Probability & Statistics for EECS: Homework #09

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Problem 1

Solution

(a) X,Y are discrete

$$P(Y = y | X = x) = \frac{P(Y = y, X = x)}{P(X = x)}$$
$$= \frac{P(X = x | Y = y)P(Y = y)}{P(X = x)}$$

This uses Bayes rule and the confine of it is P(Y = y), P(X = x) > 0.

(b) X is discrete, Y is continuous

$$\begin{split} P(Y \in (y-\varepsilon,y+\varepsilon)|X=x) &= \frac{P(X=x,Y \in (y-\varepsilon,y+\varepsilon))}{P(X=x)} \\ &= \frac{P(X=x|Y \in (y-\varepsilon,y+\varepsilon))P(Y \in (y-\varepsilon,y+\varepsilon))}{P(X=x)} \\ &\frac{P(Y \in (y-\varepsilon,y+\varepsilon)|X=x)}{2\varepsilon} &= \frac{P(X=x|Y \in (y-\varepsilon,y+\varepsilon))\frac{P(Y \in (y-\varepsilon,y+\varepsilon))}{2\varepsilon}}{P(X=x)} \end{split}$$

Because the definition of integral and when $\varepsilon = 0$

$$f_Y(y|X = x) = \frac{P(X = x|Y = y)f_Y(y)}{P(X = x)}$$

This uses Bayes rule and the confine of it is $P(Y \in (y - \varepsilon, y + \varepsilon)), P(X = x) > 0$.

(c) X is continuous, Y is discrete

$$\begin{split} P(Y=y|X\in(x-\varepsilon,x+\varepsilon)) &= \frac{P(X\in(x-\varepsilon,x+\varepsilon),Y=y)}{P(X\in(x-\varepsilon,x+\varepsilon))} \\ &= \frac{P(X\in(x-\varepsilon,x+\varepsilon))Y=y)P(Y=y)}{P(X\in(x-\varepsilon,x+\varepsilon))} \\ P(Y\in(y-\varepsilon,y+\varepsilon)|X=x) &= \frac{\frac{P(X\in(x-\varepsilon,x+\varepsilon)|Y=y)}{2\varepsilon}P(Y=y)}{\frac{P(X\in(x-\varepsilon,x+\varepsilon))}{2\varepsilon}} \end{split}$$

Because the definition of integral and when $\varepsilon = 0$

$$P(Y = y | X = x) = \frac{f_X(x | Y = y)P(Y = y)}{f_X(x)}$$

This uses Bayes rule and the confine of it is $P(Y = y), P(X \in (x - \varepsilon, x + \varepsilon)) > 0$.

(d) X,Y are continuous

$$P(Y \in (y - \varepsilon, y + \varepsilon) | X \in (x - \varepsilon, x + \varepsilon)) = \frac{P(Y \in (y - \varepsilon, y + \varepsilon), X \in (x - \varepsilon, x + \varepsilon))}{P(X \in (x - \varepsilon, x + \varepsilon))}$$

$$\frac{P(Y \in (y - \varepsilon, y + \varepsilon) | X \in (x - \varepsilon, x + \varepsilon))}{2\varepsilon} = \frac{\frac{P(Y \in (y - \varepsilon, y + \varepsilon), X \in (x - \varepsilon, x + \varepsilon))}{4\varepsilon^2}}{\frac{P(X \in (x - \varepsilon, x + \varepsilon))}{2\varepsilon}}$$

Because the definition of integral and when $\varepsilon = 0$

$$f_{Y|X}(y|x) = \frac{f_{X,Y}(x,y)}{f_X(x)}$$
$$= \frac{f_{X|Y(x|y)f_Y(y)}}{f_X(x)}$$

This uses Bayes rule and the confine of it is $P(Y \in (y - \varepsilon, y + \varepsilon)), P(X \in (x - \varepsilon, x + \varepsilon)) > 0$.

(a) Because N = X + Y, then P(N = x + y | X = x, Y = y) = 1 and we can have: P(X = x, Y = y, N = n) = P(X = x, Y = y)

 $Bayes \ rule$

Because X and Y are i.i.d., then we can have:

 $P(X=x,Y=y)=P(X=x)P(Y=y)=p*q^x*p*q^y=p^2*q^{x+y}=p^2q^n$, for nonnegative integers x, y, n, and n = x + y.

