0.0071)	-0.0320*** (0.0067)
Distance to park (100s m)	-0.0435*** (0.0076)	-0.0045* (0.0023)	-0.0037** (0.0019)	-0.1623*** (0.0118)	-0.0361*** (0.0071)	-0.0291*** (0.0068)
Neighbourhood FE (187)		Yes	Yes			
Sale year-quarter FE (18)		Yes	Yes		Yes	Yes
Building class category FE (11)			Yes			Yes
Bike-share station FE (333)					Yes	Yes
<i>Varying Slopes</i>						
Sale year-quarter (Neighbourhood)		Yes	Yes			
Sale year-quarter (Bike-share station)					Yes	Yes
Standard-Errors	Neighbourhood			Bike-share station		
Mean outcome pre-period	3,040,182	3,040,182	3,040,182	3,040,182	3,040,182	3,040,182
Observations	74,667	74,667	74,667	8,889	8,889	8,889
Adjusted R ²	0.148	0.635	0.727	0.231	0.604	0.650
Within Adjusted R ²		0.026	0.104		0.039	0.084
RMSE	0.649	0.423	0.366	0.752	0.519	0.488

Notes: Significance codes: *, 0.1, **, 0.05, ***, 0.01. Standard errors clustered at the neighbourhood-tabulation-area level in columns 1 to 3, and at the bike-share-station level for columns 4 to 6. Numbers in parenthesis next to "FE" indicate the number of fixed-effect levels for that variable.

Most of the explanatory variables have the expected sign across specifications: the larger a unit's, the closer to a subway station and a park, the higher the sales price. The coefficients for building age and distance to bus stops are less stable, but make sense in their respective specifications. In a model with no spatial nor temporal fixed effects, older buildings are associated with higher sale prices. This first reverses when controlling for time of sale and neighbourhood (younger buildings are associated with higher prices), and then disappears when controlling for building-class categories. Older buildings are associated with cachet and prestige, or might be located in areas where architecture was preserved, which may be associated with higher transaction prices. It makes sense that the impact of a building age on sale price thus disappears when comparing transactions within a given area. Distance to bus stops does not seem to explain sale prices very well when using the universe of transactions. However, being closer is a strong negative predictor of sale price for transactions within the bike-share station sample areas. A way to think about this results is that, *ceteris paribus*, a unit closer to a bus stop is probably closer to busy roadways, which are a clear disamenity.

Overall, the results from the hedonic model for both the universe of transactions and transactions within bike-share stations' sample areas indicate that the data set is reliable

and that unit attributes and distances to amenities behave in a sensible manner. In the next subsection, I present the results of the effect of bike-share stations on sale prices.

4.1.2 Treatment effect of bike-share stations

Table 6 displays the results of estimation equation 2 and 1. Column 1 and 2 use the universe of transactions and control for locations of transaction within a treatment ring (i.e., within 150 metres of a bike-share station) and within a sample ring (i.e., within 500 metres of a bike-share station). In columns 3 to 6, the sample is restricted to transactions within sample rings.

Table 6: Treatment impact on residential transaction sale prices

	Log sale price (2015 \$)					
	(1)	(2)	(3)	(4)	(5)	(6)
Treated ring \times Post-period	0.0402 (0.0472)	0.0530 (0.0511)	0.0554 (0.0474)	0.0573 (0.0396)	0.0517 (0.0329)	0.0611* (0.0328)
Treated ring	0.0058 (0.0323)	-0.0381 (0.0327)	0.1324*** (0.0487)	0.0975** (0.0416)	0.0128 (0.0235)	-0.0183 (0.0240)
Sample ring \times Post-period	-0.0371 (0.0525)	-0.0045 (0.0472)				
Sample ring	0.1340** (0.0590)	0.1025** (0.0506)				
Post-period	-0.0194 (0.1048)	-0.0499 (0.1009)	0.0884*** (0.0247)	0.1348*** (0.0235)	-0.1635*** (0.0615)	-0.1459** (0.0580)
Surface per unit (100s sqft)	0.0076*** (0.0011)	0.0205*** (0.0012)		0.0085*** (0.0015)	0.0090*** (0.0008)	0.0241*** (0.0011)
Building age (10s years)	-0.0146*** (0.0025)	-0.0020 (0.0020)		0.0144*** (0.0054)	-0.0022 (0.0035)	-0.0014 (0.0025)
Distance to bus stop (100s m)	0.0042* (0.0021)	0.0024 (0.0018)		0.0781*** (0.0180)	0.0880*** (0.0164)	0.0897*** (0.0153)
Distance to subway (100s m)	-0.0065*** (0.0021)	-0.0032* (0.0018)		-0.1075*** (0.0163)	-0.0370*** (0.0071)	-0.0319*** (0.0067)
Distance to park (100s m)	-0.0048** (0.0022)	-0.0039** (0.0018)		-0.1635*** (0.0119)	-0.0368*** (0.0072)	-0.0296*** (0.0069)
Neighbourhood FE (187)	Yes	Yes				
Sale year-quarter FE (18)	Yes	Yes			Yes	Yes
Building class category FE (11)		Yes				Yes
Bike-share station FE (333)					Yes	Yes
<i>Varying Slopes</i>						
Sale year-quarter (Neighbourhood)	Yes	Yes				
Sale year-quarter (Bike-share station)					Yes	Yes
Standard-Errors	Neighbourhood		Bike-share station			
Mean outcome pre-period	3,040,182	3,040,182	3,040,182	3,040,182	3,040,182	3,040,182
Observations	74,667	74,667	8,889	8,889	8,889	8,889
Adjusted R ²	0.636	0.727	0.008	0.241	0.605	0.650
Within Adjusted R ²	0.027	0.105			0.040	0.086
RMSE	0.423	0.366	0.854	0.747	0.518	0.487

