



Graphs and Model Selection

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Contents

- Basics of Graphical Models
 - Factorization Property
 - Markov Property
- Gaussian Graphical Model
- Graph Selection
 - Graphical Lasso Algorithm
 - Theoretical Guarantees for Graphical Lasso
 - Neighborhood Selection Algorithm



Basics of Graphical Models

- ullet A graph G=(V,E) consists of a set of vertices V and a set of edges E
- We focus exclusively on undirected graphs
- We can associate a collection of random variables $X=(X_1,X_2,\ldots,X_p)$ with the vertex set $V=\{1,2\ldots,p\}$ of some underlying graph
- Idea: see the structure of the underlying graph as a visual representation of the joint distribution of the random variables



Factorization Property

- ullet Let ${\mathcal C}$ be the set of all cliques in the graph G
- For a clique $C \in \mathcal{C}$ a compatibility function ψ_C is a function of the subvector $x_C := (x_s, s \in C)$ taking positive real values
- ullet Given a collection of compatibility functions we say that a probability distribution P factorizes over G if and only if

$$P(x_1, \dots, x_p) = \frac{1}{Z} \prod_{C \in \mathcal{C}} \psi_C(x_C)$$

where $Z = \sum_{x \in \chi^p} \prod_{c \in \mathcal{C}} \psi_C(x_C)$ ensures that P is properly normalized

 Such a factorization can lead to savings in storage and computation if the clique sizes are not too large



Factorization Property

Example:

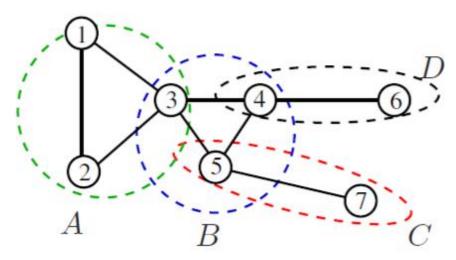


Figure: Trevor Hastie, Robert Tibshirani and Martin Wainwright (2015). Statistical Learning with Sparsity: The Lasso and Generalizations, p242.

ullet P factorizes over this graph if it has the form

$$P(x_1, \dots, x_7) = \frac{1}{Z} \psi_A(x_1, x_2, x_3) \psi_B(x_3, x_4, x_5) \psi_D(x_4, x_6) \psi_C(x_5, x_7)$$

For some choice of compatility functions $\{\psi_A, \psi_B, \psi_C, \psi_D\}$



Markov Property

- ullet Let S denote a cut set disconnection the graph into components A and B
- ullet We say that a random vector X is Markov with respect to the graph G if

$$X_A \perp \!\!\!\perp X_B \mid X_S$$
 for all cut sets $S \subset V$

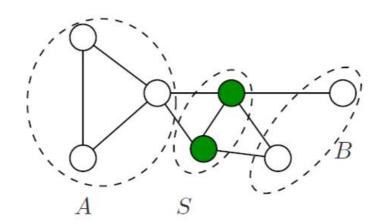


Figure: Trevor Hastie, Robert Tibshirani and Martin Wainwright (2015). Statistical Learning with Sparsity: The Lasso and Generalizations, p242.



Hammersley-Clifford Theorem

- This is the fundamental theorem of random fields and gives necessary and sufficient conditions under which a strictly positive probability distribution can be represented as a Markov network
- **Theorem:** For a strictly positive probability distribution P of a random vector X the two characterizations are equivalent; the distribution of X factorizes according to the graph G if and only if it is Markov with respect to G.



Gaussian Graphical Model

ullet Given a p dimensional Gaussian distribution with mean vector μ and covariance matrix Σ :

$$P_{\mu, \Sigma}(x) = \frac{1}{(2\pi)^{\frac{p}{2}} \det(\Sigma)^{\frac{1}{2}}} e^{-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)}$$

This can equivalently be formulated as

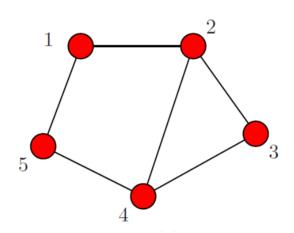
$$P_{\gamma,\Theta}(x) = \exp\{\sum_{s=1}^{p} \gamma_s x_s - \frac{1}{2} \sum_{s,t=1}^{p} \boldsymbol{\theta}_{st} x_s x_t - A(\boldsymbol{\Theta})\}$$

where $\Theta = \Sigma^{-1}$ the precision matrix, $\gamma = \Theta \mu$ and $A(\Theta) = -\frac{1}{2} \log \det(\Theta/(2\pi))$



