# Probabilistic Robotics Course Least Squares and Uncertainty

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## Maximum Likelihood Estimation

Using

 $\mathbf{x}^* = \underset{\mathbf{x}}{\operatorname{argmax}} p(\mathbf{x}|\mathbf{z})$ 

Bayes' Rule

$$p(\mathbf{x}|\mathbf{z}) = \frac{p(\mathbf{z}|\mathbf{x})p(\mathbf{x})}{p(\mathbf{z})}$$

$$\propto p(\mathbf{z}|\mathbf{x})$$

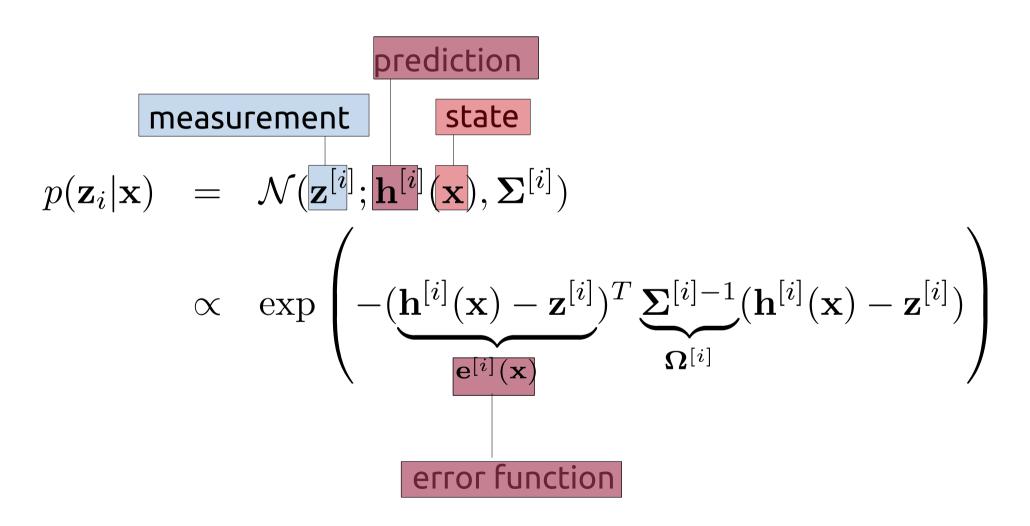
Independence,

$$= \prod_{i} p(\mathbf{z}^{[i]}|\mathbf{x})$$

We can further simplify the task

## **Gaussian Assumption**

Measurements affected by Gaussian noise



## **Gaussian Assumption**

Through Gaussian assumption

- Maximization becomes minimization
- Product turns into sum

$$\mathbf{x}^* = \underset{x}{\operatorname{argmax}} \prod_{i} p(\mathbf{z}^{[i]}|\mathbf{x})$$

$$= \underset{x}{\operatorname{argmax}} \prod_{i} \exp(-\mathbf{e}^{[i]}(\mathbf{x})^T \mathbf{\Omega}^{[i]} \mathbf{e}^{[i]}(\mathbf{x}))$$

$$= \underset{x}{\operatorname{argmin}} \sum_{i} \mathbf{e}^{[i]}(\mathbf{x})^T \mathbf{\Omega}^{[i]} \mathbf{e}^{[i]}(\mathbf{x})$$

## **Goal of the Lesson**

Use our skills with Gaussian distribution to derive an estimate of the uncertainty of the solution, around the optimum reported by Gauss-Newton

- Chain rule to get p(x,z)
- Conditioning on z

## **Around the Optimum...**

The measurement functions can be approximated by their Taylor expansion

$$\mathbf{z}^{[i]^*} = \mathbf{h}^{[i]}(\mathbf{x}^*)$$
 $\mathbf{h}^{[i]}(\mathbf{\Delta}\mathbf{x} + \mathbf{x}^*) \simeq \mathbf{z}^{[i]^*} + \underbrace{\frac{\partial \mathbf{h}^{[i]}(\mathbf{x})}{\partial \mathbf{x}}\Big|_{\mathbf{x} = \mathbf{x}^*}}_{\mathbf{J}^{[i]}} \mathbf{\Delta}\mathbf{x}$ 
 $p(\mathbf{z}^{[i]}|\mathbf{\Delta}\mathbf{x} + \mathbf{x}^*) \sim \mathcal{N}(\mathbf{J}^{[i]}\mathbf{\Delta}\mathbf{x} + \mathbf{z}^{[i]^*}, \mathbf{\Omega}^{[i]^{-1}})$ 