Notes: Significance codes: *, 0.1, **, 0.05, ***, 0.01. Standard errors clustered at the neighbourhood-tabulation-area level in columns 1 to 2, and at the bike-share-station level for columns 3 to 6. Numbers in parenthesis next to “FE” indicate the number of fixed-effect levels for that variable.

Columns 1 and 2 show no statistically significant effect of being treated by bike-share on transaction prices. Transactions in sample rings (i.e., within 500 metres of a bike-share station) are about 10% more expensive in the preferred specification (column 2). This

specification, which includes neighbourhood fixed-effects, indicates that the bike-share system might have been located in the more affluent areas within neighbourhoods.

Moving on to Columns 3 to 6, I reduce the sample to only include transactions within 500 metres of bike-share stations (i.e., within the sample rings). The specifications gradually add covariates (unit attributes and distance to amenities) and fixed-effects. In column 3, the main coefficient of interest, the interaction between the treatment and post-period dummies, is positive but not statistically significant at conventional levels (p-value 0.243). Controlling for unit attributes and distance to amenities in Column 4 does not sensibly change either the magnitude or the statistical significance of the coefficient (p-value 0.148). Column 5 adds bike-share station and quarter fixed effects: the magnitude of the coefficient remains stable and its p-value improves (0.118). Adding these fixed effects, the coefficient for being in a treated ring turns statistically insignificant. In column 6, I add building-class category fixed effects and the coefficient of interest turns statistically significant (p-value 0.063). Under this specification, the impact of having a bike-share station within 150 metres of a property increases its transaction price by 6.1% with respect to property transacted between 150 and 500 metres of the same bike-share station, or \$185,755 for an average transaction price. By adding building-class category fixed effects, we effectively compare transactions within the same building category, within the same bike-share station 500-metre radius, and within a given year-quarter, controlling for unobservable common factors in each of these dimensions. Since building categories within a given area might share a considerable number of common characteristics, which are unobserved in the current data, comparing within building class might thus be important to remove heterogeneity and sharpen the estimation of our coefficient of interest. Moreover, the coefficient for the treated ring is statistically insignificant, indicating that bike-share stations are not placed in more ex-ante expensive areas with respect to their entire sample area (i.e., 500-metre radius).

The data would allow for more disaggregated fixed effects, using for example year-month of sale (instead of year-quarter), or building class (instead of building class category). While the appeal of employing more detailed fixed effects is clear, it is important to note that they also bring their own set of disadvantages. In particular, allowing for finer levels of the current fixed effects would greatly reduce the number of observations used in the estimation. The model would compare observations within each fixed-effect “cell”, i.e., it would only compare, with each other, the observations that share the same values across all fixed effects. As we introduce more granular fixed effects, it becomes increasingly likely that a given observation (transaction) is unique in a given fixed-effect cell, and has no other observation to be compared with. In this case, the estimation ignores that observation and the coefficients are estimated using fixed-effect cells with at least two observations.⁹ Therefore, I decide not to report more granular

⁹It also follows that the model may underestimate the standard errors of the estimated coefficients, which would lead us to over-reject the null hypothesis.