Gaussian Graphical Model

- ullet This new representation allows us to discuss factorization properties in terms of the sparsity pattern of ullet
- If X factorizes according to the graph G then for $(s,t) \notin E$ we must have that $\theta_{st} = 0$
- Example:



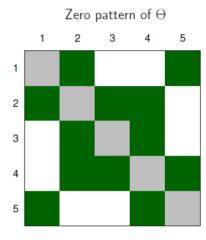


Figure: Trevor Hastie, Robert Tibshirani and Martin Wainwright (2015). Statistical Learning with Sparsity: The Lasso and Generalizations, p246.



Graph Selection

Problem: Given a collection of samples from a graphical model, where the underlying graph structure is unknown. How can we find the correct graph with high probability?



Graph Selection for Gaussian Graphical Models

- ullet Suppose ${f X}$ represents samples from a zero-mean multivariate Gaussian distribution with unknown precision matrix $oldsymbol{\Theta}$
- One can show that the log-likelihood of this distribution takes the form

$$\mathcal{L}(\mathbf{\Theta}, \mathbf{X}) = \frac{1}{N} \sum_{i=1}^{N} \log P_{\mathbf{\Theta}}(x_i) = \log \det \mathbf{\Theta} - \operatorname{trace}(\mathbf{S}\mathbf{\Theta})$$

where $\mathbf{S} = \frac{1}{N} \sum_{i=1}^{N} x_i x_i^T$ the empirical covariance matrix and

$$\log \det(\mathbf{\Theta}) = \begin{cases} \sum_{j=1}^{p} \log(\lambda_{j}(\mathbf{\Theta})) & \text{if } \mathbf{\Theta} \succ 0 \\ -\infty & \text{otherwise} \end{cases}$$

where $\lambda_j(\mathbf{\Theta})$ is the j-th eigenvalue of $\mathbf{\Theta}$



Graph Selection for Gaussian Graphical Models

- ullet This function is strictly concave, so that if the maximum is achieved it must be unique and defines the maximum likelihood estimate $\hat{m \Theta}$
- If we let N go to infinity $\hat{\mathbf{\Theta}}$ converges to the true precision matrix
- ullet But if N < p no maximum likelihood estimator exists and we need to consider suitably constrained or regularized forms
- If we are seeking Graphical models based on sparse graphs we could consider the following convex optimization problem

$$\hat{\boldsymbol{\Theta}} \in \arg \max_{\boldsymbol{\Theta} \succeq 0} \{ \log \det \boldsymbol{\Theta} - \operatorname{trace}(\mathbf{S}\boldsymbol{\Theta}) - \lambda \rho_1(\boldsymbol{\Theta}) \}$$

where
$$ho_1(\mathbf{\Theta}) = \sum_{s \neq t} |\boldsymbol{\theta}_{st}|$$



The subgradient equation corresponding to this problem is given by

$$\mathbf{\Theta}^{-1} - \mathbf{S} - \lambda \mathbf{\Psi} = \mathbf{0}$$

where Ψ has diagonal entries 0, $\psi_{ik} = \operatorname{sign}(\theta_{ik})$ if $\theta_{ik} \neq 0$ and $\psi_{ik} \in [-1,1]$ if $\theta_{ik} = 0$

 To solve this problem via blockwise coordinate descent we partition the matrices into one column versus the rest: $\mathbf{\Theta} = \begin{bmatrix} \mathbf{\Theta}_{11} & \mathbf{\theta}_{12} \\ \mathbf{\theta}_{12}^T & \mathbf{\theta}_{22} \end{bmatrix}$, $\mathbf{S} = \begin{bmatrix} \mathbf{S}_{11} & \mathbf{s}_{12} \\ \mathbf{s}_{12}^T & \mathbf{s}_{22} \end{bmatrix}$



• Then
$$\mathbf{W} = \mathbf{\Theta}^{-1} = \begin{bmatrix} \mathbf{W}_{11} & \mathbf{w}_{12} \\ \mathbf{w}_{21} & w_{22} \end{bmatrix} = \begin{bmatrix} (\mathbf{\Theta}_{11} - \frac{\theta_{12}\theta_{21}}{\theta_{22}})^{-1} & -\mathbf{W}_{11}\frac{\theta_{12}}{\theta_{22}} \\ . & . \end{bmatrix}$$

