Levi Waldron

Learning objectives and outline

Review

Overdispersion

Zero Inflation

Session 5: loglinear regression part 2

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CUNY SPH Biostatistics 2

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Learning objectives and outline

Review

Overdispersion

Zero Inflation

Learning objectives and outline

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Learning objectives and outline

Review

Overdispersion

Zero Inflation

Learning objectives

- 1 Define and identify over-dispersion in count data
- Define the negative binomial (NB) distribution and identify applications for it
- 3 Define zero-inflated count models
 - 4 Fit and interpret Poisson and NB, with and without zero inflation

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Learning objectives and outline

Review

Overdispersion

Zero Inflation

Outline

- Review of log-linear Poisson glm
- 2 Review of diagnostics and interpretation of coefficients
 - 3 Over-dispersion
 - Negative Binomial distribution
- 4 Zero-inflated models
- Vittinghoff section 8.1-8.3

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Review

Overdispersion

Zero Inflation

Review

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Review

Overdispersion

Zero Inflation

Components of GLM

- Random component specifies the conditional distribution for the response variable - it doesn't have to be normal but can be any distribution that belongs to the "exponential" family of distributions
- Systematic component specifies linear function of predictors (linear predictor)
- Link [denoted by g(.)] specifies the relationship between the expected value of the random component and the systematic component, can be linear or nonlinear

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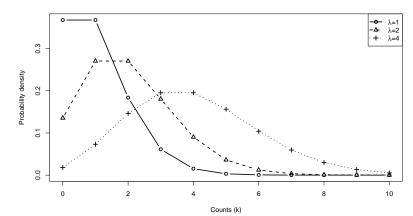
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Overdispersion

Zero Inflation

Motivating example: Choice of Distribution

- Count data are often modeled as Poisson distributed:
 - mean λ is greater than 0
 - variance is also λ
 - Probability density $P(k,\lambda) = \frac{\lambda^k}{k!} e^{-\lambda}$



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Review

Overdispersion

Zero Inflation

Poisson model: the GLM

The **systematic part** of the GLM is:

$$log(\lambda_i) = \beta_0 + \beta_1 RACE_i + \beta_2 TRT_i + \beta_3 ALCH_i + \beta_4 DRUG_i$$

Or alternatively:

$$\lambda_{i} = \exp(\beta_{0} + \beta_{1} RACE_{i} + \beta_{2} TRT_{i} + \beta_{3} ALCH_{i} + \beta_{4} DRUG_{i})$$

The **random part** is (Recall the λ_i is both the mean and variance of a Poisson distribution):

$$y_i \sim Poisson(\lambda_i)$$

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Review Over-

dispersion

Zero Inflation

##

Example: Risky Drug Use Behavior

- Load the "needle sharing" dataset
- Outcome is # times the drug user shared a syringe in the past month (shared syr)
 - Predictors: sex, ethn, homeless
 - filtered to only sex "M" or "F", ethn "White", "AA", "Hispanic"

```
## cols(
##
     id = col double(),
     sex = col character(),
##
     ethn = col character(),
##
##
     age = col double(),
##
     dprsn dx = col double(),
##
     sexabuse = col_double(),
##
     shared_syr = col_double(),
     hivstat = col double(),
##
```

-- Column specification

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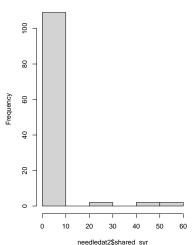
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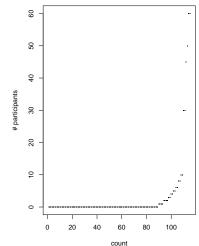
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Example: Risky Drug Use Behavior

Exploratory plots





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Review

Overdispersion

Zero Inflation

Risky Drug Use Behavior: fitting a Poisson model

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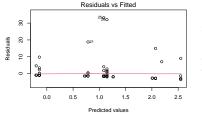
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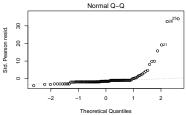
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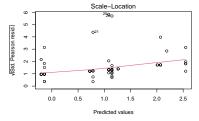
Overdispersion

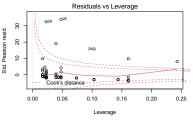
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Risky Drug Use Behavior: residuals plots









* Poisson model is definitely not a good fit.