(b) From (1), we can also have:

$$P(X = x, N = n) = P(X = x, Y = n - x, N = n)$$

$$= P(X = x, Y = n - x)$$

$$= P(X = x)P(Y = n - x)$$

$$= p * q^{x} * p * q^{n-x} = p^{2}q^{n}$$

For 0 <= x <= n.

(c)

$$P(X = x | N = n) = \frac{P(X = x, N = n)}{P(N = n)}$$

$$LOTP$$

$$= \frac{p^2 q^n}{\sum_{x=0}^n P(N = n | X = x) P(X = x)}$$

$$\begin{split} N &= X + Y \\ &= \frac{p^2 q^n}{\sum_{x=0}^n P(X + Y = n | X = x) P(X = x)} \\ &= \frac{p^2 q^n}{\sum_{x=0}^n P(Y = n - x | X = x) P(X = x)} \end{split}$$

Because of X, Y are i.i.d.

$$= \frac{p^2 q^n}{\sum_{x=0}^n P(Y = n - x) P(X = x)}$$

$$= \frac{p^2 q^n}{\sum_{x=0}^n p^2 q^n}$$

$$= \frac{p^2 q^n}{(n+1)p^2 q^n}$$

$$= \frac{1}{n+1}$$

At the same time, obviously we can get $N \sim NBin(2, p)$. That is because X, Y are i.i.d. Geom(p), and N=X+Y, then we can easily find that all we need to calculate is the number of the failture times before the second success. So we need to find one success and n failures in the first n+1 times. Then we have $P(N=n) = \binom{n+2-1}{2-1} p * p * q^n = (n+1)p^2q^n$. Therefore we can also get it. Moreover, when N is given, $P(X=x|N=n) = \frac{1}{n+1}$ is exact.

Solution

(a) From the question, we have

$$P(X \le x | X > c) = \frac{P(X \le x, X > c)}{P(X > c)}$$

$$= \frac{P(c < X <= x)}{P(X > c)}$$

$$= \frac{P(X <= x) - P(X <= c)}{1 - P(X <= c)}$$

$$= \frac{1 - e^{-\lambda x} - 1 + e^{-\lambda c}}{1 - 1 + e^{-\lambda c}}$$

$$= 1 - e^{-\lambda(x - c)}$$

Then the conditional PDF of X given X>c is $\lambda e^{-\lambda(x-c)}$.

(b) First if $x \le c$, then $P(X \le x | X \le c) = 1$. On the other hand when x > c, we have

$$\begin{split} P(X <= x | X < c) &= \frac{P(X <= x, X < c)}{P(X < c)} \\ &= \frac{P(X <= x)}{P(X < c)} \\ &= \frac{1 - e^{-\lambda x}}{1 - e^{-\lambda c}} \end{split}$$

Then the conditional PDF of X given X < c is $\frac{\lambda}{1 - e^{-\lambda c}} e^{-\lambda x}$.

Solution

(a) Because $M = max(U_1, U_2, U_3)$, then $P(M \le m) = P(U_1 \le m, U_2 \le m, U_3 \le m)$. Because U_1, U_2, U_3 are i.i.d. Unif(0,1), then

$$F_M(m) = P(M \le m) = P(U_1 \le m)P(U_2 \le m)P(U_3 \le m) = m^3$$

For, $0 \le m \le 1$.

Then

$$f_M(m) = 3m^2$$

For, $0 \le m \le 1$.

Because $P(M \le m) = P(L \le l, M \le m) + P(L > l, M \le m)$,

then
$$P(L \le l, M \le m) = P(M \le m) - P(L > l, M \le m)$$
.

So we need to calculate $P(L > l, M \le m)$ first and we also have the confine of l, which is $0 \le l \le m \le 1$.

$$\begin{split} P(L > l, M <= m) &= P(min(U_1, U_2, U_3) > l, max(U_1, U_2, U_3) <= m) \\ &= P(l < U_1 <= m, l < U_2 <= m, l < U_3 <= m) \\ &= (m - l)^3 \end{split}$$

Then

$$F_{L,M}(l,m) = (L \le l, M \le m) = m^3 - (m-l)^3$$

= $3m^2l - 3ml^2 + l^3$

And then we can also get the PDF by respectively differentiating l and m, that is

$$f_{L,M}(l,m) = 6(m-l)$$

for 0 <= l <= m <= 1.

