

The discrete Kalman filter

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So far we have developed the framework that allows us to express a random process and the covariance of its state (with usual notation ${m P}$) in the form

$$x_n = \Phi_{n-1}x_{n-1} + u_n$$
 (7.1)
$$P_n = \Phi_{n-1}P_{n-1}\Phi_{n-1}^T + Q$$

When we predict from t_{n-1} to t_n we do not have access to the actual contribution of u_n , but we are sure that (unless it is zero) it increases our state covariance. Assuming now that our process is not only driven by the model expressed in (7.1) but we take observations z_n that relate to the state vector linearly

$$z_n = H_n x_n + v_n \tag{7.2}$$

where the matrix H_n provides the (noiseless) connection between the state and the observations and v_n represents the measurement errors, which are thought to be of white Gaussian noise nature and uncorrelated with u_n .

We can now try to find a scheme that allows us to predict from t_{n-1} to t_n and then correct/update the state vector based on measurements collected at that epoch.

If we think of a method in which we first predict from t_{n-1} to t_n and then update the state and its covariance based on the measurements taken at t_n it is obvious that we we need to distinguish between a predicted state and an updated state. In order to make this distinction very clearly we use the subscript n|n-1 to indicate that the state (or its covariance) at t_n was predicted from a previous epoch t_{n-1} .

In addition we write \hat{x} when we deal with an estimated/predicted state instead of x for the true state. Thus, (7.1) becomes

$$\hat{x}_{n|n-1} = \Phi_{n-1|n-1} \cdot \hat{x}_{n-1|n-1}
P_{n|n-1} = \Phi_{n-1|n-1} \cdot P_{n-1|n-1} \cdot \Phi_{n-1|n-1}^T + Q$$
(7.3)

Now we can ask ourselves how we can correct the predicted state $x_{n|n-1}$ based on the measurements taken at t_n . Assuming that the difference $z_n - H_n \hat{x}_{n|n-1}$ gives us an indication on how we should correct $\hat{x}_{n|n-1}$ we can write

$$\hat{x}_{n|n} = \hat{x}_{n|n-1} + K_n \left(z_n - H_n \hat{x}_{n|n-1} \right)$$
 (7.4)

where K_n is a weighting factor which we call Kalman gain and need to determine in the following

The error and covariance of the predicted state is

$$e_{n|n-1} = \boldsymbol{x}_n - \hat{\boldsymbol{x}}_{n|n-1}$$

$$P_{n|n-1} = E\left(e_{n|n-1}e_{n|n-1}^T\right)$$

$$= E\left(\left(\boldsymbol{x}_n - \hat{\boldsymbol{x}}_{n|n-1}\right)\left(\boldsymbol{x}_n - \hat{\boldsymbol{x}}_{n|n-1}\right)^T\right)$$
(7.5)

In a similar way we can define the covariance of the updated state as

$$P_{n|n} = E\left(e_{n|n} \cdot e_{n|n}^T\right) = E\left(\left(x_n - \hat{x}_{n|n}\right)\left(x_n - \hat{x}_{n|n}\right)^T\right)$$
 (7.6)

Inserting (7.1) in (7.4) allows us to re-write (7.6) as

$$P_{n|n} = E\left\{ \left[\left(\boldsymbol{x}_{n} - \hat{\boldsymbol{x}}_{n|n-1} \right) - \boldsymbol{K}_{n} \left(\boldsymbol{H}_{n} \boldsymbol{x}_{n} + \boldsymbol{v}_{n} - \boldsymbol{H}_{n} \hat{\boldsymbol{x}}_{n|n-1} \right) \right] \cdot \left[\left(\boldsymbol{x}_{n} - \hat{\boldsymbol{x}}_{n|n-1} \right) - \boldsymbol{K}_{n} \left(\boldsymbol{H}_{n} \boldsymbol{x}_{n} + \boldsymbol{v}_{n} - \boldsymbol{H}_{n} \hat{\boldsymbol{x}}_{n|n-1} \right) \right]^{T} \right\}$$
(7.7)

Since the prediction error $x_n - \hat{x}_{n|n-1}$ is uncorrelated with the measurement error v_n we can write (7.7) as

$$P_{n|n} = (I - K_n H_n) P_{n|n-1} (I - K_n H_n)^T + K_n R_n K_n^T$$
(7.8)

where R_n is the covariance of the measurement noise, i.e. $R_n = E(v_n v_n^T)$.

The basic idea of the Kalman filter is to minimize the mean-square error of the a posteriori state estimation, i.e. finding an optimal state $\hat{x}_{n|n}$ that minimizes

 $E\left(\left(x_n-\hat{x}_{n|n}\right)^2\right)$. This is equivalent to minimizing the trace of the a posteriori estimate covariance matrix $P_{n|n}$, i.e.

$$J_n = \operatorname{tr}\left(P_{n|n}\right) = min. \Rightarrow \frac{\partial J_n}{\partial K_n} = \mathbf{0}$$
 (7.9)

A (useful) relation for the following computations is

$$\frac{\partial \operatorname{tr}(\boldsymbol{A}\boldsymbol{C}\boldsymbol{A}^T)}{\partial \boldsymbol{A}} = 2\boldsymbol{A}\boldsymbol{C}$$

if C is symmetric.

Thus we get from the two terms on the RHS in (7.9)

$$\frac{\partial \operatorname{tr}\left(\left(\boldsymbol{I} - \boldsymbol{K}_{n} \boldsymbol{H}_{n}\right) \boldsymbol{P}_{n|n-1} \left(\boldsymbol{I} - \boldsymbol{K}_{n} \boldsymbol{H}_{n}\right)^{T}\right)}{\partial \boldsymbol{K}_{n}} = -2\left(\boldsymbol{I} - \boldsymbol{K}_{n} \boldsymbol{H}_{n}\right) \boldsymbol{P}_{n|n-1} \boldsymbol{H}_{n}^{T} \quad (7.10)$$

and

$$\frac{\partial \operatorname{tr}\left(\boldsymbol{K}_{n}\boldsymbol{R}_{n}\boldsymbol{K}_{n}^{T}\right)}{\partial \boldsymbol{K}_{n}} = 2\boldsymbol{K}_{n}\boldsymbol{R}_{n} \tag{7.11}$$

Furthermore

$$\frac{\partial J_n}{\partial K_n} = 0 = -2\left(I - K_n H_n\right) P_{n|n-1} H_n^T + 2K_n R_n \tag{7.12}$$

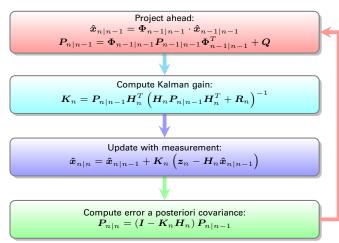
which allows to solve for K_n

$$K_n = P_{n|n-1}H_n^T \left(H_n P_{n|n-1}H_n^T + R_n\right)^{-1}$$
 (7.13)

and yields an analytic expression for the a posteriori covariance of the state

$$P_{n|n} = (I - K_n H_n) P_{n|n-1}$$
 (7.14)

In summary the Kalman filter can be sketched out as follows:



You will find the examples discussed in this lecture as Jupyter notebook under

 $\verb|https://github.com/spacegeodesy/ParameterEstimationDynamicSystems/blob/master/example07.ipynb|$