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Review

Overdispersion

Zero Inflation

Over-dispersion

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Review

Overdispersion

Zero Inflation

When the Poisson model doesn't fit

- inference from log-linear models is sensitive to assumptions on the distribution of residuals (e.g. Poisson)
- In the Poisson distribution, the variance is equal to the mean.
- *i.e.* if subjects with a particular pattern of covariates have a mean of 4 visits/yr, then variance is also 4 and the standard deviation is 2 visits / yr.
- The Poisson distribution often fails when the variance exceeds the mean
 - You can check this assumption
- Can use alternative random distributions:
 - Negative binomial distribution
- Can introduce zero-inflation

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Review

Overdispersion

Zero Inflation

Negative binomial distribution

- The binomial distribution is the number of successes in n trials:
 - Roll a die ten times, how many times do you see a 6?
- The negative binomial distribution is the number of successes it takes to observe r failures:
 - How many times do you have to roll the die to see a 6 ten times?
 - Note that the number of rolls is no longer fixed.
 - In this example, p=5/6 and a 6 is a "failure"

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Learning objectives and outline
Review
Over- dispersion
Zero Inflation

Session 5: loglinear

regression part 2

Negative binomial GLM

One way to parametrize a NB model is with a **systematic part** equivalent to the Poisson model:

$$log(\lambda_i) = \beta_0 + \beta_1 RACE_i + \beta_2 TRT_i + \beta_3 ALCH_i + \beta_4 DRUG_i$$
Or:

 $\lambda_i = exp(\beta_0 + \beta_1 RACE_i + \beta_2 TRT_i + \beta_3 ALCH_i + \beta_4 DRUG_i)$

And a random part:

$$y_i \sim \mathsf{NB}(\lambda_i, heta)$$

- θ is a **dispersion parameter** that is estimated
- When 0 0 it is a mixed at the Deisson model
- When θ = 0 it is equivalent to Poisson model
 MASS::glm.nb() uses this parametrization, dnbinom()
- does not
 The Poisson model can be considered **nested** within the Negative Binomial model

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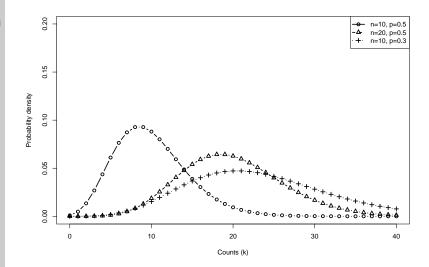
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Review

Overdispersion

Zero Inflation

Negative Binomial Random Distribution



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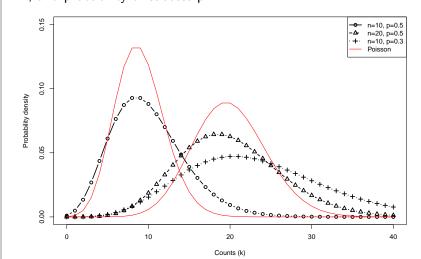
Review

Overdispersion

Zero Inflation

Compare Poisson vs. Negative Binomial

Negative Binomial Distribution has two parameters: # of trials n, and probability of success p



```
Risky drug behavior: Negative
 regression
  part 2
                                 Binomial Regression
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          library(MASS)
Learning
          fit.negbin <- glm.nb(shared_syr ~ sex + ethn + homele
objectives and
                                  data = needledat2)
Review
           summary(fit.negbin)
Over-
dispersion
Zero Inflation
          ##
           ## Call:
           ## glm.nb(formula = shared_syr ~ sex + ethn + homeles
                  init.theta = 0.07743871374, link = log)
           ##
           ##
           ## Deviance Residuals:
           ##
                  Min
                             10
                                 Median
                                                 30
                                                          Max
           ## -0.8801 -0.7787 -0.6895 -0.5748
                                                       1.5675
          ##
          ## Coefficients:
                            Estimate Std. Error z value Pr(>|z|)
           ##
```

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Learning objectives an outline
Review
Over- dispersion
Zero Inflation

Session 5: loglinear

regression part 2

Recall the Deviance:

the difference in df of the two models:

'log Lik.' -147.1277 (df=6) (ll.pois <- logLik(fit.pois))</pre>

'log Lik.' -730.0133 (df=5)

'log Lik.' 1.675949e-255 (df=6)

pchisq(2 * (11.negbin - 11.pois), df=1, lower.tail=FA

(ll.negbin <- logLik(fit.negbin))</pre>

Likelihood ratio test $\Delta(D) = -2 * \Delta(\log \text{ likelihood})$ And recall the difference in deviance under H_0 (no improvement in fit) is chi-square distributed, with df equal to

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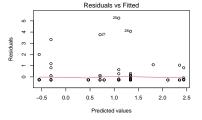
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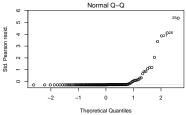
Review

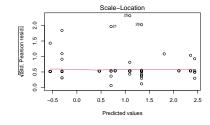
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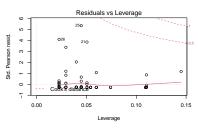
Zero Inflation

Risky Drug Use Behavior: NB regression residuals plots









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Learning objectives and outline

Review

Overdispersion

Zero Inflation

Zero Inflation

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Learning objectives and outline

Review

Overdispersion

Zero Inflation

Zero inflated "two-step" models

Step 1: logistic model to determine whether count is zero or Poisson/NB

Step 2: Poisson or NB regression distribution for y_i not set to zero by 1.

