

# Probability & Statistics for EECS:

## Homework #11

Due on April 30, 2023 at 23:59

Name: **Wang Yunfei**

Student ID: 2021533135

## Problem 1

### Solution

- (a) Yes, that is because  $t_1X + t_2Y + t_3(X + Y) = (t_1 + t_3)X + (t_2 + t_3)Y$  and  $X, Y$  are i.i.d. Normal distributions. And what we have known is that the sum of independent Normal distribution is also a Normal distribution.

So  $(t_1 + t_3)X + (t_2 + t_3)Y$  is normal for any constant number  $t_1, t_2, t_3$ .

- (b) No, that is because  $X + Y + (SX + SY) = (1 + S)X + (1 + S)Y$ , at the mean time  $S$  is a random sign, with 0.5 probability making  $(1 + S)X + (1 + S)Y$  be equal to 0 when  $S = -1$ . So from the definition of MVN, we cannot get a MVN.

- (c) Yes, that is because  $t_1SX + t_2SY = S(t_1X + t_2Y)$ , from the threom mentioned in (a), we can have  $(t_1X + t_2Y) \sim N(0, t_1^2 + t_2^2)$ .

And from what we have proved in class, we can know  $Z = SW$  and  $Z \sim N(0, 1)$  when  $W \sim N(0, 1)$  and  $S$  is sign random and is independent of  $W$ .

So under this situation, we can let  $W = \frac{(t_1X + t_2Y)}{\sqrt{t_1^2 + t_2^2}}$ , then  $W \sim N(0, 1)$ , and we have known  $S$  is independent of  $(X, Y)$ .

Therefore  $Z = SW = S \frac{(t_1X + t_2Y)}{\sqrt{t_1^2 + t_2^2}}$ , then  $\sqrt{t_1^2 + t_2^2}Z = S(t_1X + t_2Y)$  and  $Z \sim N(0, 1)$ . Finally  $S(t_1X + t_2Y) \sim N(0, t_1^2 + t_2^2)$ .

## Problem 2

(1) Method 1 MVN

$t_1(X+Y) + t_2(X-Y) = (t_1+t_2)X + (t_1-t_2)Y$ , because  $X, Y$  are i.i.d. Normal distribution and because the sum of the Normal distribution is also Normal. Therefore we can get it.

(2) Method 2 change of variables

First denote  $(X+Y, X-Y) = (Z, W)$ , then  $Z = X+Y$ ,  $W = X-Y$ .

Then we can get  $z = x+y$ ,  $w = x-y$ , and then  $x = \frac{z+w}{2}$ ,  $y = \frac{z-w}{2}$ .

Therefore we can get  $\frac{\partial(x,y)}{\partial(z,w)}$ ,  $\begin{pmatrix} 0.5 & 0.5 \\ 0.5 & -0.5 \end{pmatrix}$  Then  $\left| \frac{\partial(x,y)}{\partial(z,w)} \right| = \frac{1}{2}$ .

Then  $f_{Z,W}(z,w) = f_{X,Y}(x,y) * \left| \frac{\partial(x,y)}{\partial(z,w)} \right| = \frac{1}{2} f_{X,Y}(x,y)$ .

Because  $X, Y$  are i.i.d. Normal distribution, then we can get  $f_{Z,W}(z,w) = \frac{1}{2} f_X(x) f_Y(y) = \frac{1}{2} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} = \frac{1}{4\pi} e^{-\frac{z^2}{4}} e^{-\frac{w^2}{4}}$ .

Obviously we can simply it as  $h(z)g(w)$ , so they are independent.

## Problem 3

### Solution

(a) For this question we can use change of variables method too.

From the picture we can have  $R = \sqrt{X^2 + Y^2}$  and  $\Theta = \arctan(\frac{Y}{X})$ .

Then we have  $r = \sqrt{x^2 + y^2}$  and  $\theta = \arctan(\frac{y}{x})$ .