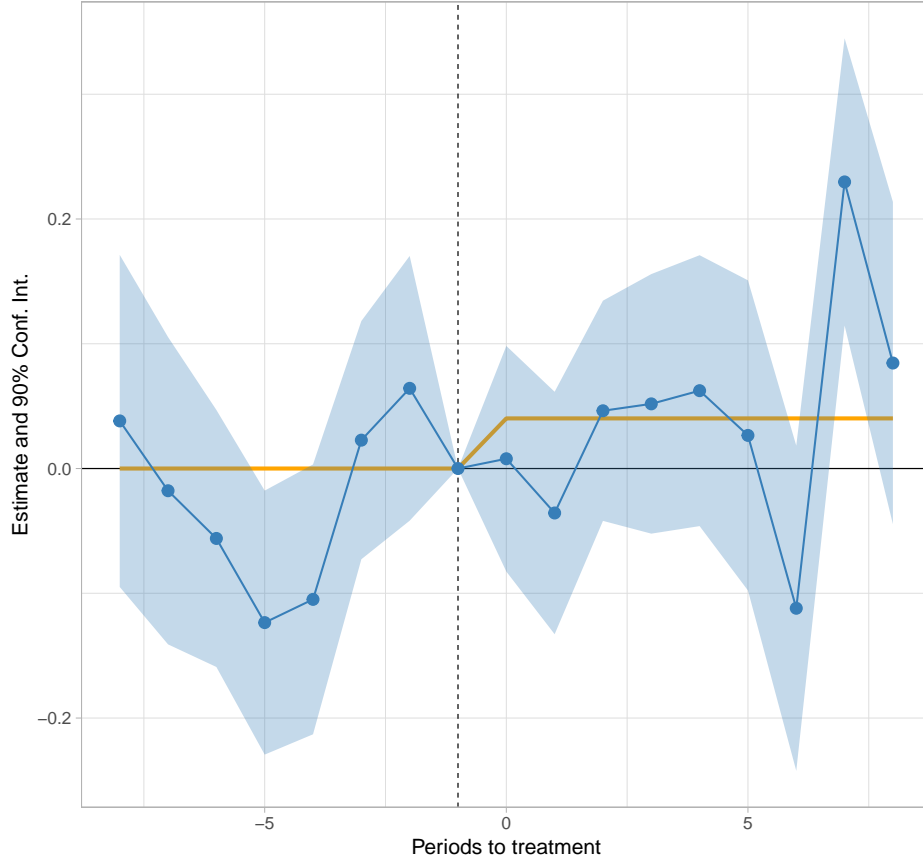


Figure 4: Dynamic effect of treatment on transaction price (event study). *Notes:* Periods to treatment is given in quarters. The shaded area indicates the coefficients' 90% confidence intervals. The orange line represents the post-period aggregated effect.

fixed effects, as those would restrict the number of observations used by the estimation.¹⁰

Figure 4 depicts the results from estimating equation 3, an event study of the effect of bike-share with respect to treatment periods. Indicative of evidence for parallel trends, pre-treatment periods display coefficients indistinguishable from zero (except for period -5): there are (almost) no statistical differences in price trends between treated and control transactions before treatment occurs. After treatment, there is a slight upward trend in the coefficients, indicating that prices in treated increased faster compared to control areas. However, the coefficient at period 7 is the only statistically significant one at the 10% significance level. Overall, the evidence from the event study does not provide very compelling evidence for the results reported in table 6.

I check the consistency of this result in Table 7. Replacing the binary treatment variable with a measure of distance to the matched bike-share station, I find that every additional hundred metres away from the station reduces transaction prices by 2.4%

¹⁰With the current set of fixed effects as in column 6, there are about five thousand observations (out of eight thousand) that belong to a fixed-effect cell with more than two observations. Introducing building class and year-month fixed effects lowers the usable observations down to two thousand. Note that when doing this, the statistical significance of the treatment effect further improves.

(preferred specification, Column 3), or \$73,816 for an average transaction price. This is in line with the result obtained with the indicator variable, and shows there exists a downward-sloping gradient of transaction prices around bike-share stations.

Table 7: Continuous treatment

	Log sale price (2015 \$)		
	(1)	(2)	(3)
Dist. to bike-share station (100s m) \times Post-period	-0.0219** (0.0110)	-0.0228** (0.0108)	-0.0243** (0.0100)
Dist. to bike-share station (100s m)	0.0033 (0.0083)	0.0014 (0.0079)	0.0076 (0.0077)
Post-period	-0.0723 (0.0687)	-0.0828 (0.0687)	-0.0608 (0.0640)
Surface per unit (100s sqft)		0.0090*** (0.0008)	0.0241*** (0.0011)
Building age (10s years)		-0.0020 (0.0035)	-0.0013 (0.0025)
Distance to bus stop (100s m)		0.0877*** (0.0165)	0.0899*** (0.0155)
Distance to subway (100s m)		-0.0369*** (0.0071)	-0.0318*** (0.0068)
Distance to park (100s m)		-0.0372*** (0.0073)	-0.0298*** (0.0070)
Bike-share station FE (333)	Yes	Yes	Yes
Sale year-quarter FE (18)	Yes	Yes	Yes
Building class category FE (11)			Yes
<i>Varying Slopes</i>			
Sale year-quarter (Bike-share station)	Yes	Yes	Yes
Mean outcome pre-period	3,040,182	3,040,182	3,040,182
Observations	8,889	8,889	8,889
Adjusted R ²	0.589	0.605	0.650
Within Adjusted R ²	0.001	0.041	0.086
RMSE	0.529	0.518	0.487

Notes: Significance codes: *: 0.1, **: 0.05, ***: 0.01. Standard errors clustered at the bike-share-station level. Numbers in parenthesis next to “FE” indicate the number of fixed-effect levels for that variable.

