

# EE5137 Semester 1 2018/9: Quiz 1 (Total 24 points)

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Matriculation Number: XX

Score: 24/24

You have 1.0 hour for this quiz. There are FOUR (4) printed pages. You're allowed 1 sheet of handwritten notes. Please provide *careful explanations* for all your solutions.

1. (8 points) [Distribution Functions] Which of the following functions is a cumulative distribution function (CDF)? For those which are, compute the probability density function (PDF). For those which are not, explain what fails.

(a)

$$F_X(x) = \begin{cases} 1 - e^{-x^2} & x \geq 0 \\ 0 & \text{else} \end{cases}$$

**Solution:** Yes. This is a distribution function

- $\lim_{x \rightarrow -\infty} F_X(x) = 0$ ;
- $\lim_{x \rightarrow +\infty} F_X(x) = \lim_{x \rightarrow +\infty} 1 - e^{-x^2} = 1$ ;
- $F_X(x)$  is continuous at every point including 0 because on the left it takes the value 0 and on the right  $1 - e^{-0^2} = 0$ .

The PDF is

$$f_X(x) = \frac{d}{dx} F_X(x) = 2xe^{-x^2}$$

for  $x \geq 0$  and 0 otherwise. This is known as the *Rayleigh distribution* with scale  $\sigma = 1/\sqrt{2}$ . See [https://en.wikipedia.org/wiki/Rayleigh\\_distribution](https://en.wikipedia.org/wiki/Rayleigh_distribution).

(b)

$$F_Y(y) = \begin{cases} 0 & y \leq 0 \\ \frac{1}{3} & 0 < y \leq \frac{1}{2} \\ 1 & y > \frac{1}{2} \end{cases}$$

**Solution:** This “CDF” is not a distribution function. It is not right-continuous. Note that  $F_Y(0) = 0$  while  $\lim_{y \downarrow 0} F_Y(y) = 1/3$ . Also  $F_Y(1/2) = 1/3$  and  $\lim_{y \downarrow 1/2} F_Y(y) = 1$ .

2. (8 points) [Strengthened Union Bound]

Let  $A_1, \dots, A_n$  be arbitrary events. Prove that

$$\Pr \left\{ \bigcup_{i=1}^n A_i \right\} \leq \min_{1 \leq k \leq n} \left( \sum_{i=1}^n \Pr\{A_i\} - \sum_{i=1: i \neq k}^n \Pr\{A_i \cap A_k\} \right).$$

*Hint: For any two sets  $C$  and  $D$ ,*

$$C = (C \cap D) \cup (C \cap D^c)$$

**Solution:** Using the hint, we have

$$\bigcup_{i=1}^n A_i = \left[ \left( \bigcup_{i=1}^n A_i \right) \cap A_k \right] \cup \left[ \left( \bigcup_{i=1}^n A_i \right) \cap A_k^c \right].$$

for any  $1 \leq k \leq n$ . But this is equivalent to

$$\bigcup_{i=1}^n A_i = A_k \cup \left[ \bigcup_{i=1}^n (A_i \cap A_k^c) \right].$$

Taking probabilities,

$$\begin{aligned} \Pr \left\{ \bigcup_{i=1}^n A_i \right\} &= \Pr \left\{ A_k \cup \left[ \bigcup_{i=1}^n (A_i \cap A_k^c) \right] \right\} \\ &\leq \Pr\{A_k\} + \sum_{i=1}^n \Pr\{A_i \cap A_k^c\} \\ &= \Pr\{A_k\} + \sum_{i=1, i \neq k}^n \Pr\{A_i \cap A_k^c\}, \quad (\text{because } A_k \cap A_k^c = \emptyset) \\ &= \Pr\{A_k\} + \sum_{i=1, i \neq k}^n [\Pr\{A_i\} - \Pr\{A_i \cap A_k\}] \\ &= \sum_{i=1}^n \Pr\{A_i\} - \sum_{i=1, i \neq k}^n \Pr\{A_i \cap A_k\} \end{aligned}$$

Since the bound holds for all  $1 \leq k \leq n$ , we can minimize the right-hand-side to yield

$$\Pr \left\{ \bigcup_{i=1}^n A_i \right\} \leq \min_{1 \leq k \leq n} \left( \sum_{i=1}^n \Pr\{A_i\} - \sum_{i=1: i \neq k}^n \Pr\{A_i \cap A_k\} \right).$$

as desired.

3. [Conditional Expectations] (8 points)

In this problem, we will calculate the expectation of a geometric random variable using the formula for iterated expectations. Let  $N$  be a geometric random variable with parameter  $p$ , i.e.,  $N$  is the number of coin flips until Head appears and  $\Pr(\text{Heads}) = p$ . In other words  $p_N(n) = (1 - p)^{n-1}p$  for  $n = 1, 2, \dots$ . Define the random variable

$$Y = \begin{cases} 1 & \text{first flip is Heads} \\ 0 & \text{else} \end{cases}$$

- (i) Calculate  $\mathbb{E}[N|Y = y]$  for  $y = 1$ .

**Solution:** If  $Y = 1$ , then we know that the first flip is a head. Thus the number of coin flips to a head is exactly 1, i.e.,

$$\mathbb{E}[N|Y = 1] = 1.$$

- (ii) Calculate  $\mathbb{E}[N|Y = y]$  for  $y = 0$  in terms of  $\mathbb{E}[N]$ .

**Solution:** If  $Y = 0$ , then we know that the first flip is a tail. Thus the number of coin flips to a head is  $\mathbb{E}[N] + 1$  (as the process is memoryless), i.e.,

$$\mathbb{E}[N|Y = 0] = \mathbb{E}[N] + 1.$$

- (iii) Now use the law of iterated expectations to deduce  $\mathbb{E}[N]$ .

**Solution:** We know that

$$\mathbb{E}[N] = \Pr(Y = 0)\mathbb{E}[N|Y = 0] + \Pr(Y = 1)\mathbb{E}[N|Y = 1].$$

Substituting the above values, we have

$$\mathbb{E}[N] = (1 - p) \times (\mathbb{E}[N] + 1) + p \times 1.$$

Solving this for  $\mathbb{E}[N]$ , we have

$$\mathbb{E}[N] = \frac{1}{p}$$

which was to be expected.