## Learning Theory Notes

immediate

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## **Learning Limits**

## The Simulation Sample $\bar{B}$ Estimator

Our goal is to approximate the function B = f(X) with some learned function  $\widehat{B} = \widehat{f}(X)$  from evolutionary simulations. In each evolutionary simulation, we sample some evolutionary parameters  $X_k$  and evolve r independent populations forward in time under these parameters and observe some genealogy. From this genealogy, we estimate the reduction in neutral diversity as:

$$\bar{B}_k = -\frac{1}{r} \sum_{i=1}^r \frac{\hat{\pi}_{k,i}}{4N\mu} \tag{1}$$

there  $\hat{\pi}$  is Tajima's estimator for pairwise diversity within a tree. When we estimate  $\hat{\pi}$  from using branch-mode tree statistics from an observed genealogy,  $\mu \to 1$ , so we can ignore mutation. If we take expectation over the evolutionary process,

$$\mathbb{E}(\bar{B}_k) = \frac{1}{4Nr} \sum_{i=1}^r \mathbb{E}(\hat{\pi}_{k,i})$$
 (2)

$$=\frac{T_k^{(2)}}{4N}\tag{3}$$

$$=\frac{4B_kN}{4N}\tag{4}$$

$$=B_k \tag{5}$$

(6)

thus, the estimated reduction in diversity from simulations  $\bar{B}$  is unbiased, since  $\mathbb{E}(\hat{\pi}_k) = 2T_k^{(2)} = 4B_kN$ . For independent evolutionary replicates, we calculate the variance of this estimator using Tajima's equation for the variance of  $\hat{\pi}$  as

$$Var(\bar{B}_k) = \frac{1}{16N^2r^2} \sum_{i=1}^r Var(\hat{\pi}_{k,i})$$
 (7)

$$=\frac{\operatorname{Var}(\hat{\pi}_{k,i})}{16N^2r}\tag{8}$$

$$=\frac{1}{16N^2r}\left(\frac{n+1}{3(n-1)}\theta + \frac{2(n^2+n+3)}{9n(n-1)}\theta^2\right) \tag{9}$$

$$= \frac{1}{16N^2r} \left( \frac{n+1}{3(n-1)} 4BN + \frac{2(n^2+n+3)}{9n(n-1)} 16B^2N^2 \right)$$
 (10)

$$= \underbrace{\frac{n+1}{12(n-1)} \frac{B}{Nr}}_{\text{sampling noise}} + \underbrace{\frac{2(n^2+n+3)}{9n(n-1)} \frac{B^2}{r}}_{\text{evolutionary variance}}$$
(11)

Now, let us look at the consistency of the estimator  $\bar{B}$  (we drop the parameter set k for clarity) both in r (over evolutionary replicates) and in n (as the sample size increases). Let  $\bar{B}_r$  be the estimator of B after r evolutionary replicates (conditioning on some n). By Chebyshev's inequality and for some  $\epsilon > 0$ ,

$$\mathbb{P}\left(|\bar{B}_r - B| \ge \epsilon\right) \le \frac{\operatorname{Var}(\bar{B}_r)}{\epsilon^2} \tag{12}$$

and since  $\lim_{r\to\infty} \operatorname{Var}(\bar{B}_r) = 0$ ,

$$\lim_{r \to \infty} \mathbb{P}\left(|\bar{B}_r - B| \ge \epsilon\right) = 0. \tag{13}$$

Thus,  $\bar{B}_r \stackrel{p}{\to} B$  as  $r \to \infty$  and  $\bar{B}_r$  is consistent in r. Now, let us look at the consistency of  $\bar{B}_n$  in sample size n. Note that  $n \le 2N$  (i.e. our sample size is bounded by the number of gametes in the population), so we imagine setting n = 2N and taking the limit  $N \to \infty$ ,

$$\lim_{N \to \infty} \mathbb{P}\left(|\bar{B}_N - B| \ge \epsilon\right) \le \lim_{N \to \infty} \frac{\operatorname{Var}(\bar{B}_N)}{\epsilon^2} \tag{14}$$

$$\lim_{N \to \infty} \mathbb{P}\left(|\bar{B}_N - B| \ge \epsilon\right) \le \frac{2B^2}{9r\epsilon^2}.\tag{15}$$

(16)

Thus, even if we sample the entire population, and let the population size  $N \to \infty$ ,  $\bar{B}_n$  is still an inconsistent estimator in n. Intuitively this is because

In our case, we estimate pairwise diversity from the genealogical tree of the entire population, so n=2N

$$\operatorname{Var}(\bar{B}_k) = \frac{n+1}{12(n-1)} \frac{B}{Nr} + \frac{2(n^2+n+3)}{9n(n-1)} \frac{B^2}{r}$$

$$= \frac{2N+1}{12(2N-1)} \frac{B}{Nr} + \frac{2(4N^2+2N+3)}{18N(2N-1)} \frac{B^2}{r}$$
(18)

$$= \frac{2N+1}{12(2N-1)} \frac{B}{Nr} + \frac{2(4N^2+2N+3)}{18N(2N-1)} \frac{B^2}{r}$$
 (18)

$$\approx \frac{B}{12Nr} + \frac{2N^2 + N}{9N^2} \frac{B^2}{r} \tag{19}$$

If we take the infinite population size limit,

$$\lim_{N \to \infty} \operatorname{Var}(\bar{B}_k) = \frac{2B^2}{9r} \tag{20}$$