# Consequences of Taylor

The conditional over **all** measurements is again a multivariate Gaussian

$$p(\mathbf{z}|\Delta\mathbf{x} + \mathbf{x}^*) \sim \mathcal{N}(\mu_{\mathbf{z}}, \mathbf{\Omega}_{\mathbf{z}}^{-1})$$

$$\mu_{\mathbf{z}} = \begin{pmatrix} \mathbf{J}^{[1]} \Delta\mathbf{x} + \mathbf{z}^{[1]*} \\ \mathbf{J}^{[2]} \Delta\mathbf{x} + \mathbf{z}^{[2]*} \\ \mathbf{J}^{[K]} \Delta\mathbf{x} + \mathbf{z}^{[K]*} \end{pmatrix} = \begin{pmatrix} \mathbf{J}^{[1]} \\ \mathbf{J}^{[2]} \\ \vdots \\ \mathbf{J}^{[K]} \end{pmatrix} \Delta\mathbf{x} + \begin{pmatrix} \mathbf{z}^{[1]*} \\ \mathbf{z}^{[2]*} \\ \vdots \\ \mathbf{z}^{[K]*} \end{pmatrix}$$

$$\mathbf{\Omega}_{\mathbf{z}} = \begin{pmatrix} \mathbf{\Omega}^{[1]} \\ \mathbf{\Omega}^{[2]} \\ \vdots \\ \mathbf{\Omega}^{[K]} \end{pmatrix}$$

## **Chain Rule**

#### We know

$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x}; \mathbf{x}^*, \mathbf{\Sigma}_{\mathbf{x}}).$$

$$p(\mathbf{z}|\mathbf{x}) = \mathcal{N}(\mathbf{z}; \mathbf{J}(\mathbf{z}, \mathbf{J}(\mathbf{z}, \mathbf{z}^*) + \mathbf{z}^*, \mathbf{\Sigma}_{\mathbf{z}})$$

### We want to compute

$$p(\mathbf{x}, \mathbf{z}) = \mathcal{N}(\mu_{\mathbf{x}, \mathbf{z}}, \Sigma_{\mathbf{x}, \mathbf{z}})$$

## The parameters are

$$egin{array}{lll} \mu_{\mathbf{x},\mathbf{z}} &=& \left(egin{array}{c} \mathbf{x}^* \ \mathbf{z}^* \end{array}
ight) \ & \Omega_{\mathbf{x},\mathbf{z}} &=& \left(egin{array}{c} \mathbf{J}^T \Omega_{\mathbf{z}} \mathbf{J} + \Omega_{\mathbf{x}} & -\mathbf{J}^T \Omega_{\mathbf{z}} \ -\Omega_{\mathbf{z}} \mathbf{J}^T & \Omega_{\mathbf{z}} \end{array}
ight) \end{array}$$

## **Applying Chain Rule**

$$p(\mathbf{\Delta}\mathbf{x}) \sim \mathcal{N}(\mathbf{0}, \mathbf{\Sigma}_{\mathbf{x}})$$
$$\mathbf{\Sigma}_{\mathbf{x}} = \infty \Rightarrow \mathbf{\Omega}_{\mathbf{x}} = \mathbf{0}$$
$$p(\mathbf{\Delta}\mathbf{x}, \mathbf{z}) \sim \mathcal{N}((0, \mathbf{z}^*)^T, \mathbf{\Omega}_{\mathbf{x}, \mathbf{z}}^{-1})$$

Instead of working in x, we work on Delta x, that has the same covariance but it is centered in 0

We know nothing a priori about x, so the information matrix is set to 0

$$egin{aligned} oldsymbol{\Omega}_{\mathbf{x},\mathbf{z}} &= \left(egin{array}{ccc} \mathbf{J}^T oldsymbol{\Omega}_{\mathbf{z}} \mathbf{J} & -\mathbf{J}^T oldsymbol{\Omega}_{\mathbf{z}} \ -oldsymbol{\Omega}_{\mathbf{z}} \mathbf{J}^T & oldsymbol{\Omega}_{\mathbf{z}} \end{array}
ight) \ &= \left(egin{array}{ccc} oldsymbol{\Omega}_{\mathbf{x}\mathbf{x}} & oldsymbol{\Omega}_{\mathbf{x}\mathbf{z}} \ oldsymbol{\Omega}_{\mathbf{z}\mathbf{x}} & oldsymbol{\Omega}_{\mathbf{z}\mathbf{z}} \end{array}
ight) \end{aligned}$$