4.1.3 Robustness checks: alternative rings

In this subsection, I perform the analysis laid out in equation 2 for alternative definitions of rings. I use sample-ring distances of 400, 500 (baseline) and 600 metres; treatment rings of 125, 150 (baseline) and 175 metres; and, for each combination of the previous two dimensions, a buffer ring between treatment and control of 0 (baseline) and 50 metres. In specifications with a 50-metre buffer, I exclude from analysis transactions that are in

Table 8: Alternative rings

<i>Treatment ring in metres →</i>	Without buffer			With 50m buffer		
	125	150	175	125	150	175
Sample ring 400m						
Treated ring × Post-period	0.0498 (0.0350) [0.1561]	0.0485 (0.0339) [0.1539]	0.0147 (0.0341) [0.6664]	0.0506 (0.0351) [0.1504]	0.0485 (0.0340) [0.1545]	0.0284 (0.0364) [0.4357]
Observations	7,635	6,078	4,724	6,746	5,278	4,117
Adj. R ²	0.6562	0.6569	0.6828	0.6575	0.6597	0.6844
Sample ring 500m						
Treated ring × Post-period	0.0610* (0.0350) [0.0825]	0.0611* (0.0328) [0.0635]	0.0357 (0.0317) [0.2616]	0.0631* (0.0353) [0.0747]	0.0678** (0.0333) [0.0423]	0.0519 (0.0336) [0.1230]
Observations	11,414	8,889	6,673	10,525	8,089	6,066
Adj. R ²	0.6542	0.6499	0.6805	0.6547	0.6522	0.6819
Sample ring 600m						
Treated ring × Post-period	0.0565 (0.0344) [0.1015]	0.0590* (0.0315) [0.0623]	0.0424 (0.0302) [0.1609]	0.0557 (0.0345) [0.1068]	0.0635** (0.0320) [0.0480]	0.0544* (0.0316) [0.0865]
Observations	16,090	12,413	9,230	15,201	11,613	8,623
Adj. R ²	0.6499	0.6460	0.6778	0.6502	0.6481	0.6787

Notes: Significance codes: *: 0.1, **: 0.05, ***: 0.01. Standard errors clustered at the bike-share-station level are given in parenthesis. *p-values* are given in square brackets.

a 50-meter ring outside the treatment ring. For example, if the treatment ring is 125 metres, transactions that are between 125 and 175 metres are excluded.

The results are summarised in table 8. All reported results follow the specification used in column 6 of table 6, with sale year-quarter, bike-share station and building-class category fixed effects, full set of covariates (unit attributes and distances to amenities), and standard errors clustered at the bike-share-station level. The column 2, second panel corresponds to the results of table 6 (i.e., the baseline case).

Table 8 shows that the results obtained earlier hold for seven out of twelve models in sample rings of 500 and 600 meters, with magnitudes remaining relatively stable.¹¹ On the other hand, the estimations for the 400-metre sample ring are inconclusive.¹² Overall, these results suggest that there is a treatment effect of bike-share stations on residential property prices between 125 and 175 metres away from stations, compared to transactions within a 500 to 600-metre radius around the station.

¹¹In addition, in the 600-metre sample ring, the treatment effect coefficients for the two 125-metre treatment ring models are significant at just over 10%.

¹²This might be explained by the loss in power when reducing the diameter of sample ring, as shown by the declining number of observations.

4.2 Commercial properties

I now turn to transactions of commercial properties. Due to the important differences that exist between building-class categories in the commercial properties realm, I perform the analysis for the different categories separately. In the main text, I only present the results for offices and retail properties (table 9). The other categories of commercial properties are commercial garages, factories and warehouses, and the results of treatment regressions on these categories can be found in appendix C.2. I do not expect these latter categories to be impacted by the implementation of the bike-share system for several reasons. First, bike-share stations were not opened near industrial areas where factories and warehouses are located. The primary goal of the system was to serve “busy” areas made up of workplaces, offices, retail and residential zoning. Second, there seems to be fewer reasons to believe that agents transacting these types of properties include “bikeability” in their decision function, and most probably give disproportionately more weight to other factors.

Starting from the hypothesis that bike-share had a positive impact on office and retail property prices, the coefficients in table 9 have the expected sign, but fall short of statistical significance at conventional levels in the preferred specification (columns 5 and 6). Interestingly, subway stations appear to be an important disamenity for both retail and office properties within bike-share-station sample areas: the closer the subway entrance, the lower the price.

As expected, the results for other categories of commercial properties (commercial garages, factories and warehouse) are also not statistically significant. The results on these other categories may be interpreted as a successful placebo analysis, showing that bike-share did not affect the sale prices of properties we would not expect it would.

Appendix C.1 presents the results of a hedonic model on all types of commercial property.

4.3 Discussion

The results displayed above indicate that the introduction of bike-share stations had a positive impact on residential real-estate transaction prices. Properties within 150 metres of a station saw their transaction price increase by 5.8% on average, compared to property transactions further away (between 150 and 500 metres), and within a given building class category. Moreover, every additional hundred metres away from a bike-share station reduces transaction price by 2.1% on average, compared to before the opening of the system.

