Probabilistic Robotics Course

Unscented Transform

$$x^{(0)} = \frac{1}{\sqrt{2}} [0, 2]^{\top} \quad x^{(1)} = -\frac{1}{\sqrt{2}} [\sqrt{3}, 1]^{\top} \quad x^{(2)} = -\frac{1}{\sqrt{2}} [-\sqrt{3}, 1]^{\top}$$

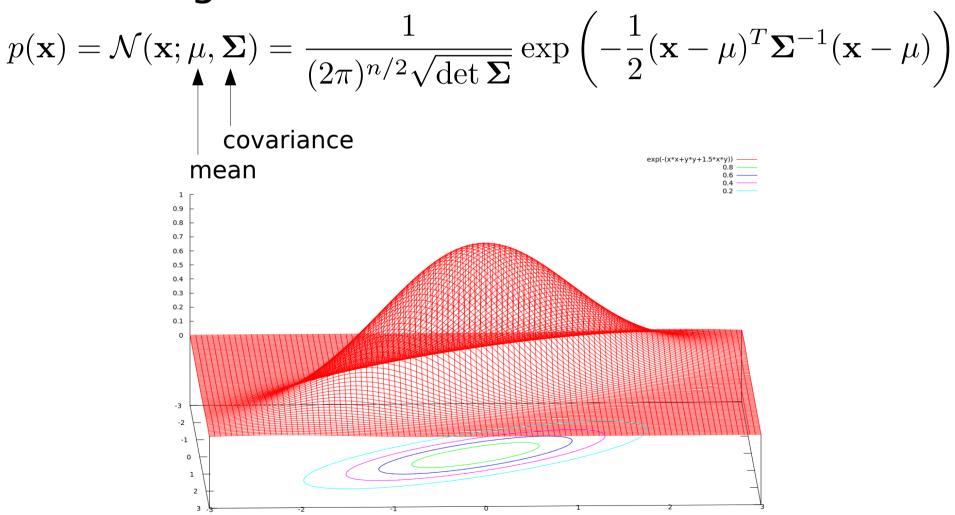
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Gaussian

The pdf of a Gaussian distribution has the following form:

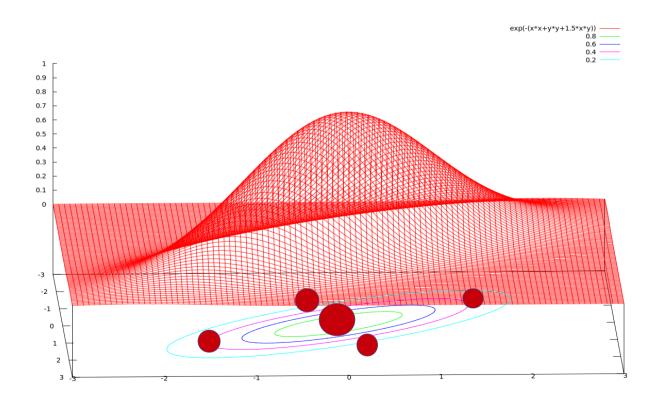




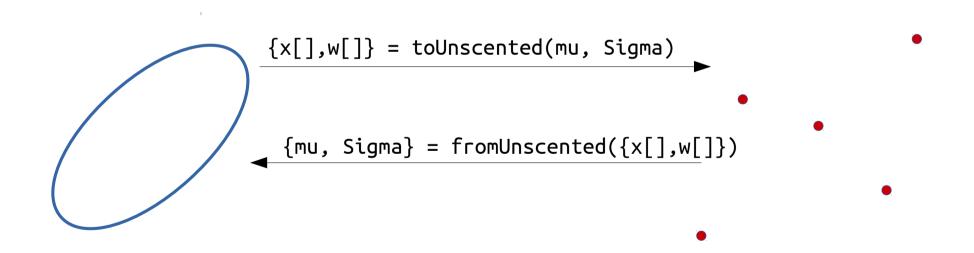
Unscented Gaussian

We can represent the Gaussian as a set of

- control points
- weights



The location of the samples and the weight should be computed such that the transformation can be inverted!



Parameters from Sigma **Points**

Each sigma point is characterized by

$$w_m^{(i)} \in \Re^+$$

a position $\mathbf{x}^{(i)}\in\Omega$ a weight for the mean $w_m^{(i)}\in\Re^+$ $w_c^{(i)}\in\Re^+$ a weight for the covariance $w_c^{(i)}\in\Re^+$

We can imagine the sigma points as weighed control points that control the shape of the Gaussian.