So for the last column of our subgradient equation we get:

$$\mathbf{w}_{12} - \mathbf{s}_{12} + \lambda \psi_{12} = \mathbf{W}_{11} \boldsymbol{\beta} - \mathbf{s}_{12} + \lambda \psi_{12} = 0$$

where
$$oldsymbol{eta} = -oldsymbol{ heta}_{12}/ heta_{22}$$

• It can be seen that this is equivalent to a modified version of the estimating equations for a lasso regression



- Recall that in the usual regression setup with outcome \mathbf{y} and predictor matrix \mathbf{Z} the lasso minimizes $\frac{1}{N} \|\mathbf{y} \mathbf{Z}\boldsymbol{\beta}\|_2^2 + \lambda \|\boldsymbol{\beta}\|_1$
- This has the subgradient equations $\frac{1}{N}\mathbf{Z}^T\mathbf{Z}\boldsymbol{\beta} \frac{1}{N}\mathbf{Z}^T\mathbf{y} + \lambda\operatorname{sign}(\boldsymbol{\beta}) = \mathbf{0}$
- Comparing to the last column of our subgradient equation shows that $\frac{1}{N}\mathbf{Z}^T\mathbf{y}$ corresponds to \mathbf{s}_{12} and $\frac{1}{N}\mathbf{Z}^T\mathbf{Z}$ corresponds to \mathbf{W}_{11}
- Hence we want to minimize $\frac{1}{2} \| \mathbf{W}_{11}^{\frac{1}{2}} \boldsymbol{\beta} \mathbf{W}_{11}^{-\frac{1}{2}} \mathbf{s}_{12} \|_2^2 + \lambda \| \boldsymbol{\beta} \|_1$

Algorithm 9.1 Graphical Lasso.

- 1. Initialize $\mathbf{W} = \mathbf{S}$. Note that the diagonal of \mathbf{W} is unchanged in what follows.
- 2. Repeat for $j = 1, 2, \dots, p, 1, 2, \dots, p, \dots$ until convergence:
 - (a) Partition the matrix W into part 1: all but the j^{th} row and column, and part 2: the j^{th} row and column.
 - (b) Solve the estimating equations $\mathbf{W}_{11}\boldsymbol{\beta} \mathbf{s}_{12} + \lambda \cdot \operatorname{sign}(\boldsymbol{\beta}) = 0$ using a cyclical coordinate-descent algorithm for the modified lasso.
 - (c) Update $\mathbf{w}_{12} = \mathbf{W}_{11}\hat{\boldsymbol{\beta}}$
- 3. In the final cycle (for each j) solve for $\hat{\boldsymbol{\theta}}_{12} = -\hat{\boldsymbol{\beta}} \cdot \hat{\theta}_{22}$, with $1/\hat{\theta}_{22} =$ $w_{22} - \mathbf{w}_{12}^T \hat{\boldsymbol{\beta}}.$

Figure: Trevor Hastie, Robert Tibshirani and Martin Wainwright (2015). Statistical Learning with Sparsity: The Lasso and Generalizations, p250.



- Example in R
 - glasso package
 - manual graphical lasso
 - plotting the coefficients for different values of λ



- ullet If we repeat the algorithm for a range of different values for λ we can plot the estimates for the entires of the precision matrix against $\rho_1(\mathbf{\Theta})$
- Example: Here the true precision matrix is

$$\mathbf{\Theta} = \begin{bmatrix} 2 & 0.6 & 0 & 0 & 0.5 \\ 0.6 & 2 & -0.4 & 0.3 & 0 \\ 0 & -0.4 & 2 & -0.2 & 0 \\ 0 & 0.3 & -0.2 & 2 & 0 \\ 0.5 & 0 & 0 & 0 & 2 \end{bmatrix}$$

• If we simulate data from the multivariate gaussian with Θ the true values are achieved at the right side of the plot

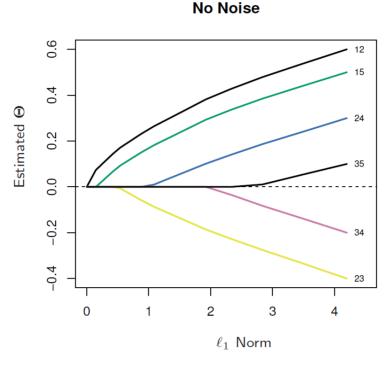


Figure: Trevor Hastie, Robert Tibshirani and Martin Wainwright (2015). Statistical Learning with Sparsity: The Lasso and Generalizations, p251.



