Canonical Estimation in a Rare Events Regime

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SILO (Oct 2011)

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- $X_1, ..., X_n$ are independent samples from p
- Law of large numbers:

$$\frac{1}{n}\sum_{i=1}^n X_i \stackrel{P}{\to} \mathbb{E}X$$



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Definition (Wagner, Viswanath and Kulkarni, IT-Trans 2011)

We say that $\{(A_n, p_n)\}_{n \in \mathbb{N}}$ is a rare-events source if

$$\frac{\check{c}}{n} \leq p_n(a) \leq \frac{\hat{c}}{n}, \quad \forall a \in \mathcal{A}_n$$

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and $\exists P$ such that $P_n \Rightarrow P$. Note $supp(P) \subseteq \mathcal{C} := [\check{c}, \hat{c}]$.



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Probabilities
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Entropy $H(p_n)$

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■ Can we estimate all reasonable quantities in a universal manner?

Canonical Estimation Problems

Let $\{Y_n\}_{n\in\mathbb{N}}$ be a sequence of real-valued random variables such that

■ There exists continuous $f_n(x)$ that converge to f(x) pointwise on C

$$\mathbb{E}[Y_n] = \int_{\mathcal{C}} f_n(x) \, dP_n(x)$$

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Definition

An estimator $\{\hat{Y}_n : \mathcal{A}_n^n \to \mathbb{R}\}_{n \in \mathbb{N}}$ is consistent if

$$\hat{Y}_n(X_{n,1},\ldots,X_{n,n}) \to \int_{\mathcal{C}} f(x) dP(x)$$

almost surely.



Probabilities

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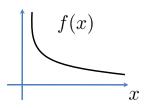
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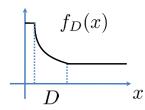
- Problems:
 - Support C isn't known
 - \blacksquare f(x) doesn't have to be bounded everywhere
 - How to get the estimate $\hat{P}_n(x)$?



Estimate the function

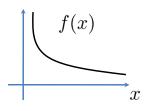
Consider a tapered version of f

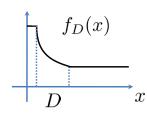




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Lemma (Ohannessian-Tan-Dahleh)

If $C \subset \mathcal{D}$, then

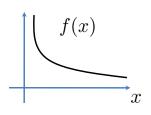
$$\hat{Y}_n := \int_{\mathcal{C}} f_{\mathcal{D}}(x) \, d\hat{P}_n(x)$$

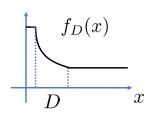
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is a consistent estimator

If $\mathcal C$ is unknown, just let $\mathcal D$ grow gradually with n_{top}



Estimate with rates

Recall the Wasserstein distance

$$d_W(P,Q) = \sup_{h \in Lip(1)} \left| \int_{\mathbb{R}^+} h \, dP - \int_{\mathbb{R}^+} h \, dQ \right|$$

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Lemma (Ohannessian-Tan-Dahleh)

If

$$Lip(f_{D_n}) d_W(\hat{P}_n, P) \rightarrow 0,$$

then,

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How to estimate the shadow P_n ?



Pseudo-Empirical Measure

Good-Turing estimator:

- lacksquare Denote the set of symbols that appear k times as $\mathcal{B}_{n,k}\subset\mathcal{A}_n$
- Denote their probabilities as

$$\gamma_{n,k} = p_n(\mathcal{B}_{n,k}) = \sum_{a \in \mathcal{B}_{n,k}} p_n(a)$$

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■ E.g.: Probability of missing mass $\approx \phi_{n,0}$ [Budianu and Tong 2004]



Strong law of large numbers gives:

Lemma (WVK)

Let the P-Poisson mixture be

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Then, $\|\gamma_n - \lambda^P\|_1 \to 0$ and $\|\phi_n - \lambda^P\|_1 \to 0$ almost surely.

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Theorem (Ohannessian-Tan-Dahleh)

For "most natural" rare event sources, there exist an s > 0 such that

$$n^{\mathbf{s}} \sup_{k \in \mathbb{N}^1} |F_{\phi_n}(k) - F_{\lambda^p}(k)| o 0, \qquad a.s.$$

(Kolmogorov-Smirnov convergence)

Estimation of $\hat{P}_n(x)$ via mixture distribution learning

Theorem

The (pseudo) maximum-likelihood estimator

$$\hat{P}_n^{ML} = \underset{Q}{\operatorname{arg\,min}} \ D(\phi_n || \ Q)$$

is a valid construction, i.e., $\hat{P}_n^{ML} \Rightarrow P$ almost surely.

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Theorem

The minimum distance estimator

$$\hat{P}_n^{MD} = rg \min_{Q} \sup_{k \in \mathbb{N}} \left| F_{\phi_n}(k) - F_{Poi(Q)}(k) \right|$$

is also valid. Furthermore, there exists s>0 such that $n^s d_W(\hat{P}_n,P) \to 0$ almost surely (with some technical conditions).

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Lemma

With $D_n = o(n^s)$,

$$\hat{Y}_n := \int_{\mathbb{R}^+} (-\log x)_{D_n} d\hat{P}_n(x)$$

is consistent

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With $D_n = o(n^{s/2})$,

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Estimating support $C = [\check{c}, \hat{c}]$

- Not quite canonical but close.
- Let Z be the weak limit of $Z_n := np_n(X_n)$ and let $P := Z_*(\mathbb{P})$. Then

$$\hat{\mathbf{c}} = \operatorname{esssup}_{\omega} Z(\omega) = \lim_{q \to \infty} \left[\int_{\mathcal{C}} x^q \, dP(x) \right]^{1/q}$$

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Lemma

With
$$q_n = \frac{\log n}{\log \log n}$$
 and $D_n = o(n^{\frac{s}{2q_n}})$,

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- Future work: Further analysis of convergence rates