We can reconstruct the parameters from the sigma points as follows:

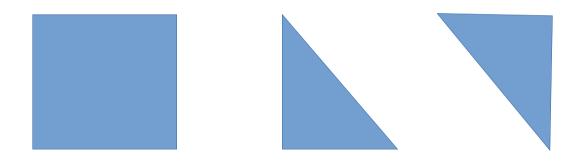
$$\mu = \sum w_m^{(i)} \mathbf{x}^{(i)}$$

$$\mathbf{\Sigma} = \sum w_c^{(i)} (\mathbf{x}^{(i)} - \mu) (\mathbf{x}^{(i)} - \mu)^T$$

Cholesky Decomposition

Can be seen as the square root of a symmetric positive definite matrix:

$$\mathbf{A} = \mathbf{L}\mathbf{L}^T$$



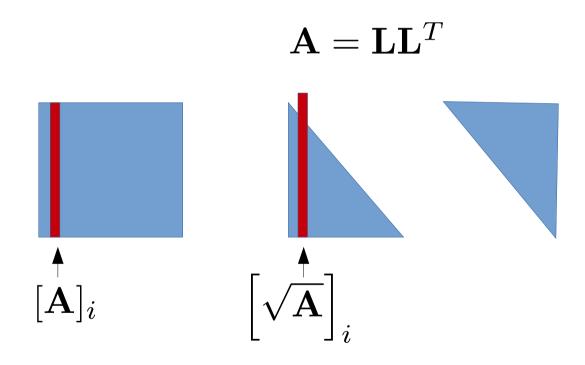
We call ${f L}$ the square root of ${f A}$

$$\mathbf{L} = \sqrt{\mathbf{A}}$$

Computed by using a variant of Gaussian elimination

Cholesky Decomposition

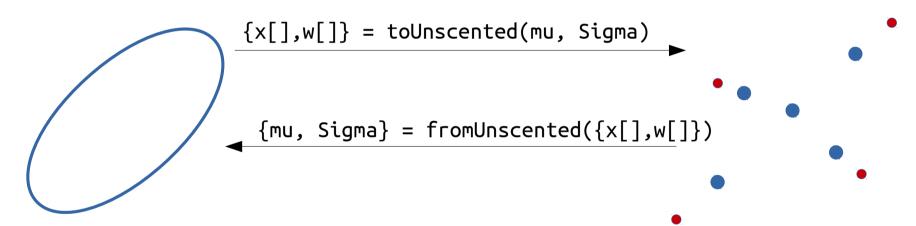
Can be seen as the square root of a symmetric positive definite matrix:



we will refer to the i-th column of the square root of a matrix as $[A]_i$.

Exercise: implement the Cholesky decomposition in C/C++, no libraries allowed

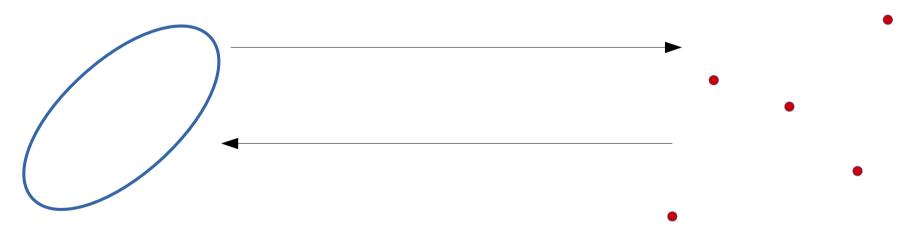
Different combination of weights/sample locations might allow to go forth and back from the representations, since in the sigma points we have more DoF than in the mean and covariance:



We can control these additional degrees of freedom through some "parameters":

$$\kappa \geq 0$$
 $\alpha \in (0,1]$) Influences how far the points are from the mean $\beta = 2$

Different combination of weights/sample locations might allow to go forth and back from the representations, since in the sigma points we have more DoF than in the mean and covariance:



