

Using Generalized Correlation to effect Variable Selection in Very High Dimensional Problems (Author: Peter HALL)

Xuelong Wang

April 25, 2018

On the Agenda

- 1 Variable Selection Problem
- 2 Solution: Generalized Correlation
- 3 Simulation study
- 4 Conclusion

Model settings and Assumptions

Model setup for variable selection

$$Y_i = \alpha + \beta_1 X_{i1} + \cdots + \beta_p X_{ip} + \text{error}$$

- We assume the model is linear between X and Y
- Variable selection is achieved by shrinking many β 's to zeros
- LASSO is one of the popular and effective approaches

What if it's not Linear between X and Y

- A key assumption of variable selection method (LASSO) is Linearity
- If the model is non-linear, some of the predictors may not be detected by a linear model-based variable selection method

Motivating Example

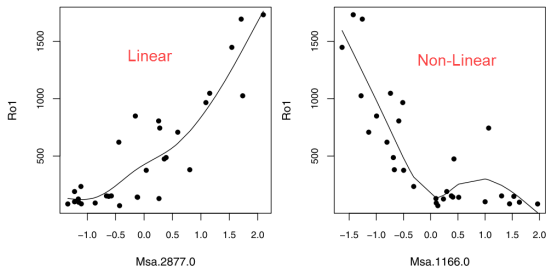


Figure 1. Top two variables with cubic-spline fits for Example 1.

- Microarray data of heart disease
- Ro1: expression level, continuous response
- Msa: different genes, continuous predictor

The Collinearity

- Even the model is perfectly linear, fitting a linear model may conceal importance components of X just because of the collinearity
- It's also called the “masking effect”, which means we only can select part of the important variables
- However, at point of variable selection, we want to be able to select all of them
- Back to the previous heat disease example,
 $\text{cor}(Msa.2877, Msa.1166) = -0.71$. So it's possible that the collinearity prevents us to find the Msa.1166

On the Agenda

- 1 Variable Selection Problem
- 2 Solution: Generalized Correlation
- 3 Simulation study
- 4 Conclusion

Generalized Correlation

Generalized Correlation coefficient between X_{ij} and Y_i

$$\sup_{h \in \mathcal{H}} \frac{\text{cov}\{h(X_{ij}), Y_i\}}{\sqrt{\text{var}\{h(X_{ij})\}, \text{var}(Y_i)}}$$

Estimated by,

$$\sup_{h \in \mathcal{H}} \frac{\sum_i \{h(X_{ij}) - \bar{h}\}(Y_i - \bar{Y})}{\sqrt{\sum_i \{h(X_{ij}) - \bar{h}\}^2 \cdot \sum_i (Y_i - \bar{Y})^2}}$$

- $(X_1, Y_1), \dots, (X_n, Y_n)$ are iid
- X is p -vectors
- Y is scalars
- \mathcal{H} is a vector space of functions
- $\bar{h} = n^{-1} \sum_i h(X_{ij})$

Simplify the computation

Removing $Var(Y_i)$

$$\psi_j = \sup_{h \in \mathcal{H}} \frac{\text{cov}\{h(X_{ij}), Y_i\}}{\sqrt{\text{var}\{h(X_{ij})\}}}$$
$$\hat{\psi}_j = \sup_{h \in \mathcal{H}} \frac{\sum_i \{h(X_{ij}) - \bar{h}\} (Y_i - \bar{Y})}{\sqrt{n \sum_i \{h(X_{ij}) - \bar{h}\}^2}}$$

- Since $Var(Y_i)$ is same for each i , we can remove it without affecting the ranking of the correlation

Simplify the computation Cont.

Theorem 1

Assume \mathcal{H} is a finite-dimensional function space include the constant function, and there exists $h \in \mathcal{H}$ that achieves $\hat{\psi}_j$,

$$\arg \min_{h \in \mathcal{H}} \sum_{i=1}^n \{Y_i - h(X_{ij})\}^2 \subseteq \arg \max_{h \in \mathcal{H}} \hat{\psi}_j$$

The maximizer of $\hat{\psi}_j$ is the solution to least squares problem in \mathcal{H}

Reduction of $\hat{\psi}_j$ in the size of squared error

$$\hat{\phi}_j = \sum_{i=1}^n (Y_i - \hat{Y})^2 - \inf_{h \in \mathcal{H}} \sum_{i=1}^n \{Y_i - h(X_{ij})\}^2$$

