Weighted Least Squares (WLS)

Intro

• The Weighted Least Squares (WLS) method is an extension of the ordinary least squares (OLS) technique. The model is formulated the same way as in OLS:

$$\hat{y}(\boldsymbol{x}) = \boldsymbol{x}^{\mathsf{T}} \hat{\boldsymbol{\beta}} + \varepsilon(\boldsymbol{x})$$

where \boldsymbol{x} is the vector of features, $\boldsymbol{\beta}$ is the vector of parameters, $\boldsymbol{y}(\boldsymbol{x})$ is the target variable, and $\varepsilon(\boldsymbol{x})$ is the error term.

- Each observation x has associated weights w(x) that reflect the importance of that particular observation.
- This method minimizes the weighted sum of squared residuals:

$$\mathrm{RSS} = \sum_{\boldsymbol{x} \in X^{\ell}} w(\boldsymbol{x}) \cdot \left(y(\boldsymbol{x}) - \hat{y}(\boldsymbol{x})\right)^2 \to \min_{\beta}.$$

• The solution to this minimization problem is given by:

$$\boldsymbol{\beta}^* = \underbrace{\left(\boldsymbol{X}^\mathsf{T} \boldsymbol{W} \boldsymbol{X}\right)^{-1} \boldsymbol{X}^\mathsf{T} \boldsymbol{W}}_{\boldsymbol{X}^\mathsf{T}_{\boldsymbol{W}}} \boldsymbol{y},$$

where W is the diagonal matrix of weights, and X_W^+ is the weighted pseudo-inverse.

FORMALISM

Weight matrix

For a weighted

$$\mathrm{RSS} = \sum_{\boldsymbol{x} \in X^\ell} w(\boldsymbol{x}) \cdot \left(y(\boldsymbol{x}) - \hat{y}(\boldsymbol{x})\right)^2$$

let's introduce the weight matrix:

$$\begin{split} W := \operatorname{diag}(w(\boldsymbol{x}_1), ..., w(\boldsymbol{x}_\ell)) \\ = \begin{pmatrix} w(\boldsymbol{x}_1) & & \\ & \ddots & \\ & & w(\boldsymbol{x}_\ell) \end{pmatrix} \end{split}$$

Matrix form

Thus, we can rewrite the RSS in matrix form as a quadratic form:

$$RSS = (\boldsymbol{y} - X\boldsymbol{\beta})^{\mathsf{T}} W(\boldsymbol{y} - X\boldsymbol{\beta}).$$

Transformation to standard LS problem

The weighted LS problem can be easily reformulated as a standard LS problem by replacing the original variables with transformed ones:

$$\boldsymbol{y}'\coloneqq W^{\frac{1}{2}}\boldsymbol{y}, \quad X'\coloneqq W^{\frac{1}{2}}X, \quad \varepsilon'\coloneqq W^{\frac{1}{2}}\varepsilon$$

Substituting these transformations into the original model, we get:

$$y' = X'\beta + \varepsilon'$$

Heteroscedasticity can be eliminated by applying weighted LS.

For a model with non-constant variance of the error term:

$$y = X\beta + \varepsilon$$
, $Var[\varepsilon(x)] = y(x)^2 \cdot \sigma^2$

To apply WLS, the weights must have a negative square unit:

$$w(\boldsymbol{x}) = \frac{1}{y(\boldsymbol{x})^2}$$

This leads to the transformations:

$$oldsymbol{y}' = rac{oldsymbol{y}}{\sqrt{w}}, \quad oldsymbol{x}' = rac{oldsymbol{arepsilon}}{\sqrt{w}}, \quad oldsymbol{arepsilon}' = rac{oldsymbol{arepsilon}}{\sqrt{w}}$$

The weight matrix is:

$$W = egin{pmatrix} rac{1}{y(oldsymbol{x}_1)^2} & & & \\ & \ddots & & \\ & & rac{1}{y(oldsymbol{x}_\ell)^2} \end{pmatrix}$$

and

$$y' = \sqrt{W}y$$
, $x' = \sqrt{W}x$, $\varepsilon' = \sqrt{W}\varepsilon$

Now, the model can be formulated as a homoscedastic least squares problem:

$$y' = X'\beta + \varepsilon', \quad \operatorname{Var}[\varepsilon'(\boldsymbol{x})] = \sigma^2$$

Quadratic form is a function of the form:

$$Q(x_1,...,x_n) = \sum_{i=1}^n \sum_{i=1}^n a_{i,j} x_i x_j.$$

Coefficients $a_{i,j}$ can be arranged in a symmetric matrix A, and the quadratic form can be written in matrix form as:

$$Q(x) = x^T A x.$$

Now, let's solve for β in the transformed model. Since W and $W^{\{\frac{1}{2}\}}$ are diagonal matrices, transposing them results in the same matrix:

$$\boldsymbol{\beta}^* = {\boldsymbol{X}'}^+ \boldsymbol{y}' = \left({\boldsymbol{X}'}^\mathsf{T} {\boldsymbol{X}'}\right)^{-1} {\boldsymbol{X}'}^\mathsf{T} \boldsymbol{y}'$$

Expanding the expressions:

$$\begin{split} \boldsymbol{\beta}^* &= \left(\left(W^{\frac{1}{2}} \boldsymbol{X} \right)^T W^{\frac{1}{2}} \boldsymbol{X} \right)^{-1} \left(W^{\frac{1}{2}} \boldsymbol{X} \right)^\mathsf{T} W^{\frac{1}{2}} \boldsymbol{y} \\ &= \left(\boldsymbol{X}^\mathsf{T} W \boldsymbol{X} \right)^{-1} \boldsymbol{X}^\mathsf{T} W \boldsymbol{y} \end{split}$$

Therefore, the solution is:

$$\boldsymbol{\beta}^* = \underbrace{\left(\boldsymbol{X}^\mathsf{T} \boldsymbol{W} \boldsymbol{X}\right)^{-1} \boldsymbol{X}^\mathsf{T} \boldsymbol{W}}_{\boldsymbol{X}_W^+} \boldsymbol{y}$$