Use these:

$$\kappa = 0$$

$$\alpha = 10^{-3}$$

$$\beta = 2$$

Let n be the dimension, we will have 2n+1 sigma points

Compute the following quantities

$$\lambda = \alpha^{2}(n+\kappa) - n$$

$$\mathbf{L} = \sqrt{(n+\lambda)\mathbf{A}}$$

Compute the positions of the sigma points

$$\mathbf{x}^{(0)} = \mu$$

$$\mathbf{x}^{(i)} = \mu + [\mathbf{L}]_i \text{ for } i \in [1..n]$$

$$\mathbf{x}^{(i)} = \mu - [\mathbf{L}]_{n-i} \text{ for } i \in [n+1..2n]$$

Let n be the dimension, we will have 2n+1 sigma points

Compute the weights:

$$w_m^{(0)} = \frac{\lambda}{n+\lambda}
 w_c^{(0)} = w_m^{(0)} + (1 - \alpha^2 + \beta)
 w_c^{(i)} = w_m^{(i)} = \frac{1}{2(n+\lambda)}$$

Exercise:

Implement both functions for an N dimensional case and verify that the implementation works

Unscented Marginalization

Let $\mathbf{x}^T = (\mathbf{x}_a^T \ \mathbf{x}_b^T)$ be a random variable represented by an UT:

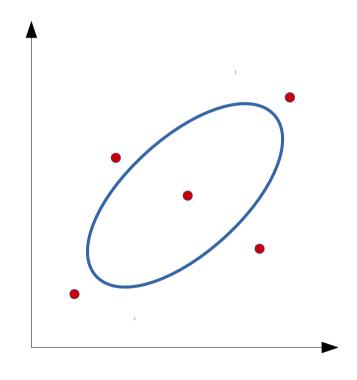
$$\mathbf{x} \sim \mathcal{UT}(\mathbf{x}; \mathbf{x}^{(i)}, w_m^{(i)}, w_c^{(i)})$$

The marginal

$$p(\mathbf{x}_a) = \int_{\mathbf{x}_b} p(\mathbf{x}_a, \mathbf{x}_b) d\mathbf{x}_b$$

is approximated by:

$$\mathbf{x}_a \sim \mathcal{UT}(\mathbf{x}_a; \mathbf{x}_a^{(i)}, w_m^{(i)}, w_c^{(i)})$$

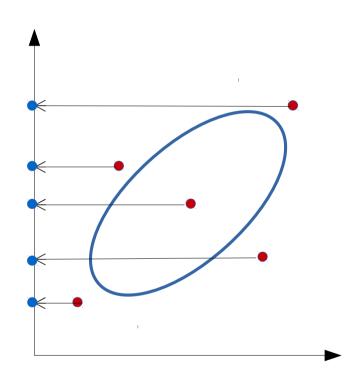


Unscented Marginalization

Keep the sigma points and suppress the marginalized dimensions.

The weights remain the same.

If there are more sigma points than needed, you can always transform it to parametric form and transform it back.



Unscented Functions

Let \mathbf{x}_a be a Gaussian random variable such that:

$$\mathbf{x}_a \sim \mathcal{UT}(\mathbf{x}_a; \mathbf{x}_a^{(i)}, w_m^{(i)}, w_c^{(i)})$$

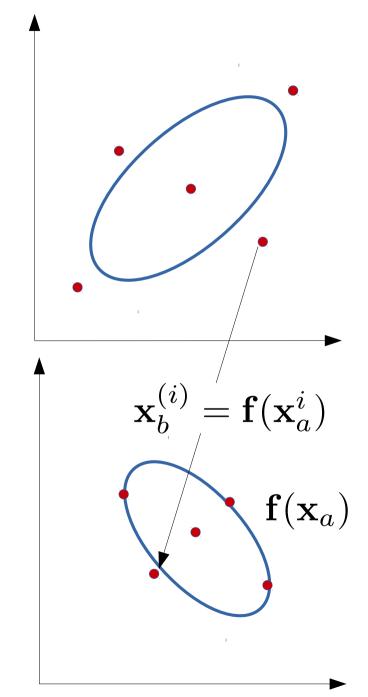
Let $\mathbf{x}_b = \mathbf{f}(\mathbf{x}_a)$ be the transformation of \mathbf{x}_a

An unscented approximation of

 \mathbf{x}_b can be obtained as:

$$\mathbf{x}_b \sim \mathcal{UT}(\mathbf{x}_b; \mathbf{x}_b^{(i)}, w_m^{(i)}, w_c^{(i)})$$

where
$$\mathbf{x}_b^{(i)} = \mathbf{f}(\mathbf{x}_a^{(i)})$$
 we write this as $\mathcal{X}_b^{(i)} = \mathbf{f}(\mathcal{X}_a^i)$



