

1. 解:

(a) 二项分布 $P(Y=1)=P, P(Y=0)=1-P \quad Y \sim B(1, P)$

设 $Y=h(X)$. ① $P(h(X)=1)=P, P(h(X)=0)=1-P$

不妨令 $h(x)=\begin{cases} 1, & (x \leq p) \\ 0, & (x > p) \end{cases}$ ② $P(h(x)=1)=P(X=h^{-1}(1))=P$
 $P(h(x)=0)=P(X=h^{-1}(0))=1-P$

$\therefore h(x)=\begin{cases} 1, & (0 \leq x \leq p) \\ 0, & (p < x \leq 1) \end{cases}$ 符合要求.

(b) $Y \sim E(\lambda), \lambda=1$. ③ 有 $P(Y \leq y)=F(y)=1-e^{-y}$

由 $Y=h(X) \Rightarrow P(h(X) \leq y)=1-e^{-y}, P(X \leq h^{-1}(y))=1-e^{-y}$

利用 $X \sim U(0, 1)$ 有 $P(X \leq h^{-1}(y))=h^{-1}(y)$ 从而有 $h^{-1}(y)=1-e^{-y}$

$\Rightarrow h(x)=-\ln(1-x) \quad (0 \leq x \leq 1)$ 可逆, 符合要求.

2. 解:

p -分位数的定义为: $\forall p \in (0, 1), F^{-1}(p)=\inf\{x \mid F(x) \geq p\}$

称 $F^{-1}(p)$ 为 X 的 p -分位数.

因此, 有

(1) $X \sim B(3, \frac{1}{3}), P(X=k)=\binom{3}{k} \left(\frac{1}{3}\right)^k \left(\frac{2}{3}\right)^{3-k} \Rightarrow$

$P(X=0)=\frac{8}{27}, P(X=1)=\frac{4}{9}, P(X=2)=\frac{2}{9}, P(X=3)=\frac{1}{27}$

$F(x)=\begin{cases} \frac{8}{27}, & (0 \leq x < 1) \\ \frac{20}{27} \approx 0.74, & (1 \leq x < 2) \\ \frac{26}{27} \approx 0.96, & (2 \leq x < 3) \\ 1, & (3 \leq x) \end{cases}$

$$\therefore F^{-1}(0.95)=\inf\{x \mid F(x) \geq 0.95\}$$

$$= 2$$

(2) $X \sim \mathcal{E}(1)$, 有 $F_X(x) = 1 - e^{-x}$

$$F^{-1}(0.05) = \inf\{x | F(x) \geq 0.05\} = \inf\{x | x \geq -\ln(0.95)\} = 0.051$$

3. 解:

(1) 证明: X 在 $\{0, 1, 2, \dots\}$ 上 随机取值.

$$\Rightarrow \sum_{i=0}^{\infty} P(X=i) = 1 ,$$

$$E(X) = \sum_{i=0}^{\infty} i P(X=i) = 0 \cdot P(X=0) + 1 \cdot P(X=1) + [P(X=2) + P(X=3)] \\ + [P(X=3) + P(X=4) + P(X=5)] + \dots$$

$$= [P(X=0) + P(X=1) + \dots] + [P(X=1) + P(X=2) + \dots] + [P(X=2) + \dots]$$

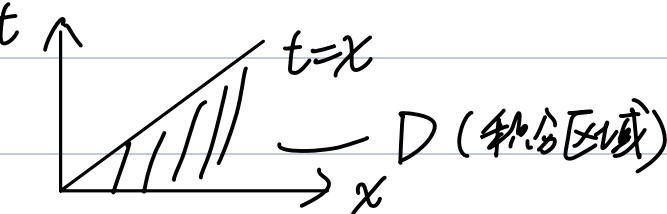
$$\text{从而有 } E(X) = \sum_{k=1}^{\infty} \sum_{i=k}^{\infty} P(X=i) = \sum_{k=1}^{\infty} P(X \geq k) = \sum_{k=0}^{\infty} P(X > k)$$

(2) 证明: 利用 $P(X > t) = \int_t^{+\infty} f(x) dx$

$$\text{则有 } \text{RHS} = \int_0^{\infty} P(X > t) dt = \int_0^{\infty} \left(\int_t^{+\infty} f(x) dx \right) dt$$

$$\text{应用 Fubini 定理, 有 } \int_0^{\infty} \int_t^{+\infty} f(x) dx dt = \int_0^{+\infty} \left(\int_0^x f(x) dt \right) dx$$

$$= \int_0^{+\infty} x f(x) dx = E(X)$$



□

4. 解: 由随机变量 $X \sim G(p)$. $P(X=k) = p(1-p)^{k-1}$

$$P(X \geq k) = P(\{\text{前 } k-1 \text{ 次均为不成功}\}) = (1-p)^{k-1}$$

$$E(X) = \sum_{k=1}^{\infty} k P(X=k) = \sum_{k=1}^{\infty} P(X \geq k) = \sum_{k=1}^{\infty} (1-p)^{k-1} = \frac{1}{p}$$

5. 解：设第一个阳性出现前阴性反应者的人数为随机变量 X

$$P(X=m) = P(\text{前 } m \text{ 个是阴性}, \text{ 第 } m+1 \text{ 个是阳性}) = \frac{A_{45}^m \times 5}{A_{50}^{m+1}} \quad (m \leq 45)$$

