

Automatic Relevance Determination in Nonnegative Matrix Factorization with the β -Divergence

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Nonnegative Matrix Factorization

• Given a nonnegative data matrix $\mathbf{V} \in \mathbb{R}_{+}^{F \times N}$, find a decomposition

$$\mathbf{V} \approx \hat{\mathbf{V}} \triangleq \mathbf{W}\mathbf{H}$$

where $\mathbf{W} \in \mathbb{R}_+^{F \times K}$ (basis) and $\mathbf{H} \in \mathbb{R}_+^{K \times N}$ (coefficients)

- Common dimension K is chosen such that $FK + KN \ll FN$
- Overall number of parameters to describe the data is reduced
- Alternating minimization usually performed

$$\min_{\mathbf{H}>0} D(\mathbf{V}|\mathbf{W}\mathbf{H}), \qquad \min_{\mathbf{W}>0} D(\mathbf{V}|\mathbf{W}\mathbf{H})$$

• The measure of fit D(V|WH) is separable, i.e.,

$$D(\mathbf{V}|\hat{\mathbf{V}}) = \sum_{f=1}^{F} \sum_{n=1}^{N} d(v_{fn}|\hat{v}_{fn})$$

• We take the scalar cost function $d(v_{fn}|\hat{v}_{fn})$ to be the so-called β -divergence. Special cases include:

 $\beta = 0$: Itakura-Saito divergence

 $\beta=1$: (Generalized) Kullback-Leibler divergence

 $\beta=2\,$: (Squared) Euclidean distance

Main Contributions

- • In practical applications, model order ${\cal K}$ is hard to choose
- If K is too large \Rightarrow overfitting \odot
- If K is too small \Rightarrow data does not fit well to the model \ominus
- Bayesian NMF model based on Automatic Relevance Determination to estimate K and get a better decomposition ⊚
- ARD has been previously employed in Bayesian PCA (Bishop 1999) and sparse Bayesian learning (Tipping 2001).

Majorization-Minimization (MM) for β -NMF

- Algorithms are based on the MM framework
- Let the cost function to be minimized be $C(\mathbf{H})$
- Build an auxiliary function $G(\mathbf{H}|\tilde{\mathbf{H}})$ such that

$$G(\mathbf{H}|\tilde{\mathbf{H}}) \ge C(\mathbf{H}), \qquad G(\tilde{\mathbf{H}}|\tilde{\mathbf{H}}) = C(\tilde{\mathbf{H}})$$

• Then, optimizing $G(\cdot | \mathbf{H}^{(i)})$ yields

$$C(\mathbf{H}^{(i+1)}) \le C(\mathbf{H}^{(i+1)}|\mathbf{H}^{(i)}) \le C(\mathbf{H}^{(i)}|\mathbf{H}^{(i)}) = C(\mathbf{H}^{(i)})$$

• Hence, the MM updates consists in performing

$$\mathbf{H}^{(i+1)} = \operatorname*{arg\,min}_{\mathbf{H} \geq 0} G(\mathbf{H}|\mathbf{H}^{(i)}).$$

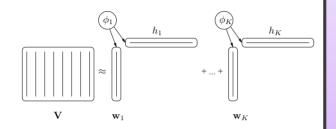
 For β-NMF, G(H|H) can be found (Févotte and Idier 2011) and this amounts to the simple multiplicative update rule

$$h_{kn} = \tilde{h}_{kn} \left(\frac{p_{kn}}{q_{kn}} \right)^{\gamma(\beta)}$$

where p_{kn}, q_{kn} are simple functions of the previous iterate and the data and $\gamma(\beta)$ is simple exponent in β .

The Model for ARD in β -NMF

 Main idea is to tie column w_k and row h_k through their prior, a common (variance-like) relevance weight φ_k



For ℓ₂-ARD, each element of the matrices W and H is assigned a Half-Normal Prior

$$p(h_{kn}|\phi_k) = \sqrt{\frac{2}{\pi\phi_k}} \exp\left(-\frac{h_{kn}^2}{2\phi_k}\right)$$

• For ℓ_1 -ARD, each element of the matrices W and H is assigned an Exponential Prior

$$p(h_{kn}|\phi_k) = \frac{1}{\phi_k} \exp\left(-\frac{h_{kn}}{\phi_k}\right)$$

• Relevance weights ϕ_k assigned inverse-Gamma priors:

$$p(\phi_k; a, b) = \frac{b^a}{\Gamma(a)} \exp\left(-\frac{b}{\phi_k}\right)$$

• Different β 's underlie different statistical noise models:

$$\begin{split} \text{IS-NMF}: & \beta = 0 & v_{fn} \sim \mathcal{G}(v_{kn}|\alpha, \hat{v}_{fn}/\alpha) \\ \text{KL-NMF}: & \beta = 1 & v_{fn} \sim \mathcal{P}(v_{kn}|\hat{v}_{fn}) \\ \text{EUC-NMF}: & \beta = 2 & v_{fn} \sim \mathcal{N}(v_{kn}|\hat{v}_{fn}, \sigma^2) \end{split}$$

• The likelihood is given by

$$-\log p(\mathbf{V}|\mathbf{W}, \mathbf{H}) = \rho D_{\beta}(\mathbf{V}|\mathbf{W}\mathbf{H}) + \text{cst}$$

where ρ is some regularization constant reflecting our belief in the noise power, e.g., $\rho=1/\sigma^2$ for $\beta=2$.