Unscented Chain Rule

We know

$$\mathbf{x}_a \sim \mathcal{UT}(\mathbf{x}_a; \mathbf{x}_a^{(i)}, w_m^{(i)}, w_c^{(i)})$$

$$p(\mathbf{x}_b|\mathbf{x}_a) = \mathcal{N}(\mathbf{x}_b; \mathbf{f}(\mathbf{x}_a), \mathbf{\Sigma}_{b|a})$$

We want to compute

$$p(\mathbf{x}_a, \mathbf{x}_b) = \mathcal{N}(\mathbf{x}_{a,b}; \mu_{a,b}, \mathbf{\Sigma}_{a,b})$$

The parameters are

$$\mu_{a,b} = \begin{pmatrix} \mu_a \\ \mu_b \end{pmatrix} = \begin{pmatrix} \mu_a \\ \mathbf{f}(\mu_a) \end{pmatrix}$$

$$\Sigma_{a,b} = \begin{pmatrix} \Sigma_a & \Sigma_{a,b} \\ \Sigma_{b,a} & \Sigma_{b|a} + \Sigma_{b,b} \end{pmatrix}$$

We need:

- the cross correlation coefficients
- •the covariance term due to the projection of \mathbf{x}_a through $\mathbf{f}(\mathbf{x}_a)$

Unscented Chain Rule

Projection of $\mathbf{x}_a \sim \mathcal{UT}(\mathbf{x}_a; \mathbf{x}_a^{(i)}, w_m^{(i)}, w_c^{(i)})$ through $\mathbf{f}(\mathbf{x}_a)$:

$$\mathbf{x}_b^{(i)} = \mathbf{f}(\mathbf{x}_a^{(i)})$$

$$\mu_b = \sum_i w_m^{(i)} \mathbf{f}(\mathbf{x}_a^{(i)})$$

$$\mathbf{x}_b^{(i)} = \mathbf{f}(\mathbf{x}_a^{(i)})$$

$$\mathbf{f}(\mathbf{x}_a)$$

$$\mathbf{x}_b^{(i)}$$

Unscented (cross) correlations:

$$\Sigma_{a,b} = \sum_{i} w_c^{(i)} (\mathbf{x}_a^{(i)} - \mu_a) (\mathbf{x}_b^{(i)} - \mu_b)^T$$

$$\Sigma_{b,b} = \sum_{i} w_c^{(i)} (\mathbf{x}_b^{(i)} - \mu_b) (\mathbf{x}_b^{(i)} - \mu_b)^T$$

Unscented Conditioning

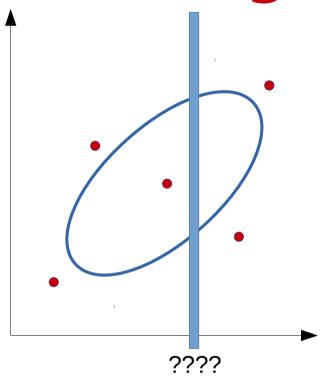
Not easy

Reason:

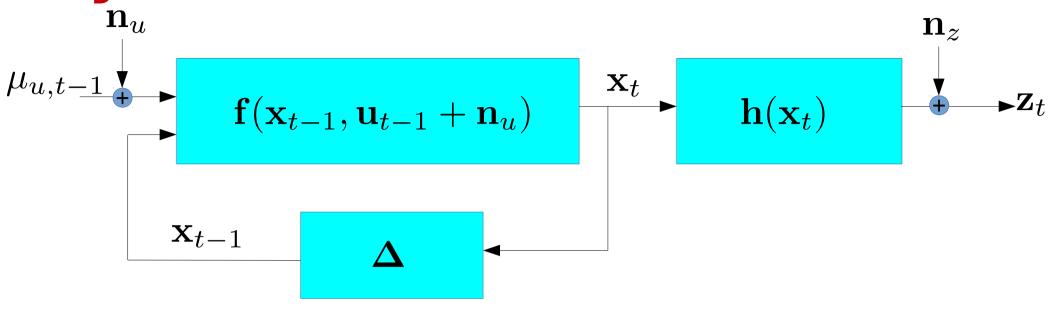
control points do not like to be sliced

What to do:

go back to the usual parameterization



System with Gaussian Noise



Inputs and Observations are affected by zero mean Gaussian noise. Initial belief is Gaussian

$$\mathbf{n}_u \sim \mathcal{N}(\mathbf{n}_u|0, \mathbf{\Sigma}_u)$$
 $p(\mathbf{x}_0) = \mathcal{N}(\mathbf{x}_0|\mu_0, \mathbf{\Sigma}_0)$ $\mathbf{n}_z \sim \mathcal{N}(\mathbf{n}_z|0, \mathbf{\Sigma}_z)$