This is the joint information matrix of measurements and states

## **Applying Chain Rule**

$$\begin{split} & \boldsymbol{\Omega}_{\mathbf{x}\mathbf{x}} = \mathbf{J}^T \boldsymbol{\Omega}_{\mathbf{z}} \mathbf{J} \\ & = \left( \begin{array}{ccc} \mathbf{J}^{[1]T} & \mathbf{J}^{[2]T} & \cdots & \mathbf{J}^{[K]T} \end{array} \right) \left( \begin{array}{c} \boldsymbol{\Omega}^{[1]} & \\ \boldsymbol{\Omega}^{[2]} & \\ & \ddots & \\ \boldsymbol{\Omega}^{[K]} \end{array} \right) \left( \begin{array}{c} \mathbf{J}^{[1]} \\ \mathbf{J}^{[2]} \\ \vdots \\ \mathbf{J}^{[K]} \end{array} \right) \\ & = \mathbf{J}^{[1]T} \boldsymbol{\Omega}^{[1]} \mathbf{J}^{[1]} + \cdots + \underbrace{\mathbf{J}^{[k]T} \boldsymbol{\Omega}^{[k]} \mathbf{J}^{[k]}}_{\mathbf{H}^{[k]}} + \cdots + \mathbf{J}^{[k]T} \boldsymbol{\Omega}^{[k]} \mathbf{J}^{[K]} \\ & = \sum_{k=1}^{K} \mathbf{H}^{[k]} \\ & = \mathbf{H} \end{split}$$

# Conditioning

Let  $(\mathbf{x}^T \mathbf{z}^T)$  be a Gaussian random variable such that

$$\left(egin{array}{c} \mathbf{x} \ \mathbf{z} \end{array}
ight) \sim \mathcal{N}\left(\left(egin{array}{c} \mu_x \ \mu_z \end{array}
ight), \mathbf{\Sigma_{x,z}}
ight)$$

The conditional  $p(\mathbf{x} \mid \mathbf{z})$ 

is Gaussian with parameters

$$N(x; \mu_{\mathbf{x}|\mathbf{z}}, \Sigma_{\mathbf{x}|\mathbf{z}})$$

The farther is the measurement from the prediction, the more the mean changes

$$\mu_{\mathbf{x}|\mathbf{z}} = \mu_{\mathbf{x}} + \mathbf{\Sigma}_{\mathbf{x}\mathbf{z}} \mathbf{\Sigma}_{\mathbf{z}\mathbf{z}}^{-1} (\mathbf{z} - \mu_{\mathbf{z}})$$
$$\mathbf{\Sigma}_{\mathbf{x}|\mathbf{z}} = \mathbf{\Sigma}_{\mathbf{x}\mathbf{x}} - \mathbf{\Sigma}_{\mathbf{x}\mathbf{z}} \mathbf{\Sigma}_{\mathbf{z}\mathbf{z}}^{-1} \mathbf{\Sigma}_{\mathbf{z}\mathbf{x}}$$

## Conditioning

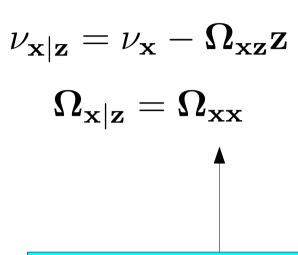
Let  $(\mathbf{x}^T \mathbf{z}^T)$  be a Gaussian random variable such that

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The conditional  $p(\mathbf{x} \mid \mathbf{z})$ 

is Gaussian with parameters

$$p(\mathbf{x}) = \mathcal{N}^{-1}(\mathbf{x}; \nu_{\mathbf{x}|\mathbf{z}}, \mathbf{\Omega}_{\mathbf{x}|\mathbf{z}})$$



The information of the conditional is just the xx block!

# Concluding

$$\mu_{\mathbf{x}|\mathbf{z}} = \mathbf{x}^* + \mathbf{\Sigma}_{\mathbf{x}\mathbf{z}}\mathbf{\Sigma}_{\mathbf{z}\mathbf{z}}^{-1}(\mathbf{z} - \mathbf{z}^*)$$

The farther is the measurement from the prediction, the more the mean changes

$$\mathbf{z} \simeq \mathbf{z}^*$$

If least squares converged well, the prediction at the optimum is close to the measurements

$$\mu_{\mathbf{x}|\mathbf{z}} \simeq \mathbf{x}^*$$

The mean of the Gaussian approximation of the solution is the optimum found by LS

$$egin{aligned} oldsymbol{\Sigma}_{\mathbf{x}|\mathbf{z}} &= oldsymbol{\Omega}_{\mathbf{x}\mathbf{x}}^{-1} \ &\simeq \mathbf{H}^{-1} \end{aligned}$$

The covariance matrix is the inverse of the H matrix