Then we can  $\frac{\partial(r,\theta)}{\partial(x,y)}$  more easily.  $\begin{pmatrix} \frac{x}{\sqrt{x^2+y^2}} & \frac{y}{\sqrt{x^2+y^2}} \\ \frac{-y}{x^2+y^2} & \frac{x}{x^2+y^2} \end{pmatrix}$  Therefore  $\left| \frac{\partial(r,\theta)}{\partial(x,y)} \right| = \frac{1}{\sqrt{x^2+y^2}}$ .

Then  $f_{R,\Theta}(r,\theta) = f_{X,Y}(x,y) * \left| \frac{1}{\frac{\partial(r,\theta)}{\partial(x,y)}} \right| = \sqrt{x^2+y^2} f_{X,Y}(x,y)$ .

Because  $X, Y$  are i.i.d. Normal distribution, then we can get  $f_{R,\Theta}(r,\theta) = \sqrt{x^2+y^2} f_X(x) f_Y(y) = r \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}}$ .

And we have  $x^2 = \frac{r^2}{1+\tan^2\theta}$ ,  $y^2 = \frac{r^2 \tan^2\theta}{1+\tan^2\theta}$ , then we have,

$$f_{R,\Theta}(r,\theta) = \frac{r}{2\pi} e^{-\frac{r^2}{2}}.$$

And obviously  $\theta$  is distributed as Uniform distribution, so they are independent.

## Problem 4

### Solution

- (a) For this question we can use change of variables method too.

From the question we can have  $T = X + Y$ ,  $W = \frac{X}{Y}$ .

Then we have  $t = x + y$ ,  $w = \frac{x}{y}$ .

Then  $x = \frac{wt}{w+1}$ ,  $y = \frac{t}{w+1}$ .

Then we can  $\frac{\partial(t,w)}{\partial(x,y)}$  more easily.  $\begin{pmatrix} 1 & 1 \\ \frac{1}{y} & -\frac{x}{y^2} \end{pmatrix}$  Therefore  $\left| \frac{\partial(t,w)}{\partial(x,y)} \right| = \frac{x+y}{y^2}$ .

Then  $f_{T,W}(t,w) = f_{X,Y}(x,y) * \left| \frac{1}{\frac{\partial(t,w)}{\partial(x,y)}} \right| = \frac{y^2}{x+y} f_{X,Y}(x,y)$ .

Because  $X, Y$  are i.i.d. exponential distribution, then we can get  $f_{T,W}(t,w) = \frac{y^2}{x+y} f_X(x) f_Y(y) = \frac{y^2}{x+y} \lambda e^{-\lambda x} \lambda e^{-\lambda y}$

Therefore  $f_{T,W}(t,w) = \frac{1}{(w+1)^2} \lambda^2 t e^{-\lambda t}$ .

Obviously it can be denoted as  $h(w)g(t)$ , then they are independent.

Because  $\int_0^\infty \frac{1}{(1+w)^2} dw = 1$ .

Finally we can have  $f_W(w) = \frac{1}{(1+w)^2}$ , and  $f_T(t) = \lambda^2 t e^{-\lambda t}$ .

- (b) First let  $T = X + Y$ , then we can have  $f_T(t) = t$  when  $0 < t < 1$ , and  $f_T(t) = 2 - t$  when  $1 < t < 2$ , otherwise  $=0$ .

Then  $W = X + Y + Z = T + Z$ , with  $0 < w - t < 1$  and by using convolution method, we have,

$$f_W(w) = \int_{-\infty}^{+\infty} f_T(t) f_Z(w-t) dt = \int_{w-1}^w f_T(t) * 1 dt.$$

When  $0 < w < 1$ ,

$$\int_{w-1}^w f_T(t) dt = \int_0^w t dt = \frac{w^2}{2}.$$

When  $1 < w < 2$ ,

$$\int_{w-1}^w f_T(t) dt = \int_{w-1}^1 t dt + \int_1^w 2 - t dt = -w^2 + 3w - \frac{3}{2}.$$

When  $2 < w < 3$ ,

$$\int_{w-1}^w f_T(t) dt = \int_{w-1}^2 2 - t dt = \frac{(w-3)^2}{2}.$$

Otherwise  $=0$ .