These impacts on real-estate transaction prices are relatively significant, but within the range we would expect from this type of intervention (see [citations]). Of particular interest to us are the estimated effects of bike-share on rents determined by Shr et al. for Taiwan. They find that six months after the opening of bike-share, rents increase on average by 1.71% for units within 150 metres of stations. While of the same order

Table 9: Treatment impact on commercial transaction sale prices, by building class category

<i>Building class category</i> →	Log sale price (2015 \$)					
	Offices (1)	Retail (2)	Offices (3)	Retail (4)	Offices (5)	Retail (6)
Treated ring × Post-period	0.5181** (0.2323)	0.2587 (0.1758)	0.2049 (0.2888)	0.2589 (0.1753)	0.3098 (0.4191)	0.1870 (0.2397)
Treated ring	-0.2149 (0.2336)	-0.3021** (0.1194)	-0.3234 (0.2552)	-0.3115*** (0.1197)	-0.5807 (0.4115)	-0.2781 (0.1778)
Post-period	-0.0583 (0.1017)	0.3162*** (0.0722)	0.4725 (0.3440)	0.3937** (0.1904)	-0.1954 (0.9012)	0.4713* (0.2643)
Surface per unit (100s sqft)	0.0010*** (0.0002)	0.0016*** (0.0006)	0.0017*** (0.0003)	0.0054*** (0.0008)	0.0017*** (0.0004)	0.0070*** (0.0011)
Building age (10s years)	0.0260 (0.0236)	0.0659*** (0.0138)	0.0194 (0.0429)	-0.0220 (0.0168)	0.0184 (0.0838)	-0.0356 (0.0234)
Distance to bus stop (100s m)	0.2728 (0.1727)	0.0622 (0.0871)	-0.0368 (0.2910)	-0.1523* (0.0904)	-0.2521 (0.4665)	-0.1480 (0.1183)
Distance to subway (100s m)	0.0526 (0.0610)	0.0579** (0.0248)	0.3327** (0.1398)	0.0393 (0.0266)	0.4851*** (0.1830)	0.0772** (0.0374)
Distance to park (100s m)	-0.0973* (0.0563)	0.0174 (0.0297)	0.1387* (0.0822)	0.0112 (0.0254)	0.2059* (0.1146)	0.0008 (0.0386)
Bike-share station FE			Yes	Yes	Yes	Yes
Sale year-quarter FE (18)			Yes	Yes	Yes	Yes
<i>Varying Slopes</i>						
Sale year-quarter (Bike-share station)					Yes	Yes
# Bike-share station	—	—	189	284	189	284
Mean outcome pre-period	2,989,510	2,950,488	2,989,510	2,950,488	2,989,510	2,950,488
Observations	492	1,211	492	1,211	492	1,211
Adjusted R ²	0.052	0.050	0.402	0.327	-0.149	0.271
Within Adjusted R ²			0.157	0.075	0.198	0.121
RMSE	1.194	1.017	0.719	0.742	0.564	0.639

Notes: Significance codes: *: 0.1, **: 0.05, ***: 0.01. Standard errors clustered at the bike-share-station level. “# [FE variable]” indicate the number of fixed-effect levels for that variable.

of magnitude, my results are relatively larger. [Should I try to give some ex-post rationalisations?].

[Should I discuss the limited statistical significance of the main coefficient more in-depth?]