For compactness we embed the noise in the control

$$\mathbf{u}_{t-1} \sim \mathcal{N}(\mathbf{n}_z | \mu_{u,t-1}, \mathbf{\Sigma}_{u,t})$$

Unscented Filtering

Some considerations

- The unscented transform behaves better under non-linear transformations
- We can replace
 - Unscented projection through function to compute the prediction
 - Unscented chain rule to compute the joint distribution over measurements and predicted states
 - The conditioning remains the same as in (E)KF

No Jacobians!

Predict

Incorporate the control by computing the Gaussian distribution of the next state given the input

The next state is a nonlinear function of the past state and the controls:

$$\mathbf{x}_t = \mathbf{f}(\mathbf{x}_{t-1}, \mathbf{u}_{t-1})$$

The previous states and controls are distributed according to

$$\begin{pmatrix} \mathbf{x}_{t-1|t-1} \\ \mathbf{u}_{t-1} \end{pmatrix} \sim \mathcal{N} \begin{bmatrix} \begin{pmatrix} \mu_{t-1|t-1} \\ \mu_{u,t-1} \end{pmatrix}, \begin{pmatrix} \mathbf{\Sigma}_{t-1|t-1} & \mathbf{0} \\ \mathbf{0} & \mathbf{\Sigma}_{u,t-1} \end{pmatrix} \end{bmatrix}$$

Predict

Compute the sigma points from the joint over previous states and controls

$$\begin{pmatrix} \mathbf{x}_{t-1|t-1} \\ \mathbf{u}_{t-1} \end{pmatrix} \sim \mathcal{N} \begin{bmatrix} \begin{pmatrix} \mu_{t-1|t-1} \\ \mu_{u,t-1} \end{pmatrix}, \begin{pmatrix} \mathbf{\Sigma}_{t-1|t-1} & \mathbf{0} \\ \mathbf{0} & \mathbf{\Sigma}_{u,t-1} \end{pmatrix} \end{bmatrix}$$

They have dimension dim(x) $\longrightarrow \mathcal{X}_{t-1|t-1}^{(i)}$ +dim(u)

$$\mathcal{X}_{t-1|t-1}^{(i)}$$

Project the point through transition function

$$\mathcal{X}_{t|t-1}^{(i)} = \mathbf{f}(\mathcal{X}_{t-1|t-1}^{(i)})$$

They have dimension dim(x)

Update: Chain Rule

Given the following distributions

$$p(\mathbf{z}_t|\mathbf{x}_t) = \mathcal{N}(\mathbf{z}_t; \mathbf{h}(\mathbf{x}_t), \mathbf{\Sigma}_{z|x})$$
$$p(\mathbf{x}_t) = \mathcal{UT}(\mathbf{x}_t; \mathcal{X}_{t|t-1}^{(i)})$$

We want to compute the joint p(x,z)

$$egin{array}{lll} \mu_{x,z} &=& \left(egin{array}{c} \mu_x \ \mu_z \end{array}
ight) \ oldsymbol{\Sigma}_{x,z} &=& \left(egin{array}{ccc} oldsymbol{\Sigma}_x & oldsymbol{\Sigma}_{x,z} \ oldsymbol{\Sigma}_{z,x} & oldsymbol{\Sigma}_{z|x} + oldsymbol{\Sigma}_{z,z} \end{array}
ight) \end{array}$$

Update Chain Rule

Projection of $\mathbf{x}_{t|t-1} \sim \mathcal{UT}(\mathbf{x}_t; \mathcal{X}_{t|t-1}^{(i)})$ through $\mathbf{h}(\mathbf{x}_t)$

$$\mathbf{z}_t^{(i)} = \mathbf{h}(\mathbf{x}_{t|t-1}^{(i)})$$

$$\mu_z = \sum_i w_m^{(i)}(\mathbf{z}_t^i)$$

Unscented (cross) correlations:

$$\Sigma_{x,z} = \sum_{i} w_c^{(i)} (\mathbf{x}_{t|t-1}^{(i)} - \mu_{t|t-1}) (\mathbf{z}_t^{(i)} - \mu_z)^T$$

$$\Sigma_{z,z} = \sum_{i} w_c^{(i)} (\mathbf{z}_t^{(i)} - \mu_z) (\mathbf{z}_t^{(i)} - \mu_z)^T$$