Since $\hat{\phi}_j$ keeps the relative relation of $\hat{\psi}_j$, we can use $\hat{\phi}_j$'s for ranking

Correlation Ranking

Some notations

- We order estimator $\hat{\psi}_j$ as $\hat{\psi}_{\hat{j}_1} \geq \cdots \geq \hat{\psi}_{\hat{j}_p}$

$$\hat{j}_1 \succeq \cdots \succeq \hat{j}_p$$

- $j \succeq j'$ means $\hat{\psi}_j \geq \hat{\psi}_{j'}$, we could say j th coefficient of X is at least as much importance as the j' th coefficient
- $r = \hat{r}(j)$ means the rank of the j th coefficient is r , in other words $\hat{j}_r = j$

Estimation of the rank

We could use Bootstrap to assess the empirical rank for each component of X

A $(1 - \alpha)$ level, two-side interval is defined as following:

Interval of rank $[\hat{r}_-(j), \hat{r}_+(j)]$

$$P\{r^*(j) \leq \hat{r}_-(j) | \mathcal{D}\} \approx P\{r^*(j) \geq \hat{r}_+(j) | \mathcal{D}\} \approx \frac{\alpha}{2}$$

- $r^*(j)$ the bootstrap version estimators of $r(j)$
- The approximation is used because of the discreteness of ranks
- Small value of $r^*(j)$ indicates large influence on Y
- In order to select variables, we could sort $r^*(j)$ or $\hat{r}_+(j)$ and set a cut off value p

On the Agenda

- 1 Variable Selection Problem
- 2 Solution: Generalized Correlation
- 3 Simulation study
- 4 Conclusion

heart disease data

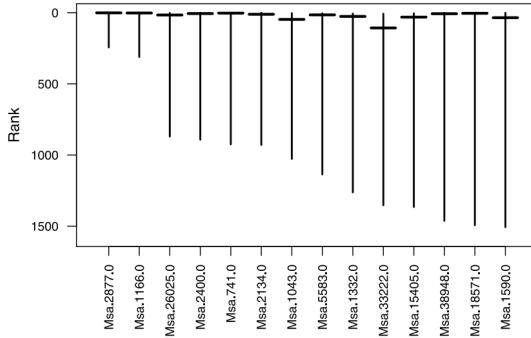


Figure 2. Variables ordered by \hat{r}_+ for Example 1.

- $n = 30$, $p = 6319$
- bootstrap sample is 400
- $\alpha = 0.02$ cutoff value is $p/4$
- ranking is based on $\hat{r}_+(j)$

Simulation study on non-linear model

simulation setup

$$Y_i = W_i^2 - 1 + \epsilon_i$$

- $W_i \sim \text{Unif}[-2, 2]$
- $X_{i1} = W_i + \delta_i$ (errors-in-variables type)
- $X_{i2}, \dots, X_{i5000} \stackrel{iid}{\sim} N(0, 1)$
- $\delta_i, \epsilon_i \stackrel{iid}{\sim} N(0, 3/4)$
- $n = 200, \alpha = 0.02, n_{bootstrap} = 500$

Simulation study on non-linear model Cont.

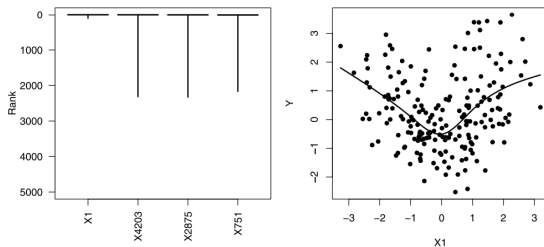


Figure 6. Top variables by \hat{r}_+ for Example 5 and the cubic-spline fit for X_1 .

- the cutoff value is $p/2$
- Conventional correlation fails to select X_{i1} as influential variable
- Generalized correlation method is able to select the X_{i1} with only 3 false positive variables

On the Agenda

- 1 Variable Selection Problem
- 2 Solution: Generalized Correlation
- 3 Simulation study
- 4 Conclusion

Potential applications

- It can be used as a “massive dimension reduction” method
- It should be a more effective variable selection method for (Generalized) Additive Model

Further topics

- Unbiasedness and Consistency of the selected Variables
- The choice of cutoff value

Thank you

Reference

Peter HALL, Hugh MILLER. 2009. "Using Generalized Correlation to Effect Variable Selection in Very High Dimensional Problems."