Note for lecture 08 In M/M/I, interpretation of $N = \frac{\rho}{I-\rho}$: think of P as CPU utilization (CPU%) → P > | : server cannot catch up with arrivals > # of items in a queue (i.e., N) the server fast enough grows much faster than the growth of CPU utilization (i.e., p)! Example 3.8: > III O KA III O KA $\begin{aligned}
\rho &= \frac{\lambda}{\mu} & \rho &= \frac{\lambda}{\mu} \\
N &= \frac{\rho}{1 - \rho} &= \frac{\lambda}{\mu - \lambda} & N &= \frac{\rho}{1 - \rho} &= \frac{\lambda}{\mu - \lambda}
\end{aligned}$ in both cases, a new arrival will see the same # of packets ahead of it statistically. Though the packets move K times faster for in the configuration on the right. *

P. P. Example 3.9: Statistically multiplexing us. TDM/FDM Consider m statistically identical and independent Poisson packet streams each with arrival rate in. The packet length for all streams ove independent and exponentially distributed. The average transmission time is -In If merging all streams into one, it is like $MM/1: m \xrightarrow{m} III$ and thus the average delay per packet is If using TDM/PDM instead, it is like in MAVi: $\frac{\lambda}{m} \to III \bigcirc$ Then the average delay

Num would be m times larger $\frac{\lambda}{m} \to III \bigcirc$ $T' = \frac{1}{M_m - M_m} = \frac{m}{M_m - \lambda} = m.T$ > why would anyone ever want to use TPM/FDM ? -> TDM/FDM guarantees a specific service rate for each stream. Plus, statistical multiplexing may introduce higher variability in latency, which is undesirable for applications such as voice/video streaming.

The PASTA theorem: Poisson Arrivals See Time Average 13 P4 (read Sections 3.3. 2-3.3.3)

—) "When arrivals are Poisson, [...] both an arriving and a departing customer in steady-state see a system that is statistically identical to the one seen by an observer looking at the system at an arbitrary time." We may use that to analyze the probability that an arrival will need to wait in queue. M/M/mlet u be the service rate, Server 2

Server 3

Server m $P = \frac{\lambda}{m\mu}$ the utilization To analyze M/M/m, it is helpful to take a closer look at how we got Pn \= Pn+1 H in M/1: recall the Markov chain in MM/1

=> Pn X = Pn+1 M

In M/M/m, we have the following Markor chain (m-1) m mp mp for $n \leq m-1$, for $n \geq m$, $P_0 \cdot \lambda = P_1 \cdot \mu$ Pmix + Pm (mp) = Pm-xx+Pm+1 (mp) P. x + Pox = Pox + P2.24 Proti / + Proti (mp) = Prox + Protiz (mp) P2: X+P, 2M= A: X+P3.3M $\Rightarrow P_n \cdot \lambda = P_{n+1}(m\mu)$ => Pn· \ = Pn+1·(n+1) / similarly, for nzm we have => Pm·mu=Pm-1· A $P_n = \frac{\lambda}{m\mu} P_{n-1}$ Pm = mu Pm-1 = (mp)" Po $\Rightarrow P_n = \frac{\lambda^n}{n! \, \mu^n} P_o$ $= \frac{\rho^n \cdot m^n}{m! \cdot m^{n-m}} P_0$ from P = mu we see that $(\frac{\lambda}{m})^n = (mp)^n$ $\Rightarrow P_n = \frac{m^m p^n}{m!} P_0$ therefore $P_n = \frac{(mp)^n}{n!} P_0$ Then Po can be obtained by $\sum_{n=0}^{\infty} P_n = 1$. Let Pa be the probability that an arrival will need to wait in queue (because all m servers are busy), and we have $PQ = \sum_{n=m}^{\infty} P_n = \sum_{n=$ $\Rightarrow P_{Q} = \frac{(m\rho)^{m}P_{o}}{m!(I-\rho)} \leftarrow \frac{\text{The Erlang C formula, named}}{\text{after the pioneer of Quencing Theory,}}$ A. K. Erlang.

Now let No be the expected # of customers P5 P6 waiting in queue. We may obtain it by In Pmtn. Alternatively, we may consider MM/m's relation to MM/1, and get $Na = Pa \cdot \frac{P}{1-P}$ from M/M/IFinally, N = Na + Ns where Ns is # of customer in service $N_S = \sum_{n=1}^{m} n \cdot P_n = \dots$ we can use Little's Theorem here, and Ns = 2. I $\Rightarrow N = m\rho + \frac{\rho P_0}{1-\rho} = m\rho$ \Rightarrow T = N/ $\lambda = \frac{1}{\mu} + \frac{Pa}{m\mu - \lambda} + using Little's Theorem.$ - To be specific, this leverages the fact that Na = E[X | queueing] in MM/m is equal to Na = E[X|queereing] in MM/1. (see the paragraph across P175-P176).
And in M/M/1, $E[X|qneneing] = \sum_{X} x \cdot \frac{P\{(X=X) \cap queueing\}}{P\{queueing\}} = \frac{N\alpha}{P\{queueing\}}$ and there Na = 7-0 and P & queueing 3 = P thus E[X|queveing] = T=P = F Therefore, in M/Wm Top = E[X|queneing] = Na = Na = Pa. -FP

Example 3.10 (compared with examples 3.8 and 3.9) M/M/m v.s. M/M/1 $\xrightarrow{\lambda} \overline{\prod} \bigcirc \}_{m}$ \xrightarrow{m} $T' = \frac{1}{m\mu - \lambda}$ $T = \frac{1}{\mu} + \frac{Pa}{m\mu - \lambda}$ $P = \frac{\lambda}{m\mu}$ \Rightarrow when $P \ll 1$, $P_{\alpha} \approx 0$ and my $\gg \lambda$, which imply $-\frac{1}{T}$, $\approx m$ \Rightarrow when $\rho \rightarrow 1$, $Pa \approx 1 \text{ and } \frac{1}{\mu} \ll \frac{1}{m\mu - \lambda}, \text{ which imply}$ \frac{1}{1} ≈ 1 The above result suggests that using one channel for statistical multiplexing will lead to a better performance in latency, compared with one using

multiple channels.