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A Building decades

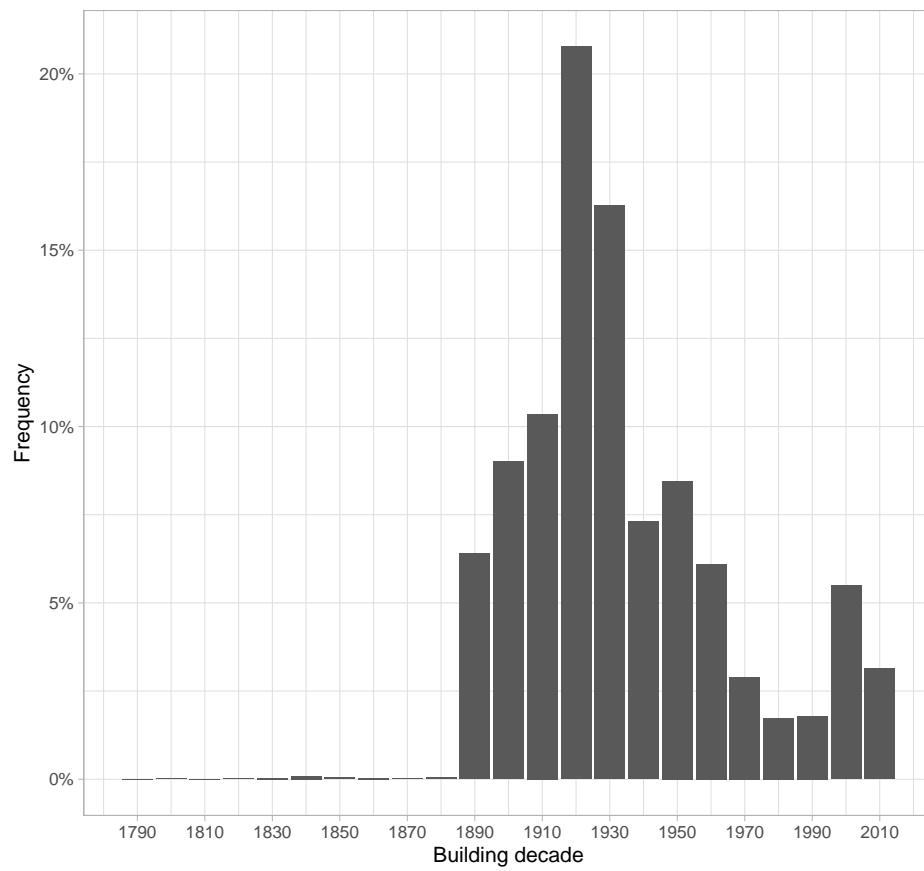


Figure A.1: Building decade of transactions in main sample data set (i.e., residential units, no price and price-per-square-feet outliers).

B Detailed treatment map

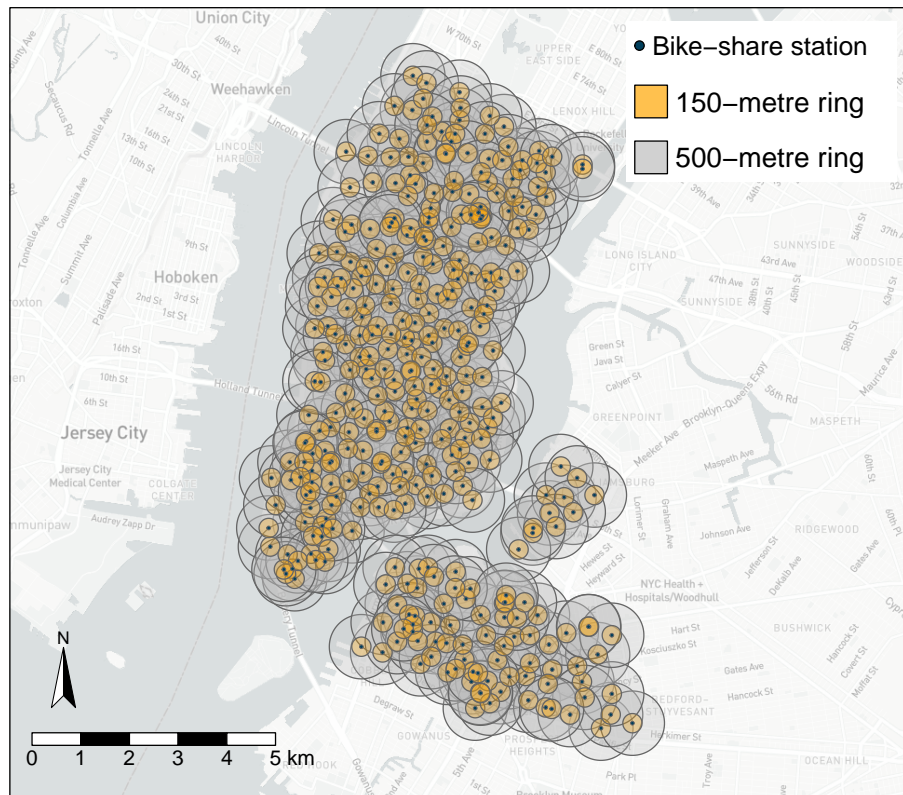


Figure B.1: Overview of the study area and treatment and control rings, not merged.
Notes: The yellow areas represent the treatment rings (i.e., within 150 metres of a bike-share station), while the grey areas indicate the controls rings (i.e., between 150 and 500 metres from a bike-share station).