- (c) method 1 convolution:

$$F_M(t) = P(M \leq t) = P(X \leq t, Y \leq t).$$

Because they are i.i.d. exponential distribution.

$$\text{Then } F_M(t) = P(X \leq t) P(Y \leq t) = (1 - e^{-\lambda t})^2.$$

Then we can get  $f_M(t) = 2\lambda e^{-2\lambda t} * (e^{\lambda t} - 1)$  by derivation.

$$\begin{aligned} f_{X+\frac{1}{2}Y}(t) &= \int_{-\infty}^{+\infty} f_X(x) f_{\frac{1}{2}Y}(t-x) dx = \int_0^t \lambda e^{-\lambda x} 2\lambda e^{-2\lambda(t-x)} dx = 2\lambda e^{-2\lambda t} \int_0^t \lambda e^{\lambda x} dx \\ &= 2\lambda e^{-2\lambda t} * (e^{\lambda t} - 1) = 2\lambda e^{-\lambda t} - 2\lambda e^{-2\lambda t}. \end{aligned}$$

Then we get it.

method 2 exponential properties:

Let  $L = \min(X, Y)$ , then we can denote  $M$  as  $M = L + (M - L)$ , and obviously  $L \sim \text{Expo}(2\lambda)$  because we have known the PDF of  $M$  and  $M + L = X + Y$ .

That is  $2\lambda e^{-\lambda t} = 2\lambda e^{-\lambda t} - 2\lambda e^{-2\lambda t} + L$ . Then  $L = 2\lambda e^{-2\lambda t}$ .

At the same time  $L$  can be seen as the subject finishing the first thing. And  $M - L \sim \text{Expo}(\lambda)$ , can be seen as the time of the second thing being finished.

Because of the property of memorylessness, then we can add them up and it is equal to  $M$ . Meanwhile,  $X \sim \text{Expo}(\lambda)$  and  $\frac{1}{2}Y \sim \text{Expo}(2\lambda)$ .

Therefore, we get it.

## Problem 5

### Solution

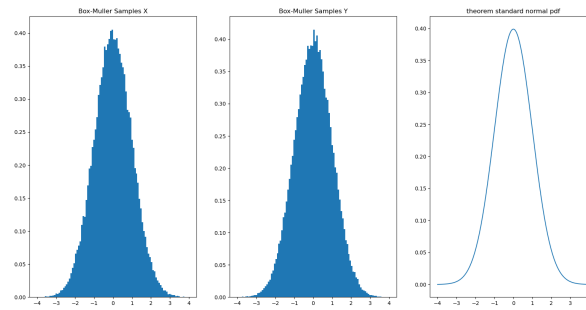


Figure 1: figure 1

(a)

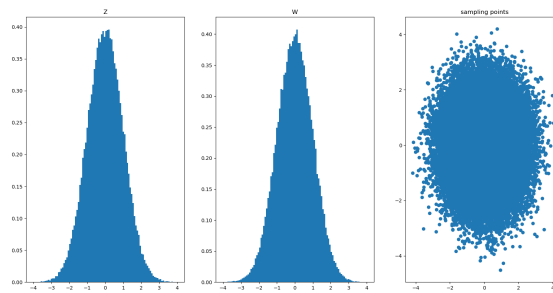


Figure 2: figure 2

(b)

(c)

