# STAT 217 homework 6

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## 1 6.7

## 1.1

Let  $X_1$ ,  $X_2$ ,  $X_3$  denote the time Ben, Max and Yolanda need to wait before being served. Then we have

$$X_1 \sim Exp(1)$$

$$X_2 \sim Exp(2)$$

$$X_3 \sim Exp(3)$$

Let  $M = min\{X_1, X_2, X_3\}$ . Then

$$P(M = X_3) = \frac{3}{1+2+3} = \frac{1}{2}.$$

## 1.2

$$P(X_1 < X_3) = \int_0^\infty P(X_1 < X_3 | X_1 = x) P(X_1 = x) dx$$

$$= \int_0^\infty P(X_3 > x) e^{-x} dx$$

$$= \int_0^\infty e^{-x} dx \int_x^\infty 3 e^{-3y} dy$$

$$= \int_0^\infty e^{-x} e^{-3x} dx$$

$$= \frac{1}{4}.$$

### 1.3

Let  $M = min\{X_1, X_2, X_3\}$ . Then

$$M \sim Exp(1+2+3) = Exp(6).$$

So

$$E(M) = \frac{1}{6}.$$

2 6.13

2.1

$$P(N_{s} = k | N_{t} = n) = \frac{P(N_{s} = k, N_{t} = n)}{P(N_{t} = n)}$$

$$= \frac{P(N_{s} = k)P(N_{t-s} = n - k)}{P(N_{t} = n)}$$

$$= \frac{\frac{e^{-\lambda s(\lambda s)^{k}}}{k!} \frac{e^{-\lambda (t-s)(\lambda (t-s))^{n-k}}}{(n-k)!}}{\frac{e^{-\lambda t(\lambda t)^{n}}}{n!}}$$

$$= \frac{n!}{k!(n-k)!} \frac{s^{k}(t-s)^{n-k}}{t^{n}}$$

$$= \binom{n}{k} (\frac{s}{t})^{k} (1 - \frac{s}{t})^{n-k}.$$

Thus

$$(N_s|N_t=n) \sim Binomial(n,\frac{s}{t}).$$

2.2

$$P(N_s = k | N_t = n) = P(N_{s-t} = k - n) = \frac{e^{-(s-t)\lambda}((s-t)\lambda)^{k-n}}{(k-n)!}.$$

3 6.14

$$P(X = x) = \int_0^\infty P(N_T = k|T = t)f_t(t)dt$$
$$= \int_0^\infty \frac{e^{bt}(bt)^k}{k!} re^{-rt}dt$$
$$= \frac{r}{r+b} (\frac{b}{r+b})^k.$$

Thus, X has a geometric distribution.

4 6.21

$$Cov(N_s, N_t) = E(N_t N_s) - E(N_t)E(N_s)$$
$$= E(N_t N_s) - \lambda^2 st.$$

$$E(N_t N_s) = E(E(N_t N_s | N_s))$$

$$= E(N_s E(N_t | N_s))$$

$$= E(N_s (N_s + E(N_{t-s})))$$

$$= E(N_s^2 + (t-s)_s)$$

$$= E(N_s^2) + (t-s)\lambda E(N_s)$$

$$= (E(N_s))^2 + Var(N_s) + (t-s)\lambda E(N_s)$$

$$= \lambda s + \lambda^2 st.$$

Thus,

$$Cov(N_t, N_s) = \lambda s + \lambda^2 st - \lambda^2 st = \lambda s.$$

Thus

$$Corr(N_t, N_s) = \frac{Cov(N_t, N_s)}{\sqrt{Var(N_t)Var(N_s)}}$$
$$= \frac{\lambda s}{\sqrt{\lambda t \cdot \lambda s}}$$
$$= \sqrt{\frac{s}{t}}.$$

Proved.

## $5 \quad 6.24$

### 5.1

Let  $X_1, X_2$  denote the time first see a meadowlark and a sparrow. Then

$$X_1 \sim Exp(\lambda)$$
  
 $X_2 \sim Exp(\mu)$ 

Then

$$P(X_1 < X_2) = \frac{\lambda}{\lambda + \mu}.$$

#### 5.2

Let  $N_m(t)$  denote the poisson process with parameter  $\lambda$  and  $N_s(t)$  denote the poisson process with parameter  $\mu$ .

Let  $N_{m+s}(t) = N_m(t) + N_s(t)$ . Then  $N_{m+s}$  is a poisson process with parameter  $(\lambda + \mu)$ .

$$P(N_{m+s}(1) = 1) = (\lambda + \mu)e^{-(\lambda + \mu)}.$$

### 5.3

Since the two process are independent,

$$P(N_s(2) = 1, N_m(2) = 2) = P(N_s(2) = 1)P(N_m(2) = 2)$$

$$= \frac{2\mu e^{(-2\mu)}}{1!} \frac{(2\lambda)^2 e^{-2\lambda}}{2!}$$

$$= 4\mu \lambda^2 e^{-2(\lambda+\mu)}.$$

## 6 3.56

An absorbing Markov chain is built on  $\{\emptyset, H, HT, HTT, HTTH, HTTHH\}$ , and

$$P(H) = \frac{1}{3}$$
$$P(T) = \frac{2}{3}$$

Thus the transition matrix is

$$P = \begin{pmatrix} states & \emptyset & H & HT & HTT & HTTH & HTTHH \\ \emptyset & \frac{2}{3} & \frac{1}{3} & 0 & 0 & 0 & 0 \\ H & 0 & \frac{1}{3} & \frac{2}{3} & 0 & 0 & 0 \\ HT & 0 & \frac{1}{3} & 0 & \frac{1}{2} & 0 & 0 \\ HTT & \frac{2}{3} & 0 & 0 & 0 & \frac{1}{3} & 0 \\ HTTH & 0 & 0 & \frac{2}{3} & 0 & 0 & \frac{1}{3} \\ HTTHH & 0 & 0 & 0 & 0 & 0 & 1 \end{pmatrix}.$$

Making HTTHH an absorbing state, then

$$F = (I - Q)^{-1}$$

$$= \begin{pmatrix} \frac{1}{3} & -\frac{1}{3} & 0 & 0 & 0\\ 0 & \frac{2}{3} & -\frac{2}{3} & 0 & 0\\ 0 & -\frac{1}{3} & 1 & -\frac{2}{3} & 0\\ -\frac{2}{3} & 0 & 0 & 1 & -\frac{1}{3}\\ 0 & 0 & -\frac{2}{3} & 0 & 1 \end{pmatrix}^{-1}$$

$$= \begin{pmatrix} 21 & 17.25 & 13.5 & 9 & 3\\ 18 & 17.25 & 13.5 & 9 & 3\\ 18 & 15.75 & 13.5 & 9 & 3\\ 18 & 15.75 & 13.5 & 9 & 3\\ 12 & 10.5 & 9 & 6 & 3 \end{pmatrix}.$$

The row sums are

Thus, the expected number of flips to get HTTHH is 63.75.