C Additional analyses on commercial properties

C.1 Hedonic regressions

Table C.1: Hedonic model — offices and retail transactions

<i>Building class category</i> →	Log sale price (2015 \$)							
	Offices (1)	Retail (2)	Offices (3)	Retail (4)	Offices (5)	Retail (6)	Offices (7)	Retail (8)
Surface per unit (100s sqft)	0.0017*** (0.0006)	0.0016 (0.0015)	0.0017*** (0.0005)	0.0016 (0.0014)	0.0017*** (0.0006)	0.0016 (0.0015)	0.0018*** (0.0004)	0.0071*** (0.0011)
Building age (10s years)	0.0367 (0.0313)	0.0501** (0.0235)	-0.0616*** (0.0192)	-0.0654*** (0.0123)	-0.0767*** (0.0244)	-0.0594*** (0.0151)	0.0476 (0.0822)	-0.0355 (0.0234)
Distance to bus stop (100s m)	-0.0067* (0.0037)	-0.0098*** (0.0029)	-0.0150 (0.0162)	0.0033 (0.0091)	-0.0118 (0.0212)	0.0009 (0.0102)	-0.2205 (0.4861)	-0.1677 (0.1179)
Distance to subway (100s m)	-0.0077** (0.0038)	-0.0020 (0.0032)	-0.0075 (0.0160)	-0.0122 (0.0081)	-0.0079 (0.0205)	-0.0117 (0.0090)	0.4931*** (0.1816)	0.0775** (0.0367)
Distance to park (100s m)	-0.1279*** (0.0328)	-0.0593*** (0.0180)	-0.0043 (0.0360)	-0.0022 (0.0135)	-0.0118 (0.0415)	0.0015 (0.0146)	0.1586 (0.1141)	0.0017 (0.0397)
Neighbourhood FE			Yes	Yes	Yes	Yes		
Sale year-quarter FE (18)			Yes	Yes	Yes	Yes	Yes	Yes
Bike-share station FE							Yes	Yes
<i>Varying Slopes</i>								
Sale year-quarter (Neighbourhood)					Yes	Yes		
Sale year-quarter (Bike-share station)							Yes	Yes
# Neighbourhood	—	—	153	183	153	183	—	—
# Bike-share station	—	—	—	—	—	—	189	284
Standard-Errors			Neighbourhood				Bike-share station	
Mean outcome pre-period	3,035,532	3,094,887	3,035,532	3,094,887	3,035,532	3,094,887	2,989,510	2,950,488
Observations	1,175	3,212	1,175	3,212	1,175	3,212	492	1,211
Adjusted R ²	0.206	0.115	0.591	0.500	0.580	0.519	-0.143	0.268
Within Adjusted R ²			0.078	0.043	0.089	0.041	0.202	0.117
RMSE	1.285	1.236	0.854	0.900	0.795	0.855	0.572	0.642

Notes: Significance codes: *: 0.1, **: 0.05, ***: 0.01. Standard errors clustered at the neighbourhood-tabulation-area level in columns 1 to 6, and at the bike-share station level for columns 7 to 8. “# [FE variable]” indicate the number of fixed-effect levels for that variable.

Table C.2: Hedonic model — commercial garage, factory and warehouse transactions

<i>Building class category</i> →	Log sale price (2015 \$)											
	Garages (1)	Factories (2)	Warehouses (3)	Garages (4)	Factories (5)	Warehouses (6)	Garages (7)	Factories (8)	Warehouses (9)	Garages (10)	Factories (11)	Warehouses (12)
Surface per unit (100s sqft)	0.0024*** (0.0007)	0.0014** (0.0006)	0.0005* (0.0002)	0.0016*** (0.0004)	0.0014** (0.0005)	0.0005** (0.0002)	0.0015*** (0.0004)	0.0015** (0.0006)	0.0005** (0.0003)	0.0030*** (0.0009)	0.0074*** (0.0022)	0.0018*** (0.0005)
Building age (10s years)	0.0128 (0.0210)	-0.0126 (0.0274)	0.0282 (0.0238)	-0.0155 (0.0164)	-0.0328 (0.0346)	-0.0025 (0.0170)	-0.0153 (0.0203)	-0.0365 (0.0345)	-0.0107 (0.0186)	0.0203 (0.0671)	-0.0067 (0.0696)	-0.0707* (0.0382)
Distance to bus stop (100s m)	-0.0099*** (0.0025)	-0.0069** (0.0033)	-0.0116*** (0.0032)	0.0179 (0.0118)	-0.0243 (0.0247)	-0.0188 (0.0164)	0.0209 (0.0145)	-0.0193 (0.0314)	-0.0125 (0.0215)	-0.3056 (0.5219)	-0.1726 (0.4458)	-0.0637 (0.4031)
Distance to subway (100s m)	0.0013 (0.0027)	-0.0079 (0.0056)	0.0025 (0.0037)	-0.0293*** (0.0112)	0.0026 (0.0183)	0.0047 (0.0133)	-0.0318** (0.0132)	0.0058 (0.0187)	0.0064 (0.0167)	-0.0978 (0.1641)	0.0257 (0.1143)	0.0010 (0.0479)
Distance to park (100s m)	-0.0236 (0.0259)	-0.0053 (0.0135)	-0.0374 (0.0243)	0.0222 (0.0147)	-0.0275* (0.0166)	-0.0242 (0.0250)	0.0267 (0.0175)	-0.0192 (0.0171)	-0.0315 (0.0259)	0.1016 (0.2422)	0.0505 (0.0642)	0.0198 (0.1346)
Neighbourhood FE				Yes	Yes	Yes	Yes	Yes	Yes			
Sale year-quarter FE				Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Bike-share station FE										Yes	Yes	Yes
<i>Varying Slopes</i>												
Sale year-quarter (Neighbourhood)							Yes	Yes	Yes			
Sale year-quarter (Bike-share station)										Yes	Yes	Yes
# Neighbourhood	—	—	—	169	104	123	169	104	123	—	—	—
# Sale year-quarter	—	—	—	18	18	18	18	18	18	18	17	17
# Bike-share station	—	—	—	—	—	—	—	—	—	113	62	117
Standard-Errors				Neighbourhood						Bike-share station		
Mean outcome pre-period	2,845,757	3,251,103	3,110,593	2,845,757	3,251,103	3,110,593	2,845,757	3,251,103	3,110,593	2,695,724	2,848,955	2,998,746
Observations	1,063	811	989	1,063	811	989	1,063	811	989	245	174	242
Adjusted R ²	0.149	0.130	0.142	0.509	0.284	0.421	0.496	0.282	0.432	9.231	-0.861	2.445
Within Adjusted R ²				0.085	0.131	0.092	0.074	0.144	0.103	1.584	0.033	0.620
RMSE	1.172	1.065	1.110	0.809	0.891	0.845	0.736	0.822	0.773	0.424	0.666	0.276