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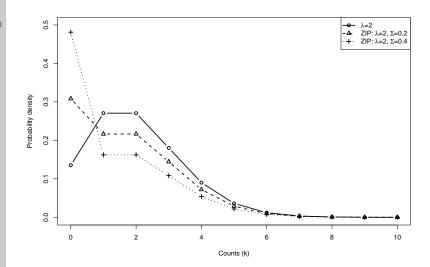
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Review

Overdispersion

Zero Inflation

Poisson Distribution with Zero Inflation



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Learning objectives and outline

Review

Overdispersion

Zero Inflation

Risky drug behavior: Zero-inflated Poisson regression

```
Learning
objectives and
outline
          ##
Review
          ## Call:
Over-
dispersion
          ## zeroinfl(formula = shared_syr ~ sex + ethn + homel
Zero Inflation
                 dist = "poisson")
          ##
          ##
          ## Pearson residuals:
          ##
                 Min
                           10 Median
                                            30
                                                    Max
          ## -1.0761 -0.5784 -0.4030 -0.3341 10.6835
          ##
          ## Count model coefficients (poisson with log link):
                           Estimate Std. Error z value Pr(>|z|)
          ##
          ## (Intercept)
                                         0.1796 17.908 < 2e-16
                             3.2168
          ## sexM
                           -1.4725
                                         0.1442 -10.212 < 2e-16
          ## ethnHispanic -0.1524
                                        0.1576 -0.968 0.333244
```

summary(fit.ZIpois)

Zero-inflated Poisson

regression - the model

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regression part 2

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Review

Overdispersion

Zero Inflation

Risky drug behavior: Zero-inflated Negative Binomial regression

- *NOTE*: zero-inflation model can include any of your variables as predictors
- WARNING Default in zerinf1() function is to use all variables as predictors in logistic model

```
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loglinear
regression
part 2
```

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Learning ##

outline

Overdispersion

Zero Inflation

Zero-inflated Negative

summary(fit.ZInegbin)

```
##
## Call:
## zeroinfl(formula = shared syr ~ sex + ethn + homeless, data = needledat2,
      dist = "negbin")
##
##
## Pearson residuals:
      Min
               10 Median
                              30
                                    Max
## -0.5402 -0.3255 -0.2714 -0.1926 5.1496
##
## Count model coefficients (negbin with log link):
##
              Estimate Std. Error z value Pr(>|z|)
## (Intercept)
                2.8410
                          1.1845 2.399 0.01646 *
## sexM
               -2.2282 0.9351 -2.383 0.01718 *
## ethnHispanic -0.4123 0.9831 -0.419 0.67492
## ethnWhite -0.4299 0.8648 -0.497 0.61908
## homelessyes 1.9460 0.7103 2.740 0.00615 **
## Log(theta) -1.1971
                           0.5159 -2.320 0.02032 *
##
## Zero-inflation model coefficients (binomial with logit link):
              Estimate Std. Error z value Pr(>|z|)
## (Intercept)
                1 6867
                           0.8465
                                  1 993
                                           0.0463 *
## seyM
                           0.8016 -1.238 0.2159
                -0.9920
                                          0.9627
## ethnHispanic -13.1868
                         281.9134 -0.047
## ethnWhite
              -0.7455
                           0.7304 -1.021
                                           0.3074
## homelessves 0.3554
                          0.7397 0.480
                                           0.6309
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
##
## Theta = 0.3021
## Number of iterations in BFGS optimization: 24
```

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Learning objectives and

Review

Overdispersion

Zero Inflation

Zero-inflated Negative Binomial regression - simplified ZI model

- Model is much more interpretable if the exposure of interest is *not* included in the zero-inflation model.
 E.g. with HIV status as the only predictor in zero inflation.
- E.g. with HIV status as the only predictor in zero-inflation model:

```
Session 5:
loglinear
regression
part 2
```

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Learning objectives and outline

