# MATH 284/FMPH 291 Survival Analysis Supplementary Learing Materials

Yuyao Wang

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#### **Tutorial**

Here is the link of a tutorial for survival analysis in R by Emily C. Zabor: https://www.emilyzabor.com/tutorials/survival\_analysis\_in\_r\_tutorial.html#Part\_1:\_Introduction\_to\_Survival\_Analysis

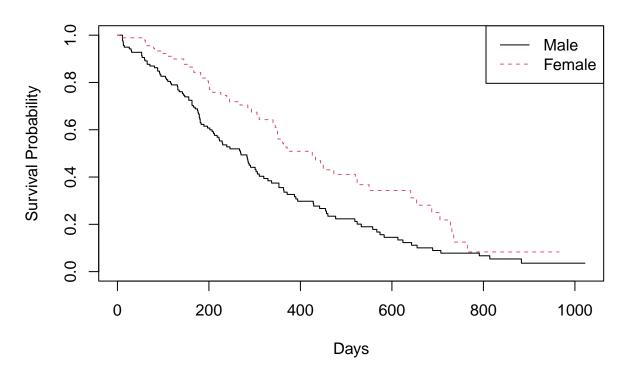
Part 1 in the above tutorial includes creating survival objects and curve, Kaplan-Meier plots, Estimating x-year survival, estimating median survival time, comparing survival times between groups (using log-rank test), and Cox regression model.

Below are some additional material for different topics.

### Kaplan-Meier Curve

The tutorial uses the 'ggsurvfit' R package for doing Kaplan-Meier plots. Alternatively, survival plots can be created using base R or the 'survminer' package. For example, below is an example of fiting KM curves for the two sex groups using the 'lung' data set from the 'survival' R package. You may find more details of the usage of the functions (for example how to add confidence intervals) by checking the R document for survfit() and plot.survfit().

## Kaplan-Meier Survival Curves by Sex



## Log-rank test

The function <code>survdiff()</code> used in the above tutorial for comparing survival times between groups is a function that conduct G-rho family tests. The default <code>rho=0</code>, which corresponds to log-rank test. More details can be found in the R documentation of this package. An alternative way of doing logrank test is using the score test for coxph score test. Try the code below, and you will find the two approaches gives exactly the same results!

BTW, survdiff() does not handle left truncated data, but coxph() can properly handle it by specifying the survival object to be Surv(Q,X,Delta), where Q is the left truncation time, X is the censored event time, and Delta is the event indicator.

```
# logrank test using G^{\t} to test with rho = 0
logrank.1 = survdiff(Surv(time, status) ~ sex, data = lung)
print(logrank.1)
## Call:
## survdiff(formula = Surv(time, status) ~ sex, data = lung)
##
##
           N Observed Expected (0-E)^2/E (0-E)^2/V
                  112
                          91.6
                                     4.55
## sex=1 138
                                               10.3
  sex=2 90
                   53
                          73.4
                                     5.68
                                               10.3
##
   Chisq= 10.3 on 1 degrees of freedom, p= 0.001
# logrank test using coxph score test
coxphfit = coxph(Surv(time, status) ~ sex, data = lung)
summary(coxphfit)
```

```
## Call:
## coxph(formula = Surv(time, status) ~ sex, data = lung)
##
##
    n= 228, number of events= 165
##
                                       z Pr(>|z|)
##
          coef exp(coef) se(coef)
                  0.5880
                           0.1672 -3.176 0.00149 **
## sex -0.5310
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
##
       exp(coef) exp(-coef) lower .95 upper .95
          0.588
## sex
                      1.701
                               0.4237
                                          0.816
##
## Concordance= 0.579 (se = 0.021)
## Likelihood ratio test= 10.63 on 1 df,
                                            p=0.001
## Wald test
                        = 10.09
                                on 1 df,
                                            p=0.001
## Score (logrank) test = 10.33 on 1 df,
                                            p=0.001
```

## The dual between estimating T distribution and estimating C distribution

The observed censored event time  $X = \min(T, C)$  not only contains information of the event time T, but also information of the censoring time C. The even indicator  $\Delta = I(T < C)$ . In other words, when  $\Delta = 1$ , X = T; when  $\Delta = 0$ , X = C.

When we focus on estimating T distribution, T is censored by C, we construct survival object using Surv(X, delta). When we focus on estimating C distribution, C can be viewed as censored by T. With this point of view think about how can we construct the survival object.

## Notes on the exponential model and weibull model fitted using survreg()

The location-scale parameterization of a Weibull distribution found in survreg is not the same as the parameterization of rweibull. See the Example for survreg() and the chunk for survreg.distributions in 'survival' R package documentation for more details on how to interpret the outputs.