Notes: Significance codes: *: 0.1, **: 0.05, ***: 0.01. Standard errors clustered at the neighbourhood-tabulation-area level in columns 1 to 9, and at the bike-share-station level for columns 10 to 12. “# [FE variable]” indicate the number of fixed-effect levels for that variable.

C.2 Treatment regressions, placebo analysis

Table C.3: Treatment impact on sale prices — commercial garage, factory and warehouse transactions

<i>Building class category</i> →	Log sale price (2015 \$)								
	Garages (1)	Factories (2)	Warehouses (3)	Garages (4)	Factories (5)	Warehouses (6)	Garages (7)	Factories (8)	Warehouses (9)
Treated ring × Post-period	0.4878 (0.3543)	-1.0683** (0.4767)	-0.4246 (0.5278)	0.2818 (0.3499)	0.0856 (0.9439)	-1.2943** (0.5214)	0.0184 (0.7088)	0.3940 (1.3875)	-2.1961** (0.8649)
Treated ring	0.1286 (0.2374)	0.7584** (0.2986)	0.1720 (0.3721)	-0.2244 (0.2524)	0.0273 (0.7015)	0.4632 (0.3719)	-0.0444 (0.4706)	-0.3329 (0.8909)	1.2072* (0.6361)
Post-period	-0.4632*** (0.1499)	0.5025*** (0.1362)	0.1817 (0.1634)	0.2908 (0.6291)	0.7233* (0.3772)		0.2754 (1.1324)	0.4847 (0.4803)	
Surface per unit (100s sqft)	0.0019*** (0.0005)	0.0049*** (0.0008)	0.0006*** (0.0001)	0.0013*** (0.0004)	0.0069*** (0.0015)	0.0007*** (0.0001)	0.0030*** (0.0009)	0.0076*** (0.0024)	0.0013** (0.0005)
Building age (10s years)	0.0500* (0.0280)	-0.0300 (0.0353)	-0.0268 (0.0302)	0.0522 (0.0352)	-0.0002 (0.0501)	-0.0730*** (0.0259)	0.0253 (0.0873)	-0.0133 (0.0781)	-0.0866** (0.0340)
Distance to bus stop (100s m)	-0.0636 (0.2489)	0.3537* (0.2083)	0.2850** (0.1186)	0.0147 (0.2258)	0.0290 (0.3974)	-0.3148** (0.1436)	-0.3104 (0.6193)	-0.1716 (0.4315)	-0.7590*** (0.2792)
Distance to subway (100s m)	-0.0653* (0.0372)	-0.1336*** (0.0281)	-0.1090*** (0.0381)	-0.1302 (0.0908)	0.0200 (0.0670)	0.0167 (0.0371)	-0.0973 (0.1759)	0.0183 (0.1130)	0.0012 (0.0451)
Distance to park (100s m)	-0.1439*** (0.0477)	-0.0385 (0.0328)	-0.0021 (0.0381)	0.0842 (0.1083)	0.0160 (0.0534)	-0.0113 (0.0804)	0.1032 (0.2735)	0.0385 (0.0724)	-0.0164 (0.1455)
Bike-share station FE				Yes	Yes	Yes	Yes	Yes	Yes
Sale year-quarter FE				Yes	Yes	Yes	Yes	Yes	Yes
<i>Varying Slopes</i>									
Sale year-quarter (Bike-share station)							Yes	Yes	Yes
# Bike-share station	—	—	—	113	62	117	113	62	117
# Sale year-quarter	—	—	—	18	17	17	18	17	17
Mean outcome pre-period	2,695,724	2,848,955	2,998,746	2,695,724	2,848,955	2,998,746	2,695,724	2,848,955	2,998,746
Observations	245	174	242	245	174	242	245	174	242
Adjusted R ²	0.189	0.316	0.093	0.674	0.272	0.716	5.113	-1.068	2.015
Within Adjusted R ²				0.057	0.120	0.336	1.292	-0.075	0.733
RMSE	1.182	0.962	0.924	0.505	0.725	0.342	0.424	0.664	0.248

Notes: Significance codes: *: 0.1, **: 0.05, ***: 0.01. Standard errors clustered at the bike-share-station level. “# [FE variable]” indicate the number of fixed-effect levels for that variable.