Review

Overdispersion

Zero Inflation

Zero-inflated Negative

summary(fit.ZInb2)

```
##
## Call:
## zeroinfl(formula = shared_syr ~ sex + ethn + homeless + hiv | hiv, data = needledat2,
##
      dist = "negbin")
##
## Pearson residuals:
      Min
               10 Median
                              30
                                     Max
## -0 4299 -0 3646 -0 3559 -0 3299 6 3053
##
## Count model coefficients (negbin with log link):
               Estimate Std. Error z value Pr(>|z|)
##
## (Intercept)
                 3.6685
                           0.9470 3.874 0.000107 ***
## sexM
                -1.7648 0.6205 -2.844 0.004454 **
## ethnHispanic -1.5807 0.7446 -2.123 0.033769 *
## ethnWhite -1.1267 0.6924 -1.627 0.103687
## homelessyes 1.0313 0.5693 1.812 0.070028 .
## hivpositive -1.0820 1.0167 -1.064 0.287235
## hivves
           2.3724
                           0.7829 3.030 0.002444 **
## Log(theta)
                0.1395
                           0.4647 0.300 0.764009
## Zero-inflation model coefficients (binomial with logit link):
               Estimate Std. Error z value Pr(>|z|)
## (Intercept) 1.2163
                           0.2851
                                    4.265
                                            2e-05 ***
## hivpositive -0.3493
                           0.9389 -0.372
                                            0.710
## hivves
               -17.9654 3065.6271 -0.006
                                            0.995
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
##
## Theta = 1.1497
## Number of iterations in BFGS optimization: 12
```

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Review

Overdispersion

Zero Inflation

Intercept-only zero-inflation

```
fit.ZInb3 <- zeroinfl(shared svr~sex+ethn+homeless|1.
                      dist="negbin", data=needledat2)
summary(fit.ZInb3)
##
## Call:
## zeroinfl(formula = shared_syr ~ sex + ethn + homeless | 1, data = needledat2,
      dist = "negbin")
##
## Pearson residuals:
      Min
               10 Median
                              30
                                     Max
## -0.3159 -0.3123 -0.3040 -0.2953 5.2940
##
## Count model coefficients (negbin with log link):
              Estimate Std. Error z value Pr(>|z|)
##
## (Intercept) 2.08542 1.42671 1.462 0.1438
## seyM
              -1.43809 0.89189 -1.612 0.1069
## ethnHispanic 0.48130 1.16642 0.413 0.6799
## ethnWhite -0.07418 0.81066 -0.092 0.9271
## homelessves 1.62076 0.67706 2.394 0.0167 *
## Log(theta) -1.12538 0.89372 -1.259 0.2080
##
## Zero-inflation model coefficients (binomial with logit link):
              Estimate Std. Error z value Pr(>|z|)
## (Intercept) 0.5211
                          0.7600 0.686
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
##
## Theta = 0.3245
## Number of iterations in BFGS optimization: 13
## Log-likelihood: -146.8 on 7 Df
```

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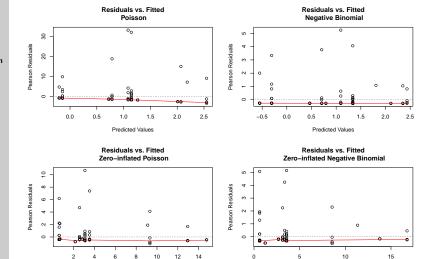
Review

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Zero Inflation

Residuals vs. fitted values

I invisibly define functions plotpanel1 and plotpanel2 that will work for all types of models (see lab). These use Pearson residuals.



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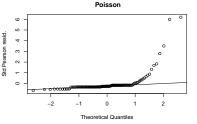
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Review

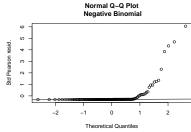
Overdispersion

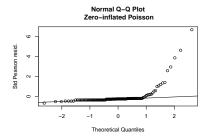
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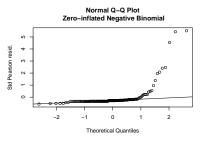
Quantile-quantile plots for residuals



Normal Q-Q Plot







still over-dispersed - ideas?

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Review

Overdispersion

Zero Inflation

(3) (4)

> (5) sexM -0.925***

Dependent variable:

shared syr

Poisson

negative

binomial count data (1) (2)

zero-inflated

Inference from the different

models

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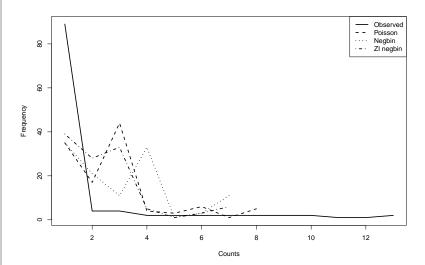
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Review

Overdispersion

Zero Inflation

Example of plotting observed and predicted counts



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Review

Overdispersion

Zero Inflation

Resources for R (and SAS)

- Short, practical tutrorials on regression in R and SAS from UCLA at http://www.ats.ucla.edu/stat/:
 - Poisson Regression: http://www.ats.ucla.edu/stat/r/dae/poissonreg.htm
 - Negative Binomial: http://www.ats.ucla.edu/stat/r/dae/nbreg.htm
 - Zero-inflated Poisson: http://www.ats.ucla.edu/stat/r/dae/zipoisson.htm
 - Zero-inflated Negative Binomial: http://www.ats.ucla.edu/stat/r/dae/zinbreg.htm