• However if we add some standard Gaussian noise to each column the true edge set is not recovered for any value of λ With Noise

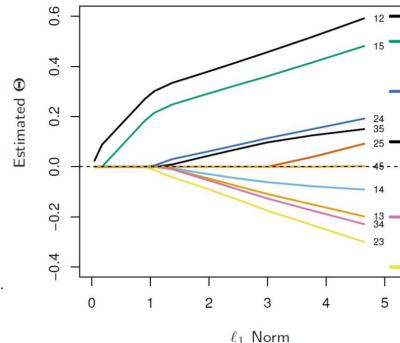


Figure: Trevor Hastie, Robert Tibshirani and Martin Wainwright (2015). Statistical Learning with Sparsity: The Lasso and Generalizations, p251.



Theoretical Guarantees for Graphical Lasso

- Plot of the operator norm $\|\hat{\Theta} \Theta\|_2$ versus the sample size N for three different graph sizes where $\lambda_N = 2\sqrt{\frac{\log p}{N}}$ was used as the regularization parameter
- We see that larger graphs require more samples for a consistent estimation

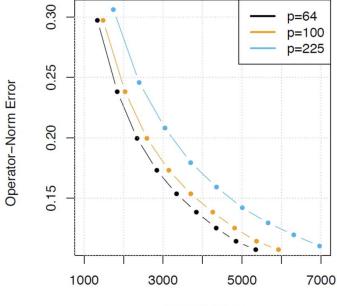


Figure: Trevor Hastie, Robert Tibshirani and Martin Wainwright (2015). Statistical Learning with Sparsity: The Lasso and Generalizations, p252.

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Sample Size

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Neighborhood Selection

- High-dimensional Graphs and Variable Selection with the Lasso; Meinshausen and Buhlmann (2006)
- It is an alternative method for graph selection that is computationally efficient and consistent for high dimensional graphs
- For a random vector $X=(X_1,\ldots,X_p)$ consider the conditional distribution of X_s given the random vector $X_{\setminus \{s\}}=(X_1,\ldots,X_{s-1},X_{s+1},\ldots,X_p)$
- ullet By the properties of a graphical model the only relevant variables are those in the neighborhood set $\mathcal{N}(s)$



Neighborhood Selection for Gaussians

ullet In the case of a multivarite Gaussian the conditional distribution of X_s given $X_{\backslash \{s\}}$ is

$$X_s = X_{\backslash \{s\}} \beta^s + W_{\backslash \{s\}}$$

where $W_{\backslash \{s\}}$ corresponds to a prediction error independent of $X_{\backslash \{s\}}$

- The key property is that the regression vector β^s satisfies $\operatorname{supp}(\beta^s) = \mathcal{N}(s)$
- It is natural to estimate β via the lasso



Neighborhood Selection

Algorithm 9.2 Neighborhood-based graph selection for Gaussian graphical models.

- 1. For each vertex $s = 1, 2, \ldots, p$:
 - (a) Apply the lasso to solve the neighborhood prediction problem:

$$\widehat{\beta}^{s} \in \underset{\beta^{s} \in \mathbb{R}^{p-1}}{\operatorname{arg\,min}} \left\{ \frac{1}{2N} \sum_{i=1}^{N} \left(x_{is} - x_{i,V \setminus \{s\}}^{T} \beta^{s} \right)^{2} + \lambda \|\beta^{s}\|_{1} \right\}. \tag{9.25}$$

- (b) Compute the estimate $\widehat{\mathcal{N}}(s) = \operatorname{supp}(\widehat{\beta}^s)$ of the neighborhood set $\mathcal{N}(s)$.
- 2. Combine the neighborhood estimates $\{\widehat{\mathcal{N}}(s), s \in V\}$ via the AND or OR rule to form a graph estimate $\widehat{G} = (V, \widehat{E})$.

Figure: Trevor Hastie, Robert Tibshirani and Martin Wainwright (2015). Statistical Learning with Sparsity: The Lasso and Generalizations, p256.



Neighborhood Selection

• Example in R unsing the glasso package setting approx=TRUE



Questions?

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