When fitting an exponential model using fit = survreg(..., dist = "exponential"), 1/exp(fit\$icoef) correspond to the rate parameter in rexp(). Try the example below.

```
n=10000
lambda = 3
TT = rexp(n, rate = lambda)
C = runif(n, 1, 2)
X = pmin(TT,C)
delta = as.numeric(TT<C)

dat = data.frame(X = X, delta = delta)

fit = survreg(Surv(X, delta)~1, data = dat, dist = "exponential")
1/exp(fit$icoef)</pre>
```

```
## (Intercept)
## 2.954834
```

## Cox regression

We will use the lung data set in the survival R package as an example.

#### summary(lung)

```
##
         inst.
                          time
                                            status
                                                              age
##
    Min.
           : 1.00
                                 5.0
                                               :1.000
                                                                :39.00
    1st Qu.: 3.00
                     1st Qu.: 166.8
                                       1st Qu.:1.000
                                                         1st Qu.:56.00
    Median :11.00
                     Median: 255.5
                                       Median :2.000
                                                         Median :63.00
                             : 305.2
##
    Mean
            :11.09
                                               :1.724
                                                         Mean
                                                                :62.45
                     Mean
                                       Mean
    3rd Qu.:16.00
                     3rd Qu.: 396.5
                                       3rd Qu.:2.000
                                                         3rd Qu.:69.00
##
    Max.
            :33.00
                     Max.
                             :1022.0
                                       Max.
                                               :2.000
                                                         Max.
                                                                :82.00
##
    NA's
            :1
##
                                          ph.karno
                                                            pat.karno
         sex
                        ph.ecog
##
    Min.
            :1.000
                     Min.
                             :0.0000
                                       Min.
                                               : 50.00
                                                          Min.
                                                                 : 30.00
                                       1st Qu.: 75.00
                                                          1st Qu.: 70.00
##
    1st Qu.:1.000
                     1st Qu.:0.0000
                                       Median: 80.00
##
    Median :1.000
                     Median :1.0000
                                                          Median: 80.00
##
    Mean
            :1.395
                     Mean
                             :0.9515
                                       Mean
                                               : 81.94
                                                          Mean
                                                                 : 79.96
##
    3rd Qu.:2.000
                     3rd Qu.:1.0000
                                       3rd Qu.: 90.00
                                                          3rd Qu.: 90.00
##
    Max.
            :2.000
                     Max.
                             :3.0000
                                       Max.
                                               :100.00
                                                          Max.
                                                                 :100.00
                             :1
##
                     NA's
                                       NA's
                                               :1
                                                          NA's
                                                                  :3
##
       meal.cal
                         wt.loss
##
    Min.
           : 96.0
                      Min.
                              :-24.000
##
    1st Qu.: 635.0
                      1st Qu.:
                                0.000
##
    Median : 975.0
                      Median : 7.000
##
    Mean
            : 928.8
                      Mean
                              : 9.832
##
    3rd Qu.:1150.0
                      3rd Qu.: 15.750
    Max.
            :2600.0
                      Max.
                              : 68.000
##
    NA's
            :47
                      NA's
                              :14
```

#### Fitting the model

Try whether you get the same result if you change the data type for sex into factor. How to interpret the results when using sex as a continuous variable and when using sex as a factor?

```
# lung$sex <- as.factor(lung$sex)
coxfit = coxph(Surv(time, status) ~ sex, data = lung)
coxfit

## Call:
## coxph(formula = Surv(time, status) ~ sex, data = lung)
##
## coef exp(coef) se(coef) z p
## sex -0.5310  0.5880  0.1672 -3.176  0.00149
##
## Likelihood ratio test=10.63 on 1 df, p=0.001111
## n= 228, number of events= 165</pre>
```

```
ss = summary(coxfit)
SS
## Call:
## coxph(formula = Surv(time, status) ~ sex, data = lung)
##
##
    n= 228, number of events= 165
##
          coef exp(coef) se(coef)
                                       z Pr(>|z|)
##
                           0.1672 -3.176 0.00149 **
## sex -0.5310
                 0.5880
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
##
##
       exp(coef) exp(-coef) lower .95 upper .95
## sex
           0.588
                      1.701
                               0.4237
                                          0.816
##
## Concordance= 0.579 (se = 0.021)
## Likelihood ratio test= 10.63 on 1 df,
                                            p=0.001
## Wald test
                      = 10.09 on 1 df,
                                            p=0.001
## Score (logrank) test = 10.33 on 1 df,
                                            p=0.001
# Extract the coefficient
coef(coxfit)
##
          sex
## -0.5310235
# or
coxfit$coefficients
## -0.5310235
\# Extract the coefficients and p-value from Wald test
coef(ss)
             coef exp(coef) se(coef)
                                                   Pr(>|z|)
## sex -0.5310235 0.5880028 0.1671786 -3.176385 0.001491229
# or
ss$coefficients
             coef exp(coef) se(coef)
                                                   Pr(>|z|)
## sex -0.5310235 0.5880028 0.1671786 -3.176385 0.001491229
```

#### Prediction

Note that when using basehaz() to get the estimator of cumulative baseline hazard, one should use the argument center = FALSE in order to get the estimate for the cumulative baseline hazard function  $\Lambda_0(t)$  in the slides. See the example below.