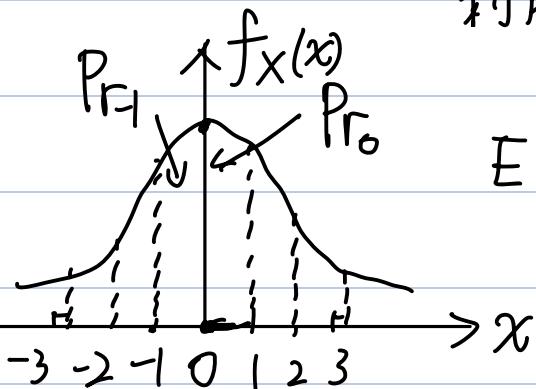
$$\text{则有 } E(X) = \sum_{m=0}^{45} m P(X=m) = \sum_{m=0}^{45} \frac{\frac{45!}{(45-m)!} \times 5}{\frac{50!}{(49-m)!}} \times m$$

$$= \sum_{m=0}^{45} \frac{m \times (49-m)(48-m)(47-m)(46-m) \times 5}{50 \times 49 \times 48 \times 47 \times 46} = \frac{15}{2}$$

6. 解： $X \sim N(0, \sigma^2) \therefore f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{x^2}{2\sigma^2}}$, $Y = \lfloor X \rfloor$

设 $P_r_i = \int_{r_i}^{r_{i+1}} f_X(x) dx$, $r_i \in \mathbb{Z}$. 则 $EY = \sum r_i P_{r_i}$ for all r_i

利用 $P_{r_i} = P_{r_{i-1}}$ 对称性可知



$$EY = \dots + (-2)P_{r_{-2}} + (-1)P_{r_{-1}} + 0P_{r_0} + 1P_{r_1} + \dots$$

$$= (-1)(P_{r_0} + P_{r_1} + P_{r_2} + \dots)$$

$$= (-1) \sum_{i=0}^{\infty} P_{r_i} = -\frac{1}{2} \quad (\text{利用 } \sum_{i=0}^{\infty} P_{r_i} = P(X \geq 0) = \frac{1}{2})$$

7. 解： $X, Y \sim N(0, 1)$ 有 $f(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$, $f(y) = \frac{1}{\sqrt{2\pi}} e^{-y^2/2}$

$X \perp Y$. 有 $T = X^2 + Y^2 \sim \chi^2_2$, 那有 $f(t) = \frac{1}{2P(1)} t^0 e^{-t/2}$, $t \geq 0$

即有 $f(t) = \frac{1}{2} e^{-t/2}$ ($t \geq 0$)

而由 $Z = T^{\alpha}$ 有 $T = g(Z) = Z^{\frac{1}{\alpha}}$, $\frac{dg(Z)}{dZ} = \frac{1}{\alpha} Z^{\frac{1}{\alpha}-1}$

$\therefore f_Z(z) = f_T(g(z)) \cdot \left| \frac{dg(z)}{dz} \right| = \frac{1}{2} e^{-\frac{1}{2} z^{\frac{1}{\alpha}}} \cdot \left| \frac{1}{\alpha} z^{\frac{1}{\alpha}-1} \right|$

$$\text{则有 } E(Z) = \int_0^{+\infty} z f_Z(z) dz = \int_0^{+\infty} \left| \frac{1}{2\alpha} \right| z^{\frac{1}{\alpha}} e^{-\frac{z^{\frac{1}{\alpha}}}{2}} dz$$

利用 Gamma 分布的性质

$$\int_0^{+\infty} \frac{\lambda^{\alpha+1}}{\Gamma(\alpha+1)} x^{\alpha} e^{-\lambda x} dx = 1 \quad . \quad \text{取 } \lambda = \frac{1}{2}, x = z^{\frac{1}{\alpha}} \text{ 有}$$

$$\int_0^{+\infty} \frac{\left(\frac{1}{2}\right)^{\alpha+1}}{\Gamma(\alpha+1)} z \cdot e^{-\frac{z^{\frac{1}{\alpha}}}{2}} \cdot \frac{1}{\alpha} z^{\frac{1}{\alpha}-1} dz = 1$$

$$\Rightarrow \int_0^{+\infty} z^{\frac{1}{\alpha}} e^{-\frac{z^{\frac{1}{\alpha}}}{2}} dz = 2^{\alpha+1} \Gamma(\alpha+1) \text{ 代入后得}$$

$$E(Z) = \frac{2^{\alpha}}{\alpha} \Gamma(\alpha+1) = \Gamma(\alpha) \cdot 2^{\alpha}$$

8. 解:

设木棍的长度为随机变量 Y . 有 $Y \sim U(1,2)$

设截出较短的一段长度为 X . 则 $X|Y \sim U(\frac{Y}{2}, Y)$

$$f_Y(y) = I_{\{1 \leq y \leq 2\}} \quad f_{X|Y}(x) = I_{\{\frac{y}{2} \leq x \leq y\}} \cdot \frac{1}{y-y/2} = I_{\{\frac{y}{2} \leq x \leq y\}} \cdot \frac{2}{y}$$

则联合分布为 $f_{X,Y}(x=y) = f_{X|Y}(x) f_Y(y)$

$$f_{X,Y}(x,y) = \frac{2}{y} \quad \left(\frac{y}{2} \leq x \leq y, 1 \leq y \leq 2 \right), \text{ otherwise } = 0$$

$$\text{则 } E(\Delta) = E(2X-Y) = \iint (2x-y) \cdot f_{X,Y}(x,y) dx dy$$

$$= \int_1^2 \left(\int_{\frac{y}{2}}^y \frac{2}{y} (2x-y) dx \right) dy = \frac{3}{4}$$