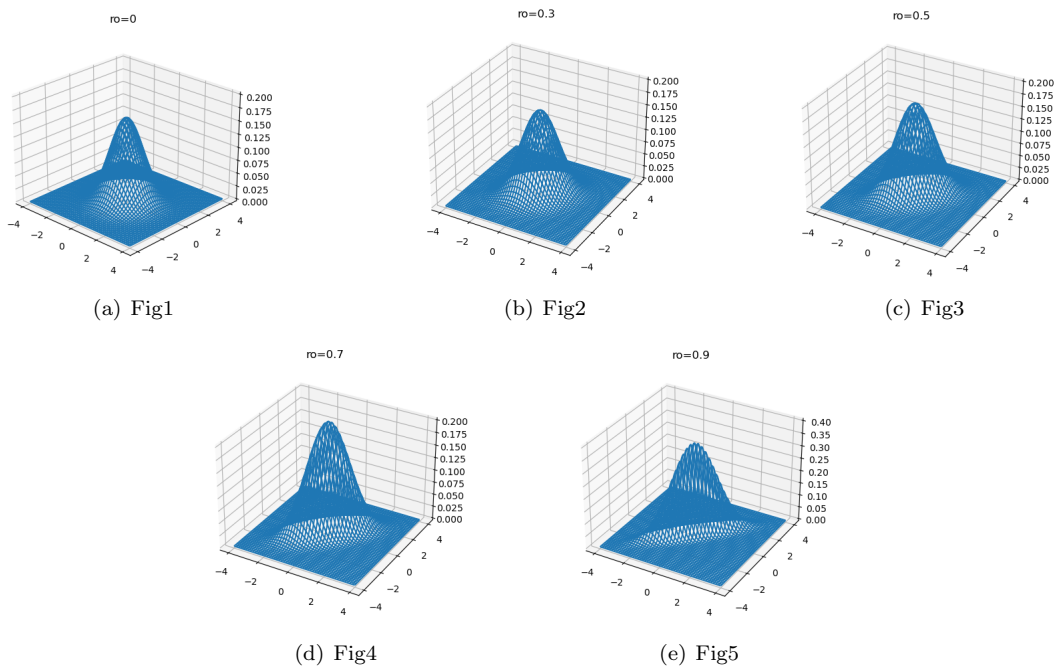


Figure 3: figure 3 3D

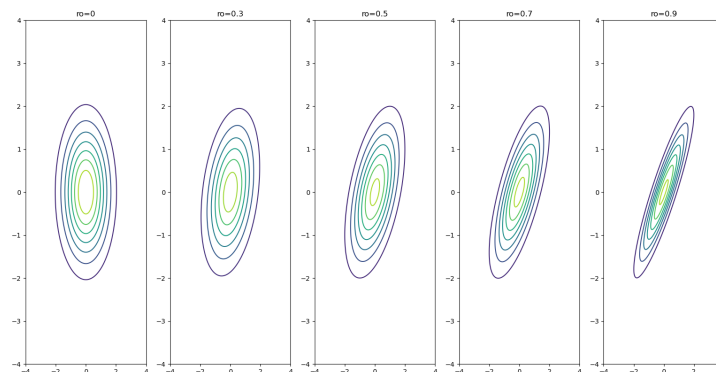


Figure 4: figure 4 contour

```
import numpy as np
import matplotlib.pyplot as plt
import math
from scipy.stats import multivariate_normal

##(1)
###unif
u1=np.random.random(100000)
u2=np.random.random(100000)
x1 = np.linspace(0 - 4 , 0 + 4, 1000)
R=np.sqrt(-2*np.log(u1))
theta=2*np.pi*u2
x=R*np.cos(theta)
y=R*np.sin(theta)
theorem=(1/np.sqrt(2*np.pi))*np.exp(-(x1**2)/2)
plt.figure(1)
plt.subplot(1,3,1)
plt.hist(x,bins=100,range=(-4,4),density=True)
plt.title("Box-Muller Samples X")
plt.subplot(1,3,2)
plt.hist(y,bins=100,range=(-4,4),density=True)
plt.title("Box-Muller Samples Y")
plt.subplot(1,3,3)
plt.title("theorem standard normal pdf")
plt.plot(x1,theorem)
plt.show()

##(2)
##because x y are independent so ro=0
ro=0
##(Z,W) is the MVN which we have learned in class
z=x
w=ro*x+np.sqrt(1-ro**2)*y
plt.figure(2)
plt.subplot(1,3,1)
plt.hist(z,bins=100,range=(-4,4),density=True)
plt.title("Z")
plt.subplot(1,3,2)

plt.hist(w,bins=100,range=(-4,4),density=True)
plt.title("W")
plt.subplot(1,3,3)
plt.scatter(z,w)
plt.title("sampling points")
plt.show()

##(3)
##set ro
ro_vector=[0,0.3,0.5,0.7,0.9]
##initial other parameters
cov=[[1,0],[0,1]]
u=[0,0]
```



```
x2 = np.linspace(0 - 4 , 0 + 4, 1000)
x3 = np.linspace(0 - 4 , 0 + 4, 1000)
Z,W=np.meshgrid(x2,x3)
space=np.empty(Z.shape+(2,))
space[:, :, 0]=Z
space[:, :, 1]=W
plt.figure(3)
for i in range(5):
    ##set parameters
    ro=ro_vector[i]
    cov[1][0]=ro
    cov[0][1]=ro
    #generate mvn
    generate_mvn=multivariate_normal(u,cov)