When using center = FALSE in basehaz(), one need to further center the covariate values when using the formula:

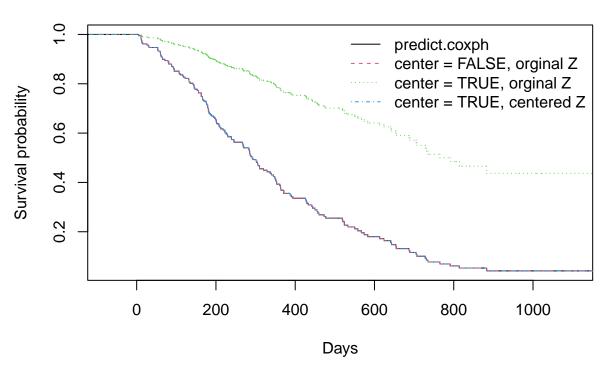
$$\hat{S}(t|Z) = e^{-\hat{\Lambda}_0(t)e^{\hat{\beta}^\top Z}}.$$

See the example below.

```
coxfit = coxph(Surv(time, status) ~ ph.karno, data = lung)
coxfit
## Call:
## coxph(formula = Surv(time, status) ~ ph.karno, data = lung)
##
##
                 coef exp(coef) se(coef)
                                              z
## ph.karno -0.016448  0.983686  0.005854 -2.81  0.00496
## Likelihood ratio test=7.56 on 1 df, p=0.005966
## n= 227, number of events= 164
      (1 observation deleted due to missingness)
ph.karno.new = 75
## Compute the baseline hazard
baseh = basehaz(coxfit, center = FALSE)
beta = coef(coxfit)
baseh_centerT = basehaz(coxfit, center = TRUE)
newdat = data.frame(time = baseh$time,
                    status = 1,
                    ph.karno = ph.karno.new)
pred = predict(coxfit, newdata = newdat, type = "survival", se.fit = TRUE)
pred_surv = pred$fit
# To see more usage of the function `prediction()` for `coxph` object, try the following.
# ?predict.coxph
# Another method - Compute the estimates from `beta` and `baseh`
# sex = 1 -> to the baseline survival, since sex = 1 is the reference group
pred2_surv_stepf <- stepfun(baseh$time, c(1, exp(-baseh$hazard*exp(beta*ph.karno.new))))</pre>
pred3_surv_stepf <- stepfun(baseh_centerT$time, c(1, exp(-baseh_centerT$hazard*exp(beta*ph.karno.new)))</pre>
pred4_surv_stepf <- stepfun(baseh_centerT$time, c(1, exp(-baseh_centerT$hazard*exp( beta*(ph.karno.new
## Plot out the step function for the estimated survival curve
pred_surv_stepf <- stepfun(newdat$time, c(1,pred_surv))</pre>
plot(pred_surv_stepf, do.points = FALSE, col = 1, lty = 1,
     main = "Predicted survival curves",
     xlab = "Days", ylab = "Survival probability")
# ?plot.stepfun # for more usage of plotting the `stepfun` object
plot(pred2_surv_stepf, do.points = FALSE, col = 2, lty = 2, add = TRUE)
plot(pred3_surv_stepf, do.points = FALSE, col = 3, lty = 3, add = TRUE)
plot(pred4 surv stepf, do.points = FALSE, col = 4, lty = 4, add = TRUE)
legend("topright",
```

```
legend = c("predict.coxph", "center = FALSE, orginal Z", "center = TRUE, orginal Z", "center = T
col = 1:4, lty = 1:4,
bty = "n")
```

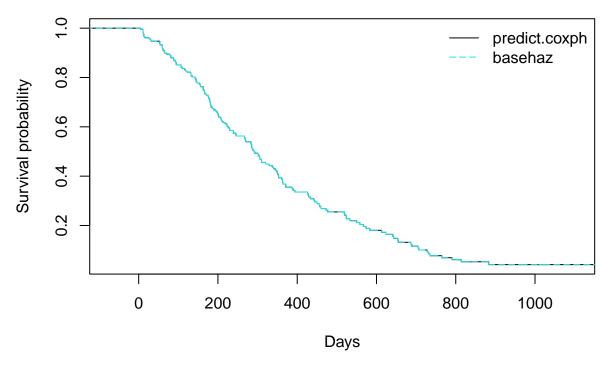
### **Predicted survival curves**



See the R documentation for  $\mathtt{predict.coxph}()$  for details about how to do prediction.

Another way to do prediction is using the hasehaz(.., newdata).

### **Predicted survival curves**



```
# To predict multiple survival curves at once
ph.karno.new = c(60, 75, 90)
newdat = data.frame(ph.karno = ph.karno.new)
cumhazard = basehaz(coxfit, newdata = newdat)
colnames(cumhazard)
```

## [1] "hazard.1" "hazard.2" "hazard.3" "time"

## **Predicted survival curves**