Update: Conditioning

Given the joint properties

$$egin{array}{lll} \mu_{x,z} &=& \left(egin{array}{c} \mu_x \ \mu_z \end{array}
ight) \ oldsymbol{\Sigma}_{x,z} &=& \left(egin{array}{ccc} oldsymbol{\Sigma}_x & oldsymbol{\Sigma}_{x,z} \ oldsymbol{\Sigma}_{z,x} & oldsymbol{\Sigma}_{z|x} + oldsymbol{\Sigma}_{z,z} \end{array}
ight) \end{array}$$

we condition on the actual measurement:

$$\mu_{t|t} = \mu_{t|t-1} + \Sigma_{x,z} \left(\Sigma_{z|x} + \Sigma_{z,z}\right)^{-1} (\mathbf{z}_t - \mu_z)$$

$$\Sigma_{t|t} = \Sigma_{t|t-1} - \Sigma_{x,z} \left(\Sigma_{z|x} + \Sigma_{z,z}\right)^{-1} \Sigma_{z,x}$$

Wrapup (UKF)

$$\begin{pmatrix} \mathbf{x}_{t-1|t-1} \\ \mathbf{u}_{t-1} \end{pmatrix} \sim \mathcal{N} \begin{bmatrix} \begin{pmatrix} \mu_{t-1|t-1} \\ \mu_{u,t-1} \end{pmatrix}, \begin{pmatrix} \boldsymbol{\Sigma}_{t-1|t-1} & \mathbf{0} \\ \mathbf{0} & \boldsymbol{\Sigma}_{u,t-1} \end{pmatrix} \end{bmatrix}$$

apply transition

compute sigma points

$$\mathcal{X}_{t|t-1}^{(i)} = \mathbf{f}(\mathcal{X}_{t-1|t-1}^{(i)}) \qquad \qquad \mathcal{X}_{t-1|t-1}^{(i)}$$

$$\mathcal{X}_{t-1|t-1}^{(i)}$$

Predict

compute sigma points and mean of

$$\mathbf{z}_t^{(i)} = \mathbf{h}(\mathbf{x}_{t|t-1}^{(i)})$$

$$\mathbf{z}_{t}^{(i)} = \mathbf{h}(\mathbf{x}_{t|t-1}^{(i)}) \sum_{i} \mathbf{\Sigma}_{x,z} = \sum_{i} w_{c}^{(i)} (\mathbf{x}_{t|t-1}^{(i)} - \mu_{t|t-1}) (\mathbf{z}_{t}^{(i)} - \mu_{z})^{T}$$

$$\mu_z = \sum_i w_m^{(i)} \mathbf{z}_t^{(i)}$$

$$\Sigma_{z,z} = \sum w_c^{(i)} (\mathbf{z}_t^{(i)} - \mu_z) (\mathbf{z}_t^{(i)} - \mu_z)^T$$

compute cross correlation of joint distribution

conditioning

$$\mu_{t|t} = \mu_{t|t-1} + \Sigma_{x,z} \left(\Sigma_{z|x} + \Sigma_{z,z} \right)^{-1} \left(\mathbf{z}_t - \mu_z \right)$$

$$oldsymbol{\Sigma}_{t|t} = oldsymbol{\Sigma}_{t|t-1} - oldsymbol{\Sigma}_{x,z} \left(oldsymbol{\Sigma}_{z|x} + oldsymbol{\Sigma}_{z,z}
ight)^{-1} oldsymbol{\Sigma}_{z,x}$$

Update

References

Further details are here (warmly recommended)

- Thomas Schoen, On Manipulating the Multivariate Gaussian Density
- Simon J. Julier, The Scaled Unscented Transform

Summary

- UKF is a <u>non-linear</u> extension of EKF
- It is reported to better deal with non-linear transition and observation functions
- Still relies on the Gaussian assumption and uses pure Gaussian conditioning
- Derivative free (better feasibility)