    U=generate_mvn.pdf(space)
    plt.subplot(1,5,i+1)
    plt.contour(Z,W,U)
    plt.title(f"ro={ro}")
plt.show()
fig=plt.figure()
ax=fig.add_subplot(projection='3d')
ax.set_zlim(0,0.2)
wframe=None
##set parameters
ro=ro_vector[0]
cov[1][0]=ro
cov[0][1]=ro
#generate mvn
generate_mvn=multivariate_normal(u,cov)
U=generate_mvn.pdf(space)
wframe=ax.plot_wireframe(Z,W,U)

plt.title(f"ro={ro}")
plt.show()
fig=plt.figure()
ax=fig.add_subplot(projection='3d')
ax.set_zlim(0,0.2)
wframe=None
##set parameters
ro=ro_vector[1]
cov[1][0]=ro
cov[0][1]=ro
#generate mvn
generate_mvn=multivariate_normal(u,cov)
U=generate_mvn.pdf(space)
wframe=ax.plot_wireframe(Z,W,U)

plt.title(f"ro={ro}")
plt.show()
```

```
fig=plt.figure()
ax=fig.add_subplot(projection='3d')
ax.set_zlim(0,0.2)
wframe=None
##set parameters
ro=ro_vector[2]
cov[1][0]=ro
cov[0][1]=ro
#generate mvn
generate_mvn=multivariate_normal(u,cov)
U=generate_mvn.pdf(space)
wframe=ax.plot_wireframe(Z,W,U)

plt.title(f"ro={ro}")
plt.show()
fig=plt.figure()
ax=fig.add_subplot(projection='3d')
ax.set_zlim(0,0.2)
wframe=None
##set parameters
ro=ro_vector[3]
cov[1][0]=ro
cov[0][1]=ro
#generate mvn
generate_mvn=multivariate_normal(u,cov)
U=generate_mvn.pdf(space)
wframe=ax.plot_wireframe(Z,W,U)

plt.title(f"ro={ro}")
plt.show()
fig=plt.figure()
ax=fig.add_subplot(projection='3d')
ax.set_zlim(0,0.4)
wframe=None
##set parameters
ro=ro_vector[4]
cov[1][0]=ro
cov[0][1]=ro
#generate mvn
generate_mvn=multivariate_normal(u,cov)
U=generate_mvn.pdf(space)
wframe=ax.plot_wireframe(Z,W,U)
plt.title(f"ro={ro}")
plt.show()
```