(b) What we want to get is $f_{M|L}(m,l) = \frac{f_{M,L}(m,l)}{f_L(l)}$, so we just need to calculate $f_L(l)$, for $0 \le l \le 1$.

$$1 - F_L(l) = P(L > l) = P(U_1 > l)P(U_2 > l)P(U_3 > l) = (1 - l)^3$$

$$F_L(l) = 1 - (1 - l)^3$$

$$F_L(l) = 3l - 3l^2 + l^3$$

$$f_L(l) = 3 - 6l + 3l^2$$

$$f_L(l) = 3(1 - l)^2$$

So
$$f_{M|L}(m,l) = \frac{2(m-l)}{(1-l)^2}$$
, for $0 <= l <= m <= 1$.

Solution

(a) First we can divide this problem into three cases.

case1,
$$l > m$$
, then $P(L = l, M = m) = 0$
case2, $l = m$, then $P(L = l, M = m) = P(X = l, Y = l)$
Because they are independent Geom distribution,
so $P(X = l, Y = l) = P(X = l)P(Y = l) = p * q^{l} * p * q^{l} = p^{2} * q^{2l}$.
case3 is as follow,

$$\begin{split} P(L=l, M=m) &= P(X=l, Y=m) + P(X=m, Y=l) \\ Because \ X, Y \ are \ i.i.d. \\ &= P(X=l)P(Y=m) + P(X=m)P(Y=l) \\ &= p*q^l*p*q^m + p*q^m*p*q^l \\ &= 2p^2q^{m+l} \end{split}$$

From the cases we divide, obviously they are not independent. Because if we have known something about L, then we can also have some information about M, due to $L \le M$.

(b) From (1), we can have

$$\begin{split} P(L=l) &= \sum_{m=0}^{\infty} P(L=l, M=m) \\ &= P(L=l, M=l) + \sum_{m=l+1}^{\infty} P(L=l, M=m) \\ &= p^2 q^{2l} + 2p^2 q^l \sum_m m = l + 1^{\infty} q^m \\ &= p^2 q^{2l} + 2pq^{2l+1} \end{split}$$

By using story, all we need to do is to translate L = min(X, Y).

L can be seemed as X,Y at least happen one for the first success when L=l. So L is also a Geom distribution and $p_l = 2p - p^2$.

That is $L \sim Geom(2p - p^2)$.

Then
$$P(L=l) = (2p-p^2)(1-2p+p^2)^l = (2p-p^2)((1-p)^2)^l = (2p-p^2)q^{2l} = (p^2+2p(1-p))q^{2l} = p^2q^{2l} + 2pq^{2l+1}$$

(c) Because E[M+L] = E[X+Y], then we have

$$\begin{split} E[M] + E[L] &= E[X] + E[Y] \\ E[M] &= E[X] + E[Y] - E[L] \\ &= \frac{1-p}{p} + \frac{1-p}{p} - \frac{1-2p+p^2}{2p-p^2} \\ &= \frac{(1-p)(3-p)}{(2-p)p} \end{split}$$

(d) From the question and (a), and because $n \ge 0$, then we have

$$P(L = l, M - L = n) = P(L = l, M = n + l)$$

= $2p^2q^{l+n+l}$
= $2p^2q^{n+2l}$

So for the joint PMF of L and M-L, and because n and l are nonnegative integers, then we can use the theorem and get

theorem and get $f_{L,M-L}(l,n) = g(l)h(n)$, and $g(l) = a * q^{2l}$, and $h(n) = b * q^n$. To get the valid PMF, $g(l) = a \sum_{l=0}^{\infty} (q^2)^l = a * \frac{1}{1-q^2} = 1$, then we can get $a = 1 - q^2$.

And for another one, this must be valid too, so $h(n) = \frac{2p^2q^n}{1-q^2}$.

Therefore L, M-L are independent.