The Overall Cost Function (Posterior)

• Combining the likelihood and the priors gives the cost function (posterior): $C(\mathbf{W}, \mathbf{H}, \lambda) = -\log p(\mathbf{W}, \mathbf{H}, \lambda | \mathbf{V}) =$

$$\rho D_{\beta}(\mathbf{V}|\mathbf{WH}) + \sum_{k=1}^{K} \frac{1}{\phi_k} (f(\mathbf{w}_k) + f(h_k) + b) + c \log \phi_k + \text{cst.}$$

where $f(\mathbf{x})=\frac{1}{2}\|\mathbf{x}\|_2^2, c=(F+N)/2+a+1$ for ℓ_2 -ARD and $f(\mathbf{x})=\|\mathbf{x}\|_1, c=F+N+a+1$ for ℓ_1 -ARD

- MAP optimization ⇒ some relevance weights converge to a small constant and corresponding components are pruned
- Optimizing only over λ leads to the cost function $C(\mathbf{W}, \mathbf{H}) =$

$$\rho D_{\beta}(\mathbf{V}|\mathbf{WH}) + c \sum_{k=1}^{N} \log(f(\mathbf{w}_k) + f(h_k) + b) + \text{cst}$$

• This cost has connections to reweighted ℓ₁ minimization [Candès et al. 2008] and group LASSO [Yuan and Lin 2006]

The Inference Algorithms

- Build auxiliary functions to optimize $C(\mathbf{W}, \mathbf{H}, \lambda)$ over \mathbf{H}
- In the end, the multiplicative updates are

$$\ell_2 - ARD: \qquad h_{kn} = \tilde{h}_{kn} \left(\frac{p_{kn}}{q_{kn} + \tilde{h}_{kn}/(\rho \phi_k)} \right)^{\xi(\beta)}$$
$$\ell_1 - ARD: \qquad h_{kn} = \tilde{h}_{kn} \left(\frac{p_{kn}}{q_{kn} + 1/(\rho \phi_k)} \right)^{\gamma(\beta)}$$

• Updates of ϕ_k proceed as follows

$$\phi_k = \frac{f(\mathbf{w}_k) + f(h_k) + b}{c}$$

- Because we use MM-based algorithms, the cost function $C(\mathbf{W}, \mathbf{H}, \lambda)$ decreases monotonically
- Choose the hyperparameter *b* using the method of moments:

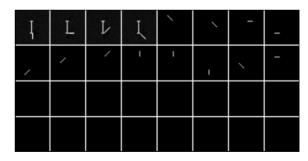
$$\ell_2 - ARD: \qquad \hat{b} = \frac{\pi(a-1)\hat{\mu}_{\mathbf{V}}}{2K}$$
$$\ell_1 - ARD: \qquad \hat{b} = \left(\frac{(a-1)(a-2)\hat{\mu}_{\mathbf{V}}}{K}\right)^{1/2}$$

Noisy Swimmer Dataset

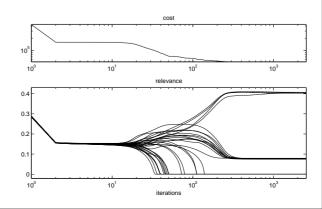
- Synthetic dataset of N=256 images size $F=32\times32=1024$
- Each image represents a swimmer composed of an invariant torso and four limbs. Each limb in four different positions.



- Dictionary learned using one run of $\ell_1\text{-}ARD$ with a=100



• Monotonic decrease in cost and evolution of relevances



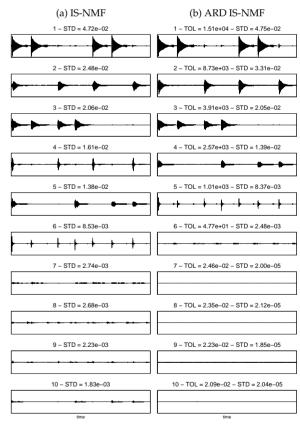
Music Decomposition Example

- Use the Itakura-Saito (IS) divergence ($\beta = 0$) which is suited for audio applications (Févotte et al. 2009)
- Underlies a generative statistical model of superimposed Gaussian components in the squared STFT domain
- Sequence is composed of 4 piano notes, played all at once in the first measure and then played by pairs in all possible combinations in the subsequent measures.

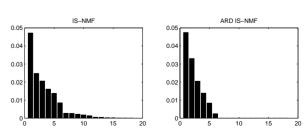


• Given an approximate factorization **WH** of the spectrogram $v_{fn} = |x_{fn}|^2$, the STFT estimate $\hat{c}_{k,fn}$ of component k is

$$\hat{c}_{k,fn} = \frac{w_{fk}h_{kn}}{\sum_{j} w_{fj}h_{jn}} x_{fn}$$



• Histogram shows that 6 components retained by ℓ_1 -ARD



• First four components extract the individual notes and the next two components extract the sound of hammer hitting the strings and the sound produced by the sustain pedal